

Exercise 1

Consider a Gaussian mixture model in which the marginal distribution $p(\mathbf{z})$ for the latent variable is given by $p(\mathbf{z}) = \prod_{k=1}^K \pi_k^{z_k}$ (where $z_k \in \{0, 1\}$ and $\sum_k z_k = 1$), and the conditional distribution $p(\mathbf{x}|\mathbf{z})$ for the observed variable is given by $p(\mathbf{x}|\mathbf{z}) = \prod_{k=1}^K \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_k}$. Show that the marginal distribution $p(\mathbf{x})$, obtained by summing $p(\mathbf{z})p(\mathbf{x}|\mathbf{z})$ over all possible values of \mathbf{z} , is a Gaussian mixture of the form $p(\mathbf{x}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$.

Solution:

$$p(\mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{x}, \mathbf{z}) = \sum_{\mathbf{z}} p(\mathbf{x}|\mathbf{z})p(\mathbf{z}) = \sum_{\mathbf{z}} \prod_{k=1}^K (\pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k))^{z_k}$$

Since z can only take 1 of K possible state, we can introduce an indicator variable I such that $I_{kj} = 1$ if $k = j$ and 0 otherwise. Now we can rewrite $p(\mathbf{x})$ as

$$p(\mathbf{x}) = \sum_{j=1}^K \prod_{k=1}^K (\pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k))^{I_{kj}} = \sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)$$

Exercise 2

Robotic arms are widely used for conducting robotics research. We designed two algorithms ($A1, A2$) for catching objects thrown towards a robotic arm. Algorithm 1 catches the objects with an unknown success rate of θ ($p_1 = P(\text{Success}|A1) = P(1|A1) = \theta$) while Algorithm 2 has a 50 percent success rate ($p_2 = p(\text{Success}|A2) = p(1|A2) = 0.5$). We ran the two algorithms several number of times and recorded their results (success $x = 1$ or failure $x = 0$). The algorithms were chosen randomly. Unfortunately after n experiments, we realize that we recorded only the results (success or failure) $\mathbf{X} = \{x^{(1)}, x^{(2)}, \dots, x^{(n)}\}$ without recording the identity ($A1$ or $A2$) of the algorithms $\mathbf{Z} = \{z^{(1)}, z^{(2)}, \dots, z^{(n)}\}$ when performing experiments.

Since repeating experiments on real robot can be a costly and time consuming process, we are interested in estimating the success rate for Algorithm 1 with EM, by only using the incomplete data.

1. Write down the complete data log-likelihood if the identity of the algorithm at each trial was also recorded in the form of a discrete vector \mathbf{z} where k -th element of \mathbf{z} can be either 0 or 1 ($z_k \in \{0, 1\}$) and $\sum_k z_k = 1$.

(If the algorithm at i^{th} trial is $A1$ then $z_1^{(i)} = 1$ and $z_2^{(i)} = 0$ or simply $\mathbf{z}^{(i)} = \begin{bmatrix} 1 \\ 0 \end{bmatrix}$)

2. In EM we have an old estimate for parameters $\boldsymbol{\theta}^{old}$ and the goal is to derive a better estimate of $\boldsymbol{\theta}$. In E-step $Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{old}) = \mathbb{E}_{\mathbf{z}} [\ln p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})|\mathbf{X}, \boldsymbol{\theta}^{old}]$ is calculated, where $\ln p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})$ is the complete data log-likelihood (which you have calculated in the previous step). Show that the Q -function for the given problem can be written as:

$$Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{old}) = \sum_{i=1}^n \sum_{k=1}^2 \gamma(z_k^{(i)}) \{ \log \pi_k + x^{(i)} \log p_k + (1 - x^{(i)}) \log(1 - p_k) \}$$

3. In M-step a revised parameter estimate is calculated as $\boldsymbol{\theta}^{new} = \arg \max_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{old})$. Calculate the update equation for the parameter θ (probability of success for Algorithm 1).

Solution:

Part 1

$$L(\theta) = p(\mathbf{X}, \mathbf{Z}|\theta) = p(\mathbf{X}|\mathbf{Z}, \theta)p(\mathbf{Z}|\theta) = \prod_{i=1}^n \prod_{k=1}^2 \pi_k^{z_k^{(i)}} (p_k^{x^{(i)}} (1-p_k)^{1-x^{(i)}})^{z_k^{(i)}}$$

$$l(\theta) = \log L(\theta) = \sum_{i=1}^n \sum_{k=1}^2 z_k^{(i)} (\log \pi_k + x^{(i)} \log p_k + (1-x^{(i)}) \log(1-p_k))$$

where $\pi_1 = \pi_2 = 0.5$, $p_1 = p(\text{Success}|A1) = \theta$ and $p_2 = p(\text{Success}|A2) = 0.5$.

Part 2

As we don't have the value of the latent variables, we consider the expectation, with respect to the posterior distribution of the latent variables, of the complete data log-likelihood. By using the property $\mathbb{E}[a\mathbf{X} + b\mathbf{Y} + c] = a\mathbb{E}[\mathbf{X}] + b\mathbb{E}[\mathbf{Y}] + c$ where a, b and c are constants.

$$\mathbb{E}_z [\ln p(\mathbf{X}, \mathbf{Z}|\theta) | \mathbf{X}, \theta^{old}] = \sum_{i=1}^n \sum_{k=1}^2 \mathbb{E}_z \{z_k^{(i)}\} \{ \log \pi_k + x^{(i)} \log p_k + (1-x^{(i)}) \log(1-p_k) \}$$

This expectation is calculated as a posterior distribution of latent variable.

$$\mathbb{E}_z \{z_k^{(i)}\} = p(z_k^{(i)} | x^{(i)}, \theta^{old})$$

Now using Bayes rule

$$p(z_k^{(i)} | x^{(i)}, \theta^{old}) = \frac{p(z_k^{(i)} | \theta^{old}) p(x^{(i)} | z_k^{(i)}, \theta^{old})}{p(x^{(i)} | \theta^{old})} = \frac{p(z_k^{(i)} | \theta^{old}) p(x^{(i)} | z_k^{(i)}, \theta^{old})}{\sum_{j=1}^2 p(z_j^{(i)} | \theta^{old}) p(x^{(i)} | z_j^{(i)}, \theta^{old})} = \frac{\pi_k (p_k^{x^{(i)}} (1-p_k)^{1-x^{(i)}})}{\sum_{j=1}^2 \pi_j (p_j^{x^{(i)}} (1-p_j)^{1-x^{(i)}})} = \gamma(z_k^{(i)})$$

$\gamma(z_k^{(i)})$ is also called the responsibility. With this the the final form of Q function is

$$Q(\theta, \theta^{old}) = \sum_{i=1}^n \sum_{k=1}^2 \gamma(z_k^{(i)}) \{ \log \pi_k + x^{(i)} \log p_k + (1-x^{(i)}) \log(1-p_k) \}$$

Part 3

The new value of parameters are the one which maximizes the Q function. For this purpose we evaluate the partial derivative of Q w.r.t. to θ and evaluate it to zero while keeping the responsibilities fixed (as they are a function of θ^{old} and will act as constants).

$$\frac{\partial Q}{\partial \theta} = 0$$

$$\frac{\partial Q}{\partial \theta} = \sum_{i=1}^n \sum_{k=1}^2 \gamma(z_k^{(i)}) \frac{\partial}{\partial \theta} \{ \log \pi_k + x^{(i)} \log p_k + (1-x^{(i)}) \log(1-p_k) \}$$

$$\frac{\partial Q}{\partial \theta} = \sum_{i=1}^n \gamma(z_1^{(i)}) \{ \frac{x^{(i)}}{\theta} + \frac{x^{(i)}-1}{1-\theta} \}$$

$$\frac{\partial Q}{\partial \theta} = \sum_{i=1}^n \gamma(z_1^{(i)}) \{ (x^{(i)})(1-\theta) + (x^{(i)}-1)(\theta) \}$$

$$\frac{\partial Q}{\partial \theta} = \sum_{i=1}^n \gamma(z_1^{(i)}) \{ x^{(i)} - \theta \} = 0$$

$$\implies \theta^{new} = \frac{\sum_{i=1}^n \gamma(z_1^{(i)}) x^{(i)}}{\sum_{i=1}^n \gamma(z_1^{(i)})}$$