Effects of COVID-19 related social media use on well-being

2 Abstract

- 3 In times of crisis such as the COVID-19 pandemic, citizens need to stay informed about
- recent political events. To this end, people increasingly use social media. However, because
- 5 social media are particularly engaging, many find it hard to disconnect, especially during
- 6 times of crisis. Using data from the Austrian Corona Panel Project consisting of 3,485
- ⁷ participants from 34 waves, controlling for several stable and varying confounders, the
- 8 results showed that COVID-19 related social media use did not meaningfully reduce
- 9 well-being. Other factors such as health, income, exercise, or internal locus of control
- showed larger and meaningful effects.
- 11 Keywords: COVID-19, well-being, social media, news use, panel study.

Effects of COVID-19 related social media use on well-being

During the COVID-19 pandemic it was critical to stay informed regarding the latest 13 developments. How dangerous is the virus? In what region is it spreading? How is it 14 transmitted? What are the current safety regulations? To obtain relevant information, 15 many people heavily relied on social media, with use being at an all time high (Statista, 16 2021). Some actually could not stop using social media to learn about COVID-19 related 17 news. A new phenomenon termed "doomscrolling" emerged (Sharma et al., 2022). Many users were glued to their screens and found it hard to pursue other relevant activities such 19 as working, taking a break, or even looking after their children (Klein, 2021). In the media 20 it was hence increasingly asked whether using social media for COVID-19 related reasons 21 would, next to all other stressors, create an additional burden on mental health (Sandstrom et al., 2021). Although research has begun addressing this question (e.g., Bendau et al., 23 2021; Eden et al., 2020; Sewall et al., 2021), it still largely unknown if COVID-19 related social media use during the pandemic has had a meaningful impact on well-being. This 25 study hence aims to 1) reveal the effect of the different types and channels of social media use on individual well-being, 2) provide generalizable and robust results by analyzing a 27 large-scale longitudinal data-set with 34 waves, and 3) determine the within-person causal 28 effects by analyzing how changes in social media use lead to changes in well-being. 29

30 Understanding Well-being and Media Use

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This study investigates how different facets of well-being are affected by different types and different channels of communication (Meier & Reinecke, 2020). Building on the typology of subjective well-being (Diener et al., 2018), three different well-being facets are analyzed: life satisfaction, positive affect, and negative affect. Because effects of social media depend on how they are used (Verduyn et al., 2015), I further distinguish three types of use and five popular channels. The types of use include reading, liking and sharing, and posting COVID-19 related content. In doing so, this study analyzes social media use focused on COVID-19 related content, which includes posting thoughts about

- the pandemic, reading posts and comments, or retweeting and liking COVID-19 related
- 40 news. Liking and sharing are combined as they both represent low-threshold,
- ⁴¹ platform-ingrained, easily quantifiable interactions. The five channels to be investigated are
- Facebook, Twitter, Instagram, WhatsApp, and YouTube, which at the time ranked among
- the most popular social media services in Austria.

44 Social Media Effects on Well-Being

How easily can well-being be affected by external influences? In general, according 45 to the set-point theory, well-being is surprisingly stable (Lykken, 1999). Although specific events such as marriage or salary increase can have significant impacts on well-being, in most cases effects are only short-term, with well-being after some time returning to prior levels (Sheldon & Lucas, 2014). Only very specific factors such as unemployment, disability, or death can cause long-term changes in well-being (Lucas, 2007). So although well-being can be affected by external events and factors, this does not happen easily. 51 Can social media use be such a factor? Current literature overviews suggest that 52 social media use on average does seem to decrease well-being (Meier & Reinecke, 2020). However, for most well-being outcomes, such as life satisfaction, general well-being, or 54 loneliness, the effects are small (Meier & Reinecke, 2020). These small findings can be 55 explained with the differential susceptibility of media effects model (Valkenburg & Peter, 2013), which states that there is substantial variation of media effects for individual users. 57 Whereas for some users social media are more beneficial, for others they are more harmful. On average, however, and this is central for this study here, effects tend to be small (Valkenburg & Peter, 2013). For example, in one study it was estimated that roughly one quarter of all users experienced negative effects, another quarter positive effects, while for 61 the rest the effects were neutral (Beyons et al., 2021). Whether or not effects are positive 62 or negative depend on (a) dispositional factors (e.g., personality, temperament, gender), (b) developmental factors (e.g., age, developmental tasks), (c) and social factors (e.g.,

environment, norms, upbringing). Finally, effects depend also on the content that is

consumed. If the content is aligned with dispositions, developmental capacities, and converging contexts, effects tend to be stronger (Valkenburg & Peter, 2013).

Why are the effects of social media use on well-being small on average? Two 68 prominent media effect theories argue implicitly against strong average negative effects. 69 First, according to mood management theory (Zillmann, 1988), using media can affect people's moods. Use can be stimulating or overwhelming, relaxing or boring. After some 71 time, users implicitly learn which media help them balance their mood and affect according 72 to their own situational needs (Zillmann, 1988). Those media that eventually become part 73 of one's media repertoire hence, on average, tend to be beneficial for users to regulate their 74 mood (Marciano et al., 2022). In conclusion, if a certain medium is used frequently, 75 mood-management theory argues that it is likely not detrimental for well-being. On the other hand, although mood management theory suggests that effects should not be overly negative, it could be that although short-term effects are positive, long-term effects are negative. Precisely because social media have so many positive consequences in the short run, this might cause problem in the long run—for example, because social media use displaces other potentially more meaningful and/or relaxing activities (Hall & Liu, 2022). 81 As with many other things, there can be too much of a good thing. It is therefore often asked whether social media can become addictive, and users sometimes express this fear 83 themselves (Yang et al., 2021). However, a recently published meta-analysis found that the 84 two most prominent measures of addiction, the Bergen Facebook Addiction Scale and the 85 Bergen Social Media Addiction Scale, have only small relations to well-being (Duradoni et al., 2020). In addition, the general idea of labeling excessive social and new media use as 87 addiction was criticized, arguing that social media use represent a new regular behavior that should not be pathologized (Galer, 2018; van Rooij et al., 2018). Second, while mood management theory considers media use mainly driven by 90

implicit learning experiences, uses and gratifications theory upholds that the process is
more explicit and rational (Katz et al., 1973). Users select those media that they expect to

have a desired effect, for example on mood, knowledge, or entertainment. If those
beneficial media effects are missing, people will spend their time elsewhere. And social
media, in general, offer several beneficial effects, explaining why they are used that much.
They help find relevant information, maintain and foster relationships, express one's
personality, and entertain oneself (Pelletier et al., 2020). In conclusion, because people
spend so much time on social media consuming COVID-19 related content, according to
both mood management theory and uses and gratifications theory this indirectly suggest
that average effects on well-being are likely not particularly negative.

101 Social Media During COVID-19

If we look at COVID-19 related use more specifically, how could the various types 102 and channels of COVID-19 related social media use affect well-being? Several uses and gratifications exist, which help explain why people used social media frequently during the 104 pandemic. Access to information and resources: Despite incorrect information, social 105 media provide a vast platform for disseminating accurate and timely information about 106 COVID-19 (John Hopkins University, 2023). It allows individuals to stay informed about 107 the latest updates, guidelines, and recommendations from reputable health organizations 108 and experts. Access to reliable information can help people make informed decisions, 109 alleviate uncertainties, and feel empowered during the pandemic. Educational 110 opportunities: Social media can be a valuable educational tool, with various organizations 111 and experts sharing informative content about COVID-19 (World Health Organization, 112 2023). This can include explanatory videos, infographics, and articles that help individuals 113 understand the virus, its transmission, prevention strategies, and vaccination information. 114 Access to educational resources on social media can empower individuals to make informed 115 decisions regarding their health. Community support and solidarity: Social media 116 platforms enable individuals to connect with others who are experiencing similar challenges 117 during the pandemic (Guazzini et al., 2022). Sharing experiences, offering support, and 118 expressing solidarity can foster a sense of community and reduce feelings of isolation. 119

Engaging in online communities and support groups can provide emotional support and create a network of like-minded individuals. Promoting mental health and well-being: 121 Many mental health organizations and professionals utilize social media to share tips, 122 strategies, and resources for maintaining mental well-being during the pandemic (Twitter, 123 2020). They provide guidance on coping mechanisms, stress reduction techniques, and 124 self-care practices. Engaging with such content might help individuals prioritize their 125 mental health and develop resilience during challenging times. Encouraging positive 126 behaviors: Social media campaigns and initiatives can promote positive behaviors related 127 to COVID-19, such as mask-wearing, physical distancing, hand hygiene, and vaccination 128 (Athey et al., 2023; Hunt et al., 2022). Public health organizations and influencers leverage 129 the power of social media to spread awareness and encourage responsible actions, 130 contributing to public health efforts and fostering a sense of collective responsibility. 131 On the other hand, the effects might be negative, perhaps best explained by the 132 following five mechanisms. Misinformation and rumors: Social media platforms can easily 133 spread false or misleading information about COVID-19 (Li et al., 2020). Due to the ease 134 of sharing and the lack of fact-checking, inaccurate information can go viral and might 135 cause confusion, anxiety, and panic among users. Overwhelm and anxiety: Constant 136 exposure to COVID-19-related content on social media can lead to information overload 137 and contribute to heightened anxiety levels (J. Fan & Smith, 2021). The rapid spread of 138 news, updates, and opinions can be overwhelming and might exacerbate existing stress or 139 fears about the pandemic (Sharma et al., 2022). Confirmation bias and echo chambers: 140 Social media algorithms are designed to show users content that aligns with their interests 141 and beliefs, potentially contributing to echo chambers (Cinelli et al., 2021). This might 142 lead to confirmation bias, where individuals are more likely to be exposed to information 143 that confirms their preexisting beliefs. In the context of COVID-19, this might reinforce 144 false or misleading information and prevent individuals from considering alternative 145 perspectives. Cyberbullying and harassment: Social media platforms are known for 146

fostering negativity, with users sometimes engaging in cyberbullying and harassment 147 (Giumetti & Kowalski, 2022). Discussions around COVID-19 can become heated and 148 polarized, leading to personal attacks and online conflicts. Such experiences threaten 149 mental well-being and might contribute to feelings of distress and isolation. Social 150 comparison and fear of missing out (FOMO): Social media often showcase the highlights 151 and accomplishments of others, encouraging social comparison (Przybylski et al., 2013). 152 During a pandemic, seeing posts about others' successes or seemingly perfect lives might 153 intensify feelings of inadequacy or FOMO, especially when individuals are unable to 154 participate in similar activities due to restrictions or personal circumstances (Sharma et al., 155 2022). This might impact mental health, as most humans have an innate need for social 156 support and physical presence, which online interactions might not fully replicate. 157 There is still little empirical research on how well-being is affected by social media 158 use that is focused on COVID-19 specifically. Echoing the theoretical rationales outlines 159 above, studies have yielded mixed results. Some studies found negative effects, indicating 160 that excessive social media use for COVID-19 news led to compulsive behavior and 161 increased stress levels, particularly due to upward social comparison (Klein, 2021; 162 Stainback et al., 2020; Yue et al., 2022). Individuals who relied on social media as their 163 primary information source reported higher levels of anxiety and depression symptoms 164 (Bendau et al., 2021). Similarly, increased COVID-19-related media consumption was 165 associated with higher psychological distress. On the other hand, some studies reported 166 positive outcomes. Certain individuals experienced increased virtual community and social 167 connectedness during the pandemic through social media, which contributed to their 168 well-being (Guazzini et al., 2022). Additionally, engaging more on social media was 169 associated with reduced feelings of loneliness (Latikka et al., 2022). A couple of studies 170 reported mostly neutral effects of social media use on well-being indicators (Bradley & 171 Howard, 2021; Eden et al., 2020; Sewall et al., 2021). Overall, the literature demonstrates a 172 mixed picture, highlighting both positive and negative effects of social media use focused 173

on or during COVID-19 on well-being (see also Bradley & Howard, 2021; Dörnemann et al., 2021; Liu & Tong, 2020; Riehm et al., 2020; Sewall et al., 2021).

In conclusion, given these mixed empirical results, together with the observation
that social media effects on well-being are very small in general, and that several plausible
theoretical mechanisms exist for both positive and negative effects, I expect that COVID-19
related communication on social media should not be decidedly positive or negative. It
seems most likely that both positive and negative coexist, but that on average using social
media for COVID-19 related reasons should not have substantial effects on well-being.

Hypothesis: The within-person effects of all measures of COVID-19 related social media use (types: reading, liking and sharing, posting; channels: Twitter, Instagram, Facebook, YouTube, WhatsApp) on all measures of well-being indicators (positive affect, negative affect, life satisfaction)—while controlling for several stable and varying covariates such as sociodemographic variables and psychological dispositions (see below)—will be trivial.

188 Method

89 Preregistration

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The hypotheses, the sample, the measures, the analyses, and the inference criteria 190 (SESOI, p-value) were preregistered on the Open Science Framework, accessible here: 191 https://osf.io/87b24/?view only=b2289b6fec214fa88ee75a18d45c18f3. Because in this 192 study I analyze data from an already existing large-scale data set, the preregistration was 193 done prior to accessing the data. The preregistration was designed on the basis of the 194 panel documentation online (Kittel et al., 2020). In some cases, it was impossible to 195 execute the analyses as I had originally planned, for example because some properties of 196 the variables only became apparent when seeing the actual data. The most relevant 197 deviations are reported below, and a complete list of all changes can be found in the online 198 companion website (https://XMtRA.github.io/Austrian Corona Panel Project). 199

Sample 200

The data come from the Austrian Corona Panel Project (Kittel et al., 2021), which 201 is a large-scale standalone panel study. The data are hosted on AUSSDA, are publicly 202 available here (https://doi.org/10.11587/28KQNS), and consist of 34 waves. Participants 203 were sampled from a pre-existing online access panel provided by the company 204 Marketagent, Austria. Panel members were incentivized with 180 credit points for each 205 wave of the study. The study was conducted between March 2020 and February 2023. 206 Between March 2020 and July 2020, the intervals between waves were weekly, until May 207 2022 (wave 32) monthly, and afterward after 5 months. Each wave consists of at least 1,500 208 respondents. Panel mortality was compensated through a continuous re-acquisition of new 209 participants. The sample size was N = 3,641, with overall 123,794 observations. For an overview of the study set-up, see Figure 1. 211 Achieved via quota sampling, the sample matched the Austrian population in terms 212 of age, gender, region/state, municipality size, and educational level. In order to 213 participate in the study, the respondents needed to be Austrian residents and had to be at 214 least 14 years of age. All respondents needed to have access to the internet (via computer 215 or mobile devices such as smartphones or tablets). Ethical review and approval was not 216 required for the study in accordance with the local legislation and institutional 217 requirements. The participants provided their written informed consent to participate in 218 this study. The average age was 40 years, 49 percent were male, 14 percent had a 219 University degree, and 5 percent were currently unemployed.

Smallest Effect Size of Interest

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Testing the hypothesis necessitates defining what is considered a "trivial effect size". 222 To this end, we need to define a so-called smallest effect size of interest (SESOI) (Lakens et 223 al., 2018). A trivial effect would then need to be smaller than the SESOI (see below). 224 What could be a minimally interesting, nontrivial effect? Being a normative question, 225 finding a clear, single, or unanimous answer is impossible. However, it is still necessary and

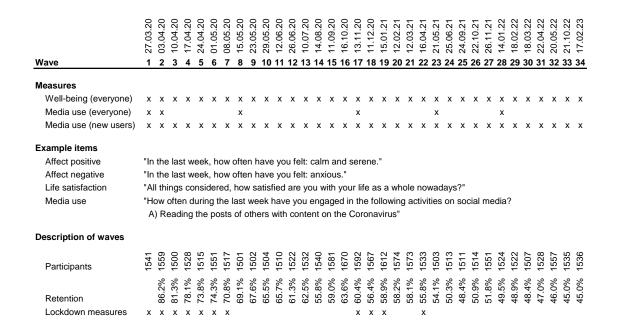


Figure 1 Overview of study set-up

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helpful to work toward a plausible benchmark. I suggest the following SESOI for this research question:

SESOI: If a heavy user of COVID-19 related social media news suddenly *stops* using social media altogether, this should have a *noticeable* impact on their overall well-being.

What does this mean practically and how can it be operationalized? In this study,

COVID-19 related social media use was measured on a 5-point scale, ranging from 1 =

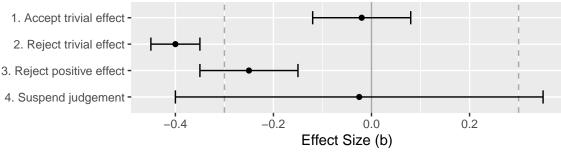
never to 5 = several times a day. Thus, a change of four units in social media use (e.g., a

complete stop) should correspond to a noticeable change in well-being. What is a noticeable

change in well-being? According to Norman et al. (2003), people can reliably distinguish

seven levels of satisfaction with health. So if satisfaction is measured on a 7-point scale, a 237 four unit change in social media use should result in a one unit change in life satisfaction. 238 In this study, life satisfaction was measured on an 11-point scale. If people can 239 reliably differentiate 7 levels, this corresponds to 11 / 7 = 1.57 unit change on an 11-point 240 scale. Hence, a four-point change in media use (e.g., a complete stop) should result in a 241 1.57-point change in life satisfaction. In a statistical regression analysis, b estimates the 242 change in the dependent variable if the independent variable increases by one point. For 243 life satisfaction, we would therefore define a SESOI of b = 1.57 / 4 = 0.39. For positive or 244 negative affect, which was measured on a 5-point scale, our SESOI would be b = 0.71 / 4 =245 0.18. Because we are agnostic as to whether the effects are positive or negative, the null 246 region includes both negative and positive effects. Finally, in order not to exaggerate 247 precision and to be less conservative, these numbers are reduced to nearby thresholds.¹ Together, this leads to a null region ranging from b = -.30 to b = .30 for life satisfaction, and b = -.15 to b = .15 for positive and negative affect. 250 The hypothesis is analyzed using the interval testing approach as proposed by 251 Dienes (2014). To illustrate, let us consider the case of life satisfaction [SESOI: -.30: 252 +.30]. If the 95% confidence interval falls completely within the null-region (e.g., b = -.05, 253 [95% CI: -.15, .05]), the hypothesis that the effect is trivial is supported. If the confidence 254 interval falls completely outside of the null-region (e.g., b = -.40, [95% CI: -.45, -.35]), the 255 hypothesis is rejected and the existence of a meaningful negative effect is supported. If the 256 confidence interval and the null region overlap (e.g., b = -.30, [95% CI: -.35, -.25]), the 257 hypothesis is not supported and the results are considered inconclusive, while a meaningful 258 positive effect is rejected. If the confidence interval exceeds both sides of the null region 259 (e.g., b = -.025, [95% CI: -.40, .35]), the hypothesis is not supported and judgement is 260 suspended. For an illustration, see Figure 2. 261

¹ Note that other researchers also decreased or recommended decreasing thresholds for effect sizes when analyzing within-person or cumulative effects (Beyens et al., 2021; Funder & Ozer, 2019).



Smallest effect size of interest: b = |.30|Null region: b = -.30, .30

Figure 2

Using confidence intervals to test a null region. In this study, a trivial effect of social media use on life satisfaction is defined as ranging from b = -.30 to b = .30. Figure adapted from Dienes (2014).

When using longitudinal designs to analyze causality, it is important to (a) focus on

Data Analysis

263 Causality

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within-person effects (Hamaker, 2014); to (b) control for confounders (Rohrer & 265 Murayama, 2021); and to (c) test a plausible interval between measures (Dormann & 266 Griffin, 2015). First, in non-experimental designs it makes much sense to analyze causal 267 effects from an internal, within-person perspective (Hamaker, 2014). If a specific person 268 changes their media diet, we need to measure how this behavior affects their well-being. 269 Between-person comparisons from longitudinal data cannot provide such insights (Lucas, 270 2022). To test the hypothesis, I thus consider only the within-person effects. 271 Second, to identify confounders we should control for variables that affect both 272 media use and well-being, which helps isolate the actual effect (Rohrer, 2018). Because we 273 are adopting a within-person perspective, we need to implement time-varying confounders 274 (Rohrer & Murayama, 2021). And because we are determining the *overall* causal effect, we 275 need to make sure not to control for mediating variables (Rohrer, 2018), for doing so would 276 bias our assessment of the causal effect. In this study, I hence preregistered to control for 277 the following variables, which either have already been shown or are likely to affect both

social media use and well-being, and which also are not mediators: gender, age, education,
Austria country of birth, Austria country of birth of parents, residency Vienna, text-based
news consumption, video-based news consumption, household size, health, living space,
access to garden, access to balcony, employment, work hours per week, being in
home-office, household income, outdoor activities, disposition to take risks, and locus of
control (Eger & Maridal, 2015; Ward et al., 2016).

Finally, one precondition of causality is temporal order and finding a plausible interval (Dormann & Griffin, 2015). If variables are stable, longer intervals are needed; if they fluctuate, shorter intervals. In the case of well-being, we need shorter intervals for the more fluctuating positive and negative affect, and longer ones for the more stable life satisfaction (Dienlin & Johannes, 2020). Whereas using social media can have instant effects on mood (Marciano et al., 2022), effects on life satisfaction often take longer to manifest. For example, because media use leads to actual changes in specific behaviors, which then in turn affect life satisfaction (Dienlin et al., 2017).

In this study, I hence analyze how changes in using social media during the last week 293 affected changes in positive and negative affect during the same week. In other words, if 294 people during the last week engaged in more COVID-19 related social media use than 295 usual, did they feel better or worse during that week than usual? For life satisfaction, I 296 implemented a longer interval. If people during the last week used COVID-19 related social 297 media more than they usually do, were they at the end of the week more or less satisfied 298 with their lives than they usually are? This way it is analyzed if when a person changes 299 their social media diet, are there (a) simultaneous changes in their affect and (b) 300 subsequent changes in their life satisfaction? For the main analyses, the interval is 301 implemented via the wording of the items (see below), not by using lagged measures 302 coming from prior waves waves. In additional analyses, I also tested how media use affects 303 well-being one month or four months later. All analyses will be controlled for varying 304 confounders (see below), which fosters a causal interpretation. 305

306 Statistical model

The hypothesis was analyzed using random effect within-between models (REWB, 307 Bell et al., 2019). Altogether three models were run, one for each dependent variable. The 308 data were hierarchical, and responses were separately nested in participants and waves 309 (i.e., participants and waves were implemented as random effects). Nesting in participants 310 accounts for the longitudinal design. Nesting in waves controls for general exogenous 311 developments, such as general decreases in well-being in the population, for example due to 312 lockdown measures. Thus, there was no need additionally to control for specific phases or 313 measures of the lockdown. Predictors were modeled as fixed effects. They included social 314 media communication types and channels, separated into within and between-person 315 factors, as well as stable and varying covariates. Between-person predictors are the predictors centered on the grand mean; within-person predictors are the predictors centered 317 on the person's mean. Between-person predictors (which, measuring relations, are not of 318 particular interest in this study) represent how the mean of one respondent differs from the 319 mean of all the other respondents. The within-person predictors represent how much a 320 person at one specific wave differs from their own mean. For example, we could find that on 321 Wave 3 a person used social media more than usual, while also experiencing more negative 322 affect than usual. All predictors were included simultaneously in each of the three models. 323 The factorial validity of the scales were tested with confirmatory factor analyses 324 (CFA). Because Mardia's test showed that the assumption of multivariate normality was 325 violated, I used the more robust Satorra-Bentler scaled and mean-adjusted test statistic 326 (MLM) as estimator. Mean scores were used for positive and negative affect. Missing 327 responses were imputed using multiple imputation with predictive mean matching (five 328 iterations, 30 data-sets), including categorical variables. All variables were imputed except 320 the social media use measures, as they were not collected on each wave. All variables 330 included in the analyses presented here were used to impute missing data. For the main 331 analyses, results were pooled across all thirty data-sets. 332

To contextualize the results, I conducted additional exploratory analyses. I reran 333 the analyses (a) with additional not-preregistered covariates such as trust in media or 334 government, (b) without covariates, (c) with single imputation, and (d) without 335 imputation. For more information on the analyses, a complete documentation of the 336 models and results, and all additional analyses, see companion website. 337

Measures 338

For the variables' means, range, and variance, see Table 1. For a complete list of all 339 items and item characteristics, see companion website. 340

Well-being

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341 Life satisfaction was measured with the item "All things considered, how satisfied 342 are you with your life as a whole nowadays?", which comes from the European Social 343 Survey (European Social Survey, 2021). The response options ranged from 0 (extremely 344 dissatisfied) to 10 (extremely satisfied). 345 To capture positive affect, respondents were asked how often in the last week they 346 felt (a) calm and relaxed, (b) happy, and (c) full of energy (World Health Organization, 347 1998). The response options were 1 (never), 2 (on some days), 3 (several times per week), 4 348 (almost every day), and 5 (daily). The scale showed good factorial fit, $\chi^2(66) = 69.42$, p =349 .363, CFI = 1.00, RMSEA < .01, 90% CI [< .01, .02], SRMR = .01. Reliability was high, ω 350 = .85.351 For negative affect, respondents were asked how often in the last week they felt (a) 352 lonely, (b) aggravated, (c) so depressed, that nothing could lift you up, (d) very nervous, 353 (e) anxious, and (h) glum and sad (World Health Organization, 1998). The response 354 options were 1 (never), 2 (on some days), 3 (several times per week), 4 (almost every day), 355 and 5 (daily). The scale showed good factorial fit, $\chi^2(471) = 4012.14$, p < .001, CFI = .98, RMSEA = .07, 90% CI [.07, .08], SRMR = .03. Reliability was high, $\omega = .91$. 357 All three variables were measured on each wave.

COVID-19 related social media use

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COVID-19 related social media use focused on communication types was measured 360 with the three dimensions of (a) reading, (b) liking and sharing, and (c) posting. The items 361 come from Wagner et al. (2018) and were adapted for the context of this study. The 362 general introductory question was "How often during the last week have you engaged in the 363 following activities on social media?". The three items were "Reading the posts of others 364 with content on the Coronavirus", "When seeing posts on the Coronavirus, I clicked 'like', 365 'share' or 'retweet'", "I myself wrote posts on the Coronavirus on social media." Answer 366 options were 1 (several times per day), 2 (daily), 3 (several times per week), 4 (weekly), 5 367 (never). The items were inverted for the analyses. 368 COVID-19 related social media use focused on channels was measured with five variables from Wagner et al. (2018), adapted for this study. The general introductory 370 question was "How often in the last week have you followed information related to the 371 Corona-crisis on the following social media?" The five items were (a) Facebook, (b) 372 Twitter, (c) Instagram, (d) Youtube, and (e) WhatsApp. Again, the answer options were 1 373 (several times per day), 2 (daily), 3 (several times per week), 4 (weekly), 5 (never). Again, 374 the items were inverted for the analyses. 375 Social media use was measured for all participants on waves 1, 2, 8, 17, 23, and 28 376 (see Figure 1). Freshly recruited respondents always answered all questions on COVID 377 19-related social media use. Because new respondents always provided data on media use. 378 it was possible to include these data into the analyses. Hence, for the main analyses data 379 from all 34 waves were used. In the additional analyses I tested longer intervals, namely if 380 changes in social media use were associated with changes in well-being either one month of 381 four months later. For these analyzes I used the predictors from waves 1, 2, 8, 17, 23, and 382 28, to see if they predicted changes in well-being either one month or four months later.

Table 1

Descriptives of the main variables.

	sd	min	max	mean
Well-being				
Life satisfaction	2.23	6.32	6.60	6.49
Positive affect	0.94	3.09	3.22	3.16
Negative affect	0.77	1.75	1.86	1.81
Social media use				
Read	1.38	1.92	2.88	2.35
Like & share	1.19	1.54	1.94	1.74
Posting	0.89	1.36	1.42	1.40
Social media channel				
Facebook	1.58	2.02	2.72	2.37
Twitter	0.95	1.36	1.43	1.39
Instagram	1.34	2.00	2.08	2.05
WhatsApp	1.66	2.27	2.60	2.46
YouTube	1.28	1.91	1.98	1.95

384 Control variables

The effects of COVID-19 related social media use were controlled for the following stable variables: gender (female, male, diverse), age, education (ten options), Austria country of birth (yes/no), Austria parents' country of birth (no parent, one parent, both parents), and household size. I also controlled for the following varying covariates: five items on current living conditions, including self-reported physical health, whether participants contracted COVID-19 since the last wave, current household income, working in home-office, and overall work hours; nine items measuring use of specific national text-based and video-based news outlets; five items measuring outdoor activities such as exercise or meeting friends; and two more psychological measures including locus of control and disposition to take risks.

Results

Descriptive Analyses

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Looking at the variables from a descriptive perspective, aligned with set-point 397 theory we can see that the level of all well-being measures were surprisingly stable during 398 data collection (see Figure 3). COVID-19 related social media use, however, showed 399 changes. Reading, sharing and liking COVID-19 related content decreased substantially 400 (almost one scale point from 3 to 2). Posting about COVID-19 related content stayed the 401 same. Using Facebook and WhatsApp for COVID-19 related content decreased. 402 Instagram, YouTube, and Twitter stayed the same. The general initial decrease could be 403 explained by the fact that the collection of data began at the end of March 2020, hence approximately three months after the pandemic's onset. After an initial uptick, COVID-19 related social media use might have already been declining at the time.

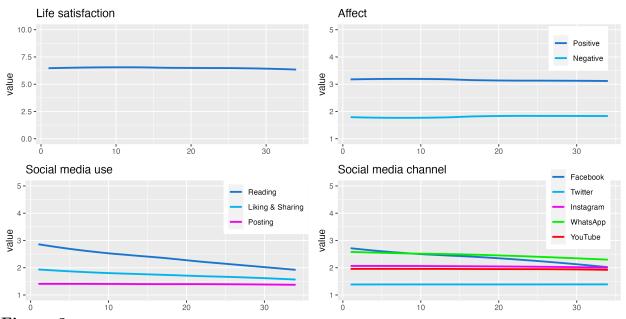


Figure 3

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Well-being and media use across the 34 waves. Note. Values obtained from mixed effect models, with participants and waves as grouping factors and without additional predictors.

Using the average values across all waves, which provides a stable picture of the

general relations, I next looked at the correlations between social media use and well-being 408 (see Figure 4). Several interesting patterns emerged. In general, people who spend more 409 time engaging with COVID-19 related content on social media reported reduced well-being. 410 Users who spend more time reading, liking and sharing, and posting COVID-19 related 411 content were less satisfied with their lives. They also showed slightly less positive affect. 412 This overall negative picture was even more pronounced for negative affect. People who 413 engaged more with COVID-19 related content, including all types and channels of 414 communication, reported substantially higher levels of negative affect. For example, people 415 who were more likely to post COVID-19 content had much higher levels of negative affect 416 (r = .61). Note that these results represent between-person correlations, not causal 417 within-person effects. 418

Preregistered Analyses 419

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Social media communication types 420

The study's main hypothesis was that the causal effects of all types and channels of social media use on all facets of well-being would be trivial. Regarding the effects of 422 different communication types (i.e., reading, sharing, of posting about COVID-19 related 423 content), all within-person effects fell completely within the a-priori defined null region (see 424 Figure 5). For example, respondents who used social media more frequently than usual to 425 like or share COVID-19 related content did not show a simultaneous change in life 426 satisfaction (b = -0.02 [95% CI -0.06, 0.01]). As a result, the hypothesis of trivial effects 427 was supported for all COVID-19 related types of social media communication. 428 However, several effects stood out, as statistically they were significantly different 429 from zero. Users who read more COVID-19 related content than usual reported slightly 430 reduced levels of positive affect (b = -0.03 [95% CI -0.05, -0.02]). Users who liked and 431 shared more COVID-19 related content than usual also experienced slightly more negative affect than usual (b = 0.05 [95% CI 0.04, 0.07]). Posting COVID-19 related content 433 affected all types of well-being. Users who wrote more COVID-19 related posts than usual

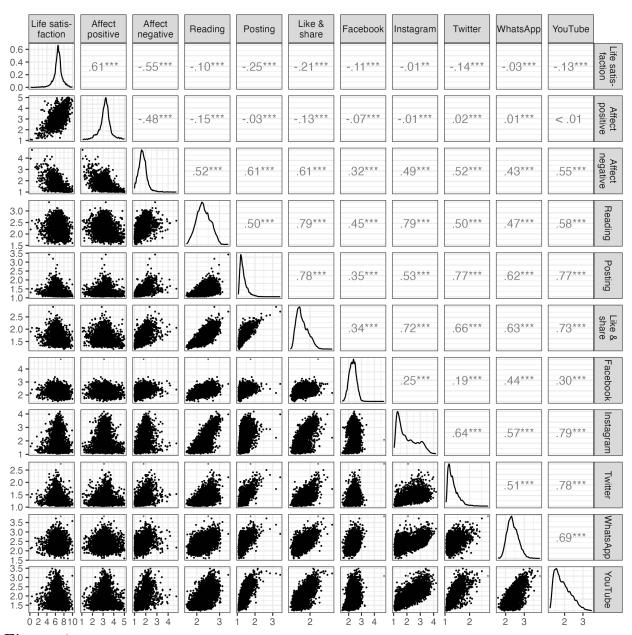


Figure 4

Descriptives of the main variables, capturing well-being and social media use with their average values across all waves. Upper triangle: correlation coefficients; diagonal: density plots; lower triangle: scatter plots.

also reported slightly less life satisfaction than usual (b = -0.04 [95% CI -0.08, -0.01]) and slightly more negative affect than usual (b = 0.05 [95% CI 0.04, 0.07]). Interestingly, however, users who wrote more COVID-19 related posts than usual also experienced slightly *higher* levels of positive affect than usual (b = 0.02 [95% CI 0.01, 0.04]).

439 Social media communication channels

Regarding the COVID-19 related use of social media channels (i.e., Facebook,
Instagram, WhatsApp, YouTube, and Twitter) the results were comparable (see Figure 5).
Changes in the frequency of using different social media channels to attain information
regarding COVID-19 were unrelated to meaningful changes in well-being. For example,
respondents who used Facebook more frequently than usual to learn about COVID-19 did
not show a simultaneous change in life satisfaction (b -0.01 [95% CI -0.04, 0.02]). In sum,
the hypothesis of trivial effects was supported also for the COVID-19 related use of
important social media channels.

That said, two effects differed statistically from zero. Respondents who used Twitter more frequently than usual to attain COVID-19 related news reported slightly higher levels of negative affect than usual (b = 0.02 [95% CI 0.01, 0.04]). Likewise, respondents who used YouTube more frequently than usual for COVID-19 related issues reported slightly higher levels of negative affect than usual (b = 0.01 [95% CI < 0.01, 0.02]). However, both effects were still completely inside of the null region, hence likely not large enough to be considered meaningful.

For an overview of all within-person effects, see Table 2 and Figure 5.

Exploratory Analyses

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To contextualize the results reported above and to see if the study included any meaningful effects at all, I also looked at the effect sizes of the covariates. Because each variable featured different response options, which would require defining a SESOI for each variable, I hence report the results of the standardized scales, which allows for a better comparison across the differently scaled variables. Here, we can build on Cohen's

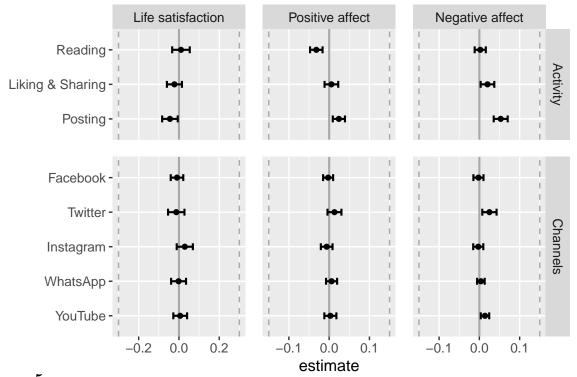


Figure 5

Unstandardized within-person effects of COVID-19 related social media use on well-being.

Note. The SESOI was b = |0.30| for life satisfaction and b = |0.15| for affect. Hence, all of

the reported effects are not considered large enough to be meaningful.

convention that small effects begin at r = |.10|.

The results showed that several effects crossed or fell completely outside of the 463 SESOI, and can hence be considered meaningful. For example, if physical health decreased, this had a meaningful detrimental impact on life satisfaction ($\beta = .19$ [95% CI .18, .20]), 465 positive affect ($\beta = .18$ [95% CI .17, .19]), and negative affect ($\beta = -.19$ [95% CI -.20, -.18]). Spending more time outside to exercise meaningfully increased positive affect ($\beta = .12$ [95% CI .11, .14]). The strongest aspect affecting well-being was internal locus of control. 468 If people felt more in control of their lives, this strongly increased both life satisfaction (β 469 = .33 [95% CI .31, .35]) and positive affect (β = .28 [95% CI .27, .30]), while decreasing 470 negative affect ($\beta = -.29$ [95% CI -.31, -.27]). For an overview, see Figure 6. 471 Because life satisfaction is more stable than affect, the effects of communication 472

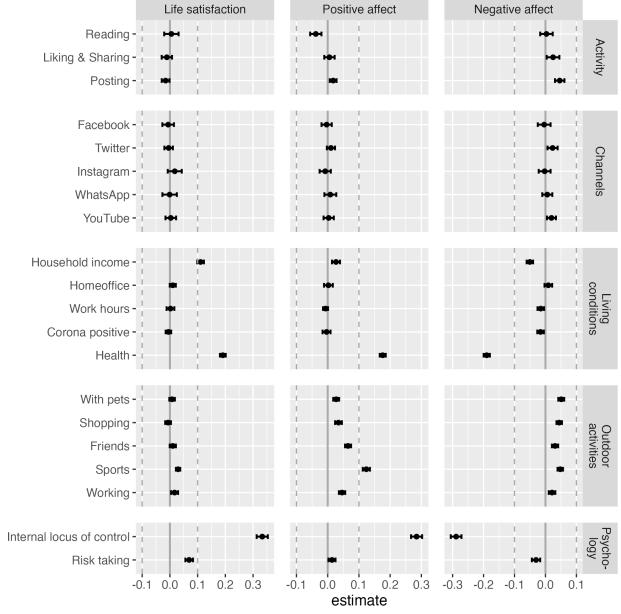


Figure 6

Results of main variables together with covariates to provide context. All variables standardized. SESOI: beta = |.10|

Table 2

Overview of all within-person effects.

		Confide	nce interval		
Predictor	b	Lower	Higher	beta	p
Life satisfaction					
Reading	0.01	-0.03	0.05	0.01	.639
Liking & Sharing	-0.02	-0.06	0.01	-0.01	.227
Posting	-0.04	-0.08	-0.01	-0.02	.025
Facebook	-0.01	-0.04	0.02	-0.01	.527
Instagram	0.03	-0.01	0.07	0.02	.149
WhatsApp	0.00	-0.04	0.04	0.00	.917
YouTube	0.01	-0.03	0.04	0.00	.713
Twitter	-0.01	-0.05	0.03	0.00	.503
Positive affect					
Reading	-0.03	-0.05	-0.02	-0.04	< .001
Liking & Sharing	0.01	-0.01	0.02	0.01	.508
Posting	0.02	0.01	0.04	0.02	.003
Facebook	0.00	-0.02	0.01	0.00	.671
Instagram	-0.01	-0.02	0.01	-0.01	.390
WhatsApp	0.01	-0.01	0.02	0.01	.374
YouTube	0.00	-0.01	0.02	0.00	.686
Twitter	0.01	0.00	0.03	0.01	.130
Negative affect					
Reading	0.00	-0.01	0.02	0.00	.747
Liking & Sharing	0.02	0.00	0.04	0.02	.022
Posting	0.05	0.04	0.07	0.05	< .001
Facebook	0.00	-0.01	0.01	0.00	.710
Instagram	0.00	-0.02	0.01	0.00	.654
WhatsApp	0.00	-0.01	0.01	0.01	.417
YouTube	0.01	0.00	0.02	0.02	.011
Twitter	0.02	0.01	0.04	0.02	.008

might materialize some time later. I hence also tested the effects across the longer intervals
of one month and four months. Results showed that all effects disappeared. No effect
remained significant, implying that at least in this case in this case effects take place on a
shorter interval.

Finally, as suggested by the differential susceptibility of media effects model, media
effects can depend on dispositional factors, developmental stages, or cultural norms
(Valkenburg & Peter, 2013), such as gender and age (Orben et al., 2022). I hence reran the
analyses, differentiating effects for boys and girls and for age cohorts. The results showed

that effects did not differ across genders. The effects also did not depend on age. However, one effect stood out and was significant. Compared to the middle age category Generation X, results showed that if users from Generation Z posted more COVID-19 content than usual this lead to significantly more negative affect ($\beta = .04$ [95% CI .01, .06]).

485 Discussion

Based on a panel study with 34 waves largely representative of the Austrian 486 population, this study analyzed the effects of COVID-19 related social media use on 487 well-being. Between person correlation analyses showed that more active users of 488 COVID-19 related content on social media also reported decreased well-being. For 489 example, respondents who read more COVID-19 related content than others reported slightly lower levels of life satisfaction, somewhat lower levels of positive affect, and substantially higher levels of negative affect than others. To see if these between person 492 correlations would translate to within-person effects, I analyzed if changes in a person's 493 media use led to changes in their well-being. The within-person relations showed a different 494 pattern. If people consumed more COVID-19 content on social media than usual, this did 495 not meaningfully reduce their well-being. Although several statistically significant effects 496 were found, these were very small. For example, people who read more COVID-19 related 497 posts than usual reported slightly decreased positive affect. People who liked and shared 498 more COVID-19 related posts than usual reported slightly higher levels of negative affect. 490 Posting more content about COVID-19 than usual slightly decreased life satisfaction, while 500 increasing both negative affect and positive affect. Using Twitter for COVID-19 related 501 content slightly increased negative affect, as did YouTube. Again, although all of these 502 within-person effects were statistically significant, they were very small, smaller than the 503 predefined smallest effect size of interest. According to the preregistered procedure, they 504 should hence be considered irrelevant. Additional analyses revealed that other factors, for 505 which we would expect to find meaningful effects, such as health or sports, indeed showed substantial and meaningful impacts on well-being. In addition, when testing for the longer 507

intervals of one month and four months, again no meaningful effects were found. In conclusion, COVID-19 related activity on social media was not a particularly strong influence on peoples' well-being. The results do not support the popular fears that "doomscrolling" or overusing social media during times of crises constitutes a prominent risk for well-being.

These specific observations notwithstanding, several general trends can be observed. 513 First, overall the results do suggest that effects of COVID-19 related social media use on 514 well-being tend to take place in the negative as opposed to the positive spectrum. 515 Although very small, five statistically significant negative results of COVID-19 related 516 social media use on well-being were found. Only one positive effect emerged. Also note 517 that in the analyses several control variables were included, ruling out plausible alternative 518 explanations for the negative results. For example, it was controlled for as to whether or not participants contracted a COVID-19 infection during a specific wave. Hence, we can rule out the alternative explanation that having an infection was the root cause of increased communication and reduced well-being. 522

Second, six significant outcomes emerged for positive or negative affect, but only 523 one for life satisfaction. Life satisfaction is more stable and not that easily affected by any type or channel of social media communication. The more fluctuating positive and negative 525 affect, however, were affected (albeit only slightly). Liking, sharing, and posting COVID-19 526 related content, and spending more time on Twitter and YouTube to browse COVID-19 527 related content, all slightly negatively influenced affect. This is aligned with prior findings 528 which showed that social media use can trigger negative affect, but that it is less likely to 520 determine life satisfaction (Huang, 2017). Conversations about COVID-19 on social media 530 are often extreme, negative, or aggressive (L. Fan et al., 2020). More deeply engaging with 531 this type of content could negatively affect active authors. The hypothesis that tonality 532 could explain the negative effects is especially supported by the observation that spending 533 more time on Twitter and YouTube than usual increased negative affect. Communication 534

on both channels is often found to be negative and impolite (e.g., Mueller & Saeltzer, 2022), also when compared to other SNSs (Halpern & Gibbs, 2013). Consuming more negative and misleading information could hence explain the (slightly) increased levels of negative affect.

Third, the results show that it makes sense to analyze different communication 539 types and communication channels. Reading slightly reduced positive affect, while liking, 540 sharing, and posting slightly increased negative affect. Interestingly, posting COVID-19 541 related comment slightly increased negative affect, while at the same time it also slightly 542 increased positive affect. Posting content is often met with strong reaction, both positive 543 by means of likes and negative by means of critical comments. Overall, though, posting led 544 to slightly reduced levels of life satisfaction. In conclusion, whereas it was often stated that 545 passive use is bad and active use good (Verduyn et al., 2015), this pattern was only partially found here. The results are aligned with the findings from Valkenburg et al. (2022), who could not confirm that active use is good and that passive use is bad. Focusing on communication channels, Twitter and YouTube seem to be more negative, as has often been observed (Halpern & Gibbs, 2013), while Instagram, WhatsApp, and Facebook were 550 neutral. But, again, all of these effects are very small. Future research might elaborate on 551 these specific relations to probe their stability and relevance. 552

Taken together, the results are hence aligned with the underlying theoretical models 553 and prior empirical results. The findings support the differential susceptibility of media 554 effects model (Valkenburg & Peter, 2013), such that effects are generally small and that 555 they depend on the type and channel of communication. Additional analyses did not reveal 556 that effects depended on gender. Age also large did not play a significant moderation role, 557 but effects of posting COVID-19 related content were found to be more negative for 558 Generation Z. Indeed, it has often been argued that effects of social media use are more 550 negative for Gen Z than for prior generations, and this finding can be seen a further 560 tentative support for this hypothesis. From a broader perspective, the results are 561

well-aligned with mood management theory (Zillmann, 1988) and the uses and 562 gratifications approach (Katz et al., 1973), whose premises preclude particularly negative 563 effects of routine and widespread media consumption. Both theories posit that if the effects 564 of social media were indeed profoundly negative on average, then people likely would not 565 spend so much time on social media engaging with COVID-19 content. Finally, recent 566 empirical studies and meta-analyses reported rather small negative effects, too. Several 567 studies found that the effects of various types of social media use on well-being are small, 568 often too small to matter (Bendau et al., 2021; Ferguson et al., 2021; Meier & Reinecke, 569 2020; Orben, 2020), echoing the results obtained here. 570

71 Limitations

Focusing on within-person effects and controlling for several potential confounders, 572 this study provides an improved perspective on assessing causality. However, several 573 challenges remain. In order to correctly establish causality in non-experimental designs, it 574 is necessary to control for all relevant confounding third variables (Rohrer, 2018). 575 Although this study included are large list of confounders, it could still be that crucial 576 variables were missed. More thought needs to be invested in which factors to control for 577 and, equally important, for which factors not to control for. I hope this study provides a 578 first step into this direction. 579

Although I had already reduced the predefined SESOIs to be less conservative, one could argue they were still too large. Media use is only one aspect of several factors that simultaneously affect well-being. Is it realistic to expect that changing only *one* of these aspects should already manifest in a detectable change in well-being? Or would it make more sense to expect that thoroughly committing to say *two* activities (e.g. regularly exercising *and* establishing a reading habit) should then cause a detectable improvement in well-being? Practically, this would imply a SESOI half the size defined here, namely b = |.15| for life satisfaction and b = |.075| for affect. In the case of this study, however, even halving the SESOI would not make a difference. All but one effect would still be completely

in the null region, and no effect would fall completely outside of the null region. I encourage future research to elaborate on what effect sizes are considered meaningful and what not.

Both media use and well-being were measured using self-reports. Because assessing well-being necessarily requires introspection, using self-reports for affect and life satisfaction is adequate. However, for social media use objective measures are clearly preferable, as people often cannot reliably estimate their use (Scharkow, 2016). At the same time, most objective measures cannot capture the content or the motivation of use. Hence, for the type of research question analyzed here, it still seems necessary to use self-reported measures. In many cases they can still be informative (Verbeij et al., 2021).

Being collected in a single country, the generalizability of the results is limited. The results apply primarily to the more Western sphere. They might not hold true in other cultures, especially cultures with a different media landscape or alternative social media channels.

Conclusion

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In this study, COVID-19 related social media use did not meaningfully affect 603 well-being. Very small negative effects were found for writing COVID-19 related posts, 604 sharing COVID-19 related content, and spending more time than usual on Twitter. Factors 605 other than social media use, however, were meaningfully related to well-being, including 606 physical health, exercise, satisfaction with democracy, or believing that one is in control of 607 one's life. In light of the overall very small effects, engaging in COVID 19-related social 608 media use should not be considered a major concern for one's well-being. Hence, when 609 trying to improve well-being during a pandemic, instead of focusing on social media it 610 seems more fruitful to address other, more pertinent societal problems related to health 611 care, regular exercise, or psychological resilience. 612

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Competing Interests

I declare no competing interests.

Supplementary Material

- All the stimuli, presentation materials, analysis scripts, and a reproducible version
- 813 of the manuscript can be found on the companion website
- 814 (https://XMtRA.github.io/Austrian_Corona_Panel_Project).

Data Accessibility Statement

The data are shared on AUSSDA, see https://doi.org/10.11587/28KQNS. The data

can only be used for scientific purposes.

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