

Applied Econometrics project:

A Structural Model of Birth Weight:

Factors Influencing Birth Weight

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We declare on our honors that this econometrics project on Excel has been written personally, without any unauthorized external help, and that it has never been presented before or published anywhere, even partially.

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The factors which influence the birth weight of a newborn

I. INTRODUCTION

When a newborn arrives, one of the first things you are told is the baby's weight. Knowing the baby's weight not only helps to inform others but is also a clear indicator of the baby's health, but not all babies weigh the same, in well to being a measure of health condition, birth weight is also an indication of economic and social welfare, so, we were interested in looking at what factors affect the baby's weight.

Using representative data collected in the early 1990s, this paper investigates the birth weight of infants and several factors that influence it and discusses their relevance. We found that maternal behavior during pregnancy was strongly associated with infant birth weight, and other factors significantly associated with birth weight included the mother's living environment, such as family income level, and the gender of the infant.

In this paper, we have selected several variables related to the birth weight of infants, including family income, the gender of the infant, the mother's smoking level and the mother's education level, and we investigate the correlation between each explanatory variable and the explanatory variable through some simple linear regression equations.

The remainder of the paper is organized as follows. Section 2 reviews the relevant literature on the determinants of birth weight, followed by the justification of the econometric model, presentation of the data, description of the data, and econometric analysis which includes the discussion of the results in sections 3 to 5 respectively. This is followed by the conclusion of the paper.

II. RELATED LITERATURE

The literature on birth weight is immense and we have found a selection of economic studies on this topic. In general, since the structural production model of birth weight was first proposed, many studies have been carried out along the same lines.

The first is the variable that underlies our study, cigarette, and the variable that is of most interest to us. The number of mothers' daily smoking during pregnancy was used as one of the explanatory variables to explain not only his negative effect on baby's birth weight, cause the harmful effects of smoking on the mother's body can affect the health of the fetus (Shengfang et.al, 2020), but also progresses to analyze the effect on other health indicators of the infant from a medical professional perspective (Magdalena Chelchowska et.al, 2020).

The baby's birth weight also depends on the level of the family income, there is some variation in birth weight due to differences in family income, and the cost of raising a child is one of the main factors determining the fertility rate (Esther Toledano et al., 2013). For example, having a family allowance is equivalent to a certain increase in family income, which reduces the cost of raising a child, so that the quality of life of the mother during pregnancy is improved and is a certain possibility to increase the baby's birth weight, with low-birth-weight babies usually coming from low-income families (Roni Frish et al., 2013).

Gender is another explanatory variable that influences baby's birth weight, some biological differences between males and females result in most male infants weighing more than female infants, and its correlates different national situation which affects the whole social environment and economic development, presenting the social nature of gender differences between men and women, and finally affecting the biological nature of gender (Germano Mwabu, 2008).

Finally, maternal education is strongly associated with poor perinatal results, including preterm birth, and poorer health at birth among infants of mothers with low levels of education (Arnaud

Chevalier and Vincent O’Sullivan, 2007). Further, higher educated pregnant women are more likely than lower educated pregnant women to change their minds about the physical, psychosocial, and behavioral factors that influence pregnancy, for example, poor nutrition, smoking and manual labor (B. K. L. Walawalkar Rural Medical College et al., 2019).

III. Justification of the econometric model

Simple regression linear equation:

1) Variable: cigarette

We identified the first simple linear regression model based on the explanatory variable X = the number of cigarettes smoked per day by the mother, the explanatory variable Y = birth weight, and the error term U :

$$Y_i = \beta_0 + \beta_1 X_{1i} + U_i \quad i = 1, 2, \dots, n \quad (1)$$

The hypothesis of this relation:

We know that smoking is harmful to health. Based on reading the relevant literature and general knowledge of life, we can propose the hypothesis that the number of cigarettes smoked per day by the mother has a negative effect on the baby’s birth weight, and we will test the hypothesis and the significance in the following by analyzing specific data from models of Cigarette Smoking Behavior.

Multiple Linear Regression:

2) Variable: family income

As family income is also a factor that influences birth weight, we added the variable of family income to examine the relationship between the variable of family income and the birth weight and go on to discuss whether and what correlation still existed between the number of cigarettes smoked per day by the mother and the birth weight.

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + U_i \quad i = 1, 2, \dots, n \quad (2)$$

The hypothesis of this relation:

Based on our reading of the literature, we suggest that family income is positively associated

with a baby's birth weight. With the addition of a variable, at the same time, we hypothesized the correlation between the number of cigarettes smoked per day by the mother and the baby's birth weight would change. However, the coefficient for the smoking correlation will change with the addition of the second variable, and we suspect that the correlation will be relatively lower but does not change the positive or negative correlation itself.

3) Variable: male

It is well known that there is a certain difference in body size between males and females, so it is reasonable to suppose that there is a difference in the weight basis between males and females at birth.

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{3i} + U_i \quad i = 1, 2, \dots, n \quad (3)$$

The hypothesis of this relation:

In this model, in addition to the number of cigarettes and family income already studied, we add a new variable - the sex of the child. We know from extensive data that the birth weight of male infants is generally greater than that of female infants. Then, we predict that the effect of the other two variables on the study objective of birth weight will decrease with the addition of the child's sex variable but does not change the positive or negative correlation itself.

4) Variable: mother's education

The last variable we selected to correlate with the birth weight of the infant was the mother's level of education.

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{3i} + \beta_4 X_{4i} + U_i \quad i = 1, 2, \dots, n \quad (4)$$

The hypothesis of this relation:

Based on the literature we read, we suggest that the lower the mother's level of education, the higher the possibility that the infant will be born in poor health, so we suspect that there is a positive correlation between the two, and we will continue to examine whether the correlation between the previous variables on birth weight changes with the inclusion of this variable. We predict that the effect of the other variables will decrease but does not change the positive or negative correlation itself.

IV. DATA

1) Presentation of data

The data we use are derived from Wooldridge Source: J. Mullahy (1997), “Instrumental-Variable Estimation of Count Data Models: Applications to Models of Cigarette Smoking Behavior,” Review of Economics and Statistics 79, 596-593. Professor Mullahy kindly provided the data. He obtained them from the 1988 National Health Interview Survey. Data loads lazily. A data. Frame with 1388 observations on 14 variables. It contains 13 explanatory variables and 1 explanatory variable (bwght), and some log transformations of the variables.

| Variable | Description of the variable |
|----------|---------------------------------|
| bwght | birth weight, ounces |
| cigs | cigs smoked per day while preg |
| cigtax | cig. tax in home state, 1988 |
| cigprice | cig. price in home state, 1988 |
| faminc | 1988 family income, \$1000s |
| male | =1 if male child |
| motheduc | mother's yrs of educ |
| fatheduc | father's yrs of educ |
| parity | birth order of child |
| white | =1 if white |
| lbwght | log of bwght |
| bwghtlbs | birth weight, pounds |
| packs | packs smoked per day while preg |
| lfaminc | log(faminc) |

Table 1

In this project, we have selected four variables out of these fourteen variables, which are: 10th cigarettes; 1st family income; 8th male; 6th mother's education. Because we think about these four variables are of interest to us, particularly the number of cigarettes smoked by the mother in a day during pregnancy. We will create four equations to explain these four variables with the explanatory variable baby's birth weight.

2) Description of data

| Dictionary of variables | Definition | Nature | Unit |
|-------------------------|---|----------------------------------|-----------------|
| bwght | birth weight | discrete quantitative[23, 271] | ounces |
| cig | cigarettes smoked per day while pregnancy | discrete quantitative[0, 50] | 1 piece |
| faminc | 1988 family income | continuous quantitative[0.5, 65] | \$1000 |
| male | gender | nominal qualitative | 1 if male child |
| motheduc | mother's years of education | discrete quantitative[2, 18] | 1 year |

Table 2

| Y: bwght | |
|--|--------------|
| Moyenne | 118.6995677 |
| Erreur-type | 0.546329028 |
| Médiane | 120 |
| Mode | 120 |
| Écart-type | 20.35396434 |
| Variance de l'échantillon | 414.2838643 |
| Kurtosis (Coefficient d'aplatissement) | 3.163342525 |
| Coefficient d'asymétrie | -0.146023516 |
| Plage | 248 |
| Minimum | 23 |
| Maximum | 271 |
| Somme | 164755 |
| Nombre d'échantillons | 1388 |
| Niveau de confiance(95.0%) | 1.07172044 |

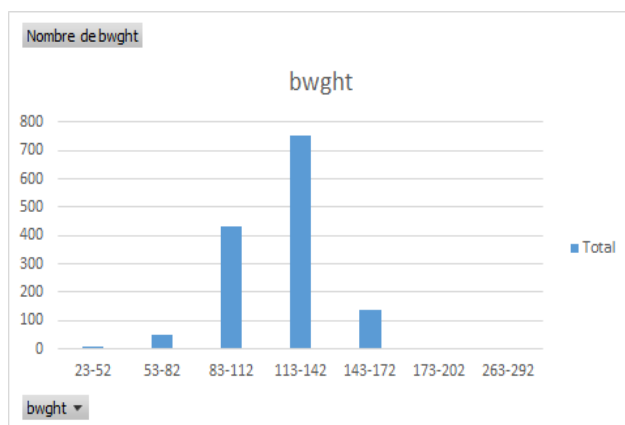


Table 3

Graph 1

According to descriptive statistical analysis, the mean weight of the 1388 newborns observed was 118.69 ounces, with a maximum value of 271 ounces and a minimum value of 23 ounces. Half of the infants had a birth weight of fewer than 120 ounces, while the other half were greater than 120 ounces.

| X1: cig | |
|--|-------------|
| Moyenne | 2.087175793 |
| Erreur-type | 0.160315342 |
| Médiane | 0 |
| Mode | 0 |
| Écart-type | 5.972687881 |
| Variance de l'échantillon | 35.67300053 |
| Kurtosis (Coefficient d'aplatissement) | 14.99224562 |
| Coefficient d'asymétrie | 3.564300836 |
| Plage | 50 |
| Minimum | 0 |
| Maximum | 50 |
| Somme | 2897 |
| Nombre d'échantillons | 1388 |
| Niveau de confiance(95.0%) | 0.31448673 |

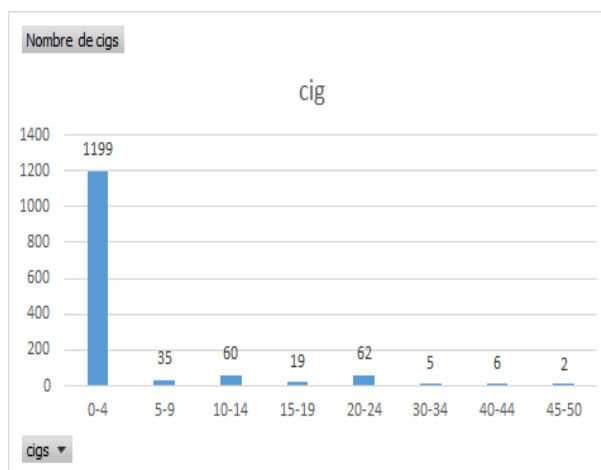


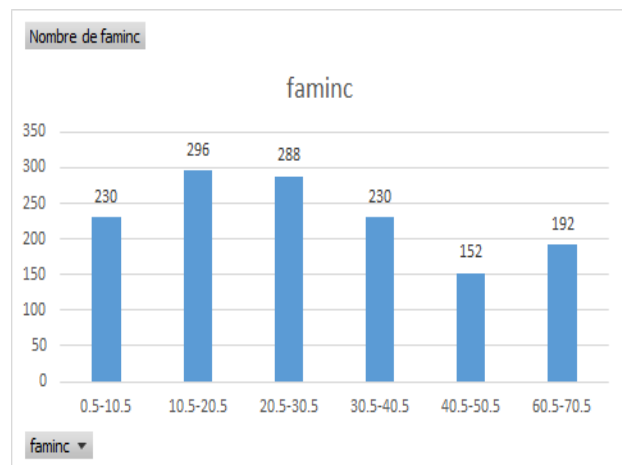
Table 3

Graph 2

The average number of cigarettes smoked per day during pregnancy was 2.09 for 1388 mothers, with a maximum of 50 cigarettes per day and a minimum of no smoking, with a large variation. The median was 0 indicating that half of the mothers did not smoke during pregnancy, while the other half had a daily smoking habit.

| X2: faminc | |
|--|--------------|
| Moyenne | 29.02665706 |
| Erreur-type | 0.502988754 |
| Médiane | 27.5 |
| Mode | 65 |
| Écart-type | 18.73928463 |
| Variance de l'échantillon | 351.1607885 |
| Kurtosis (Coefficient d'aplatissement) | -0.524170256 |
| Coefficient d'asymétrie | 0.618288847 |
| Plage | 64.5 |
| Minimum | 0.5 |
| Maximum | 65 |
| Somme | 40289 |
| Nombre d'échantillons | 1388 |
| Niveau de confiance(95.0%) | 0.986700873 |

Table 4

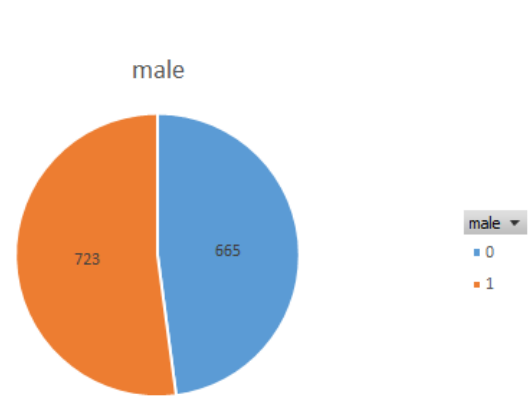


Graph 3

The data show that in 1988 the average annual income of 1388 households was \$29,026.66, with a maximum annual income of \$65,000 and a minimum annual income of \$500. Half of the households had annual incomes below \$27,500, while the other half had incomes greater than \$27,500, with large differences in income between households.

| X3: male | |
|--|--------------|
| Moyenne | 0.520893372 |
| Erreur-type | 0.013413814 |
| Médiane | 1 |
| Mode | 1 |
| Écart-type | 0.499743332 |
| Variance de l'échantillon | 0.249743397 |
| Kurtosis (Coefficient d'aplatissement) | -1.995866072 |
| Coefficient d'asymétrie | -0.083737069 |
| Plage | 1 |
| Minimum | 0 |
| Maximum | 1 |
| Somme | 723 |
| Nombre d'échantillons | 1388 |
| Niveau de confiance(95.0%) | 0.026313554 |

Table 5

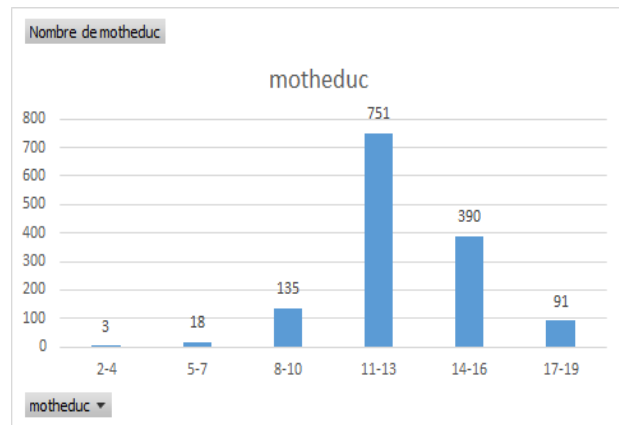


Graph 4

This is a binary variable, which is a quality variable. When the infant is male, the condition is met and is denoted by 1. When the infant is female, the target condition is not met and is denoted by 0. The mean is 0.52, with a similar ratio of males to females, and 1388 infants are roughly equally divided about males and females.

| X4: motheduc | |
|---|--------------|
| Moyenne | 12.93299712 |
| Erreur-type | 0.063834742 |
| Médiane | 12 |
| Mode | 12 |
| Écart-type | 2.378218986 |
| Variance de l'échantillon | 5.655925546 |
| Kurstosis (Coefficient d'aplatissement) | 0.648383261 |
| Coefficient d'asymétrie | -0.031761988 |
| Plage | 16 |
| Minimum | 2 |
| Maximum | 18 |
| Somme | 17951 |
| Nombre d'échantillons | 1388 |
| Niveau de confiance(95.0%) | 0.12522307 |

Table 6



Graph 5

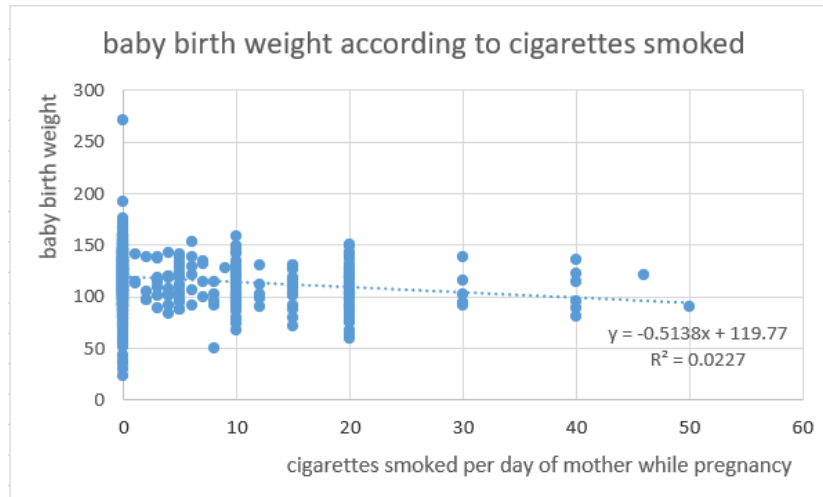
The average number of years of education for the 1388 mothers was 12.93 years, with the highest level of education being 18 years (PhD level) and the lowest level of education being 2 years (not completed elementary school). The median was 12, indicating that half of the mothers had less than 12 years of education and the other half had more than 12 years (high school graduation).

V. Econometric analysis

1) The first equation: simple linear regression (1)

$$Y_i = \beta_0 + \beta_1 X_{1i} + U_i \quad \text{avec } i = 1, 2, \dots, n$$

There are 2 ways to analyze the equation of the fitting line by the OLS (ordinary least squares) between the 2 variables: by the scatterplot or by using the analysis utility in Excel, because we have just one independent variable in this model.



Graph 6

According to the scatterplot (Graph 6), birth weight is negatively correlated with the number of cigarettes smoked per day during the mother's pregnancy: the higher the number of cigarettes smoked, the lower the birth weight. If the mother did not smoke during pregnancy, the infant's birth weight was 119.77 ounces, and the infant's weight decreased by 0.5138 ounces for each additional cigarette smoked, all other things being equal. The coefficient of determination is 0.0227, which means that the model can explain 2.27% of the variations in the situation.

In other words, we can analyze the fit line and the graph of residuals by the analysis utility. In this case, we get more detailed information between the variables of the model.

| | | | | | | |
|---------------------------------------|------------------|------------------|--------------------|-------------|--------------------------|---------------------------|
| RAPPORT DÉTAILLÉ | | | | | | |
| <i>Statistiques de la régression</i> | | | | | | |
| Coefficient de détermination multiple | 0.150761803 | | | | | |
| Coefficient de détermination R^2 | 0.022729121 | | | | | |
| Coefficient de détermination R^2 | 0.022024019 | | | | | |
| Erreur-type | 20.1285784 | | | | | |
| Observations | 1388 | | | | | |
| ANALYSE DE VARIANCE | | | | | | |
| | Degré de liberté | Somme des carrés | Moyenne des carrés | F | Valeur critique de F | |
| Régression | 1 | 13060.41937 | 13060.41937 | 32.23524054 | 1.66154E-08 | |
| Résidus | 1386 | 561551.3004 | 405.1596684 | | | |
| Total | 1387 | 574611.7197 | | | | |
| | Coefficients | Erreur-type | Statistique t | P-value | Limite inférieure α = 5% | Limite supérieure α = 95% |
| Constante | 119.7719004 | 0.572340664 | 209.2668018 | 0 | 118.6491529 | 120.8946479 |
| cig | -0.513772093 | 0.09049093 | -5.677608699 | 1.66E-08 | -0.691286074 | -0.336258112 |

Table 6

We have the following equation of the regression line:

$$\widehat{bwght}_i = 119.771 - 0.514cig_i$$

$$N = 1388$$

Interpretation of the estimate:

if the number of cigarette consumption increases by one, the birth weight of the baby decreases by 0.514 ounces, all other things being equal. The weight of a baby whose mother does not smoke during pregnancy is 119.771 ounces.

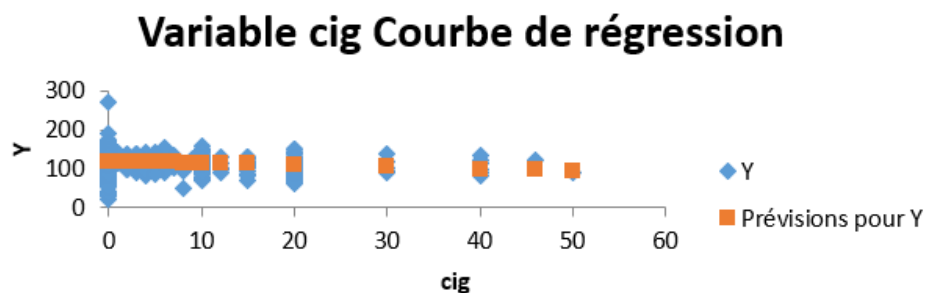
Student's t-test:

$$H_0: \beta_1 = 0$$

$$H_1: \beta_1 \neq 0$$

In table 6 we can find t-statistics for the variable of cigarette which is -5.678, its absolute value is higher than 1,96 (according to the student test for $N > 120$), it is not in the confidence interval [-0.691, -0.336], so we reject H_0 at the 5% risk and this variable is significant.

Since the p-value is 1.66E-08 much less than 0.05, indicating that this variable is significant in this simple linear regression model. Then maternal smoking status during pregnancy has a negative effect on the weight of the newborn.



Graph 7

According to graph 7 of the residuals of the variable cig in relation to the birth weight, it can be observed that the true values of the weight are not perfectly equal to the predicted values of the weight, and therefore the differences between the two values are the residuals.

$$u_i = Y_i - \hat{Y}_i$$

In the analysis of the variance table (table 6), we have $R^2 = 0.0227$. So, the goodness of fit of the estimated model is not good enough, it explains only 2.27% of the situations. The OLS method offers an imperfect fit, and the knowledge of the values of x (cigarette) does not allow us to predict well the values of y (baby's birth weight). We then try to use a "log-level" model to see if the quality of fit of the model improves.

$$\log(Y_i) = \beta_0 + \beta_1 X_{1i} + U_i$$

We were able to get a new analysis table:

| RAPPORT DÉTAILLÉ | | | | |
|---------------------------------------|-------------------------|-------------------------|---------------------------|--------------------|
| <i>Statistiques de la régression</i> | | | | |
| Coefficient de détermination multiple | 0.140674235 | | | |
| Coefficient de détermination R^2 | 0.019789241 | | | |
| Coefficient de détermination R^2 | 0.019082018 | | | |
| Erreur-type | 0.188834348 | | | |
| Observations | 1388 | | | |
| ANALYSE DE VARIANCE | | | | |
| | <i>Degré de liberté</i> | <i>Somme des carrés</i> | <i>Moyenne des carrés</i> | <i>F</i> |
| Régression | 1 | 0.997780189 | 0.997780189 | 27.9816224 |
| Résidus | 1386 | 49.42255745 | 0.035658411 | |
| Total | 1387 | 50.42033763 | | |
| | <i>Coefficients</i> | <i>Erreur-type</i> | <i>Statistique t</i> | <i>Probabilité</i> |
| Constante | 4.769403816 | 0.00536936 | 888.2630622 | 0 |
| cig | -0.004490652 | 0.000848932 | -5.289765819 | 1.4218E-07 |

Table 7

We have the following equation of the regression line:

$$\widehat{\log bwght}_i = 4.77 - 0.0045cig_i$$

$$N = 1388$$

Interpretation of the estimate:

If the number of cigarette consumption increases by one, the birth weight of the baby decreases by 0.45 %, all other things being equal.

Since the p-value is 1.42E-07 much less than 0.05, indicating that the effect of variable cig on the log of the baby's birth weight is very significant in this simple linear regression model.

As the 2 models have the same observations of y and the same number of explanatory variables and, therefore, we can consider the same variance to be explained and directly compare the 2 determination coefficients. In this log-level model, the $R^2 = 0.0198$ is almost equal to that of the level-level model (0.0227), even smaller, which means the number of cigarettes smoked explains 1.98% of the total variation of the log of baby's birth weight observed on the 1388 individuals of the sample. So, the goodness of fit of the model to the empirical data is not very good either. Therefore, we choose the model with the larger coefficient of determination, the level-level model which measures the better quality of adjustment of the model in this case.

In both the original and the changed models, we can conclude that the negative effect of the number of cigarettes smoked by the mother during pregnancy, as the independent variable X_1 , on the dependent variable Y - baby birth weight is significant. However, the shortcomings of this model are also evident in that R^2 is small, so the variation explained is minimal. This becomes the reason for us to continue studying other factors next.

2) The second equation: multiple linear regression with 2 variables (2)

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + U_i \quad \text{avec } i = 1, 2, \dots, n$$

As we have added a second variable, different from the first simple linear regression equation, we cannot get the linear equation and R^2 directly from the scatterplot, so we directly use the

analysis utility.

| RAPPORT DÉTAILLÉ | | | | | | |
|---|------------------|------------------|--------------------|-------------|----------------------|--------------|
| Statistiques de la régression | | | | | | |
| Coefficient de détermination multiple | 0,172640775 | | AIC | 403,3866043 | | |
| Coefficient de détermination R ² | 0,029804837 | | SIC | 407,9773254 | | |
| Coefficient de détermination R ² | 0,028403833 | | | | | |
| Erreur-type | 20,06281671 | | | | | |
| Observations | 1388 | | | | | |
| ANALYSE DE VARIANCE | | | | | | |
| | Degré de liberté | Somme des carrés | Moyenne des carrés | F | Valeur critique de F | |
| Régression | 2 | 17126,20883 | 8563,104417 | 21,27391544 | 7,942E-10 | |
| Résidus | 1385 | 557485,5109 | 402,5166144 | | | |
| Total | 1387 | 574611,7197 | | | | |
| | Coefficients | Erreur-type | Statistique t | P-value | lower 95% | upper 95% |
| Constante | 116,9741305 | 1,04898413 | 111,5118209 | 0 | 114,9163611 | 119,0318999 |
| Cig | -0,46340754 | 0,091576823 | -5,060314648 | 4,74744E-07 | -0,643051806 | -0,283763274 |
| Faminc | 0,092764738 | 0,029187869 | 3,178194982 | 0,001514708 | 0,03550753 | 0,150021947 |

Table 8

We have the following equation of the multiple linear regression:

$$\widehat{\text{bwght}} = 116.974 - 0.463\text{Cig} + 0.093\text{faminc}$$

$$N=1388$$

Interpretation of the estimate:

The mother doesn't smoke during her pregnancy and her baby's birth weight increases by 0.093 ounces when her family income increases by \$1,000; if family income is 0, the infant's birth weight decreases by 0.463 ounces with one unit increase in the number of cigarettes smoked; if a mother does not smoke and her family income is zero then the birth weight of her child is 116.974 ounces, all other things being equal. Compared to the first equation containing only one variable of cigarette smoking, the absolute value of the cigarette smoking coefficient decreases a bit, which indicates that the effect of maternal cigarette smoking on infant birth weight is less significant than before, influenced by the variable of family income.

Fisher's test:

$H_0: \beta_1 = 0, \beta_2 = 0$, non-significant model

$H_1: \beta_1 \neq 0, \beta_2 \neq 0$, significant model

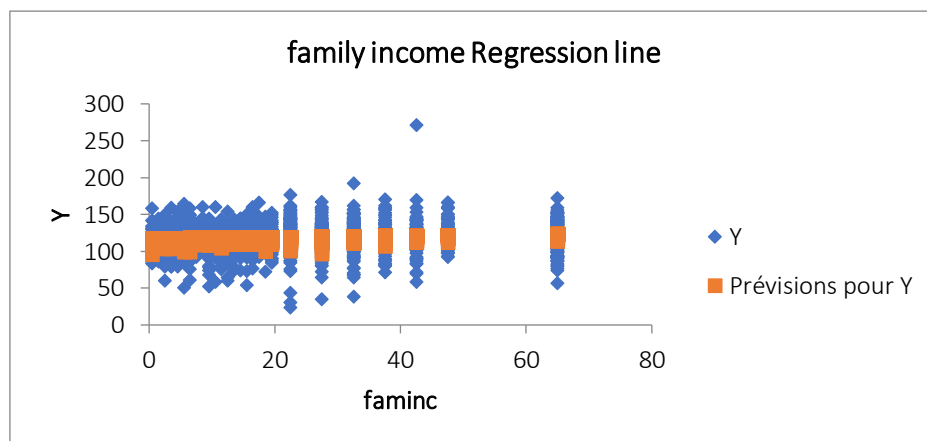
In table 8 we can find F for the variable of family income which is 21.274 > 2.99 (according to the Fisher for $N > 120, k=2$) so we reject H_0 at the 5% risk and this model is significant.

The P-value of the variable cigarette is 4.74744E-07 and the P-value of the variable family

income is 0.001514708. The p-value of every variable is much less than 0.05, indicating that 2 variables are significant in this linear regression model.

We then compare the bivariate and univariate level-level models. Even though the 2 models have the same observations of y, they do not have the same number of explanatory variables, so, we cannot directly compare the 2 determination coefficients. But we can calculate other predictive quality indicators like the adjusted R-squared. The adjusted R-Squared of this equation is 0.0284, The greater adjusted R-Squared compared to the previous equation indicates that with the inclusion of the family income variable, this bivariate model is somewhat more significantly representative than the univariate.

Then maternal smoking status during pregnancy has a negative effect on the weight of the newborn, and it is the opposite for family income. But as influenced by the addition of the second variable, the absolute value of the correlation coefficient between smoking and birth weight changed from 0,51 to 0,46, indicating a decrease in the correlation between the mother's smoking during pregnancy and the baby's birth weight, influenced by the family income.



Graph 8

From table 8 and graph 8 we can see that there is some difference between the estimated and real values of birth weight, and the goodness of fit of the estimated model is not good enough, it explains only 2.29% of the situations.

We then try to use a "log-level" model to see if the quality of the model improves. We will use a logistic function model to investigate the percentages that affect birth weight:

$$\text{Log}(Y_i) = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + U_i \quad \text{avec } i = 1, 2, \dots, n$$

We were able to get a new analysis table:

| RAPPORT DÉTAILLÉ | | | | | | |
|---------------------------------------|------------------|------------------|--------------------|-------------|--------------------------|--------------------------|
| Statistiques de la régression | | | | | | |
| Coefficient de détermination multiple | 0,162666362 | | | | | |
| Coefficient de détermination R^2 | 0,026460345 | | AIC | 0,0355179 | | |
| Coefficient de détermination R^2 | 0,025054512 | | SIC | 0,03592211 | | |
| Erreur-type | 0,188258594 | | | | | |
| Observations | 1388 | | | | | |
| ANALYSE DE VARIANCE | | | | | | |
| | Degré de liberté | Somme des carrés | Moyenne des carrés | F | Valeur critique de F | |
| Régression | 2 | 1,334139543 | 0,667069772 | 18,821821 | 8,60795E-09 | |
| Résidus | 1385 | 49,08619809 | 0,035441298 | | | |
| Total | 1387 | 50,42033763 | | | | |
| | Coefficients | Erreur-type | Statistique t | Probabilité | Limite inférieure α = 5% | Limite supérieure α = 5% |
| Constante | 4,743956539 | 0,009843098 | 481,95765 | 0 | 4,724647547 | 4,763265532 |
| Cigs | -0,004032558 | 0,000859307 | -4,692801271 | 2,9616E-06 | -0,005718243 | -0,002346874 |
| faminc | 0,000843747 | 0,000273883 | 3,080682519 | 0,00210595 | 0,000306476 | 0,001381018 |

Table 9

$$\log(\widehat{\text{bwght}}) = 4.744 - 0.004\text{Cig} + 0.00084\text{faminc}$$

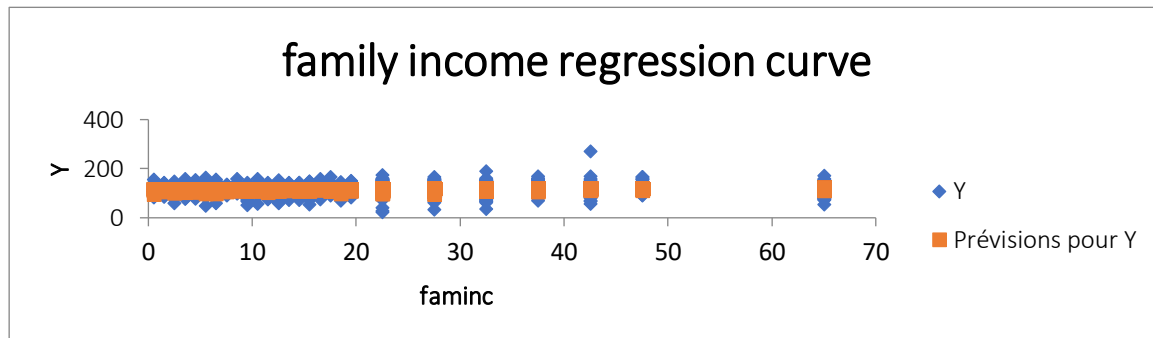
$$N=1388$$

Interpretation of the estimate:

The mother does not smoke during her pregnancy and her baby's birth weight increases by 0.084% when her family income increases by \$1,000; if family income is zero, the infant's birth weight decreases by 0.04% with one unit increase in the number of cigarettes smoked, all other things being equal.

The effect of cigarettes on the log of birth weight is significant in this simple linear regression model because the p-value is 2.9616E-06, and family income on the log of birth weight is significant also because the p-value is 0.00210595.

At the same time, we can see that the difference between the $R^2 = 0,0265$ of the log-level model and the $R^2 = 0,0298$ of the level-level model is not too much, or even smaller, which means that the fit of the level-level model is relatively higher, and the error term is smaller and closer to the true value.



Graph 9

To investigate whether there are equations that make the estimates more like the real values, so that these explanatory variables can better explain the variation in Y, we tried to use a ‘log–log’ model:

$$\text{Log}(Y_i) = \beta_0 + \beta_1 X_{1i} + \beta_2 \log(X_{2i}) + U_i \quad \text{avec } i = 1, 2, \dots, n$$

We were able to get a new analysis table:

| RAPPORT DÉTAILLÉ | | | | | | |
|---------------------------------------|------------------|------------------|--------------------|------------|----------------------|------------|
| Statistiques de la régression | | | | | | |
| Coefficient de détermination multiple | 0,160496985 | | | | | |
| Coefficient de détermination R^2 | 0,025759282 | | | | | |
| Coefficient de détermination R^2 | 0,024352436 | | | | | |
| Erreur-type | 0,188326366 | | | | | |
| Observations | 1388 | | | | | |
| ANALYSE DE VARIANCE | | | | | | |
| | Degré de liberté | Somme des carrés | Moyenne des carrés | F | Valeur critique de F | |
| Régression | 2 | 1,298791701 | 0,649395851 | 18,3099541 | 1,41709E-08 | |
| Résidus | 1385 | 49,12154593 | 0,03546682 | | | |
| Total | 1387 | 50,42033763 | | | | |
| | Coefficients | Erreur-type | Statistique t | P-value | lower 95% | upper 95% |
| Constante | 4,71859367 | 0,018244497 | 258,631057 | 0 | 4,682803835 | 4,7543835 |
| Cig | -0,004081578 | 0,000858214 | -4,755900294 | 2,1819E-06 | -0,005765117 | -0,002398 |
| lfaminc | 0,016265688 | 0,005583314 | 2,913267922 | 0,00363405 | 0,005313023 | 0,02721835 |

Table 10

$$\log(\widehat{\text{bwght}}) = 4.7186 - 0.004\text{Cig} + 0.0163\log\text{faminc}$$

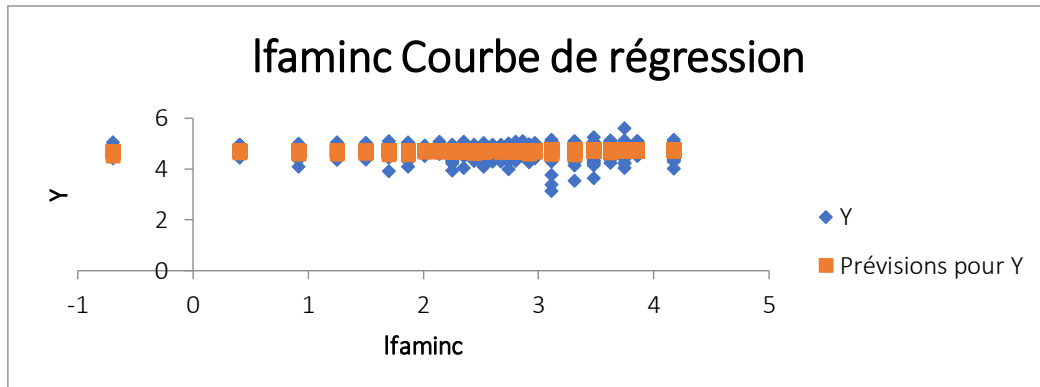
N=1388

Interpretation of the estimate:

The mother doesn't smoke during her pregnancy and her baby's birth weight increases by 0.0163% when her family income increases by 1% (\$10); if family income is zero, infant birth weight decreases by 0.04% with one unit increase in number of cigarettes smoked also, all other things being equal.

The effect of cigarettes on the log of birth weight is very significant in this simple linear

regression model because the p-value is 2.9616E-06, and family income on the log of birth weight is significant also because the p-value is 0.00210595.



Graph 10

From table 10 and graph 10, we can see that $R^2 = 0.02576$, which is still in a decreasing situation compared to the previous log level model with $R^2 = 0.02646$, therefore, by transforming the three models, we believe that the level-level model is relatively accurate and the goodness fit of the regression equation is better.

In both the original and the changed models, we can conclude that the negative effect of the number of cigarettes smoked by the mother during pregnancy, as the independent variable X_1 , on the dependent variable Y — baby birth weight is significant, and the positive effect of the family income on baby's birth weight. The shortcomings of this model are also evident in that R^2 is small, so the variation explained is minimal, although the coefficient of determination has increased compared to the previous two models. In the next step, we continue to add new independent variables and investigate whether the new model can further improve the quality.

3) The third equation with 3 variables: **(3)**

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{3i} + U_i$$

We have the following equation of the regression line:

$$\widehat{bwght}_i = 115.23 - 0.461cig_i + 0.0969faminc_i + 3.114male_i$$

$$N = 1388$$

| | | | | | | |
|---|------------------|------------------|--------------------|------------|----------------------------------|----------------------------------|
| RAPPORT DÉTAILLÉ | | | | | | |
| <i>Statistiques de la régression</i> | | | | | | |
| Coefficient de détermination multiple | 0.188776182 | | | | | |
| Coefficient de détermination R ² | 0.035636447 | AIC | 401.540116 | | | |
| Coefficient de détermination R ² | 0.033546063 | SIC | 407.644587 | | | |
| Erreur-type | 20.00965446 | | | | | |
| Observations | 1388 | | | | | |
| ANALYSE DE VARIANCE | | | | | | |
| | Degré de liberté | Somme des carrés | Moyenne des carrés | F | Valeur critique de F | |
| Régression | 3 | 20477.11998 | 6825.70666 | 17.0478039 | 7.10494E-11 | |
| Résidus | 1384 | 554134.5998 | 400.3862715 | | | |
| Total | 1387 | 574611.7197 | | | | |
| | Coefficients | Erreur-type | Statistique t | P-value | Limite inférieure $\alpha = 5\%$ | Limite supérieure $\alpha = 5\%$ |
| Constante | 115.2277082 | 1.207880175 | 95.39663828 | 0 | 112.8582344 | 117.597182 |
| cig | -0.4610457 | 0.091337813 | -5.047698062 | 5.0662E-07 | -0.640221217 | -0.281870183 |
| faminc | 0.096879835 | 0.02914526 | 3.324034017 | 0.00091045 | 0.039706175 | 0.154053494 |
| male | 3.113967891 | 1.076396465 | 2.892956258 | 0.00387607 | 1.002422981 | 5.225512801 |

Table 11

Interpretation of the estimate:

The weight of a female baby whose mother does not smoke during pregnancy and has zero family income is 115.23 ounces, if the number of cigarette consumption increases by one, the family income is zero and the baby is female, the birth weight of the baby decreases by 0.46 ounces; if the family income increases by \$1,000, the mother does not smoke and baby is female, the birth weight of baby increases by 0.097 ounces; if the baby is male, the mother does not smoke and the family income is zero, the birth weight of baby increases by 3.11 ounces, all other things being equal.

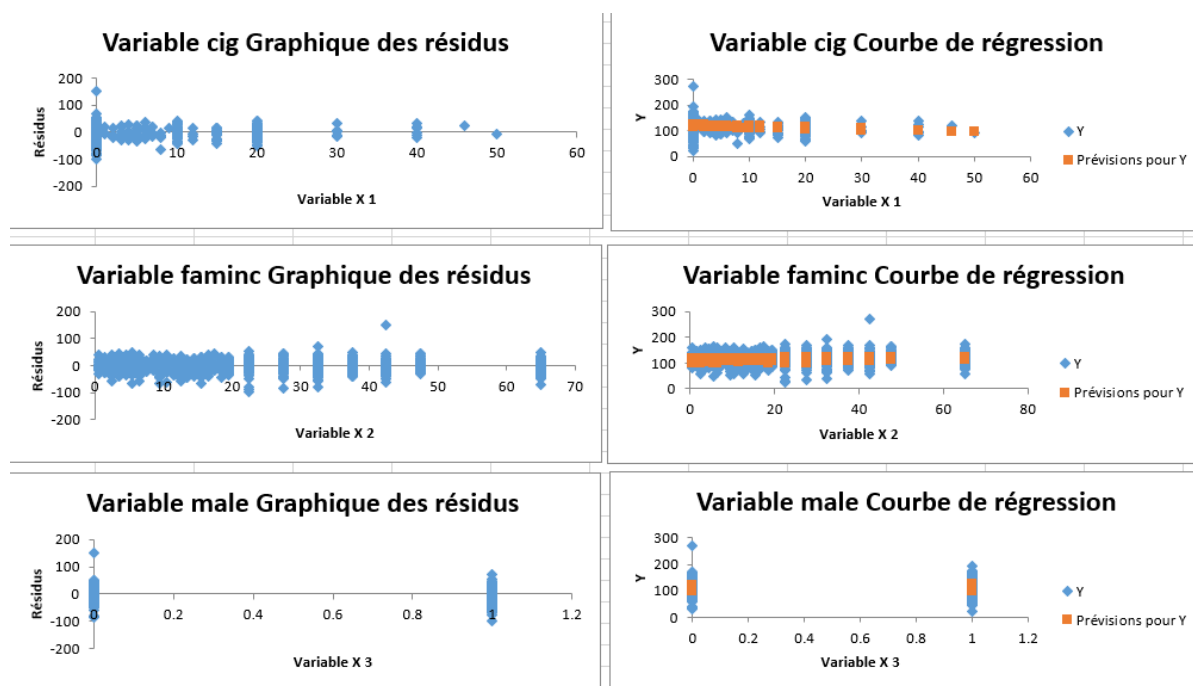
Fisher's test:

$H_0: \beta_1 = 0, \beta_2 = 0, \beta_3 = 0$ non-significant model

$H_1: \beta_1 \neq 0, \beta_2 \neq 0, \beta_3 \neq 0$ significant model

In table 11 we can find F for the variable of family income which is $17.048 > 2.60$ (according to the Fisher for $N > 120, k=3$) so we reject H_0 at the 5% risk and this model is significant.

Since the p-value of every variable is much less than 0.05, indicating that 3 variables cig, faminc and male are significant in this linear regression model. Then maternal smoking status during pregnancy has a negative effect on the weight of the newborn, and it's the opposite for family income and male infants.



Graph 11

According to the regression curve and the graph of the residuals of the variables---cig, faminc and male in relation to the baby's birth weight, it can be observed that the true values of the weight are not perfectly equal to the predicted values of the weight, and therefore the differences between two values are the residuals.

In the internal comparisons of this model, gender has the most direct and important effect on weight, followed by the number of mothers smoking and finally by family income. This result can be interpreted as the infant's weight is the most correlated with its own nature, followed by the close physiological correlation between mother and infant, where the mother's health status tends to be reflected in the child, and finally by external factors similar to family status, where a higher family income the more likely the mother is to have a healthy and stable life status, then there is some positive effect on the birth weight of the newborn.

Compared to the previous two univariate and bivariate models, the trivariate model has increased quality and the ability to explain the changes explained. Comparing the coefficients of different variables across models, the coefficient on smoking decreased continuously from -0.51 to -0.4634 to -0.4610, and the coefficient on family income increased from 0.0928 to

0.0969, indicating that the addition of new variables reduced the effect of smoking on body weight and increased the effect of income and gender on body weight, but the overall correlation remained the same.

In the analysis of the variance table, we have $R^2 = 0.0356$. So, the goodness of fit of the estimated model is not good enough, it explains only 3.56% of the situations. The OLS method offers an imperfect fit, and the knowledge of the values of cigarettes, family income, and gender of the baby does not allow for predict well the values of the baby's birth weight. We then try to use a "log-level" model to see if the quality of the model improves.

$$\log(Y_i) = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{3i} + U_i$$

| RAPPORT DÉTAILLÉ | | | | |
|---|------------------|------------------|--------------------|-------------|
| Statistiques de la régression | | | | |
| Coefficient de détermination multiple | 0.176333088 | | | |
| Coefficient de détermination R ² | 0.031093358 | | | |
| Coefficient de détermination R ² | 0.028993127 | | | |
| Erreur-type | 0.187877943 | | | |
| Observations | 1388 | | | |
| ANALYSE DE VARIANCE | | | | |
| | Degré de liberté | Somme des carrés | Moyenne des carrés | F |
| Régression | 3 | 1.567737614 | 0.522579205 | 14.8047313 |
| Résidus | 1384 | 48.85260002 | 0.035298121 | |
| Total | 1387 | 50.42033763 | | |
| | Coefficients | Erreur-type | Statistique t | Probabilité |
| Constante | 4.729375031 | 0.011341227 | 417.0073345 | 0 |
| cig | -0.004012838 | 0.000857604 | -4.679127179 | 3.163E-06 |
| faminc | 0.000878105 | 0.000273655 | 3.208799012 | 0.00136346 |
| male | 0.025999638 | 0.010106679 | 2.572520411 | 0.01019956 |

Table 12

We have the following equation of the regression line:

$$\widehat{\log bwght}_i = 4.73 - 0.004cig_i + 0.0009faminc_i + 0.0259male_i$$

$$N = 1388$$

Interpretation of the estimate:

If the number of cigarette consumption increases by one, the family income is zero and baby is female, the birth weight of the baby decreases by 0.4 %; if the family income increases by \$1,000, the mother does not smoke and baby is female, the birth weight of baby increases by 0.09 %; if the baby is male, the mother does not smoke and the family income is zero, the birth weight of baby increases by 2.59 %, all other things being equal.

Since p-value of every variable much less than 0.05, indicating that the effect of 3 variables cig, faminc and male on the log of the baby's birth weight is very significant in this linear regression model. The different coefficients represent the closeness of the association between the different factors and the dependent variable: the effect of the infant's sex on its birth weight is the most direct, with male infants usually being 3.11 ounces heavier than female infants; followed by the negative effect of the mother's smoking habit during pregnancy on the health of the newborn; in contrast, the effect of household income on the infant's weight is present, but the relationship is not as strong.

As the 2 models have the same observations of y and the same number of explanatory variables and, therefore, we can consider the same variance to be explained and directly compare the 2 determination coefficients. In this log-level model, the $R^2 = 0.0311$ is almost equal to that of the level-level model (0.0356), even smaller, which means the model explains 3.11% of the total variation of the log of baby's birth weight observed on the 1388 individuals of the sample. So, the goodness of fit of the model to the empirical data is not very good either. Therefore, we choose the model with the larger coefficient of determination, the level-level model which measures the better quality of adjustment of the model in this case.

We then compare the bivariate and trivariate level-level models. Even the 2 models have the same observations of y , but they do not have the same number of explanatory variables, so, we cannot directly compare the 2 determination coefficients. But calculate other predictive quality indicators just like the adjusted R-squared, Akaike's Information Criterion (AIC) and Schwartz's Information Criterion (SIC) could be a solution.

The adjusted R^2 of bivariate models is 0.0284, and the adjusted R^2 of trivariate models is 0.0335. The AIC of bivariate models is 403.39, and the AIC of trivariate models is 401.54. The SIC of bivariate models is 407.98, and the SIC of trivariate models is 407.64.

Increasing the number of explanatory variables improved the explanatory power of the linear

fit, and the more the model minimizes the variance of the residual, the better the fit will be. So, the preferred model will be the model with the highest adjusted R-squared, lowest AIC and SIC. The consideration of these three additional criteria confirms our choice of model with 3 explanatory variables.

In both the original and the changed models, we can conclude that the negative effect of the number of cigarettes smoked by the mother during pregnancy, as the independent variable X_1 , on the dependent variable Y— baby birth weight is significant, and the positive effect of the family income and the sex of baby on baby's birth weight. The shortcomings of this model are also evident in that R^2 is small, so the variation explained is minimal, although the coefficient of determination has increased compared to the previous two models. In the next step, we continue to add new independent variables and investigate whether the new model can further improve the quality.

4) The fourth equation with 4 variables: (4)

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{3i} + \beta_4 X_{4i} + U_i$$

| RAPPORT DÉTAILLÉ | | | | | | |
|---------------------------------------|------------------|------------------|--------------------|-------------|----------------------------------|----------------------------------|
| Statistiques de la régression | | | | | | |
| Coefficient de détermination multiple | 0,188776189 | | AIC | 402,11912 | | |
| Coefficient de détermination R^2 | 0,03563645 | | SIC | 409,7751778 | | |
| Coefficient de détermination R^2 | 0,032847256 | | | | | |
| Erreur-type | 20,01688727 | | | | | |
| Observations | 1388 | | | | | |
| ANALYSE DE VARIANCE | | | | | | |
| | Degré de liberté | Somme des carrés | Moyenne des carrés | F | Valeur critique de F | |
| Régression | 4 | 20477,12157 | 5119,280393 | 12,77661566 | 3,24688E-10 | |
| Résidus | 1383 | 554134,5982 | 400,675776 | | | |
| Total | 1387 | 574611,7197 | | | | |
| | Coefficients | Erreur-type | Statistique t | Probabilité | Limite inférieure $\alpha = 5\%$ | Limite supérieure $\alpha = 5\%$ |
| Constante | 115,2335505 | 3,168783834 | 36,36522922 | 9,4124E-204 | 109,0174082 | 121,4496928 |
| cigs | -0,46107391 | 0,092459191 | -4,986782866 | 6,91679E-07 | -0,642449328 | -0,279698493 |
| faminc | 0,096908074 | 0,032412054 | 2,989877581 | 0,002840202 | 0,033325971 | 0,160490177 |
| male | 3,114025115 | 1,077167752 | 2,890937934 | 0,003900939 | 1,000965851 | 5,22708438 |
| motheduc | -0,00051287 | 0,257154407 | -0,001994404 | 0,998408985 | -0,504967724 | 0,503941985 |

Table 12

The equation of the regression linear equation is as follows:

$$\widehat{\text{bwght}} = 115.2336 - 0.4611\text{Cig} + 0.0969\text{faminc} + 3.1140\text{male} - 0.0005\text{motheduc}$$

$$N=1388$$

Interpretation of the estimate:

If a mother is a non-smoker, her family income is 0 and the baby is female, the baby's weight decreases by 0.0005 ounces when her year of education increases by one unit (one year). If the mother is a non-smoker, the family income is zero, the sex of the infant is male, and the mother's education level is zero, the infant's birth weight increases by 3.114 ounces from 115.2336 ounces, and if the mother is a non-smoker and the sex of the infant is female, when the family income increases by \$1,000, the infant's birth weight increases by 0,0969 ounces, and when the family income is zero, the mother's education level is also zero and the infant is female, the infant's birth weight decreases by 0,4611 ounces when her number of cigarettes per day increases by one, all other things being equal.

The negative coefficient on the variable related to mother's education in this model was surprising to us because in our earlier reviews of the literature, our assumption was that mother's education was positively related to the birth weight of the infant, whereas our data showed a negative relationship between mother's education and birth weight in the absence of testing errors.

Fisher's test:

$H_0: \beta_1 = 0, \beta_2 = 0, \beta_3 = 0, \beta_4 = 0$ non-significant model

$H_1: \beta_1 \neq 0, \beta_2 \neq 0, \beta_3 \neq 0, \beta_4 \neq 0$ significant model

In the table 12 we can find F for the variable of family income which is $12.7766 > 2.37$ (according to the Fisher for $N > 120, k=4$) so we reject H_0 at the 5% risk and this model which have four variables is significant.

However, since the p-value of the variable cig, faminc and male less than 0.05, indicating that variables cigarette, family income and sex of are significant in this linear regression model. But the p-value of variable motheduc is more than 0.05, so the variable the mother's education is non-significant of this model. And when compared with the previous equation, we can see that the variable of the mother's education does not affect the correlation coefficients of the other three variables with infant birth weight.

Therefore, although we hypothesized in our speculation that the higher the education of the mother, the higher the birth weight of the baby would also increase, as more educated pregnant women are more likely to change their perceptions of the physical, psychosocial and behavioral factors that influence pregnancy than less educated pregnant women, we did not reach this conclusion through our data analysis and we were unable to conclude that there was a definite correlation between the two.

We suppose that the reason for the difference may be a limitation of the data, which comes from 1988, when perhaps most women were at essentially the same level of education and all at a low level, so it is possible that the social environment was such that the level of education of the mothers of that period did not affect the birth weight of their babies.

Since the Student t test concludes that the coefficients are not significant coefficients while the Fisher test concludes that the model is significance of the model, it is the index of multicollinearity, so we make the test of Klein as the detection of the multicollinearity.

The general model:

$$\text{bwght}=\beta_0+\beta_1\text{cigs}+\beta_2\text{faminc}+\beta_3\text{male}+\beta_4\text{motheduc}+U_i$$

The auxiliary models:

- a) $\text{Cigs}=\beta_0+\beta_1\text{faminc}+\beta_2\text{male}+\beta_3\text{motheduc}$
- b) $\text{Faminc}=\beta_0+\beta_1\text{cigs}+\beta_2\text{male}+\beta_3\text{motheduc}$
- c) $\text{Male}=\beta_0+\beta_1\text{cigs}+\beta_2\text{faminc}+\beta_3\text{motheduc}$
- d) $\text{Mohteduc}=\beta_0+\beta_1\text{cigs}+\beta_2\text{faminc}+\beta_3\text{male}$

| | R^2 | TOL | VIF |
|---------|--------------|------------|------------|
| general | 0,0356364496 | 0.96436355 | 1.03695333 |
| a) | 0.0527234083 | 0.94727659 | 1.05565788 |
| b) | 0.2169355637 | 0.7830644 | 1.27703412 |
| c) | 0.0030900653 | 0.99690993 | 1.00309964 |
| d) | 0.2276307072 | 0.77236929 | 1.29471745 |

The R^2 of the auxiliary models for cigarette, family income and mother's education are greater than the R^2 of the general model, so there is a multicollinearity problem among the three variables. An R^2 of the auxiliary model with male as the explained variable which close to 0 indicates that the independent variables do not explain most of the variation in the dependent variable, which confirms the suspicion that the male variable is not multicollinear with the other variables.

But after calculating the results of the tolerance and variance inflation factor (VIF) for each auxiliary regression, we see that the TOL are close to 1 and VIF are small. This means that a large part of the variance of an explanatory variable that is not explained by the other explanatory variables which does not provide evidence on multicollinearity.

Heteroscedasticity of disturbances

The initial model is estimated by the OLS:

$$bwght = \beta_0 + \beta_1 cigs + \beta_2 faminc + \beta_3 male + \beta_4 motheduc + U_i$$

The auxiliary model is estimated:

$$\hat{u}_i^2 = \lambda_0 + \lambda_1 cigs + \lambda_2 faminc + \lambda_3 male + \lambda_4 motheduc + v_i$$

Breusch – Pagan test :

$H_0 : V(U_i) = \sigma^2, \forall n$ (Homoscedasticity hypothesis)

$H_1 : V(U_i) = \sigma_i^2 = \lambda_0 + \lambda_1 cigs + \lambda_2 faminc + \lambda_3 male + \lambda_4 motheduc$ (Heteroscedasticity hypothesis)

| | |
|--------------|-------------|
| R^2 | 0,00027650 |
| BP_C | 0,38378347 |
| $Chi2(4ddl)$ | 9,487729037 |

$BP_C < Chi2$, the null hypothesis of homoscedasticity is not rejected.

The random error terms in the overall regression function satisfy the homoscedasticity, that is, they all have the same variance, which conforms to one of the important assumptions of the classical linear regression model and ensures that the regression parameter estimators have good statistical properties.

VI. Conclusion

The purpose of this paper is to study the factors influencing the birth weight of newborns and to select four out of a large amount of data as the main study subjects: the number of mothers smoking during pregnancy, household income, infant gender, and mother's education level. The first simple linear regression equation was established with the number of cigarettes smoked as the base variable, and then some new multiple linear regressions were established by adding variables one after another, and a comparative analysis was made based on the differences within and between the models to draw conclusions and explain the causes.

From the analysis of the data, we conclude that the effects of the first three selected explanatory variables on the study population were significant same as our hypothesis, and that a reduction in smoking by the mother during pregnancy, an increase in household income and male infants were responsible for increasing the birth weight of the infant. The underlying linear relationship was constant regardless of model changes. However, as the variables increase, the parameters of each variable change, either increasing or decreasing, representing a change in the degree of influence of each variable on the study population, and the contribution may increase or decrease. Unexpectedly the fourth variable, the mother's education, unlike our hypothesis, does not reflect the significance of the variable and has the opposite sign of correlation to what we expected.

But in the process of analyzing the results, we found a common problem, the values of R^2 are small, regardless of the number of variables or the type of model (level-level, log-level, log-log). Among all the models studied, the largest one does not exceed 5%, indicating that the variables do not explain more than 5% of the variation. This is not only a conclusion for us but also a limitation and a question for our study. Another shortcoming is that the data empirical tests differ from the hypothesis conclusions, excluding the reasons for analytical errors, and the solution may lie in the selection of data and model building. These will be the questions to be addressed with attention as we continue our study.

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