Machine Learning Handbook

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2018-2-28

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Part I High-level Views

Math Review

1.1 Linear Algebra

Concepts:

- scalar, vector, matrix, tensor(n-rank tensor, matrix is a rank 2 tensor)
- Gaussian Elimination, rank
- p-norm

$$|X|_p = (\sum_i |x_i|^p)^{\frac{1}{p}}$$

- inner product $< x_i, y_i >$, outer product
- orthogonal dimension, basis, orthogonal basis
- linear transformation Ax = y
- eigenvalue, eigenvector $Ax = \lambda x$ (transformation and speed)
- vector space, linear space(with summation, scalar production), inner product space(inner product space)

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1.2 Probability

Concepts:

- Classic Probability Model: Frequentist
- Bayesian Probability Theory
- Random variable, continuous RV, discrete RV, probability mass function, probability density function, cumulative density function
- expectation, moments, variance, covariance, correlation coefficient

Theorems:

- Law of Total Probability
- Bayes' Rule

$$P(H|D) = \frac{P(D|H)P(H)}{P(D)}$$

P(H)-prior probability, P(D|H)-likelihood, P(H|D)-posterior probability,

Important Distributions:

- 1. Bernoulli distribution
- 2. Binomial distribution(n,p)

$$P(X = k) = \binom{N}{k} p^k (1 - p)^{(n-k)}$$

3. Poisson distribution

$$P(X = k) = \lambda^k \frac{e^{-\lambda}}{k!}$$

- 4. Normal Distribution, See next chapter
- 5. Bernoulli Distribution

- 4
- 6. Uniform Distribution,
- 7. Exponential distribution

$$e^{-\frac{x}{\theta}}\theta$$

$$P(x > s + t | X > s) = P(x > t)$$

,

- 8. Poisson Distribution
- 9. normal distribution
- 10. t-distribution

Moment Generating Functions:

1.3 Information Theory

Concepts:

• Information

$$h(A) = -log_2 p(A)$$

(bit)

• (Information Source) Entropy

$$H(X) = -\sum_{i=1}^{n} p(a_i) log_2 p(a_i) \le log_2 n$$

Maximize under equal probability

• Conditional Entropy

$$H(Y|X) = -\sum_{i=1}^{n} p(x_i)H(Y|X = x_i) = -\sum_{i=1}^{n} p(x_i)\sum_{j=1}^{n} p(y_j|x_i)\log_2 p(y_j|x_i)$$
$$= \sum_{i=1}^{n} \sum_{j=1}^{n} p(x_i, y_j)\log_2 p(y_j|x_i)$$

• Mutual Information/Information Gain

$$I(X;Y) = H(Y) - H(Y|X)$$

• Kullback-Leibler Divergence (K-L) Divergence

$$D_{KL}(P||Q) = \sum_{i=1}^{n} p(x_i) log_2 \frac{p(x_i)}{q(x_i)} \neq D_{KL}(Q||P)$$

$$D_{KL}(f,\hat{f})) = \int_{-\infty}^{\infty} log(\frac{f_X(x)}{f(x)}) f_X(x) dx$$

K-L Divergence Measures the Distance of two distributions. The optimal encoding of information has the same bits as the entropy. Measures the extra bits if the real distribution is q rather than p. (Using P to approximate Q) K-L divergence plays an important role in both information theory and MLE theory. MLE $\hat{\theta}$ is actually finding the closest K-L Distance approximation of $f(x;\theta)$ to sample distribution.

Theorems:

• The Maximum Entropy Principle. Without extra assumption, max entropy/equal probability has the minimum prediction risk.

1.4 Optimization Theory

- Objective function/Evaluation function, constrained/unconstrained optimizationFeasible Set, Optimal Solution, Optimal Value, Binding Constraints, Shadow Price, Infeasible Price, Infeasibility, Unboundedness
- Linear Programming
- Lagrange Multiplier

$$L(x, y, \lambda) = f(x, y) + \lambda \varphi(x, y)$$

• Convex Set, Convex Function $f: S \to R$ is convex if and only if $\nabla^2 f(\mathbf{x})$ is positive semidefinite

Optimization Methods:

- Linear Search Method: Direction First, Step Size second
 - Gradient Descent: Batch Processing(Use all samples) vs
 Stochastic Gradient Descent(Use one sample)
 - Newton's Method: Use Curvature Information

$$\boldsymbol{\beta}^{t+1} = \boldsymbol{\beta}^t - (\frac{\partial^2 Loss(\boldsymbol{\beta})}{\partial \boldsymbol{\beta} \partial \boldsymbol{\beta}^T})^{-1} \frac{\partial Loss(\boldsymbol{\beta})}{\partial \boldsymbol{\beta}}$$

- Trust Region: Step first, direction second. Find optimal direction of second-order approximation. If the descent size is too small, make step size smaller.
- Heuristics Method
 - Genetic Algorithm
 - Simulated Annealing
 - Partical Swarming/Ant Colony Algorithm

Theorems:

1.5 Formal Logic

Concepts

- Generative Expert System: Rule+Facts+Deduction Engine
- Godel's incompleteness theorems

Statistics

2.1 Concepts

2.1.1 Basic

- parameter(constant for probability model), statistic (model of sample data), data, sample, population
- point estimation, interval estimation, Confidence Interval($P(L \le \theta \le U)$, notice: θ is not random, L, U is random! (We repeat constructing confidence interval a n times, α percent of the times, it will contain *theta*.

2.1.2 Estimator and Estimation

• Method of Moments: $E(X^k)$ based on LOLN. If We have p parameters, we can use p moments to form a system of equations to solve $\theta_1, ... \theta_p$

$$\sum_{i=1}^{n} X_i^j = E(X^j)$$

, for
$$j = 1,...,p$$

• Maximum Likelihood Estimation. Multiply p.m.f/p.d.f since every sample is independent. Maximize the likelihood of finding samples. If $X_1, ... X_n \stackrel{i.i.d}{\sim} f_x(x, \theta)$,

$$l(\theta) = \prod_{i=1}^{n} f_{X_i}(x_i; \theta), L(\theta) = logl(\theta)$$

$$\hat{\theta}_{MLE} = argmax_{\theta} f_x(x; \theta) = argmax_{\theta} L(\theta)$$

Analytical or Numerically solved.

$$\frac{\partial}{\partial \theta}[logL(\theta)] = 0, \frac{\partial^2}{\partial \theta^2}[logL(\theta)] < 0$$

, for multiple parameters, we need the Hessian matrix to be negative definite $x^t Hx < 0, \forall x$

- Properties of MLE
 - 1. Invariance $\hat{\theta}$ is MLE of θ , then $g(\hat{\theta})$ is MLE of $g(\theta)$
 - 2. Consistency

$$P(\hat{\theta} - \theta) \to 0$$

as $n \to 0, \forall \epsilon > 0$ Under the conditions

- (a) $X_1,...X_n \stackrel{i.i.d}{\sim} f_x(x|\theta)$
- (b) parameters are identifiable, $\theta \neq \theta'$, $f_x(x|\theta) \neq f_x(x|\theta')$
- (c) densities $f_x(x|\theta)$ has common support(set of x with positive density/probability), $f_x(x|\theta)$ is differentiable at θ
- (d) parameter space Ω contains open set ω where true θ_0 is an interior point
- 3. Asymptotic Normality

$$\sqrt{n}(\hat{\theta}_{MLE} - \theta_0) \to N(0, I^{-1}(\theta_0))$$

$$I(\theta_0) = E(-(\frac{\partial}{\partial \theta}[log f(x, \theta)])^2) = E(-\frac{\partial^2}{\partial \theta^2}[log f(x, \theta)])$$

called the Fisher Information

$$\hat{\theta}_{MLE} pprox N(\theta_0, \frac{1}{nI(\theta_0)})$$

$$nI(\theta_0) = E(-\frac{\partial^2}{\partial \theta^2}logL(\theta))$$

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– So the Variance of MLE($1/E(-\frac{\partial^2}{\partial \theta^2}logL(\theta))$) is the reciprocal of amount of curvature at MLE. Usually, We can just use the *observed Fisher Information* (curvature near θ_{MLE}) instead. $(I(\theta_{MLE}))$ $\frac{1}{nI(\theta_0)}$ is called Cramer-Rao Lower Bound. Under Multi-dimensional Case,

$$I(\theta_0)_{ij} == E(-\frac{\partial^2}{\partial \theta_i \partial \theta_i} [log f(x, \theta)])$$

 $Hessian \approx nI(\theta_0) \ Hessian^{-1} \approx nI(\theta_0)$ when we use numerical approach.

- Under the above four conditions plus
 - (a) $\forall x \in \chi$, $f_x(x|\theta)$ is three times differentiable with respect to θ , and third derivative is continuous at θ , and $\int f_x(x|\theta)dx$ can be differentiated three times under integral sign
 - (b) $\forall \theta \in \Omega, \exists c, M(x)$ (both depends on θ_0) such that

$$\frac{\partial^3}{\partial \theta^3}[log f(x,\theta)] \leq M(x), \forall x \in \chi, \theta_0 - c < \theta < \theta_0 + c, E_{\theta_0}[M(x)] < \infty$$

• Δ -Method: $g(t\hat{heta}_{MLE})$ is approximately

$$N(g(\theta), (g'(\theta))^2 \frac{1}{nI(\theta)})$$

if asymptotic normality is satisfied. In Multivariate Case:

$$\hat{\theta} \sim N(\theta, \Sigma/n), \theta, \hat{\theta} \in R^{p}$$

$$g: R^{p} \to R^{m}$$

$$g(\hat{\theta}) \sim N(g(\theta), G\Sigma G^{T}/n)$$

$$G = \begin{pmatrix} \frac{\partial g_{1}(\theta)}{\partial \theta_{1}} & \cdots & \frac{\partial g_{1}(\theta)}{\partial \theta_{p}} \\ \vdots & \ddots & \vdots \\ \frac{\partial g_{m}(\theta)}{\partial \theta_{1}} & \cdots & \frac{\partial g_{m}(\theta)}{\partial \theta_{p}} \end{pmatrix}$$

• Estimation criteria

- Unbiased $E(\hat{\theta}) = \theta$
- Minimum Variance (MVUE, minimum variance unbiased estimator) $Var(\hat{\theta}) < Var(\theta')$
- Efficient
- Coherent

2.1.3 Model Selection

AIC - Akaike Information Criterion

By K-L Distance

$$\begin{split} D_{KL}(f,\hat{f})) &= \int_{-\infty}^{\infty} log(\frac{f_X(x)}{f(x)}) f_X(x) dx \\ &= const + \frac{1}{2} \int (-2log\hat{f}(x)) f(x) dx = const + AIC \\ A(f,\hat{f}) &= -2logL(\theta) + 2p(\frac{n}{n-p+1}) \end{split}$$

2.1.4 Hypothesis Testing

- Hypothese, Test Statistic(T), Rejection Region
- p-value (chance of rejecting, largest choice of α that we would fail to reject H_0)
- type-I error(wrongly reject), type-II error(wrongly accept)

Hypothesis Testings (Based on the distribution of $\hat{\theta}$)

• Wald Test

$$\begin{split} T &= \frac{\hat{\theta} - \theta_0}{Se(\hat{\theta})} \\ \hat{\theta}_{MLE} &\approx N(\theta_0, \frac{1}{nI(\theta_0)}) \\ T &= \frac{\hat{\theta} - \theta_0}{\sqrt{\frac{1}{nI(\theta_0)}}} \end{split}$$

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- Likelihood Ratio Test
- Score Test

*Computation-based hypothesis testing approach

• Permutation tests:

Test $X_1,...X_n \sim F, Y_1,...Y_n \sim G, if F = G$. Use $T = Mean(X_i) - Mean(Y_i)$, each time scramble X and V labels and should not not change the distributions of vectors $X_1,...X_n, Y_1,..., Y_n$

• Bootstrapping:

 $X_1,...X_n \sim F$ with $T = T(X_1,...,X_n)$, to get the distribution of T, **sample with replacement.** The belief is $(\hat{\theta} - \theta)$ should behave the same as $(\theta * -t \hat{h} \hat{e} t a)$. The first quantity can be treated like a pivot. (use $(\theta *_1 - \hat{\theta}_1),...(\theta *_n - \hat{\theta}_n)$ to test.

Multiple Testing

- Family-wise Error Rate(FWER) the probability of rejecting at least one of at least one null hypothesis Under independence, the probability of making mistake when all null are true: P(any type I mistake) = 1-P(no type I mistake for all) = $1 (1 \alpha)^M = \beta$
- Bonferroni correction, assuming independence

$$P(\bigcup_{i=1}^{n} \text{typeI mistake}) \leq \sum_{i=1}^{n} P(\text{typeI mistake}) \leq M\alpha$$

,control at $\alpha = \frac{\alpha}{M}$ α being to small will impact power of the individual tests!

• False Discovery Rate(FDR): bound the fraction of type-I errors. R be the total number of hypotheses rejected. V be the number of rejected hypotheses that were actually null. Let FDR = V/max(R,1), control $E(FDR) \le \alpha$.

2.2 Theorems

- Law of Large Number
- Central Limit Theorem
- Bias/Variance decomposition (error = bias + variance + noise)

$$MSE(\mu(X)) = E[(Y - \hat{\mu}(X))^{2}] == E[(Y - f(x) + f(x) - \hat{\mu}(X))^{2}]$$

$$= E[(Y - f(x))^{2} + 2E[(Y - f(x))(f(x) - \hat{\mu}(X))] + E[(f(x) - \hat{\mu}(X))^{2}]$$

$$= E[(Y - f(x))^{2} + 2E[(Y - f(x))(f(x) - \hat{\mu}(X))] + (f(x) - \hat{\mu}(X))^{2}$$

$$= \sigma_{x}^{2} + Bias(\hat{\mu}(X))^{2} + Var(\hat{\mu}(X))$$

2.3 Important Distributions

- 1. Normal Distribution, $X_1,...X_n \sim N(\mu,\sigma^2)$ then
 - (a) \bar{X} and s^2 are independent
 - (b) $\frac{\bar{X}-\mu}{\sigma/\sqrt{n}} \sim N(0,1)$
 - (c) $\frac{(n-1)s^2}{\sigma^2} \sim \chi_{n-1}^2$

(d)
$$\frac{\bar{X}-\mu}{s/\sqrt{n}} = \frac{\frac{\bar{X}-\mu}{\sigma/\sqrt{n}}}{\frac{(n-1)s^2}{\sigma/2} \frac{1}{\sigma-1}} \sim t_{n-1}$$

2. Multi-variate normal distribution

$$f_x(x) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} exp(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu))$$

- (a) $X_1,...X_n$ normal $\Leftarrow (X_1,...X_n)$ is multivariate normal. (Not equivalent)
- (b) $E(X) = \mu$, $Var(X) = \Sigma$
- (c) Linear transformations $AX + b \sim N(A\mu + b, A\Sigma A^T)$ remain multivariate normal
- (d) Marginals are multivariate normal, each sub-vector is multivariate normal, the parameters are just sub-matrices.
- (e) All conditionals are multivariate normal

2.3. IMPORTANT DISTRIBUTIONS

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- 3. t-distribution: like normal distribution, but heavier tails
 - (a) $Z \sim N(0,1), Y \sim \chi^2_{\nu}, Z, Y$ independent,

$$X = Z/\sqrt{Y/\nu} \sim t_{\nu}$$

- (b) pdf has polynomial tails (decays much slower than exponential ones)
- (c) $\nu = 1$, it is the **Cauchy Distribution**, with very heavy tails (no expectation)
- (d) The MCF not exist. $E(|X|^k) < \infty$ for $k < \nu$, $E(|X|^k) = \infty$ for $k > \nu$
- (e) $X \sim t_{\nu}$, E(X) = 0, $Var(X) = \frac{\nu}{\nu 2}$

$$f_X(x) = \frac{1}{\pi(1+x^2)}$$

4. χ^2 distribution

$$f_x(x) = \frac{1}{(2^{k/2}\Gamma(k/2)} x^{\frac{k}{2}-1} e^{-\frac{x}{2}}, x \in [0, \infty) \sim Gamma(\frac{k}{2}, \frac{1}{2})$$

(a)
$$E(X) = k$$
, $Var(X) = 2k$, $M_X(t) = (\frac{1}{1-2^t})^{k/2}$

(b)
$$X \sim N(0,1) \Rightarrow X^2 \sim \chi^2, X_1, ... X_n \sim N(0,1) i.i.d \Rightarrow \sum X_i^2 \sim \chi^2,$$

$$f_X(x) = \frac{1}{\pi(1+x^2)}$$

5. F-Distribution

More Generalized Distributions

- 1. Generalized Error Distribution (symmetric)
- 2. Non-standard t-distribution (shift and scaling, heavy tailed, symmetric)
- 3. Theodossious skewed t-distribution
- 4. Theodossious skewed t-distribution plus shift

2.4 Practice/Examples

- 1. sample mean(\bar{X}) is unbiased. Sample variance $(\frac{1}{n-1}\sum_{i=1}^{n}x_{i}^{n})$ is unbiased. But sample std is not unbiased. $SE(\bar{X}) = \frac{\sigma^{2}}{n}$
- 2. $\hat{Cov}(X.Y) = \frac{1}{n-1} \sum_{i=1}^{n} (X_i \bar{X})(X_i \bar{Y})$ is unbiased
- 3. Distributions with Expectation not exist? (Cauchy)
- 4. Common Confidence Intervals:

$$μ: P(-t_{\alpha/2,n-1} \le \frac{\bar{X}-\mu}{s/\sqrt{n}} \le t_{\alpha/2,n-1}) = 1 - \alpha,$$
 $σ: P(a \le \frac{(n-1)s^2}{\sigma^2} \le b) = 1 - \alpha$

- 5. Solve MLE/MOM for beta, exponential $(n/\sum X_i$, normal
- 6. * prove Asymptotic Normality of MLE(hint: using Taylor Expansion for θ , $\hat{\theta}$)
- 7. * Use t^{th} quantile to approximate c.d.f, what's the distribution? $(Y_n = \frac{1}{n} \sum I(X_i < x))$, a Bernoulli distribution with $p = F_x(x)$, $\sqrt{n}[Y_n(x) F_x(x)] \sim N(0, F(x)(1 F(x)))$.
- 8. $X_1,...X_n \sim Binomial(n,p)$, What's the MLE for p and Fisher Information? ($\hat{p} = \frac{x_i}{n}$, I(p) = 1/p(1-p), $var(p) = \frac{p(1-p)}{n}$)
- 9. $(x_i, y_i) \sim N(\mu_i, \sigma^2)$, find MLE for σ ($\frac{1}{4N} \sum (x_i y_i)$)
- 10. How can you get N(0,1) random variables from U[0,1]? (Method1: Inverse Transformation, Method2; Use $SumZ_i^2 \sim \chi_k^2, k=2, F^{-1}(u)=-2log(1-u), \\ R^2 \sim \chi^2, Z_1=Rcos\theta, Z_2=Rsin\theta, \theta \in [0,2\pi]$
- 11. (Permutation test) how can you test $X_1, ..., X_n \sim F$, how can you test if F is symmetric? (Multiply -1 on all two form two sample groups)
- 12. Draw a bootstrap sample, what fraction of original data points appear in this sample on average? Define I be the indicator is it is in the sample. $E(\frac{1}{n}\sum I_i = E(I_i) = P(\text{ith point shows up}) = 1 (1 \frac{1}{n})^n$

Bayesian Statistical Theory

Bayes' Rule

$$f_{Y|X}(y|x) = \frac{f_{(X,Y)}(x,y)}{f_{X}(x)} = \frac{f_{(X,Y)}(x,y)}{\int f(x|y)f(y)dy}$$

Byesian Inference:

All parameters are random variables,

$$\pi(\theta|x) = \frac{f(x|\theta)\pi(\theta)}{\int f(x|\theta)\pi(\theta)d\theta}$$
$$\pi(\theta|x) \sim f(x|\theta)\pi(\theta)$$

 $\pi(\theta)$ is the prior distribution, $\pi(\theta|x)$ is the posterior distribution for θ given x.

Bayes Estimator

$$\hat{\theta}_{Bayes} = E(\theta|X) = \int \theta \pi(\theta|X) d\theta$$

Conjugate Distribution: f(x), π is called conjugate distributions if model $\pi(\theta|x)$, $\pi(\theta)$ follows the same Distribution

eg. Bernoulli (θ) and Beta (α, β) , ($\pi(\theta|x) \sim \text{Beta}(\alpha + \sum X_i, \beta + n - \sum X_i)$) $(f(x|\theta) = \prod_{i=1}^n f_{X_i}(X_i|\theta))$

$$\hat{\theta}_{Bayes} = E(\theta|X) = \frac{\alpha + \sum X_i}{\alpha + \beta + n}$$

$$= \frac{\sum X_i}{n} \frac{n}{\alpha + \beta + n} + \frac{\alpha}{\alpha + \beta} \frac{\alpha + \beta}{\alpha + \beta + n}$$

The prior mean (second term) influences less as n grows.

Poisson(θ) and $Gamma(\alpha + \sum X_i, \beta + n)$

3.1 Bayesian Decision Theory

In State ω we take action $a \in A$, incurr Loss $L(\omega, a)$, how to choose a? Risk:

$$R(a|x) = \sum_{j=1}^{k} L(\omega_j, a) P(\omega_j | x)$$

Decision Rule $d \in A$

$$d^*(x) = \operatorname*{arg\,min}_{a \in A} R(a|x)$$

Computational Learning Theory

Part II Supervised Learning Models

Basic Data Science and Data Wrangling Knowledge

Basic Transformations

Box-Cox power Transformation -useful when response is strictly postitive

$$y = \begin{cases} \frac{y^{\lambda} - 1}{\lambda}, & \text{if } \lambda \neq 0\\ \log(y), & \text{if } \lambda = 0 \end{cases}$$

 λ could be selected via MLE

• Yeo-Johnson Transformation

$$y = \begin{cases} \frac{(y+1)^{\lambda} - 1}{\lambda}, & \text{if } \lambda \neq 0, y \ge 0\\ \log(y+1), & \text{if } \lambda = 0 \text{ if } \lambda = 0, y \ge 0\\ \frac{(-y+1)^{2-\lambda} - 1}{\lambda}, & \text{if } \lambda \neq 2, y < 0\\ \log(-y+1), & \text{if } \lambda = 0 \text{ if } \lambda = 20, y < 0 \end{cases}$$

Optimal Transformation can be found with the MLE Method

Regression Overview and Linear Regression

6.1 Overview

6.1.1 Type of Models

All Basic Models begins with Linear Regression Because

- Linear relationship is the simplest relationship other than constant relationship or "null" model (average)
- It's a global model
- Data Invariance: Simple linear model don't do any pre-processing or transformation on the covariants.
- Very Explainable, limited interpretation power.

So, the alternation/improvements also focuses on these aspects

- Nonlinear features-Introduction of basis function
 - Polynomial Regression

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- Spline Models(eg. Cubic Spline, Smoothing Spline)
- Nonlinear parameters: Parameters Self-adjusting.(activation function is an example of basis function as well)
 - Neutral-Network
- global nonlinear: global nonlinear on both parameters and features achieved by linkage function, extends regression models to classification.
 - Generalized Linear Model
- Change the global model to a local model
 - Local Regression (Regression + KNN)
 - Nonparametric Regression
 - Kernel Function
 - Distance Based Learning
- Data Preprocessing (Transformation) and Dimension Reduction
 - PCA
 - LDA
 - Manifold Learning
- Improve Generalization Capability from outside (not from inside the model)
 - Regularization Methods(eg. Ridge, Lasso)
 - Ensemble Learning(Stacking, Aggregating): Random Forest, Boosting(GBDT), Deep Learning...

6.1.2 The Key Questions

- What assumptions are the model making
- How will we access the validity of those assumptions
- How can we be confident about out-of-sample fitting (overfitting problem)
- How do we make predictions and quantify the uncertainty in models?

6.2 Linear Regression

Common Terms

- 1. Independent Variable, Features, Covariates, Predictors
- 2. Dependent Variable, Response, Output (variable)
- 3. Scaling transform a variable to have mean zero and variance one

6.2.1 Assumptions

Classic Assumptions for Statistics:

- 1. Linear Relationship between covariates and dependent variable
- 2. $E(\varepsilon) = 0$
- 3. $Var(\varepsilon) = \sigma^2$: Homoscedasticity
- 4. ε is independent with covariates
- 5. x is observed without error (and no perfect multicollinearity in multivariate case)
- 6. (optional, Gauss-Markov Theorem) ε is normal when it is, OLS and MLE agrees and to be BLUE(Best Linear Unbiased Estimator)

Testing the Assumptions of Linear Regression

- Scatter Plot Linear Relationship and Outliers
- Residual Analysis $\hat{\varepsilon} = y \hat{y}$ Diagnostic Plots:
 - 1. Plot of Residuals vs. Fitted Values
 - 2. Normal Probability Plot
 - 3. Plot Residuals versus time (see any trend of fit)

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• Cook's Distance

$$D_j = \frac{\sum_{i=1}^n (\hat{y}_i - \hat{y}_{i(-j)})^2}{(p+1)\hat{\sigma}^2}$$

Test Against $F_{(p+1),(n-p-1)}$ degrees of freedom, over 50th percentile will definitely become a problem

• Detect Multicollinearity (two or more predictors are strongly related to one) - Use **Variance Inflation Factor**

$$VIF_k = \frac{1}{1 - R_k^2}$$

fit feature k against other predictors. Note VIF does not give any information of specific predictors

6.2.2 Resolutions of Assumption Violations

- Verify the Linear Relationships again. (non-linear regression, generalized linear models)
- Transformations (for outliers, heteroskesticities, etc)
- Use different models on different periods/data
- Weighted Least Squares regression, (for outliers, heteroskesticity)
- Robust Regression and Huber Loss Function

$$\sum_{i=1}^{n} \rho(\frac{y_i - x_i^T \beta}{\sigma})$$

Huber Loss Function

$$\rho(x) = \begin{cases} x^2, & \text{if } |x| < k \\ k(2(|x| - k), & \text{otherwise} \end{cases}$$

(default k=1.345) (when k=0, it is an L1-regression, $K \to \infty$, the regression goes back to a linear regression model. It is effective in down-weighting the extreme examples.

Special Situations

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Inputs are discrete variable - Factor Inputs (discrete features): a
factor of k levels adds k-1 terms into the regression function.(k-1
different betas)

6.2.3 Interpretation

Under Normal Condition, we have

$$y \sim N(\beta_0 + \beta x_i, \sigma^2), L(\theta) = (\frac{1}{\sqrt{2\pi}\sigma})^n exp(-\frac{\sum_{i=1}^n (y_i - (\beta_0 + \beta_1 x_i))^2}{2\sigma^2})$$

Equivalent to minimize

$$RSS(\theta) = \sum_{i=1}^{n} (y_i - (\beta_0 + \beta_1 x_i))^2$$
$$\partial_{\beta_i} RSS = 0, i = 1, 2$$

, we get

$$r_{xy} = \frac{s_{xy}}{s_x s_y}, \beta_1 = r_{xy} \frac{s_y}{s_x} = \frac{s_{xy}}{s_x^2}, \beta_0 = \bar{y} - \hat{\beta}\bar{x}$$

In Multi-variate Case:

$$f(x) = \mathbf{w}^T \mathbf{x} = \sum_{i=1}^n w_i x_i$$

$$\mathbf{w}^* = \underset{\hat{\mathbf{w}}}{\operatorname{arg\,min}} (\mathbf{y}) - \mathbf{X} \hat{\mathbf{w}})^T (\mathbf{y}) - \mathbf{X} \hat{\mathbf{w}})$$

$$\frac{\partial E}{\partial \hat{\mathbf{w}}} = 2\mathbf{X}^T (\mathbf{X} \hat{\mathbf{w}} - \mathbf{y})$$

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

Assuming noise is normal, maximize

$$p(\mathbf{x_1}, \mathbf{x_2}...\mathbf{x_n}|\mathbf{w}) = \prod_{k} \frac{1}{\sqrt{2\pi}\sigma} exp[-\frac{1}{2\sigma^2} (y_k - \mathbf{w_t} \mathbf{x_k})^2]$$

Another matrix representation

$$f(\beta) = \min(Y - X\beta)^{T}(Y - X\beta), f'(\beta) = 2X^{T}(Y - X\hat{\beta}) = 0$$

to solve $\hat{\beta}$

$$min||y_k - \mathbf{w}^{\mathsf{T}} \mathbf{x}_k||^2 + \lambda ||\mathbf{w}||_1$$

Variance Error In Prediction

$$\begin{split} V(\hat{y^*} - y^*) &= \sigma^2 + \sigma^2 [\frac{1}{n} + \frac{x^* - \bar{x})^2}{(n-1)s_x^2}] \\ &= V(E(y^*) - y^*) + V(\hat{y^*} - E(y^*)) + 2cov(\hat{y^*} - y^*, \hat{y^*} - y^*) \end{split}$$

The cross term is zero, the first term is variance with ε^* , second term is variance in β .

The confidence interval is $\hat{y^*} \pm t_{\alpha/2,n-2}SE(\hat{y^*})$.

 R^2 , the coefficient of determination: The proportion of the sum of squared response which is accounted by the model relative to the model with no covariance. (take mean of response)

$$R^2 = 1 - \frac{\sum (y_i - \hat{y}_i)^2}{\sum (y_i - \bar{y})}$$

Note that $0 \le R^2 \le 1$ It only tells predictive power if the model is a good fit.

Adjusted R^2 : R^2 + penalty P

Hat Matrix: The relationship of predicted value and response

$$Y = H\hat{Y}$$
$$H = X(X^TX)^{-1}X^T$$

The Diagonal Entires h_{ii} are the Leverages.

6.2.4 Model Selection

- Exhaustive Search by AIC or BIC (more stable than LOOCV)
- Stepwise Regression/Stepwise Variable Selection (At each step one covariate is added or dropped)
- Cross-Validation
 Leave-one-Out cross Validation of Linear Regression: Prediction
 Error Sum of Squares

$$PRESS = \frac{\sum (y_i - \hat{y}_{-i})^2}{n}$$
$$y_i - \hat{y}_{-i} = \frac{\hat{\varepsilon}_i}{1 - h_{ii}}$$

h is the leverage (hat matrix)

6.2.5 Regularization, Ridge, Lasso

Ridge

$$Rss + \lambda \sum_{i=1}^{p} \beta^2$$

 λ is the regularization parameter. The result of Ridge is a **Shrinkage** of $\hat{\beta}$ towards zero.

Note

- 1. No penalty for $\beta + 0$ or b.
- 2. The predictors should usually be standardized prior to fitting
- 3. Choose λ by cross-validation

Lasso(Least Absolute Shrinkage and Selection Operator)

(Tibshirani)

$$Rss + \lambda \sum_{i=1}^{p} |\beta|$$

Can be extended to

$$-loglikelihood + \lambda \sum_{i=1}^{p} |\beta|$$

Group Lasso

group predictors together to be either included or excluded.

Elastic Net

$$Rss + \lambda \sum_{i=1}^{p} (\alpha |\beta_j| + (1 - \alpha)\beta_j^2)$$

$$0 \le \alpha \le 1$$

6.3 Nonlinear Regression Models

- Nonparametric Regression: Complexity controlled by the smoothing parameter (bandwidth). model complexity interpreted in Degrees of Freedom/Effective degrees of freedom/equivalent degrees of freedom
 - Residual Degrees of freedom is n minus model degrees of freedom.
- Local polynomial Regression: only fit a **neighborhood** of a target point. parameter α to control the span-traditionally, 0.5. When weighting the data in the neighborhood, Fit by weighted sum of squares

$$\sum_{i=1}^{n} w_i (y_i - (\beta_0 + \beta_1 x_i))^2$$

Splines

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 Penalized (Smoothing) Splines: find twice differentiable x to minimize

$$\sum_{i=1}^{n} (y_i - f(x_i))^2 + \lambda \int [f^{(2)}(x)]^2 dx$$

 λ penalty for wiggy function. search of x can be a combination of **basis functions** (n + 4 basis functions, n is the knots)

- Cubic Splines

6.4 Generalized Additive Models

Additive Model: no interactions/cross terms

6.5 Practice/Examples

1. What is Anscombe's Quartet

Logistic Regression and Generalized Linear Model

7.1 Logistic Regression

7.1.1 Logistic Regression

Sigmoid/ Log Probability Function

$$\sigma(z) = \frac{1}{1 + e^{-z}} = \frac{1}{1 + e^{-w^T x}}$$

Loss Function

$$J(z) = -ylogy + (1-y)log(1-y)$$

Training of the Model

MLE of w based on a Bernoulli Distribution

$$L(\mathbf{w}|\mathbf{x}) = \prod_{i=1}^{N} [p(y=1|\mathbf{x}, \mathbf{w})]^{y_i} [1 - p(y=1|\mathbf{x}, \mathbf{w})]^{1-y_i}$$

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Take Log to get the Loss Function

$$logL(\mathbf{w}|\mathbf{x}) = \sum_{i=1}^{N} -y_i log y_i + (1 - y_i) log (1 - y_i)$$
$$l(\mathbf{w}|\mathbf{x}) = logL(\mathbf{w}|\mathbf{x}) = \sum_{i=1}^{N} (y_i \mathbf{w}^T \mathbf{x}_i - log(1 + e^{\mathbf{w}^T \mathbf{x}_i}))$$

Intuition

• The log odds (logit function)

$$ln\frac{y}{1-y} = \mathbf{w}^T \mathbf{x} + b$$
$$p(y=1|x) = \frac{e^{\mathbf{w}^T \mathbf{x}}}{1 + e^{\mathbf{w}^T \mathbf{x}}}$$
$$p(y=0|x) = \frac{1}{1 + e^{\mathbf{w}^T \mathbf{x}}}$$

- Exponential Family
- Maximum Entropy (See Loss function) of the exponential family

$$2\sum_{i|1}^{N} -y_{i}log\frac{y_{i}}{\hat{p}_{i}} + (1-y_{i})log(\frac{1-y_{i}}{1-\hat{p}_{i}})$$

Another Intuition about minimizing the loss function is to minimize the **K-L Divergence** with Maximum-Entropy Model

*Connection with Naive Bayesian

• Naive Bayesian assumes $p(x_i|Y=y_k)$ follows a normal distribution. Then the posterior probability is

$$P(Y = 0|x) = \frac{P(Y = 0)P(X|Y = 0)}{P(Y = 0)P(X|Y = 0) + P(Y = 0)P(X|Y = 1)}$$

$$= \frac{1}{1 + exp(ln\frac{P(Y = 0)P(X|Y = 1)}{P(Y = 0)P(X|Y = 0)})}$$

$$= \frac{1}{exp(ln\frac{1 - p_0}{p_0} + \sum(\frac{\mu_{i1} - \mu_{i0}}{\sigma_i^2}X_i + \frac{\mu_{i0}^2 - \mu_{i1}^2}{2\sigma_i^2}))}$$

7.1. LOGISTIC REGRESSION

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- Though the solution follows the exact same pattern, Logistic Regression does not have the assumption of independence. When assumptions differ, the results differ. Generally, logistic regression results less bias, more variance(more flexible)
- The rate of convergence is also different, logistic regression needs more data feeding to perform better.

7.1.2 Extension: Softmax

$$P(Y = k|x) = \frac{e^{w^T x}}{\sum_{i=1} K e^{w^T x}}$$

Neutral-Network

Distance and Kth Nearest Neighbors

Naive Bayesian

Tree Models and Ensemble Learning

Classification Overview and Support Vector Machine

Stacking Aggregating: Random Forest, Boosting(GBDT), Deep Learning...

Part III

Unsupervised Learning Models

Clustering

Dimension Reduction

14.1 Principal Component Analysis(PCA)

$$\mathbf{x}_i = (x_{i1}, x_{i2}, x_{ip})^T$$

find u to maximize variance $v_i = \sum_{j=1}^p u_j x_{ij}$

subject to

$$\sum_{j=1}^p u_j^2 = 1$$

u are like the new axes

$$v_i = \sum_{j=1}^p u_j x_{ij}$$

same as finding the eigenvalue Sigma

X is samples stacking in columns,

$$\mathbf{X} = \mathbf{X} - \mathbf{1}\boldsymbol{\mu}^T$$

be X shifted to mean zero

Then $\mathbf{v} = \mathbf{X}\mathbf{u}$ is a transformation (new position in axis)

$$\mathbf{v}^T 1 = (\mathbf{X}\mathbf{u})^T \mathbf{1} = \mathbf{u}^T \mathbf{X}^T 1 = \mathbf{u}^T \mathbf{0} = 0$$

the mean of v entries is zero

The variance of entries of v:

$$\frac{1}{n}\mathbf{v}^{\mathsf{t}}\mathbf{v} = \frac{1}{n}\mathbf{u}^{\mathsf{T}}\mathbf{X}^{\mathsf{T}}\mathbf{X}\mathbf{u}$$

equivalent to maximizing

$$\frac{\mathbf{u}^T \mathbf{X}^T \mathbf{X} \mathbf{u}}{\mathbf{u}^T \mathbf{u}}$$

$$\mathbf{X}^T\mathbf{X}$$

is the covariance matrix, this is equivalent to find the first eigenvector of covariance matrix if we scale X to have unit standard deviation, the problem is finding eigenvalues for correlation matrix

14.2 LDA

Part IV

Deep Learning and Enhanced Learning Theory

Deep Learning, Neural Network and Multi-layer Perceptron

15.1 Neural Networks

15.1.1 History Class

• MP Neuron

$$y = \phi(\sum_{i=1}^{N} w_i x_i)$$

$$w_i(t+1) = w_i(t) + \eta[d_j - y_j(t)]x_{j,i}$$

can only solve linear-separable problems

Multilayer Perceptron: Including one hidden layer. The first
Feedforward Neural Network. Errors passes by Back Propagation.
Its it proven that a single-hidden layer multilayer perceptron can
approximate any continuous functions at arbitrary error
level.(universal approximation)

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15.1.2 Neural Networks

Radial Basis Function Network

In the hidden layer, use activation function as activation function

Radial Basis Function

$$\rho(\mathbf{x}, \mathbf{w_i}, \sigma) = exp(\frac{-\|\mathbf{x} - \mathbf{w_i}\|^2}{2\sigma^2})$$

As long as a feature's distance to the center vector (here $\mathbf{w_i}$) the same, the function value is the same. $\mathbf{w_i}$ separates different hidden unit band with bandwidth σ

The Gaussian function $exp(-\|\mathbf{x} - \mathbf{u_i}\|^2)$ (like Kernel) can help to transform linear inseparable case as-if projecting to a high-dimension space (same as SVM), to a linear separable case.

Alternatively, treat an RBF as a interpolation solution. It tries to data hyperplane. It reduces the noise by interpolation among the data points. The interpolated hyperplane still passes all data points.

Training of RBF

1. Initialization of $\mathbf{w_i}$ by random initialization or **unsupervised learning** like K-means.

Usually, we have $\sigma = d_m ax / \sqrt{2K}$, $d_m ax$ is the maximum distance between centers. (make sure bandwidth is not too small or too big)

2. Training w_i. Use Recursive Least Square

$$\mathbf{R}(n)\mathbf{\hat{w}}(n) = \mathbf{r}(n)$$

 $\mathbf{R}(n)$ is the covariance matrix between hidden layer outputs (\hat{y}) , $\mathbf{r}(n)$ is the covariance vector between hidden layer outputs (\hat{y}) and

r(n) is the covariance vector between fluden layer outputs (y) and model response.

Training by solving $\mathbf{R}^{-1}(n)$

3. After training, use Back propagation to train all parameters one more time. (train the whole network after training layers)

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Compare with Neural Network: both can achieve universal approximation, while RBF network uses a local approximation approach.

Deep Learning sees most results in supervised Learning.

Feedforward neural network

Part V

Bayesian Learning and Probabilistic Graphical Model

Markov Chain Hidden Markov Chain Markov Blanket Bayesian Network EM Algorithm