

Derivation of Gaussian Distribution

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Suppose we have n random variables that are independent and identically distributed

$$X_1, X_2, \dots, X_n \sim iid \quad (1)$$

which are observations of the true parameter μ (itself unobservable). Then the observation errors would be

$$X_1 - \mu, X_2 - \mu, \dots, X_n - \mu \sim iid$$

We denote the density function of a single error by $f(\cdot)$, then the joint density of errors would be

$$F(x_1, x_2, \dots, x_n) = \prod_{i=1}^n f(x_i - \mu)$$

which, from the Maximum Likelihood point of view, can also be seen as a likelihood function of the parameter μ

$$\mathcal{L}(\mu; x_1, x_2, \dots, x_n) = \prod_{i=1}^n f(x_i - \mu) \quad (2)$$

where the observation values x_1, x_2, \dots, x_n are considered as given.

Now we make three basic assumptions for derivation:

1. $f(x) > 0, x \in \mathbb{R}$
2. $f(\cdot)$ is continuous differentiable
3. $\mathcal{L}(\mu)$ is maximized at $\mu = \bar{x} := \frac{1}{n} \sum_{i=1}^n x_i$

The first two assumptions are only for the convenience of deduction, *i.e.* the ability to take log and derivative. The last assumption is really the essential one here. It suggests that the mean of the observations should be a good estimate of the true value μ , as it maximizes the likelihood function $\mathcal{L}(\mu)$. Then, because $\log(\cdot)$ is a monotonically increasing function, we are able to perform logarithmic transformation on \mathcal{L} without changing the maximum point $\mu = \bar{x}$. According to Fermat's theorem, $\mu = \bar{x}$ should be a stationary point of $\log(\mathcal{L}(\mu))$, which means

$$\left. \frac{d \log(\mathcal{L}(\mu))}{d\mu} \right|_{\mu=\bar{x}} = 0 \quad (3)$$

From the chain rule and the computation rules of logarithm, we can deduct

$$\begin{aligned}
0 &= \frac{d \log \left(\prod_{i=1}^n f(x_i - \mu) \right)}{d\mu} \Big|_{\mu=\bar{x}} \\
&= \frac{d \sum_{i=1}^n \log(f(x_i - \mu))}{d\mu} \Big|_{\mu=\bar{x}} \\
&= \sum_{i=1}^n \frac{d \log(f(x_i - \mu))}{d\mu} \Big|_{\mu=\bar{x}} \\
&= \sum_{i=1}^n \frac{1}{f(x_i - \mu)} \frac{df(x_i - \mu)}{d\mu} \Big|_{\mu=\bar{x}} \\
&= \sum_{i=1}^n \frac{f'(x_i - \bar{x})}{f(x_i - \bar{x})}
\end{aligned}$$

Denote $g(\cdot) := \frac{f'(\cdot)}{f(\cdot)}$, then $g(\cdot)$ is a continuous function. And now we have

$$\sum_{i=1}^n g(x_i - \bar{x}) = 0 \quad (4)$$

Before going further, we need to note here that equation (4) holds for any arbitrary values of $\vec{x} = (x_1, x_2, \dots, x_n) \in \mathbb{R}^n$ and $\forall n \in \mathbb{Z}^+$. The reason is that we assume \bar{x} to be the maximum point of $\mathcal{L}(\mu)$ regardless of the number and values of observations \vec{x} . The only hidden constraint here is $\bar{x} = \frac{1}{n}(x_1 + x_2 + \dots + x_n)$.

Consider the case $n = 2$, it is straightforward that

$$x_1 - \bar{x} = \frac{1}{2}(x_1 - x_2) = -\frac{1}{2}(x_2 - x_1) = x_2 - \bar{x}$$

together with the arbitrariness of the values of x_1 and x_2 , equation (4) turns into

$$g(x) + g(-x) = 0, \forall x \in \mathbb{R} \quad (5)$$

which means $g(\cdot)$ is an odd function. Then consider the case $n = 3$, it is easy to see that

$$\begin{aligned}
x_1 - \bar{x} &= \frac{1}{3}(2x_1 - x_2 - x_3) \\
x_2 - \bar{x} &= \frac{1}{3}(2x_2 - x_1 - x_3) \\
x_3 - \bar{x} &= \frac{1}{3}(2x_3 - x_1 - x_2) \\
\implies -(x_3 - \bar{x}) &= (x_1 - \bar{x}) + (x_2 - \bar{x})
\end{aligned} \quad (6)$$

given the arbitrariness of x_1 , x_2 and x_2 , and that $g(\cdot)$ is odd, equation (4) turns into

$$g(x) + g(y) = g(x + y), \forall x, y \in \mathbb{R} \quad (7)$$

Equation (7) is **Cauchy's functional equation**. Given that $g(\cdot)$ is continuous, the equation has a single family of solutions

$$g(x) = ax, a \in \mathbb{R} \quad (8)$$

The proof is in the appendix.

By the definition of $g(\cdot)$, we have the following ordinary differential equation

$$\frac{f'(x)}{f(x)} = ax \quad (9)$$

$$\implies \left(\log(f(x)) \right)' = ax$$

$$\implies f(x) = c \exp\left(\frac{a}{2}x^2\right), c > 0 \quad (10)$$

where c is the constant of integral. In order to be a density function, f needs to integrate to 1 on the real axis, therefore a must be negative for the (infinite limited) integral to converge, and also

$$\begin{aligned} 1 &= \int_{-\infty}^{\infty} f(x) dx = \sqrt{\frac{2\pi c^2}{|a|}} \\ \implies c &= \sqrt{\frac{|a|}{2\pi}} \end{aligned} \quad (11)$$

The computation of the integral is in the appendix. Now we have

$$f(x) = \sqrt{\frac{|a|}{2\pi}} \exp\left(\frac{|a|}{2}x^2\right), a < 0$$

And we want to denote $\sigma^2 := \frac{1}{|a|}$. In that way,

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{x^2}{2\sigma^2}\right), \forall \sigma^2 \in \mathbb{R} \quad (12)$$

and for any random variable X with density f , we have

$$\text{var}(X) = \int_{-\infty}^{\infty} x^2 f(x) dx = \sigma^2$$

Voilà, there is our Gaussian distribution.

Appendix

A. Suppose $g(\cdot)$ is a continuous function which satisfies

$$g(x) + g(y) = g(x + y), \forall x, y \in \mathbb{R}$$

prove that function g is scalar multiplication.

Proof: First, we want to show that the conclusion holds for any given rational number r .

(a) $r = 0$, obviously $g(0) = 0$

(b) $r > 0$, $\exists n, m \in \mathbb{Z}^+$ s.t. $r = \frac{n}{m}$

$$\begin{aligned} mg(r) &= mg\left(\frac{n}{m}\right) = g\left(\frac{n}{m}\right) + g\left(\frac{n}{m}\right) + \cdots + g\left(\frac{n}{m}\right) = g\left(\frac{n}{m} + \frac{n}{m} + \cdots + \frac{n}{m}\right) \\ &= g(n) = g(1) + g(1) + \cdots + g(1) = ng(1) \\ &\implies g(r) = \frac{n}{m}g(1) = rg(1) \end{aligned}$$

therefore we have $g(r) = g(1) \cdot r, \forall r \in \mathbb{Q}^+$

(c) $r < 0$, can be derived directly from the (b) and the fact that g is an odd function.

Then with the continuity of g and the fact that \mathbb{Q} is dense in \mathbb{R} , we can derive the conclusion.

B. Calculate $\int_{-\infty}^{\infty} c \exp\left(\frac{a}{2}x^2\right)dx$ where $c > 0, a < 0$.

With the polar coordinates

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp\left(-\frac{x^2 + y^2}{2}\right) dx dy = \int_0^{2\pi} \int_{-\infty}^{\infty} \exp\left(-\frac{r^2}{2}\right) r dr d\theta = \int_0^{2\pi} d\theta = 2\pi$$

And the left hand side

$$\begin{aligned} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp\left(-\frac{x^2 + y^2}{2}\right) dx dy &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp\left(-\frac{x^2}{2}\right) \exp\left(-\frac{y^2}{2}\right) dx dy \\ &= \left(\int_{-\infty}^{\infty} \exp\left(-\frac{x^2}{2}\right) dx\right) \left(\int_{-\infty}^{\infty} \exp\left(-\frac{y^2}{2}\right) dy\right) \\ &= \left(\int_{-\infty}^{\infty} \exp\left(-\frac{x^2}{2}\right) dx\right)^2 \end{aligned}$$

Therefore,

$$\int_{-\infty}^{\infty} \exp\left(-\frac{x^2}{2}\right) dx = \sqrt{2\pi}$$

Then with the substitution of $u = \sqrt{|a|}x$, we can easily derive that

$$\int_{-\infty}^{\infty} c \exp\left(\frac{a}{2}x^2\right) dx = c \sqrt{\frac{2\pi}{|a|}}$$

References

Cauchy's functional equation - Wikipedia
<https://zhuanlan.zhihu.com/p/24437232>