# Analysis of Variance for Balanced Two-Factor Experiments

# An Example Two-Factor Experiment

Researchers were interested in studying the effects of 2 diets (low fiber, high fiber) and 3 drugs (D1, D2, D3) on weight gained by Yorkshire pigs. A total of 12 pigs were assigned to the 6 diet-drug combinations using a balanced and completely randomized experimental design. Pigs were housed in individual pens, injected with their assigned drugs once per week, and fed their assigned diets for a 6-week period. The amount of weight gained during the 6-week period was recorded for each pig.

#### A Model for the Data

For i = 1, 2; j = 1, 2, 3; and k = 1, 2; let  $y_{ijk}$  denote the weight gain of the  $k^{th}$  pig that received diet i and drug j, and suppose

$$y_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijk} \quad (i = 1, 2; \ j = 1, 2, 3; \ k = 1, 2) \quad \text{where}$$
 
$$\mu, \alpha_1, \alpha_2, \beta_1, \beta_2, \beta_3, \gamma_{11}, \gamma_{12}, \gamma_{13}, \gamma_{21}, \gamma_{22}, \gamma_{23}$$

are unknown real-valued parameters and

$$\epsilon_{111}, \epsilon_{112}, \epsilon_{121}, \epsilon_{122}, \epsilon_{131}, \epsilon_{132}, \epsilon_{211}, \epsilon_{212}, \epsilon_{221}, \epsilon_{222}, \epsilon_{231}, \epsilon_{232} \stackrel{i.i.d.}{\sim} N(0, \sigma^2)$$

for some unknown  $\sigma^2 > 0$ .

#### Model in Matrix and Vector Form

$$y = X\beta + \epsilon, \quad \epsilon \sim N(\mathbf{0}, \sigma^2 \mathbf{I})$$

#### Same Model with Cell Means Parameterization

$$\begin{bmatrix} y_{111} \\ y_{112} \\ y_{121} \\ y_{122} \\ y_{131} \\ y_{132} \\ y_{211} \\ y_{212} \\ y_{221} \\ y_{222} \\ y_{231} \\ y_{232} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} \mu_{11} \\ \mu_{12} \\ \mu_{13} \\ \mu_{21} \\ \mu_{22} \\ \mu_{23} \end{bmatrix} + \begin{bmatrix} \epsilon_{111} \\ \epsilon_{112} \\ \epsilon_{122} \\ \epsilon_{211} \\ \epsilon_{212} \\ \epsilon_{221} \\ \epsilon_{222} \\ \epsilon_{231} \\ \epsilon_{232} \end{bmatrix}$$

$$y = X\beta + \epsilon, \quad \epsilon \sim N(\mathbf{0}, \sigma^2 \mathbf{I})$$

#### Matrices with Nested Column Spaces

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$$X_4 = \left[ egin{array}{ccccccc} 1 & 0 & 0 & 0 & 0 & 0 & 0 \ 1 & 0 & 0 & 0 & 0 & 0 & 0 \ 0 & 1 & 0 & 0 & 0 & 0 & 0 \ 0 & 1 & 0 & 0 & 0 & 0 & 0 \ 0 & 0 & 1 & 0 & 0 & 0 & 0 \ 0 & 0 & 0 & 1 & 0 & 0 & 0 \ 0 & 0 & 0 & 0 & 1 & 0 & 0 \ 0 & 0 & 0 & 0 & 0 & 1 & 0 \ 0 & 0 & 0 & 0 & 0 & 1 & 0 \ 0 & 0 & 0 & 0 & 0 & 0 & 1 \ 0 & 0 & 0 & 0 & 0 & 0 & 1 \end{array} 
ight].$$

## ANOVA Table for Our Two-Factor Example

Source	Sum of Squares	DF
Diets 1	$y'(P_2-P_1)y$	2 - 1 = 1
Drugs 1, Diets	$\boldsymbol{y}'(\boldsymbol{P}_3-\boldsymbol{P}_2)\boldsymbol{y}$	4 - 2 = 2
$Diets \times Drugs   1, Diets, Drugs$	$\boldsymbol{y}'(\boldsymbol{P}_4-\boldsymbol{P}_3)\boldsymbol{y}$	6 - 4 = 2
Error	$oldsymbol{y}'(oldsymbol{I}-oldsymbol{P}_4)oldsymbol{y}$	12 - 6 = 6
C. Total	$y'(I - P_1)y$	12 - 1 = 11

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$Diets \times Drugs$	$\mathbf{y}'(\mathbf{P}_4 - \mathbf{P}_3)\mathbf{y}$	6 - 4 = 2
Error	$oldsymbol{y}'(oldsymbol{I}-oldsymbol{P}_4)oldsymbol{y}$	12 - 6 = 6
C. Total	$\mathbf{y}'(\mathbf{I} - \mathbf{P}_1)\mathbf{y}$	12 - 1 = 11

#### The Diet-Drug Dataset

```
>
 d
   diet drug weightgain
                      41.3
                      43.7
3
                      40.9
                      39.2
5
                      37.4
6
                      37.9
                      36.8
8
       2
                      34.6
9
                      33.6
10
                      34.3
             3
11
                      35.8
12
                      35.1
```

```
> d$diet=factor(d$diet)
> d$drug=factor(d$drug)
>
> a=d$diet
> b=d$drug
> y=d$weightgain
```

```
> x1=matrix(1, nrow=nrow(d), ncol=1)
> x1
     [,1]
[1,] 1
[2,] 1
[3,] 1
[4,] 1
[5,] 1
[6,] 1
[7,] 1
[8,] 1
[9,] 1
[10,] 1
[11,] 1
[12,] 1
```

```
> x2=cbind(x1, model.matrix(~0+a))
> x2
   x1 a1 a2
3 1 1 0
5
8
10
11
12
```

```
> x3=cbind(x2, model.matrix(~0+b))
> x3
      a1 a2 b1 b2 b3
3
5
6
8
10
11
12
```

```
> x4=model.matrix(~0+b:a)
> x4
   b1:a1 b2:a1 b3:a1 b1:a2 b2:a2 b3:a2
10
11
12
```

```
> SumOfSquares=c(
+ t(v) % * % (p2-p1) % * % v,
+ t(v) % * % (p3-p2) % * % v,
+ t(y) % * % (p4-p3) % * % y,
+ t(y) % * 8 (I-p4) % * 8 y,
+ t(y) % * % (I-p1) % * % y)
>
> Source=c(
+ "Diet | 1",
+ "Drug|1, Diet",
+ "Diet x Drug|1, Diet, Drug",
+ "Error",
+ "C. Total")
```

```
> o=lm(weightgain~diet+drug+diet:drug,data=d)
>
> anova(o)
Analysis of Variance Table
Response: weightgain
         Df Sum Sq Mean Sq F value Pr(>F)
diet 1 76.003 76.003 61.9592 0.0002226 ***
drug 2 14.820 7.410 6.0408 0.0365383 *
diet:drug 2 12.287 6.143 5.0082 0.0525735 .
Residuals 6 7.360 1.227
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
```

#### What do the *F*-tests in this ANOVA table test?

Recall the null hypothesis for  $F_j$  is true if and only if

$$\beta' X' (\mathbf{P}_{j+1} - \mathbf{P}_j) X \beta = 0.$$

We have the following equivalent conditions

$$\beta' X' (\mathbf{P}_{j+1} - \mathbf{P}_j) X \beta = 0 \iff \beta' X' (\mathbf{P}_{j+1} - \mathbf{P}_j)' (\mathbf{P}_{j+1} - \mathbf{P}_j) X \beta = 0$$

$$\iff || (\mathbf{P}_{j+1} - \mathbf{P}_j) X \beta ||^2 = 0$$

$$\iff (\mathbf{P}_{j+1} - \mathbf{P}_j) X \beta = \mathbf{0}$$

$$\iff C\beta = \mathbf{0},$$

where C is any full-row-rank matrix with the same row space as  $(P_{i+1} - P_i)X$ .

#### What do the *F*-tests in this ANOVA table test?

Let's take a look at  $(P_{j+1} - P_j)X$  for each test in the ANOVA table.

When computing  $(P_{j+1} - P_j)X$ , we can use any model matrix X that specifies one unrestricted treatment mean for each of the six treatments.

The entries in any rows of  $(P_{j+1} - P_j)X$  are coefficients defining linear combinations of the elements of the parameter vector  $\beta$  that corresponds to the chosen model matrix X.

## Our Choice for X and $\beta$

$$\boldsymbol{\beta} = \begin{bmatrix} \mu_{11} \\ \mu_{12} \\ \mu_{13} \\ \mu_{21} \\ \mu_{22} \\ \mu_{23} \end{bmatrix}$$

#### **ANOVA Diet Test**

```
> x = x4
> fractions((p2-p1)%*%x)
  b1:a1 b2:a1 b3:a1 b1:a2 b2:a2 b3:a2
  1/6 1/6 1/6 -1/6 -1/6 -1/6
2 1/6 1/6 1/6 -1/6 -1/6 -1/6
3 1/6 1/6 1/6 -1/6 -1/6 -1/6
4 1/6 1/6 1/6 -1/6 -1/6 -1/6
5 1/6 1/6 1/6 -1/6 -1/6
6 1/6 1/6 1/6 -1/6 -1/6
7 -1/6 -1/6 -1/6 1/6 1/6 1/6
8 -1/6 -1/6 -1/6 1/6 1/6 1/6
9 -1/6 -1/6 -1/6 1/6 1/6 1/6
10 -1/6 -1/6 -1/6 1/6 1/6 1/6
11 -1/6 -1/6 -1/6 1/6 1/6 1/6
12 -1/6 -1/6 -1/6 1/6 1/6
                           1/6
```

#### **ANOVA Diet Test**

$$(\mathbf{P}_2 - \mathbf{P}_1)X\beta = \mathbf{0} \iff \mathbf{C}\beta = \mathbf{0},$$

where

$$C\beta = \begin{bmatrix} \frac{1}{3} & \frac{1}{3} & \frac{1}{3} & -\frac{1}{3} & -\frac{1}{3} & -\frac{1}{3} \end{bmatrix} \begin{bmatrix} \mu_{11} \\ \mu_{12} \\ \mu_{13} \\ \mu_{21} \\ \mu_{22} \\ \mu_{22} \end{bmatrix} = \bar{\mu}_{1} - \bar{\mu}_{2}.$$

# **ANOVA Drug Test**

```
> fractions((p3-p2)%*%x)
  b1:a1 b2:a1 b3:a1 b1:a2 b2:a2 b3:a2
1 1/3 -1/6 -1/6 1/3 -1/6 -1/6
2 1/3 -1/6 -1/6 1/3 -1/6 -1/6
3 - 1/6  1/3 - 1/6  - 1/6  1/3 - 1/6
4 - 1/6  1/3 - 1/6  -1/6  1/3 -1/6
5 -1/6 -1/6 1/3 -1/6 -1/6 1/3
6 -1/6 -1/6 1/3 -1/6 -1/6 1/3
7 	 1/3 	 -1/6 	 -1/6 	 1/3 	 -1/6 	 -1/6
8 1/3 -1/6 -1/6 1/3 -1/6 -1/6
9 - 1/6  1/3 - 1/6  - 1/6  1/3 - 1/6
10 -1/6 1/3 -1/6 -1/6 1/3 -1/6
11 -1/6 -1/6 1/3 -1/6 -1/6 1/3
12 -1/6 -1/6 1/3 -1/6 -1/6 1/3
```

# **ANOVA Drug Test**

# **ANOVA Drug Test**

$$(\mathbf{P}_3 - \mathbf{P}_2)X\beta = \mathbf{0} \iff \mathbf{C}\beta = \mathbf{0},$$

where

$$\mathbf{C}\boldsymbol{\beta} = \begin{bmatrix} \frac{1}{2} & -\frac{1}{2} & 0 & \frac{1}{2} & -\frac{1}{2} & 0 \\ \frac{1}{2} & 0 & -\frac{1}{2} & \frac{1}{2} & 0 & -\frac{1}{2} \end{bmatrix} \begin{bmatrix} \mu_{11} \\ \mu_{12} \\ \mu_{13} \\ \mu_{21} \\ \mu_{22} \\ \mu_{23} \end{bmatrix} \\
= \begin{bmatrix} \bar{\mu}_{\cdot 1} - \bar{\mu}_{\cdot 2} \\ \bar{\mu}_{\cdot 1} - \bar{\mu}_{\cdot 3} \end{bmatrix}.$$

# ANOVA Test for Diet × Drug Interactions

```
> fractions((p4-p3)%*%x)
  b1:a1 b2:a1 b3:a1 b1:a2 b2:a2 b3:a2
1 1/3 -1/6 -1/6 -1/3 1/6 1/6
2 1/3 -1/6 -1/6 -1/3 1/6 1/6
3 -1/6 1/3 -1/6 1/6 -1/3 1/6
4 -1/6 1/3 -1/6 1/6 -1/3 1/6
5 -1/6 -1/6 1/3 1/6 1/6 -1/3
6 -1/6 -1/6 1/3 1/6 1/6 -1/3
7 -1/3 1/6 1/6 1/3 -1/6 -1/6
8 -1/3 1/6 1/6 1/3 -1/6 -1/6
9 1/6 -1/3 1/6 -1/6 1/3 -1/6
10 1/6 -1/3 1/6 -1/6 1/3 -1/6
11 1/6 1/6 -1/3 -1/6 -1/6 1/3
12 1/6 1/6 -1/3 -1/6 -1/6 1/3
```

# ANOVA Test for Diet × Drug Interactions

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$$(\mathbf{P}_4 - \mathbf{P}_3)X\beta = \mathbf{0} \iff \mathbf{C}\beta = \mathbf{0},$$

where

$$C\beta = \begin{bmatrix} 1 & -1 & 0 & -1 & 1 & 0 \\ 1 & 0 & -1 & -1 & 0 & 1 \end{bmatrix} \begin{bmatrix} \mu_{11} \\ \mu_{12} \\ \mu_{13} \\ \mu_{21} \\ \mu_{22} \\ \mu_{23} \end{bmatrix}$$
$$= \begin{bmatrix} \mu_{11} - \mu_{12} - \mu_{21} + \mu_{22} \\ \mu_{11} - \mu_{13} - \mu_{21} + \mu_{23} \end{bmatrix}.$$

# **ANOVA for Balanced Two-Factor Experiments**

The diet-drug experiment is *balanced* in the sense that every treatment (defined by a diet-drug combination) has the same number of experimental units.

Each experimental unit provided a single response measurement (weight gain), so the resulting diet-drug dataset is *balanced* in the sense that each treatment has the same number of independent, constant variance observations of the response.

Due to this balance, the tests for diets, drugs, and diets  $\times$  drugs in the ANOVA table turn out to be exactly the same as the tests for diet main effect, drug main effects, and diet  $\times$  drug interactions we discussed previously as tests of  $C\beta = 0$ .