**Version info**: Code for this page was tested in SAS 9.3.

Multinomial logistic regression is for modeling nominal outcome variables, in which the log odds of the outcomes are modeled as a linear combination of the predictor variables.

Please Note: The purpose of this page is to show how to use various data analysis commands.

It does not cover all aspects of the research process which researchers are expected to do. In

particular, it does not cover data cleaning and checking, verification of assumptions, model

diagnostics and potential follow-up analyses.

### Examples of multinomial logistic regression

Example 1. People's occupational choices might be influenced by their parents' occupations and their own education level. We can study the relationship of one's occupation choice with education level and father's occupation. The occupational choices will be the outcome variable which consists of categories of occupations.

Example 2. A biologist may be interested in food choices that alligators make. Adult alligators might have difference preference than young ones. The outcome variable here will be the types of food, and the predictor variables might be the length of the alligators and other environmental variables.

Example 3. Entering high school students make program choices among general program, vocational program and academic program. Their choice might be modeled using their writing score and their social economic status.

### Description of the data

For our data analysis example, we will expand the third example using the **hsbdemo** data set. You can download the data <a href="https://stats.idre.ucla.edu/wp-content/uploads/2016/02/hsbdemo.sas7bdat">https://stats.idre.ucla.edu/wp-content/uploads/2016/02/hsbdemo.sas7bdat</a>).

```
proc contents data = "c:\hsbdemo";
run;
```

The CONTENTS Procedure

Data Set Name c:\datahsbdemo Observations 200
Member Type DATA Variables 13

Engine	V9				Indexes	0	
Created	Thursday, August 2	29, 2013	09:42:59	AM	Observation Length	40	
Last Modified	Thursday, August 2	29, 2013	09:42:59	AM	Deleted Observations	0	
Protection					Compressed	NO	
Data Set Type					Sorted	YES	>

Label Written by SAS

Data Representation WINDOWS\_64

Encoding wlatin1 Western (Windows)

# Engine/Host Dependent Information

Data Set Page Size	4096
Number of Data Set Pages	3
First Data Page	1
Max Obs per Page	101
Obs in First Data Page	42
Number of Data Set Repairs	0

Filename c:\datahsbdemo.sas7bdat

Release Created 9.0301M1 Host Created X64\_7PRO

# Alphabetic List of Variables and Attributes

#	Variable	Туре	Len	Label
12	AWARDS	Num	3	
13	CID	Num	3	
2	FEMALE	Num	3	
11	HONORS	Num	3	honores eng
1	ID	Num	4	
8	MATH	Num	3	math score
5	PROG	Num	3	type of program
6	READ	Num	3	reading score
4	SCHTYP	Num	3	type of school
9	SCIENCE	Num	3	science score
3	SES	Num	3	
10	SOCST	Num	3	social studies score
7	WRITE	Num	3	writing score

```
Sort Information

Sortedby PROG

Validated YES

Character Set ANSI
```

The data set contains variables on 200 students. The outcome variable is **prog**, program type. The predictor variables are social economic status, **ses**, a three-level categorical variable and writing score, **write**, a continuous variable. Let's start with getting some descriptive statistics of the variables of interest.

```
proc freq data = "c:\hsbdemo";
tables prog*ses / chisq norow nocol nofreq;
run;
```

```
The FREQ Procedure
Table of PROG by SES
PROG(type of program) SES
Percent | 1| 2| 3| Total
-----+
    1 | 8.00 | 10.00 | 4.50 | 22.50
-----+
    2 | 9.50 | 22.00 | 21.00 | 52.50
-----+
    3 | 6.00 | 15.50 | 3.50 | 25.00
-----+
       47 95
Total
                       58
                             200
        23.50 47.50 29.00
                           100.00
Statistics for Table of PROG by SES
Statistic
                      DF
                           Value
                                     Prob
Chi-Square
                      4
                                  0.0023
                           16.6044
Likelihood Ratio Chi-Square
                      4
                           16.7830
                                   0.0021
Mantel-Haenszel Chi-Square 1
                            0.0598
                                  0.8068
Phi Coefficient
                            0.2881
Contingency Coefficient
                            0.2769
Cramer's V
                            0.2037
Sample Size = 200
proc sort data = "c:\hsbdemo";
by prog;
run;
```

```
proc means data = "c:\hsbdemo";
var write;
by prog;
run;
type of program=1
The MEANS Procedure
              Analysis Variable : WRITE writing score
  Ν
                            Std Dev
                                             Minimum
               Mean
                                                             Maximum
                         9.3977754
 45
         51.3333333
                                          31.0000000
                                                          67.0000000
type of program=2
              Analysis Variable : WRITE writing score
               Mean
                            Std Dev
                                            Minimum
                                                             Maximum
         56.2571429
                          7.9433433
                                          33.0000000
type of program=3
              Analysis Variable : WRITE writing score
                            Std Dev
  Ν
               Mean
                                            Minimum
                                                             Maximum
 50
                          9.3187544
                                          31.0000000
```

# Analysis methods you might consider

- Multinomial logistic regression: the focus of this page.
- Multinomial probit regression: similar to multinomial logistic regression but with independent normal

error terms.

- Multiple-group discriminant function analysis: A multivariate method for multinomial outcome variables
- Multiple logistic regression analyses, one for each pair of outcomes: One problem with this approach is
  that each analysis is potentially run on a different sample. The other problem is that without
  constraining the logistic models, we can end up with the probability of choosing all possible outcome
  categories greater than 1.
- Collapsing number of categories to two and then doing a logistic regression: This approach suffers from loss of information and changes the original research questions to very different ones.
- Ordinal logistic regression: If the outcome variable is truly ordered and if it also satisfies the assumption
  of proportional odds, then switching to ordinal logistic regression will make the model more
  parsimonious.
- Alternative-specific multinomial probit regression: allows different error structures therefore allows to relax the independence of irrelevant alternatives (IIA, see below "Things to Consider") assumption. This requires that the data structure be choice-specific.
- Nested logit model: also relaxes the IIA assumption, also requires the data structure be choice-specific.

## Multinomial logistic regression

Below we use **proc logistic** to estimate a multinomial logistic regression model. The outcome **prog** and the predictor **ses** are both categorical variables and should be indicated as such on the **class** statement. We can specify the baseline category for **prog** using (ref = "2") and the reference group for **ses** using (ref = "1"). The **param=ref** option on the **class** statement tells SAS to use dummy coding rather than effect coding for the variable **ses**. Note that the levels of **prog** are defined as:

1=general

2=academic (reference group)

3=vocational

Data Set c:\datahsbdemo Written by SAS
Response Variable PROG type of program
Number of Response Levels 3
Model generalized logit
Optimization Technique Newton-Raphson

Number of Observations Read 200
Number of Observations Used 200

#### Response Profile

Total		Ordered
Frequency	PROG	Value
45	1	1
105	2	2
50	3	3

Logits modeled use PROG=2 as the reference category.

#### Class Level Information

		Design			
Class	Value	Varia	bles		
SES	1	0	0		
	2	1	0		
	3	0	1		

#### Model Convergence Status

Convergence criterion (GCONV=1E-8) satisfied.

#### Model Fit Statistics

		Intercept
	Intercept	and
Criterion	Only	Covariates
AIC	412.193	375.963
SC	418.790	402.350
-2 Log L	408.193	359.963

Testing Global Null Hypothesis: BETA=0

Test Chi-Square DF Pr > ChiSq

Likelihood Ratio	48.2299	6	<.0001
Score	45.1588	6	<.0001
Wald	37.2946	6	<.0001

Type 3 Analysis of Effects

		Wald	
Effect	DF	Chi-Square	Pr > ChiSq
SES	4	10.8162	0.0287
WRITE	2	26.4633	<.0001

# Analysis of Maximum Likelihood Estimates

					Standard	Wald	
Parameter		PROG	DF	Estimate	Error	Chi-Square	Pr > ChiSq
Intercept		1	1	2.8522	1.1664	5.9790	0.0145
Intercept		3	1	5.2182	1.1635	20.1128	<.0001
SES	2	1	1	-0.5333	0.4437	1.4444	0.2294
SES	2	3	1	0.2914	0.4764	0.3742	0.5407
SES	3	1	1	-1.1628	0.5142	5.1137	0.0237
SES	3	3	1	-0.9827	0.5956	2.7224	0.0989
WRITE		1	1	-0.0579	0.0214	7.3200	0.0068
WRITE		3	1	-0.1136	0.0222	26.1392	<.0001

#### Odds Ratio Estimates

Effect	Ξ			PROG	Point Estimate	95% Wa Confidence	
SES	2	vs	1	1	0.587	0.246	1.400
SES	2	vs	1	3	1.338	0.526	3.404
SES	3	vs	1	1	0.313	0.114	0.856
SES	3	vs	1	3	0.374	0.116	1.203
WRITE				1	0.944	0.905	0.984
WRITE				3	0.893	0.855	0.932

- In the output above, the likelihood ratio chi-square of48.23 with a p-value < 0.0001 tells us that our model as a whole fits significantly better than an empty model (i.e., a model with no predictors)
  - Several model fit measures such as the AIC are listed under Model Fit Statistics

Two models are tested in this multinomial regression, one comparing membership to general versus academic program and one comparing membership to vocational versus academic program. They correspond to the two equations below:

$$ln\left(rac{P(prog=general)}{P(prog=academic)}
ight) = b_{10} + b_{11}(ses=2) + b_{12}(ses=3) + b_{13}write$$

$$ln\left(rac{P(prog=vocation)}{P(prog=academic)}
ight) = b_{20} + b_{21}(ses=2) + b_{22}(ses=3) + b_{23}write$$

where bs are the regression coefficients.

- A one-unit increase in the variable write is associated with a .058 decrease in the relative log odds of being in general program vs. academic program.
- A one-unit increase in the variable write is associated with a .1136 decrease in the relative log odds of being in vocation program vs. academic program.
- The relative log odds of being in general program vs. in academic program will decrease by 1.163 if moving from the lowest level of ses (ses==1) to the highest level of ses (ses==3).
- The overall effects of ses and write are listed under "Type 3 Analysis of Effects", and both are significant.
- The ratio of the probability of choosing one outcome category over the probability of
  choosing the baseline category is often referred to as relative risk (and it is also sometimes
  referred to as odds as we have just used to described the regression parameters above).
   Relative risk can be obtained by exponentiating the linear equations above, yielding
  regression coefficients that are relative risk ratios for a unit change in the predictor variable.
   In the case of two categories, relative risk ratios are equivalent to odds ratios, which are
  listed in the output as well.

- The odds ratio for a one-unit increase in the variable **write** is .944 (exp(-.0579) from the regression coefficients above the odds ratios) for being in general program vs. academic program.
- The odds ratio of switching from **ses** = 1 to 3 is .313 for being in general program vs. academic program. In other words, the expected risk of staying in the general program is lower for subjects who are high in **ses**.

Using the **test** statement, we can also test specific hypotheses within or even across logits, such as if the effect of **ses=3** in predicting general versus academic equals the effect of **ses=3** in predicting vocational versus academic. Use of the **test** statement requires the unique names SAS assigns each parameter in the model. The option **outest** on the **proc logistic** statement produces an output dataset with the parameter names and values. We can get these names by printing them, and we transpose them to be more readable. The **noobs** option on the **proc print** statement suppresses observation numbers, since they are meaningless in the parameter dataset.

```
proc logistic data = "c:\hsbdemo" outest = mlogit_param;
class prog (ref = "2") ses (ref = "1") / param = ref;
model prog = ses write / link = glogit;
```

```
run;
proc transpose data = mlogit_param;
run;
proc print noobs;
run;
NAME
               LABEL
                                             PROG
Intercept 1
               Intercept: PROG=1
                                            2.852
Intercept 3
               Intercept: PROG=3
                                            5.218
SES2 1
               SES 2: PROG=1
                                           -0.533
SES2 3
               SES 2: PROG=3
                                            0.291
               SES 3: PROG=1
SES3 1
                                           -1.163
SES3 3
               SES 3: PROG=3
                                           -0.983
               writing score: PROG=1
WRITE 1
                                           -0.058
               writing score: PROG=3
WRITE_3
                                           -0.114
               Model Log Likelihood
_LNLIKE_
                                         -179.982
```

Here we see the same parameters as in the output above, but with their unique SAS-given names. We are interested in testing whether SES3\_general is equal to SES3\_vocational, which we can now do with the test statement. The code preceding the ":" on the test statement is a label identifying the test in the output, and it must conform to SAS variable-naming rules (i.e., 32 characters in length or less, letters, numerals, and underscore).

```
proc logistic data = "c:\hsbdemo" outest = mlogit_param;
class prog (ref = "2") ses (ref = "1") / param = ref;
model prog = ses write / link = glogit;
second are second ar
```

```
run;

***SOME OUTPUT OMITTED***

Linear Hypotheses Testing Results

Wald

Label Chi-Square DF Pr > ChiSq

SES3_general_vs_SES3_vocational 0.0772 1 0.7811
```

The effect of **ses=3** for predicting general versus academic is not different from the effect of **ses=3** for predicting vocational versus academic.

You can also use predicted probabilities to help you understand the model. You can calculate predicted probabilities using the **Ismeans** statement and the **ilink** option. For multinomial data, **Ismeans** requires **glm** rather than **reference** (dummy) coding, even though they are essentially the same, so be sure to respecify the coding on the **class** statement. However, **glm** coding only allows the last category to be the reference group (**prog** = vocational and **ses** = 3)and will ignore any other reference group specifications. Below we use **Ismeans** to calculate the predicted probability of choosing program type academic or general at each level of **ses**, holding **write** at its means.

```
proc logistic data = "c:\hsbdemo" outest = mlogit_param;
class prog ses / param = glm;
model prog = ses write / link = glogit;
lamana and / o ilink all;
```

2

nun•									
run;									
***SOME C	OUTPUT (	MITTED***							
***SOME OUTPUT OMITTED***  Coefficients for SES Least Squares Means  type of  Parameter program SES Row1 Row2 Row3 Row4 Row5  Intercept 1 1 1 1 1  Intercept 2 1 1 1  SES 1 1 1 1 1  SES 1 2 1 1 1  SES 2 1 2 1  SES 2 2 2 2 1  SES 3 1 3 1 3  SES 3 2 3  writing score 1 52.775 52.775  ***SOME OUTPUT OMITTED***									
		type of							
Danamotor	,	_	SES	Row1	Row2	Row3	Row/	Row5	Row6
r ar ameter		pi ogi alli	JLJ	KOWI	NOWZ	NOWS	NOW4	ROWS	KOWO
Intercept		1		1	1	1			
Intercept	:	2					1	1	1
SES 1		1	1	1					
SES 1		2	1				1		
SES 2		1	2		1				
SES 2		2	2					1	
SES 3		1	3			1			
SES 3		2	3						1
writing s	core	1		52.775	52.775	52.775			
writing s	core	2					52.775	52.775	52.775
***COME C	NITOLIT (	MTTTED***							
SOME C	JUIPUI (	וובוייים ווברייים							
		SES Lea	st Squa	res Means	5				
			Sta	ndard					
type of			Erro	or of	Lower	Upper			
program	SES	Mean		Mean	Mean	Mean			
1	1	0.3582	0.0	<b>27264</b>	0.2158	0.5006			
1	2	0.2283	0.0	<b>04512</b>	0.1399	0.3168			
1	3	0.1785	0.0	05405	0.07256	0.2844			
2	1	0.4397	0.0	<b>27799</b>	0.2868	0.5925			
2	2	0.4777	0.0	ð5526	0.3694	0.5861			

The predicted probabilities are in the "Mean" column. Thus, for ses = 3 and write = 52.775, we see that the probability of being the academic program (program type 2) is 0.7009; for the general program (program type 1), the probability is 0.1785.

0.5709

0.8309

0.06630

3

0.7009

using the **descending** option on the **proc logistic** statement. This will make academic the reference group for **prog** and 3 the reference group for **ses**.

```
proc logistic data = "c:\hsbdemo" outest = mlogit_param descending;
class prog ses / param = glm;
model prog = ses write / link = glogit;
```

run;									
i uii,									
***SOME (	OUTPUT (	OMITTED***							
			efficien	ts for S	SES Least So	uares Mean:	5		
		type of							
Parameter	•	program	SES	Row1	Row2	Row3	Row4	Row5	Rowe
Intercept	Ξ.	3		1	1	1			
Intercept	-	2					1	1	1
SES 1		3	1	1					
SES 1		2	1				1		
SES 2		3	2		1				
SES 2		2	2					1	
SES 3		3	3			1			
SES 3		2	3						1
writing s	score	3		52.775	52.775	52.775			
writing s	score	2					52.775	52.775	52.775
***SOME (	OUTPUT (	OMITTED***							
		SES Lea	st Squar	es Means	;				
			Stan	dand					
tuno of				r of	Louon	Unnon			
type of	CEC	Moan			Lower	Upper			
program	SES	Mean		Mean	Mean	Mean			
3	1	0.2021	0.0	5996	0.08459	0.3197			
3	2	0.2939	0.0	5036	0.1952	0.3926			
3	3	0.1206	0.0	4643	0.02960	0.2116			
2	1	0.4397	0.0	7799	0.2868	0.5925			
2	2	0.4777		5526	0.3694	0.5861			
2	3	0.7009	0.0	6630	0.5709	0.8309			

Here we see the probability of being in the vocational program when ses = 3 and write = 52.775 is 0.1206, which is what we would have expected since (1 - 0.7009 - 0.1785) = 0.1206, where 0.7009 and 0.1785 are the probabilities of being in the academic and general programs under the same conditions.

Things to consider

#### HIIIIYS TO CONSIDER

- The Independence of Irrelevant Alternatives (IIA) assumption: Roughly, the IIA assumption
  means that adding or deleting alternative outcome categories does not affect the odds
  among the remaining outcomes.
- Diagnostics and model fit: Unlike logistic regression where there are many statistics for performing model diagnostics, it is not as straightforward to do diagnostics with multinomial logistic regression models. Some model fit statistics are listed in the output.
- Pseudo-R-Squared: The R-squared offered in the output is basically the change in terms of log-likelihood from the intercept-only model to the current model. It does not convey the same information as the R-square for linear regression, even though it is still "the higher, the better".
- Sample size: Multinomial regression uses a maximum likelihood estimation method.

  Therefore, it requires a large sample size. It also uses multiple equations. Therefore, it requires an even larger sample size than ordinal or binary logistic regression.
- Complete or quasi-complete separation: Complete separation implies that only one value of a predictor variable is associated with only one value of the response variable. You can tell from the output of the regression coefficients that something is wrong. You can then do a two-way tabulation of the outcome variable with the problematic variable to confirm this and then rerun the model without the problematic variable.
- Empty cells or small cells: You should check for empty or small cells by doing a crosstab between categorical predictors and the outcome variable. If a cell has very few cases (a small cell), the model may become unstable or it might not run at all.
- Sometimes observations are clustered into groups (e.g., people within families, students within classrooms). In such cases, you may want to see our page on <u>non-independence</u> within clusters (/stata/dae/non-independence-within-clusters/).

#### See Also

SAS Annotated Output: Multinomial Logistic Regression
 (https://stats.idre.ucla.edu/sas/output/multinomial-logistic-regression/)

#### References

Hosmer, D. and Lemeshow, S. (2000) <u>Applied Logistic Regression (Second Edition)</u>
 (/examples/alr2/). New York: John Wiley & Sons, Inc..

Agresti, A. (1996) <u>An Introduction to Categorical Data Analysis (/examples/icda/</u>). New York:
 John Wiley & Sons, Inc.

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