# STA457 Project

# Xing Yu Wang

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```
library(dplyr)
library(tidyverse)
library(readr)
library(lubridate)
library(forecast)
## Warning: package 'forecast' was built under R version 4.3.3
library(astsa)
## Warning: package 'astsa' was built under R version 4.3.3
library(tseries)
## Warning: package 'tseries' was built under R version 4.3.3
library(mgcv)
library(Metrics)
## Warning: package 'Metrics' was built under R version 4.3.3
library(ggplot2)
library(xgboost)
## Warning: package 'xgboost' was built under R version 4.3.3
# library(XGBClassifier)
library(rugarch)
## Warning: package 'rugarch' was built under R version 4.3.3
library(tibble)
library(xts)
## Warning: package 'xts' was built under R version 4.3.3
## Warning: package 'zoo' was built under R version 4.3.3
1. EDA
```

```
price = read.csv("./Daily Prices_ICCO.csv")
weather = read.csv("./Ghana_data.csv")
USD_GHS_Historical_Data = read.csv("./USD_GHS Historical Data.csv")
```

### 1.1 Clean Data

```
weather <- weather |> dplyr::select(DATE, TAVG)
exchangerate <- USD_GHS_Historical_Data |> dplyr::select(Date, Price)

colnames(price) [colnames(price) == 'ICCO.daily.price..US..tonne.'] <- 'Daily_Price'
colnames(weather) [colnames(weather) == 'DATE'] <- 'Date'
colnames(weather) [colnames(weather) == 'TAVG'] <- 'Avg_Temp'
colnames(exchangerate) [colnames(exchangerate) == 'Price'] <- 'exchange_rate'</pre>
```

# 1.2 Check duplicated values

```
price |> group_by(Date) |> filter(n() > 1) |> ungroup()
## # A tibble: 8 x 2
##
            Daily_Price
    Date
     <chr>
                <chr>
## 1 31/01/2024 4,798.20
## 2 31/01/2024 10,888.05
## 3 30/01/2024 4,775.17
## 4 30/01/2024 10,676.42
## 5 09/01/2024 4,171.24
## 6 09/01/2024 4,171.24
## 7 15/12/2023 4,272.15
## 8 15/12/2023 4,272.15
price <- price |> filter(!(Date == "31/01/2024" & Daily_Price == "10,888.05"))
price <- price |> filter(!(Date == "30/01/2024" & Daily_Price == "10,676.42"))
price <- distinct(price)</pre>
```

### 1.3 Convert to Time Series Data

### 1.3.1 price Dataset

```
price$Date <- as.Date(price$Date, format="%d/%m/%Y")
price$Daily_Price <- as.numeric(gsub(",", "", price$Daily_Price))
price_month <- price |> mutate(Time = floor_date(Date, "month")) |> group_by(Time) |>
    summarise(month_Price = mean(Daily_Price, na.rm = TRUE)) |> ungroup()

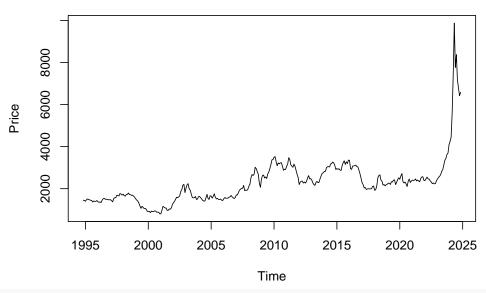
summary(price)

## Date Daily_Price
```

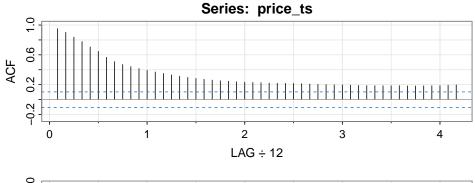
```
## Min. :1994-10-03 Min. : 774.1
## 1st Qu.:2002-05-16 1st Qu.: 1557.8
## Median :2009-12-17 Median : 2202.0
## Mean :2009-12-17 Mean : 2350.1
## 3rd Qu.:2017-07-24 3rd Qu.: 2738.1
## Max. :2025-02-27 Max. :11984.7
price_ts <- ts(price_month$month_Price, start = c(1994, 11), end = c(2024, 11), frequency = 12)</pre>
```

# plot(price\_ts, main="Monthly Price Time Series", ylab="Price", xlab="Time")

# **Monthly Price Time Series**



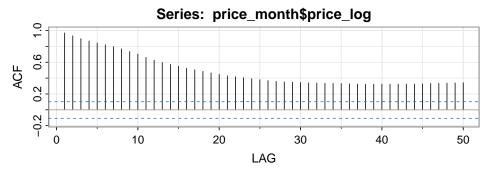
acf2(price\_ts, 50)

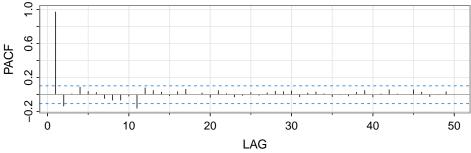




```
[,2]
                                                                                                               [,3]
                                                                                                                                                   [,4]
                                                                                                                                                                                   [,5] [,6]
                                                                                                                                                                                                                                                   [,7] [,8] [,9] [,10] [,11] [,12] [,13]
##
                                              [,1]
                                            0.95 \quad 0.90 \quad 0.84 \quad 0.78 \quad 0.71 \quad 0.65 \quad 0.57 \quad 0.51 \quad 0.47 \quad 0.44 \quad 0.42 \quad 0.39 \quad 0.37 \quad 0.51 \quad 0.47 \quad 0.44 \quad 0.42 \quad 0.39 \quad 0.37 \quad 0.44 \quad 0.42 \quad 0.39 \quad 0.37 \quad 0.44 \quad 0.42 \quad 0.39 \quad 0.37 \quad 0.44 \quad 0.42 \quad 0.49 \quad 
## ACF
## PACF 0.95 -0.04 -0.16 -0.01 -0.14 0.08 -0.24 0.18 0.23 0.00 0.01 -0.13 0.08
                                              [,14] [,15] [,16] [,17] [,18] [,19] [,20] [,21] [,22] [,23] [,24] [,25]
## ACF
                                                  0.35 0.33 0.31
                                                                                                                                                   0.30 0.28
                                                                                                                                                                                                                      0.27
                                                                                                                                                                                                                                                          0.26 0.25 0.25 0.24 0.23
                                                                                                                                                  0.05 0.10 -0.07
                                                                                                                                                                                                                                                          0.00 0.03 -0.04 0.00 -0.02 0.10
## PACF -0.07 -0.07 0.09
##
                                              [,26] [,27] [,28] [,29] [,30] [,31] [,32] [,33] [,34] [,35] [,36] [,37]
## ACF
                                                                                   0.22
                                                                                                                       0.22
                                                                                                                                                     0.22
                                                                                                                                                                                        0.22 0.21
                                                                                                                                                                                                                                                            0.21
                                                                                                                                                                                                                                                                                                0.20
                                                                                                                                                                                                                                                                                                                                   0.20 0.20 0.20
## PACF 0.02 -0.04
                                                                                                                    0.04 0.00 0.03 -0.08 0.00 0.05 0.01 0.03 -0.02 0.04
```

```
[,38] [,39] [,40] [,41] [,42] [,43] [,44] [,45] [,46] [,47] [,48] [,49]
         0.19 \quad 0.19 \quad 0.19 \quad 0.18 \quad 0.18 \quad 0.19 \quad 0.18 \quad 0.18 \quad 0.18 \quad 0.18 \quad 0.19 \quad 0.19
## ACF
## PACF -0.02 -0.02 0.01 0.00 0.06 0.00 -0.04 0.03 0.04 -0.01 0.00 0.05
##
         [,50]
## ACF
         0.20
## PACF 0.03
ndiffs(price_ts)
## [1] 1
price_month$price_log <- log(price_month$month_Price)</pre>
adf.test(price month$price log)
##
##
    Augmented Dickey-Fuller Test
##
## data: price_month$price_log
## Dickey-Fuller = -1.736, Lag order = 7, p-value = 0.6883
## alternative hypothesis: stationary
acf2(price_month$price_log, 50)
```





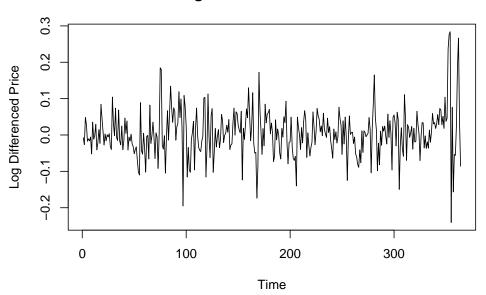
## [,1][,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11] [,12] [,13]  $0.97 \quad 0.93 \ 0.90 \ 0.87 \ 0.84 \ 0.82 \quad 0.80 \quad 0.77 \quad 0.74 \quad 0.70 \quad 0.66 \quad 0.63$ PACF 0.97 -0.14 0.01 0.09 0.04 0.02 -0.05 -0.07 -0.07 -0.02 -0.16 0.08 0.05 ## [,14] [,15] [,16] [,17] [,18] [,19] [,20] [,21] [,22] [,23] [,24] [,25] ## ACF 0.57 0.55 0.53 0.51 0.49 0.47 0.45 0.43 0.42 0.40 0.39 0.38 PACF 0.03 -0.01 0.03 0.06 0.00 0.02 -0.03 0.05 0.01 -0.03 -0.01 0.02 [,26] [,27] [,28] [,29] [,30] [,31] [,32] [,33] [,34] [,35] [,36] [,37] 0.36 0.35 0.35 0.34 0.34 0.34 0.34 0.33 0.33 ## ACF 0.37 0.33 0.32 ## PACF -0.01 0.02 0.04 0.03 0.04 -0.03 0.02 0.03 0.01 -0.02 0.00 -0.02 [,38] [,39] [,40] [,41] [,42] [,43] [,44] [,45] [,46] [,47] [,48] [,49] ## ACF 0.32 0.32 0.32 0.32 0.33 0.33 0.33 0.33 0.33 0.34 ## PACF 0.03 0.05 -0.03 0.01 0.06 0.01 0.00 0.06 0.03 -0.02 0.00 0.03

```
## [,50]
## ACF 0.34
## PACF 0.00
```

Hence, we want to difference the price data.

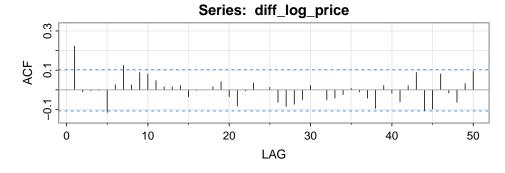
```
diff_log_price = diff(price_month$price_log)
ts.plot(diff_log_price, main = "Log Differenced Price Data", ylab = "Log Differenced Price")
```

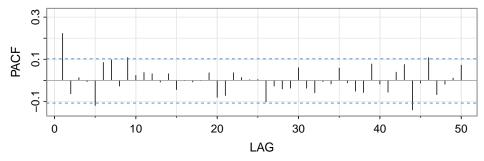
# **Log Differenced Price Data**



```
adf.test(diff_log_price)
```

```
## Warning in adf.test(diff_log_price): p-value smaller than printed p-value
##
## Augmented Dickey-Fuller Test
##
## data: diff_log_price
## Dickey-Fuller = -6.1385, Lag order = 7, p-value = 0.01
## alternative hypothesis: stationary
acf2(diff_log_price, 50)
```





```
##
       [,1] [,2] [,3] [,4] [,5] [,6] [,7]
                                            [,8] [,9] [,10] [,11] [,12] [,13]
## ACF 0.22 -0.01 0.00 0.00 -0.11 0.03 0.12 0.02 0.09 0.08 0.05 0.02 0.02
## PACF 0.22 -0.06 0.01 -0.01 -0.12 0.09 0.10 -0.03 0.11 0.02 0.04 0.03 -0.01
       [,14] [,15] [,16] [,17] [,18] [,19] [,20] [,21] [,22] [,23] [,24] [,25]
##
                       0 0.00 0.02 0.04 -0.03 -0.08 -0.01 0.03
        0.02 - 0.04
                       0 -0.01 0.00 0.04 -0.08 -0.07 0.04 0.01
  PACF 0.03 -0.04
       [,26] [,27] [,28] [,29] [,30] [,31] [,32] [,33] [,34] [,35] [,36] [,37]
## ACF
       -0.06 -0.09 -0.07 -0.05 0.02 0.00 -0.05 -0.04 -0.02 0.01 -0.01 -0.04
## PACF -0.10 -0.03 -0.04 -0.04 0.06 -0.04 -0.06 -0.01 -0.02 0.06 -0.01 -0.05
       [,38] [,39] [,40] [,41] [,42] [,43] [,44] [,45] [,46] [,47] [,48] [,49]
## ACF -0.09 0.02 -0.02 -0.06 0.02 0.09 -0.11 -0.10 0.08 -0.02 -0.06 0.03
## PACF -0.06 0.08 -0.02 -0.06 0.04 0.08 -0.14 -0.01 0.11 -0.07 -0.02 0.01
##
       [,50]
        0.09
## ACF
## PACF 0.07
```

# 1.3.2 ghana Dataset

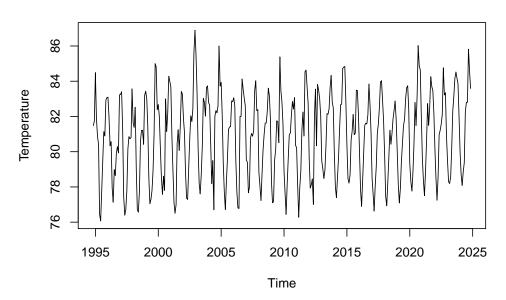
```
weather$Date <- as.Date(weather$Date)
weather$Avg_Temp <- as.numeric(gsub("", "", weather$Avg_Temp))
weather_month <- weather |> mutate(Time = floor_date(Date, "month")) |> group_by(Time) |>
summarise(Avg_Temp = mean(Avg_Temp, na.rm = TRUE)) |> ungroup()
```

### summary(weather month)

```
Time
##
                             Avg_Temp
    Min.
           :1990-01-01
                          Min.
                                  :76.07
##
    1st Qu.:1998-09-23
                          1st Qu.:78.90
    Median :2007-07-16
                          Median :81.20
           :2007-06-22
                          Mean
                                  :80.97
##
    Mean
                          3rd Qu.:82.82
    3rd Qu.:2016-03-08
##
    Max.
           :2024-11-01
                          Max.
                                 :86.90
```

```
weather_ts <- ts(weather_month$Avg_Temp, start = c(1994, 11), end = c(2024, 11), frequency = 12)
ts.plot(weather_ts, main="Monthly Average Temperature Time Series", ylab="Temperature", xlab="Time")</pre>
```

# **Monthly Average Temperature Time Series**

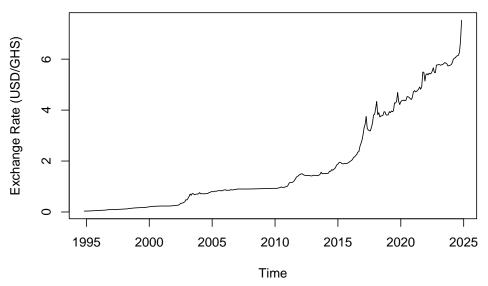


### 1.3.3 exchange Data

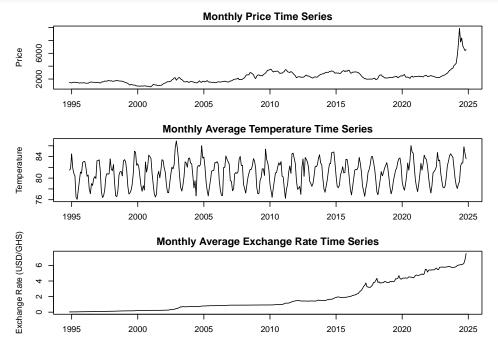
```
:1992-03-01
                                : 0.0338
##
                          Min.
   Min.
    1st Qu.:2000-06-01
                          1st Qu.: 0.5400
##
                         Median : 1.1595
##
   Median :2008-09-01
                                 : 2.8314
##
   Mean
           :2008-08-31
                          Mean
##
    3rd Qu.:2016-12-01
                          3rd Qu.: 4.2805
## Max.
           :2025-03-01
                          Max.
                                 :16.2500
rate_ts \leftarrow ts(rate_month$exchange_rate, start = c(1994, 11), end = c(2024, 11), frequency = 12)
```

ts.plot(rate\_ts, main="Monthly Average Exchange Rate Time Series", ylab="Exchange Rate (USD/GHS)", xlab

# **Monthly Average Exchange Rate Time Series**

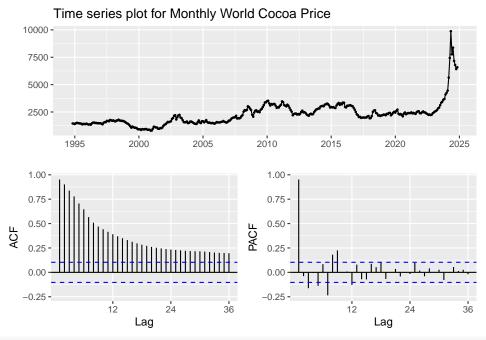


```
par(mfrow=c(3,1), mar = c(3, 4, 2, 2))
# price
plot(price_ts, main="Monthly Price Time Series", ylab="Price", xlab="Time")
#temperature
ts.plot(weather_ts, main="Monthly Average Temperature Time Series", ylab="Temperature", xlab="Time")
# exchange rate
ts.plot(rate_ts, main="Monthly Average Exchange Rate Time Series", ylab="Exchange Rate (USD/GHS)", xlab
```

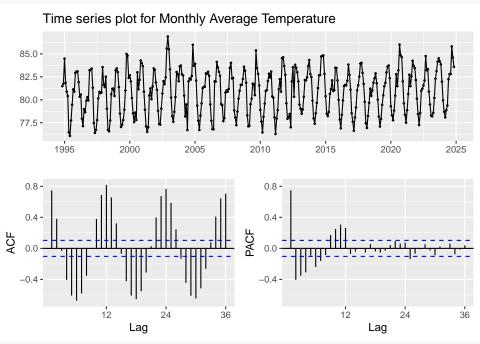


# 1.4 Time series plots for data

ggtsdisplay(price\_ts, main="Time series plot for Monthly World Cocoa Price")

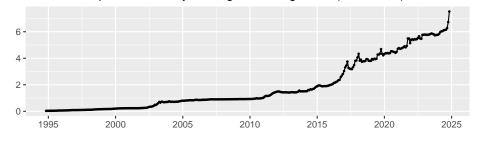


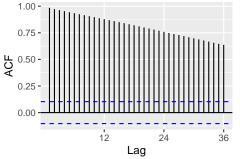
ggtsdisplay(weather\_ts, main="Time series plot for Monthly Average Temperature")

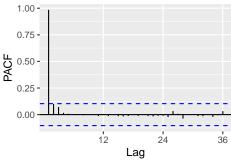


ggtsdisplay(rate\_ts, main="Time series plot for Monthly Average Exchange Rate(USD/GHS)")

# Time series plot for Monthly Average Exchange Rate(USD/GHS)

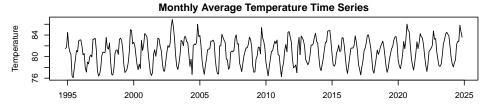


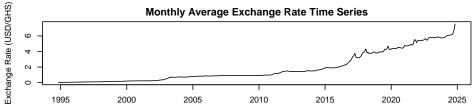




```
par(mfrow=c(3,1), mar = c(3, 4, 2, 2))
# price
plot(price_ts, main="Monthly Price Time Series", ylab="Price", xlab="Time")
#temperature
ts.plot(weather_ts, main="Monthly Average Temperature Time Series", ylab="Temperature", xlab="Time")
# exchange rate
ts.plot(rate_ts, main="Monthly Average Exchange Rate Time Series", ylab="Exchange Rate (USD/GHS)", xlab
```



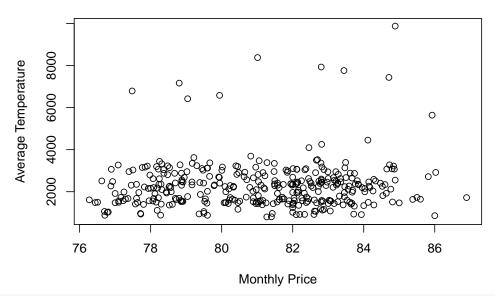




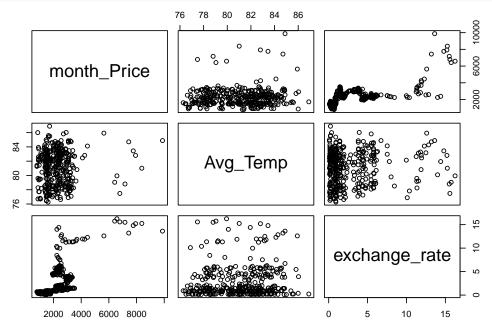
# 1.5 Combine Datasets

```
data <- price_month |> left_join(weather_month, by = "Time") |> left_join(rate_month, by = "Time")
data <- data |> mutate(log_price = log(month_Price), diff_log_price =
```

# **Daily Price vs. Avg Temperature**





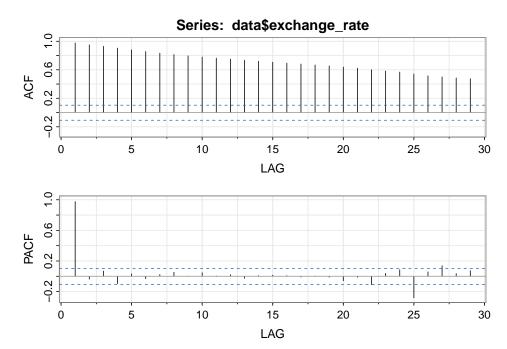


# 1.6 Stationary Check

```
adf.test(data$Avg_Temp)
```

```
## Warning in adf.test(data$Avg_Temp): p-value smaller than printed p-value
##
##
   Augmented Dickey-Fuller Test
##
## data: data$Avg_Temp
## Dickey-Fuller = -12.411, Lag order = 7, p-value = 0.01
## alternative hypothesis: stationary
adf.test(data$exchange rate)
## Warning in adf.test(data$exchange_rate): p-value greater than printed p-value
    Augmented Dickey-Fuller Test
##
##
## data: data$exchange_rate
## Dickey-Fuller = 0.7342, Lag order = 7, p-value = 0.99
## alternative hypothesis: stationary
adf.test(log(data$exchange_rate))
##
    Augmented Dickey-Fuller Test
##
##
## data: log(data$exchange_rate)
## Dickey-Fuller = -2.8782, Lag order = 7, p-value = 0.2063
## alternative hypothesis: stationary
adf.test(diff(log(data$exchange_rate)))
## Warning in adf.test(diff(log(data$exchange_rate))): p-value smaller than
## printed p-value
##
##
    Augmented Dickey-Fuller Test
## data: diff(log(data$exchange_rate))
## Dickey-Fuller = -5.3335, Lag order = 7, p-value = 0.01
## alternative hypothesis: stationary
Since monthly average temperature is already stationary, we would do take the differenced and log-transformed
exchange rate as our exogenous factors.
```

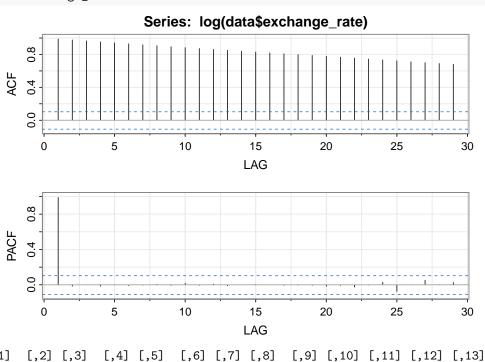
```
acf2(data$exchange rate)
```



```
[,1]
##
             [,2] [,3] [,4] [,5]
                                 [,6] [,7] [,8] [,9] [,10] [,11] [,12] [,13]
                                                0.8 0.78
       0.98 0.95 0.93 0.9 0.88 0.86 0.83 0.81
                                                          0.76 0.75 0.73
## PACF 0.98 -0.04 0.07 -0.1 0.03 -0.03 0.02 0.05
                                                0.0 0.05 0.00
                                                                 0.02 -0.03
       [,14] [,15] [,16] [,17] [,18] [,19] [,20] [,21] [,22] [,23] [,24] [,25]
##
        0.72 0.71 0.69 0.68
                              0.67 0.65 0.64 0.62 0.60 0.58
                                                                 0.57 0.54
## ACF
                               0.00 -0.01 -0.06 -0.01 -0.11 0.04
## PACF
                                                                 0.09 -0.28
       0.01
             0.02 0.01 -0.01
       [,26] [,27] [,28] [,29]
##
## ACF
        0.52
             0.50
                   0.49
                        0.47
## PACF 0.06 0.14 0.03 0.07
```

### acf2(log(data\$exchange\_rate))

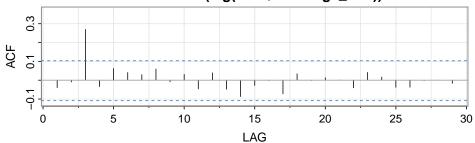
##

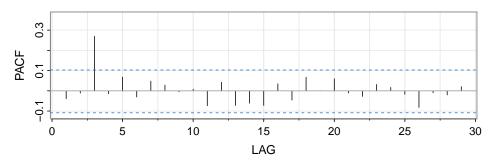


```
## ACF 0.99 0.98 0.97 0.95 0.94 0.93 0.92 0.91 0.89 0.88 0.87 0.86 0.85
  PACF 0.99 -0.02 0.00 -0.02 0.00 -0.01 0.00 0.01 -0.01 0.02 -0.01 0.01 -0.01
       [,14] [,15] [,16] [,17] [,18] [,19] [,20] [,21] [,22] [,23] [,24] [,25]
              0.83 0.82 0.81
                                0.8 0.79 0.78 0.77 0.76 0.74 0.73 0.72
## ACF
        0.84
              0.00
                   0.00 -0.01
                                0.0 -0.01 -0.02 -0.01 -0.03 0.00 0.03 -0.08
       [,26] [,27] [,28] [,29]
##
              0.70
                    0.69
## ACF
        0.71
## PACF
       0.00
             0.05
                   0.00 0.03
```

### acf2(diff(log(data\$exchange\_rate)))

# Series: diff(log(data\$exchange\_rate))





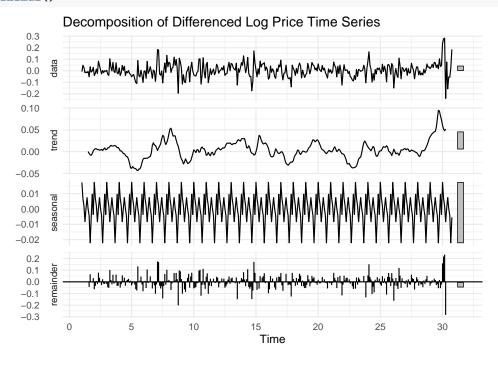
```
##
        [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10] [,11] [,12] [,13]
       -0.04 -0.01 0.27 -0.03 0.06 0.04 0.03 0.06 -0.01 0.03 -0.05 0.04 -0.05
  PACF -0.04 -0.01 0.27 -0.01 0.07 -0.03 0.05 0.03 0.00 0.01 -0.07 0.04 -0.07
       [,14] [,15] [,16] [,17] [,18] [,19] [,20] [,21] [,22] [,23] [,24] [,25]
##
## ACF
       -0.09 -0.03 0.00 -0.07
                               0.03
                                        0 0.01 0.00 -0.04 0.04 0.02 -0.04
## PACF -0.06 -0.07 0.03 -0.05
                               0.07
                                        0 0.06 -0.01 -0.03 0.03 0.02 -0.02
       [,26] [,27] [,28] [,29]
## ACF
       -0.04 0.00 0.00 -0.02
## PACF -0.08 -0.01 -0.02 0.02
```

ACF shows similar trend, where only differenced log-transformed exchange rate is stationary. Hence, this differenced and log-transformed exchange rate will be used as one of the external(exogenous) regressors in ARIMAX and GARCHX.

### adf.test(data\$month\_Price)

```
##
## Augmented Dickey-Fuller Test
##
## data: data$month_Price
## Dickey-Fuller = -1.7041, Lag order = 7, p-value = 0.7017
## alternative hypothesis: stationary
```

```
adf.test(data$log_price)
##
##
    Augmented Dickey-Fuller Test
##
## data: data$log_price
## Dickey-Fuller = -2.3875, Lag order = 7, p-value = 0.4133
## alternative hypothesis: stationary
adf.test(data$diff_log_price)
## Warning in adf.test(data$diff_log_price): p-value smaller than printed p-value
##
##
    Augmented Dickey-Fuller Test
##
## data: data$diff_log_price
## Dickey-Fuller = -6.2103, Lag order = 7, p-value = 0.01
## alternative hypothesis: stationary
Since only the diff_log_price is stationary, we choose differenced monthly log price when fitting the model.
diff_price_ts <- ts(data$diff_log_price, frequency = 12)</pre>
autoplot(decompose(diff_price_ts, type="additive")) +
  ggtitle("Decomposition of Differenced Log Price Time Series") +
  theme_minimal()
```



# 1.7 Split data

```
data <- data[order(data$Time), ]
cutoff <- floor(0.7 * nrow(data))
trainSet <- data[1:cutoff, ]
testSet <- data[(cutoff+1):nrow(data), ]</pre>
```

# 2. Method

### 2.1 ETS Model

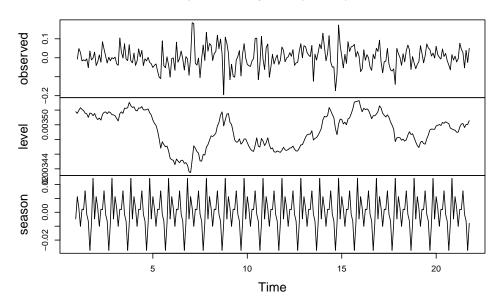
ETS is a purely univariate model and cannot directly handle external regressors.

```
data_train_ts <- ts(trainSet$diff_log_price, frequency = 12)</pre>
```

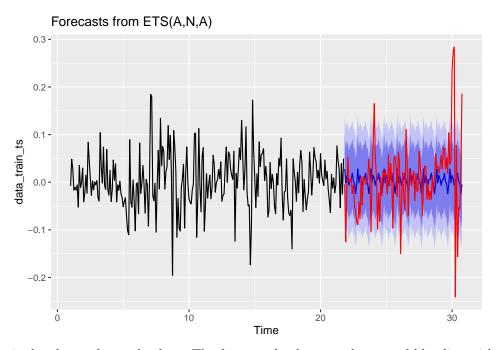
```
2.1.1 Fit Model
ets_model <- ets(data_train_ts, model = "ANA")</pre>
ets_zmodel <- ets(data_train_ts, model = "ZZZ") # Automatically selects best model
summary(ets_model)
## ETS(A,N,A)
##
## Call:
## ets(y = data_train_ts, model = "ANA")
##
     Smoothing parameters:
##
       alpha = 1e-04
##
       gamma = 1e-04
##
##
     Initial states:
       1 = 0.0035
##
       s = -0.0048 \ 0.0244 \ -0.008 \ -0.0274 \ -0.0064 \ -0.0014
##
              0.0154 0.0019 0.0022 -0.0101 0.0029 0.0112
##
##
##
     sigma: 0.057
##
                              BIC
##
         AIC
                   AICc
  -36.76439 -34.71311 16.05752
##
## Training set error measures:
##
                                 RMSE
                                              MAE
                                                        MPE
                                                                MAPE
                                                                           MASE
## Training set -0.000482889 0.05534 0.04218605 116.6964 180.0324 0.6834589
##
## Training set 0.1729102
summary(ets_zmodel)
## ETS(A,N,N)
##
## Call:
## ets(y = data_train_ts, model = "ZZZ")
##
##
     Smoothing parameters:
##
       alpha = 1e-04
##
##
     Initial states:
##
       1 = 0.0029
##
##
     sigma: 0.0569
##
                              BIC
##
         AIC
                   AICc
```

```
## -48.96308 -48.86552 -38.39869
##
## Training set error measures:
## Training set 1.567182e-05 0.05666171 0.04285329 109.1957 114.6766 0.694269
## ACF1
## Training set 0.1682833
plot(ets_model)
```

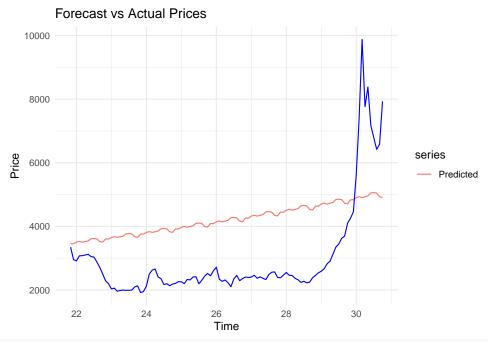
# Decomposition by ETS(A,N,A) method



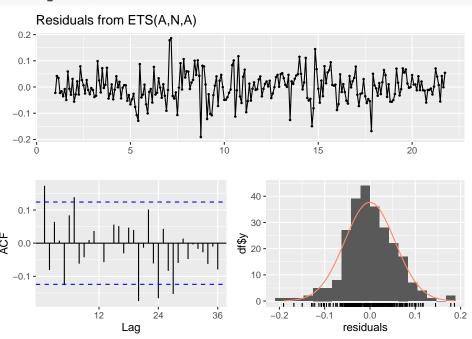
# 2.1.2 Forecasting and Plotting



The red line is the observed actual values. The forecasted values are the central blue line within the blue shaded prediction intervals.

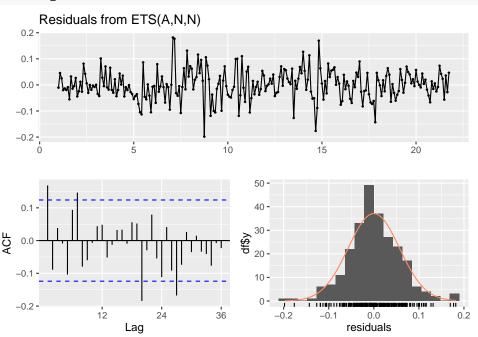


# checkresiduals(ets\_model)



```
##
## Ljung-Box test
##
## data: Residuals from ETS(A,N,A)
## Q* = 46.672, df = 24, p-value = 0.003672
##
## Model df: 0. Total lags used: 24
```

### checkresiduals(ets\_zmodel)



```
##
## Ljung-Box test
##
## data: Residuals from ETS(A,N,N)
## Q* = 42.424, df = 24, p-value = 0.01156
##
## Model df: 0. Total lags used: 24
```

# 2.2 ARIMAX Model

Recall that in Section 1.3.1, we have tested the acf and adf.test, and determined that we would be using the differenced price data. To fit the trainset, we evaluate p and q for ARIMA model.

```
adf.test(trainSet$log_price)
```

```
##
## Augmented Dickey-Fuller Test
##
## data: trainSet$log_price
## Dickey-Fuller = -2.5744, Lag order = 6, p-value = 0.334
## alternative hypothesis: stationary
```

Next, we check if applying 1st differencing is good enough

```
adf.test(diff(trainSet$month_Price))
```

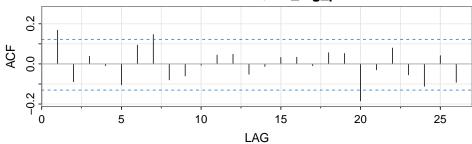
```
## Warning in adf.test(diff(trainSet$month_Price)): p-value smaller than printed
## p-value
##
## Augmented Dickey-Fuller Test
##
## data: diff(trainSet$month_Price)
## Dickey-Fuller = -5.2038, Lag order = 6, p-value = 0.01
```

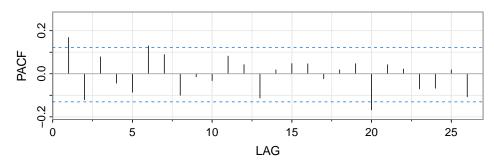
### ## alternative hypothesis: stationary

P-value is smaller than 0.01 for differenced log price, and we are

### acf2(trainSet\$diff\_log\_price)

# Series: trainSet\$diff\_log\_price





### adf.test(data\$Avg\_Temp)

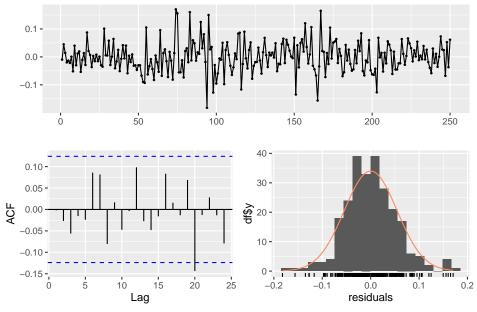
```
## Warning in adf.test(data$Avg_Temp): p-value smaller than printed p-value
##
## Augmented Dickey-Fuller Test
##
## data: data$Avg_Temp
## Dickey-Fuller = -12.411, Lag order = 7, p-value = 0.01
## alternative hypothesis: stationary
adf.test(diff(log(data$exchange_rate)))
```

```
## Warning in adf.test(diff(log(data$exchange_rate))): p-value smaller than
## printed p-value
##
## Augmented Dickey-Fuller Test
##
## data: diff(log(data$exchange_rate))
## Dickey-Fuller = -5.3335, Lag order = 7, p-value = 0.01
```

### 2.2.1 Fit ARIMAX Model

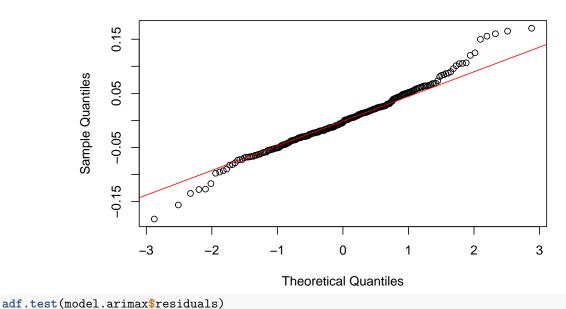
```
dl.rate.train <- c(0, diff(log(trainSet$exchange_rate)))</pre>
xreg_matrix <- cbind(trainSet$Avg_Temp, dl.rate.train)</pre>
colnames(xreg_matrix) <- c("Avg_Temp", "dl_exchange_rate")</pre>
p <- 0:10
q < -0:10
aic.arimax <- matrix(0, length(p), length(q))</pre>
for (i in 1:length(p)) {
 for (j in 1:length(q)) {
      modij = Arima(trainSet$diff_log_price, order = c(p[i], 0, q[j]),
                    method = "ML", xreg=xreg_matrix)
      aic.arimax[i, j] = AIC(modij)
 }
}
j.arimax <- ceiling(which.min(aic.arimax) / length(p))</pre>
i.arimax <- which.min(aic.arimax) - (j.arimax-1)*length(p)
sprintf("Selected order for ARIMAX: p = %d, q = %d", p[i.arimax], q[j.arimax])
## [1] "Selected order for ARIMAX: p = 2, q = 3"
model.arimax <- Arima(trainSet$diff_log_price, order=c(2,0,3), xreg = xreg_matrix)</pre>
summary(model.arimax)
## Series: trainSet$diff_log_price
## Regression with ARIMA(2,0,3) errors
##
## Coefficients:
##
             ar1
                      ar2
                              ma1
                                       ma2
                                               ma3 intercept Avg_Temp
##
         -0.2295 -0.8933 0.4381 0.9099 0.2695
                                                      -0.0090
                                                                 0.0001
## s.e.
        0.0716
                  0.0884 0.0907 0.0974 0.0622
                                                       0.1392
                                                                  0.0017
         dl_exchange_rate
##
                   0.0398
                   0.1033
## s.e.
##
## sigma^2 = 0.003047: log likelihood = 373.33
## AIC=-728.67 AICc=-727.92 BIC=-696.98
## Training set error measures:
                                     RMSE
                                                 MAE
                                                          MPE
                                                                   MAPE
## Training set -1.229987e-06 0.05430941 0.04164443 140.2458 203.7395 0.7523938
                       ACF1
## Training set -0.00203552
checkresiduals(model.arimax)
```

# Residuals from Regression with ARIMA(2,0,3) errors



```
##
## Ljung-Box test
##
## data: Residuals from Regression with ARIMA(2,0,3) errors
## Q* = 7.1814, df = 5, p-value = 0.2075
##
## Model df: 5. Total lags used: 10
qqnorm(model.arimax$residuals)
qqline(model.arimax$residuals, col="red")
```

# Normal Q-Q Plot



```
## Warning in adf.test(model.arimax$residuals): p-value smaller than printed
## p-value
```

```
##
## Augmented Dickey-Fuller Test
##
## data: model.arimax$residuals
## Dickey-Fuller = -5.2873, Lag order = 6, p-value = 0.01
## alternative hypothesis: stationary
```

- ADF Test on ARIMAX Model Residuals: Failed to reject H\_0, indicating that the residuals do not exhibit significant autocorrelation.
- Histogram and QQ-Plot of Residuals: Residuals align well with the 45-degree line, suggesting normality.
- ACF of Residuals: Appears random, with all lags within the range of -0.15 to 0.1, indicating no strong autocorrelations.
- Standardized Residuals Plot: No discernible trend observed, further supporting the model's adequacy.
- Ljung-Box Test (Residuals from ARIMA(2,0,3) model):
  - $Q^* = 7.1814, df = 5, p-value = 0.2075$
  - Model degrees of freedom: 5, Total lags used: 10 Conclusion: The ARIMAX model effectively captures the trend of the training dataset.

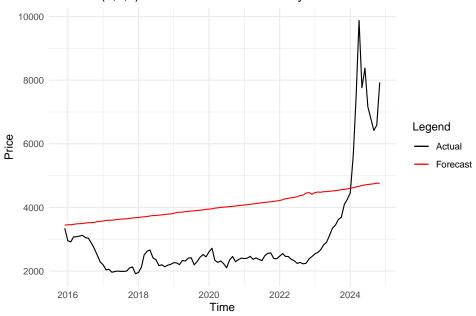
### 2.2.2 Forecasting With ARIMAX Model

Next we try to fit this ARIMAX model to forecast on testing set.

```
dl.rate.test <- c(0, diff(log(testSet$exchange_rate)))</pre>
forecast.arimax.xreg <- cbind(testSet$Avg_Temp, dl.rate.test)</pre>
colnames(forecast.arimax.xreg) <- c("Avg_Temp", "dl_exchange_rate")</pre>
forecast.arimax <- forecast(model.arimax, xreg=forecast.arimax.xreg,</pre>
                             h=nrow(testSet))
last_log_price <- tail(trainSet$log_price, 1)</pre>
# Convert back to actual price
forecast.arimax.final <- exp(cumsum(forecast.arimax$mean) + last log price)</pre>
model.arimax.fitted <- as.numeric(model.arimax$fitted)</pre>
model.arimax.fitted.converted <- exp(log(trainSet$month Price[1])</pre>
                                       + cumsum(model.arimax.fitted))
rmse(trainSet$month_Price, model.arimax.fitted.converted)
## [1] 376.4071
mae(trainSet$month_Price, model.arimax.fitted.converted)
## [1] 302.0143
mape(trainSet$month_Price, model.arimax.fitted.converted)
## [1] 0.1877152
rmse(testSet$month_Price, forecast.arimax.final)
## [1] 1701.38
mae(testSet$month_Price, forecast.arimax.final)
## [1] 1559.875
mape(testSet$month_Price, forecast.arimax.final)
## [1] 0.5865196
```

```
forecast.arimax.df <- tibble(</pre>
  Time = testSet$Time,
  Price = forecast.arimax.final
)
test.arimax.df <- tibble(</pre>
  Time = testSet$Time,
  Price = testSet$month_Price
)
ggplot() +
  geom_line(data = test.arimax.df, aes(x = Time, y = Price, color = "Actual")) +
  geom_line(data = forecast.arimax.df, aes(x = Time, y = Price, color = "Forecast")) +
    title = "ARIMAX(2,0,3) Forecast vs Actual Monthly Prices",
    y = "Price",
    x = "Time",
    color = "Legend"
  ) +
  theme_minimal() +
  scale_color_manual(values = c(
    "Actual" = "black",
    "Forecast" = "red"))
```

# ARIMAX(2,0,3) Forecast vs Actual Monthly Prices



### 2.3 GARCH Model

### 2.3.1 GARCH Parameters (With Xreg)

```
# xreg_matrix is the same as arimax
p = 0:3
q = 0:3
## select ARMA order
aic.armax.garch1 <- matrix(0, length(p), length(q))
for (i in 1:length(p)) {</pre>
```

```
for (j in 1:length(q)) {
      modij = Arima(trainSet$diff_log_price, order = c(p[i], 0, q[j]),
                     method = "ML", xreg=xreg_matrix)
      aic.armax.garch1[i, j] = AIC(modij)
 }
}
j.armax <- ceiling(which.min(aic.armax.garch1) / length(p))</pre>
i.armax <- which.min(aic.armax.garch1) - (j.armax-1)*length(p)</pre>
sprintf("Selected order for ARMA: %d, %d", p[i.armax], q[j.armax])
## [1] "Selected order for ARMA: 2, 3"
This is the same as what we have for ARIMAX
m = 1:3
n = 1:3
# dl.rate.train <- c(0, diff(log(trainSet$exchange_rate)))</pre>
# xreq_matrix <- cbind(trainSet$Avq_Temp, dl.rate.train)</pre>
# colnames(xreg_matrix) <- c("Avg_Temp", "dl_exchange_rate")</pre>
## select GARCH order
aic.armax.garch2 <- matrix(0, length(m), length(n))</pre>
for (i in 1:length(m)) {
  for (j in 1:length(n)) {
      spec = ugarchspec(variance.model=list(model="sGARCH",
                                              garchOrder=c(m[i],n[j])),
                mean.model=list(armaOrder=c(2, 3),
                                 include.mean=T,
                                 external.regressors = xreg_matrix),
                distribution.model="std")
      modij = ugarchfit(spec=spec, data = trainSet$diff_log_price,
                         solver = 'hybrid', trace = FALSE)
      aic.armax.garch2[i, j] = infocriteria(modij)[1]
 }
}
j.garch <- ceiling(which.min(aic.armax.garch2) / length(m))</pre>
i.garch <- which.min(aic.armax.garch2) - (j.garch-1)*length(m)</pre>
sprintf("Selected order for GARCH: %d, %d", m[i.garch], n[j.garch])
## [1] "Selected order for GARCH: 1, 1"
spec <- ugarchspec(variance.model=list(garchOrder=c(1,1)),</pre>
                mean.model=list(armaOrder=c(2, 3),
                                 include.mean=T,
                                 external.regressors = xreg_matrix),
                distribution.model="std")
model.armax.garch <- ugarchfit(spec, data = trainSet$diff_log_price,</pre>
                                solver = 'hybrid', trace = FALSE)
model.armax.garch
##
## *
              GARCH Model Fit
##
## Conditional Variance Dynamics
```

```
## GARCH Model : sGARCH(1,1)
## Mean Model : ARFIMA(2,0,3)
## Distribution : std
## Optimal Parameters
        Estimate Std. Error t value Pr(>|t|)
## mu
       0.020917
                0.131217
                           0.159405 0.873350
## ar1
       ## ar2
      -0.894392 0.031402 -28.482399 0.000000
      ## ma1
## ma2
      0.258470 0.064191
## ma3
                          4.026596 0.000057
0.000126 1.353592 0.175867
## omega
        0.000171
## alpha1 0.092076 0.048015 1.917657 0.055155
## beta1
                  0.065243 13.156584 0.000000
        0.858378
## shape
        6.610146
                  3.242426 2.038642 0.041486
##
## Robust Standard Errors:
        Estimate Std. Error t value Pr(>|t|)
##
                0.150659 0.138834 0.889581
## mu
        0.020917
## ar1
       ## ar2

      0.180533
      0.061615
      2.930006
      0.003390

      0.953060
      0.016170
      58.939192
      0.000000

## ma1
## ma2
## ma3
        0.258470 0.063494
                          4.070783 0.000047
## mxreg1 -0.000227 0.001881 -0.120407 0.904160
## alpha1 0.092076 0.040354 2.281704 0.022507
## beta1
                0.043445 19.757817 0.000000
         0.858378
## shape
         6.610146
                 2.907748 2.273287 0.023009
##
## LogLikelihood: 383.864
##
## Information Criteria
## -----
            -2.9749
## Akaike
## Baves
            -2.8059
## Shibata
           -2.9792
## Hannan-Quinn -2.9069
##
## Weighted Ljung-Box Test on Standardized Residuals
## -----
##
                       statistic p-value
## Lag[1]
                       0.004822 0.9446
## Lag[2*(p+q)+(p+q)-1][14] 4.789881 1.0000
## Lag[4*(p+q)+(p+q)-1][24] 9.750395 0.8540
## d.o.f=5
## HO : No serial correlation
```

```
##
## Weighted Ljung-Box Test on Standardized Squared Residuals
## -----
##
                     statistic p-value
## Lag[1]
                       0.3107 0.5772
## Lag[2*(p+q)+(p+q)-1][5] 1.5176 0.7356
## Lag[4*(p+q)+(p+q)-1][9] 3.1776 0.7293
## d.o.f=2
##
## Weighted ARCH LM Tests
## -----
            Statistic Shape Scale P-Value
## ARCH Lag[3] 0.6921 0.500 2.000 0.4054
## ARCH Lag[5] 2.3180 1.440 1.667 0.4051
## ARCH Lag[7] 3.3988 2.315 1.543 0.4416
##
## Nyblom stability test
## -----
## Joint Statistic: 1.955
## Individual Statistics:
## mu
      0.21143
## ar1 0.33766
      0.21801
## ar2
## ma1
      0.09681
## ma2 0.14306
## ma3 0.09043
## mxreg1 0.20585
## mxreg2 0.27455
## omega 0.09041
## alpha1 0.08057
## beta1 0.09023
## shape 0.12346
## Asymptotic Critical Values (10% 5% 1%)
## Joint Statistic: 2.69 2.96 3.51
## Individual Statistic: 0.35 0.47 0.75
##
## Sign Bias Test
## -----
##
                 t-value prob sig
## Sign Bias
                0.65210 0.5149
## Negative Sign Bias 0.01309 0.9896
## Positive Sign Bias 0.29202 0.7705
## Joint Effect 1.49230 0.6840
##
##
## Adjusted Pearson Goodness-of-Fit Test:
## -----
  group statistic p-value(g-1)
## 1 20 12.56 0.8603
## 2
      30 25.28
                     0.6636
## 3 40 35.12
                    0.6475
## 4 50 44.40
                     0.6599
##
```

```
## Elapsed time : 0.175719
model.armax.garch
##
## *----*
      GARCH Model Fit
## *----*
##
## Conditional Variance Dynamics
## -----
## GARCH Model : sGARCH(1,1)
## Mean Model : ARFIMA(2,0,3)
## Distribution : std
##
## Optimal Parameters
         Estimate Std. Error t value Pr(>|t|)
##
## mu
        ## ar1 0.023795 0.040488 0.587703 0.556732
      -0.894392 0.031402 -28.482399 0.000000
## ar2

      0.180533
      0.066946
      2.696704
      0.007003

      0.953060
      0.013923
      68.451358
      0.000000

## ma1
## ma2
## ma3
      0.258470 0.064191 4.026596 0.000057
## omega 0.000171 0.000126 1.353592 0.175867
## alpha1 0.092076 0.048015 1.917657 0.055155
                 0.065243 13.156584 0.000000
## beta1
         0.858378
## shape
         6.610146
                    3.242426
                            2.038642 0.041486
##
## Robust Standard Errors:
                            t value Pr(>|t|)
##
        Estimate Std. Error
## mu
        ## ar1
      0.023795 0.049430 0.481394 0.630236
## ar2
      -0.894392 0.034242 -26.120005 0.000000

      0.180533
      0.061615
      2.930006
      0.003390

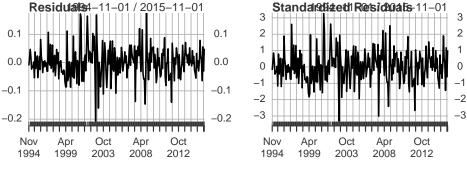
      0.953060
      0.016170
      58.939192
      0.000000

## ma1
## ma2
## ma3
         0.258470 0.063494 4.070783 0.000047
## omega 0.000171 0.000092 1.858395 0.063113
## alpha1 0.092076 0.040354 2.281704 0.022507
         ## beta1
                  2.907748 2.273287 0.023009
## shape
         6.610146
##
## LogLikelihood : 383.864
##
## Information Criteria
## Akaike
             -2.9749
## Bayes
             -2.8059
## Shibata
            -2.9792
```

##

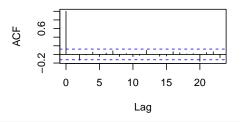
```
## Hannan-Quinn -2.9069
##
## Weighted Ljung-Box Test on Standardized Residuals
## -----
                      statistic p-value
## Lag[1]
                       0.004822 0.9446
## Lag[2*(p+q)+(p+q)-1][14] 4.789881 1.0000
## Lag[4*(p+q)+(p+q)-1][24] 9.750395 0.8540
## d.o.f=5
## HO : No serial correlation
## Weighted Ljung-Box Test on Standardized Squared Residuals
## -----
##
                     statistic p-value
## Lag[1]
                        0.3107 0.5772
## Lag[2*(p+q)+(p+q)-1][5] 1.5176 0.7356
## Lag[4*(p+q)+(p+q)-1][9] 3.1776 0.7293
## d.o.f=2
##
## Weighted ARCH LM Tests
## -----
     Statistic Shape Scale P-Value
## ARCH Lag[3] 0.6921 0.500 2.000 0.4054
## ARCH Lag[5] 2.3180 1.440 1.667 0.4051
## ARCH Lag[7] 3.3988 2.315 1.543 0.4416
## Nyblom stability test
## -----
## Joint Statistic: 1.955
## Individual Statistics:
## mu
       0.21143
## ar1
      0.33766
## ar2 0.21801
## ma1 0.09681
      0.14306
## ma2
## ma3
      0.09043
## mxreg1 0.20585
## mxreg2 0.27455
## omega 0.09041
## alpha1 0.08057
## beta1 0.09023
## shape 0.12346
## Asymptotic Critical Values (10% 5% 1%)
## Joint Statistic: 2.69 2.96 3.51
## Individual Statistic: 0.35 0.47 0.75
## Sign Bias Test
## -----
                  t-value prob sig
##
## Sign Bias
                 0.65210 0.5149
## Negative Sign Bias 0.01309 0.9896
## Positive Sign Bias 0.29202 0.7705
## Joint Effect 1.49230 0.6840
```

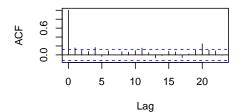
```
##
##
## Adjusted Pearson Goodness-of-Fit Test:
##
     group statistic p-value(g-1)
## 1
        20
               12.56
                            0.8603
## 2
        30
               25.28
                            0.6636
## 3
               35.12
                            0.6475
        40
## 4
        50
               44.40
                            0.6599
##
##
## Elapsed time : 0.175719
model.armax.garch@fit$coef
                                          ar2
                                                                        ma2
##
              mu
                            ar1
                                                         ma1
##
    0.0209166566
                  0.0237951944 -0.8943915639
                                               0.1805327148
                                                              0.9530604439
##
                                                                     alpha1
             ma3
                         mxreg1
                                       mxreg2
                                                       omega
##
    0.2584695098 -0.0002265352 -0.0003053705 0.0001710718
                                                              0.0920755280
##
           beta1
                          shape
   0.8583775746 6.6101460622
garch_time_index <- as.POSIXct(trainSet$Time)</pre>
residuals_armax_garch_xts <- xts(residuals(model.armax.garch),</pre>
                                  order.by = garch_time_index)
std_resid_armax_garch_xts <- xts(model.armax.garch@fit$z,</pre>
                                  order.by = garch_time_index)
# Residual Analysis
par(mfrow = c(2, 2))
# Residual plots
plot(residuals_armax_garch_xts, main = "Residuals")
plot(std_resid_armax_garch_xts, main = "Standardized Residuals")
# ACF plots
acf(na.omit(as.numeric(residuals(model.armax.garch))), main = "ACF of Residuals")
acf(na.omit(as.numeric(residuals(model.armax.garch)^2)), main = "ACF of Squared Residuals")
```



### **ACF of Residuals**

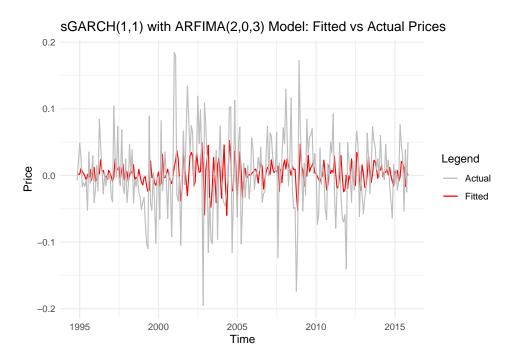
### **ACF of Squared Residuals**





```
# Extract the actual data (assuming you are using `trainSet` or the original series)
armax.garch.actual.values <- trainSet$diff_log_price</pre>
# Extract fitted values from the GARCH model (use `fitted` for the model residuals)
armax.garch.fitted.values <- fitted(model.armax.garch)</pre>
armax.garch.fit.df <- tibble(</pre>
 Time = trainSet$Time,
  Price = armax.garch.fitted.values
armax.garch.train.df <- tibble(</pre>
 Time = trainSet$Time,
  Price = trainSet$diff_log_price
ggplot() +
  geom_line(data = armax.garch.fit.df, aes(x = Time, y = Price, color = "Fitted")) +
 geom_line(data = armax.garch.train.df, aes(x = Time, y = Price, color = "Actual")) +
 labs(
    title = "sGARCH(1,1) with ARFIMA(2,0,3) Model: Fitted vs Actual Prices",
    y = "Price",
    x = "Time",
    color = "Legend"
  ) +
  theme minimal() +
  scale_color_manual(values = c("Actual" = "grey", "Fitted" = "red"))
```

## Don't know how to automatically pick scale for object of type <xts/zoo>. ## Defaulting to continuous.

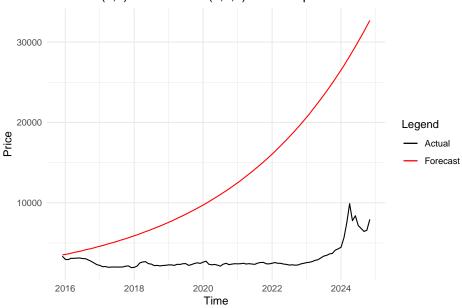


### 2.3.2 GARCH Forecast with param selection (With Xreg)

```
# multi-step forecast
ngarchfore = length(testSet$diff log price)
xreg_test_matrix <- cbind(testSet$Avg_Temp, diff(log(testSet$exchange_rate)))</pre>
## Warning in cbind(testSet$Avg_Temp, diff(log(testSet$exchange_rate))): number of
## rows of result is not a multiple of vector length (arg 2)
colnames(xreg_test_matrix) <- c("Avg_Temp", "dl_exchange_rate")</pre>
fore.garch.dl = ugarchforecast(model.armax.garch, n.ahead = ngarchfore,
                               external.forecasts = list(mreg=xreg test matrix))
fore.garch.dl.data <- fore.garch.dl@forecast$seriesFor</pre>
last_log_price <- tail(trainSet$log_price, 1)</pre>
forecast.armax.garch.multi <- exp(cumsum(fore.garch.dl.data) + last_log_price)</pre>
forecast.garch.multi.df <- tibble(</pre>
  Time = testSet$Time,
  Price = forecast.armax.garch.multi
test.garch.df <- tibble(</pre>
  Time = testSet$Time,
  Price = testSet$month_Price
)
ggplot() +
  geom_line(data = test.garch.df, aes(x = Time, y = Price, color = "Actual")) +
  geom_line(data = forecast.garch.multi.df, aes(x = Time, y = Price, color = "Forecast")) +
    title = "sGARCH(1,1) with ARFIMA(2,0,3) Multi-steps Forecast vs Actual Monthly Prices",
    v = "Price",
    x = "Time",
    color = "Legend"
```

```
) +
theme_minimal() +
scale_color_manual(values = c("Actual" = "black", "Forecast" = "red"))
```

sGARCH(1,1) with ARFIMA(2,0,3) Multi-steps Forecast vs Actual Monthl



RMSE below shows the forecast ability of the final model with one-step forecast

```
## [1] 383.9596
```

```
mae(trainSet$month_Price, model.armax.garch.fitted.converted)
```

```
## [1] 299.082
```

```
mape(trainSet$month_Price, model.armax.garch.fitted.converted)
```

```
## [1] 0.1682952
```

```
rmse(testSet$month_Price, forecast.armax.garch.multi[1:length(testSet$Time)])
```

### ## [1] 12425.02

```
mae(testSet$month_Price, forecast.armax.garch.multi[1:length(testSet$Time)])
```

### ## [1] 10119.37

```
mape(testSet$month_Price, forecast.armax.garch.multi[1:length(testSet$Time)])
```

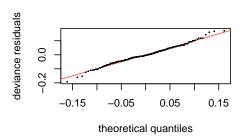
## [1] 3.424806

# 2.5 GAM Model

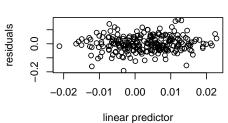
# 2.5.1 Fit Model

### 2.5.1.1 Basic Model

## Family: gaussian



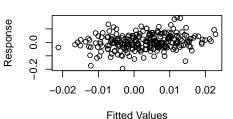
# Resids vs. linear pred.



### Histogram of residuals

# -0.2 -0.1 0.0 0.1 0.2 Residuals

# Response vs. Fitted Values



```
##
## Method: REML
                  Optimizer: outer newton
## full convergence after 10 iterations.
## Gradient range [-0.0001167052,0.0001905195]
## (score -350.1765 & scale 0.003138989).
## Hessian positive definite, eigenvalue range [3.540616e-06,123.5188].
## Model rank = 29 / 29
## Basis dimension (k) checking results. Low p-value (k-index<1) may
## indicate that k is too low, especially if edf is close to k'.
##
##
                            edf k-index p-value
                       k'
## s(month_num)
                    10.00
                           3.05
                                   1.05
                                           0.800
## s(Avg_Temp)
                     9.00
                           1.00
                                   1.06
                                           0.820
## s(exchange_rate)
                           1.00
                                   0.89
                                           0.045 *
                     9.00
                     '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
summary(gam_basic)
```

```
## Link function: identity
##
## Formula:
## diff_log_price ~ s(month_num, bs = "cc", k = 12) + s(Avg_Temp) +
##
       s(exchange_rate)
##
## Parametric coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 0.002929 0.003543 0.827
##
## Approximate significance of smooth terms:
                     edf Ref.df
##
                                    F p-value
                   3.054 10.000 0.795 0.0278 *
## s(month_num)
## s(Avg_Temp)
                   1.001 1.001 1.289 0.2571
## s(exchange_rate) 1.000 1.000 0.694 0.4055
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## R-sq.(adj) = 0.0261 Deviance explained = 4.59%
## -REML = -350.18 Scale est. = 0.003139 n = 250
trainSet$dateInt = as.integer(trainSet$Time)
trainSet$Time_num <- as.numeric(trainSet$Time) / 365 # Convert to years
gam_year <- gam(diff_log_price ~ s(month_num, bs="cc", k=12) +</pre>
                 sinpi(dateInt / 182.625) + cospi(dateInt / 182.625) +
                 sinpi(dateInt / 91.3125) + cospi(dateInt / 91.3125) +
                  s(Avg_Temp) + s(exchange_rate), data = trainSet, method = "REML")
summary(gam_basic)
```

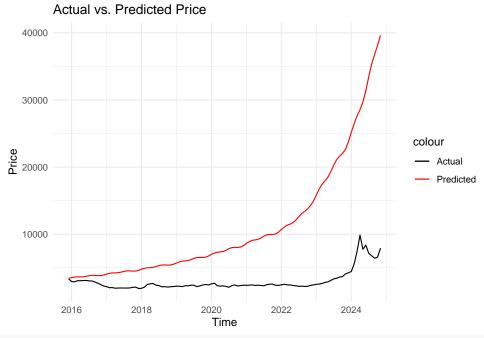
### 2.5.1.2 Complex Model

```
##
## Family: gaussian
## Link function: identity
##
## Formula:
## diff_log_price ~ s(month_num, bs = "cc", k = 12) + s(Avg_Temp) +
##
      s(exchange_rate)
##
## Parametric coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.002929 0.003543
                                   0.827
                                             0.409
## Approximate significance of smooth terms:
##
                     edf Ref.df
                                    F p-value
## s(month num)
                   3.054 10.000 0.795 0.0278 *
                   1.001 1.001 1.289 0.2571
## s(Avg_Temp)
## s(exchange_rate) 1.000 1.000 0.694 0.4055
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## R-sq.(adj) = 0.0261 Deviance explained = 4.59\%
```

```
## -REML = -350.18 Scale est. = 0.003139 n = 250
summary(gam_year)
##
## Family: gaussian
## Link function: identity
## Formula:
## diff_log_price ~ s(month_num, bs = "cc", k = 12) + sinpi(dateInt/182.625) +
##
      cospi(dateInt/182.625) + sinpi(dateInt/91.3125) + cospi(dateInt/91.3125) +
##
      s(Avg_Temp) + s(exchange_rate)
##
## Parametric coefficients:
##
                           Estimate Std. Error t value Pr(>|t|)
                          2.936e-03 3.542e-03 0.829
## (Intercept)
                                                        0.4081
## sinpi(dateInt/182.625) 8.070e-03 8.529e-03 0.946
                                                        0.3450
## cospi(dateInt/182.625) 9.264e-03 1.058e-02 0.876
                                                        0.3820
## sinpi(dateInt/91.3125) -7.343e-05 5.084e-03 -0.014
                                                        0.9885
## cospi(dateInt/91.3125) 1.449e-02 6.045e-03 2.397
                                                        0.0173 *
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Approximate significance of smooth terms:
                         edf Ref.df
                                       F p-value
                   0.0003453 10.000 0.000 0.707
## s(month_num)
                   1.0003821 1.001 0.289
                                           0.592
## s(Avg Temp)
## s(exchange rate) 1.0000069 1.000 0.820
                                           0.366
##
## R-sq.(adj) = 0.0271 Deviance explained = 5.06%
## -REML = -338.66 Scale est. = 0.0031356 n = 250
# plot(gam_complex, pages = 1, shade = TRUE)
# gam.check(gam_complex)
```

### 2.5.2 Forecast and Plot

### theme\_minimal()



rmse(testSet\$month\_Price, testSet\$pred\_price)

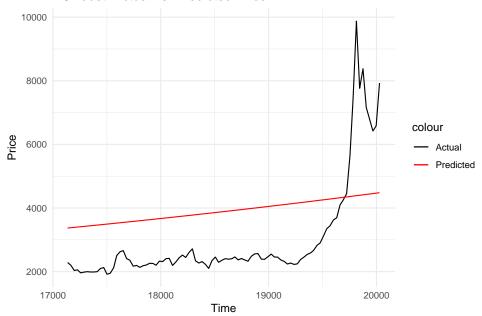
## [1] 10954.54

### 2.6 XGBoost Model

```
trainSet$Lag1 <- lag(trainSet$diff_log_price, 1)</pre>
trainSet$Lag12 <- lag(trainSet$diff_log_price, 12)</pre>
testSet$Lag1 <- lag(testSet$diff_log_price, 1)</pre>
testSet$Lag12 <- lag(testSet$diff_log_price, 12)</pre>
predictors <- c("Time", "Avg_Temp", "exchange_rate", "Lag1", "Lag12")</pre>
target <- "diff_log_price"</pre>
train_data <- trainSet[complete.cases(trainSet[, c(predictors, target)]), ]</pre>
test_data <- testSet[complete.cases(testSet[, predictors]), ]</pre>
train_data <- train_data |>
  mutate(across(all_of(predictors), as.numeric))
test_data <- test_data |>
  mutate(across(all_of(predictors), as.numeric))
# Convert to DMatrix (XGBoost's optimized format)
dtrain <- xgb.DMatrix(</pre>
  data = as.matrix(train_data[, predictors]),
  label = train_data[[target]]
params <- list(</pre>
  objective = "reg:squarederror",
                                      # For regression
  eta = 0.05,
                                      # Learning rate (lower for time series)
```

```
max_depth = 6,
                                   # Tree depth (avoid overfitting)
  subsample = 0.8,
                                   # Random subset of data per tree
                                   # Random subset of features per tree
  colsample_bytree = 0.8,
  gamma = 1,
                                   # Minimum loss reduction for splits
  min_child_weight = 5
                                   # Prevent overfitting to small groups
set.seed(123)
xgb_model <- xgb.train(</pre>
 params,
  data = dtrain,
 nrounds = 1000,
                                   # Large number (early stopping will handle)
 watchlist = list(train = dtrain),
                                   # Stop if no improvement for 50 rounds
 early_stopping_rounds = 50,
 print_every_n = 10
)
## [1] train-rmse:0.475762
## Will train until train_rmse hasn't improved in 50 rounds.
## [11] train-rmse:0.289145
## [21] train-rmse:0.179561
## [31] train-rmse:0.117399
## [41] train-rmse:0.084142
## [51] train-rmse:0.068493
## [61] train-rmse:0.061738
## [71] train-rmse:0.059145
## [81] train-rmse:0.058259
## [91] train-rmse:0.057915
## [101] train-rmse:0.057826
## [111]
         train-rmse:0.057774
## [121] train-rmse:0.057760
         train-rmse:0.057755
## [131]
         train-rmse:0.057752
## [141]
## [151] train-rmse:0.057749
## [161] train-rmse:0.057748
## [171] train-rmse:0.057748
         train-rmse:0.057749
## [181]
## [191] train-rmse:0.057748
## [201]
         train-rmse:0.057748
## [211]
           train-rmse:0.057748
## [221]
          train-rmse:0.057750
## [231]
         train-rmse:0.057749
## [241]
           train-rmse:0.057748
## [251]
           train-rmse:0.057749
## [261]
           train-rmse:0.057748
## Stopping. Best iteration:
           train-rmse:0.057748
## [214]
dtest <- xgb.DMatrix(as.matrix(test_data[, predictors]))</pre>
test_data$pred_diff_log <- predict(xgb_model, dtest)</pre>
# Convert to actual price predictions
test_data <- test_data %>%
```

### XGBoost: Actual vs. Predicted Price



```
# xgb = XGBClassifier(random_state=42)
# xgb.fit(predictors, target)
```