

Probabilistic Artificial Intelligence

Problem Set 3

Oct 21, 2019

1. Markov chains and detailed balance

Assume that you are given a Markov chain with finite state space Ω and transition matrix T , which is defined for all $x, y \in \Omega$ and $t \geq 0$ as $T(x, y) := P(X_{t+1} = y \mid X_t = x)$. Furthermore, let π be the stationary distribution of the chain.

- (i) Show that, if for some t the current state X_t is distributed according to the stationary distribution and additionally the chain satisfies the detailed balance equations

$$\pi(x)T(x, y) = \pi(y)T(y, x), \text{ for all } x, y \in \Omega,$$

then the following holds for all $k \geq 0$ and $x_0, \dots, x_k \in \Omega$:

$$P(X_t = x_0, \dots, X_{t+k} = x_k) = P(X_t = x_k, \dots, X_{t+k} = x_0).$$

(This is why a chain that satisfies detailed balance is called *reversible*.)

- (ii) Show that, if T is a symmetric matrix, then the chain satisfies detailed balance, and the uniform distribution on Ω is stationary for that chain.

2. Convergence of the Metropolis-Hastings algorithm

We use Markov Chain Monte Carlo (MCMC) methods to sample from a target distribution $P(x) = Q(x)/Z$ using a proposal distribution $R(x|x')$, without computing the normalization constant Z . A famous MCMC method is the METROPOLIS-HASTINGS algorithm, given below:

Algorithm 1 METROPOLIS-HASTINGS

Input: Unnormalized target distribution $Q(x)$, proposal distribution $R(x|x')$

Initialize: x_1 arbitrary

For $t = 1, 2, \dots, T$:

1. Sample proposal x from the proposal distribution $R(x|x_t)$.
 2. Compute the *acceptance probability* $\alpha = \min\left(1, \frac{Q(x)R(x_t|x)}{Q(x_t)R(x|x_t)}\right)$.
 3. With probability α , set $x_{t+1} = x$, else $x_{t+1} = x_t$.
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The algorithm defines a Markov chain with transition kernel $T(x, x') = P(x_{t+1} = x' \mid x_t = x)$. In this exercise, we prove that the stationary distribution of this Markov chain is equal to the target distribution $P(x)$. *Remark:* While we show, that Metropolis-Hastings converges to the correct distribution, the proof doesn't tell us how fast it converges. In practice, we typically use samples only after a 'burn-in' period, which allows the chain to converge.

- (i) Show that if a unnormalized distribution Q on Ω satisfies the detailed balance equations,

$$Q(x)T(x, y) = Q(y)T(y, x), \text{ for all } x, y \in \Omega, \quad (1)$$

then $\pi(x) = \frac{1}{Z}Q(x)$ is the stationary distribution of the Markov chain defined by the transition kernel $T(x, x')$.

- (ii) Show that if METROPOLIS-HASTINGS transitions to a new state, i.e. $x_{t+1} \neq x_t$, then the transition probability $T(x_t, x_{t+1})$ can be written as

$$T(x_t, x_{t+1}) = \frac{1}{Q(x_t)} \min(Q(x_t)R(x_{t+1}|x_t), Q(x_{t+1})R(x_t|x_{t+1})) . \quad (2)$$

Use this to show that the detailed balance equation for Q is satisfied if $x_{t+1} \neq x_t$.

- (iii) Finally, show that if $x_t = x_{t+1}$, the detailed balance condition is trivially satisfied. *Remark:* You can still compute the transition probability $T(x_t, x_{t+1})$ for this case, but the result follows independent of the exact transition probability.

3. Forward and Reverse KL for Variational Inference

In variational inference we seek to find a distribution Q in a class of distributions \mathcal{Q} , that minimizes the KL-distance to a target distribution P , i.e.

$$Q \in \arg \min_{Q \in \mathcal{Q}} \text{KL}(Q\|P) . \quad (3)$$

The KL-distance (for finite support) is defined as $\text{KL}(Q\|P) = \sum_x Q(x) \log(Q(x)/P(x))$. The KL-distance is not symmetric, so in general, $\text{KL}(Q\|P) \neq \text{KL}(P\|Q)$. We refer to $\text{KL}(P\|Q)$ as *forward KL* and $\text{KL}(Q\|P)$ as *reverse KL*. It is possible to use both, forward and reverse KL for variational inference, but the approximations will be different. In the plots below, we fit the true distribution $P(x)$ with a Gaussian $Q(x)$, using either forward or reverse KL. Explain which KL was used in either case!

