

Lecture 06

* $X_i \stackrel{iid}{\sim} N(0, 1)$, $\bar{X} = \frac{1}{n}(X_1 + \dots + X_n) \sim N(0, \frac{1}{n})$] indep.
 $\sum_{i=1}^n (X_i - \bar{X})^2 \sim \chi^2_{n-1}$

$$X = \begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix} \sim N\left(\begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & & & \\ & \ddots & & \\ & & 1 & \end{bmatrix}\right) \stackrel{d}{=} N(0, I)$$

$$\begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} = Y = PX \quad \text{where } P = \begin{bmatrix} \frac{1}{\sqrt{n}}, \frac{1}{\sqrt{n}}, \frac{1}{\sqrt{n}}, \dots, \frac{1}{\sqrt{n}} \\ \frac{1}{\sqrt{n(n-1)}}, \frac{1}{\sqrt{n(n-1)}}, \frac{1}{\sqrt{n(n-1)}}, \dots, -\frac{n-1}{\sqrt{n(n-1)}} \\ \vdots & \vdots & \vdots & \vdots \\ \frac{1}{\sqrt{3 \times 4}}, \frac{1}{\sqrt{3 \times 4}}, \frac{1}{\sqrt{3 \times 4}}, \dots, 0 \\ \frac{1}{\sqrt{2 \times 3}}, \frac{1}{\sqrt{2 \times 3}}, -\frac{1}{\sqrt{2 \times 3}}, \dots, 0 \\ \frac{1}{\sqrt{1 \times 2}}, -\frac{1}{\sqrt{1 \times 2}}, 0, \dots, 0 \end{bmatrix} \quad \text{norm 1}$$

$(P^T P = P P^T = I)$
orthogonal

$$Y \sim N(P E(X), P \text{Var}(X) P^T) \stackrel{d}{=} N(0, I)$$

Multiply by P : Rotation

Specifically, $Y_1 = \frac{1}{\sqrt{n}}X_1 + \dots + \frac{1}{\sqrt{n}}X_n = \sqrt{n}\bar{X}$
 $Y_1^2 + Y_2^2 + \dots + Y_n^2 = \|Y\|_2^2 = \|X\|_2^2 = X_1^2 + X_2^2 + \dots + X_n^2$

$$\begin{aligned} \text{So } Y_2^2 + \dots + Y_n^2 &= X_1^2 + \dots + X_n^2 - Y_1^2 \\ &= X_1^2 + \dots + X_n^2 - n\bar{X}^2 \\ &= \sum_{i=1}^n (X_i - \bar{X})^2 \end{aligned}$$

We have shown that $Y_1 \stackrel{iid}{\sim} N(0, 1)$,

$$\text{so } \sqrt{n}\bar{X} \sim N(0, 1), \quad \bar{X} \sim N(0, \frac{1}{n})$$

$$\text{Plus, } Y_2^2 + \dots + Y_n^2 \sim \chi^2_{n-1}, \text{ so } \sum_{i=1}^n (X_i - \bar{X})^2 \sim \chi^2_{n-1}$$

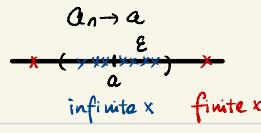
Since Y_1 is independent of (Y_2, \dots, Y_n) ,
we know \bar{X} is independent of $\sum_{i=1}^n (X_i - \bar{X})^2$.

$$X \perp\!\!\!\perp Y \Rightarrow g(X) \perp\!\!\!\perp g(Y)$$

↓

$$P(X \in A, Y \in B) = P(X \in A) \cdot P(Y \in B), \quad \forall A, B$$

$$\begin{aligned} \text{Then } P(g(X) \in A', g(Y) \in B') &= P(X \in g^{-1}(A'), Y \in g^{-1}(B')) \\ &= P(X \in g^{-1}(A')) \cdot P(Y \in g^{-1}(B')) \\ &= P(g(X) \in A') \cdot P(g(Y) \in B') \end{aligned}$$



* Modes of convergence

$Y_1, Y_2, \dots, Y_n, \dots$

① Convergence in probability

$$\lim_{n \rightarrow \infty} P(|Y_n - Y| > \varepsilon) = 0, \quad \forall \varepsilon > 0.$$

$[\forall \varepsilon > 0, \exists N \in \mathbb{N}, \text{ such that } P(|Y_n - Y| > \varepsilon) < \varepsilon \text{ for any } n \geq N]$

Ex. $Y_n = \begin{cases} 1 & \text{w/ prob. } \frac{1}{n} \\ 0 & \text{w/ prob. } 1 - \frac{1}{n} \end{cases}, \quad Y = 0 \text{ w/ prob. 1}$

$$P(|Y_n - Y| > \frac{1}{2}) = P(Y_n = 1) = \frac{1}{n}$$

* Weak law of large numbers

$$X_1, X_2, \dots, \stackrel{\text{iid}}{\sim} f(x), \quad \mu = \mathbb{E}(X_i), \quad \sigma^2 = \text{Var}(X_i)$$

$$\text{let } Y_n = \bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i, \quad n \in \{1, 2, 3, \dots\}$$

WLLN: $\bar{X}_n \xrightarrow{\text{conv in prob.}} \mu$

Pf. $\mathbb{E}(\bar{X}_n) = \mathbb{E}(X_i) = \mu, \quad \text{Var}(\bar{X}_n) = \frac{1}{n} \text{Var}(X_i)$

$$P(|\bar{X}_n - \mu| > \varepsilon) \leq \frac{\text{Var}(\bar{X}_n)}{\varepsilon^2} = \frac{\sigma^2}{n\varepsilon^2} \xrightarrow{n \rightarrow \infty} 0, \quad \forall \varepsilon$$



Chebychev's

$$P(|w - \mathbb{E}w| > \varepsilon) \leq \frac{\text{Var}(w)}{\varepsilon^2}$$

② Convergence in distribution

$$Y_n \Rightarrow Y \text{ if } \lim_{n \rightarrow \infty} P(Y_n \in A) = P(Y \in A)$$

for every $A \in \mathcal{B}$ such that $P(Y \in \partial A) = 0$

Y cannot be on the boundary.

eg. $Y_n \sim \text{Beta}(\frac{1}{n}, \frac{1}{n})$,

$$Y_n \Rightarrow Y, \quad Y \sim \text{Beta}(\frac{1}{2})$$

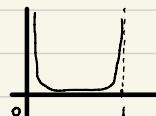
which is $Y_n \Rightarrow \text{Beta}(\frac{1}{2})$

which is $\text{Beta}(\frac{1}{n}, \frac{1}{n}) \Rightarrow \text{Beta}(\frac{1}{2})$

$$\text{We can show: } P(Y_n \in (0.5, 0.8)) \rightarrow P(Y \in (0.5, 0.8))$$

$$P(Y_n \in (0.5, 1.5)) \rightarrow P(Y \in (0.5, 1.5))$$

Check all intervals.



Shortcut:

$Y_n \Rightarrow Y$ if and only if $F_{Y_n}(y) \xrightarrow{n \rightarrow \infty} F_Y(y)$ for every $y \in \mathbb{R}$, s.t. $P(Y=y)=0$.

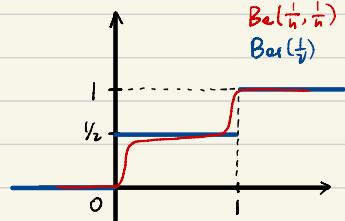
WTS:

$F_{Y_n}(y) \rightarrow F_Y(y)$ for all $y \notin \{0, 1\}$.

e.g. $(X_1, \dots, X_d) \sim \text{Multi}(n, p)$

$$(p_1, \dots, p_d) \in \Delta^d$$

$$Y_n = \sum_{j=1}^d \frac{(X_j - np_j)^2}{np_j} \Rightarrow \chi_{d-1}^2 \text{ Should show: } F_{Y_n}(y) \rightarrow F_{\chi_{d-1}^2}(y), \forall y \in \mathbb{R}.$$



* CLT

$X_1, X_2, \dots \stackrel{iid}{\sim} f(x), \mu = \mathbb{E}(X_i), \sigma^2 = \text{Var}(X_i)$

$$\sqrt{n}(\bar{X}_n - \mu) \xrightarrow{} N(0, \sigma^2)$$

$$\text{which is } \sqrt{n} \frac{\bar{X} - \mu}{\sigma} \xrightarrow{} N(0, 1)$$

$$\text{which is } \bar{X} \sim AN(\mu, \frac{\sigma^2}{n}) \quad [\text{NOT allowed: } \bar{X}_n \xrightarrow{} N(\mu, \frac{\sigma^2}{n})]$$

not fixed.
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