

Simultaneous Treatment of Random and Systematic Errors in the Historical Radiosonde Temperature Archive

Joshua M. Browning¹ and Amanda S. Hering¹

February 20, 2015

Abstract

The historical radiosonde temperature archive, and indeed any large and lengthy observational dataset, must be quality controlled before it can be used properly. Most research on quality control for such data focuses on the identification and removal of either systematic errors (homogenization) or random errors without considering an optimal process for treatment of both. Additionally, little has been done to evaluate homogenization methods applied to sub-daily data, and no research exists on using robust estimators in homogenization procedures. In this paper, we simulate realistic radiosonde temperature data and contaminate it with both systematic and random errors. We then evaluate (1) the performance of several homogenization algorithms and (2) the sequence in which the random and systematic errors are identified and corrected. In our simulations we find that the robust Standard Normal Homogeneity Test (SNHT) that we introduce performs better than the traditional SNHT, and it is better than several other modern alternatives. Moreover, we find that systematic errors present in the data lead to poorer performance of random error removal algorithms, but the presence of random errors in the data is not as detrimental to homogenization algorithms.

Some keywords: Change Point Detection; Homogenization; Outlier Detection; Radiosonde Temperature Data

Short title: Simultaneous Random and Systematic Error Detection

¹Department of Applied Mathematics and Statistics, Colorado School of Mines, Golden, CO 80401, USA.
303.384.2462,
E-mail: {jbrownin, ahering}@mines.edu

1 Introduction

Any large dataset whose observations reach far back in time may require treatment for both systematic and random errors. Datasets such as the International Surface Temperature Initiative (ISTI) global land surface databank [21] with over 32,000 stations, and the Integrated Global Radiosonde Archive (IGRA) housed at the National Climatic Data Center (NCDC) [5] are examples of such large datasets. Systematic errors can occur when the station location changes; the area surrounding the station becomes urbanized; or the instrumentation is changed. Random errors can occur due to faulty data transmission; sporadic instrumentation problems; keystroke entries; or errors in data management. To illustrate, Figure 1 plots the temperature recorded by radiosondes at the 50 mb pressure level at Station 70219 (Bethel, Alaska, USA); this time series appears to have random and possibly systematic errors. It is important to treat both sources of errors in large historical datasets as robustly and automatically as possible. In most published research, methods for handling systematic and random errors are treated separately, and opinions among climate and weather scientists differ in terms of which type of error should be handled first. The purpose of this study is to shed light on the order in which systematic and random error methods should be applied to such large datasets when considering both sources of error simultaneously. In addition, robust estimators in homogenization algorithms when random errors are present have not yet been considered, so these are proposed and investigated as well.

In this paper, we focus on the Upper Air Database (UADB) housed at the National Center for Atmospheric Research (NCAR). This archive differs from the IGRA archive in that it contains some different stations, and many of the records are older. Since the radiosonde data are the only measured values of the upper atmosphere, it is an important resource for studies in climate change [6, 7] and for use as an input to global reanalysis datasets [13, 14]. Currently over 2,000 station locations exist, and atmospheric variables are collected at standard pressure levels as the radiosonde rises through the atmosphere. In large datasets such as these, the error detection methods must be automated since the archives are so large that visual inspections of every station are not feasible.

Many methods have been developed to homogenize radiosonde data, but most are not tested

on simulated data with known contamination errors [7, 9, 17, 18, 23]. However, a study was recently conducted by the European Cooperation in Science and Technology to compare many different homogenization methods [23]. A single large, realistic dataset with known change points was simulated, and then researchers were asked to test their homogenization algorithm on the dataset. As the researchers did not have knowledge of the true change point locations, this experiment provided a way to compare the performance of these methods.

Most homogenization techniques are designed for monthly or annual time series, but radiosonde observations occur, on average, twice daily. Some of these techniques rely on optimizing an objective function over all possible change point configurations [16, 19, 20, 22], and many of these approaches are too computationally expensive for daily data. Additionally, some methods may only locate proposed change points and may not correct for the difference in means, which is a necessary homogenization step. In this paper, we compare the Standard Normal Homogeneity Test (SNHT) [1], the PELT algorithm [16], and binary segmentation [22]. We also propose a robust version of the SNHT.

Automated random error detection methods for radiosonde temperature data have not been investigated as thoroughly [5, 17]. Models such as [12] for the entire vertical column using pairs of locations could be adapted for random error detection, but recently in [2], several random error detection methods for a given location and pressure level are proposed and evaluated via simulated datasets. The authors find that the optimal error detection algorithm requires two steps: first scanning for observations that are too many standard deviations from the global mean and secondly scanning for observations that are too many standard deviations from their local mean. Robust estimators of mean and standard deviation are used in both cases to mitigate the influence of errors, and a robust, asymmetric estimate of standard deviation is introduced to account for skewness in temperature data.

However, to our knowledge, no research has been done to date describing which type of error should be handled first when a dataset contains both types of errors. We do a simulation study in which data is contaminated with both known random errors and with known change points so that we can evaluate the performance of our robust SNHT and the sequence in which

different quality control algorithms are applied. We henceforth refer to the choice of performing random error detection or systematic error detection first as “the sequence of the quality control method” or simply “the sequence.” In Section 2, we discuss the details of our data simulation and contamination. Section 3 evaluates the performance of the homogenization algorithms we use, and Section 4 gives the results from the sequencing study. In Section 5, we present a case study of this method applied to a real dataset, and some conclusions are offered in Section 6.

2 Simulation Method

Observational data cannot be used to evaluate the performance of quality control (QC) and homogenization methods directly since we cannot know exactly where true change points and errors occur. Therefore, a rigorous simulation study is developed in order to accurately compare methods and their sequence. Evaluation of methodology via simulation is commonplace in the statistics literature, and our approach bases the simulation on actual data. In order for this simulation study to validate methods for radiosonde data, it is crucial that we simulate data that is similar in structure to true radiosonde data.

2.1 Modeling Radiosonde Data

In order to capture seasonal and hourly trends, we fit a Generalized Additive Model (GAM) to the radiosonde temperature data. GAMs are flexible, non-parametric models that allow the response variable to be a linear combination of smoothed functions of the input variables [10]. In our case, we model temperature (for a fixed location and pressure level) as a function of hour of day, day of year, and year. We model the annual trend with a linear term to capture long term increases or decreases in the series. Thus, the model we fit is

$$t_i = \beta_0 + s_1(h_i) + s_2(d_i) + \beta_1 y_i + \epsilon_i, \quad (1)$$

where t_i is the temperature at a given station and pressure level; h_i , d_i and y_i are the hour, day, and year of the i -th observation, respectively; β_0 is the intercept; β_1 is the coefficient for the long

term trend; and $s_1(\cdot)$ and $s_2(\cdot)$ are cubic regression splines.

Typically the error term, ϵ_i , in Equation (1) would be modeled as normal with some unknown variance, but the distribution of the error terms could be skewed or have heavier tails than a normal distribution. Thus, we use a skew- t distribution for the errors of this model, which has 4 parameters, ξ, σ, α , and ν , which are useful in controlling the first four moments of the distribution [3]. This distribution is very flexible and can handle skewed and heavy-tailed data.

Additionally, we expect there to be temporal correlation in the error terms. However, since we have already included hourly and seasonal terms in the model, we expect most of this autocorrelation to be explained, so an AR(1) time series model is sufficient to account for the remaining structure in the residuals. This model assumes that each error term has some fixed correlation with the error one time step in the past, and thus can be estimated by simply computing the correlation between t_i and t_{i+1} when the observations are equally spaced.

However, for radiosonde data, observations are not equally spaced in time. Launches are scheduled globally at 0 and 12 UTC; however, many deviations from this pattern are observed, especially in the historic record. Most observations are within an hour or two of the scheduled launches, but in some instances, no launches occur on a given day, and on others, more than two radiosondes are launched. Thus, to estimate the lag- h autocorrelation, $\phi(h)$, in hours, we must use only those observations that are h time steps apart:

$$\hat{\phi}(h) = \frac{1}{|\mathcal{P}_h|} \sum_{(\hat{\epsilon}_i, \hat{\epsilon}_j) \in \mathcal{P}_h} \frac{(\hat{\epsilon}_i - \bar{\epsilon}_i)(\hat{\epsilon}_j - \bar{\epsilon}_j)}{\sqrt{s_{\epsilon_i} s_{\epsilon_j}}}, \quad (2)$$

where \mathcal{P}_h is the set of all pairs of residuals that are h hours apart (or within some window), and $\hat{\epsilon}_i$ is the observed residual from Equation (1). For an AR(1) model, we need only estimate $\phi(\cdot)$ at $h = 12$ hours, and we use a window of 5% of 12 hours, or 0.6 hours.

2.2 Data Simulation

The data simulation procedure has seven steps:

1. Fit Equation (1) to radiosonde temperature data at a given location and pressure level.

Then, fit a skew- t distribution to the error terms, and model the autocorrelation with Equation (2).

2. We choose a fixed time period and assume that two observations occur for each day within that time period: one in the morning and one in the evening. The time of each morning (evening) observation is simulated by sampling a time from the morning (evening) subset of the observed data. This process is done to ensure that variability in the simulated hour of observation is comparable with that of the observed data.
3. We use the GAM model fit in step 1 to determine the expected value of temperature at the simulated time, denoted \hat{t}_i .
4. To simulate the noise in the observations, we randomly draw values δ_i from a skew- t distribution with parameters as fit in step 1.
5. We wish to introduce autocorrelation in these δ_i . Thus, we simulate an AR(1) model via

$$\epsilon_i = \hat{\phi}(12)^{\Delta_{i-1}/12}\epsilon_{i-1} + \delta_i,$$

where ϵ_i is the simulated noise in the model at time i , and Δ_{i-1} is the time difference, in hours, between the $(i-1)$ th and i th observation. Note that the k th term in this series will depend on all the previous $k-1$ values. To ensure the correct correlation structure, we simulate 1,000 more values than we need and discard the first 1,000.

6. Then, \hat{t}_i is added to ϵ_i to construct a simulated series that is similar to real radiosonde temperatures.
7. Lastly, we contaminate this data with systematic and random errors. (a) Random errors are generated by sampling 1, 2, 5 or 10% of the observations and adding or subtracting a random error following a distribution of $N(10\sigma, 1\sigma^2)$, where σ is the standard deviation of the simulated series, estimated from Equation (1). (b) Systematic errors are generated by sampling 1, 2, or 3 observations uniformly per simulated decade and then drawing a break

size from a $N(0, 0.04\sigma^2)$. The break size is then added to all observations after the change point. Both the contaminated and uncontaminated datasets are stored for comparison. Figure 2 shows an example of one of the radiosonde temperature datasets as well as three different realizations of simulated and contaminated datasets.

We vary several additional factors within our data simulation to understand the effect that each factor has on homogenization algorithms and the sequence in which the algorithms are applied.

Climate Zones: Radiosonde temperature data from different climate zones can be dramatically different, and so we analyze data from many different climate zones. In [2], ten representative stations are chosen and analyzed from ten different climate types, and we simulate data based on models fit to these ten stations.

Pressure Level: Radiosonde temperature data can also vary over pressure level, and so we analyze the pressure levels chosen in [2]: 100 mb, 300 mb, and 850 mb.

Sample Size: For the sequencing study, sample sizes of twice-daily data simulated for 20, 40, and 80 years are used. The study comparing homogenization algorithms, however, is much more computationally expensive, so we use sample sizes of 10, 20, and 40 years.

3 Homogenization Algorithms

Radiosonde observations are collected over long periods of time, as long as 100 years for some stations, and therefore systematic changes in the mean temperature are not uncommon. These errors can happen for one of many reasons: changes in instrumentation, relocation of a station, post-processing of data, etc. Methods which detect and/or correct these errors are referred to as homogenization algorithms, and many such techniques have been developed by the meteorological community [1, 4, 8, 9, 18, 19, 20, 23]. Many of the homogenization algorithms make use of metadata, which document changes in the data collection process and/or compare data from neighboring stations. We do not evaluate such algorithms since we simulate data from one station and pressure level at a time. In [9], SNHT is applied by combining both metadata and

the ERA-40, but we use a simplified version that operates purely on the observed data.

In this section, we compare the abilities of four different homogenization algorithms to detect systematic errors when random errors are also present in the data. We summarize the methods we investigate, namely Binary Segmentation (BinSeg) [22], Pruned Exact Linear Time (PELT) [16], SNHT, and a new robust SNHT. Additionally, we study the performance of these algorithms when the seasonal trend is modeled and removed or simply ignored. We simulate data as described in Section 2 and then introduce change points and random errors, and we evaluate the ability of the algorithms to detect the known change points.

3.1 Methodology

Two algorithms, BinSeg and PELT, detect the number and location of change points by optimizing a cost function of the form

$$\sum_{i=1}^{m+1} [\mathcal{C}(y_{(\tau_{i-1}+1):\tau_i})] + \beta f(m), \quad (3)$$

where τ_i is the i th change point; m is the number of change points; \mathcal{C} is a cost function; $y_{(\tau_{i-1}+1):\tau_i}$ is the observed data between the $(i-1)$ and i th change point; and $\beta f(m)$ is a penalty term on the number of change points to prevent overfitting [16]. Note that, for notational convenience, τ_{m+1} is defined to be the last observation. Often, \mathcal{C} is chosen to be twice the negative log likelihood, and $f(\cdot)$ is linear.

Optimization of Equation (3) can be done in several ways. BinSeg uses a divide-and-conquer algorithm: each observation is considered a candidate change point, and the one which leads to the largest reduction in the cost function is chosen as a change point. This change point then segments the data into two groups, and the same procedure is repeated on each segment. If no observations lead to a reduction in the cost function, then the procedure is terminated. BinSeg is known to be computationally efficient but is not guaranteed to reach the global minimum of the cost function.

PELT is another algorithm for optimizing Equation (3), but it computes the exact minimum.

It proceeds recursively as follows: first, the optimal number and location of change points is determined for observations 1 and 2 only. The optimal number and location of change points for the first three observations is then determined using this information, and more generally the optimal number and location of change points for the first $k + 1$ observations is determined by considering the optimal configurations for the first $2, 3, \dots, k$ observations. PELT is computationally efficient. For our analysis, we used the BinSeg and PELT algorithms implemented in the `changepoint` package in R [15].

The SNHT test works as follows. For each observation, two means are computed: one for the N days prior to observation i , $\bar{X}_{L,i}$, and one for the N days following, $\bar{X}_{R,i}$. Then, the test statistic

$$T_i = \frac{N}{s_i^2} \left((\bar{X}_{L,i} - \bar{X}_i)^2 + (\bar{X}_{R,i} - \bar{X}_i)^2 \right), \quad (4)$$

is computed where \bar{X}_i is the mean of $\bar{X}_{L,i}$ and $\bar{X}_{R,i}$, and s_i is the estimated standard deviation over the N days prior and N days following observation i . If there are not N observations both before and after the current observation, no test is performed. If the largest T_i exceeds some threshold at time $i = i^*$, we conclude that a change point occurred at time i^* , and we adjust all observations after time i^* by $\bar{X}_{L,i^*} - \bar{X}_{R,i^*}$. Homogenization now proceeds iteratively. T_i is recomputed for all i that are sufficiently far away from the current change points, $i \in \{1, \dots, n\} \setminus \{i^* - k, \dots, i^* + k\}$, and the test is performed again until no T_i exceed the threshold, and we use $k = N$. Note that in practice, it is generally preferable to homogenize to the most recent data, as that data is considered to be more reliable, and some follow this convention [4].

The test statistic in [9] differs from the one here in that a denominator of s_i is used in place of s_i^2 . However, we believe that to be a typographical error, as the statistic in [9] is derived from [1]. In [1], $\bar{X}_{L,i}$, $\bar{X}_{R,i}$ and \bar{X}_i are first standardized, and thus this is equivalent to our formulation. However, for comparison purposes, we homogenize the simulated datasets with both statistics. For the statistic with s_i in the denominator, a threshold of 100 is recommended, and we find that value to work well in our simulations.

Now, a different threshold for the statistic with s_i^2 in the denominator is required. We first note that the SNHT algorithms generally make the assumption that the data is approximately

normal with some fixed mean, and under this assumption, the statistic follows a chi-squared distribution with one degree of freedom [Josh: I still need to check this, but I'm pretty sure the Alexandersson paper states this.]. This assumption is unreasonable with radiosonde temperature data, as temperatures tend to follow seasonal trends. Thus, homogenization will likely be improved if the seasonal components of the data are removed. To do this, we fit a GAM model to the temperature data using only time of year as a predictor. We then subtract this fit from the data, giving errors about a seasonal mean. This simple model will leave the SNHT statistic unchanged, as we use a period of 365 of 730 days. It is now reasonable to assume that the SNHT statistic is chi-squared with one degree of freedom, and we pick the 99% quantile of this distribution as our threshold. Moreover, the SNHT statistic performs many dependent tests, and thus we use the Bonferroni-Hochberg adjustment for multiple testing.

We propose an alternative estimator that replaces the means and standard deviation in Equation (4) with the Huber M-estimator of the mean and standard deviation [11]. These robust estimators of center and spread are computed as follows:

1. First, the estimates of the mean, $\hat{\mu}$, and standard deviation, $\hat{\sigma}$, are initialized to

$$\hat{\mu} = \text{median}(\mathbf{x})$$

$$\hat{\sigma} = \text{MAD}(\mathbf{x}),$$

where \mathbf{x} is a vector of the data, and MAD is the median absolute deviation, defined as

$$MAD = \text{median}(|x_i - \text{median}(\mathbf{x})|).$$

2. Then, Winsorized values, y_i , are computed. These are defined as

$$y_i = \begin{cases} \hat{\mu} - k\hat{\sigma} & : x_i \leq \hat{\mu} - k\hat{\sigma} \\ x_i & : \hat{\mu} - k\hat{\sigma} < x_i \leq \hat{\mu} + k\hat{\sigma} \\ \hat{\mu} + k\hat{\sigma} & : x_i > \hat{\mu} + k\hat{\sigma}, \end{cases}$$

and k is commonly taken to be 1.5, which is what we use.

3. Updated estimates of $\hat{\mu}$ and $\hat{\sigma}$ are computed as the mean of \mathbf{y} and the standard deviation of \mathbf{y} , respectively.
4. Steps 2 and 3 are repeated until $\hat{\mu}$ changes by less than $10^{-6}\hat{\sigma}$.

This definition forces unusually large observations to have little to no influence on the estimators of the mean and standard deviation, and they are robust against random errors, which may be present during homogenization. BinSeg and PELT are not robust against random errors when \mathcal{C} is chosen to be twice the negative Gaussian log likelihood. An R package implementing this method is available on CRAN at <http://cran.r-project.org/web/packages/snht/index.html>.

Evaluation of homogenization algorithms can be done by computing the number of simulated change points in the data that were accurately detected. However, it is unlikely that a homogenization algorithm will detect the exact time of the change point, and thus hit rate is not a very useful metric. Instead, we use efficiency as defined in [4]. Let \mathbf{x} , \mathbf{c} , and \mathbf{h} be the original, contaminated, and contaminated and homogenized time series, respectively and let the i -th observation be denoted by x_i , c_i , and h_i respectively. The Root Mean Square Error (RMSE) of \mathbf{h} is then defined as follows:

$$\text{RMSE}(\mathbf{h}) = \sqrt{\frac{1}{n} \sum_{i=1}^n (h_i - x_i)^2}.$$

Then, the efficiency of the homogenized series, where 1 means perfect skill, 0 means no improvement, and negative values indicate degradation is

$$\text{Eff}(\mathbf{h}) = \frac{\text{RMSE}(\mathbf{c}) - \text{RMSE}(\mathbf{h})}{\text{RMSE}(\mathbf{c})}.$$

The homogenization algorithm is not designed to locate or correct random errors. Furthermore, random errors in the data introduce variability in the estimate of efficiency, so we remove the random errors in \mathbf{c} and \mathbf{h} before computing the RMSE scores.

We compare the efficiency of all four different homogenization algorithms on simulated

datasets. The simulated datasets are selected to be either 10, 20, or 40 years long, and change point locations are simulated uniformly at random across the entire time series excluding the first and last year. We simulate either one, two, or three change points per decade. All of the homogenization algorithms considered have tuning parameters: for PELT and BinSeg we must choose penalty functions and the β constant, and for SNHT and its robust variant we must specify the period N . Thus, in our simulations we vary the following tuning parameters to observe their effect on the overall performance:

- PELT: We consider penalties of $\beta = n/2, n, 2n, 4n$, and $8n$.
- BinSeg: We use no penalty term and instead restrict the maximum number of change points that can occur, varying from 1 to 10.
- SNHT: This algorithm computes means of seasonal data, so periods which are multiples of a year should be considered. Thus, we use one and two year averaging windows with $N = 365$ or $N = 730$.
- Robust SNHT: We use $N = 365$ and $N = 730$.

In summary, the simulation process is as follows:

1. Simulate and contaminate data as described in Section 2.2.
2. Apply each homogenization algorithm to the contaminated dataset.
3. Store the efficiency of each of the homogenization algorithms.
4. Steps 1-3 are repeated 1,000 times for each climate zone and pressure level.

3.2 Results

Figure 3 depicts a boxplot of the efficiencies measured for each of the different algorithms across all 90,000 simulations (500 simulations for each of 3 pressure levels, 10 stations, 3 sample sizes, and 2 seasonal removal configurations). The robust version of the SNHT appears to achieve

the best efficiency among all homogenization algorithms considered. The BinSeg algorithms perform best when we force the algorithm to choose a small number of change points. However, in practice, we will not know the true number of change points, and the BinSeg algorithm is very sensitive to this choice. The PELT algorithm appears to perform best with a penalty of $n/2$, but its performance is worse than the alternative algorithms.

To further understand the performance of these algorithms, and to understand their sensitivity to different simulation parameters, we fit a logistic regression model to the simulation results. The response variable is 1 if efficiency is positive and 0 otherwise, and the independent variables we use are the sample size n , the outlier contamination rate, the station, the pressure level, the homogenization algorithm, and if a seasonal adjustment was applied. We fit 4 different logistic regression models: first one with main effects and then models with k -way interactions, where $k = 2, 3, 4$. Table 1 reports the deviance for each model. As the deviance does not decrease substantially after k increases beyond 2, we use the model with only 2-way interaction terms.

Table 2 displays the average fitted efficiency as a function of n , outlier contamination, and the homogenization method. The first number indicates the fitted efficiency, averaged across all station and pressure level combinations, and the number in parentheses indicates the proportion of station and pressure level combinations where this model attains the highest fitted efficiency. The robust SNHT with seasonal trend removed is the superior model in almost all scenarios. However, the traditional SNHT is occasionally better when the outlier contamination rate is small (0% or 1%). The robust SNHT with a 730 day period performs slightly better on the 40 year datasets, but not enough to warrant its use. However, this suggests that it may be optimal if many years of data are available. Based on these results, we use the robust SNHT for the remainder of this paper with $N = 365$ and seasonal trend removed.

Plots of the fitted probability that efficiency is positive are given in Figure 4. Each individual line represents a different simulation configuration (number of years simulated and outlier contamination rate). As seen previously, this plot shows that the SNHT and robust SNHT attain the highest fitted probabilities in all cases. Also, the efficiency generally improves as the number of years simulated increases.

4 Sequencing Study

Many radiosonde temperature datasets have observations collected over long periods of time. Thus, it is possible and likely that both systematic and random errors exist in the data. It is not clear if random errors should be removed from the data prior to systematic errors, or vice versa. Thus, this simulation study investigates the performance of different sequences of these quality control methods.

4.1 Random Error Detection

We follow the random error identification process developed and tested in [2]. Given that errors are present in the data, traditional methods of computing the mean and standard deviation are known to perform poorly. Thus, the authors use the two-sided Huber estimator, which produces a robust measure of the location and measures of scale for both the left and right sides of the distribution [11]. The two-sided Huber estimator differs from the estimator described in Section 3.1 only in that it produces two estimates of scale, σ_R and σ_L . The estimate for σ_R (σ_L) is computed using only the data to the right (left) of $\hat{\mu}$.

Anderson et al. [2] investigate several different strategies for selecting subsets of observations with which to estimate the Huber mean and standard deviations. The *Global* set uses all of the observations to estimate the parameters, and the *Hourly Combined* set takes all observations within a 45 day and 12 hour window of each observation across all years and computes parameter estimates for each one. Their final algorithm first removes observations whose z -scores based on the Global parameter estimates are greater than 6, and then removes observations whose z -scores based on the Hourly Combined parameter estimates exceed 5.

4.2 Sequencing Simulation

We apply four different sequencings of homogenization and random error identification to the data: homogenization followed by error detection; error detection followed by homogenization; homogenization followed by error detection followed by homogenization; and error detection

followed by homogenization followed by error detection. We refer to these approaches as “Sys-Ran,” “Ran-Sys,” “Sys-Ran-Sys,” and “Ran-Sys-Ran,” respectively.

We hypothesize that many random errors will not be detected if the data is not homogenized and that the homogenization procedure will not perform as well if random errors are not first removed. For these reasons, we included the two additional methods “Sys-Ran-Sys” and “Ran-Sys-Ran”. In both of these approaches, a homogenization procedure is performed after random error detection. Likewise, a random error detection will be performed after a homogenization algorithm as well.

In summary, the simulation process is as follows:

1. Simulate and contaminate data as described in Section 2.2.
2. Apply each sequencing of the quality control process.
3. Store the true and false positive rate for random error detection as well as the efficiency of the homogenization algorithm. The true positive rate, TPR, is defined as the percent of identified errors that are random errors, and the false positive rate, FPR, is defined as the percent of identified errors that are not random errors.
4. Steps 1-3 are repeated 1,000 times for each climate zone and pressure level.

4.3 Sequencing Results

The percent of error contamination as well as the number of simulated change points in the data can strongly influence TPR and FPR. Thus, we again fit logistic regression models to the simulation results, where we model each of TPR, FPR, and the probability that efficiency is positive as the response variables. The dependent variables are sample size, outlier contamination rate, station, pressure level, and sequencing.

We begin by fitting five logistic models for each of the three responses, first with only main effects and then models with all k -way interaction terms, where $k = 2, 3, 4, 5$. The deviances are given in Table 3. We again find that the deviance does not decrease much when 3-way interaction

terms are included in the model, and thus we use models with 2-way interaction terms for all three responses.

For the TPR model, we find that there is no significant difference between the sequencings “Ran-Sys-Ran,” “Sys-Ran,” and “Sys-Ran-Sys.” Table 4 shows the fitted TPR averaged across all stations and pressure levels, and the number in parentheses shows the percent of the time when that model attains the highest fitted TPR. Additionally, these effects are plotted in Figure 5. The table shows that the “Ran-Sys” sequencing performs best when the outlier contamination rate is small and when the sample size is relatively small. The other sequences perform quite similarly.

For the FPR model, we found no significant difference between the sequencings “Sys-Ran” and “Sys-Ran-Sys”. Table 5 shows the fitted false positive rates averaged across all stations and pressure levels. Additionally, these effects are plotted in Figure 6. In almost all simulations, “Ran-Sys” attained the lowest FPR; however, FPR is quite low across all models. Due to this fact, and the conclusions from the TPR model, we recommend one of the sequencings “Ran-Sys-Ran”, “Sys-Ran”, or “Sys-Ran-Sys” if the end goal is to maximize FPR and minimize TPR.

Lastly, results from fitting the efficiency model are shown in Table 6. The largest fitted efficiency is almost always obtained with the sequence “Sys-Ran-Sys”. However, as shown in Figure 7, the difference among the four sequencings is not statistically significant. Thus, we conclude that the sequencing chosen does not appear to have a large effect on the efficiency of the final homogenized data.

5 Case Study

In [9], station 70219 (Bethel, Alaska, USA) is analyzed for change points. We perform a similar analysis, although we use raw radiosonde temperatures rather than the difference between the observed temperature and a reanalysis. We apply the robust SNHT to this time series followed by the outlier detection algorithm described in Section 4.1. As the dataset contains 53 years of data, we use a window of 2 years.

The algorithm detected a total of 15 change points and 19 outliers. The detected change points

have mean shifts ranging from 1.01°C to 2.80°C , and the average is 1.55°C . The global random error detection algorithm identified 12 errors, with test statistics ranging from 6.91 to 28.51 and averaging 15.57. The windowed random error detection algorithm detected an additional 7 errors, with test statistics ranging from 5.63 to 6.95 and averaging 6.12. Figure 8 depicts a plot of the time series before and after the quality control algorithm has been applied. Only a subset of the data is plotted in the middle panel in order to show the effects of homogenization more clearly. In the 10 year period shown, three change points are detected as indicated by the dashed vertical lines. The detected change points shift the mean of the process, and the corrections made in the quality controlled dataset seem to improve the homogeneity of the final product.

6 Conclusion

In this study we evaluate several different homogenization techniques, and we find that the robust SNHT method with seasonality removed performs well. It attains a high efficiency, indicating that this method is reasonably effective at returning the data to its uncontaminated state. It attains higher efficiencies than the BinSeg and PELT algorithms, and it outperforms the non-robust SNHT even when the outlier contamination rate is small. The optimal period may be a function of the size of the dataset, but the 365 day period is optimal for datasets with 40 years or less.

We also evaluate the effect that the sequence in which the random error detection and homogenization algorithms are applied have on the final performance of the overall quality control routine. We find that failing to remove systematic errors before searching for random errors leads to a much lower true positive rate of the error removal algorithm in most cases. However, the removal of random errors first does not have a large influence on the detection of systematic errors.

In the absence of metadata regarding a station’s history, we recommend performing data homogenization first followed by random error detection. This two step procedure performs significantly better than its reversal, and it performs similarly to three step procedures. The three step procedures do not perform significantly better than “Sys-Ran.” Since the three step proce-

dures are more computationally expensive, especially given the size of the radiosonde archive, we recommend against their use. Our recommended approach can be applied throughout the archive with minimal human intervention; however, gains in random and systematic error detection may be realized if metadata, multiple pressure levels, or multiple station locations are combined, as in [12]. More work is needed to determine how to best handle both types of errors when such additional information is incorporated.

Acknowledgments

The authors wish to thank the organizers and participants of the SAMSI Surface Temperature Initiative Workshop held at the National Center for Atmospheric Research in Boulder, Colorado on July 8-16, 2014 for helpful discussions regarding the content of this paper.

References

- [1] H. Alexandersson. A homogeneity test applied to precipitation data. *Journal of Climatology*, 6(6): 661–675, 1986.
- [2] A. Anderson, J. M. Browning, J. Comeaux, A. S. Hering, and D. Nychka. A comparison of automated statistical quality control methods for error detection in historical radiosonde temperatures. *Submitted*, 2014.
- [3] A. Azzalini and A. Capitanio. Distributions generated by perturbation of symmetry with emphasis on a multivariate skew t -distribution. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 65(2): 367–389, 2003.
- [4] P. Domonkos. Measuring performances of homogenization methods. *Quarterly Journal of the Hungarian Meteorological Service*, 117(1): 91–112, 2013.
- [5] I. Durre, R. S. Vose, and D. B. Wuertz. Overview of the integrated global radiosonde archive. *Journal of Climate*, 19(1): 53–68, 2006.
- [6] W. P. Elliott and D. J. Gaffen. On the utility of radiosonde humidity archives for climate studies. *Bulletin of the Meteorological Society*, 72(10): 1507–1520, 1991.
- [7] R. E. Eskridge, O. A. Alduchov, I. V. Chernykh, Z. Panmao, A. C. Polansky, and S. R. Doty. A comprehensive aerological reference data set (CARDS): Rough and systematic errors. *Bulletin of the American Meteorological Society*, 76(10): 1759–1775, 1995.
- [8] C. Gruber and L. Haimberger. On the homogeneity of radiosonde wind time series. *Meteorologische Zeitschrift*, 17(5): 631–643, 2008.
- [9] L. Haimberger. Homogenization of radiosonde temperature time series using innovation statistics. *Journal of Climate*, 20(7): 1377–1403, 2007.

- [10] T. J. Hastie and R. J. Tibshirani. *Generalized Additive Models*, volume 43. CRC Press, 1990.
- [11] P. J. Huber. *Robust Statistics*. Springer, 2011.
- [12] R. Ignaccolo, M. Franco-Villoria, and A. Fassò. Modelling collocation uncertainty of 3d atmospheric profiles. *Stochastic Environmental Research and Risk Assessment*, 2014. doi: DOI10.1007/s00477-014-0890-7.
- [13] E. Kalnay, M. Kanamitsu, R. Kistler, W. Collins, D. Deaven, L. Gandin, M. Iredell, S. Saha, G. White, J. Woollen, et al. The NCEP/NCAR 40-year reanalysis project. *Bulletin of the American Meteorological Society*, 77(3): 437–471, 1996.
- [14] M. Kanamitsu, W. Ebisuzaki, J. Woollen, S.-K. Yang, J. Hnilo, M. Fiorino, and G. Potter. NCEP-DOE AMIP-II reanalysis (r-2). *Bulletin of the American Meteorological Society*, 83(11): 1631–1643, 2002.
- [15] R. Killick and I. A. Eckley. changepoint: An R package for changepoint analysis. *Journal of Statistical Software*, 58(3): 1–19, 2014. URL <http://www.jstatsoft.org/v58/i03/>.
- [16] R. Killick, P. Fearnhead, and I. Eckley. Optimal detection of changepoints with a linear computational cost. *Journal of the American Statistical Association*, 107(500): 1590–1598, 2012.
- [17] J. R. Lanzante. Resistant, robust and non-parametric techniques for the analysis of climate data: Theory and examples, including applications to historical radiosonde station data. *International Journal of Climatology*, 16(11): 1197–1226, 1996.
- [18] J. R. Lanzante, S. A. Klein, and D. J. Seidel. Temporal homogenization of monthly radiosonde temperature data. part I: Methodology. *Journal of Climate*, 16(2): 224–240, 2003.
- [19] Y. Li and R. Lund. Bayesian multiple changepoint detection using metadata. (*submitted*), 2014.
- [20] Q. Lu, R. Lund, and T. C. Lee. An MDL approach to the climate segmentation problem. *The Annals of Applied Statistics*, 4(1): 299–319, 2010.
- [21] J. Rennie, J. Lawrimore, B. Gleason, P. Thorne, C. Morice, M. Menne, C. Williams, W. G. Almeida, J. Christy, M. Flannery, et al. The international surface temperature initiative global land surface databank: Monthly temperature data release description and methods. *Geoscience Data Journal*, 2014. doi: 10.1002/gdj3.8.
- [22] A. Scott and M. Knott. A cluster analysis method for grouping means in the analysis of variance. *Biometrics*, pages 507–512, 1974. doi: 10.2307/2529204.
- [23] V. K. Venema, O. Mestre, E. Aguilar, I. Auer, J. A. Guijarro, P. Domonkos, G. Vertacnik, T. Szentimrey, P. Stepanek, P. Zahradnicek, et al. Benchmarking homogenization algorithms for monthly data. *Climate of the Past*, 8(1): 89–115, 2012.

Model Type	Deviance	% Reduction
Intercept Only	1973919	—
Main Effects	1564061	20.76%
2-Way Interactions	1505738	3.73%
3-Way Interactions	1483785	1.46%
4-Way Interactions	1475014	0.59%

Table 1: Deviance for the efficiency logistic regression models.

Number of Years	Outlier Contamination	No Adjustment			Seasonal Adjustment		
		SNHT-365	robust-365	robust-730	SNHT-365	robust-365	robust-730
10	0%	95.7% (3%)	97.5% (3%)	96.5% (0%)	98.0% (3%)	98.7% (90%)	97.3% (0%)
	1%	95.6% (0%)	97.8% (7%)	96.8% (0%)	97.9% (0%)	98.8% (93%)	97.4% (0%)
	2%	95.5% (0%)	98.1% (10%)	97.0% (0%)	97.7% (0%)	98.9% (90%)	97.5% (0%)
	5%	95.1% (0%)	98.8% (17%)	97.6% (0%)	97.1% (0%)	99.2% (83%)	97.7% (0%)
	10%	94.4% (0%)	99.4% (30%)	98.3% (0%)	95.8% (0%)	99.5% (70%)	98.0% (0%)
	0%	96.0% (0%)	97.6% (3%)	97.2% (7%)	98.0% (0%)	98.7% (90%)	97.6% (0%)
20	1%	95.9% (0%)	97.9% (10%)	97.4% (3%)	97.8% (0%)	98.8% (87%)	97.7% (0%)
	2%	95.7% (0%)	98.2% (17%)	97.6% (0%)	97.6% (0%)	98.9% (83%)	97.7% (0%)
	5%	95.3% (0%)	98.8% (23%)	98.0% (0%)	96.9% (0%)	99.1% (77%)	97.9% (0%)
	10%	94.5% (0%)	99.4% (30%)	98.6% (0%)	95.4% (0%)	99.5% (70%)	98.1% (0%)
	0%	96.5% (0%)	97.9% (0%)	98.2% (27%)	97.8% (0%)	98.5% (70%)	98.1% (3%)
	1%	96.3% (0%)	98.1% (3%)	98.3% (27%)	97.6% (0%)	98.6% (70%)	98.1% (0%)
40	2%	96.2% (0%)	98.4% (7%)	98.4% (23%)	97.3% (0%)	98.7% (70%)	98.1% (0%)
	5%	95.7% (0%)	98.9% (17%)	98.6% (13%)	96.5% (0%)	99.0% (70%)	98.2% (0%)
	10%	94.6% (0%)	99.4% (53%)	99.0% (0%)	94.5% (0%)	99.3% (47%)	98.3% (0%)
	0%						

Table 2: Fitted efficiency averaged over all station and pressure level combinations. Numbers in parentheses indicate the percent of station and pressure level combinations where the given model obtained the highest fitted efficiency. Bolded numbers are the highest within each row. Note that only 6 of the 38 models are considered here. For brevity, models were included only if they attained the highest efficiency for some configuration.

Model Type	TPR		FPR		Efficiency	
	Deviance	% Reduction	Deviance	% Reduction	Deviance	% Reduction
Intercept Only	6240630	—	5835529	—	159942	—
Main Effects	4631181	25.79%	1244423	78.68%	136467	14.68%
2-Way Interactions	4101383	11.44%	279232	77.56%	132980	2.56%
3-Way Interactions	4058965	1.03%	258216	7.53%	132585	0.30%
4-Way Interactions	4049724	0.23%	255145	1.19%	132421	0.12%
5-Way Interactions	4047592	0.05%	254845	0.12%	132411	0.01%

Table 3: Deviance for the sequencing logistic regression models.

Number of Years	Outlier Contamination	Ran-Sys	Ran-Sys-Ran	Sys-Ran	Sys-Ran-Sys
20	0%	56.9% (100%)	48.4% (0%)	48.4% (0%)	48.4% (0%)
	1%	55.2% (100%)	48.6% (0%)	48.6% (0%)	48.6% (0%)
	2%	53.4% (100%)	48.8% (0%)	48.8% (0%)	48.8% (0%)
	5%	48.0% (0%)	49.4% (63%)	49.4% (0%)	49.4% (37%)
	10%	39.2% (0%)	50.4% (63%)	50.4% (0%)	50.4% (37%)
40	0%	53.8% (100%)	49.2% (0%)	49.2% (0%)	49.2% (0%)
	1%	51.9% (100%)	49.3% (0%)	49.3% (0%)	49.3% (0%)
	2%	50.0% (83%)	49.4% (3%)	49.4% (0%)	49.4% (13%)
	5%	44.3% (0%)	49.6% (63%)	49.6% (0%)	49.6% (37%)
	10%	35.2% (0%)	50.0% (60%)	50.0% (0%)	50.0% (40%)
80	0%	47.5% (0%)	50.8% (63%)	50.8% (0%)	50.8% (37%)
	1%	45.4% (0%)	50.7% (63%)	50.7% (0%)	50.7% (37%)
	2%	43.3% (0%)	50.6% (63%)	50.6% (0%)	50.6% (37%)
	5%	37.1% (0%)	50.1% (60%)	50.1% (0%)	50.1% (40%)
	10%	27.8% (0%)	49.4% (40%)	49.4% (0%)	49.4% (60%)

Table 4: Fitted TPR averaged over all station and pressure level combinations. Numbers in parentheses indicate the percent of station and pressure level combinations where the given sequencing obtained the highest fitted TPR. Bolded numbers are the highest within each row.

Number of Years	Outlier Contamination	Ran-Sys	Ran-Sys-Ran	Sys-Ran	Sys-Ran-Sys
20	0%	0.059% (97%)	0.065% (3%)	0.065% (0%)	0.065% (0%)
	1%	0.055% (97%)	0.063% (3%)	0.063% (0%)	0.063% (0%)
	2%	0.052% (97%)	0.060% (0%)	0.060% (3%)	0.060% (0%)
	5%	0.043% (100%)	0.054% (0%)	0.054% (0%)	0.054% (0%)
	10%	0.032% (100%)	0.046% (0%)	0.046% (0%)	0.046% (0%)
40	0%	0.054% (97%)	0.064% (3%)	0.064% (0%)	0.064% (0%)
	1%	0.051% (97%)	0.061% (3%)	0.061% (0%)	0.061% (0%)
	2%	0.048% (100%)	0.059% (0%)	0.059% (0%)	0.059% (0%)
	5%	0.039% (100%)	0.053% (0%)	0.053% (0%)	0.053% (0%)
	10%	0.030% (100%)	0.046% (0%)	0.045% (0%)	0.045% (0%)
80	0%	0.046% (100%)	0.061% (0%)	0.061% (0%)	0.061% (0%)
	1%	0.043% (100%)	0.059% (0%)	0.059% (0%)	0.059% (0%)
	2%	0.040% (100%)	0.057% (0%)	0.057% (0%)	0.057% (0%)
	5%	0.034% (100%)	0.051% (0%)	0.051% (0%)	0.051% (0%)
	10%	0.025% (100%)	0.044% (0%)	0.043% (0%)	0.043% (0%)

Table 5: Fitted FPR averaged over all station and pressure level combinations. Numbers in parentheses indicate the percent of station and pressure level combinations where the given sequencing obtained the lowest fitted FPR. Bolded numbers are the lowest within each row.

Number of Years	Outlier Contamination	Ran-Sys	Ran-Sys-Ran	Sys-Ran	Sys-Ran-Sys
20	0%	61.4% (0%)	61.4% (0%)	61.3% (0%)	61.8% (100%)
	1%	65.2% (0%)	65.2% (0%)	65.1% (0%)	65.6% (100%)
	2%	68.8% (0%)	68.8% (0%)	68.7% (0%)	69.2% (100%)
	5%	78.4% (0%)	78.4% (0%)	78.3% (0%)	78.6% (100%)
	10%	89.3% (0%)	89.3% (0%)	89.2% (0%)	89.4% (100%)
40	0%	73.3% (0%)	73.3% (0%)	73.2% (0%)	73.6% (100%)
	1%	77.8% (0%)	77.8% (0%)	77.7% (0%)	78.0% (100%)
	2%	81.7% (0%)	81.7% (0%)	81.6% (0%)	81.9% (100%)
	5%	90.2% (0%)	90.2% (0%)	90.2% (0%)	90.4% (100%)
	10%	97.0% (0%)	97.0% (0%)	97.0% (0%)	97.0% (100%)
80	0%	88.7% (0%)	88.7% (0%)	88.7% (0%)	88.9% (100%)
	1%	92.1% (0%)	92.1% (0%)	92.1% (0%)	92.2% (100%)
	2%	94.6% (0%)	94.6% (0%)	94.6% (0%)	94.7% (100%)
	5%	98.3% (0%)	98.3% (0%)	98.3% (0%)	98.4% (100%)
	10%	99.8% (0%)	99.8% (0%)	99.8% (0%)	99.8% (100%)

Table 6: Fitted efficiency averaged over all station and pressure level combinations. Numbers in parentheses indicate the percent of station and pressure level combinations where the given sequencing obtained the highest fitted efficiency. Bolded numbers are the highest within each row.

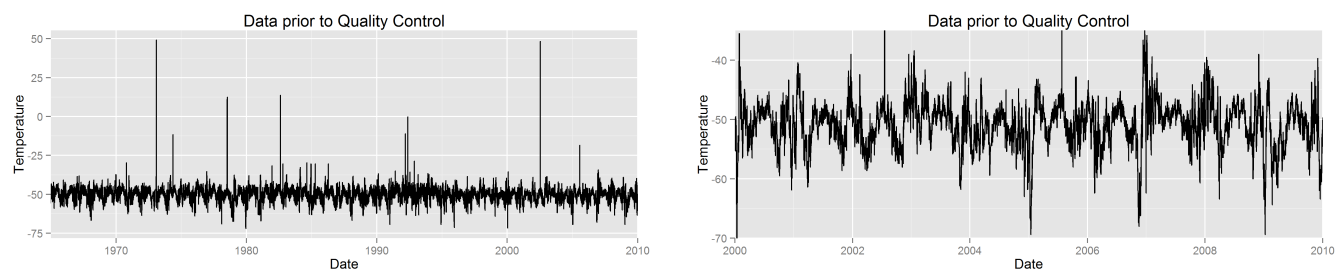


Figure 1: Temperature for Station 70219 plotted over time. The second image is the same as the first but zoomed in to the last ten years to show more detail.

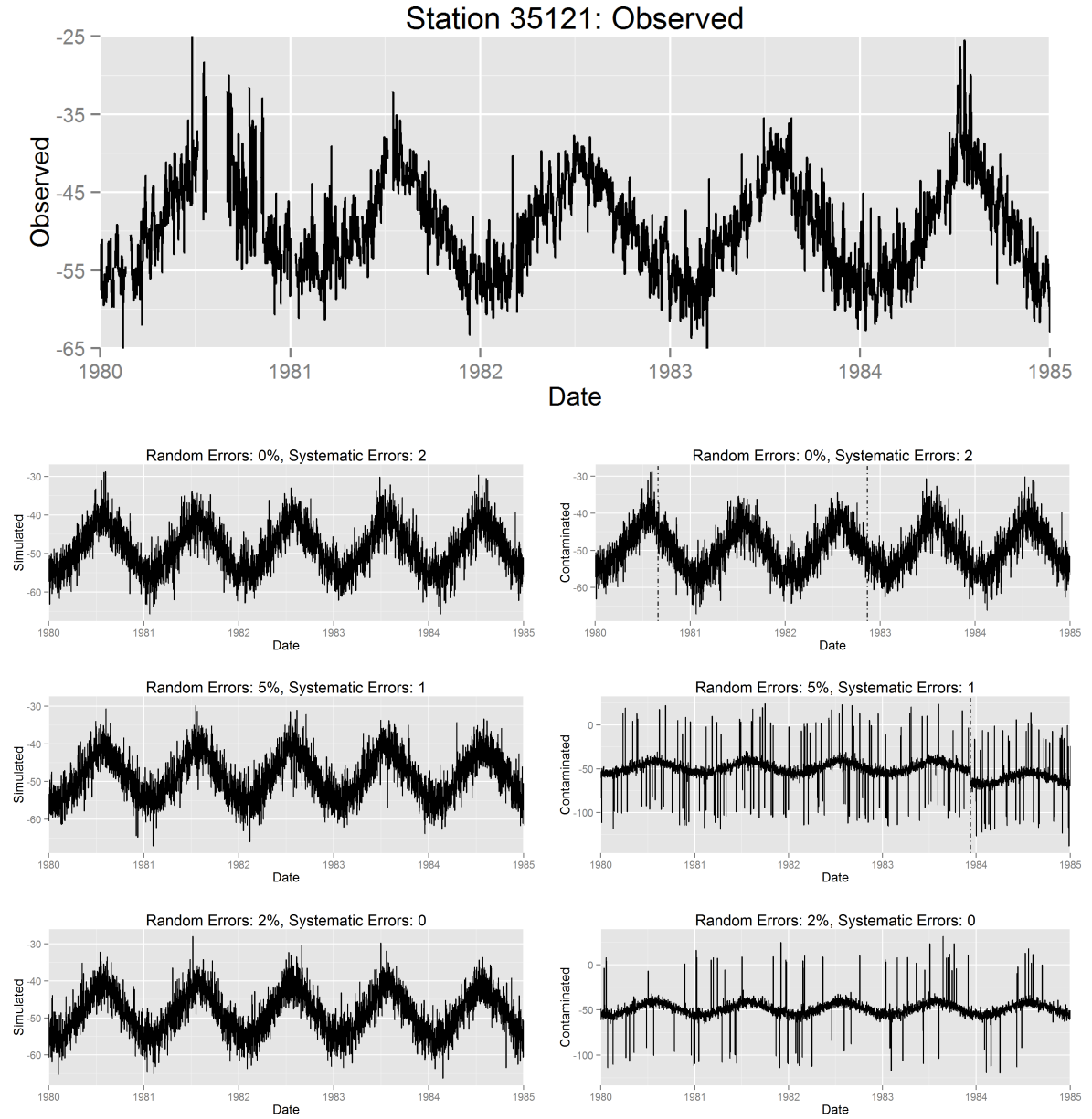


Figure 2: Time series plots of radiosonde temperature data from station 35121. The top plot shows the observed time series, and the following 3 pairs show three realizations of simulated datasets. The plots on the left show the simulated data prior to contamination, and the right plots show the data after contamination with random and systematic errors.

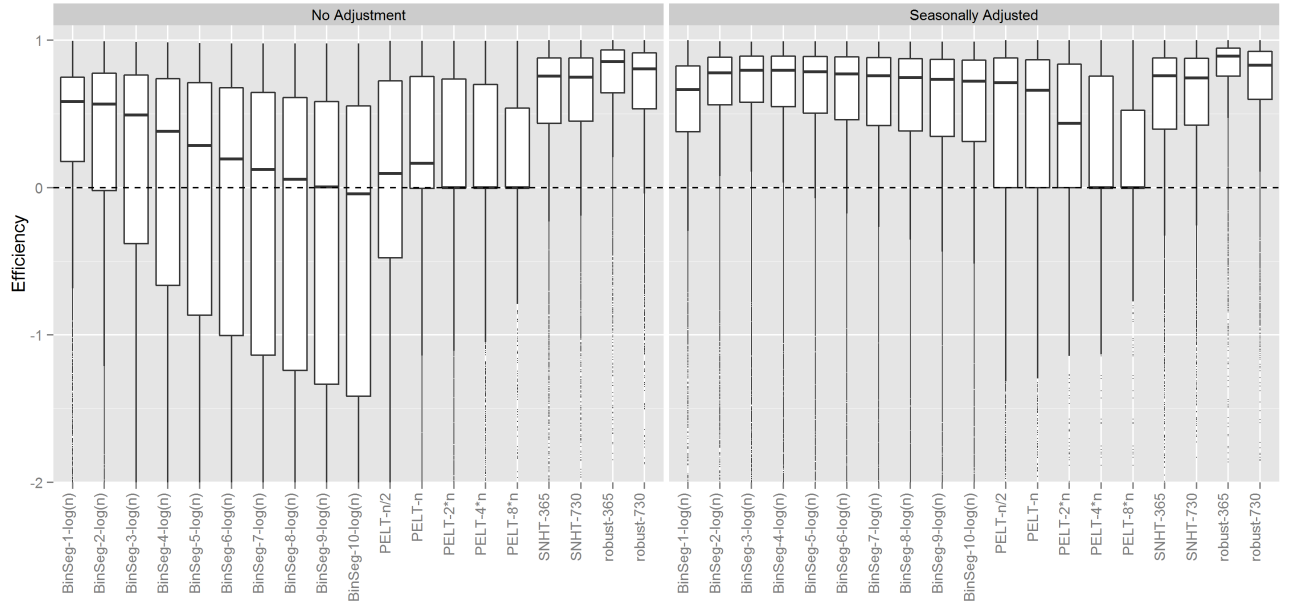


Figure 3: Boxplot of efficiency scores for the various homogenization algorithms. Note that this graph is constrained to the efficiency range of $(-3, 1)$ in order to show more detail. The SNHT and the robust SNHT perform substantially better than their alternatives.

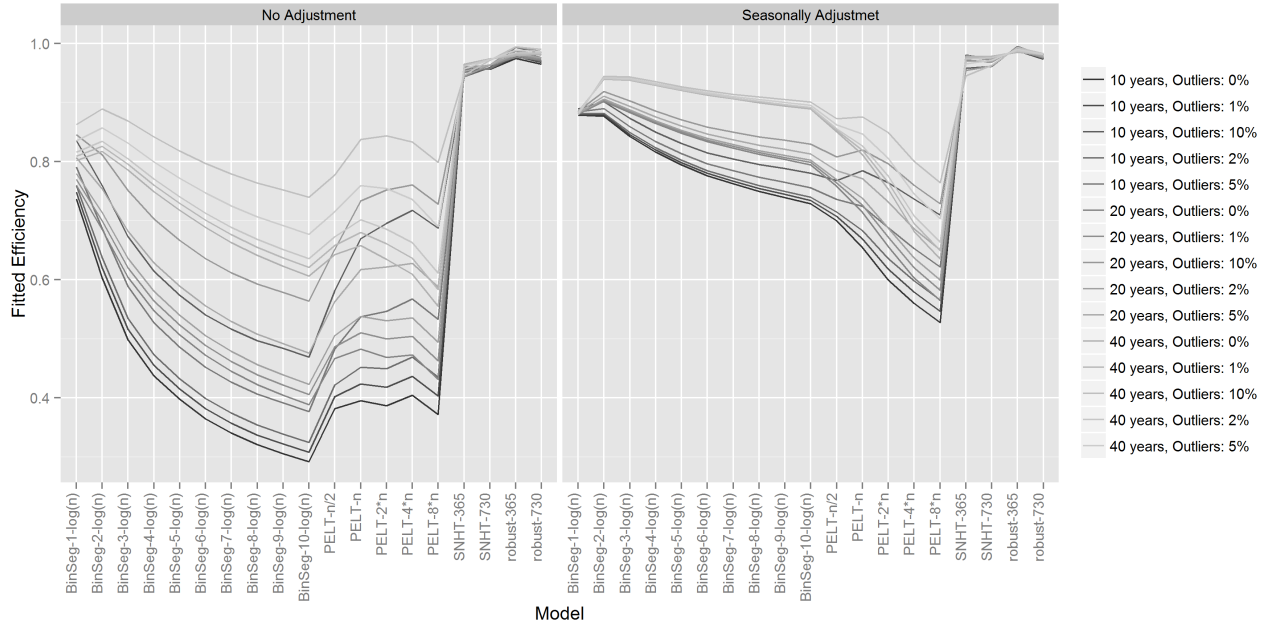


Figure 4: This graph depicts the estimated probability from the efficiency logistic regression model. Each line represents a different simulation configuration (number of years simulated and outlier contamination rates). This shows that there is substantial variability among the different simulations but that the SNHT and robust SNHT models perform better than the alternatives in all simulated configurations.

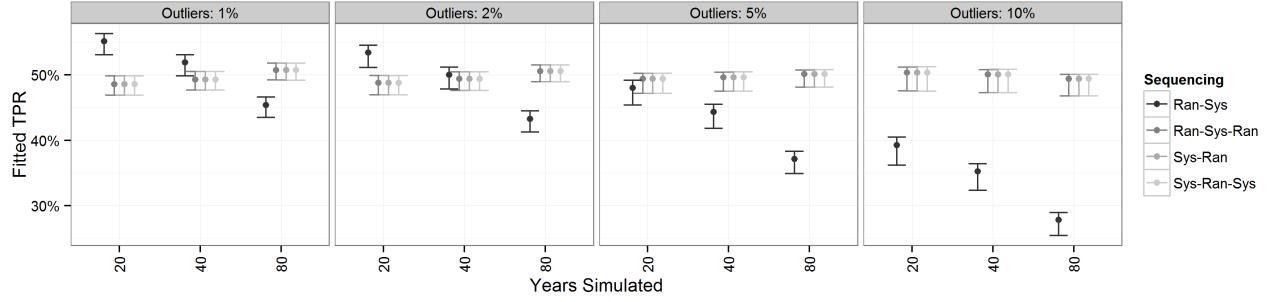


Figure 5: This graph depicts the estimated TPR from the logistic regression model. The middle of each error bar is the estimated TPR averaged over all station and pressure level combinations. The maximum (minimum) of the error bar is the highest (lowest) TPR obtained across all station and pressure level combinations.

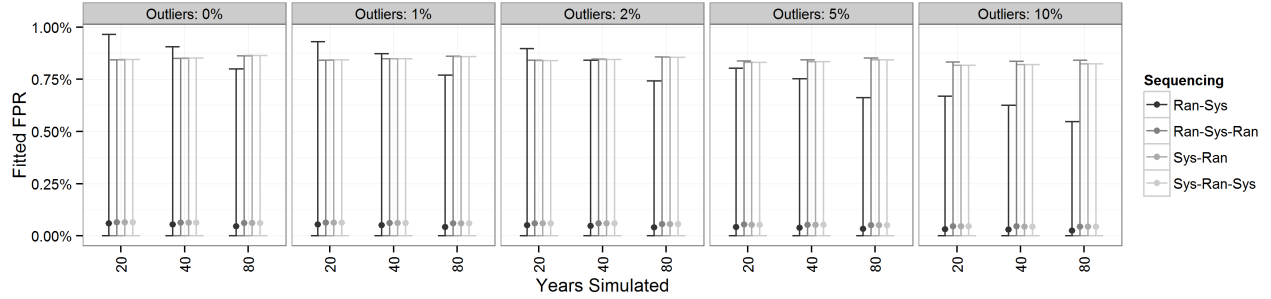


Figure 6: This graph depicts the estimated FPR from the logistic regression model. The middle of each error bar is the estimated FPR averaged over all station and pressure level combinations. The maximum (minimum) of the error bar is the highest (lowest) FPR obtained across all station and pressure level combinations.

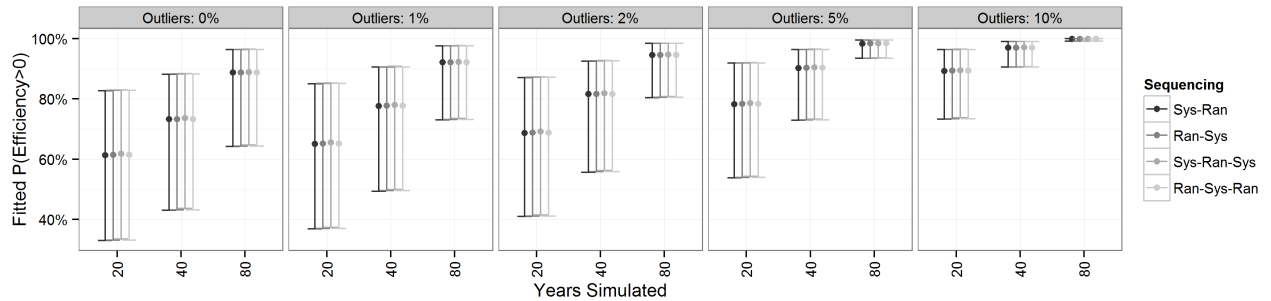


Figure 7: This graph depicts the estimated efficiency from the logistic regression model. The middle of each error bar is the estimated efficiency averaged over all station and pressure level combinations. The maximum (minimum) of the error bar is the highest (lowest) efficiency obtained across all station and pressure level combinations.

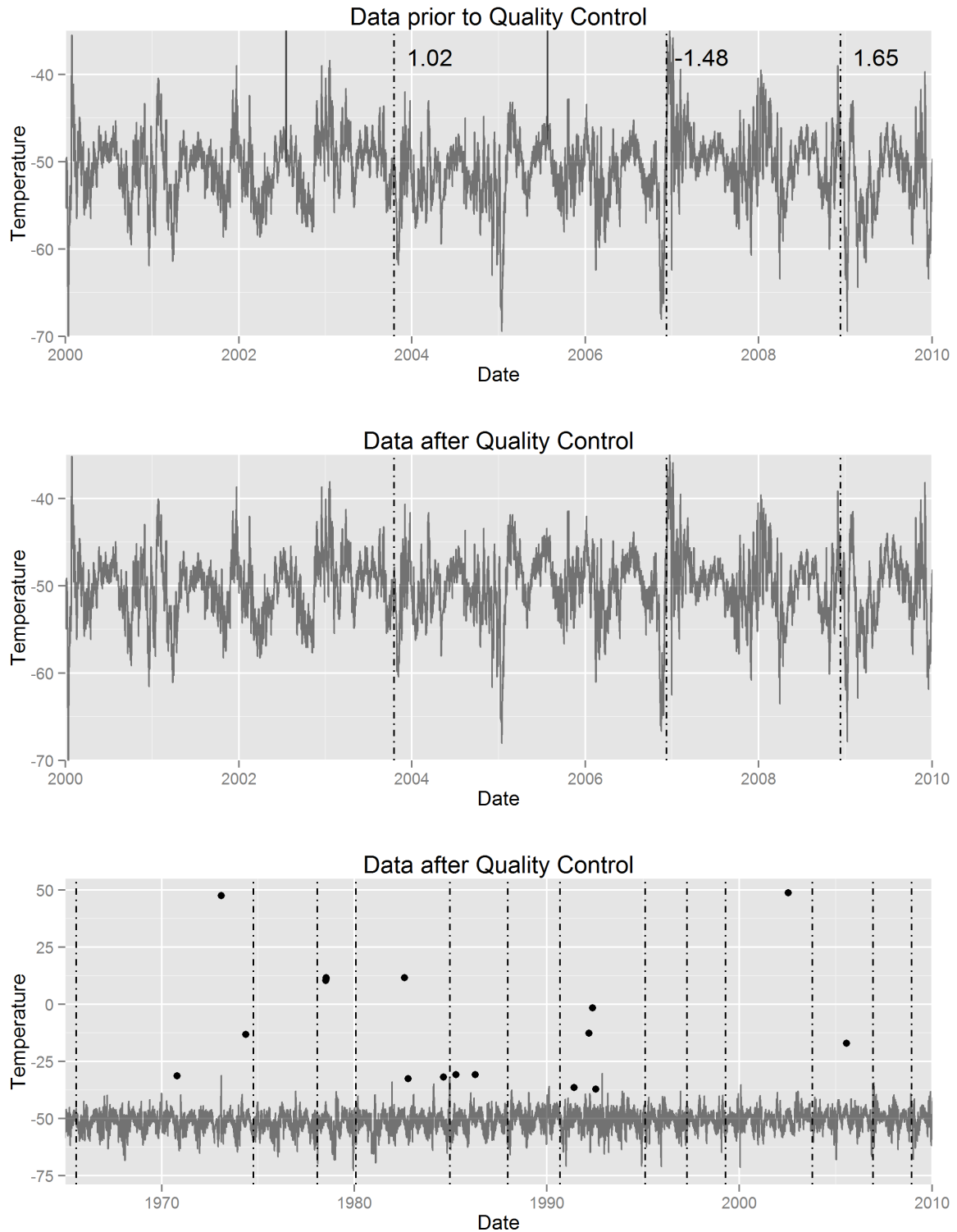


Figure 8: Time series plot of the radiosonde temperature data at station 70219 (Bethel, Alaska, USA). The top plot shows the data prior to the quality control algorithm, and the other two show the data after. Dashed lines indicate detected change points, numbers indicate the magnitude of the corresponding change point, and dots indicate detected random errors.