Bayesian Learning Lecture 10 - Bayesian Model Comparison

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Overview

- Bayes factor and posterior model probabilities
- Marginal likelihood
- **■** Log Predictive Score

Using likelihood for model comparison

- Consider two models for the data $y = (y_1, ..., y_n)$: M_1 and M_2 .
- Let $p_i(y|\theta_i)$ denote the data density under model M_i .
- If we know θ_1 and θ_2 , the likelihood ratio is useful

$$\frac{p_1(y|\theta_1)}{p_2(y|\theta_2)}.$$

The likelihood ratio with ML estimates plugged in:

$$\frac{p_1(y|\hat{\theta}_1)}{p_2(y|\hat{\theta}_2)}.$$

- Bigger models always win in estimated likelihood ratio.
- Hypothesis tests are problematic for non-nested models. End results are not very useful for analysis.

Bayes factor and posterior model probabilities

- I Just use your priors $p_1(\theta_1)$ och $p_2(\theta_2)$.
- lacksquare The marginal likelihood for model M_k with parameters $heta_k$

$$p_k(y) = \int p_k(y|\theta_k)p_k(\theta_k)d\theta_k.$$

- \blacksquare θ_k is 'removed' by the averaging wrt prior. Priors matter!
- The Bayes factor

$$B_{12}(y) = \frac{p_1(y)}{p_2(y)}.$$

■ Posterior model probabilities

$$\underbrace{\Pr(M_k|y)}_{\text{posterior model prob.}} \propto \underbrace{p(y|M_k)}_{\text{marginal likelihood prior model prob.}} \cdot \underbrace{\Pr(M_k)}_{\text{prior model prob.}}$$

Bayesian hypothesis testing - Bernoulli

Hypothesis testing is just a special case of model selection:

$$\begin{aligned} M_0:&x_1,...,x_n \overset{iid}{\sim} Bernoulli(\theta_0) \\ M_1:&x_1,...,x_n \overset{iid}{\sim} Bernoulli(\theta), \theta \sim Beta(\alpha,\beta) \\ &p(x_1,...,x_n|M_0) = \theta_0^s(1-\theta_0)^f, \\ &p(x_1,...,x_n|M_1) &= \int_0^1 \theta^s(1-\theta)^f B(\alpha,\beta)^{-1} \theta^{\alpha-1}(1-\theta)^{\beta-1} d\theta \\ &= B(\alpha+s,\beta+f)/B(\alpha,\beta), \end{aligned}$$

where $B(\cdot, \cdot)$ is the Beta function.

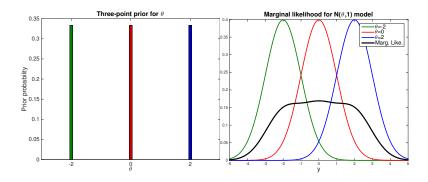
Posterior model probabilities

$$Pr(M_k|x_1,...,x_n) \propto p(x_1,...,x_n|M_k)Pr(M_k)$$
, for $k = 0, 1$.

The Bayes factor

$$BF(M_0; M_1) = \frac{p(x_1, ..., x_n | H_0)}{p(x_1, ..., x_n | H_1)} = \frac{\theta_0^s (1 - \theta_0)^t B(\alpha, \beta)}{B(\alpha + s, \beta + f)}.$$

Priors matter



Example: Geometric vs Poisson

- Model 1 Geometric with Beta prior:
 - \triangleright $y_1, ..., y_n | \theta_1 \stackrel{iid}{\sim} Geo(\theta_1)$
 - \blacktriangleright $\theta_1 \sim Beta(\alpha_1, \beta_1)$
- Model 2 Poisson with Gamma prior:
 - \downarrow $y_1, ..., y_n | \theta_2 \stackrel{iid}{\sim} Poisson(\theta_2)$
 - \blacktriangleright $\theta_2 \sim Gamma(\alpha_2, \beta_2)$
- \blacksquare Marginal likelihood for M_1

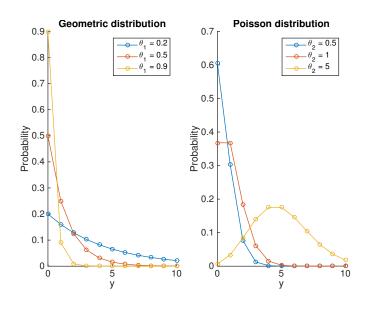
$$p_1(y_1, ..., y_n) = \int p_1(y_1, ..., y_n | \theta_1) p(\theta_1) d\theta_1$$

$$= \frac{\Gamma(\alpha_1 + \beta_1)}{\Gamma(\alpha_1) \Gamma(\beta_1)} \frac{\Gamma(n + \alpha_1) \Gamma(n\bar{y} + \beta_1)}{\Gamma(n + n\bar{y} + \alpha_1 + \beta_1)}$$

 \blacksquare Marginal likelihood for M_2

$$p_2(y_1,...,y_n) = \frac{\Gamma(n\bar{y} + \alpha_2)\beta_2^{\alpha_2}}{\Gamma(\alpha_2)(n + \beta_2)^{n\bar{y} + \alpha_2}} \frac{1}{\prod_{i=1}^n y_i!}$$

Geometric and Poisson



Geometric vs Poisson

Priors match prior predictive means:

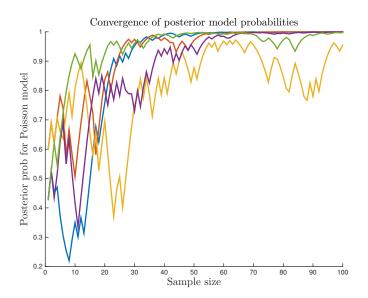
$$E(y_i|M_1) = E(y_i|M_2) \iff (\alpha_1 - 1)\alpha_2 = \beta_1\beta_2$$

Data: $y_1 = 0$, $y_2 = 0$.

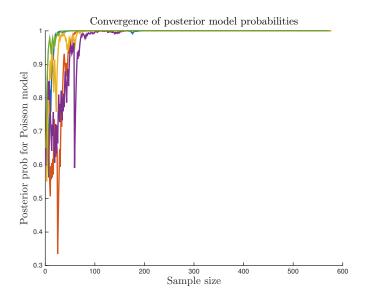
Data: $y_1 = 3$, $y_2 = 3$

Data: $y_1 = 5$, $y_2 = 5$.								
	$\alpha_1=2$, $\beta_1=2$	$\alpha_1 = 11, \beta_1 = 20$	$lpha_1=$ 101, $eta_1=$ 200					
	$\alpha_2 = 2$, $\beta_2 = 1$	$lpha_2=$ 20, $eta_2=$ 10	$lpha_2=$ 200, $eta_2=$ 100					
BF_{12}	0.21	0.28	0.30					
$\Pr(M_1 y)$	0.18	0.22	0.23					
$Pr(M_2 y)$	0.82	0.78	0.77					

Geometric vs Poisson for Pois(1) data



Geometric vs Poisson for Pois(1) data



Model choice in multivariate time series¹

■ Multivariate time series

$$\mathbf{x}_t = \alpha \beta' \mathbf{z}_t + \Phi_1 \mathbf{x}_{t-1} + ... \Phi_k \mathbf{x}_{t-k} + \Psi_1 + \Psi_2 t + \Psi_3 t^2 + \varepsilon_t$$

- Need to choose:
 - **Lag length**, (k = 1, 2.., 4)
 - ► Trend model (s = 1, 2, ..., 5)
 - ▶ Long-run (cointegration) relations (r = 0, 1, 2, 3, 4).

The most prob	SABLE	(k, r, s)) сом	BINATI	ONS IN	THE	Danisi	H MON	ETARY	DATA.
k	1	1	1	1	1	1	1	1	0	1
r	3	3	2	4	2	1	2	3	4	3
s	3	2	2	2	3	3	4	4	4	5
p(k, r, s y, x, z)	.106	.093	.091	.060	.059	.055	.054	.049	.040	.038

¹Corander and Villani (2004). Statistica Neerlandica.

Some properties

Coherence of pair-wise comparisons

$$B_{12}=B_{13}\cdot B_{32}$$

Consistency when true model is in $\mathcal{M} = \{M_1, ..., M_K\}$

$$\Pr\left(M = M_{TRUE}|\mathsf{y}\right) o 1 \quad \text{as} \quad n o \infty$$

■ "KL-consistency" when $M_{TRUF} \notin \mathcal{M}$

$$\Pr\left(M = M^* | \mathsf{y}\right) \to 1 \quad \mathsf{as} \quad n \to \infty,$$

 M^* minimizes KL divergence between $p_M(y)$ and $p_{TRUE}(y)$.

- Smaller models always win when priors are very vague.
- Improper priors cannot be used for Bayes factors.



Marginal likelihood measures out-of-sample predictive performance

■ The marginal likelihood can be decomposed as

$$p(y_1,...,y_n) = p(y_1)p(y_2|y_1)\cdots p(y_n|y_1,y_2,...,y_{n-1})$$

Assume that y_i is independent of $y_1, ..., y_{i-1}$ conditional on θ :

$$p(y_i|y_1,...,y_{i-1}) = \int p(y_i|\theta)p(\theta|y_1,...,y_{i-1})d\theta$$

- **Prediction** of y_1 is based on the prior of θ . Sensitive to prior.
- Prediction of y_n uses almost all the data to infer θ . Not sensitive to prior when n is not small.

Normal example

- Model: $y_1, ..., y_n | \theta \stackrel{iid}{\sim} N(\theta, \sigma^2)$ with σ^2 known.
- Prior: $\theta \sim N(0, \kappa^2 \sigma^2)$.
- Intermediate posterior at time i-1

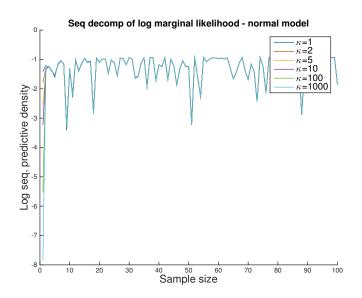
$$\theta|y_1,...,y_{i-1} \sim N\left[\frac{(i-1)}{\kappa^{-2} + (i-1)}\bar{y}_{i-1}, \frac{\sigma^2}{\kappa^{-2} + (i-1)}\right]$$

Intermediate predictive density at time i-1

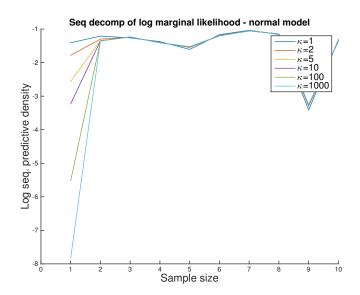
$$y_i|y_1,...,y_{i-1} \sim N\left[\frac{(i-1)}{\kappa^{-2}+(i-1)}\bar{y}_{i-1},\sigma^2\left(1+\frac{1}{\kappa^{-2}+(i-1)}\right)\right]$$

- For $i=1,\ y_1\sim N\left[0,\sigma^2\left(1+\frac{1}{\kappa^{-2}}\right)\right]$ can be very sensitive to κ .
- For large $i: y_i|y_1,...,y_{i-1} \stackrel{approx}{\sim} N\left(\bar{y}_{i-1},\sigma^2\right)$, not sensitive to κ .

First observation is sensitive to κ



First observation is sensitive to κ - zoomed



Log Predictive Score - LPS

- Reduce sensitivity to the prior: sacrifice n^* observations to train the prior into a posterior.
- Predictive (Density) Score (PS). Decompose $p(y_1, ..., y_n)$ as $\underbrace{p(y_1)p(y_2|y_1)\cdots p(y_{n^*}|y_{1:(n^*-1)})}_{training} \underbrace{p(y_{n^*+1}|y_{1:n^*})\cdots p(y_n|y_{1:(n-1)})}_{test}$
- Usually report on log scale: Log Predictive Score (LPS).
- Time-series: obvious which data are used for training.
- Cross-sectional data: training-test split by cross-validation:

Fold 1	Fold 2	Fold 3	Fold 4	Fold 5
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5

Be careful with Bayes factors

- Be especially careful with Bayes factors (posterior model probabilities) in the following situations:
 - ► The compared models are
 - very different in structure
 - severly misspecified
 - very complicated (black boxes).
 - ▶ The **priors** for the parameters in the models are
 - not carefully elicited
 - only weakly informative
 - not matched across models.
 - ► The data
 - has outliers (in all models)
 - has a multivariate response.