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**Emotion regulation strategy use and forecasting in response to dynamic, multimodal stimuli**

**William J. Mitchell a,** billy.mitchell@temple.edu **\***

**Joanne Stasiak b,** [joanne.stasiak@psych.ucsb.edu](mailto:joanne.stasiak@psych.ucsb.edu)

**Steven Martinez a,** [stevent.martinez@temple.edu](mailto:stevent.martinez@temple.edu)

**Katelyn Cliver c,** [katelyn.cliver@temple.edu](mailto:katelyn.cliver@temple.edu)

**David Gregory a,** [david.gregory@temple.edu](mailto:.gregory@temple.edu)

**Samantha Reisman d,** [Reisman@brown.edu](mailto:Reisman@brown.edu)

**Helen Schmidt a,** [helen\_schmidt@temple.edu](mailto:helen_schmidt@temple.edu)

**Vishnu P. Murty a,** [vishnu.murty@temple.edu](mailto:.murty@temple.edu)

**Chelsea Helion a,** [chelsea.helion@temple.edu](mailto:chelsea.helion@temple.edu)

1. **Department of Psychology & Neuroscience**

Weiss Hall, Temple University, 1701 N 13th St. Philadelphia, PA, USA 19122

1. **Department of Psychological & Brain Sciences**

Building 251, University of California, Santa Barbara, Santa Barbara, CA, 93106

1. **Department of Psychological & Brain Sciences**

Disque Hall, 3201 Chestnut St, Philadelphia, PA 19104

1. **Department of Cognitive, Linguistic, and Psychological Sciences**

Box 1821, Brown University, Providence, RI, 02912

**\*** Corresponding author.

*E-mail address:* [billy.mitchell@temple.edu](mailto:billy.mitchell@temple.edu)

*Address:* 717 Weiss Hall, Temple University,

1701 N 13th St. Philadelphia, PA 19122

**ABSTRACT (248 / 250 Words):** Successful emotion regulation (ER) requires effective strategy selection. Prior research suggests that disengagement strategies (e.g., distraction) are more likely to be selected than engagement strategies (e.g., reappraisal) as emotional experiences intensify. However, the extent to which ER strategy *choice* relationships in ideal or controlled circumstances reflect strategy *usage* during complex, multimodal events is not well understood. The present research uses dynamic, multimodal stimuli (i.e., a haunted house, horror movies) to examine the association between affective intensity and regulatory strategy usage among untrained participants – individuals given no prior regulation instructions or direction. Study 1 (*n =* 118) failed to find a relationship between emotional intensity and strategy usage to downregulate emotions as participants navigated a haunted house. Contrary to expectations, the success of distraction decreased as emotional intensity increased in this context. Participants in Study 2 (*n* = 152) forecasted regulation strategy usage based upon descriptions of emotionally-regulated experiences from the haunted house. Affective intensity predicted which strategies forecasters predicted they would use; though, forecasters overpredicted how often distraction was used in practice. Study 3 (*n* = 242) incorporated experiencing and forecasting within the same design by showing untrained participants video stimuli of varying-intensity and capturing their regulatory responses. Forecasters again predicted using distraction more often than experiencers did in practice. Forecasters also overpredicted how effectively distraction would reduce negative affective intensity relative to what experiencers reported. This pattern of results may highlight a disconnect between self-regulation strategy fittedness while planning and executing, especially in dynamic, complex settings.

**SIGNIFICANCE:** How individuals manage their emotions during dynamic, complex experiences is an important question for anxiety- and trauma-based disorder prevention and treatment. This study found distraction, but not reappraisal, to be less effective than anticipated as emotional intensity increases in these difficult-to-regulate environments, which contrasts findings from relatively less demanding contexts.

**CONSTRAINTS ON GENERALITY:** We aimed to generalize our findings to the downregulation of negative emotions by non-clinical US populations in dynamic, multimodal emotional situations. However, the setting of Study 1 (i.e., a haunted house) likely will not generalize to all dynamic, multimodal emotional situations. Studies 2 and 3 contain reasonably representative samples in terms of education, socioeconomic status, and race, but were collected online, thus limiting generalizability to non-optimal settings. Our recruitment added additional constraints as well. Using horror-related stimuli likely resulted in a self-selection bias. Participants in Study 1 skewed young, falling between the ages of 18 and 34, and were recruited from an urban setting, which may limit the applicability of our conclusions to children, older age groups, or individuals living in non-urban settings. Participants skewed more female than male in Studies 1 and 2 and more male than female in Study 3, which may result in an underrepresentation of men and women, respectively. Lastly, racial demographic information was not captured in Study 1, which leaves questions of racial representation unanswered. Future research should explore these phenomena in other novel settings with a wider age range and balanced gender ratio to enhance the external validity of our findings.

**KEYWORDS:** emotion, self-regulation, naturalistic stimuli, decision making

**INTRODUCTION**

Overwhelming situations can generate profoundly intense emotions which often do not match our ideal emotional states. We can try to change our emotional responses in loud or crowded spaces, unexpected confrontations, and circumstances that we lack control over, but our ability to successfully regulate our emotions in these situations may be limited. A key factor in determining regulation efficacy is the strategy someone uses or chooses. Regulation strategy efficacy can be highly context-dependent, leading many to emphasize studying emotion regulation (ER) in different contexts as the next crucial direction for the field (Aldao, 2013; Dixon-Gordon et al., 2015; English et al., 2017; Rottweiler et al., 2018; Tang & Huang, 2019).

**The Process Model of Emotion Regulation.** Foundational ER research (i.e., the PROCESS Model) has identified categories of common strategies that people use to regulate their emotions (Gross, 1998). Characteristics of the regulator and the context can determine the effectiveness of these strategies (Young & Suri, 2020) and people often choose strategies to match their present circumstances (Opitz et al., 2015; Sheppes et al., 2011). For example, engagement strategies, like altering the meaning or interpretation of the emotion-eliciting stimulus (i.e., reappraisal), often have high cognitive demands. These can be contrasted against disengagement strategies, like diverting attention (i.e., distraction) or inhibiting the expression of emotion (i.e., suppression), which typically have relatively low cognitive demands (Sheppes & Gross, 2011).

How might these regulatory patterns be reflected in strategy choice? Although not explicitly directed to do so, the anticipation of a scary moment during a horror movie might prompt a person to look away from the screen (distraction), think about the actors in a different light (reappraisal), or limit the expression of their fear, all in an effort to reduce, or downregulate, an unwanted feeling. People in situations with few cognitive resources may compensate by selecting strategies that demand less cognitive effort according to the selection, optimization, and compensation (ER-SOC) hypothesis (Opitz et al., 2012). For example, if someone were watching this horror movie while they were tired or under high cognitive load, it may be easier to look away from the screen than to generate alternative, less fear-inducing reinterpretations by which to view the film’s events. Attempting to reappraise would constitute a high-risk strategy in this context because it might be less likely to work (Ford & Troy, 2019) than a comparatively low-effort strategy like distraction.

This supposition is bolstered by the especially robust influence of emotional intensity upon strategy choice, as distraction is chosen more often and is more effective than reappraisal in response to high intensity stimuli (Hay et al., 2015; Orejuela-Dávila et al., 2019; Shafir et al., 2016; Sheppes & Gross, 2011; Young & Suri, 2020), at least in part because it is less effortful when emotional intensity is high (Silvers et al., 2014). This effect has been thoroughly replicated in lab studies and ecological momentary assessment (EMA) studies (Colombo et al., 2020; Haines et al., 2016; Heiy & Cheavens, 2014). However, this association may not readily translate to more complex and demanding environments (Sheppes, 2020), like those which we test directly here, due to differences in how we commonly measure and manipulate ER in these studies.

**External Validity in Extant Emotion Regulation Paradigms.** Lab ER paradigms differ from the everyday experience of ER in a few key ways that may limit generalizability. Lab ER paradigms (e.g., Sheppes et al., 2011, 2014) usually train participants to use regulatory strategies before a task begins, which may prime more introspection and metacognition (Carver & Scheier, 1981) than what occurs in the typical ER experience. Lab ER paradigms also often necessarily show previews of emotional stimuli to allow participants to prepare their regulatory responses, but dynamic, high-intensity events in our everyday lives are often unexpected or difficult to anticipate. Following stimulus previews, lab ER paradigms also frequently prompt individuals to select a strategy, but aversive experiences in everyday life, like sharing the road with an erratic driver or being bullied, may not explicitly prompt the implementation of self-regulation.

Such study design decisions prioritize internal validity over external validity. Self-report capture of emotion in ER studies is often either assumed based upon standardized ratings associated with the stimuli (e.g., IAPS picture set) (Bradley & Lang, 2007), measured through unidimensional Likert scales (e.g., Valence) (Shafir et al., 2016), or is captured through established measures (i.e., Positive and Negative Affect Schedule) (Watson et al., 1988; Weiss et al., 2021). These approaches offer an efficient, reliable, and standardized means of assessing self-regulation but might not accurately reflect the multidimensionality of emotional experience. Contemporary constructivist theories posit that emotions are not fixed reactions but are instead shaped dynamically by cognitive and social processes, encompassing one’s personal interpretations, beliefs, and social interactions (Lindquist et al., 2012). Study designs that prioritize external validity promote participant- or stimulus-level idiosyncratic experiences that are pivotal to developing working generative models of how the world around us works (Lee et al., 2021). Applying natural language processing techniques to free-response data can improve our ability to capture the multifaceted, idiosyncratic emotional experiences that constructivist frameworks emphasize (Mohammad, 2018) without sacrificing accuracy (Diamond et al., 2020) or quantifiability. Free-response capture, more than discretely categorized self-reports or unipolar scales, may require fewer a priori assumptions from researchers about a participant’s emotional experience, thus reducing unintended researcher influence (Gendron et al., 2012; Lindquist et al., 2006) and improving generalizability (Miller et al., 2019).

EMA studies – another common means of studying ER – do more directly assess the external validity of ER strategy choice relationships by capturing emotionally evocative events within the everyday lives of trained research participants (e.g., Colombo et al., 2020; Haines et al., 2016; Heiy & Cheavens, 2014). However, most people are not trained to consider their ER strategies in their daily lives and are not prompted or primed to engage regulatory control before an emotional event occurs (Friedman & Gustavson, 2022). As such, training participants may introduce important but often underappreciated deviations in regulatory behaviors from how untrained counterparts might respond in the same situation.

Research designs that incorporate more features of naturalistic regulation, such as not instructing or prompting participants to regulate (e.g., Heiy & Cheavens, 2014; Opitz et al., 2015) often find people explore and flexibly apply multiple strategies that blur the boundaries of typical strategy classifications (Aldao & Nolen-Hoeksema, 2013; Ford et al., 2019; Heiy & Cheavens, 2014; Opitz et al., 2015; Szasz et al., 2018). These approaches also capture meaningful variance in self-regulatory behaviors that more controlled designs cannot (Friedman & Gustavson, 2022; Kamradt et al., 2014; Malanchini et al., 2019). For example, overstimulation from complex, multimodal contexts may simultaneously be aversive and more cognitively demanding (i.e., better suited for disengagement strategies). However, attention may also be challenging to control in a context with so much attention-demanding stimuli, reducing the likelihood of observing the high-intensity-distraction association characterized in laboratory studies (Draheim et al., 2022). Spontaneous or untrained ER in these contexts may rely more heavily on person-specific features like habits than how emotionally arousing the situation may be (Christou-Champi et al., 2015; Koole et al., 2015; Norem, 2008). Yet, challenging situations may be precisely when adaptive regulatory control (i.e., the pattern between intensity and regulation choice observed in most laboratory studies) may be most valuable, as maladaptive ER tendencies predict more severe manifestations of post-traumatic stress disorders (Hannan & Orcutt, 2020) and related post-traumatic stress disorder outcomes (Specker & Nickerson, 2022). Thus, identifying whether the established association between intensity and effort-related strategy usage occurs in high stress contexts is important for the development of potential interventions.

**The Present Study.** To our knowledge, no study has examined whether affective intensity predicts strategy usage in aversive, high-intensity, multimodal contexts using untrained participants (i.e., individuals given no prior training in how to use or classify ER strategies and no direction to regulate their experiences). While affective intensity represents a particularly prominent predictor of ER behavior (i.e., *r+* = 0.46 – 0.61; “a very large effect size”, according to recent meta-analyses (Matthews et al., 2021)), the extent to which features of a dynamic, multimodal situation could overshadow this effect remain unclear. The goal of the present research is to examine whether these well-established regulatory patterns emerge in samples of untrained participants exposed to a dynamic, feature-rich stimuli high in emotional variability, such as a haunted house. Haunted house experiences have been used with marked success in recent research to study emotion and self-regulation (Clasen, 2019; Stasiak et al., 2023; Tashjian et al., 2022). While haunted houses only represent a small proportion of the variability which emotionally-relevant experiences could materialize as, they nonetheless offer safe and controlled but high-arousal events to elicit a wide range of emotional experiences (i.e., positive and negative emotions), intensities, and responses (i.e., regulation behaviors). This variance would be difficult to generate in any other complex multimodal context outside of the lab while offering the same level of safety to the participants and control to the researchers.

Because the comparison between reappraisal and distraction is nearly ubiquitous in the extant regulation literature (Heiy & Cheavens, 2014), Experiment 1 aimed to replicate this effect using untrained participants who navigated a haunted house and reported their undirected emotional and regulatory behaviors in a surprise recall task immediately after and one week after exposure. This granted high levels of fidelity in capturing subjective participant emotion and experience. We additionally anticipated that participants experiencing high cognitive load as a result of exposure would more often use distraction, in line with the ER-SOC hypothesis. Surprisingly, we did not find that intensity or cognitive load predicted strategy usage. To minimize researcher bias, we then applied a multiverse approach to the data, systematically expanding and constraining our data inclusion criteria and adjusting for important covariates such as regulation tendencies and emotional goals, in an effort to identify conditions in which the relationship between affective intensity and strategy usage might emerge as statistically significant. We again failed to find an association.

This surprising finding motivated Experiment 2, which aimed to determine whether participants exposed to similar experiences as the haunted house but in a less dynamic, lower-intensity context (i.e., forecasting rather than experiencing) would more often forecast, or predict, using distraction in response to descriptions of high intensity regulated events and reappraisal in response to descriptions of low intensity regulated events from the haunted house. We did observe the canonical association between emotional intensity and regulatory strategy selection with this design. However, many study design differences between Studies 1 and 2 limited our ability to draw conclusions.

Study 3 tasked untrained participants with watching videos of varying negative intensity and subsequently either reporting the regulatory strategies that they used (experiencing) or the strategies that they predict might be used (forecasting) to downregulate their negative emotions. For both forecasters and experiencers, increasing self-reported affective intensity predicted a greater likelihood of using distraction. However, forecasters were less likely than regulators to use reappraisal at low intensities. A trend that emerged across all three studies was an inconsistency between how effective participants predicted distraction might be within these situations and how effective distraction was actually reported to be. Methods and analyses supporting these findings are discussed.

**EXPERIMENT 1 METHODS**

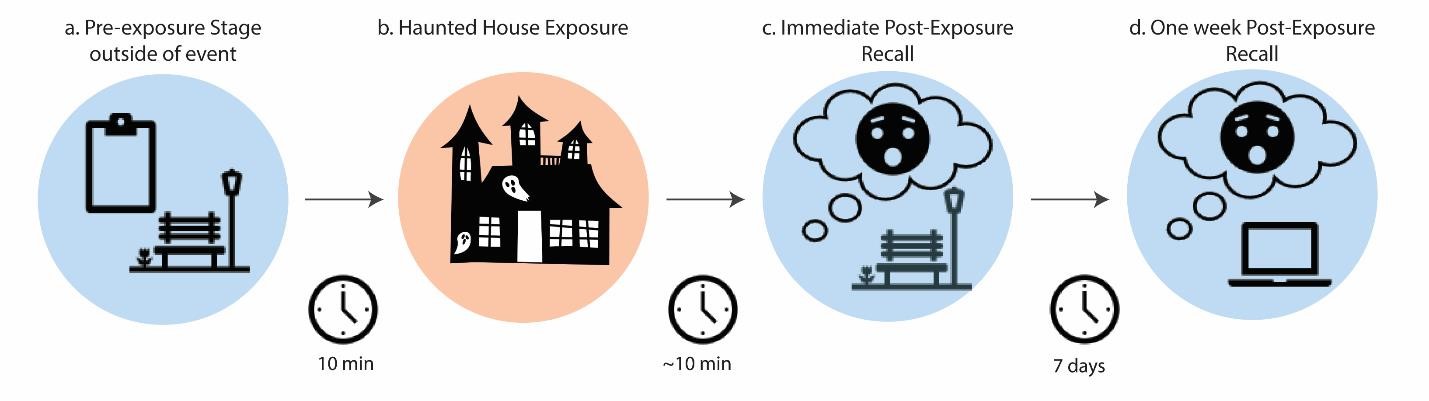
Experiment 1 tested whether the emotional intensity of negatively-valenced events was associated with the likelihood of using a low-effort or high-effort regulatory strategy in a dynamic, feature-rich setting with an untrained sample of participants. To assess emotional intensity, participants performed a surprise free recall task immediately after exposure during which they self-reported the emotions they experienced via free-response and the intensity of those emotions during self-selected events from the haunted house. Participants also noted the direction of their regulation attempts (i.e., upregulation, downregulation, no regulation), how effective those regulation attempts were, and the means by which they attempted to regulate in their own words. We report how we determined our sample size, all data exclusions, all manipulations, and all measures in the study.

**PARTICIPANTS:**

**Pilot Study.** A pilot study was conducted in October 2019 consisting of 54 participants (age: *x̄ age* = 24.22 yrs, range = 18 – 34 yrs, *sdage* = 3.97 yrs; gender: 26 female, 27 male, 1 non-binary; Ethnicity: 18.51% Hispanic, 81.49% Not Hispanic) who were recruited from a large northeastern city via flyers for an IRB-approved fear and memory study. Piloting allowed us to determine the distribution of ER strategy usage in this context and determine a more appropriate sample size for our primary study, as a review of the extant literature did not yield analogous study designs within the ER space. *A priori* power analyses for this pilot were conducted using the smallest effect size reported by Sheppes et al.’s 2011 examination of emotional intensity and regulatory choice (*ηp2* = 0.43). Using “WebPower” (Zhang & Mai, 2019) in R 3.6.1 (R Core Team, 2022), we determined 18 participants would sufficiently power our main effect in a typical lab context. Given the additional complications our study design introduces which would likely reduce the effect size between our variables of interest and the resources we had available, we surmised a sample of 54 participants, three times the minimum sample size, might be sufficient to identify an effect. This analysis failed to find a statistically significant association between affective intensity and strategy usage, but our best performing multilevel binary logistic regression model [strategy ~ intensity + (1 | Participant)], as determined by an AIC comparison information theoretic approach, produced an odds ratio of *OR* = 1.83 (*95% CI* = [0.65, 3.2], *p* = 0.079) (*See* **Supplementary Materials** for more pilot design and analyses details). This observed model from the pilot was used to generated an *a priori* power curve for Experiment 1 via “simr” (Green & MacLeod, 2016) in R 3.6.1, which estimates the proportion of simulated datasets in which the null hypothesis is rejected given the target model. This approach to calculating power is preferrable for hierarchically-structured data because it does not assume independence of observations. We determined that at least 76 participants with an average of 3 observations each (228 observations total) would be required to sufficiently power our experiment based upon simulations with the observed data (Two-tailed, α = 0.05, 1–β= 0.80, *Pr(Y =1|X=1)* *H0* = 0.615). We increased the target sample size to 120 participants, due to attrition concerns and the needs of a concurrently-ran experiment on fear and memory. However, we were only able to recruit 98.33% of our recruitment goal due to time constraints (i.e., all data collection must occur before the haunted house closes for the season).

**Experiment 1.** In October 2021, 118 participants (age: *x̄ age* = 20.80 yrs, *range* = 18 – 34 yrs, *sdage* = 2.87 yrs; gender: 73 female, 40 male, 5 non-binary) were recruited from a large northeastern city via flyers for an IRB-approved study on fear processing at a local haunted house attraction. A description of Study 1’s design is illustrated in **Figure 1**. Eligible participants were native English speakers between the ages of 18 and 35, had normal or corrected-to-normal vision, were not pregnant, had no history of seizures, cardiovascular issues, or neurological disorders, could comfortably walk for at least one hour, and had not been to this haunted house in the past. On average, participants were more educated (*x̄ Years of Education* = 16.90 yrs, *sd Years of Education* = 2.75 yrs) than the average US adult, who according the U.S. Census Bureau’s American Community Survey, has completed 13.7 years of education. Categorically, 81.9% reported having completed some college (58.6%), a 4-year degree (12.9%), some post-graduate studies (03.4%), or a post-graduate degree (06.9%). Socioeconomic status was slightly negatively skewed, with 14.5% of respondents reporting making less than $15,000 per year, 07.7% reporting between $15,001 and $25,000, 07.7% reporting $25,001 to $35,000, 05.1% reporting between $35,001 and $50,000, 22.2% reporting between $50,001 and $75,000, 12.8% reporting between $75,001 and $100,000, 17.9% reporting between $100,001 and $150,000, and 12.0% reporting greater than $150,000. The racial and ethnic identities of participants were not assessed. Participants were compensated $60.00 in Visa debit cards.

**Fig 1.** Study 1: Task Overview - One hundred and eighteen (118) participants traversed a haunted house in small groups. a. Prior to the haunted house, participants completed baseline questionnaires outside of the event at a local park. b. The haunted house lasted for ~37 minutes. c) Participants then immediately recalled three events, and their attempts to regulate them post exposure. d) They then again recalled the same three events and an additional six events at an online follow-up session.



**MATERIALS AND PROCEDURE:**

**Pre-Exposure**. Participants reported to a provisional headquarters (i.e., tent, tables, chairs, computers, etc.) constructed just outside the haunted house property to complete individual difference questionnaires, questionnaires assessing prior knowledge of the haunted house, expectations, and motivations for participating, as well as a measure of cognitive load. To assess prior knowledge, participants indicated the number of times they had attended other haunted houses, how much information they feel they know about this specific haunted house (1 = ‘None at all’ to 5 = ‘A great deal’), and from which sources had they learned information about this specific haunted house (e.g., advertisements, friends, news reports, etc.). To assess expectations, participants were asked how much positive and negative emotion they anticipated experiencing prior to exposure on a 5-point Likert scale, with 1 being ‘None at all’ and 5 being ‘A great deal’. Participants were also asked how fearful they felt, how sensitive they were to startling stimuli, how much they enjoy haunted houses, and how much they enjoy feeling fear (all assessed on Likert scales). Motivations to participate were assessed using 100-point sliding scales and included payment, thrills, novel experiences, challenges, social pressure, scientific interest or duty, and boredom. Cognitive load was assessed prior to exposure, immediately after exposure and at a later follow-up using a 15-item Remote Associates Test (RAT). Forty-five RAT items were selected for their difficulty as measured by Bowden’s 15-second trials, such that each item had two equally difficult counterparts which could be randomly assigned across the three timepoints (Bowden & Jung-Beeman, 2003). Following instructions, participants completed three practice trials with feedback. During the RAT task, participants had 15 seconds to identify the target word and did not receive feedback. Participants were then fitted with heartrate monitors and escorted to the haunted house entrance. All pre-exposure questionnaires and materials are available within our OSF repository (*See* **Open Practices**).

**Exposure (Haunted House).** Participants experienced the haunted house in 31 groups across 11 nights (*x̄ size* = 3.81 participants; *sd*size = 1.12 participants). This specific haunted house was chosen because: 1) it uses professional actors renowned for eliciting a range of affective responses; 2) it contained four themed sections each with a unique aesthetic providing variability to the stimuli; 3) it provides a remarkably consistent experience across sessions; and 4) coordination with the facility granted us special privileges to use equipment (i.e., computers, heart-rate monitors) and better guarantee consistency across exposures (i.e., entering before other guests, keeping groups together).

Participants received minimal instructions to promote ecological validity (participants were to walk through the haunted house in a single file line and avoid sharing thoughts, reactions, and experiences with other participants). However, they were encouraged to act and react as naturally as possible without interacting with each other (i.e., grabbing, holding, touching, etc.). As part of a concurrently-run memory experiment, some participants (*n =* 58) did receive additional instruction to navigate specific sections as if they would later have to complete a memory test or write a review of that section. However, these groups did not statistically differ significantly from the control group in reported affective intensity (*F*(2,81) = 1.41, *p* = 0.25), regulation strategy usage (*F*(2,73) = 0.05, *p* = 0.95), or regulation strategy success (*F*(2,81) = 1.93, *p* = 0.15) as determined by ANOVA. The data relevant to this memory experiment is beyond the purview of this study and is better captured within another forthcoming manuscript. The accompanying research assistant led the group through each section. The approximate exposure time was 37.4 minutes.

**Post-Exposure.** Following exposure, participants completed immediate post-exposure assessments at a remote site outside of the haunted house. Participants were tasked with identifying three emotionally salient events that occurred from a randomly selected haunted house section and reporting affective and regulatory details of that event. Events from other haunted house sections were not tested during the immediate exposure session to avoid conflicts with the aforementioned concurrently-running memory experiment. Participants were not trained in ER strategies, nor were they primed to consider their ER strategies prior to these questionnaires. Though experiential sampling methods (ESM) and ecological momentary assessments (EMA) are often applied in naturalistic settings to capture moment-to-moment fluctuations in emotion and regulation with considerable success (*See* Colombo et al., 2020; Shahane et al., 2023), we chose post-exposure assessments as our means of assessment in part for the following reasons: 1) assessing participant reactivity mid-exposure can lead participants to alter their affective and regulatory behaviors in response to the assessment prompts (Stone et al., 2003), 2) training participants to use ESM or EMA technology effectively in this context would require training in ER strategy categorization in violation of the goals of this study, and 3) the use of such technology during exposure would violate the immersive, high-intensity nature of the context (Shiffman et al., 2008). For each of the three events, participants wrote a detailed description of what occurred. Participants then noted which emotions (up to a total of five) they felt during this event via free response. Participants rated how intense their emotion felt on a 5-point Likert scale (labels included, in order of increasing intensity: ‘None at all’, ‘A little’, ‘A moderate amount’, ‘A lot’, and ‘A great deal’). Participants were also tasked with describing how they tried to regulate each emotion, if at all, via free response. People’s accuracy in recalling details of similar real-world experiences via free response was surprisingly high in recent investigations using similar methodology (Diamond et al., 2020). Participants were also asked directly whether they attempted to down- or up-regulate their experiences, how successful their regulatory efforts were, and regulatory responses were assessed in response to each emotion rather than each event. Thus, rather than exploring this phenomenon at the event-level, which might require regressing the probability of using a strategy upon the average intensity of all emotions experienced in that event – an assumption we would not make in confidence – we draw associations between regulation strategy usage and the emotions that participants identify as directly motivating them. We refer to data captured at this time point as being “immediately reported”. Following completion of immediate post-exposure measures, participants were dismissed, instructed to not discuss their experiences, and to remain in contact with researchers for a follow-up session which was conducted remotely (*time since exposure: x̄ delay* = 7.01 days; *sddelay* = 0.91 days) to assess how memory of self-regulation and memory accuracy in this context alters over time. At this delayed follow-up session, participants were tasked with identifying six additional events and their affective and regulatory responses to each.

**Strategy Usage Coding**. During piloting, two hypotheses-blind raters classified strategy descriptions into one or more strategy categories: Reappraisal, Distraction, Suppression, Situation Selection, Situation Modification, or ‘None of the above’ (IRR Agreement = 0.880). Raters were also provided the participant’s description of the event and the emotions they experienced which they indicated having downregulated. Raters were undergraduate research assistants who trained by first reviewing examples of landmark literature which defined the strategies of interest as commonly used in the field (Gross, 1998, 2002). Raters then reviewed select methodological excerpts from experimental papers to see how cognitive reappraisal, attention deployment, and other Process Model strategies were defined within past studies (Shafir et al., 2016; Sheppes et al., 2011). Lastly, raters independently completed classification exercises using examples of regulation strategy descriptions from the same context but which were collected during piloting. Through the training and classification process, raters were instructed not to collaborate or discuss their ratings with each other during the rating process. After individually classifying each description, a moderator (W.M.) compared the ratings for disagreements (i.e., cases in which raters disagreed on how a regulation event should be classified). The moderator then met with both raters remotely using a digital video conferencing platform and moderated a review of the classifications, asking raters to compromise in cases of classification disagreement. The moderator’s role was to facilitate discussion of classifications and document their conclusions, but was not involved in the discussion and disconnected during them (i.e., muted their microphone; turned off camera) to avoid unduly influencing the outcome.

This approach revealed that distraction and reappraisal were by far the most commonly used strategies in this context. Of the 182 self-reported downregulated events recorded in the pilot, 30.7% used reappraisal and 61.5% used distraction, with the other three strategies (i.e., suppression, situation modification, situation selection) combined appearing in fewer than 20% of events (Note that the total percentage sum is greater than 100% due to the occurrence of multi-strategy experiences). Two hypotheses-blind raters classified each observation’s strategy description in Study 1 into one or more strategy categories: Reappraisal, Distraction, Suppression, a combination of the three, or none of the above (IRR Agreement = 0.877). Situation Modification and Selection were excluded due to the infrequency of their use. Suppression was also used infrequently in Study 1, but was categorized by Study 1 raters due to its large presence within the broader literature. Raters were undergraduate research assistants who were trained using the same methodology described in the pilot study, but were not the same raters from the pilot study. The training materials and instructions generated for this purpose have been made available within our OSF repository (*See* **Open Practices**).

**Questionnaire response processing and coding**. Emotion responses were processed by: 1) removing entries lacking intelligible affective information (e.g., “-“, “nothing”, “idk man”), 2) removing unnecessary punctuation, hyphenation, and qualitative modifiers (e.g., “very sad” becomes “sad”, 3) splitting compound emotion response (e.g., “sad / angry” becomes “sad” and “angry”, 4) correcting spelling errors according to the top suggestions recommended by R’s native spell checking software, 5) lemmatization (e.g., “annoyance”, “annoying”, and “annoyed” become “annoy”). These modified emotion responses were then merged with the NRC lexicon which contains over 20,000 English emotion words human rated by valence, arousal, and dominance (Mohammad, 2018). Valence was determined using NRC lexicon valence scores. Observations without an associated NRC lexicon entry were dropped due to lack of valence data.

**Analysis.** To explore our primary question, the effect of emotional intensity upon regulatory strategy usage, we specified mixed effect binary logistic regressions accounting for the random effect of participants using the “lme4” package (Bates et al., 2015) in R (R Core Team, 2022) and followed an information theoretic approach via AIC comparison, which prioritizes model parsimony and penalizes models with excessive variables. All data and scripts used to produce this analysis are publicly available at OSF (*See* **Open Practices**). Preregistration for Experiment 1 methods and hypotheses is publicly available at As Predicted (https://aspredicted.org/DP1\_453).

**EXPERIMENT 1 RESULTS**

A subset of 298 observations in which a negative emotion was downregulated by either distraction or reappraisal was used for analyses related to our primary hypothesis. These observations were reported by 77 participants. Of the 298 total observations, 175 (58.72%) reported using distraction to regulate their emotions. The average emotional intensity of observations was 2.41 (*range* = 0 – 4, *sd* = 0.932). **Figure 2b** illustrates the variance in all emotions and **Figure 2a** illustrates the emotions endorsed within the dataset of 298 observations.

**Thrill-seeking motivated study participation.** Although horror attractions may feature aspects of emotional experiences that are difficult to capture in more traditional paradigms, they may have limitations as well. For example, such a setting may generate self-selection biases in our sample. To monitor this, participants were asked about their expectations and motivations to participate in this study. Participants were asked how much they enjoyed fear on a 0 to 6 scale, with the average response sitting above the midpoint of the scale (*x̄* = 3.40, *median* = 4, *sd* = 1.82). Participants were also asked, “To what extent are the following items motivating your participation?” on a 0 to 100 scale with items including: the payment received for participating (*x̄* = 59.2, *median* = 59, *sd* = 31.1), the thrills they may feel in the haunted house (*x̄* = 63.3, *median* = 73, *sd* = 31.7), the opportunity for new experiences (*x̄* = 70.1, *median* = 78, *sd* = 28.0), the opportunity for challenging experiences (*x̄* = 53.3, *median* = 50, *sd* = 30.3), social pressures (*x̄* = 15.5, *median* = 02, *sd* = 25.7), a desire to help science (*x̄* = 59.4, *median* = 58, *sd* = 29.9), and boredom (*x̄* = 28.0, *median* = 15, *sd* = 31.4). Thrill-seeking motivations strongly correlated with fear enjoyment (*r* = 0.582, *p* < 0.001) and Bonferroni-adjusted contrasts did determine that thrill-seeking as a motivation was significantly greater than boredom (*t*(144) = 7.17, *p* < 0.001), social pressure (*t*(144) = 9.70, *p* < 0.001), and the pooled average of all motivations (*t*(504) = 4.19, *p* < 0.001), but not challenge-seeking (*t*(144) = 2.05, *p* =0.289), novelty-seeking (*t*(144) = -1.38, *p* = 1.000), payment (*t*(144) = 0.84, *p* = 1.000), or science participation (*t*(144) = 0.81, *p* = 1.000). This may suggest that our participants were slightly higher in thrill-seeking motivations relative to what we might expect to find in an average population.



**Fig 2.** Frequency-weighted word clouds illustrating: **a.)** allemotions endorsed during Study 1 and **b.)** all negatively-valenced emotions which were downregulated using either reappraisal or distraction during Study 1. The size of the term positively correlates with the frequency with which it was self-reported by participants.

1. **B.**

**Study participants regulated their emotions without prompt.** Another limitation may be a lack of motivation to self-regulate. Because haunted houses are entertainment, participants may be less inclined to self-regulate, regardless of which strategy they use. Within the pilot study, we found a positive association between state anxiety measured immediately before the haunted house and the average extent to which participants attempted to regulate their emotions (assessed through the question, “To what extent did you attempt to change or regulate how you felt during this event?”) across events (*β* = 0.34, *95% CI* = [0.054, 0.626], *p* = 0.021), as well as a positive association between the average negative affective intensity of an event and extent to which participants attempted to regulate their emotions during that event (*β* = 0.31, *95% CI* = [0.17, 0.46], *p* < 0.001) (*See* **Supplementary Materials** for more pilot design and analyses details). Both suggest, as expected, that participants within this setting experiencing negative emotions do self-regulate these emotions without researcher-prompting. Self-regulation extent or efforts were not assessed within Study 1 to reassess this relationship.

**Intensity did not predict regulatory strategy usage.** To test our primary hypotheses, models using either z-scored emotional intensity or person-centered emotional intensity as the primary predictor were constructed iteratively (i.e., adding one fixed effect to the model at a time), but across all model comparisons, no model performed better than our null model which did not feature any fixed effects (*ICC* = 0.40). Our best performing non-null model, including only intensity as a fixed effect (*p* = 0.10 when compared to null), did not find a relationship between emotional intensity and strategy usage (*OR* = 1.36, *95% CI* = [0.95, 1.95], *p* = 0.10) (**Fig. 3**). An odds ratio of 1.36 suggests that for every one standard deviation unit increase in emotional intensity, the odds of choosing distraction to regulate increase by approximately 36%, but importantly, this association is not statistically significant and markedly smaller than what might be observed in more controlled strategy selection paradigms. As such, we did not find evidence to support that emotional intensity predicts strategy usage in this dynamic, high-intensity situation.

**Fig 3.** Across almost all tested mixed effects binary logistic regression models, emotional intensity failed to predict strategy usage. Visualized is our model using only emotional intensity (*x̄* = 2.41, *sd* = 0.932) to predict regulation strategy choice between reappraisal and distraction among negative emotions. Regression line represents likelihood of selecting distraction as opposed to reappraisal at any given emotional intensity value. Points represent individual observations. Regression ribbon represents standard error.



**Cognitive load did not predict regulatory strategy usage.** The association between strategy choice and cognitive load has been well documented in the literature (Dorman Ilan et al., 2019), but its replication beyond lab settings has been limited. Given the high-intensity, often overwhelming nature of this setting, we had preregistered the hypothesis that cognitive load should demonstrate a positive association with the probability of using distraction as an ER strategy. However, we failed to find evidence that cognitive load post-exposure, as assessed by the RAT test predicted strategy usage during exposure (*b* = - 0.02, *95% CI =* [-0.044, 0.010], *p* = 0.21), even when adjusting for baseline cognitive load, as assessed both prior to exposure and one-week later.

**Nuisance variables did not predict regulatory strategy usage.** In addition to our primary analyses, we evaluated the impact of several nuisance variables. These are variables that, according to existing research, might confound the relationship between affective intensity and strategy usage, but were not of primary theoretical focus for these experiments. Examples measured in this study include emotion expectations (Denny et al., 2014), motivations to participate (Tamir, 2016), attitudes towards fear and haunted houses (Argyriou & Lee, 2020), sex (McRae et al., 2008), age (Blanchard-Fields et al., 2004), depression (BDI-II), anxiety (STAI), intolerance of uncertainty (IUS) (Aldao et al., 2010), regulation tendencies (ERQ) (Gross & John, 2003), time of day, and presence of peers.

We did not find significant associations between the proportion of events in which distraction was used and how positively (*b* = 0.035, *95% CI =* [-0.032, 0.102], *p* = 0.30) or negatively (*b* = 0.047, *95% CI =* [-0.017, 0.112], *p* = 0.14) participants expected to feel during the study. We also did not find an association between distraction versus reappraisal usage and the motivations cited for participation, including payment (*b =* -0.000, *95% CI =*[-0.003, 0.002 ], *p =* 0.629), thrill-seeking (*b =* -0.000, *95% CI =*[-0.003, 0.002], *p =* 0.801), novelty-seeking (*b =* 0.001, *95% CI =*[-0.002, 0.004], *p =* 0.454), peer influence (*b =* 0.001, *95% CI =*[ -0.002, 0.003], *p =* 0.595), boredom (*b =* -0.001, *95% CI =*[ -0.004, 0.001], *p =* 0.341), contributing to science (*b =* 0.000, *95% CI =*[-0.002, 0.002], *p =* 0.858), and seeking a challenge (*b =* 0.000, *95% CI =*[ -0.003, 0.002], *p =* 0.935). We additionally did not find any associations between distraction usage and how much participants self-reported enjoying fear (*b =* -0.012, *95% CI =*[-0.050, 0.027], *p =* 0.55) or haunted houses (*b =* 0.021, *95% CI =*[ -0.024, 0.066], *p =* 0.35), sex (t(73) = 1.54, *95% CI =*[ -0.029, 0.225], *p =* 0.13), age (*b =* 0.000, *95% CI =*[-0.024, 0.023], *p =* 0.98), depression (*b =* -0.004, *95% CI =*[-0.024, 0.016], *p =* 0.70), anxiety (*b =* 0.000, *95% CI =*[-0.009, 0.008], *p =* 0.97), intolerance of uncertainty (*b =* 0.000, *95% CI =*[-0.005, 0.006], *p =* 0.92), tendency to use reappraisal (*b =* 0.000, *95% CI =*[-0.012, 0.013], *p =* 0.97) or suppression (*b =* 0.002, *95% CI =*[-0.011, 0.015], *p =* 0.76), time of day (*F*(2,73) = 0.04, *p =* 0.96), presence of familiar peers (*t*(60) = -0.40, *p =* 0.700), or which group participants traversed the haunted house with (*F*(30,45) = 0.93, *p =* 0.57). Despite the lack of association, these variables were included as covariates in some models as part of our subsequent multiverse analysis.

**Multiverse approach also failed to explain strategy usage.** We expanded the scope of our primary analyses and conducted additional exploratory analyses to determine whether a stronger association between strategy usage and affective intensity could be found using different inclusion criteria, comparing engagement strategies (i.e., reappraisal) to disengagement strategies (i.e., suppression, distraction) as defined in the broader literature (e.g., Dixon-Gordon et al., 2015) and including contra-hedonic regulation activity (i.e., downregulating positive emotion). We found only a single model which surpassed nominal statistical thresholds of significance in model fit (*ICC* = 0.37; *p* = 0.04 when compared to null), but which did not maintain significance after adjusting to maintain a family-wise error rate (*p* = 0.32). Details and results of this analysis can be found in the **Supplementary Materials**.

**Regulatory strategy usage and intensity interact to predict regulatory success.** Following our emotional intensity analyses, we explored how strategy usage moderated the relationship between intensity and success, as using distraction during high-intensity emotional states should more successfully regulate emotions than using reappraisal (Sheppes et al., 2011). After constructing a series of multilevel linear models and again following an information theoretic approach, we found that our best-performing model did indeed include an interaction between strategy usage and emotional intensity (*ICC* = 0.42, *p* = 0.003) and found that interaction to be significant (*β* = 0.25, *95% CI* = [0.09, 0.42], *p* = 0.003). However, a simple slopes analysis revealed a surprising finding: no relationship was observed between regulatory success and emotional intensity for events regulated via reappraisal (*β* = -0.03, *95% CI* = [-0.16, 0.10], *p* = 0.70), but regulatory success was *negatively* associated with emotional intensity for distraction-regulated events (*β* = -0.28, *95% CI* = [-0.40, -0.16], *p* < 0.001) (**Fig. 4**). In other words, each standard deviation unit increase in emotional intensity yielded a -0.28 standard deviation decrease in the reported effectiveness of distraction, but not reappraisal, as an ER strategy. Though the extant literature from comparable lab studies should motivate us to expect the efficacy of distraction to increase and reappraisal to decrease as affective intensity increases, our data seems to document a deviation from this pattern in a high-intensity, quasi-naturalistic setting: distraction appeared to be less – not more – successful as affective intensity increased.



**Fig 4.** Strategy moderated the relationship between emotional intensity and regulatory success (β = 0.25, p = 0.003). While the success of reappraisal was relatively unrelated to emotional intensity, distraction demonstrated a negative association with emotional intensity, contrary to what extant literature might suggest. Given the frequency with which distraction was reported, the relative underperformance of distraction at high emotional intensities may partially explain the absence of an association between strategy choice and emotional intensity within our study.



**EXPERIMENT 2 METHODS**

Hypothesized models across a pilot study and Experiment 1 both failed to find an association between affective intensity and strategy usage in a complex, high-intensity, multimodal, environment using untrained, undirected participants, even when adjusting for noted moderators like cognitive load. Exploratory models found that the relationship may appear, albeit weakly, when including positive and negative emotions, and that this lack of an effect may be due to distraction being less successful than hypothesized in this complex, multimodal experience. We hypothesized that participants exposed to similar information outside of the complex, multimodal environment would likely still demonstrate ER usage patterns in line with the extant literature. We hypothesized that generating a decontextualized representation of the experience with only the relevant information (i.e., description of event, emotions felt, intensity of emotions) present and asking participants to simulate or forecast how they might self-regulate would reproduce the positive association between the emotional intensity of an experience and the frequency of choosing disengagement over engagement regulation strategies.

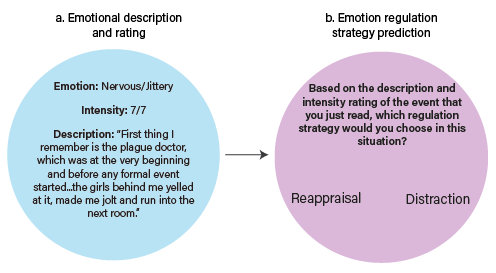
Although using audiovisual recordings from the experience would have been ideal, we were unable to obtain permission to record such data during the previous studies. However, we had access to the descriptive information reported by participants who experienced these events (i.e., their text-formatted memories of the events). If there is a difference between participants simulating (i.e., forecasting) self-regulation and participants performing (i.e., experiencing) self-regulation, it may further emphasize the complications that dynamic, multimodal contexts introduce to the ER space. In Experiment 2, a novel set of participants were presented with information about the events that motivated ER reported by pilot participants (events available within OSF repository), but not the regulatory behaviors participants used, and asked to predict which regulation strategy they would employ based upon the information provided. We report how we determined our sample size, all data exclusions, all manipulations, and all measures in the study.

**PARTICIPANTS:** In July 2021, 170 participants (age: *x̄ age* = 34.34 yrs, *range* = 18 -75 yrs, *sdage* = 14.31 yrs; gender: 100 female, 68 male, 2 non-binary) consented to an IRB-approved online study described as measuring individual differences in choice predictions. Participants completed the study on Qualtrics and were recruited/filtered via Prolific. Eligible participants were native English speakers residing in the US between the ages of 18 and 85, had normal or corrected-to-normal vision, had no history of reading-related disorders or literacy difficulties, as well as no history of mild cognitive impairment, head injury leading to unconsciousness, or unregulated mental health diagnosis. The racial identity of participants were as follows: 13.6% Asian, 06.8% Black, 04.3% Mixed, 03.7% Other, and 71.6% White. Although socioeconomic status data is not available, 45.2% of participants reported working full-time, 19.2% reported working part-time, 24.7 % reported not working full- or part-time, and 11.0% did not specify their work status.

Sample size was determined via a priori power analyses assuming an attrition rate of 10%, *r2* ≤ 0.10 for covariates, and a small effect size (*OR* = 1.68), which suggested we must recruit at least 163 participants to achieve 1-β = 0.80 (*α* = 0.05, two-tailed); however, this approach had not taken into account the hierarchical nature of our observations and likely underestimates our true power. Eighteen participants were excluded for failing attention checks (*n* = 7), failing to complete the study (*n* = 9), and scoring a Q Recaptcha Score lower than 0.7, indicating significant bot activity (*n* = 2). Participants were paid at a rate of $10.25/hr.

**MATERIALS AND PROCEDURE:** Details from seventy-eight (78) negatively-valenced pilot study events regulated through either reappraisal or distraction were presented to online participants. Participants first read definitions of both reappraisal (thinking about the experience in a way that reduces the intensity of the negative emotions) and distraction (looking or thinking about something else that is emotionally neutral) and reviewed examples of how both strategies might be employed. These haunted house-specific examples included: a.) making an effort to remind oneself that the people are just actors who are using props, rather than zombies trying to hurt them, and b.) choosing to look down at one’s feet or focusing on what one ate for lunch rather than focusing on the zombies coming after you, for reappraisal and distraction specifically. Participants performed a brief practice task which required successfully defining and applying both categories before the primary task began (**Fig. 5**). Participants were provided an opportunity to pause participation and contact research staff if they had questions about definitions or strategy application before proceeding. All 78 events were randomized and serially presented. These events were screened for information pertinent to reappraisal and distraction that may unduly influence participant decisions. For each event, the emotions experienced, the intensity of each emotion, how the experiencer described the event, and definitions for both strategies were displayed. Participants were then asked to predict which strategy they would choose to reduce the emotional intensity of the situation. Participants were granted as much time as needed to complete the task. Following the primary task, participants completed individual difference measures, including the Emotion Regulation Questionnaire (ERQ), the Difficulties with Emotion Regulation Survey (DERS), and the Intolerance of Uncertainty Scale (IUS). Participants completed the study in 33.7 minutes on average (*median* = 31.5 minutes, *sd* = 14.2 minutes, *range* = 10.9 – 88.4 minutes).

**Fig 5.** Study 2: Task Overview - One hundred and seventy (170) participants (forecasters) read the descriptions that pilot study participants (experiencers) wrote about their emotional experience in the haunted house. a. Forecasters read the experiencers’ emotional descriptions and intensity rating. b. Forecasters indicated what regulation strategy (distraction or reappraisal) they would use to regulate their emotions in the described event.



**Analysis.** To explore whether the affective intensity experiencers reported influenced the strategies forecasters chose, we again specified mixed effect binary logistic regressions accounting for the random effect of participant (both forecaster and experiencer) using the “lme4” package (Bates et al., 2015) in R (R Core Team, 2022) and followed an information theoretic approach via AIC comparison. Fixed effects models were built iteratively from and compared to our null model (ICC = 0.14). All data and scripts used to produce this analysis are publicly available at OSF (*See* **Open Practices**). The design and hypotheses of Study 2 were preregistered with AsPredicted (<https://aspredicted.org/XXH_W1V>), though please note a discrepancy exists in the number of events used, as fewer met our outlined criteria than initially determined.

**EXPERIMENT 2 RESULTS**

**Study samples were similar across individual difference measures.** Our first analyses aimed to determine whether relevant trait differences existed between the online sample of participants and the participants who experienced the haunted house. If such differences exist, they would limit our ability to associate differences in regulatory forecasting or usage to differences in presentation and context. Both pilot and experiment 2 participants completed the ERQ and IUS. If differences exist in ERQ scores, the groups may differ in their underlying propensity to choose specific strategies. Using a Welch’s Two Sample T-Test, we did not find significant differences between the groups in their likelihood of using reappraisal (*xpilot* = 29.9, *xexp2*= 31.0, *95% CI* = [-3.22, 1.03], *t*(44.6 = -1.04, *p* = 0.305) according to the ERQ reappraisal subscale. However, significant differences were observed between the two groups for the expressive suppression ERQ subscale (*xpilot* = 12.5, *xexp2* = 15.7, *95% CI* = [-5.03, -1.32]. *t*(48.1) = -3.44, *p* = 0.001). The relevancy of the suppression subscale is unclear in this context, as suppression is not directly tested and neither subscale proved to be predictive of strategy usage during our pilot. Furthermore, a bivariate linear model found that neither the ERQ reappraisal subscale (*b* = -0.001, *se* = 0.002, *p* = 0.58) nor the suppression subscale (*b* = 0.002, *se* = 0.002, *p* = 0.35) predicted the proportion of trials in which participants selected reappraisal rather than distraction in this study as well. Additionally, if differences exist in IUS scores, the groups may differ in how they respond to ambiguity or uncertain situations; perhaps especially relevant given that “confused” was the most cited negative emotion in Study 1. However, no differences were observed in IUS scores across groups (*xpilot* = 33.6, *xexp2*= 34.3, *95% CI*= [-7.07, 5.59], *t*(34.9) = -0.238, *p* = 0.813). Assessing differences in difficulties in applying emotion regulation strategies via DERS subscales was not possible because it was not administered to pilot participants.

**Intensity predicts regulatory strategy forecasts.** Our first model, containing only the affective intensity of pilot participants as a predictor to predict online participants’ strategy forecasts, performed better than our null model (*χ2*(1) = 8.39), demonstrating a small positive effect (*OR* = 1.06, *95% CI* = [1.02, 1.10], *p* = 0.004) such that more intense events were associated with an increased probability that participants chose distraction to regulate them. Additional models included age, gender, IUS score, the DERS limited access to strategies subscale, and a combination of all four as covariates. However, each model failed to outperform our initial model in chi square tests of the models (age: *χ2*(1)= 0.84*, p* = 0.36; gender: *χ2*(2)= 2.81, p = 0.25; IUS: *χ2*(1)= 1.25, p = 0.26; DERS: *χ2*(1)= 0.54, p = 0.46; combo: *χ2*(5)= 4.52, p = 0.48; respectively).

**Online participants matched pilot participants less than chance when selecting distraction.** In signal detection theory, d prime (*d’*) is a measure of sensitivity that quantifies the ability to distinguish between signal and noise in a binary decision task, and thus, can be used to determine whether congruency in strategy forecast and usage between those who forecasted ER and those who experienced it is greater than chance. D prime is calculated as the normalized value of the proportion of hits (i.e., when a strategy was used by a pilot participant and selected by an Experiment 2 participant) minus the proportion of false alarms (i.e., when a strategy was not used by an experiencer but was selected by forecaster). Importantly, d’ is robust to unequal prior probabilities in binary outcomes, as occurs in our strategy selection (73.1% of pilot observations used distraction) through the incorporation of a bias parameter. Using this approach, we found that forecasting participants matched the experiences of pilot participants in their selection of reappraisal nearly at chance (*d’* = -0.08) but were below chance in matching distraction (d’ = -0.41) (**Fig. 6**). The difference in selection congruency between these strategies was significant as determined by a paired samples t-test (*xdiff* = 0.328, *95% CI* = [0.313, 0.342], *t*(149) = 43, *p* < 0.001), suggesting that differences in the deployment of distraction between the two groups may be driving differences in how predictive affective intensity was towards strategy forecast or usage across these two contexts. The distribution of strategy selection differed between our online sample and haunted house sample, as distraction was only selected in 48.8% of observations for Experiment 2. Taken together, we found that participants used emotional intensity to inform their ER strategy predictions but not strategy usage and that the predicted utility, and thus frequency, of using distraction in such a context might not reflect what is experienced in practice.



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**Fig 6.** The strategies participants predicted they would use were assessed against the strategies that were reported as used in each situation. While neither strategy was reported greater than chance, participants were significantly more likely to predict using reappraisal on trials when distraction was actually used when compared to predicting distraction when reappraisal was actually used.

**EXPERIMENT 3 METHODS**

Though affective intensity failed to predict ER usage in a high-intensity, quasi-naturalistic setting, participants presented with descriptions of events that were regulated by haunted house ‘experiencers’ were more likely to forecast using distraction to regulate events described as high intensity and reappraisal for events described as being of a lower intensity. The different results observed in these experiments are difficult to compare, though, as a number of features differ between the approaches. Participants in Study 1 experienced the haunted house, were not trained or prompted to self-regulate, experienced the emotions that they reported, and reported the regulatory strategies that they used. Participants in Study 2 simulated the experiences of the haunted house, were instructed in ER strategies and prompted to make a forecast, did not experience the emotions associated with the event, and reported what strategy that they forecasted they might use in the given situation. Furthermore, Study 1 did not attempt to manipulate the emotions of participants. Study 3 attempted to rectify these discrepancies using a 2 between (condition: forecast v. experience) x 2 within (intensity: low v. high) mixed study design. As we could not incorporate an immersive experiential component such as in Study 1 in this experimental design[[1]](#footnote-2), we instead used clips from lesser-known horror films to elicit regulatory response to feature-rich, dynamic representations of negative emotion. This design allows us to directly compare the differences in regulatory behavior among untrained and unprompted participants when either reporting actual usage or forecasted selection in response to high and low intensity stimuli.

We originally hypothesized that we would replicate the results of Study 1 and Study 2: that participants placed within the “experiencer” condition would not demonstrate a relationship between stimulus intensity and regulation strategy usage, while participants in the “forecaster” condition would demonstrate the canonical relationship between stimulus intensity and regulation strategy usage (i.e., as affective intensity increases the likelihood of relying upon distraction also increases). Counter to our predictions, we found that both conditions demonstrated characteristics of this canonical relationship; however, important differences in ER strategy usage and effectiveness were also observed across the two conditions. We report how we determined our sample size, all data exclusions, all manipulations, and all measures in the study.

**PARTICIPANTS:** In October 2023, 247 participants (age: *x̄ age* = 38.57 yrs, *range* = 20 -76 yrs, *sdage* = 12.24 yrs; gender: 94 female, 152 male, 1 non-binary) consented to an IRB-approved online study described as using clips from horror films to explore emotions and behavior. Participants completed the study on Pavlovia and were recruited/filtered via Prolific. Eligible participants were native English speakers residing in the US between the ages of 18 and 85, had normal or corrected-to-normal vision, had no history of reading-related disorders or literacy difficulties, as well as no history of mild cognitive impairment, head injury leading to unconsciousness, or unregulated mental health diagnosis. The racial identity of participants were as follows: 08.9% Asian, 06.9% Black, 05.7% Hispanic, 09.3% Mixed, 04.6% Other, and 64.6% White. Socioeconomic status was well distributed with 09.3% reporting a household income of under $15,000 per year, 09.3% between $15,001 and $25,000, 10.5% between $25,001 and $35,000, 13.4% between $35,001 and $50,000, 19.0% between $50,001 and $75,000, 13.8% between $75,001 and $100,00, 08.1% between $100,001 and $150,000, 06.9% above $150,000 and an additional 04.9% who did prefer not to say. In response to the question, “How much do you enjoy watching horror movies”, the average participant response was a 2.95 on a scale from 0 to 6, with 0 corresponding to “Not at all” and 6 corresponding to “Extremely” (*median*  = 3, *sd*  = 2.16).

Sample size was determined via a priori power analyses assuming an attrition rate of 10%, *r2* ≤ 0.10 for covariates, and the most conservative effect size observed in our previous studies (*OR* = 1.30). To achieve 1-β = 0.80 (*α* = 0.05, two-tailed), at least 240 participants were required, or 264 when accounting for attrition. Up to 300 slots were made available on Prolific to account for returned or incomplete research participation. Five participants were excluded for failing attention checks (*n* = 1) and familiarity with the stimuli (*n* = 4). Though we did not meet our recruitment goal of *n* = 264, our sample (*n* = 242) still surpassed the threshold calculated to achieve sufficient power. Participants were paid at a rate of $12.00/hr.

**MATERIALS AND PROCEDURE:** Participants completed a web browser-based task built primarily with jsPsych v7.0 (De Leeuw et al., 2023). Following consent and a disclaimer regarding the content of the stimuli (i.e., gore, violence, mature language, harm to self or others), participants were instructed to complete the study in an isolated, distraction free space without others present to improve immersion. Participants were also instructed to set audio to a comfortable but audible level, to silence their phones, and to minimize or close other programs. To further standardize the experience and minimize distractions, the task was programmed to automatically run the browser in full screen mode and the task was programmed to stop if full screen mode was existed. Mobile devices, tablets, and screens with a resolution of less than 700 x 1250 pixels were eligible to participate.

Participants were instructed to watch and react to each of the four video clips as they naturally would. These instructions appeared as a reminder before each video played. Before each stimulus, participants were asked, “How intense are the negative emotions that you feel before starting this clip?” with responses captured on a linear sliding scale ranging in values from 0 (labelled “Not at all Intense”) to 100 (labelled “Extremely Intense”). Which side of the scale corresponded to which label was counter-balanced across participants. The default value of the slider thumb was set to 50 and participants were required to interact with the slider thumb before progressing. As participants moved the slider thumb, the value corresponding to its position waws visible and updated accordingly. Video order was randomized. After each video, participants were asked, “How intense were the negative emotions that you felt while watching this clip?”. Responses were captured using the same scale as previously described. However, participants were also reminded of the value that they had selected prior to starting the video (e.g., a participant who indicated a 77 on the scale before watching the video saw the message: “Before this video, you reported your negative emotions were a: **77**”).

All video clips were 120 seconds in length and pulled from relatively lesser-known independent horror films, including *Vicious Fun* (Particular Crowd, Black Fawn Films), *The Marshes* (28 Productions), *Head Count* (Godmother Industries), and *Superhost* (Superchill). Forty hypothesis-blind independent raters watched and rated these clips, as well as six others on a scale from 0 to 100 across metrics including arousal, valence, narrative coherence, and familiarity. These four clips were chosen due to their low average familiarity values (*range x̄ = 0.25 – 6.05),* indicating that few participants would have likely seen them before, and negative median valence scores (i.e., values below 50; *x̄Vicious Fun* = 15, *x̄The Marshes*  = 01.5, *x̄Head Count*  = 33.5, *x̄Superhost*  = 25.5). They were also chosen for having either relatively high or relatively low median arousal scores (*x̄Vicious Fun* = 77.5, *x̄The Marshes*  = 70, *x̄Head Count*  = 50.7, *x̄Superhost*  = 58.5) and minimal variance around those values. This allowed us to attempt to manipulate the affective experiences of participants to influence self-regulation behaviors.

After viewing and rating all videos, Study 3 participants were assigned to either an “experience” (*n* = 130) or “forecast” (*n* = 112) condition. Both conditions were given descriptions and examples of distraction and reappraisal that mirrored those used in Study 2 (though, ‘haunted house’ was replaced with ‘horror movie’). Whether participants received the description and examples of distraction or reappraisal first was counterbalanced across participants. However, following these descriptions participants were asked either, “Which strategy, if any, did you use to regulate your negative emotions while watching this clip?” (experiencer condition) or “Which strategy, if any, would you predict the average person would use to regulate their negative emotions while watching this clip?” (forecaster condition). Participants responded to this question by selecting either “Reappraisal”, “Distraction”, or “Neither”. The order in which these options appeared were randomized for each participant. Reminders of what the three options referred to appeared below these options. Reminders for participants in the experience condition read, “CHOOSE REAPPRAISAL if you changed how you thought about the video; CHOOSE DISTRACTION if you diverted your attention from the video; CHOOSE NEITHER if you did not implement either of the other two strategies”. Reminders for participants in the forecast condition read, “CHOOSE REAPPRAISAL if they should change how they thought about the video; CHOOSE DISTRACTION if they should divert their attention from the video; CHOOSE NEITHER if they should not implement either of the other two strategies”. If participants selected neither, they were not asked any further questions for that stimulus. If participants selected a non-neither option, they were asked how much the strategy they used reduced or the strategy they forecasted would reduce negative emotions while watching the stimulus on a 0 to 100 scale as previously described. Participants also either answered how effortful it was to use a strategy and how successfully they could use it, or how effortful that they predict it might be and how successfully they predict it could be used. These were each captured on 7-point Likert scales ranging from “Not at all” to “Extremely”. Participants responded to questions for each clip in the order that they clips were watched and a still from each clip was centrally-placed on the screen for each question to ensure participants understood which stimulus that they question was referring to. Following all questions, participants were asked to provide a brief example of both reappraisal and distraction in a free-response textbox to further assess task comprehension. Following the primary task, participants completed individual difference measures, including the Emotion Regulation Questionnaire, the Difficulties with Emotion Regulation Survey, and the Intolerance of Uncertainty Scale. Participants completed the study in 21.9 minutes on average (*median* = 20.5 minutes, *sd* = 56.0 minutes, *range* = 13.6 – 50.7 minutes). For additional information, see our OSF repository containing all task materials, including instructions and texts (*See* **Open Practices**).

**Analysis.** To explore whether affective intensity and condition influenced the strategies participants used or selected, we again specified mixed effect binary logistic regressions accounting for the random effect of participant and stimulus using the “lme4” package (Bates et al., 2015) in R (R Core Team, 2022) and followed an information theoretic approach via AIC comparison. Fixed effects models were iteratively built from and compared to our null model (ICC = 0.162), which did not contain any fixed effects. All data and scripts used to produce this analysis are publicly available at OSF (*See* **Open Practices**). The design and hypotheses of Study 3 were preregistered with AsPredicted (<https://aspredicted.org/n3ne3.pdf>).

**EXPERIMENT 3 RESULTS**

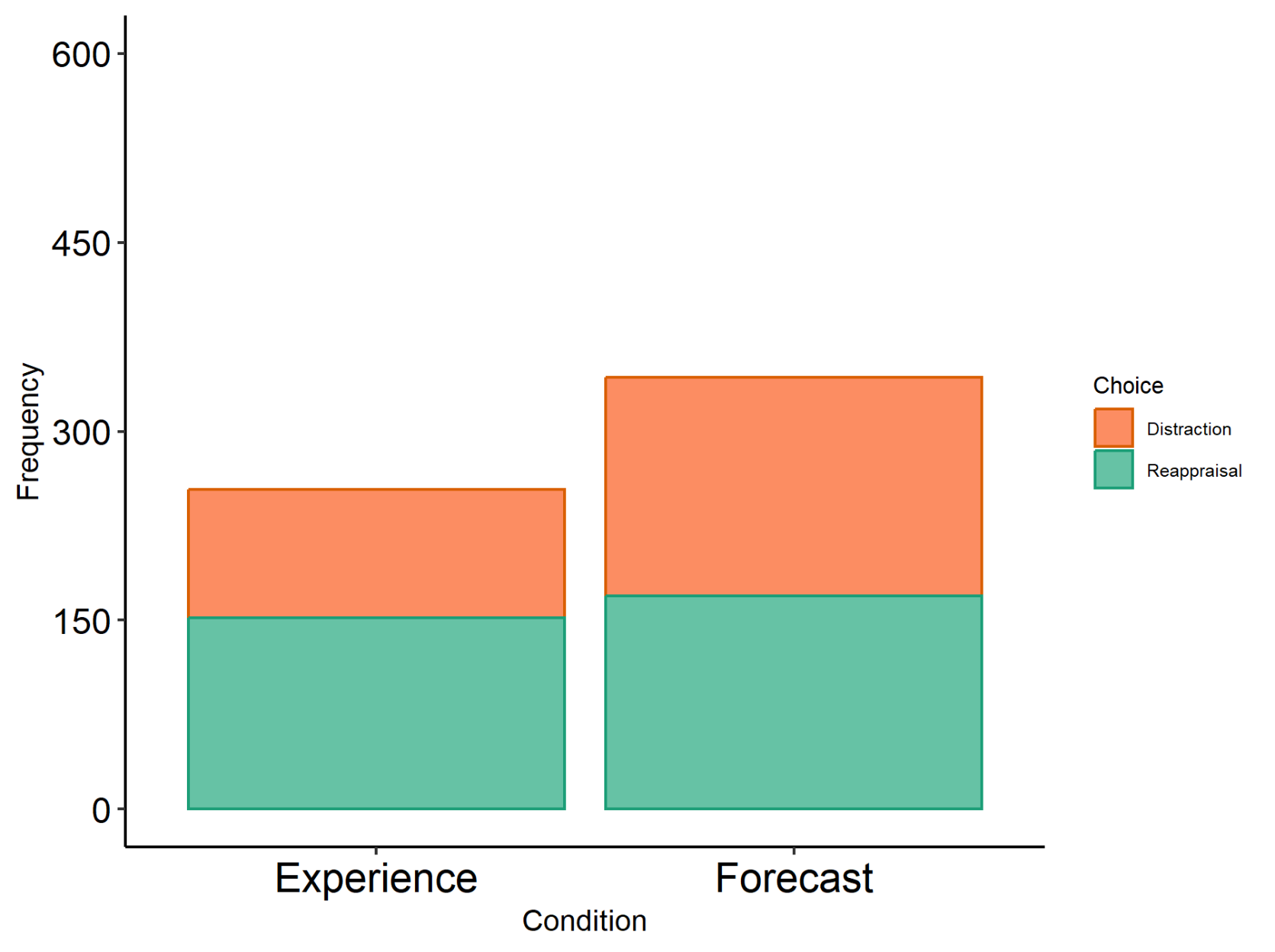
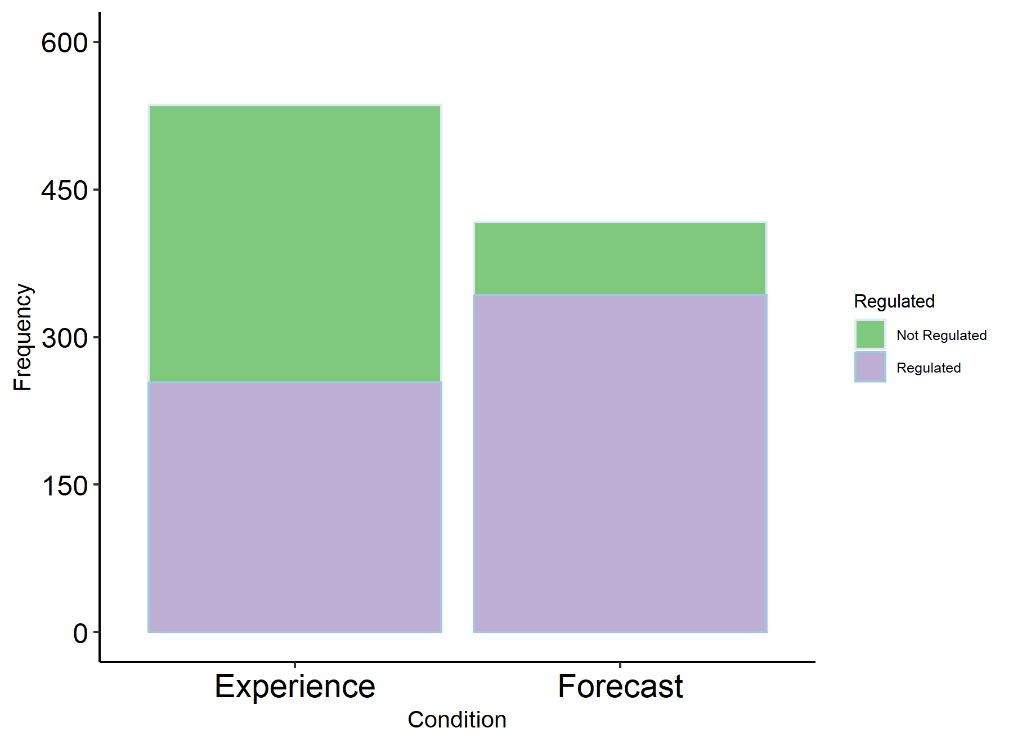
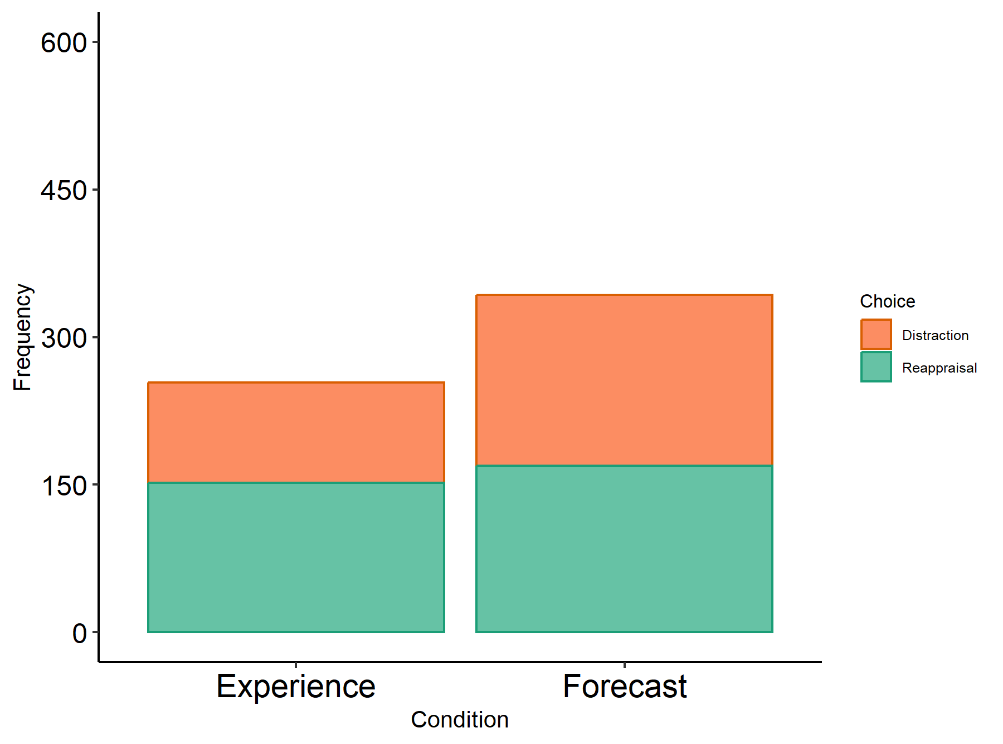
**Video stimuli elicited predictable emotional responses.** To ensure that our emotion manipulation was successful, negative affective intensity ratings following each low- and high-intensity stimuli were compared using Welch’s Two Sample t-test, which found that our manipulation was successful (*x̄ High* = 53.5*, x̄ Low*= 40.6, *95% CI* = [9.1, 16.8], *t*(948.1) = 6.61, *p* < 0.001). The average baseline intensity (i.e., intensity assessed prior to each video) was 27.4 pts (*median* = 19, *sd*  = 27.0) while the average post-exposure intensity was 47.1 pts (*median* = 50, *sd* = 30.9).

**Forecasters and experiencers reported no differences in emotional intensity.** Though participants were randomly assigned to conditions after having provided affective ratings, we wanted to ensure that any effects discovered between conditions could not be attributable to differences in experienced affective intensity. A Welch’s Two Sample t-test comparing affective intensity ratings between experiencers and forecasters both before (*x̄ Experience* = 26.2*, x̄ Forecast*= 28.8, *95% CI* = [-6.1, 0.8], *t*(908.3) = -1.50, *p* = 0.134) and after (*x̄ High* = 46.5*, x̄ Low*= 48.0, *95% CI* = [-5.4, 2.4], *t*(917.9) = -0.73, *p* = 0.460) stimulus exposure found no statistically significant differences in affective ratings. Forecasters and experiencers also did not differ across any relevant individual difference measures or demographics, including the reappraisal (*t*(198.6) = 0.14, *p* = 0.890) and suppression (*t*(222.57) = -1.94, *p* = 0.054) ERQ subscales, both Factor 1(*t*(235.0) = -0.94, *p* = 0.349) and Factor 2 (*t*(220.4) = -0.96, *p* = 0.336) of the IUS scale, the limited access to strategies subscale of the DERS (*t*(221.4) = 0.08, *p* = 0.938), age (*t*(220.0) = -0.14, *p* = 0.888), or horror enjoyment (*t*(225.9) = 0.12, *p* = 0.908).

**Intense emotional responses were more likely to be regulated.** To add greater ecological validity to a forced choice paradigm, we provided participants the option to indicate whether a stimulus should be regulated (i.e., reappraisal or distraction) or not regulated (i.e., neither). In line with hedonic motivations to self-regulate, we expect that videos of greater self-reported negative affective intensity should be more likely to be regulated; if that is not the case, then our design may generate noisy or counterintuitive results. To that end, we did find that videos in which participants selected neither distraction nor reappraisal were of a lower intensity (*x̄ Regulated* = 53.3*, x̄ Unregulated*= 36.7, *95% CI* = [12.7, 20.], *t*(735.6) = 8.30, *p* < 0.001) and a mixed effect binary logistic regression found that each standard deviation unit increase in affective intensity results in a 2.86 fold increase in the odds of self-regulation occurring (*95% CI* = [2.16, 3.78], *p* < 0.001).

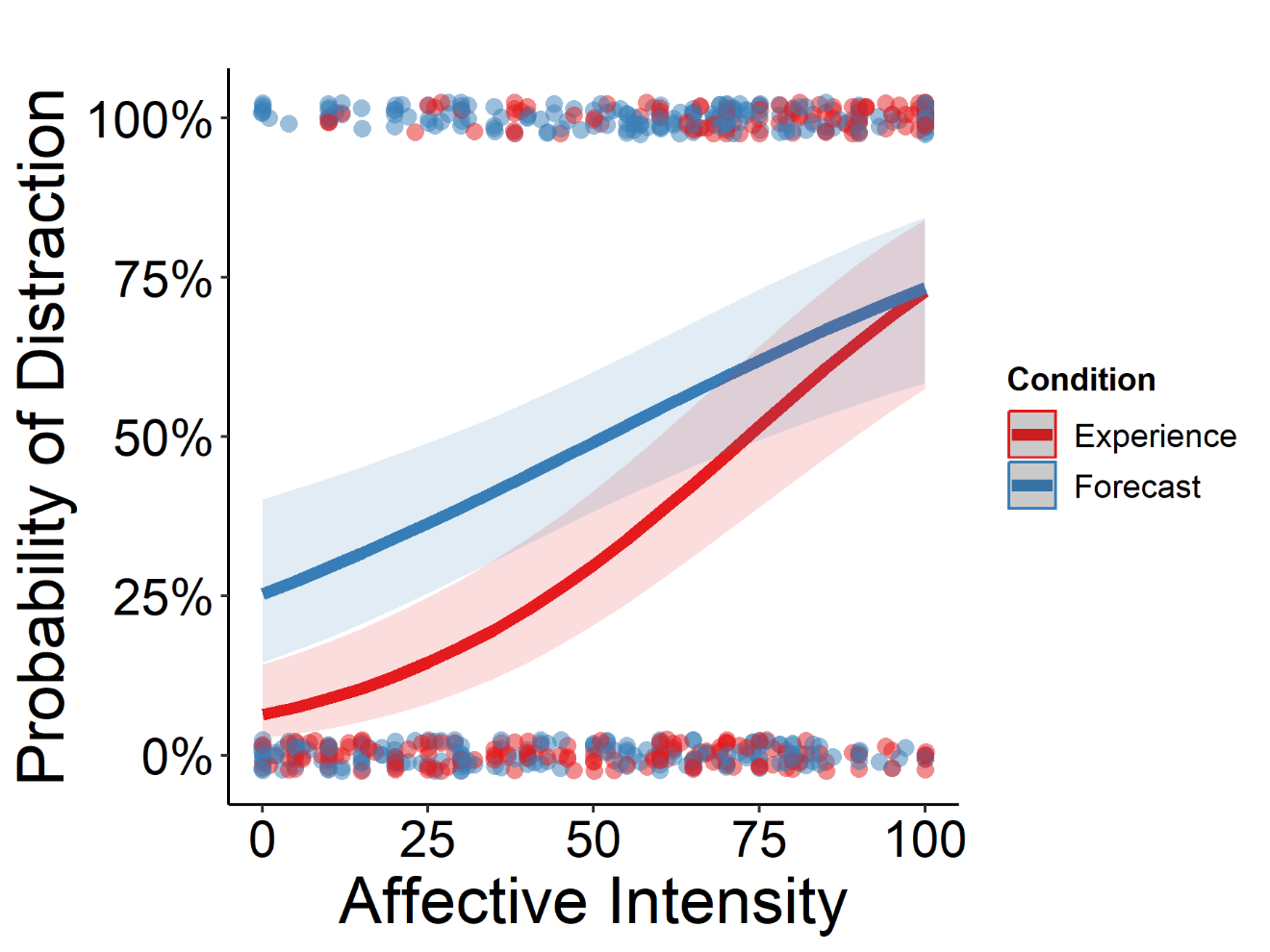
**Forecasters regulated more often than experiencers; Experiencers reappraised more often than forecasters.** Overall, reappraisal was used or forecasted in 33.1% of trials, distraction was used or forecasted in 28.5% of trials, and 36.7% of trials were left unregulated or without either option being forecasted. To assess whether there were differences in the distribution of strategy usage or forecast by condition, which could have implications for our primary analysis, we conducted chi-square test on strategy use/forecast and condition, which found differences in the frequencies with which forecasters and experiencers left trials unregulated (*χ2*(1, *N* = 953) = 120.35, *p* < 0.001) and used or forecasted a regulation strategy (*χ2*(1, *N* = 613)= 6.14, *p* = 0.013). More precisely, experiencers reported regulating (52.6%) and not regulating (47.4%) about equally while forecasters heavily favored regulating (82.2%) over not regulating (17.7%) across trials (**Fig. 7a**). Additionally, experiencers reported using reappraisal (59.8%) slightly more often than distraction (40.2%) while forecasters chose both distraction (49.2%) and reappraisal (50.7%) about evenly (**Fig. 7b**).

**Forecasters and experiencers differ in strategy selection probability at low, not high, intensities.** Following the analysis plan outlined within the Study 3 methods sections, we compared models via chi-square tests that included fixed effects for post-exposure affective intensity, baseline affective intensity, condition, and an interaction between post-exposure affective intensity and condition. These models were built iteratively, adding one effect at a time. Our best fitting model included all four terms (*χ2*(1) = 4.205, *p* = 0.040) and found the interaction was predictive of strategy usage/forecasting (*OR* = 0.56, *95% CI* = [0.36, 0.87], *p* = 0.010). Having tested the model that we hypothesized, we identified a selection of possible covariates (age, horror, enjoyment, ERQ subscales, IUS subscales, and the DERS limited access to strategies subscale) and iteratively added them to the model in order of most correlated with the outcome variable to pursue the model of best fit. Only the addition of the reappraisal ERQ subscale improved model fit above that of our hypothesized model (*χ2*(1)= 9.3587, *p* = 0.002). This model yielded a significant interaction term (*OR* = 0.61, *95% CI* = [0.39, 0.95], *p* = 0.029) and suggests that condition is moderating the relationship between affective intensity and strategy usage/forecasting. More specifically, at high intensities both forecasters and experiencers are similar in their likelihood of using or forecasting distraction. However, in congruence with forecasters predicting to use reappraisal less than experiencers used it in practice, at lower intensities, forecasters are less likely than experiencers to use reappraisal to regulate emotions (**Fig. 8**).



**Fig 7.** Chi-square tests highlighted that experiencers were more likely to not regulate a stimulus than forecasters (**A**), but that forecasters more evenly used both strategies when selecting a strategy (**B)**.

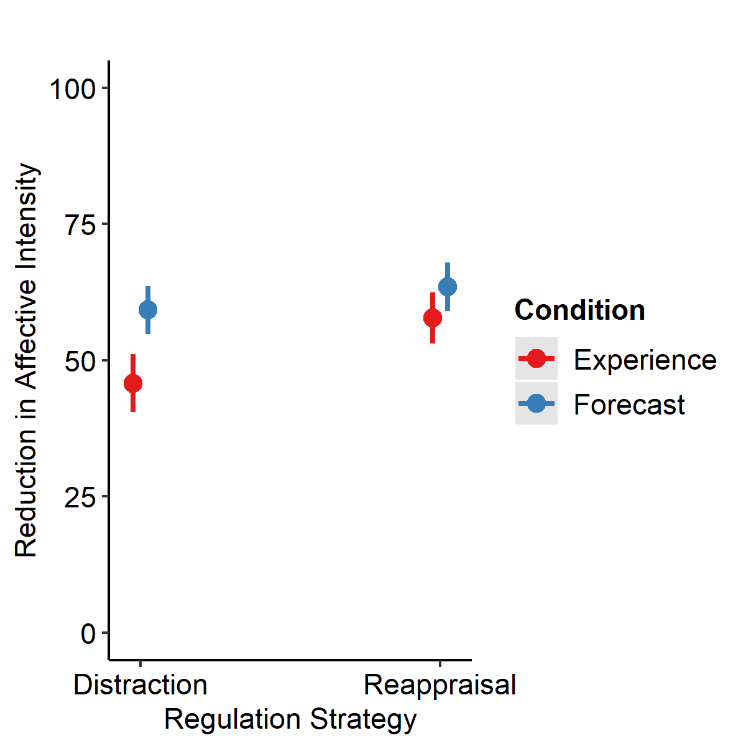
**A. B.**



**Fig 8.** Mixed effects binary logistic regression found that intensity predicts strategy usage and choice for both experiencers and forecasters, respectively. However, at lower intensities, forecasters are less likely to select reappraisal than experiencers were to use reappraisal.

**Forecasters and experiencers differ in regulation effort.** Participants in both conditions failed to generate significant differences in how much effort they exerted or predicted that they would exert to use reappraisal and distraction (*F*(1, 215) = 0.392, p = 0.532). However, the participants in the experience condition did differ from participants in the forecasting condition in how effortful regulation was across strategies, even when adjusting for strategy usage/forecast (*F*(1,215) = *3.820*, *p* = 0.045). Bonferroni-adjusted post-hoc contrasts determined that this difference is primarily driven by reappraisal; experiencers reported that reappraisal was more effortful than forecasted predicted it would be (*x̄ Experience* = 3.11*, x̄ Forecast*= 2.60, *95% CI* = [0.13, 0.88], *t*(296.5) = .66, *p* = 0.016).

**Forecasters and experiencers differ in regulation intensity reduction.** When examining how effective participants thought the strategies would be at reducing affective intensity, adjusting for the random effect of participant, we found a significant main effect of strategy (*x̄ Reappraisal* = 61.9*, x̄ Distraction*= 54.5, *95% CI* = [3.31, 11.48], *t*(582.12) = 3.55, *p* < 0.001) and condition (*x̄ Experience* = 55.0*, x̄ Forecast*= 61.1, *95% CI* = [-10.32, -1.85], *t*(494.6) = -2.83, *p* = 0.004), but more intriguingly, in an interaction model we found a significant interaction between the two variables (*F*(1,377) = 4.31, *p* = 0.038). A subsequent Bonferroni-adjusted contrast revealed a significant difference between experiencers and forecasters in how effectively distraction reduces affective intensity (*x̄ Experience* = 48.0*, x̄ Forecast*= 58.3, *95% CI* = [-16.62, -4.13], *t*(198.2) = -3.28, *p* = 0.001) **(Fig 9)**. This finding is congruent with findings from both Study 1 and Study 2 in which individuals seem to overestimate or mispredict the effectiveness of distraction, but not reappraisal, within these settings.



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**Fig 9.** Forecasters predicted that distraction would reduce negative affective intensity by a greater magnitude than experiencers reported that it actually had.

**GENERAL DISCUSSION**

Three experiments examined the association between emotional intensity and regulation strategy usage in response to dynamic, multimodal stimuli of varying intensity. Experiments 1 tasked untrained participants to recall emotional and spontaneous regulatory behaviors in a surprise recall task after exposure. Affective intensity did not predict strategy usage. Though distraction was used more often than reappraisal, distraction was less successful at regulating in response to high affective intensities in this context. Experiment 2 found an association between affective intensity and strategy usage among participants forecasting regulatory behavior in response to descriptions of regulated events that haunted house participants actually experienced. However, forecasters overpredicted how often distraction was used by experiencers. Study 3 measured differences between experiencers and forecasters regarding ER within the same study and found significant differences between forecasters and experiencers in how often distraction was used or forecasted and how effective the groups reported and predicted distraction to be. The present findings highlight challenges in translating ER theory to real-world application, as complex, dynamic situations in everyday life may strongly influence the efficacy and frequency with which regulatory strategies are used.

These results offer nuance to our understanding of affective intensity’s influence upon regulation strategy application and complements research on environmental affordances, or the extent to which features of a situation lend themselves to either distraction or reappraisal (Suri et al., 2018; Young & Suri, 2020). It is possible that individuals initially regulated their emotions with the strategy they believed had the greatest value (i.e., the most effective strategy with the lowest effort investment, typically distraction) based upon models of regulation they had previously generated from experiences with similar media, as outlined by computational decision-making frameworks of ER (Etkin et al., 2015). However, the unrelenting, attention-grabbing nature of challenging high-intensity situations may grant few affordances by which to distract oneself. On the other hand, reappraisals may appear to be less valuable initially, resulting in lower usage, but could be more effective when used due to a relatively greater volume of environmental affordances (i.e., things to repurpose or reconstrue (Uusberg et al., 2019); e.g., actors, props, goals, etc.). This explanation mirrors the strategy-selection relationship hypothesized by Etkin and colleagues (2015) who posited that a more explicit, higher-cost, model-based approach to ER may be applied more effectively when implicit, or model-free, ER tendencies were not arriving at their desired goal via prediction error adjustment alone. Distilling events down to a text-based representation or video, as had been done in Experiments 2 and 3, may provide greater distraction affordances and rebalance the likelihood of individuals choosing reappraisal or distraction in situations when it would canonically make sense to do so.

How a person passively assesses the emotionally-relevant features of an environment may guide regulatory behaviors, but so too might the action affordances – opportunities to move, modify, or engage with the environment – that a situation presents. Action-oriented perspectives on ER emphasize that emotion regulation behaviors are a product of active processes, such as forward modelling (Bramson et al., 2023). Common associations between self-regulation, emotion, and activity in the sensorimotor and pre-motor systems illustrate how entangled action is with regulation (Bramson et al., 2018; Mobbs et al., 2007; Saarimäki et al., 2016). The degree to which one is free to physically navigate their space or interact with evocative stimuli may have important implications on situational appraisals, thus informing which strategies participants predict may or may not work (Ridderinkhof, 2017). **Figure 5** highlights one example of this within the haunted house as a participant was able to mitigate a negative emotional reaction by physically circumventing the stimulus. Stimulus-response paradigms which situate participants in stationary positions or lack contexts enveloping the emotionally-evocative stimulus may not be well suited for modeling the effect that these factors have upon typical or daily ER strategy usage.

Cold-hot empathy gap research, which measures forecasting differences between how people think they will feel hypothetically and how they feel in practice (Loewenstein, 1996), are relevant as well. Individuals in “cold states” consistently underpredict the challenges associated with meeting affectively-relevant goals during “hot states” (Sayette et al., 2008; Van Boven & Loewenstein, 2003). Such a pattern mirrors the differences observed between Experiments 1 and 2, wherein decontextualizing events (i.e., shifting from a hot state to a cold state) yielded a predictable pattern in strategy choice not observed during hot state ER usage. They also mirror differences in strategy usage and success. Such patterns highlight that emotion self-regulation is a complex, multi-faceted construct and different proportions of its variability may be better captured by different approaches (Friedman & Gustavson, 2022). Though this study is the first to our knowledge that has utilized a high-intensity, dynamic, feature-rich paradigm to demonstrate this in the domain of ER, similar approaches have demonstrated similar discrepancies in the moral domain (FeldmanHall et al., 2012). Computerized lab tasks have been theorized to assess regulatory performance in optimal conditions (Wennerhold & Friese, 2020) and this may contrast study designs such as Study 1, which may be closer to performance in typical conditions, and Studies 2 and 3, which may be somewhere in-between (Friedman & Gustavson, 2022). Though participants may be capable of performing self-regulation at high levels in optimal conditions, they may not feel motivated to do so in typical conditions (Grund & Carstens, 2019).

**Limitations.** There are several limitations in our experimental approach that demand attention. First, our Study 1 aims were relatively narrow in comparison to the vast ER behavior variability captured by this dataset and resulted in excluding many observations that did not meet our inclusion criteria (*See* **Fig. 2**). More analyses would be required to fully explore this space. Additionally, though many features of our design mirror regulation of emotions in everyday life, haunted houses may have limited generalizability to other high-intensity settings that we commonly experience (Clasen, 2019; Tashjian et al., 2022) and the purpose haunted houses commonly serve (i.e., entertainment) could also circumscribe their generalizability to graver situations. The use of a haunted house as our setting also may have resulted in self-selection biases within our sample. It must be noted that we did not directly manipulate emotional intensity within this design and the self-report data that we did collect was captured post-exposure, not during exposure. Additionally, we unfortunately did not have measures of startle sensitivity or fear-enjoyment from Experiment 2 participants. Thus, we cannot compare our Study 1 and Study 2 samples along these metrics, which explain some of the differences in regulatory behaviors between samples. Experiments 2 and 3 were also conducted entirely online. Though means of standardizing the experience were attempted, we have less control than possible in an in-person context and cannot guarantee that participants were fully focused on study tasks.

A small but important contingent of the existing ER literature has highlighted intra-strategy heterogeneity in the regulation techniques that we had examined (Uusberg et al., 2019; Webb et al., 2012). For example, Webb and colleagues identified at least three distinct ER approaches that could be categorized as reappraisal (i.e., reappraising the stimulus, reappraising the emotional response, reappraisal via perspective-taking) and these reappraisal subtypes demonstrate varying effectiveness in modulating emotion depending upon situational and personal factors. It may be the case that specific types of reappraisal and distraction demonstrate an intensity-selection relationship that better mirrors the canonical association observed in lab studies, but testing such a question with sufficient power would necessitate far more observations than are available in this dataset.

Due to resource constraints, participants also necessarily experienced the haunted house in groups. Although participants were instructed to not discuss their experiences and post-hoc analyses failed to find any association between group membership and strategy usage, the group context in which the experience occurred could have influenced behavior and cognitive perceptions in unknown ways. The presence and strength of friendship among group members was also assessed and was not predictive of regulation. Lastly, some events, particularly in the pilot and multiverse analyses, had to be captured at a one-week delay due to collaboration with a memory study, which calls into question the reliability of some responses. Future research should limit the delay between experience and report as much as possible without interfering with emotional events. Future research might also take interest in the order of strategies reported in regulation-challenging situations; that is to say, does the likelihood of using a reappraisal strategy for a given person in this context increase with each failed attempt to engage distraction? Although we did not have sufficient observations to examine the data with such granularity, examining associations between the intensity of specific emotions and regulation usage might also be of particular interest to future researchers in this space (Young & Suri, 2020). Despite its limitations, this dataset and approach may be of interest to those exploring spontaneous regulation tendencies from untrained participants in response to both positive and negative events.

Taken together, the present experiments represent what we believe to be the first attempt to extend the association observed between affective intensity and different ER behaviors to dynamic, multimodal and high-intensity settings using untrained participants. This approach offers an alternative means of exploring ER usage while pursuing greater ecologically-valid in study design. Our data and results may be of particular interest to other emotion, self-regulation, and cognitive control researchers interested in quasi-naturalistic design. In failing to replicate lab results with Studies 1 but finding a modest association in Studies 2 and 3, we may offer greater support for the importance of context in determining the fit of ER strategies, though the limitations inherent to this study leave room for other possibilities. Ultimately, though, this study highlights the importance of extending what we know about regulation in relatively mundane, controlled situations to those crowded, loud, and perhaps uncontrollable contexts in which emotion regulation success could be of dire consequence.

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The authors declared that they had no conflicts of interest with respect to their authorship or the publication of this article.

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**Author’s Note**:

Earlier versions of this manuscript have been available on PsyArxiv (<https://osf.io/preprints/psyarxiv/23wtz>) since November 28th, 2022. Posters containing some of this data and analyses have been presented at the 2021 and 2022 Society for Personality & Social Psychology Conference, as well as the 2022 Society for Affective Science Conference. Invited presentations using some of these data and analyses occurred in 2020 at the Temple University Motivational Behavior Seminar Series and in 2023 to the March Lab at Florida State University. Publications using the same dataset include Stasiak et al., 2023.

**Transparency and Openness Practices:**

The preregistration for Experiments 1, 2, and 3 can be found at [https://aspredicted.org/DP1\_453,](https://aspredicted.org/DP1_453) <https://aspredicted.org/XXH>[\_W1V,](https://aspredicted.org/XXH_W1V) and <https://aspredicted>.org/n3ne3.pdf, respectively. Deidentified data, analysis and data cleaning scripts, questionnaires, and data collection files to replicate the findings have been made publicly available via OSF and can be accessed at <https://osf.io/j5sku/?view_only=89d87669e7674096819c439ca109c483>. All products generated from this data, including powerpoints and posters highlighted under the Author’s Note have been made available within this repository as well.

**CRediT:**

**William J. Mitchell:** Conceptualization, Methodology, Software, Validation, Formal analysis, Investigation, Data Curation, Writing - Original Draft, Writing - Review & Editing, Visualization, Project administration. **Joanne Stasiak:** Conceptualization, Methodology, Investigation, Resources, Data Curation, Writing - Review & Editing, Project administration. **Steven Martinez:** Conceptualization, Investigation, Data Curation, Project administration. **Katelyn Cliver:** Conceptualization, Investigation, Data Curation, Project administration. **David Gregory:** Conceptualization, Investigation, Data Curation, Project administration. **Samantha Reisman:** Conceptualization, Investigation, Data Curation, Project administration. **Helen Schmidt:** Software, Validation, Formal analysis, Writing - Review & Editing. **Vishnu P. Murty:** Conceptualization, Resources, Writing - Review & Editing, Supervision, Project administration, Funding acquisition. **Chelsea Helion:** Conceptualization, Methodology, Resources, Writing - Review & Editing, Visualization, Supervision, Project administration.

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**Supplementary Materials**

**PILOT STUDY METHODS**

A pilot study tested whether the emotional intensity of negatively-valenced events was associated with the likelihood of using a low-effort or high-effort regulatory strategy in a quasi-naturalistic setting with an untrained sample. To assess emotional intensity, participants self-reported the emotional intensity of events from the haunted house one week after exposure during a surprise recall task. Participants also noted whether they wanted to reduce the intensity of these emotions and, if so, how they attempted to do so in their own words. We report how we determined our sample size, all data exclusions, all manipulations, and all measures in the study.

**PARTICIPANTS:** In October 2019, 54 participants (*x̄ age* = 24.22 yrs, range = 18 – 34 yrs, *sdage* = 3.97 yrs, 26 female, 1 non-binary, 18.51% Hispanic) were recruited from a large northeastern city via flyers for an IRB-approved fear and memory study. Participants were predominantly well-educated (*x̄ Years of Education* = 15.3 yrs, *sd Years of Education* = 2.4 yrs), with 92.6% reporting having completed some college (35.2%), a 4-year degree (29.6%), some post-graduate studies (16.7%), or a post-graduate degree (11.1%). Socioeconomic status was more normally distributed, with 12.7% of respondents reporting making less than $15,000 per year, 16.3% reporting between $15,001 and $25,000, 09.1% reporting $25,001 to $35,000, 23.6% reporting between $35,001 and $50,000, 12.7% reporting between $50,001 and $75,000, 16.4% reporting between $75,001 and $100,000, 09.1% reporting between $100,001 and $150,000, and no one reporting greater than $150,000. The racial identity of participants was not assessed.

A priori power analyses using the WebPower (Zhang, Z., & Mai, Y., 2019) In R 3.6.1 (R Core Team, 2022) determined 18 participants would sufficiently power our main effect using the smallest effect size reported by Sheppes et al.’s 2011 examination of emotional intensity and regulatory choice (ηp2 = 0.43). Participants were excluded for previously visiting the haunted house (*n* = 1), not completing the study (*n* = 1), identifying English as their second-language (*n* = 2), or not following instructions (*n* = 3). Participants received $70.00 in Visa debit cards for participating.

**MATERIALS AND PROCEDURE:** Our design consisted of an exposure session and a follow-up. Following consent, participants completed computerized questionnaires and were fitted with physiological monitors which are beyond the purview of this study (*See* Stasiak et al., 2023). Participants were then escorted by two research assistants to the remotely-located haunted house which was the same as used within Experiment 1.

**Session 1, Haunted House.** Participants navigated the haunted house in twelve groups (*x̄* *size* = 4.50 participants; sd*size*= 0.79 participants) for approximately 55.40 minutes (*sd* = 5.05 minutes) and were provided with minimal instructions to promote ecological validity (participants were to walk through the haunted house in a single file line and avoid sharing thoughts, reactions, and experiences with other participants). However, they were encouraged to act and react as naturally as possible. Each participant was randomly assigned to lead the group through one section. The accompanying research assistant led the group through any sections without a participant-leader. Following exposure, participants were scheduled for an individual follow-up and were instructed to not discuss their experience with anyone.

**Session 2, Laboratory follow-up session.** At follow-up (*time since exposure:* x̄ *delay* = 5.98 days; sd*delay*= 0.79 days), participants completed a surprise free-recall memory task and questionnaires. Notably, participants identified, described, and chronologically ordered ten (10) discrete events from within the haunted house. For each event, participants identified which of 13 emotion categories they had felt, the intensity of each of those emotions, the effort participants exerted trying to regulate each of those emotions, and to describe how they attempted to regulate them (if at all). It must be noted that while participants were able to endorse multiple emotions of differing intensities for any one event, participants were only asked about their regulatory behavior once per event and not whether that regulatory behavior was directed towards any specific emotions endorsed during that event. Thus, because we have greater granularity of emotion than we do the regulatory responses to those emotions, the association between these variables was assessed through multiple approaches to account for this discrepancy. To avoid confusion, we refer to this approach as capturing regulatory behaviors at the “event-level”. During Experiment 1, we ask participants about their regulatory responses to each emotion they endorsed, which we refer to as capturing regulatory responses at the “emotion-level”. For any one event, one or more emotions may have been endorsed by the participant.

Participants were not trained in emotion regulation strategies, nor were they primed to consider their emotion regulation strategies prior to these questionnaires. Though EMA approaches could also be applied in such a setting, the training and/or consideration of experience required would violate the immersive, naturalistic experience we aimed to model. Emotion categories were adapted from the Positive and Negative Affect Schedule (PANAS) (Watson et al., 1988). Some noted additions relevant to a typical haunted house experience included “tense”, and “disgusted”. Ten of the 13 options were negatively-valenced emotions (i.e., *Disgusted/Grossed Out, Fearful/Afraid, Hostile/Aggressive, Irritable/Annoyed, Nervous/Jittery, Overwhelmed, Panicked, Shocked/Surprised, Tense, Upset/Distressed*). Applying Cronbac’'s alpha to the emotional intensity value of negative emotions yielded a value of α = 0.91 (*95% CI* = [0.89, 0.92]), suggesting excellent internal consistency. Both emotional intensity and regulation effort were captured on a 7-point Likert scale, with 1 representing “Not at all” and 7 representing “A great deal”. Regulation strategies were captured for each event using free-response to the prompt: *If you did attempt to change or regulate your emotions, how did you do so?* Participants were subsequently debriefed and paid for their participation.

**Analysis.** To explore our primary question, the effect of emotional intensity upon regulatory strategy usage, we specified multilevel binary logistic regressions accounting for the random effect of participants using the “lme4” package (Bates et al., 2015) in R (R Core Team, 2022). We followed an information theoretic approach via AIC comparison. Our model comparisons included a: A) a null model without fixed effects, B) a model containing only z-scored emotional intensity as a fixed effect, C) a model containing only person-centered, z-scored emotional intensity and person-mean, z-scored emotional intensity as fixed effects, and D) a model including person-centered and person-mean z-scored emotional intensity with covariates. All data and scripts used to produce this analysis are publicly available at OSF (*See* Open Practices). This study was not pre-registered.

**PILOT STUDY RESULTS**

As we primarily aimed to determine how analogous associations observed in ER strategy choice paradigms were to ER usage in quasi-naturalistic contexts, observations that did not contain negative emotions regulated by either distraction or reappraisal were beyond the purview of our study. Our full dataset consisted of 469 unique events from 47 participants with an average of 2.43 emotions (SD) endorsed per event. Of the 1138 endorsed emotions, 603 (52.99%) were classified as being negatively valenced. Of the 603 negative emotions endorsed, there were 166 observations in which a negative emotion was downregulated by either distraction or reappraisal. These 166 observations came from 78 unique events reported by 32 participants. Of the 78 unique events, 57 events (74.36%), or 130 observations (78.31%) reported using distraction to regulate their emotions. The average emotional intensity of observations was 5.55 (range: 1 – 7, Likert scale).

**Intensity and state anxiety predicts regulatory extent.** Pro-hedonic trends in emotion regulation (i.e., minimizing aversive experiences) suggest that a positive linear relationship should exist between negative affective states and efforts to regulate those experiences. If we did not find associations between how intense an emotional experience was or how anxiety-provoking an experience was and how much effort someone devoted to regulating that event – regardless of *how* they regulate it – it may call into question whether our paradigm is well suited to elicit emotion regulation behaviors. For each event, we asked participants how intense the emotions that they endorse were as well as the extent to which they attempted to regulate that event (i.e., “To what extent did you attempt to change or regulate how you felt during this event?”), which may be comparable to effort. We also assessed state anxiety prior to entering the haunted house. To assess an association between event intensity and event regulation extent, we ran multilevel linear models specifying regulatory extent as our criterion variable, participant as a random intercept, and building fixed effects from a null model (ICC = 0.35). Our best performing model as determined by AIC comparison found person-centered emotional intensity to be a significant predictor of regulation extent (*β* = 0.31, *95% CI* = [0.17, 0.46], *p* < 0.001). To assess an association between state anxiety and regulatory extent, we built a simple linear model regressing each participant’s average self-reported regulatory extent across events onto State STAI scores and again found a significant positive association (*β* = 0.34, *95% CI* = [0.054, 0.626], *p* = 0.021). These results suggest our paradigm elicited regulatory behaviors from participants that follow a logical, predictable pattern.

**Intensity did not predict regulatory strategy usage.** Using multilevel binary logistic regression models specifying regulation strategy usage as our criterion variable, participant as a random intercept, and building fixed effects from a null model (ICC = 0.70), we failed to find any model that performed better than our null model by AIC comparison. Our model using only emotional intensity to predict regulation strategy usage trended significant in the model comparison (p = 0.063). However, even if traditional statistical thresholds were loosened and the model was deemed superior to our null, within that model, we did not find that emotional intensity predicted strategy usage in that model (OR = 1.83, 95% CI = [0.65, 3.2], p = 0.079).

Because of the previously noted discrepancy in granularity of capturing emotion and regulation, analyzing the relationship between individual emotions and event-level regulatory strategies may be missing stronger relationships between emotions and regulation that exist when analyzing data at the event-level only. As such, additional analyses were conducted using the average emotional intensity of each event to predict strategy choice as well as the sum of emotional intensity for each event to predict strategy usage. In both cases, multilevel binary logistic regressions failed to perform better than the null model (ICC = 0.28; Emotion Sum: p = 0.130; Emotion Average: p = 0.430) and none of the affective variables predicted strategy usage, reinforcing the results of our primary analysis. Thus, regardless of whether emotions are considered individually or concurrently, we do not find evidence to support an association between affective intensity and ER strategy usage in this context.

**STUDY 1 MULTIVERSE METHODS & RESULTS**

**Multiverse approach also failed to explain strategy usage.** We expanded the scope of our primary analyses and conducted additional exploratory analyses to determine whether a stronger association between strategy choice and affective intensity could be found using different inclusion criteria, comparing engagement strategies (i.e., reappraisal) to disengagement strategies (i.e., suppression, distraction) as defined in the broader literature (e.g., Dixon-Gordon et al., 2015). Expanding our groups yielded a subset of 360 observations in which a negative emotion was downregulated by either disengagement or engagement strategy. These observations were reported by 89 participants. Of the total observations, 237 (65.80%) reported using distraction or suppression to regulate their emotions. The average emotional intensity of observations was 2.40. However, our best performing non-null model, including only intensity as a fixed effect (*p* = 0.32 when compared to null), again did not find an association between emotional intensity and strategy usage (*OR* = 1.18, *95% CI* = [0.85, 1.63], *p* = 0.32). Again, though not significant, this statistic suggests that every one standard deviation unit increase in emotional intensity increases the odds of choosing a disengagement strategy by approximately 18%.

Because it could be argued that a haunted house setting could elicit greater contra-hedonic regulation activity (i.e., downregulating positive emotion), we then constructed a series of additional models which were beyond the purview of our initial aims and hypotheses to determine whether any statistically significant relationship could be observed between affective intensity and regulation strategy usage in this context. We iteratively modified predictors and outcome variables across 14 additional models, including covariate models adjusting for sex, cognitive load, and ERQ reappraisal subscale scores, which despite failing to demonstrate significance in this dataset often predict regulation choice in lab settings. We found only a single model which surpassed nominal statistical thresholds of significance in model fit (*ICC* = 0.37; *p* = 0.04 when compared to null), but which did not maintain significance after adjusting to maintain a family-wise error rate (*p* = 0.32). This model included a random intercept for participant and a single predictor, affective intensity of positive and negative emotions, regressed upon distraction versus reappraisal strategy usage with data from unique events reported both immediately after and one-week after exposure (*OR* = 1.42, *95% CI* = [1.03, 1.98], *p* = 0.04). The model composition, comparison and results of all of these models can be found in **Table 1**.

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| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| **Model Details** | | | | | **Participants** (n) | **Observations** | **Null ICC** | **Model Comparison** (*x2*) | **Model Significance** | **Bonferroni Adjustment** | **Affective Intensity (z)** | | |
| *Outcome* | | *Emotions Included* | *Data Collection Time* | *Covariates* † | *Odds Ratio* | *Lower Bound ^* | *Upper Bound ^* |
| Disengage. V. Engage. | | Positive & Negative | Immediate & Delayed | Included | 90 | 397 | 0.34 | - | - | - | 1.27 | 0.93 | 1.73 |
| Not Included | 2.64 | 0.10 | 0.80 | 1.28 | 0.95 | 1.71 |
| Disengage. V. Engage. | | Negative | Immediate & Delayed | Included | 89 | 360 | 0.36 | - | - | - | 1.17 | 0.83 | 1.64 |
| Not Included | 0.98 | 0.32 | 1.00 | 1.18 | 0.85 | 1.63 |
| Distract. V. Reappraisal | | Positive & Negative | Immediate & Delayed | Included | 78 | 328 | 0.37 | - | - | - | 1.45 | 1.03 | 2.05 |
| Not Included | 4.45 | 0.04 \* | 0.32 | 1.42 | 1.03 | 1.98 |
| Distract. V. Reappraisal | | Negative | Immediate & Delayed | Included | 77 | 298 | 0.40 | - | - | - | 1.38 | 0.95 | 1.99 |
| Not Included | 2.70 | 0.10 | 0.80 | 1.36 | 0.95 | 1.95 |
| Disengage. V. Engage. | | Positive & Negative | Immediate | Included | 79 | 213 | 0.42 | - | - | - | 1.02 | 0.65 | 1.58 |
| Not Included | 0.09 | 0.76 | 1.00 | 1.07 | 0.71 | 1.61 |
| Disengage. V. Engage. | | Negative | Immediate | Included | 77 | 194 | 0.39 | - | - | - | 1.05 | 0.65 | 1.68 |
| Not Included | 0.14 | 0.71 | 1.00 | 1.09 | 0.70 | 1.68 |
| Distract. V. Reappraisal | | Positive & Negative | Immediate | Included | 64 | 171 | 0.45 | - | - | - | 1.28 | 0.78 | 2.11 |
| Not Included | 0.98 | 0.32 | 1.00 | 1.26 | 0.79 | 2.01 |
| Distract. V. Reappraisal | | Negative | Immediate | Included | 63 | 155 | 4.01 | - | - | - | 1.33 | 0.77 | 2.28 |
| Not Included | 1.17 | 0.28 | 1.00 | 1.30 | 0.81 | 2.09 |
| *\* = p < 0.05* | | |  | † *Covariates included Sex, Cognitive Load, and ERQ reappraisal subscale* | | | |  |  | *^ Bounds represent 95% confidence intervals* | | | |
|  | **Table 1.** *Results of Study 1 Exploratory Multiverse Models –* Hierarchical binary logistic regression models were constructed to explore how predictive affective intensity is of strategy usage and compared against null models allowing each participants’ intercept to vary randomly. The outcome variable used, data inclusion criteria, and variables included are listed under the first four columns. Data size and model comparison results are listed under the subsequent five columns. Model comparisons against the null for not conducted for covariate models because comparing models with such a disproportionate number of terms can be misleading and is not statistically recommended. The odds ratio of affective intensity within each model is listed in the latter three columns. Only a single model performed better than its null and found affective intensity to predict regulation strategy usage, but did not survive family-wise adjustment. | | | | | | | | | | | | |

1. The haunted house has a limited seasonal run time, and we cannot experimentally manipulate the intensity of the events in the haunted house as it is run by a private company. [↑](#footnote-ref-2)