MMF1928 PRICING THEORY LECTURE NOTES

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December 1, 2024

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1.1 Setup

Financial Market ⇒ Time Series ⇒ Information Technology

Space: $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \in \mathcal{T}}, \mathbb{P})$

- 1. Probability Triple:
 - (1) Ω : Sample space, all possible outcomes
 - (2) \mathcal{F} : σ -algebra, all information, collection of event sets
 - (3) \mathbb{P} : Probability measure
- 2. $\{\mathcal{F}_t\}$: A collection of event sets, indexed by t.
 - (1) \mathcal{F}_t : σ -algebra, information being available at time $t, \forall s < t, s, t, \in \mathcal{T}, \mathcal{F}_s \subset \mathcal{F}_t \subset \mathcal{F}$.
 - (2) \mathcal{T} = Time Set
 - $\mathcal{T} = \{0, T\}$: Single-period model
 - $\mathcal{T} = \{0, 1, 2, \dots, T\}$: Multi-period model
 - $\mathcal{T} = \{0, 1, 2, \dots\}$: Discrete-time model
 - $\mathcal{T} = [0, T]$: Continuous-time model
- 3. $\{\mathcal{F}\}_{t\in\mathcal{T}}$: filtration, i.e., increasing collection of σ -algebra

1.2 Financial Assets

- 1. Cash: Corporate/Government bonds, interest rates (r). Usually no risk, deterministic, money market (numéraire).
- 2. Stock: Usually very risky, stochastic, S_t , adaptive to \mathcal{F}_t .
- 3. Option: Maturity T, strike price K. European call at $T:(S_T-K)^+$. American call before $T:(S_t-K)^+$
 - (1) X_t : Wealth of portfolio (to replicate the payoff of option).
 - (2) Δ_t : Number of shares in risky assets. $\Delta_t > 0$: Long position. $\Delta_t < 0$: Short position (Assumption in this course: continuous, short selling allowed.)

1.3 Single Period Model

Question: $X_0 = V_0 \Longrightarrow \forall \omega \in \Omega, X_T(\omega) = V_T(\omega)$

Key feature: no extra input/output, self-financing strategy (Δ_t)

Definition 1.3.1. A self-financing strategy $\{\Delta_t\}_{t\in\{0,T\}}$ is an adaptive stochastic process such that

- i. $\{\Delta_t\}_{t\in\{0,T\}}$ is adaptive, Δ_t is \mathcal{F}_t -measurable
- ii. $X_0 = \Delta_0 S_0 + (X_0 \Delta_0 S_0)$

Note: As time goes on, it becomes $X_1(\omega) = \Delta_0 S_1(\omega) + (1+r)(X_0 - \Delta_0 S_0) = \Delta_1(\omega) S_1(\omega) + X_1(\omega) - \Delta_1(\omega) S_1(\omega)$

Example 1.3.2. Assume r = 0.

At time t = 0, $\Delta_0 = 10$ and $S_0 = \$100$, we have $X_0 = 10 \times \$100 + \$1000 - 10 \times \$100 = \1000 . At this time point, money in stock is $10 \times \$100 = \1000 and money in cash is $\$1000 - 10 \times \$100 = \$0$.

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As time goes to t=1, $\Delta_1=9$ and $S_1=\$105$, we have $X_1=10\times\$105+\$1000-10\times\$100=\1050 and $X_1=9\times\$105+\$1050-9\times\$105=\1050 . At this time point, money in stock is $9\times\$105=\945 and money in cash is $\$1050-9\times\$105=\$105$.

Definition 1.3.3. An arbitrage strategy $\{\Delta_t\}_{t\in\{0,T\}}$ is a self-financing strategy such that

- i. $X_0 = 0$ portfolio at time 0
- ii. $\mathbb{P}[(X_T \ge 0] = 1 \text{ and } \mathbb{P}[(X_T > 0] > 0, X_T \text{ portfolio at } T$

1.4 Binomial Model

Space: $(u, d, r, p = \mathbb{P}[\text{outcome} = H])$

- 1. Stock: The initial stock price is S_0 . If the stock price goes up, then $S_1(H) = uS_0 \Rightarrow u = \frac{S_1(H)}{S_0}$; if the stock price goes down, then $S_2(H) = dS_0 \Rightarrow d = \frac{S_2(T)}{S_0}$, where $u \neq d$.
- 2. Cash: Compounding at the interest rate of r.
- 3. Arbitrage: Suppose $d < u \le 1 + r$, then we have

$$X_0 = 0 = -\Delta_0 S_0 + \Delta_0 S_0$$

$$X_1(H) = -\Delta_0 S(H) + (1+r)\Delta_0 S_0 = (1+r-u)\Delta_0 S_0$$
, where $1+r-u \ge 0$

$$X_1(T) = -\Delta_0 S(T) + (1+r)\Delta_0 S_0 = (1+r-d)\Delta_0 S_0$$
, where $1+r-d>0$

$$\mathbb{P}(X_1 \ge 0) = 1, \mathbb{P}(X_1 > 0) = \mathbb{P}(\text{outcome} = T) > 0$$

To prevent arbitrage, we need p > 0, u > 1 + r > d.

For example, $S_0=4,\,u=2,\,d=\frac{1}{2},\,r=\frac{1}{4},\,\mathbb{P}(\text{outcome}=H)=\frac{1}{2}.$

1.5 Risk-Neutral Pricing

Under what measure is the discount stock price being a martingale?

Idea: every option is compounding with respect to the interest rate r.

Consider the conditional expectation, we have

$$\tilde{\mathbb{E}}[S_1 \mid S_0] = (1+r)S_0$$

$$\Rightarrow \quad \tilde{\mathbb{E}}[S_1 \mid S_0] = \tilde{p}uS_0 + \tilde{q}dS_0$$

$$\Rightarrow \quad \begin{cases} \tilde{p} + \tilde{q} = 1 \\ u\tilde{p} + d\tilde{q} = 1 + r \end{cases} \Rightarrow \begin{cases} \tilde{p} = \frac{1+r-d}{u-d} \\ \tilde{q} = \frac{u-(1+r)}{u-d} \end{cases}$$

Since d < 1 + r < u, then we know $\tilde{p}, \tilde{q} \in (0, 1)$.

Recall a self-financing strategy $\{\Delta_t\}_{t\in\{0,1\}}$, we know

$$X_0 = x = \Delta_0 S_0 + x - \Delta_0 S_0 \Rightarrow X_1 = \Delta_0 S_1 + (1+r)(x - \Delta_0 S_0)$$

then we can derive that, in Q-world,

$$\tilde{\mathbb{E}}[X_1 \mid S_0, X_0 = x] = \Delta_0 \tilde{\mathbb{E}}[S_1 \mid S_0] + (1+r)(x - \Delta_0 S_0) = \Delta_0 S_0 + (1+r)(x - \Delta_0 S_0) = (1+r)x$$

Example 1.5.1. Assume $S_0 = 4$, u = 2, $d = \frac{1}{2}$, $r = \frac{1}{4}$, $V_1 = (S_1 - K)^+$, K = 4.

We calculate
$$\tilde{p} = \frac{1+r-d}{u-d} = \frac{1+\frac{1}{4}-\frac{1}{2}}{2-\frac{1}{2}} = \frac{1}{2}$$
 and $\tilde{q} = 1 - \tilde{p} = \frac{1}{2}$.

To replicate the portfolio, for some unknown Δ_0, X_0 , we have

$$\begin{cases} X_1(H) = V_1(H) = \Delta_0 S_1(H) + (1+r)(X_0 - \Delta_0 S_0) = 4 \\ X_1(T) = V_1 T H) = \Delta_0 S_1(T) + (1+r)(X_0 - \Delta_0 S_0) = 0 \end{cases} \Rightarrow \Delta_0 = \frac{V_1(H) - V_1(T)}{S_1(H) - S_1(T)} = \frac{4}{8-2} = \frac{2}{3}$$

Therefore, the delta-hedging formula is $\Delta_0 = \frac{V_1(H) - V_1(T)}{S_1(H) - S_1(T)}$

1.6 Fundamental Theorem of Asset Pricing I (FTAP I)

Theorem 1.6.1. (Fundamental Theorem of Asset Pricing I (FTAP I)). No arbitrage in a financial market \Leftrightarrow There exists a risk-neutral measure $\mathbb{Q} \sim \mathbb{P}$ such that the discounted stock price is a martingale.

Remark 1.6.2. $\mathbb{Q} \sim \mathbb{P}$ means that $\forall A \in \mathcal{F}, \mathbb{Q}[A] = 0 \Leftrightarrow \mathbb{P}[A] = 0$

Theorem 1.6.3. (First fundamental theorem of asset pricing from Steve Shreve, Stochastic Calculus for Finance Volume II Theorem 5.4.7). If a market model has a risk-neutral probability measure, then it does not admit arbitrage.

Example 1.6.4. Assume u=2, $d=\frac{1}{2}$, r=0, $S_0=\$4$. If the stock price goes up, $S_u=\$8$. If the stock price remains the same, $S_n=\$4$. If the stock price goes down, $S_d=\$2$. Let p denote the probability of stock price going up and q denote the probability of stock pricing going down. Assume the probability of going up and going down are $\tilde{p}=\frac{1}{4}$ and $\tilde{q}=\frac{1}{4}$. Then we have $\Pr[S_1=\$4]=1-\frac{1}{4}-\frac{1}{4}=\frac{1}{2}$.

For $A=\varnothing$, we have $\mathbb{P}[A]=\mathbb{Q}[A]=0$. For $A=\{\varnothing,\{S_1=\$4\}\},\,\hat{p}=\hat{q}=\frac{1}{2}$, however, we know that $\tilde{p}=\frac{1}{4}$ and $\tilde{q}=\frac{1}{4}$, which is not equivalent. We also calculate

$$\widetilde{\mathbb{E}}[S_1 \mid S_0] = \frac{1}{4} \times \$8 + \frac{1}{4} \times \$2 + \frac{1}{2} \times \$4 = \$4.5 \neq \$4$$

Therefore, it is not a risk-neutral measure.

If
$$\tilde{p}^{\mathbb{Q}} = \frac{1}{6}$$
, $\tilde{q}^{\mathbb{Q}} = \frac{1}{3}$, $\Pr^{\mathbb{Q}}[S_1 = \$4] = \frac{1}{2}$, then

$$\tilde{\mathbb{E}}^{\mathbb{Q}}[S_1 \mid S_0] = \frac{1}{6} \times \$8 + \frac{1}{3} \times \$2 + \frac{1}{2} \times \$4 = \$4 = \$4$$

Therefore, it is a risk-neutral measure.

If
$$\tilde{p}^{\mathbb{Q}} = \frac{1}{5}$$
, $\tilde{q}^{\mathbb{Q}} = \frac{2}{5}$, $\Pr^{\mathbb{Q}}[S_1 = \$4] = \frac{2}{5}$, then

$$\tilde{\mathbb{E}}^{\mathbb{Q}}[S_1 \mid S_0] = \frac{1}{5} \times \$8 + \frac{2}{5} \times \$2 + \frac{2}{5} \times \$4 = \$4 = \$4$$

Therefore, it is a risk-neutral measure.

1.7 Fundamental Theorem of Asset Pricing II (FTAP II)

Theorem 1.7.1. (Fundamental Theorem of Asset Pricing II (FTAP II)). There exists a unique risk-neutral measure \Leftrightarrow The financial market is complete, i.e., for every option there exists a replicating portfolio.

Theorem 1.7.2. (Second fundamental theorem of asset pricing from Steve Shreve, Stochastic Calculus for Finance Volume II Theorem 5.4.9). Consider a market model that has a risk-neutral probability measure. The model is complete if and only if the risk-neutral probability measure is unique.

2.1 Review of FTAP I and FTAP II

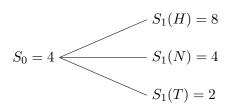
Theorem 2.1.1. (FTAP I). There is no arbitrage in the financial market if and only if there exists a probability measure $\mathbb{Q} \sim \mathbb{P}$ such that under \mathbb{Q} , the discount asset prices are martingales, i.e., $\mathbb{Q} \sim \mathbb{P}$ is a risk-neutral measure.

Theorem 2.1.2. (FTAP II). There exists a unique risk-neutral measure $\mathbb{Q} \sim \mathbb{P}$ if and only if the market is complete, i.e., all options have a replicating portfolio.

Remark 2.1.3. "~" denotes the equivalence of probability measures.

Remark 2.1.4. The time of a martingale is "today".

Example 2.1.3. (Application of FTAP I and FTAP II). On a probability space of $(\Omega, \mathcal{F}, \mathbb{P})$, we have $\mathcal{F}_0 = \{\varnothing, \Omega\}$ and $\mathcal{F}_1 = \mathcal{F}$. For $\Omega = \{H, N, T\}$, $\omega \in \Omega$, and $S_1(\omega)$, we have $N_{\Omega} = 3$, $\mathcal{F} = \{\varnothing, \Omega, \{H\}, \{N\}, \{T\}, \{H, N\}, \{H, T\}, \{N, T\}\}$, and $N_{\mathcal{F}} = 2^{N_{\Omega}} = 2^3 = 8$.



For a physical measure \mathbb{P} , $\mathbb{P}[\Omega] = 1$, $\mathbb{P}[H] = \mathbb{P}[N] = \mathbb{P}[T] = \frac{1}{3}$, $\mathbb{P}[\{H, N\}] = \mathbb{P}[S_1 \neq 2] = \frac{2}{3} = \mathbb{P}[\{H, T\}] = \mathbb{P}[\{N, T\}]$, and $\mathbb{P}[\varnothing] = 0$. Are the following portfolios risk-neutral? Assume that r = 0.

(1) For a new probability measure $\tilde{\mathbb{P}}$, $\tilde{\mathbb{P}}[H] = \tilde{\mathbb{P}}[T] = \frac{1}{4}$ and $\tilde{\mathbb{P}}[N] = \frac{1}{2}$:

For any $A\in\mathcal{F}$, we can easily get $\tilde{\mathbb{P}}[A]=0\Leftrightarrow A=\varnothing$, so $\tilde{\mathbb{P}}\sim\mathbb{P}$.

Then we calculate the discounted asset price as $\tilde{\mathbb{E}}[S_1\mid S_0]=\frac{1}{4}\cdot 8+\frac{1}{4}\cdot 2+\frac{1}{2}\cdot 4=4.5\neq S_0.$

Therefore, $\tilde{\mathbb{P}}$ is not a risk-neutral measure.

(2) For a new probability measure $\hat{\mathbb{P}}$, $\hat{\mathbb{P}}[H] = \hat{\mathbb{P}}[T] = \frac{1}{2}$ and $\hat{\mathbb{P}}[N] = 0$:

For $\{N\}=\{S_1=4\}\in\mathcal{F}$, we have $\hat{\mathbb{P}}[N]=0$, however, $\mathbb{P}=\frac{1}{3}$, so $\hat{\mathbb{P}}$ is not equivalent to \mathbb{P} .

(3) For a new probability measure \mathbb{Q} , $\mathbb{Q}[H] = \frac{1}{6}$, $\mathbb{Q}[T] = \frac{1}{3}$, and $\mathbb{Q}[N] = \frac{1}{2}$:

For any $A \in \mathcal{F}$, we can easily get $\mathbb{Q}[A] = 0 \Leftrightarrow A = \emptyset$, so $\mathbb{Q} \sim \mathbb{P}$.

Then we calculate the discounted asset price as $\tilde{\mathbb{E}}[S_1 \mid S_0] = \frac{1}{6} \cdot 8 + \frac{1}{3} \cdot 2 + \frac{1}{2} \cdot 4 = 4 = S_0$.

Therefore, \mathbb{Q} is a risk-neutral measure.

Note: The standard mechanism is

- (1) Check $\mathbb{O} \sim \mathbb{P}$
- (2) Check if discounted asset prices are martingale

 \mathbb{Q} is not unique since $\hat{\mathbb{Q}}[H] = \frac{2}{5}$, $\hat{\mathbb{Q}}[T] = \frac{1}{5}$, and $\hat{\mathbb{Q}}[N] = \frac{2}{5}$ is also a risk-neutral measure. Then by FTAP II, there exists unhedgable options in the market.

To find an unhedgable option in the market, consider the following option

$$V_1(H) = (S_1(H) - 3)^+ = 5 = X_0 - \Delta_0 S_0 + S_1(H) \Delta_0 = X_0 + 4\Delta_0$$

$$V_1(N) = (S_1(N) - 3)^+ = 1 = X_0 - \Delta_0 S_0 + S_1(N) \Delta_0 = X_0$$

$$V_1(T) = (S_1(T) - 3)^+ = 0 = X_0 - \Delta_0 S_0 + S_1(T) \Delta_0 = X_0 - 2\Delta_0$$

There is no solution to these equations, therefore, $V_1 = (S_1 - 3)^+$ is unhedgable.

2.2 Martingale

Definition 2.2.1. On space $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \in \mathcal{T}}, \mathbb{P})$, a stochastic process $\{X_t\}_{t \in \mathcal{T}}$ is a *martingale* provided

- i. $\{X_t\}_{t\in\mathcal{T}}$ is adaptive to $\{\mathcal{F}_t\}_{t\in\mathcal{T}}$
- ii. $\mathbb{E}[|X_t|] < \infty$
- iii. $\forall s < t, t, s \in \mathcal{T}, \mathbb{E}[X_t \mid \mathcal{F}_s] = X_s$ almost surely

Example 2.2.2. Let $X \sim N(\mu, \sigma^2)$ and we only know X at time 1.

Define $\{\mathcal{F}_t\}_{t \in [0,1]}$ and $X_t = \tilde{\mathbb{E}}[X \mid \mathcal{F}_t]$. Then, $\forall 0 \leq s < t \leq 1$, we have

$$\mathbb{E}[X_t \mid \mathcal{F}_s] = \mathbb{E}[\mathbb{E}[X \mid \mathcal{F}_t] \mid \mathcal{F}_s]$$

$$= \mathbb{E}[X \mid \mathcal{F}_s] \quad \text{by tower property}$$

$$= X_s \quad \text{almost surely}$$

Therefore, $\{X_t = \mathbb{E}[X \mid \mathcal{F}_t]\}_{t \in [0,1]}$ is a martingale with respect to $\{\mathcal{F}_t\}_{t \in \mathcal{T}}$ by definition of martingale.

Definition 2.2.3. On space $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \in \mathcal{F}}, \mathbb{P})$, a stochastic process $\{X_t\}_{t \in \mathcal{T}}$ is a *sub-martingale* provided

- i. $\{X_t\}_{t\in\mathcal{T}}$ is adaptive to $\{\mathcal{F}_t\}_{t\in\mathcal{T}}$
- ii. $\mathbb{E}[|X_t|] < \infty$
- iii. $\forall s < t, t, s \in \mathcal{T}, \mathbb{E}[X_t \mid \mathcal{F}_s] \geq X_s$ almost surely

2.3 Multi-period Binomial Model

A multi-period binomial model is

$$S_t(\omega_1\omega_2\dots\omega_t) = uS_t(\omega_1\omega_2\dots\omega_t)$$
$$S_t(\omega_1\omega_2\dots\omega_t) = dS_t(\omega_1\omega_2\dots\omega_t)$$

where u, d, r does not depend on $t, \omega_i \in \{H, T\}, i = 1, 2, \dots, t$. For an N-period binomial model, $\omega = \omega_1 \omega_2 \cdots \omega_N \in \Omega$.

Example 2.3.1. (2-Period). Consider a 2-period binomial model. Assume $r = \frac{1}{4}$ and $u = \frac{1}{2} = \frac{1}{d}$.

$$S_{1}(H) = 8$$

$$S_{2}(HH) = 16$$

$$S_{2}(HT) = S_{2}(TH) = 4$$

$$S_{1}(T) = 2$$

$$S_{2}(TT) = 1$$

Under the physical measure, the branching probabilities are $p = \frac{1}{2} = q$.

Under the risk-neutral measure, the discounted stock price is a martingale.

The probability sets are $\Omega = \{HH, HT, TH, TT\}, N_{\Omega} = 4, \mathcal{F} = \{\emptyset, \Omega, \{HH\}, \{HT\}, \{TH\}, \{TT\}, \{HH, HT\}, \{HH, HT\}$ $\{HH, TH\}, \{HH, TT\}, \{HT, TH\}, \{HT, TT\}, \{TH, TT\}, \{HH, HT, TH\}, \{HH, TH, TT\}, \{HT, TH, TT\}\},$ and $N_{\mathbb{F}} = 2^4 = 16.$

Then we know $\mathcal{F}_0 = \{\varnothing, \Omega\}, \mathcal{F}_1 = \{\varnothing, \Omega, \{HH, HT\}, \{TH, TT\}\}, \mathcal{F}_2 = \mathcal{F}$. Therefore, $\mathcal{F}_0 \subset \mathcal{F}_1 \subset \mathcal{F}_2$.

From t = 0 to t = 1, we have

$$\begin{cases} \tilde{p}_1 + \tilde{q}_1 = 1 \\ \tilde{\mathbb{E}} \left[\frac{S_1}{1+r} \mid S_0 \right] = \frac{1}{1+r} \left[\tilde{p}_1 S_1(H) + \tilde{q}_1 S_1(T) \right] \end{cases}$$

$$\Rightarrow \begin{cases} \tilde{p}_1 + \tilde{q}_1 = 1 \\ \frac{S_0}{1+r} (u \tilde{p}_1 + d \tilde{q}_1) = S_0 \end{cases}$$

$$\Rightarrow \begin{cases} \tilde{p}_1 = \frac{1+r-d}{u-d} > 0 \\ \tilde{q}_1 = \frac{u-(1+r)}{u-d} > 0 \end{cases} \text{ with } d < 1 + r < u \end{cases}$$

From t = 1 to t = 2, we have

$$\begin{cases} \tilde{p}_{2H} + \tilde{q}_{2H} = 1 \\ \tilde{\mathbb{E}} \left[\frac{S_2}{(1+r)^2} \mid S_1(H) \right] = \frac{S_1(H)}{(1+r)^2} (u \tilde{p}_{2H} + d \tilde{q}_{2H}) \end{cases} \Rightarrow \begin{cases} \tilde{p}_{2H} = \frac{1+r-d}{u-d} \\ \tilde{q}_{2H} = \frac{u-(1+r)}{u-d} \end{cases}$$

and

$$\begin{cases} \tilde{p}_{2T} + \tilde{q}_{2T} = 1 \\ \tilde{\mathbb{E}} \left[\frac{S_2}{(1+r)^2} \mid S_1(T) \right] = \frac{S_1(T)}{(1+r)^2} (u\tilde{p}_{2T} + d\tilde{q}_{2T}) \end{cases} \Rightarrow \begin{cases} \tilde{p}_{2T} = \frac{1+r-d}{u-d} \\ \tilde{q}_{2T} = \frac{u-(1+r)}{u-d} \end{cases}$$

Theorem 2.3.2. (Risk-Neutral Probabilities). For each of the branch, d < 1 + r < u, the risk-neutral probabilities for the next period are $\tilde{p} = \frac{1+r-d}{u-d}$ and $\tilde{q} = \frac{u-(1+r)}{u-d}$.

Remark 2.3.3. *u*, *d*, *r* remain unchanged with respect to *t*. This risk-neutral measure is unique.

Example 2.3.4. (3-Period European Put Option). Consider an European put option whose option payoff is calculated as $V_t = (K - S_t)^+$ with $u = 2 = \frac{1}{2}$, $r = \frac{1}{4}$, and K = 6.

$$V_t=(K-S_t)^+ \text{ with } u=2=\frac{1}{2}, r=\frac{1}{4}, \text{ and } K=6.$$

$$S_3(HHH)=32$$

$$S_2(HH)=16$$

$$S_3(HHT)=S_3(HTH)=S_3(THH)=8$$

$$S_1(T)=2$$

$$S_2(TT)=1$$

$$S_3(TTT)=\frac{1}{2}$$
 Then we know $\tilde{p}=\frac{1}{2}=\tilde{q}$. The option payoffs at $t=3$ are

Then we know $\tilde{p} = \frac{1}{2} = \tilde{q}$. The option payoffs at t = 3 are

$$V_3(HHH) = (6-32)^+ = 0$$

$$V_3(HHT) = V_3(HTH) = V_3(THH) = (6-8)^+ = 0$$

$$V_3(HTT) = V_3(THT) = V_3(TTH) = (6-2)^+ = 4$$

$$V_3(TTT) = (6-\frac{1}{2})^+ = 5.5$$

We discount the option value back to t=2: $\tilde{\mathbb{E}}\left[\frac{V_3}{1+r}\mid \mathcal{F}_2\right]$

$$V_{2}(HH) = \frac{1}{1+r} \left(\tilde{p}V_{3}(HHH) + \tilde{q}V_{3}(HHT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0 + \frac{1}{2} \cdot 0 \right) = 0$$

$$V_{2}(HT) = \frac{1}{1+r} \left(\tilde{p}V_{3}(HTH) + \tilde{q}V_{3}(HTT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0 + \frac{1}{2} \cdot 4 \right) = 1.6$$

$$V_{2}(TH) = \frac{1}{1+r} \left(\tilde{p}V_{3}(THH) + \tilde{q}V_{3}(THT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0 + \frac{1}{2} \cdot 4 \right) = 1.6$$

$$V_{2}(TT) = \frac{1}{1+r} \left(\tilde{p}V_{3}(TTH) + \tilde{q}V_{3}(TTT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 4 + \frac{1}{2} \cdot 5.5 \right) = 3.8$$

We discount the option value back to t=1: $\tilde{\mathbb{E}}\left[\frac{V_2}{1+r}\mid \mathcal{F}_1\right]$

$$V_1(H) = \frac{1}{1+r} \left(\tilde{p}V_2(HH) + \tilde{q}V_2(HT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0 + \frac{1}{2} \cdot 1.6 \right) = 0.64$$

$$V_1(T) = \frac{1}{1+r} \left(\tilde{p}V_2(TH) + \tilde{q}V_2(TT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 1.6 + \frac{1}{2} \cdot 3.8 \right) = 2.16$$

We discount the option value back to t=0: $\tilde{\mathbb{E}}\left[\frac{V_1}{1+r}\mid \mathcal{F}_0\right]$

$$V_0 = \frac{1}{1+r} \left(\tilde{p}V_1(H) + \tilde{q}V_1(T) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0.64 + \frac{1}{2} \cdot 2.16 \right) = 1.12$$

Example 2.3.5. (3-Period Lookback Option). Consider a lookback option whose option payoff is calculated as $V_t = \max_{0 < t \le T} S_t - S_T$ with $u = 2 = \frac{1}{2}$ and $r = \frac{1}{4}$.

$$S_{3}(HHH) = 32$$

$$S_{3}(HHH) = 32$$

$$S_{3}(HHH) = S_{3}(HHH) = S_{3}($$

Then we know $\tilde{p} = \frac{1}{2} = \tilde{q}$. The option payoffs at t = 3 are

$$V_3(HHH) = 32 - 32 = 0$$

$$V_3(HHT) = 16 - 8 = 8$$

$$V_3(HTH) = 8 - 8 = 0$$

$$V_3(HTT) = 8 - 2 = 6$$

$$V_3(THH) = 8 - 8 = 0$$

$$V_3(THT) = 4 - 2 = 2$$

$$V_3(TTH) = 4 - 2 = 2$$

$$V_3(TTT) = 4 - 0.5 = 3.5$$

We discount the option value back to t=2: $\tilde{\mathbb{E}}\left[\frac{V_3}{1+r}\mid \mathcal{F}_2\right]$

$$V_{2}(HH) = \frac{1}{1+r} \left(\tilde{p}V_{3}(HHH) + \tilde{q}V_{3}(HHT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0 + \frac{1}{2} \cdot 8 \right) = 3.2$$

$$V_{2}(HT) = \frac{1}{1+r} \left(\tilde{p}V_{3}(HTH) + \tilde{q}V_{3}(HTT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0 + \frac{1}{2} \cdot 6 \right) = 2.4$$

$$V_{2}(TH) = \frac{1}{1+r} \left(\tilde{p}V_{3}(THH) + \tilde{q}V_{3}(THT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0 + \frac{1}{2} \cdot 2 \right) = 0.8$$

$$V_{2}(TT) = \frac{1}{1+r} \left(\tilde{p}V_{3}(TTH) + \tilde{q}V_{3}(TTT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 2 + \frac{1}{2} \cdot 3.5 \right) = 2.2$$

We discount the option value back to t=1: $\tilde{\mathbb{E}}\left[\frac{V_2}{1+r}\mid \mathcal{F}_1\right]$

$$V_1(H) = \frac{1}{1+r} \left(\tilde{p}V_2(HH) + \tilde{q}V_2(HT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 3.2 + \frac{1}{2} \cdot 2.4 \right) = 2.24$$

$$V_1(T) = \frac{1}{1+r} \left(\tilde{p}V_2(TH) + \tilde{q}V_2(TT) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 0.8 + \frac{1}{2} \cdot 2.2 \right) = 1.2$$

We discount the option value back to t=0: $\tilde{\mathbb{E}}\left[\frac{V_1}{1+r}\mid \mathcal{F}_0\right]$

$$V_0 = \frac{1}{1+r} \left(\tilde{p}V_1(H) + \tilde{q}V_1(T) \right) = \frac{1}{1+\frac{1}{4}} \left(\frac{1}{2} \cdot 2.24 + \frac{1}{2} \cdot 1.2 \right) = 1.376$$

Then we replicate the portfolio.

From t = 0 to t = 1, we compute

$$\begin{cases} V_1(H) = (1+r)(X_0 - \Delta_0 S_0) + \Delta_0 S_1(H) \\ V_1(T) = (1+r)(X_0 - \Delta_0 S_0) + \Delta_0 S_1(T) \end{cases} \Rightarrow \Delta_0 = \frac{V_1(H) - V_1(T)}{S_1(H) - S_1(T)}$$

To check $V_0 = X_0$, we calculate

$$\begin{split} V_0 &= \frac{1}{1+r} \left[\tilde{p} V_1(H) + \tilde{q} V_1(T) \right] \\ &= \frac{1}{1+r} \left[\tilde{p} \left((1+r) (X_0 - \Delta_0 S_0) + \Delta_0 S_1(H) \right) + \tilde{q} \left((1+r) (X_0 - \Delta_0 S_0) + \Delta_0 S_1(T) \right) \right] \\ &= \frac{1}{1+r} \left[(1+r) (X_0 - \Delta_0 S_0) + \Delta_0 (\tilde{p} S_1(H) + \tilde{q} S_1(T)) \right] \\ &= \frac{1}{1+r} \left[(1+r) (X_0 - \Delta_0 S_0) + \Delta_0 (1+r) S_0 \right] \\ &= X_0 - \Delta_0 S_0 + \Delta_0 S_0 \\ &= X_0 \end{split}$$

From t = 1 to t = 2, we compute

$$\begin{cases} V_2(HH) = (1+r)(X_1(H) - \Delta_1(H)S_1(H)) + \Delta_1(H)S_2(HH) \\ V_2(HT) = (1+r)(X_1(H) - \Delta_1(H)S_1(H)) + \Delta_1(H)S_2(HT) \end{cases} \Rightarrow \Delta_0 = \frac{V_2(HH) - V_2(HT)}{S_2(HH) - S_2(HT)}$$

$$\begin{cases} V_2(TH) = (1+r)(X_1(T) - \Delta_1(T)S_1(T)) + \Delta_1(T)S_2(TH) \\ V_2(TT) = (1+r)(X_1(T) - \Delta_1(T)S_1(T)) + \Delta_1(T)S_2(TT) \end{cases} \Rightarrow \Delta_0 = \frac{V_2(TH) - V_2(TT)}{S_2(TH) - S_2(TT)}$$

3.1 Brownian Motion

Definition 3.1.1. On space $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t\geq 0}, \mathbb{P})$, we say $\{B_t\}_{t\geq 0}$ is a *Brownian Motion (B.M.)* provided

- i. $B_0 = 0$
- ii. $B_t(\omega)$ has continuous path, $\forall \omega \in \Omega$
- iii. $B_t B_s \sim N(0, t s), \forall t \geq s \geq 0$
- iv. $B_t B_s$ is independent of $B_u B_v$, $\forall 0 \le s \le t \le v \le u$

Remark 3.1.2. For simulation, we have $dB_t = \Delta B_t \stackrel{\text{i.i.d.}}{\sim} N(0, \Delta t)$, where Δt is the length of time discretization.

Remark 3.1.3. In *Steve Shreve, Stochastic Calculus for Finance Volume II*, the notation is the Wiener process $\{W_t\}_{t\geq 0}$. In this course, Brownian Motion is the same as Wiener process.

Remark 3.1.4. The mean of a B.M. process is $\mathbb{E}[B_t] = 0$. The variance of a B.M. process is $\text{Var}[B_t] = \mathbb{E}[B_t^2] = t$, i.e., growing linearly with respect to time.

Remark 3.1.5. B_t has continuous sample path with $+\infty$ variation.

Claim:
$$\sum_{t=0}^{N-1} \left| B_{t_{i+1}}(\omega) - B_{t_i}(\omega) \right| \xrightarrow{\max|t_{i+1} - t_i| \to 0} + \infty$$

Aside: For $Z \sim N(0,1)$, the moments are $\mathbb{E}[Z] = 0$, $\mathbb{E}[Z^2] = 1$, $\mathbb{E}[Z^3] = 0$, $\mathbb{E}[Z^4] = 3$. The moment generating function for a standard normal variable is $\mathbb{E}[e^{xZ}] = e^{\frac{1}{2}x^2}$.

Proof. For discretization $(t_0 = 0 < t_1 < \dots < t_N = T)$, $(B_{t_0} = B_0 = 0, B_{t_1}, B_{t_2}, \dots, B_{t_\omega} = B_T)$ follows a joint normal distribution. Then the moments are

$$\mathbb{E}\left[B_{t_{i+1}} - B_{t_i}\right] = 0, \quad \mathbb{E}\left[\left(B_{t_{i+1}} - B_{t_i}\right)^2\right] = t_{i+1} - t_i,$$

$$\mathbb{E}\left[\left(B_{t_{i+1}} - B_{t_i}\right)^3\right] = 0, \quad \mathbb{E}\left[\left(B_{t_{i+1}} - B_{t_i}\right)^4\right] = 3\left(t_{i+1} - t_i\right)^2$$

Let $QV(t_0,\ldots,t_N):=\sum_{i=0}^{N-1}\left(B_{t_{i+1}}-B_{t_i}\right)^2$ be a random variable and its variance is

$$\begin{split} \mathbb{E}\left[\left(\mathbf{QV}(t_0,\ldots,t_N)-T\right)^2\right] &= \mathbb{E}\left[\sum_{i=0}^{N-1}\left((B_{t_{i+1}}-B_{t_i})^2-(t_{i+1}-t_i)\right)^2\right] \quad \text{since } T = \sum_{i=0}^{N-1}\left(t_{i+1}-t_i\right) \\ &= \sum_{i=0}^{N-1}\mathbb{E}\left[\left(B_{t_{i+1}}-B_{t_i}\right)^4-2(t_{i+1}-t_i)(B_{t_{i+1}}-B_{t_i})^2+(t_{i+1}-t_i)^2\right] \\ &= \sum_{i=0}^{N-1}\left[3(t_{i+1}-t_i)^2-2(t_{i+1}-t_i)^2+(t_{i+1}-t_i)^2\right] \\ &= 2\sum_{i=0}^{N-1}\left(t_{i+1}-t_i\right)^2 \\ &\leq 2\cdot\max|t_{i+1}-t_i|\cdot\sum_{i=0}^{N-1}(t_{i+1}-t_i) \\ &= 2T\max|t_{i+1}-t_i| \end{split}$$

Therefore, as $\max |t_{i+1} - t_i| \to 0$, $QV(t_0, \dots, t_N) \to T$.

We can easily get
$$0 < \text{QV}(t_0, \dots, t_N) = \sum_{i=0}^{N-1} (B_{t_{i+1}} - B_{t_i})^2 \le \max |B_{t_{i+1}} - B_{t_i}| \sum_{i=0}^{N-1} |B_{t_{i+1}} - B_{t_i}|$$
.

We also know that, as $\max |t_{i+1} - t_i| \to 0$, $\mathrm{QV}(t_0, \dots, t_N) \to T$ but $\max |B_{t_{i+1}} - B_{t_i}| \to 0$ due to continuity of B.M., then $\sum_{i=0}^{N-1} |B_{t_{i+1}} - B_{t_i}|$ must goes to $+\infty$.

Remark 3.1.6. B_t has a continuous path that is no where differentiable.

Remark 3.1.7. $(dB_t)^2 = dt$ due to the Quadrative Variation (QV) that $\langle B, B \rangle_t = [B, B]_t = \int_0^t (dB_u)^2 = t$.

Note that QV for B.M. is deterministic.

Remark 3.1.8. $dB_t \sim O(\sqrt{dt})$

3.2 Itô's Formula for Brownian Motion

Theorem 3.2.1. (*Itô's Formula for Brownian Motion*). For a function $f : \to \mathbb{R}$, f is twice differentiable, then we have

$$df(B_t) = f'(B_t)dB_t + \frac{1}{2}f''(B_t)dt$$

Proof. Using Taylor's expansion of $f(B_t + dB_t)$, we have

$$df(B_t) = f(B_t + dB_t) - f(B_t)$$

$$= f(B_t) + dB_t f'(B_t) + \frac{1}{2} (dB_t)^2 f''(B_t) + O((dB_t)^2) - f(B_t)$$

$$= f(B_t) + dB_t f'(B_t) + \frac{1}{2} dt f''(B_t) + O((dB_t)^2) - f(B_t)$$

$$= f'(B_t) dB_t + \frac{1}{2} f''(B_t) dt$$

3.3 Stochastic Differential Equations (SDEs)

1. The stochastic differential equation (SDE) for a diffusion process X_t is $dX_t = \mu_t dt + \sigma_t dB_t$.

The solution is $X_t = X_0 + \int_0^t \mu_t du + \int_0^t \sigma_u dB_u$

2. The SDE for a Bachelier process S_t is $dS_t = \mu_t dt + \sigma_t dB_t$.

The solution is $S_t = S_0 + \mu t + \sigma dB_t$.

3. The SDE for a Geometric Brownian Motion (GBM) process S_t is $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$, where \tilde{B}_t represents a B.M. process under risk-neutral measure.

The solution is $S_t = S_0 \exp\left\{\left(r - \frac{1}{2}\sigma^2\right)t + \sigma B_t\right\}$

3.4 Itô's Formula for SDE

Theorem 3.4.1. (Itô's Formula for SDE). Suppose X_t is a stochastic process defined by the SDE:

$$dX_t = \mu(t, X_t)dt + \sigma(t, X_t)dB_t$$

Let $f(t, X_t)$ be a function that is twice differentiable. The differential $df(t, X_t)$ is given by

$$df(t, X_t) = f_t(t, X_t)dt + f_x(t, X_t)dX_t + \frac{1}{2}f_{xx}(t, X_t)(dX_t)^2$$

$$= \left[f_t(t, X_t) + f_x(t, X_t)\mu(t, X_t) + \frac{1}{2}f_{xx}(t, X_t)\sigma^2(t, X_t) \right] dt + f_x(t, X_t)\sigma(t, X_t)dB_t$$
(2)

Proof. Note that since $(dt)^2 = 0$, $(dt)(dB_t) = 0$, and $(dB_t)^2 = dt$, then we have

$$\begin{split} df(t,X_{t}) &= f_{t}(t,X_{t})dt + f_{x}(t,X_{t})dX_{t} + \frac{1}{2}f_{xx}(t,X_{t})(dX_{t})^{2} \\ &= f_{t}(t,X_{t})dt + f_{x}(t,X_{t})\left[\mu(t,X_{t})dt + \sigma(t,X_{t})dB_{t}\right] + \frac{1}{2}f_{xx}(t,X_{t})\left[\mu(t,X_{t})dt + \sigma(t,X_{t})dB_{t}\right]^{2} \\ &= f_{t}(t,X_{t})dt + f_{x}(t,X_{t})\mu(t,X_{t})dt + f_{x}(t,X_{t})\sigma(t,X_{t})dB_{t} + \\ &\frac{1}{2}f_{xx}(t,X_{t})\left[(\mu(t,X_{t})dt)^{2} + (\mu(t,X_{t})dt\sigma(t,X_{t})dB_{t}) + (\sigma(t,X_{t})dB_{t})^{2}\right] \\ &= f_{t}(t,X_{t})dt + f_{x}(t,X_{t})\mu(t,X_{t})dt + f_{x}(t,X_{t})\sigma(t,X_{t})dB_{t} + \frac{1}{2}f_{xx}(t,X_{t})(\sigma^{2}(t,X_{t})dt + f_{x}(t,X_{t})\mu(t,X_{t})dt + f_{x}(t,X_{t})\sigma^{2}(t,X_{t})\right]dt + f_{x}(t,X_{t})\sigma(t,X_{t})dB_{t} \end{split}$$

Example 3.4.2. Consider the SDE for a GBM process, $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$, $\mu(t, S_t) = rS_t$, $\sigma(t, S_t) = \sigma S_t$, $f(x) = \log x$.

We calculate the first and second derivatives for f(x): $f'(x) = \frac{1}{x}$, $f''(x) = -\frac{1}{x^2}$.

By Itô's Formula for SDE (1), we have

$$df(t, S_t) = f_t(t, S_t)dt + f_s(t, S_t)dS_t + \frac{1}{2}f_{ss}(t, dS_t)(dS_t)^2$$

$$\Rightarrow d\log S_t = 0 + \frac{1}{S_t} \left(rS_t dt + \sigma S_t d\tilde{B}_t \right) + \frac{1}{2} \left(-\frac{1}{S_t^2} \right) \left(r^2 S_t^2 (dt)^2 + 2 \cdot rS_t dt \cdot \sigma S_t d\tilde{B}_t + \sigma^2 S_t^2 \left(d\tilde{B}_t \right)^2 \right)$$

$$d\log S_t = r dt + \sigma d\tilde{B}_t - \frac{1}{2}\sigma^2 S_t dt$$

$$d\log S_t = \left(r - \frac{1}{2}\sigma^2 \right) dt + \sigma d\tilde{B}_t$$

$$\Rightarrow \log S_t = \log S_0 + \left(r - \frac{1}{2}\sigma^2 \right) t + \sigma \tilde{B}_t$$

$$\Rightarrow S_t = S_0 \exp \left\{ \left(r - \frac{1}{2}\sigma^2 \right) t + \sigma \tilde{B}_t \right\}$$

By Itô's Formula for SDE (2), we have

$$df(t, S_t) = \left[f_t(t, S_t) + f_s(t, S_t) \mu(t, S_t) + \frac{1}{2} f_{ss}(t, S_t) \sigma^2(t, S_t) \right] dt + f_s(t, S_t) \sigma(t, S_t) dB_t$$

$$\Rightarrow d \log S_t = \left[\frac{1}{S_t} r S_t + \frac{1}{2} \left(-\frac{1}{S_t} \right) \sigma^2 S_t^2 \right] dt + \frac{1}{S_t^2} \sigma S_t d\tilde{B}_t$$

$$d \log S_t = \left(r - \frac{1}{2} \sigma^2 \right) dt + \sigma d\tilde{B}_t$$

$$\Rightarrow \log S_t = \log S_0 + \left(r - \frac{1}{2} \sigma^2 \right) t + \sigma \tilde{B}_t$$

$$\Rightarrow S_t = S_0 \exp \left\{ \left(r - \frac{1}{2} \sigma^2 \right) t + \sigma \tilde{B}_t \right\}$$

Example 3.4.3. Derive the differential for $e^{-qt}S_t$ under physical measure, where $dS_t = \mu S_t dt + \sigma S_t dB_t$.

First, we know that $\mu(t, S_t) = \mu S_t$ and $\sigma(t, S_t) = \sigma S_t$.

We calculate the derivatives for $f(t, S_t) = e^{-qt}S_t$:

$$f_t(t, S_t) = -qe^{-qt}S_t$$

$$f_s(t, S_t) = e^{-qt}$$

$$f_{ss}(t, S_t) = 0$$

Substituting into Itô's Formula (2), we have

$$df(t, S_t) = \left[f_t(t, S_t) + f_s(t, S_t) \mu(t, S_t) + \frac{1}{2} f_{ss}(t, S_t) \sigma^2(t, S_t) \right] dt + f_s(t, S_t) \sigma(t, S_t) dB_t$$

$$= \left(-qe^{-qt} S_t + e^{-qt} \mu S_t + 0 \cdot \sigma^2 S_t^2 \right) dt + e^{-qt} \sigma S_t dB_t$$

$$= e^{-qt} S_t \left[(\mu - q) dt + \sigma \right] dB_t$$

Example 3.4.4. Derive the differential for $S_t = S_0 \exp \left\{ \left(\mu - \frac{1}{2} \sigma^2 \right) t + \sigma B_t \right\}$.

First we calculate the derivatives for $S_t = f(t, B_t) = S_0 \exp\left\{\left(\mu - \frac{1}{2}\sigma^2\right)t + \sigma B_t\right\}$:

$$f_t(t, B_t) = \left(\mu - \frac{1}{2}\sigma^2\right) S_0 \exp\left\{\left(\mu - \frac{1}{2}\sigma^2\right)t + \sigma B_t\right\} = \left(\mu - \frac{1}{2}\sigma^2\right) f(t, B_t)$$

$$f_b(t, B_t) = \sigma S_0 \exp\left\{\left(\mu - \frac{1}{2}\sigma^2\right)t + \sigma B_t\right\} = \sigma f(t, B_t)$$

$$f_{bb}(t, B_t) = \sigma^2 S_0 \exp\left\{\left(\mu - \frac{1}{2}\sigma^2\right)t + \sigma B_t\right\} = \sigma^2 f(t, B_t)$$

Substituting into Itô's Formula (1), we have

$$dS_t = df(t, B_t) = f_t(t, B_t)dt + f_b(t, B_t)dB_t + \frac{1}{2}f_{bb}(t, dB_t)(dB_t)^2$$

$$= \left(\mu - \frac{1}{2}\sigma^2\right)f(t, B_t)dt + \sigma f(t, B_t)dB_t + \frac{1}{2}\sigma^2 f(t, B_t)(dB_t)^2$$

$$= \left(\mu - \frac{1}{2}\sigma^2\right)S_t dt + \sigma S_t dB_t + \frac{1}{2}\sigma^2 S_t dt$$

$$= \mu S_t dt + \sigma S_t dB_t$$

3.5 Martingale in Continuous Time

Definition 3.5.1. An adaptive stochastic process $\{X_t\}_{t\geq 0}$ on $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t\geq 0}, \mathbb{P})$, is a martingale provided

i.
$$\mathbb{E}[|X_t|] < +\infty$$

ii.
$$\mathbb{E}[X_t \mid \mathcal{F}_s] = X_s$$
 a.s. $\forall 0 \leq s \leq t$

 \Leftrightarrow A diffusion process $dX_t = \mu_t dt + \sigma_t dB_t$ is a martingale provided

i.
$$\mu_t = 0$$
 a.s. $\forall t \ge 0$

ii.
$$\mathbb{E}\left[\int_0^T \sigma_t^2 dt\right] < +\infty$$

Remark 3.5.2. The moment generating function (MGF) for $X \sim N(\mu, \sigma^2)$ is $\mathbb{E}\left[e^{\lambda x}\right] = e^{\lambda \mu + \frac{1}{2}\lambda^2 \sigma^2}$

Exercise 3.5.3. Show that the discounted stock price, $e^{-rt}S_t$, where $S_t = S_0 e^{\left(r - \frac{1}{2}\sigma^2\right)t + \sigma B_t}$, is a martingale. For an interest rate r, $dS_t = rS_t dt + \sigma S_t dB_t$, where \tilde{B}_t is a B.M. under risk-neutral measure.

Proof. Method 1: First we need to show that $\mathbb{E}\left[\left|e^{-rt}S_t\right|\right]<+\infty$:

$$\begin{split} \mathbb{E}\left[\left|e^{-rt}S_{t}\right|\right] &= \mathbb{E}\left[e^{-rt}S_{0}e^{\left(r-\frac{1}{2}\sigma^{2}\right)t+\sigma B_{t}}\right] \\ &= \mathbb{E}\left[S_{0}e^{-\frac{1}{2}\sigma^{2}t+\sigma B_{t}}\right] \\ &= S_{0}e^{-\frac{1}{2}\sigma^{2}t}\mathbb{E}\left[e^{\sigma B_{t}}\right] \\ &= S_{0}e^{-\frac{1}{2}\sigma^{2}t}e^{\frac{1}{2}\sigma^{2}t} \quad \text{by MGF and since } B_{t} \sim N(0,t) \\ &= S_{0} < +\infty \end{split}$$

Then for all $t \ge s \ge 0$, we have

$$\mathbb{E}\left[e^{-rt}S_{t} \mid \mathcal{F}_{s}\right] = \mathbb{E}\left[e^{-rt}S_{0}e^{\left(r-\frac{1}{2}\sigma^{2}\right)t+\sigma B_{t}} \middle| \mathcal{F}_{s}\right]$$

$$= e^{-rt}S_{0}e^{\left(r-\frac{1}{2}\sigma^{2}\right)t}\mathbb{E}\left[e^{\sigma B_{t}+\sigma B_{s}-\sigma B_{s}} \middle| \mathcal{F}_{s}\right]$$

$$= e^{-rt}S_{0}e^{\left(r-\frac{1}{2}\sigma^{2}\right)t+\sigma B_{s}}\mathbb{E}\left[e^{\sigma(B_{t}-B_{s})} \middle| \mathcal{F}_{s}\right]$$

$$= e^{-rt}S_{0}e^{\left(r-\frac{1}{2}\sigma^{2}\right)t+\sigma B_{s}}\mathbb{E}\left[e^{\sigma(B_{t}-B_{s})}\right] \quad \text{since } B_{t}-B_{s} \perp \mathcal{F}_{s}$$

$$= e^{-rt}S_{0}e^{\left(r-\frac{1}{2}\sigma^{2}\right)t+\sigma B_{s}}e^{\frac{1}{2}\sigma^{2}(t-s)} \quad \text{by MGF and since } B_{t}-B_{s} \sim N(0,t-s)$$

$$= S_{0}e^{-\frac{1}{2}\sigma^{2}s+\sigma B_{s}}$$

$$= e^{-rs}S_{0}e^{\left(r-\frac{1}{2}\sigma^{2}\right)s+\sigma B_{s}}$$

$$= e^{-rs}S_{s}$$

Therefore, $e^{-rt}S_t$ is a martingalge by definition. \square

Method 2: By Itô's Formula, we calculate the differential for $f(t, S_t) = e^{-rt}S_t$:

$$df(t, S_t) = f_t(t, S_t)dt + f_s(t, S_t)dS_t + \frac{1}{2}f_{ss}(t, dS_t)(dS_t)^2$$
$$= -re^{-rt}S_tdt + e^{-rt}(rS_tdt + \sigma S_tdB_t) + \frac{1}{2} \cdot 0dt$$
$$= \sigma e^{-rt}S_tdB_t$$

Thus, $\mu_t = 0$ for all $t \geq 0$.

Then we need to show that
$$\mathbb{E}\left[\int_0^T \sigma_t^2 \, dt\right] = \mathbb{E}\left[\int_0^T \sigma^2 e^{-2rt} S_t^2 \, dt\right] < +\infty:$$

$$\mathbb{E}[S_t^2] = \mathbb{E}\left[S_0^2 e^{2rt - \sigma^2 t + 2\sigma B_t}\right]$$

$$= S_0^2 e^{2rt - \sigma^2 t} \mathbb{E}\left[e^{2\sigma B_t}\right]$$

$$= S_0^2 e^{2rt - \sigma^2 t + \frac{1}{2} \cdot 4\sigma^2 t} \quad \text{by MGF and since } B_t \sim N(0,t)$$

$$= S_0^2 e^{2rt + \sigma^2 t}$$

$$\mathbb{E}\left[\int_0^T \sigma^2 e^{-2rt} S_t^2 \, dt\right] = \int_0^T \sigma^2 e^{-2rt} \mathbb{E}[S_t^2] \, dt$$

$$= \int_0^T \sigma^2 e^{-2rt} S_0 e^{2rt + \sigma^2 t} \, dt$$

$$= \int_0^T \sigma^2 S_0 e^{\sigma^2 t} \, dt$$

$$= \sigma^2 S_0 \int_0^T e^{\sigma^2 t} \, dt$$

$$= \sigma^2 S_0 \left[\frac{e^{\sigma^2 t}}{\sigma^2}\right]_0^T$$

$$= S_0 e^{\sigma^2 T} < +\infty$$

Therefore, $e^{-rt}S_t$ is a martingale by definition. \square

4.1 Review of Brownian Motion

Definition 4.1.1. On space $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t\geq 0}, \mathbb{P})$, we say $\{B_t\}_{t\geq 0}$ is a *Brownian Motion (B.M.)* provided

- i. $B_0 = 0$
- ii. $B_t(\omega)$ is continuous, $\forall \omega \in \Omega$
- iii. $B_t B_s \sim N(0, t s), \forall 0 \le s \le t$
- iv. $B_t B_s$ is independent of \mathcal{F}_s . In particular, $B_t B_s$ is independent of $B_u B_v$, $\forall 0 \le s \le t \le v \le u$

Remark 3.1.2. $\{B_t\}_{t\geq 0}$ is a Gaussian process with $\mathbb{E}[B_t]=0$ and $\text{Var}[B_t]=t$.

Remark 3.1.3. $B_t(\omega)$ is continuous but nowhere differentiable for all $\omega \in \Omega$.

Remark 3.1.4. $\{B_t\}_{t\geq 0}$ is a martingale on $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t\geq 0})$.

Proof. To prove that the B.M. process $\{B_t\}_{t\geq 0}$ is a martingale, we need to show the two conditions by the definition of martingale.

First, by Cauchy-Schwarz Inequality, we have $\mathbb{E}[|B_t|] \leq \sqrt{\mathbb{E}[B_t^2]} = \sqrt{\mathrm{Var}(B_t)} = \sqrt{t} < +\infty$

Then we evaluate

$$\mathbb{E}[B_t \mid \mathcal{F}_s] = \mathbb{E}[B_t - B_s + B_s \mid \mathcal{F}_s]$$

$$= \mathbb{E}[B_t - B_s] + B_s \quad \text{since } B_t - B_s \perp \mathcal{F}_s$$

$$= B_s \quad \text{since } B_t - B_s \sim N(0, 1)$$

Therefore, a Brownian Motion $\{B_t\}_{t>0}$ is a martingale on $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t>0}, \mathbb{P})$. \square

Example 3.1.5. Prove that $\{B_t^2 - t\}_{t \geq 0}$ is a martingale on $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \geq 0})$.

Proof. To prove $\{B_t^2 - t\}_{t \ge 0}$ is a martingale, we need to show the two conditions in definition of martingale.

First, we can show that $\mathbb{E}[|B_t^2 - t|] \leq \mathbb{E}[|B_t^2|] + |t| = 2t < +\infty$. Then we calculate

$$\mathbb{E}[B_t^2 \mid \mathcal{F}_s] = \mathbb{E}[(B_t - B_s + B_s)^2 \mid \mathcal{F}_s]$$

$$= \mathbb{E}[(B_t - B_s)^2 + 2B_s(B_t - B_s) + B_s^2 \mid \mathcal{F}_s]$$

$$= \mathbb{E}[(B_t - B_s)^2] + 2B_s\mathbb{E}[B_t - B_s] + B_s^2$$

$$= t - s + B_s^2 \quad \text{since } B_t - B_s \sim N(0, t - s)$$

$$\Rightarrow \mathbb{E}[B_t^2 - t \mid \mathcal{F}_s] = B_s^2 - s$$

Therefore, $\{B_t^2 - t\}$ is a martingale. \square

Remark 3.1.6. Quadratic Variation: $\langle B \rangle_t = [B]_t = t = \int_0^t (dB_s)^2, (dB_s)^2 \sim dt.$

4.2 Risk-Neutral Asset Pricing - Probability Approach

- Pros: Closed-form
- Cons: Restricted (GBM-based)

4.2.1 Black-Scholes Equation for European Call Option

Define call option as $C(T, S_T) = (S_T - K)^+$.

Suppose $\{\tilde{B}_t\}_{t\geq 0}$ is the Brownian Motion under risk-neutral measure \mathbb{Q} .

Suppose $dS_t = S_t(rdt + \sigma d\tilde{B}_t)$, where r is the interest rate.

We have proved in the last lecture that the discounted stock price, $\{e^{-rt}S_t\}_{t\geq 0}$, is a martingale under \mathbb{Q} . Specially with GBM, we have $S_t = S_0 e^{\left(r-\frac{1}{2}\sigma^2\right)t+\sigma \tilde{B}_t}$.

To realize risk-neutral pricing, we need to prove $\{e^{-rt}C(t,S_t)\}_{t\in[0,T]}$ is a martingale, i.e., to calculate $C(t,S_t)=e^{-r(T-t)}\mathbb{E}\left[(S_T-K)^+\mid \mathcal{F}_t\right]$.

Recall that for $Z \sim N(0,1)$, the probability density function (pdf) is $\phi(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}z^2}$, and the cumulative distribution function (CDF) is $N(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-\frac{1}{2}y^2} dy$.

Let Y = Z + c. Then we have $\mathbb{P}[Y \ge y] = \mathbb{P}[Z \ge y - c] = N(c - y)$, and we calculate

$$\mathbb{E}\left[Z\mathbb{1}_{[Z\geq c]}\right] = \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2} \cdot y \mathbb{1}_{[y\geq c]} \, dy$$

$$= \int_{c}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2} \cdot y \, dy$$

$$= \frac{1}{\sqrt{2\pi}} \left[e^{-\frac{1}{2}y^2} \cdot \right]_{+\infty}^{c} \quad \text{since } \frac{\partial}{\partial y} e^{-\frac{1}{2}y^2} = -y e^{-\frac{1}{2}y^2}$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}c^2} \quad \text{for any } c \geq 0$$

$$\mathbb{E}\left[e^Z \mathbb{1}_{[Z\geq c]} \right] = \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2} \cdot e^y \mathbb{1}_{[y\geq c]} \, dy$$

$$= \int_{c}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y^2 - 2y)} \, dy$$

$$= \int_{c}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y^2 - 2y + 1) + \frac{1}{2}} \, dy$$

$$= e^{\frac{1}{2}} \int_{c}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y - 1)^2} \, dy$$

$$= \sqrt{e} \int_{c-1}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \, dx \quad \text{by changing variable of } x = y - 1$$

$$= \sqrt{e} N(-(c-1))$$

$$= \sqrt{e} N(-c+1)$$

Then we calculate the call option price:

$$\begin{split} C(t,S_t) &= e^{-r(T-t)} \mathbb{E}[(S_T - K)^+ \mid \mathcal{F}_t] \\ &= e^{-r(T-t)} \mathbb{E}[(S_t e^{\left(r - \frac{1}{2}\sigma^2\right)(T-t) + \sigma(B_T - B_t)} - K)^+ \mid \mathcal{F}_t] \quad \text{by substituting } S_T = S_t e^{\left(r - \frac{1}{2}\sigma^2\right)(T-t) + \sigma(B_T - B_t)} \\ &= e^{-r(T-t)} \mathbb{E}[(S_t e^{\left(r - \frac{1}{2}\sigma^2\right)(T-t) + \sigma\sqrt{T-t}z} - K)^+ \mid \mathcal{F}_t] \end{split}$$

since $B_T - B_t \sim N(0, T - t)$, then $B_T - B_t = \sqrt{T - t}z$, where $Z \sim N(0, 1)$. Also, since $B_T - B_t \perp \mathcal{F}_t$, then $z \perp \mathcal{F}_t$.

Then we calculate the stock price part and strike price part inside the expectation:

$$\left(S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)+\sigma\sqrt{T-t}z}-K\right)^{+} = \left(S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)+\sigma\sqrt{T-t}z}-K\right)\mathbb{1}_{\left[S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)+\sigma\sqrt{T-t}z}\geq K\right]} \\
= \left(S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)+\sigma\sqrt{T-t}z}-K\right)\mathbb{1}_{\left[\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)+\sigma\sqrt{T-t}z\geq \log\left(\frac{K}{S_{t}}\right)\right]} \\
= \left(S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)+\sigma\sqrt{T-t}z}-K\right)\mathbb{1}_{\left[z\geq \frac{\log\frac{K}{S_{t}}-\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)}{\sigma\sqrt{T-t}}\right]}$$

$$\mathbb{E}\left[K\mathbb{1}_{\left[z \ge \frac{\log \frac{K}{S_t} - \left(r - \frac{1}{2}\sigma^2\right)(T - t)}{\sigma\sqrt{T - t}}\right]}\right| \mathcal{F}_t\right] = KN\left(-\frac{\log \frac{K}{S_t} - \left(r - \frac{1}{2}\sigma^2\right)(T - t)}{\sigma\sqrt{T - t}}\right)$$

$$= KN\left(\frac{\log \frac{S_t}{K} + \left(r - \frac{1}{2}\sigma^2\right)(T - t)}{\sigma\sqrt{T - t}}\right)$$

$$= KN(d_-(T - t, S_t))$$

where $d_{-}(z,x) = \frac{\log \frac{x}{K} + \left(r - \frac{1}{2}\sigma^{2}\right)z}{\sigma\sqrt{z}}$.

Then we calculate the other part:

$$\begin{split} &\mathbb{E}\left[S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)+\sigma\sqrt{T-t}z}\mathbb{1}_{[z\geq -d_{-}(T-t,S_{t})]}\Big|\mathcal{F}_{t}\right] \\ &=S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)}\int_{-d_{-}(T-t,S_{t})}^{+\infty}\frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}y^{2}}e^{\sigma\sqrt{T-t}y}\,dy \\ &=S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)}\int_{-d_{-}(T-t,S_{t})}^{+\infty}\frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}\left(y^{2}-2\sigma\sqrt{T-t}y+\sigma^{2}(T-t)\right)+\frac{1}{2}\sigma^{2}(T-t)}\,dy \\ &=S_{t}e^{\left(r-\frac{1}{2}\sigma^{2}\right)(T-t)}e^{\frac{1}{2}\sigma^{2}(T-t)}\int_{-d_{-}(T-t,S_{t})}^{+\infty}\frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}\left(y^{2}-2\sigma\sqrt{T-t}y+\sigma^{2}(T-t)\right)}\,dy \\ &=S_{t}e^{r(T-t)}N(d_{+}(T-t,S_{t})) \end{split}$$

where $d_+(z,x) = d_-(z,x) + \sigma\sqrt{z} = \frac{\log\frac{x}{K} + \left(r + \frac{1}{2}\sigma^2\right)z}{\sigma\sqrt{z}}$

Substiting to the original call option price formula, we get the Black-Scholes-Merton formula for European call option price as $C(t,S_t)=S_tN(d_+(T-t,S_t))-Ke^{-r(T-t)}N(d_-(T-t,S_t))$. Aside, the Black-Scholes-Merton formula for European put option price is $P(t,S_t)=Ke^{-r(T-t)}N(-d_-(T-t,S_t))-S_tN(-d_+(T-t,S_t))$.

4.2.2 Put-Call Parity

We use the put-call parity to derive put option price:

$$(S_T - K)^+ - (K - S_T)^+ = S_T - K$$

$$e^{-rt} (C(t, S_t) - P(t, S_t)) = \tilde{\mathbb{E}} \left[e^{-rT} \left((S_T - K)^+ - (K - S_T)^+ \right) \middle| \mathcal{F}_t \right]$$

$$e^{-rt} (C(t, S_t) - P(t, S_t)) = e^{-rt} S_t - e^{-rT} K$$

Therefore, the put option price is $P(t, S_t) = C(t, S_t) - S_t + e^{-r(T-t)}K$

4.2.3 The Greeks

For the call option formula $C(t,x) = xN(d_+(T-t,x)) - Ke^{-r(T-t)}N(d_-(T-t,x))$, the Greeks are

- 1. Delta (Δ): $C_x(t,x) = N(d_+(T-t,x)) \ge 0$
- 2. Theta (Θ): $C_t(t,x) = -\frac{\sigma x}{2\sqrt{T-t}}N'(d_+(T-t,x)) rKe^{-r(T-t)}N(d_-(T-t,x))$
- 3. Gamma (Γ): $C_{xx}(t,x) = \frac{1}{\sigma\sqrt{T-t}x}N'(d_+(T-t,x)) \geq 0$

4.3 Self-Financing and Arbitrage

Definitions:

- 1. Self-Financing Strategy: $V_t^{\varphi} = \varphi_t S_t + V_t^{\varphi} \varphi_t S_t$
 - (1) Continuous: $dV_t^{\varphi} = \varphi_t dS_t + r(V_t^{\varphi} \varphi_t S_t) dt$
 - (2) Discrete: $\Delta V_t \varphi = \varphi(S_{t+1} S_t) + r(V_t^{\varphi} \varphi_t S_t)$
- 2. Arbitrage Strategy:
 - (1) $\{\varphi_t\}_{t\geq 0}$ is self-financing and $V_0^{\varphi}=0$
 - (2) $\exists t \in [0, +\infty)$ s.t. $\mathbb{P}[V_t^{\varphi} \ge 0] = 1$ and $\mathbb{P}[V_t^{\varphi} > 0] > 0$

Example 4.3.1. Suppose a risk-neutral measure $\mathbb{Q} \sim \mathbb{P}$, $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$.

Show that a discounted self-financing process is a martingale.

Let $\{\varphi_t\}_{t>0}$ be a self-financing process such that

$$\begin{split} dV_t^\varphi &= \varphi_t dS_t + r(V_t^\varphi - \varphi_t S_t) dt \quad \text{from the continuous model of self-financing strategy} \\ &= \varphi_t (rS_t dt + \sigma S_t d\tilde{B}_t) + (rV_t^\varphi - r\varphi_t S_t) dt \quad \text{by substituting } dS_t \\ &= rV_t^\varphi dt + \sigma S_t \varphi_t d\tilde{B}_t \\ \Rightarrow de^{-rt}V_t^\varphi &= -re^{-rt}V_t^\varphi dt + e^{-rt}dV_t^\varphi + \frac{1}{2} \cdot 0 \cdot dt \\ &= -re^{-rt}V_t^\varphi dt + e^{-rt}(rV_t^\varphi dt + \sigma S_t \varphi_t d\tilde{B}_t) \quad \text{by substituting } dV_t^\varphi \\ &= e^{-rt}\sigma S_t \varphi_t d\tilde{B}_t \end{split}$$

Therefore, the discounted wealth $\{e^{-rt}V_t^{\varphi}\}_{t\geq 0}$ is a martingale under the risk-neutral measure \mathbb{Q} .

Example 4.3.2. Replicate the portfolio for an European call, where $C(t, S_t) = V_t^{\varphi}$, in particular, $V_T \varphi = (S_T - K)^+$.

Recall the Black-Scholes equation for an European call is

$$C(t, S_t) = S_t N(d_+(T - t, S_t)) - Ke^{-r(T - t)} N(d_-(T - t, S_t))$$

Then we calculate the differentials for the call option price and the discounted call option price

$$\begin{split} dC(t,S_t) &= C_t(t,S_t)dt + C_s(t,S_t)dS_t + \frac{1}{2}C_{ss}(t,S_t)(dS_t)^2 \\ de^{-rt}C(t,S_t) &= (-re^{-rt}C(t,S_t) + e^{-rt}C_t(t,S_t))dt + e^{-rt}C_s(t,S_t)(rS_tdt + \sigma S_td\tilde{B}_t) + \frac{1}{2}e^{-rt}C_{ss}(t,S_t)(\sigma^2S_t^2dt) \\ &= e^{-rt}\left(-rC(t,S_t) + C_t(t,S_t) + rS_tC_s(t,S_t) + \frac{1}{2}\sigma^2S_t^2C_{ss}(t,S_t)\right)dt + e^{-rt}\sigma S_tC_s(t,S_t)d\tilde{B}_t \\ &= e^{-rt}\sigma S_tC_s(t,S_t)d\tilde{B}_t \quad \text{according to formula of Greeks} \end{split}$$

Therefore, from Example 4.3.1, we know $\varphi_t = C_s(t, S_t)$, replicating the portfolio and also indicating the delta hedging.

4.4 Risk-Neutral Asset Pricing - PDE Approach

PDE Approach:

• $\{e^{-rt}C(t,S_t)\}_{t\geq 0}$ is a martingale under risk-measure \mathbb{Q} .

• PDE Approach: Using Black-Scholes PDE

Pros and Cons:

• Pros: General

• Cons: Limited closed-form solution

Theorem 4.4.1. (4-Step Procedure to Find a Maringale (General)). Suppose we have the SDE

$$dX_t = \mu(t, X_t)dt + \sigma(t, X_t)d\tilde{B}_t$$

and $Y_T = h(X_T)$. We want to find $Y_t := \mathbb{E}[h(X_T) \mid \mathcal{F}_t]$. Then we follow the following steps:

- (1) Assume that there exists $g \in C^2$, g(t, x) such that $g(t, X_t) = Y_t = \mathbb{E}[h(X_T) \mid \mathcal{F}_t]$. Then $\{g(t, X_t)\}_{t>0}$ is a martingale.
- (2) To find $g(t, X_t)$, compute its differential via Itô's Formula:

$$dg(t, X_t) = g_t(t, X_t)dt + g_x(t, X_t)dX_t + \frac{1}{2}g_{xx}(t, X_t)(dX_t)^2$$

$$= (g_t(t, X_t) + g_x(t, X_t)\mu(t, X_t) + \frac{1}{2}g_{xx}(t, X_t)\sigma^2(t, X_t))dt + g_x(t, X_t)\sigma(t, X_t)d\tilde{B}_t$$

(3) Set dt term equal to 0:

$$g_t(t, X_t) + g_x(t, X_t)\mu(t, X_t) + \frac{1}{2}g_{xx}(t, X_t)\sigma^2(t, X_t) = 0$$

(4) Replace (t, X_t) to (t, x) and write down the terminal formula

$$\begin{cases} g_t(t,x) + g_x(t,x)\mu(t,x) + \frac{1}{2}g_{xx}\sigma^2(t,x) = 0\\ g(T,x) = h(x) \end{cases}$$

Example 4.4.2. Apply the 4-step procedure to find a martingale to the discounted call option price.

Suppose we have the SDE

$$dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$$

and
$$e^{-rT}C(T, S_T) = e^{-rT}(S_T - K)^+$$
.

We want to find $e^{-rt}C(t, S_t) = \mathbb{E}[e^{-rT}(S_T - K)^+ \mid \mathcal{F}_t]$. Then we follow the following steps.

Step 1: Assume that there exists $e^{-rt}C(t,s)$ such that $e^{-rt}C(t,S_t)=\mathbb{E}[e^{-rT}(S_T-K)^+\mid \mathcal{F}_t]$.

Then $\{e^{-rt}C(t,S_t)\}_{t\geq 0}$ is a martingale.

Step 2: To find $e^{-rt}C(t,S_t)$, we compute its differential via Itô's Formula:

$$de^{-rt}C(t, S_t) = (-re^{-rt}C(t, S_t) + e^{-rt}C_t(t, S_t))dt + e^{-rt}C_s(t, S_t)dS_t + \frac{1}{2}e^{-rt}C_{ss}(t, S_t)(dS_t)^2$$

$$= e^{-rt}\left(-rC(t, S_t) + C_t(t, S_t) + rS_tC_s(t, S_t) + \frac{1}{2}\sigma^2S_t^2C_{ss}(t, S_t)\right)dt + e^{-rt}\sigma S_tC_s(t, S_t)d\tilde{B}_t$$

Step 3: Setting dt term equal to 0, we have

$$e^{-rt} \left(-rC(t, S_t) + C_t(t, S_t) + rS_tC_s(t, S_t) + \frac{1}{2}\sigma^2 S_t^2 C_{ss}(t, S_t) \right) = 0$$
$$-rC(t, S_t) + C_t(t, S_t) + rS_tC_s(t, S_t) + \frac{1}{2}\sigma^2 S_t^2 C_{ss}(t, S_t) = 0$$

Step 4: Replacing (t, S_t) to (t, x), the terminal value is

$$\begin{cases} -rC(t,x) + C_t(t,x) + rxC_x(t,x) + \frac{1}{2}\sigma^2 x^2 C_{xx}(t,x) = 0\\ C(T,x) = (x-K)^+ \end{cases}$$

by using Black-Scholes PDE.

5.1 Review of Black-Scholes Equation for European Call Option

- 1. Expression: $C(t, S_t) = S_t N(d_+(T t, S_t)) Ke^{-r(T-t)} N(d_-(T t, S_t))$, where
 - (1) $N(\cdot)$ is the CDF of a standard normal distribution.

(2)
$$d_+(z,x) = \frac{\log \frac{x}{K} + \left(r + \frac{1}{2}\sigma^2\right)z}{\sigma\sqrt{z}}$$

(3)
$$d_{-}(z,x) = \frac{\log \frac{x}{K} + \left(r - \frac{1}{2}\sigma^{2}\right)z}{\sigma\sqrt{z}}$$

- 2. Assumptions:
 - (1) No arbitrage \Leftrightarrow There exists risk-neutral measure (FTAP I), and the B.M. process $\{\tilde{B}_t\}_{t\geq 0}$ is under risk-neutral measure
 - (2) Under risk-neutral measure, stock price dynamic follows a GBM, i.e., $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$.
 - (3) The interest r is constant.
- 3. Properties:
 - (1) Discounted option price is a martingale under the risk-neutral measure.

That is,
$$\tilde{\mathbb{E}}[e^{-rt}C(t,S_t) \mid \mathcal{F}_u] = e^{-ru}C(u,S_u)$$
, for all $-\leq u \leq t \leq T$ (or $de^{-rt}C(t,S_t) = e^{-rt}\sigma S_tC_s(t,S_t)d\tilde{B}_t$). In particular, $\tilde{\mathbb{E}}[e^{-rT}(S_T-K)^+ \mid \mathcal{F}_t] = e^{-rt}C(t,S_t)$.

(2) If no trading costs, no liquidity impact, market is frictionless.

The replicating portfolio is $X_t = \Delta_t S_t + (X_t - \Delta_t S_t) = C(t, S_t)$.

Delta-hedging for continuous model is

$$de^{-rt}X_t = de^{-rt}C(t,S_t)$$

$$e^{-rt}\sigma S_t\Delta_t d\tilde{B}_t = e^{-rt}\sigma S_tC_s(t,S_t)d\tilde{B}_t \quad \text{from Example 4.3.1 and 4.3.2}$$

$$\Delta_t = C_s(t,S_t) = N(d_+(T-t,S_t)) \geq 0 \quad \text{for European call, by option Greeks}$$

Recall that delta-hedging for discrete model is

$$\Delta_t = \frac{C(t+1,H) - C(t+1,T)}{S(t+1,H) - S(t+1,T)} \quad \text{and}$$

$$C_s(t,S_t) \approx \frac{C(t+\Delta t, S_{t+\Delta t}(H)) - C(t+\Delta t, S_{t+\Delta t}(T))}{S_{t+\Delta t}(H) - S_{t+\Delta t}(T)}$$

- (3) Put-call parity: $P(t, S_t) = C(t, S_t) + e^{-r(T-t)}K S_t$ At time t = T, it becomes $P(T, S_T) - C(T, S_T) = K - S_T$
- 4. Parameters: T, K, r, σ, t , stock price S_t treated as a random variable.

In Reality: $T, K, r, t, C(0, S_0), \sigma$ unobservable in the market. The Black-Scholes equation in reality is

$$C(0, S_0; T, K) = S_0 N(d_+(T, S_0; K)) - Ke^{-rT} N(d_-(T, S_0; K))$$

5.2 Dupire's Formula

To arrive at the Dupire's formula, we need the first and second order partial derivatives of Black-Scholes equation for European call option with respect to K and the first-order derivative with respect to T.

1. The first-order derivative with respect to K is

$$\frac{\partial C(0, S_0; T, K)}{\partial K} = S_0 N'(d_+(T, S_0; K)) \frac{\partial d_+(T, S_0; K)}{\partial K} - e^{-rT} N(d_-(T, S_0; K)) - Ke^{-rT} N'(d_-(T, S_0; K)) \frac{\partial d_-(T, S_0; K)}{\partial K}$$

First, we calculate the $\frac{\partial d_+(T,S_0;K)}{\partial K}$ and $\frac{\partial d_-(T,S_0;K)}{\partial K}$ parts:

$$\frac{\partial d_{+}(T, S_{0}; K)}{\partial K} = \frac{\partial d_{-}(T, S_{0}; K)}{\partial K} = \frac{\partial}{\partial K} \left(\frac{\log \frac{S_{0}}{K} + \left(r \pm \frac{1}{2}\sigma^{2}\right)T}{\sigma\sqrt{T}} \right) = -\frac{1}{\sigma\sqrt{T}K}$$
(1)

Secondly, we verify that $Ke^{-rT}N'(d_{-}(T, S_0; K)) = S_0N'(d_{+}(T, S_0; K))$:

$$Ke^{-rT}N'(d_{-}(T, S_{0}; K)) = e^{-rT}K\frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}d_{-}^{2}(T, S_{0}; K)}$$

$$= e^{-rT}K\frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}\frac{\left(\log\frac{S_{0}}{K} + \left(r - \frac{1}{2}\sigma^{2}\right)T\right)^{2}}{\sigma^{2}T}}$$

$$= e^{-rT}K\frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}\frac{\left(\log\frac{S_{0}}{K} + \left(r - \frac{1}{2}\sigma^{2}\right)T\right)^{2} - \left(\log\frac{S_{0}}{K} + \left(r + \frac{1}{2}\sigma^{2}\right)T\right)^{2} + \left(\log\frac{S_{0}}{K} + \left(r + \frac{1}{2}\sigma^{2}\right)T\right)^{2}}\right]$$

Using the identity $A^2 - B^2 = (A - B)(A + B)$, we get

$$\begin{split} \left(\log\frac{S_0}{K} + \left(r - \frac{1}{2}\sigma^2\right)T\right) - \left(\log\frac{S_0}{K} + \left(r + \frac{1}{2}\sigma^2\right)T\right) &= -\sigma^2T \\ \left(\log\frac{S_0}{K} + \left(r - \frac{1}{2}\sigma^2\right)T\right) + \left(\log\frac{S_0}{K} + \left(r + \frac{1}{2}\sigma^2\right)T\right) &= 2\left(\log\frac{S_0}{K} + rT\right) \\ \Rightarrow \left(\log\frac{S_0}{K} + \left(r - \frac{1}{2}\sigma^2\right)T\right)^2 - \left(\log\frac{S_0}{K} + \left(r + \frac{1}{2}\sigma^2\right)T\right)^2 &= -2\sigma^2T\left(\log\frac{S_0}{K} + rT\right) \end{split}$$

Thus, we continue to calculate that

$$Ke^{-rT}N'(d_{-}(T, S_{0}; K)) = e^{-rT}K \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2\sigma^{2}T} \left[-2\sigma^{2}T\left(\log\frac{S_{0}}{K} + rT\right) + \left(\log\frac{S_{0}}{K} + \left(r + \frac{1}{2}\sigma^{2}\right)T\right)^{2}\right]}$$

$$= K \frac{1}{\sqrt{2\pi}} e^{\log\frac{S_{0}}{K}} e^{-\frac{1}{2}d_{+}^{2}(T, S_{0}; K)}$$

$$= K \frac{S_{0}}{K}N'(d_{+}(T, S_{0}; K))$$

$$= S_{0}N'(d_{+}(T, S_{0}; K)) \qquad (2)$$

Hence, together by equation (1) and (2), the equation is only left with

$$\frac{\partial C(0, S_0; T, K)}{\partial K} = -e^{-rT} N(d_{-}(T, S_0; K)) \tag{*}$$

2. The second-order derivative with respect to K is

$$\frac{\partial^2}{\partial K^2}C(0, S_0; T, K) = -e^{-rT}N'(d_-(T, S_0; K))\frac{\partial}{\partial K}d_-(T, S_0; K) = \frac{e^{-rT}}{\sigma\sqrt{T}K}N'(d_-(T, S_0; K)) \tag{**}$$

3. The first-order derivative with respect to T is

$$\frac{\partial C(0, S_0; T, K)}{\partial T} = S_0 N'(d_+(T, S_0; K)) \frac{\partial d_+(T, S_0; K)}{\partial T} + rKe^{-rT} N(d_-(T, S_0; K)) - Ke^{-rT} N'(d_-(T, S_0; K)) \frac{\partial d_-(T, S_0; K)}{\partial T}$$

From equation (2), we only needs to calculate the $\frac{\partial d_+(T,S_0;K)}{\partial T} - \frac{\partial d_-(T,S_0;K)}{\partial T}$ parts:

$$\begin{split} \frac{\partial d_{+}(T,S_{0};K)}{\partial T} - \frac{\partial d_{-}(T,S_{0};K)}{\partial T} &= \frac{\partial}{\partial T} \left(\frac{1}{\sigma\sqrt{T}} \log \frac{S_{0}}{K} + \left(r + \frac{1}{2}\sigma \right) \sqrt{T} - \left(\frac{1}{\sigma\sqrt{T}} \log \frac{S_{0}}{K} + \left(r - \frac{1}{2}\sigma \right) \sqrt{T} \right) \right) \\ &= \frac{\partial}{\partial T} \sigma \sqrt{T} \\ &= \frac{\sigma}{2\sqrt{T}} \end{split}$$

Then we continue to calculate that

$$\frac{\partial C(0, S_0; T, K)}{\partial T} = Ke^{-rT}N'(d_{-}(T, S_0; K))\frac{\sigma}{2\sqrt{T}} + rKe^{-rT}N(d_{-}(T, S_0; K)) \qquad (***)$$

Combining steps 1, 2, and 3, we can observe that

$$\frac{\partial C(0,S_0;T,K)}{\partial T} = Ke^{-rT}N'(d_-(T,S_0;K))\frac{\sigma}{2\sqrt{T}} + rKe^{-rT}N(d_-(T,S_0;K))$$

$$\frac{\partial}{\partial T}C(0,S_0;T,K) = \frac{1}{2}K^2\sigma^2\frac{\partial^2}{\partial K^2}C(0,S_0;T,K) - rK\frac{\partial}{\partial K}C(0,S_0;T,K)$$

$$\sigma_{\rm BS}^2 = \frac{\frac{\partial}{\partial T}C(0,S_0;T,K) + rK\frac{\partial}{\partial K}C(0,S_0;T,K)}{\frac{1}{2}K^2\frac{\partial^2}{\partial K^2}C(0,S_0;T,K)} \quad \text{by substituting equations (*) and (**)}$$

$$= \frac{Ke^{-rT}N'(d_-(T,S_0;K))\frac{\sigma}{2\sqrt{T}}}{\frac{1}{2}K^2\frac{e^{-rT}}{\sigma\sqrt{T}K}N'(d_-(T,S_0;K))} \quad \text{verifying}$$

$$= \sigma^2$$

Thus, the implied volatility σ_{imp}^2 is to find the solution of σ such that $C(0, S_0; T, K) = \text{observation}$.

Note that the assumptions for Dupire's formula are: 1) under Black-Scholes, 2) r constant, and 3) GBM for S.

5.3 Different Volatility Models

- 1. Physical Measure: $dS_t = \mu_t dt + \sigma_t dB_t$, where S_t is the observed stock price at time t.
- 2. Market Volatility: $\hat{\sigma}_t = \frac{\text{std}(S_t)}{S_t}$ within a short time horizon, usually around 20%. $(dt \sim (dB_t)^2 \Rightarrow dB_t = \sqrt{dt} >> dt)$ Note that in $\mu_t = \frac{S_{t+\Delta t} - S_t}{\Delta t}$, where Δt is relatively large.
- 3. Implied Volatility (1 data): $C(0, S_0; T, K) = S_0 N(d_+(T, S_0; K)) Ke^{-rT} N(d_-(T, S_0; K))$, which is obtained from reverse calculation from Black-Scholes pricing formula in reality:
- 4. Black-Scholes Volatility (> 1 data): $\sigma_{BS}^2 = \frac{\frac{\partial}{\partial T}C(0,S_0;T,K) + rK\frac{\partial}{\partial K}C(0,S_0;T,K)}{\frac{1}{2}K^2\frac{\partial^2}{\partial K^2}C(0,S_0;T,K)}$, which is obtained from Dupire's formula.
- 5. Local Volatility: $dS_t = rS_t dt + \sigma(t, S_t) S_t d\tilde{B}_t$
- 6. Stochastic Volatility:
 - (1) Cox-Ingersoll-Ross (CIR) Model: $dS_t = rS_t dt + \sqrt{v_t} S_t d\tilde{B}_t$
 - (2) Heston Model: $dV_t = b(t, v_t)dt + \alpha(t, v_t)d\tilde{B}_t + \beta(t, v_t)(d\tilde{B}_t)^2$

5.4 Review of Risk-Neutral Asset Pricing Approaches

- 1. Probability Approach:
 - (1) Find the distribution of S_T under the risk-neutral measure.
 - (2) Calculate $\tilde{\mathbb{E}}[e^{-r(T-t)}f(S_T) \mid \mathcal{F}_t]$.
- 2. PDE Approach:
 - (1) Assume that there exists $g \in C^2$, g(t, x) such that $g(t, X_t) = Y_t = \mathbb{E}[h(X_T) \mid \mathcal{F}_t]$. Then $\{g(t, X_t)\}_{t>0}$ is a martingale.
 - (2) To find $g(t, X_t)$, compute its differential via Itô's Formula:

$$dg(t, X_t) = g_t(t, X_t)dt + g_x(t, X_t)dX_t + \frac{1}{2}g_{xx}(t, X_t)(dX_t)^2$$

= $(g_t(t, X_t) + g_x(t, X_t)\mu(t, X_t) + \frac{1}{2}g_{xx}(t, X_t)\sigma^2(t, X_t))dt + g_x(t, X_t)\sigma(t, X_t)d\tilde{B}_t$

(3) Set dt term equal to 0:

$$g_t(t, X_t) + g_x(t, X_t)\mu(t, X_t) + \frac{1}{2}g_{xx}(t, X_t)\sigma^2(t, X_t) = 0$$

(4) Replace (t, X_t) to (t, x) and write down the terminal formula

$$\begin{cases} g_t(t,x) + g_x(t,x)\mu(t,x) + \frac{1}{2}g_{xx}\sigma^2(t,x) = 0\\ g(T,x) = h(x) \end{cases}$$

Exercise 5.4.1. Consider the Asian option of the stock S_t with maturity T, and the terminal payoff is given by $\left(\frac{1}{T}\int_0^T S_t \, dt - K\right)^+$. Using the 4-step procedure to establish the PDF for the price of the option.

Let $Y_t = \int_0^t S_u du \Rightarrow dY_t = S_t dt$, where $Y_0 = 0$.

The stock price S_t follows $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$ under risk-neutral measure.

Step 1: Assume there exists a martingale:

$$g(t, S_t, Y_t) = \tilde{\mathbb{E}} \left[e^{-r(\mathbf{T} - t)} \left(\frac{1}{T} \int_0^T S_t dt - K \right)^+ \middle| \mathcal{F}_t \right] = \tilde{\mathbb{E}} \left[e^{-r(\mathbf{T} - t)} \left(\frac{1}{T} Y_{\mathbf{T}} - K \right)^+ \middle| \mathcal{F}_t \right]$$

Step 2: Apply Itô's formula for discounted option price:

$$\begin{split} dg(t,S_t,Y_t) &= g_t(t,S_t,Y_t)dt + g_s(t,S_t,Y_t)dS_t + \frac{1}{2}g_{ss}(t,S_t,Y_t)(dS_t)^2 + g_y(t,S_t,Y_t)dY_t + \frac{1}{2}g_{yy}(t,S_t,Y_t)(dY_t)^2 \\ &+ g_{sy}(t,S_t,Y_t)(dS_tdY_t) \quad \text{by Itô's formula for } C^2 \text{ function} \\ &= g_t(t,S_t,Y_t)dt + g_s(t,S_t,Y_t)(rS_tdt + \sigma S_td\tilde{B}_t) + \frac{1}{2}g_{ss}(t,S_t,Y_t)(rS_tdt + \sigma S_td\tilde{B}_t)^2 \\ &+ g_y(t,S_t,Y_t)(S_tdt) + \frac{1}{2}g_{yy}(t,S_t,Y_t)(S_tdt)^2 + g_{sy}(t,S_t,Y_t)(rS_tdt + \sigma S_td\tilde{B}_t)(S_tdt) \\ &= \left(g_t(t,S_t,Y_t) + rS_tg_s(t,S_t,Y_t) + \frac{1}{2}\sigma^2S_t^2g_{ss}(t,S_t,Y_t) + S_tg_y(t,S_t,Y_t)\right)dt + \sigma S_tg_s(t,S_t,Y_t)d\tilde{B}_t \\ de^{-rt}g(t,S_t,Y_t) &= e^{-rt}(-rg(t,S_t,Y_t)dt + dg(t,S_t,Y_t)) \quad \text{by product rule} \\ &= e^{-rt}\left(-rg(t,S_t,Y_t) + g_t(t,S_t,Y_t) + rS_tg_s(t,S_t,Y_t) + \frac{1}{2}\sigma^2S_t^2g_{ss}(t,S_t,Y_t) + S_tg_y(t,S_t,Y_t)\right)dt \\ &+ e^{-rt}\sigma S_tg_s(t,S_t,Y_t)d\tilde{B}_t \end{split}$$

Step 3: Set dt term equal to 0:

$$-rg(t, S_t, Y_t) + g_t(t, S_t, Y_t) + rS_t g_s(t, S_t, Y_t) + \frac{1}{2}\sigma^2 S_t^2 g_{ss}(t, S_t, Y_t) + S_t g_y(t, S_t, Y_t) = 0$$

Step 4: Replace (t, S_t, Y_t) to (t, x, y):

$$\begin{cases} -rg(t,x,y) + g_t(t,x,y) + rxg_x(t,x,y) + \frac{1}{2}\sigma^2 x^2 g_{xx}(t,x,y) + xg_y(t,x,y) = 0\\ g(T,x,y) = (\frac{1}{T}y - K)^+ \end{cases}$$

Exercise 5.4.2. Suppose there is an option of the stock S_t with maturity T, and the terminal payoff is given by $\left(S_T + \int_0^T S_u \, du - K\right)^+$. Using the 4-step procedure to establish the PDE for the price of the option. Can we still have a delta-hedging formula similar to the discrete-time binomial model?

Let
$$Y_t = \int_0^t s_u du \Rightarrow dY_t = S_t dt$$
.

The stock price follows $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$.

Step 1: Assume there exists a martingale:

$$g(t, S_t, Y_t) = \tilde{\mathbb{E}} \left[e^{-r(T-t)} \left(S_T + \int_0^T S_u \, du - K \right)^+ \middle| \mathcal{F}_t \right] = \tilde{\mathbb{E}} [e^{-r(T-t)} (S_T + Y_T - K)^+ \mid \mathcal{F}_t]$$

Step 2: Apply Itô's formula for discounted option price:

$$\begin{split} dg(t,S_t,Y_t) &= g_t(t,S_t,Y_t)dt + g_s(t,S_t,Y_t)dS_t + \frac{1}{2}g_{ss}(t,S_t,Y_t)(dS_t)^2 + g_y(t,S_t,Y_t)dY_t + \frac{1}{2}g_{yy}(t,S_t,Y_t)(dY_t)^2 \\ &+ g_{sy}(t,S_t,Y_t)(dS_tdY_t) \quad \text{by Itô's formula for } C^2 \text{ function} \\ &= g_t(t,S_t,Y_t)dt + g_s(t,S_t,Y_t)(rS_tdt + \sigma S_td\tilde{B}_t) + \frac{1}{2}g_{ss}(t,S_t,Y_t)(rS_tdt + \sigma S_td\tilde{B}_t)^2 \\ &+ g_y(t,S_t,Y_t)(S_tdt) + \frac{1}{2}g_{yy}(t,S_t,Y_t)(S_tdt)^2 + g_{sy}(t,S_t,Y_t)(rS_tdt + \sigma S_td\tilde{B}_t)(S_tdt) \\ &= \left(g_t(t,S_t,Y_t) + rS_tg_s(t,S_t,Y_t) + \frac{1}{2}\sigma^2S_t^2g_{ss}(t,S_t,Y_t) + S_tg_y(t,S_t,Y_t)\right)dt + \sigma S_tg_s(t,S_t,Y_t)d\tilde{B}_t \\ de^{-rt}g(t,S_t,Y_t) &= e^{-rt}(-rg(t,S_t,Y_t)dt + dg(t,S_t,Y_t)) \quad \text{by product rule} \\ &= e^{-rt}\left(-rg(t,S_t,Y_t) + g_t(t,S_t,Y_t) + rS_tg_s(t,S_t,Y_t) + \frac{1}{2}\sigma^2S_t^2g_{ss}(t,S_t,Y_t) + S_tg_y(t,S_t,Y_t)\right)dt \\ &+ e^{-rt}\sigma S_tg_s(t,S_t,Y_t)d\tilde{B}_t \end{split}$$

Step 3: Set dt term equal to 0:

$$-rg(t, S_t, Y_t) + g_t(t, S_t, Y_t) + rS_tg_s(t, S_t, Y_t) + \frac{1}{2}\sigma^2 S_t^2 g_{ss}(t, S_t, Y_t) + S_tg_y(t, S_t, Y_t) = 0$$

Step 4: Replace (t, S_t, Y_t) to (t, x, y):

$$\begin{cases} -rg(t,x,y) + g_t(t,x,y) + rxg_x(t,x,y) + \frac{1}{2}\sigma^2 x^2 g_{xx}(t,x,y) + xg_y(t,x,y) = 0\\ g(T,x,y) = (x+y-K)^+ \end{cases}$$

Yes, we have delta-hedging $\Delta_t = g_x$.

5.5 Numerical Methods

Finite Difference Methods (Pricing/PDE in Deep Learning):

1. For the time derivative:

$$\begin{split} g_t(t,x,y) &\approx \frac{g(t+\Delta t,x,y) - g(t,x,y)}{\Delta t} \quad \text{(forward difference)} \\ &\approx \frac{g(t,x,y) - g(t-\Delta t,x,y)}{\Delta t} \quad \text{(backward difference)} \\ &\approx \frac{g(t+\Delta t,x,y) - g(t-\Delta t,x,y)}{2\Delta t} \quad \text{(central difference)} \end{split}$$

2. For the spatial derivatives:

$$g_x(t, x, y) \approx \frac{g(t, x + \Delta x, y) - g(t, x - \Delta x, y)}{2\Delta x}$$

$$g_y(t, x, y) \approx \frac{g(t, x, y + \Delta y) - g(t, x, y - \Delta y)}{2\Delta y}$$

$$g_{xx}(t, x, y) \approx \frac{g(t, x + \Delta x, y) + g(t, x - \Delta x, y) - 2g(t, x, y)}{(\Delta x)^2}$$

- 3. Start at *T*:
 - (1) Set g(T, x, y) for all x, y.
 - (2) Calculate $g_x(T, x, y)$, $g_{xx}(T, x, y)$, $g_y(T, x, y)$, and use the PDE to obtain $g_t(T, x, y)$.
 - (3) Step backward in time to find $g(T \Delta t, x, y)$.

6.1 Connections between SDE and PDE

The stock price dynamic under risk-neutral measure is $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$, which is the Black-Scholes Model.

The generalized Black-Scholes Model is $dS_t = rS_t dt + S_t \sigma(t, S_t) d\tilde{B}_t$, where $\sigma(t, S_t)$ is a local volatility.

An European-type option, i.e., can only exercise at time T of the underlying asset has terminal payoff G(T). For example, European call payoff is $G(S_T) = (S_T - K)^+$.

Let $g(t, S_t)$ denote the time t price of an option, then by 4-step procedure, we have

- (1) $g(t, S_t) = \tilde{\mathbb{E}}\left[e^{-r(T-t)}G(S_T) \mid \mathcal{F}_t\right]$, where $\{e^{-rt}g(t, S_t)\}_{0 \le t \le T}$ is a martingale.
- (2) $de^{-rt}g(t, S_t)$ via Itô's formula is

$$de^{-rt}g(t,S_t) = e^{-rt}\left(-rg(t,S_t) + g_t(t,S_t) + rS_tg_t(t,S_t) + \frac{1}{2}\sigma^2 S_t^2 g_{ss}(t,S_t)\right)dt + e^{-rt}\sigma S_tg_s(t,S_t)d\tilde{B}_t$$

- (3) Set dt term to 0.
- (4) Replace (t, S_t) to (t, x) and write down PDE system together with the terminal condition

$$\begin{cases} -rg(t,x) + g_t(t,x) + rxg_x(t,x) + \frac{1}{2}x^2\sigma^2(t,x)g_{xx}(t,x) = 0\\ g(T,x) = G(x) \end{cases}$$

which is a generalized Black-Scholes PDE.

Therefore, from a stochastic process martingale, we can have the PDE system, and from the PDE system, we can derive SDE, which returns back to a stochastic process.

6.2 Feynman-Kac Theorem

Theorem 6.2.1. (Feynman-Kac Theorem). Let g(t,x) be a C^2 function satisfying the following PDE:

$$\begin{cases} g_t(t,x) + a(t,x)g_x(t,x) + \frac{1}{2}b^2(t,x)g_{xx}(t,x) = c(t,x)g(t,x) \\ g(T,x) = G(x) \end{cases}$$

where a(t, x), b(t, x), and c(t, x) are known functions.

The solution g(t, x) of the PDE can be represented as

$$g(t,x) = \mathbb{E}\left[e^{-\int_t^T c(u,X_u) du} G(X_T) \middle| \mathcal{F}_t, X_t = x\right]$$

Then a stochastic process $\{X_t\}_{t\geq 0}$ on $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$ satisfies the SDE

$$dX_t = a(t, X_t)dt + b(t, X_t)dB_t, X_0 = x$$

where $\{B_t\}_{t\geq 0}$ is a B.M. process on $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$.

Example 6.2.2. The SDE $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$, $S_0 = s$ has the solution $S_t = se^{\left(r - \frac{1}{2}\sigma^2\right)t + \sigma \tilde{B}_t}$.

Corollary 6.2.3. 4-step approach (PDE)
Feynman-Kac Formula calculating conditional expectation (SDE)

6.3 Radon-Nikodym Derivative

For discrete-time model: $\mathbb{Q} \sim \mathbb{P}$, \tilde{p} , \tilde{q} s.t. $\tilde{p} + \tilde{q} = 1$.

For continuous-time model: $\mathbb{Q} \sim \mathbb{P}$, $\{\tilde{B}_t\}_{t>0}$ under \mathbb{Q} .

For every continuous adaptive process $\{X_t\}_{t\geq 0}$ on $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t\geq 0}, \mathbb{P})$ with the SDE $dX_t = \mu_t dt + \sigma_t dB_t$, where $\{B_t\}_{t\geq 0}$ is a B.M. on \mathbb{P} , then we can apply the martingale representation theorem.

Under physical measure \mathbb{P} , we have the SDE

$$dS_t = S_t \mu(t, S_t) dt + S_t \sigma(t, S_t) dB_t$$

and the bank account

$$dX_t = rX_tdt$$

By subtracting two equations, we get the Sharpe ratio, $\frac{\mu(t,S_t)-r}{\sigma(t,S_t)}$. Aggregating into \mathbb{Q} , we have

$$dS_t = rS_t dt + S_t \sigma(t, S_t) \left(dB_t - \frac{\mu(t, S_t) - r}{\sigma(t, S_t)} dt \right)$$

Note that $dB_t - \frac{\mu(t,S_t) - r}{\sigma(t,S_t)} dt = d\tilde{B}_t$, where $\{\tilde{B}_t\}$ is a B.M. under \mathbb{Q} , and the Sharpe ratio is an adjusted drift term w.r.t. dt.

Example 6.3.1. Let $X \sim N(0,1)$ on $(\Omega, \mathcal{F}, \mathbb{P})$, $\mathbb{P}[X \leq x] = N(x)$.

Let
$$Y = X + \theta \sim N(0, 1), \theta \in \mathbb{R}, \mathbb{P}[Y \le y] = \mathbb{P}[X \le y - \theta] = N(y - \theta).$$

Try to find $\mathbb{Q}(\tilde{\mathbb{P}})$ s.t.

- $(1) \ \tilde{\mathbb{P}}[Y \le y] = N(y)$
- (2) $\tilde{\mathbb{P}} \sim \mathbb{P}$

Solution. First, we let $Z=e^{-\theta X-\frac{1}{2}\theta^2}>0$, $\mathbb{P}[Z=0]=0$. $\forall A\in\mathcal{F}$, we define $\tilde{\mathbb{P}}[A]=\tilde{\mathbb{E}}[\mathbb{1}_A]=\mathbb{E}[Z\mathbb{1}_A]=\int_A Zd\mathbb{P}$.

Then we calculate (1):

$$\begin{split} \tilde{\mathbb{P}}\left[Y \leq y\right] &= \tilde{\mathbb{P}}\left[X \leq y - \theta\right] \\ &= \mathbb{E}\left[Z\mathbbm{1}_{\left[X \leq y - \theta\right]}\right] \\ &= \mathbb{E}\left[e^{-\theta X - \frac{1}{2}\theta^2}\mathbbm{1}_{\left[X \leq y - \theta\right]}\right] \\ &= \frac{1}{\sqrt{2\pi}}\int_{-\infty}^{y - \theta}e^{-\theta x - \frac{1}{2}\theta^2 - \frac{1}{2}x^2}\,dx \\ &= \frac{1}{\sqrt{2\pi}}\int_{-\infty}^{y - \theta}e^{-\frac{1}{2}(x + \theta)^2}\,dx \\ &= \frac{1}{\sqrt{2\pi}}\int_{-\infty}^{y}e^{-\frac{1}{2}z^2}dz \\ &= N(y) \end{split}$$

which shows that Y is a standard normal random variable under the probability measure \mathbb{P} . And Z is the *Radon-Nikodym derivative* of $\tilde{\mathbb{P}}$ with respect to \mathbb{P} , and we write $Z = \frac{d\tilde{\mathbb{P}}}{d\mathbb{P}}$.

Secondly, we evaluate the equivalence. On one hand, $\forall A \in \mathcal{F}, \mathbb{P}[A] = 0 \Rightarrow \tilde{\mathbb{P}}[A] = \mathbb{E}\left[Z\mathbb{1}_A\right] = 0.$

On the other hand, $\tilde{\mathbb{P}}[A] = 0 = \mathbb{E}[Z\mathbb{1}_A]$. Since $Z(\omega) > 0, \forall \omega \in \Omega$, suppose $\mathbb{P}[A] > \epsilon \Rightarrow \tilde{\mathbb{P}}[A] = \mathbb{Z}\mathbb{1}_A > \tilde{\epsilon}$, then $Z\mathbb{1}_A > \tilde{\epsilon}$, which is a contradiction against $\tilde{\mathbb{P}} = 0$, so $\mathbb{P}[A] = 0$ proved by contradiction. Therefore, $\tilde{\mathbb{P}} \sim \mathbb{P}$.

When $Z>0\Rightarrow \frac{1}{Z}>0$, we have $\hat{\mathbb{P}}[A]=\hat{\mathbb{E}}[\mathbb{1}_A]=\tilde{\mathbb{E}}[\frac{1}{Z}\mathbb{1}_A]=\mathbb{E}[\frac{1}{Z}\mathbb{1}_AZ]=\mathbb{P}[A], \forall A\in\mathcal{F}\Rightarrow \frac{1}{Z}=\frac{d\mathbb{P}}{d\hat{\mathbb{P}}}$.

Note that for random variable ξ , $\mathbb{E}[\xi] = \tilde{\mathbb{E}}[\frac{\xi}{Z}] \Rightarrow \tilde{\mathbb{E}}[\xi] = \mathbb{E}[Z\xi]$.

For continuous-time model, $Z = \frac{d\tilde{\mathbb{P}}}{d\mathbb{P}}$, we need to find that for all $t \geq 0$ on space $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t \geq 0}, \mathbb{P})$ such that $Z_t = \mathbb{E}[Z \mid \mathcal{F}_t]$. Note that \mathbb{E} is with respect to \mathbb{P} .

Definition 6.3.2. Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space.

Let $\tilde{\mathbb{P}}$ be another probability measure on (Ω, \mathcal{F}) that is equivalent to \mathbb{P} , i.e., $\tilde{\mathbb{P}} \sim \mathbb{P}$.

Let Z be an almost surely positive random variable that relates $\mathbb P$ and $\tilde{\mathbb P}$ via $\tilde{\mathbb P}(A)=\int_A Z(\omega)d\mathbb P(\omega).$

Then Z is called the *Radon-Nikodym derivative* of $\tilde{\mathbb{P}}$ with respect to \mathbb{P} , and we write

$$Z = \frac{d\tilde{\mathbb{P}}}{d\mathbb{P}}$$

Definition 6.3.3. On probability space space $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$, for $\mathcal{F}_s \subset \mathcal{F}$, we say that $\mathbb{E}[X \mid \mathcal{F}_s]$ is the *conditional expectation* of X on \mathcal{F}_s provided $\forall A \in \mathcal{F}_s$,

$$\int_{A} X \, d\mathbb{P} = \int_{A} \mathbb{E}[X \mid \mathcal{F}_{s}] \, d\mathbb{P}$$

Lemma 6.3.3. Let Z be a Radon-Nikodym derivative of $\tilde{\mathbb{P}}$ and \mathbb{P} on space $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t>0}, \mathbb{P})$, then

- (1) $Z_t = \mathbb{E}[Z \mid \mathcal{F}_t], \{Z_t\}_{t \geq 0}$ is a \mathbb{P} -martingale.
- (2) For an adaptive process $\{Y_t\}_{t\geq 0}$ on space $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$, for $0\leq s\leq t$,

$$\widetilde{\mathbb{E}}[Y_t \mid \mathcal{F}_s] = \frac{1}{Z_s} \mathbb{E}[Y_t Z_t \mid \mathcal{F}_s]$$

In particular, when s = 0, $\tilde{\mathbb{E}}[Y_t] = \mathbb{E}[Y_t Z_t]$.

(3) Let $\{Y_t\}_{t\geq 0}$ be a $\tilde{\mathbb{P}}$ -martingale. Then $\{Y_tZ_t\}$ is a \mathbb{P} -martingale.

Proof.

(1) Since $\mathbb{E}[Z] = \mathbb{E}[Z\mathbbm{1}_{\Omega}] = \tilde{\mathbb{P}}[\Omega] = 1$ and $Z \geq 0$ by definition of the Radon-Nikodym derivative, we have

$$0 \leq \mathbb{E}[|Z|] = \mathbb{E}[Z] = 1 < +\infty$$

For $0 \le s \le t \le T$, we calculate

$$\mathbb{E}[Z_t \mid \mathcal{F}_s] = \mathbb{E}[\mathbb{E}[Z \mid \mathcal{F}_t] \mid \mathcal{F}_s] = \mathbb{E}[Z \mid \mathcal{F}_s] = Z_s$$

Therefore, $Z_t = \mathbb{E}[Z \mid \mathcal{F}_t]$, $\{Z_t\}_{t \geq 0}$ is a \mathbb{P} -martingale.

(2) Consider when s = 0, $\tilde{\mathbb{E}}[Y_t] = \mathbb{E}[Y_t Z_t]$ and $\mathbb{E}[Y_t Z \mid \mathcal{F}_t] = Y_t \mathbb{E}[Z \mid \mathcal{F}_t] = Y_t Z_t$ by taking out what is known. Then we have $\tilde{\mathbb{E}}[Y_t] = \mathbb{E}[\mathbb{E}[Y_t Z \mid \mathcal{F}_t]] = \mathbb{E}[Y_t Z_t]$ by tower property of conditional expectation.

For \mathcal{F}_t -measurable random variable, we only need to know Z_t to change measure.

For $0 \le s \le t$, consider $\forall A \in \mathcal{F}_s, Y_t \mathbb{1}_A$. By definition of conditional expectation, we have

$$\tilde{\mathbb{E}}[Y_t \mathbb{1}_A] = \int_A Y_t d\tilde{\mathbb{P}} = \int_A \tilde{\mathbb{E}}[Y_t \mid \mathcal{F}_s] d\tilde{\mathbb{P}} \quad (1)$$

$$\mathbb{E}[Y_t Z_t \mathbb{1}_A] = \int_A Y_t Z_t d\mathbb{P} = \int_A \mathbb{E}[Y_t Z_t \mid \mathcal{F}_s] d\mathbb{P} \quad (2)$$

When s = 0, above two equations are equivalent.

Let $\eta_s := \tilde{\mathbb{E}}[Y_t \mid \mathcal{F}_s]$. Using similar idea as the unconditional expectation result, we obtain that

$$\begin{split} \tilde{\mathbb{E}}[Y_t \mathbb{1}_A] &= \int_A \eta_s \, d\tilde{\mathbb{P}} \quad \text{from equation (1)} \\ &= \tilde{\mathbb{E}}[\eta_s \mathbb{1}_A] \\ &= \mathbb{E}[Z_s \mathbb{1}_A \eta_s] \quad \text{by change of measure} \\ &= \int_A Z_s \eta_s \, d\mathbb{P} \end{split}$$

By definition of conditional expectation and definition of η_s , we have

$$\mathbb{E}[Y_t Z_t \mid \mathcal{F}_s] = Z_s \eta_s = Z_s \tilde{\mathbb{E}}[Y_t \mid \mathcal{F}_s] \Rightarrow \tilde{\mathbb{E}}[Y_t \mid \mathcal{F}_s] = \frac{1}{Z_s} \mathbb{E}[Y_t Z_t \mid \mathcal{F}_s] \quad (3)$$

(3) Let $\{Y_t\}_{t\geq 0}$ be a $\tilde{\mathbb{P}}$ -martingale and $Z_t = \mathbb{E}[Z \mid \mathcal{F}_t]$. Then we have

$$Y_s = \tilde{\mathbb{E}}[Y_t \mid \mathcal{F}_s] = \frac{1}{Z_s} \mathbb{E}[Y_t Z_t \mid \mathcal{F}_s] \quad \text{from equation (3)}$$

$$\Rightarrow Y_s Z_s = \mathbb{E}[Y_t Z_t \mid \mathcal{F}_s]$$

Since $\mathbb{E}[Y_t Z_t \mid \mathcal{F}_s]$ is well-defined, then $\mathbb{E}[|Y_t Z_t|] < +\infty$.

Therefore, $\{Y_t Z_t\}_{t \geq 0}$ is a \mathbb{P} -martingale.

6.4 Girsanov's Theorem for a Single Brownian Motion

Fact 6.4.1. For an adaptive process $\{X_t\}_{t\geq 0}$ with

$$\mathbb{E}\left[e^{\frac{1}{2}\int_0^t X_u^2 \, du}\right] < +\infty,$$

then we have

$$Y_t := e^{\int_0^t X_u \, dB_u - \frac{1}{2} \int_0^t X_u^2 \, du}$$

is a martingale.

Fact 6.4.2. For adaptive processes $\{X_t^{(1)}\}_{t\geq 0}$, $\{X_t^{(2)}\}_{t\geq 0}$ with

$$\mathbb{E}\left[e^{\frac{1}{2}\int_0^t \left(X_u^{(1)}\right)^2 + \left(X_u^{(2)}\right)^2 du}\right] < +\infty,$$

then we have

$$Y_t := e^{\int_0^t X_u^{(1)} dB_u + \mathbf{i} \int_0^t X_u^{(2)} dB_u - \frac{1}{2} \int_0^t \left(\left(X_u^{(1)} \right)^2 + \mathbf{i} \left(X_u^{(2)} \right)^2 \right) du}$$

is a martingale.

Theorem 6.4.3. (Girsanov's Change of Measure (1-dim)). Let $\{B_t\}_{t\geq 0}$ be a Brownian Motion on the probability space $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$ and $\{\Theta_t\}_{t\geq 0}$ be an adaptive process on $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$.

Focus on [0,T] and assume $\mathbb{E}[e^{\frac{1}{2}\int_0^T\Theta_u^2\,du}]<+\infty$. Define

$$\tilde{B}_t = B_t + \int_0^t \Theta_u \, du$$

$$Z_t = e^{-\int_0^t \Theta_u \, dB_u - \frac{1}{2} \int_0^t \Theta_u^2 \, du}$$

Then for $\tilde{\mathbb{P}}$ defined by $\tilde{\mathbb{E}}[X] := \mathbb{E}[Z_T X]$, for all X being \mathcal{F}_T -measurable, we have that $\{\tilde{B}_t\}_{t\geq 0}$ is a Brownian Motion under $\tilde{\mathbb{P}}$ -measure.

Proof. Since $\tilde{B}_t = B_t + \int_0^t \Theta_u du$, then we have

$$\tilde{B}_0 = B_0 + 0 = 0$$

$$\tilde{B}_t(\omega) = B_t(\omega) + \int_0^t \Theta_u(\omega) du$$

where $\int_0^t \Theta_u(\omega) du$ is continuous, thus satisfying the first two conditions in definition of Brownian Motion.

To arrive at the third and fourth condition in the Brownian Motion definition, we consider $\tilde{B}_t - \tilde{B}_s = B_t - B_s + \int_s^t \Theta_u \, du \, (*)$ and use the characteristic function to capture the probability distribution of $\tilde{B}_t - \tilde{B}_s$:

$$\widetilde{\mathbb{E}}[e^{ik(\tilde{B}_t - \tilde{B}_s)} \mid \mathcal{F}_s] = \frac{1}{Z_s} \mathbb{E}[Z_t e^{ik(\tilde{B}_t - \tilde{B}_s)} \mid \mathcal{F}_s] \quad \text{by Lemma 6.3.3 (2)}$$

Taking out Z_t , we have

$$\begin{split} \frac{Z_t}{Z_s} e^{ik(\tilde{B}_t - \tilde{B}_s)} &= e^{ik(B_t - B_s) + ik \int_s^t \Theta_u \, du - \int_s^t \Theta_u \, dB_u - \frac{1}{2} \int_s^t \Theta_u^2 \, du} \quad \text{from equation } (*) \text{ and definition of } Z_t \\ &= e^{\int_s^t (ik - \Theta_u) \, dB_u - \frac{1}{2} \int_s^t (\Theta_u^2 - ik\Theta_u) \, du} \quad \text{by taking } B_t - B_s \text{ into integral} \\ &= e^{\int_s^t (ik - \Theta_u) \, dB_u - \frac{1}{2} \int_s^t (\Theta_u^2 - ik\Theta_u + (ik\Theta_u)^2 - (ik\Theta_u)^2) \, du} \quad \text{to create a square term of } \Theta_u - ik \\ &= e^{\int_s^t (ik - \Theta_u) \, dB_u - \frac{1}{2} \int_s^t (\Theta_u - ik)^2 \, du - \frac{1}{2} k^2 (t - s)} \end{split}$$

Factor 6.4.1 yields that $Y_t := e^{\int_0^t (ik - \Theta_u) dB_u - \frac{1}{2} \int_0^t (\Theta_u - ik)^2 du}$ is a martingale. Then we can calculate

$$\mathbb{E}[Y_t \mid \mathcal{F}_s] = Y_s$$

$$\Rightarrow \frac{1}{Y_s} \mathbb{E}[Y_t \mid \mathcal{F}_s] = 1$$

$$\Rightarrow \mathbb{E}[\frac{Y_t}{Y_s} \mid \mathcal{F}_s] = 1$$

$$\Rightarrow \mathbb{E}\left[e^{\int_s^t (ik - \Theta_u) dB_u - \frac{1}{2} \int_s^t (\Theta_u - ik)^2 du}\right] = 1 \quad (**)$$

Therefore, we have

$$\begin{split} \tilde{\mathbb{E}}[e^{ik(\tilde{B}_t - \tilde{B}_s)} \mid \mathcal{F}_s] &= \frac{1}{Z_s} \mathbb{E}[Z_t e^{ik(\tilde{B}_t - \tilde{B}_s)} \mid \mathcal{F}_s] \\ &= \mathbb{E}\left[\frac{Z_t}{Z_s} e^{ik(\tilde{B}_t - \tilde{B}_s)} \mid \mathcal{F}_s\right] \\ &= e^{-\frac{1}{2}k^2(t-s)} \mathbb{E}\left[e^{\int_s^t (ik - \Theta_u)dB_u - \frac{1}{2}\int_s^t (\Theta_u - ik)^2 du}\right] \\ &= e^{-\frac{1}{2}k^2(t-s)} \quad \text{from equation (**)} \\ \Rightarrow \tilde{B}_t - \tilde{B}_s \mid \mathcal{F}_s \sim N(0, t-s) \end{split}$$

 $\Rightarrow \tilde{B}_t - \tilde{B}_s \sim N(0, t - s), \tilde{B}_t - \tilde{B}_s \perp \mathcal{F}_s$ satisfying third and fourth conditions in Brownian Motion definition

7.1 Self-Financing and Arbitrage

Recall:

- 1 risky asset, $dS_t = \alpha_t S_t dt + \sigma_t S_t dB_t$
- 1 safe asset, interest rate r
- 1. Self-Financing Strategy:
 - (1) $\{\varphi_t\}_{t>0}, V_0=0$
 - (2) $V_t^{\varphi} = V_t^{\varphi} \varphi_t S_t + \varphi_t S_t$
 - (3) $dV_t^{\varphi} = r(V_t^{\varphi} \varphi_t S_t)dt + \varphi_t dS_t$: instantaneous change in stock
- 2. Arbitrage Strategy:
 - (1) $\{\varphi_t\}_{t\geq 0}$ self-financing, $V_0=0$
 - (2) $\exists t \in [0, +\infty), \mathbb{P}[V_t^{\varphi} \ge 0] = 1, \mathbb{P}[V_t^{\varphi}] > 0$
 - (3) $dV_t^{\varphi} = r(V_t^{\varphi} \varphi_t S_t) dt + \varphi_t dS_t$

Example 7.1.1. r interest rate, $dS_t = \alpha S_t dt + 0 \cdot S_t dB_t$, $\alpha > r$.

 $V_0=0,\,V_t^{arphi}=V_t^{arphi}-arphi_tS_t+arphi_tS_t.$ Calculate differential for discounted self-financing strategy.

Solution.

$$\begin{split} dV_t^\varphi &= r(V_t^\varphi - \varphi_t S_t) dt + \varphi_t dS_t \quad \text{from instantaneous change in stock} \\ &= r(V_t^\varphi - \varphi_t S_t) dt + \varphi_t (\alpha S_t dt) \\ &= (rV_t^\varphi + (\alpha - r)\varphi_t S_t) dt \\ de^{-rt}V_t^\varphi &= -re^{-rt}V_t^\varphi dt + e^{-rt}dV_t^\varphi \quad \text{by Chain Rule} \\ &= -re^{-rt}V_t^\varphi dt + e^{-rt}(rV_t^\varphi + (\alpha - r)\varphi_t S_t) dt \quad \text{by substituting } dV_t^\varphi \\ &= e^{-rt}(\alpha - r)\varphi_t S_t dt \end{split}$$

Since $\alpha > r$ and $\varphi_t S_t > 0$, then the discounted self-financing strategy is positive, thus having arbitrage if we

- borrow $\varphi_t S_t$ from bank
- invest all $\varphi_t S_t$ into stock

7.2 Girsanov's Change of Measure (1-dim)

Example 7.2.1. Suppose 2 risky assets:

$$\begin{split} dS_t^{(1)} &= \alpha^{(1)} S_t^{(1)} dt + \sigma^{(1)} S_t^{(1)} dB_t, & \text{number of stock } 1: \varphi_t^{(1)} \\ dS_t^{(2)} &= \alpha^{(2)} S_t^{(2)} dt + \sigma^{(2)} S_t^{(2)} dB_t, & \text{number of stock } 2: \varphi_t^{(2)} \end{split}$$

and 1 safe asset with interest rate r.

Suppose
$$\alpha^{(1)}>r,$$
 $\alpha^{(2)}>r,$ $\frac{\alpha^{(1)-r}}{\sigma^{(1)}}>\frac{\alpha^{(2)}-r}{\sigma^{(2)}}.$

Solution.

$$\begin{split} V_0 &= 0 \\ V_t^\varphi &= V_t^\varphi - \varphi_t^{(1)} S_t^{(1)} - \varphi_t^{(2)} S_t^{(2)} + \varphi_t^{(1)} S_t^{(1)} + \varphi_t^{(2)} S_t^{(2)} \\ dV_t^\varphi &= r (V_t^\varphi - \varphi_t^{(1)} S_t^{(1)} - \varphi_t^{(2)} S_t^{(2)}) dt + \varphi_t^{(1)} dS_t^{(1)} + \varphi_t^{(2)} dS_t^{(2)} \quad \text{from instantaneous change in stocks} \\ &= r (V_t^\varphi - \varphi_t^{(1)} S_t^{(1)} - \varphi_t^{(2)} S_t^{(2)}) dt + \varphi_t^{(1)} (\alpha^{(1)} S_t^{(1)} dt + \sigma^{(1)} S_t^{(1)} dB_t) + \varphi_t^{(2)} (\alpha^{(2)} S_t^{(2)} dt + \sigma^{(2)} S_t^{(2)} dB_t) \\ &= (r V_t^\varphi + (\alpha^{(1)} - r) \varphi_t^{(1)} S_t^{(1)} + (\alpha^{(2)} - r) \varphi_t^{(2)} S_t^{(2)}) dt + (\sigma^{(1)} \varphi_t^{(1)} S_t^{(1)} + \sigma^{(2)} \varphi_t^{(2)} S_t^{(2)}) dB_t \end{split}$$

For arbitrage, setting dB_t term = 0 and dt term > 0:

$$\begin{split} &\sigma^{(1)}\varphi_t^{(1)}S_t^{(1)} + \sigma^{(2)}\varphi_t^{(2)}S_t^{(2)} = 0 \Rightarrow \varphi_t^{(2)}S_t^{(2)} = -\frac{\sigma^{(1)}}{\sigma^{(2)}}\varphi_t^{(1)}S_t^{(1)} \quad (*) \\ &rV_t^{\varphi} + (\alpha^{(1)} - r)\varphi_t^{(1)}S_t^{(1)} + (\alpha^{(2)} - r)\varphi_t^{(2)}S_t^{(2)} > 0 \\ &rV_t^{\varphi} + \left[(\alpha^{(1)} - r) - \frac{\sigma^{(1)}}{\sigma^{(2)}}(\alpha^{(2)} - r) \right] \varphi_t^{(1)}S_t^{(1)} > 0 \quad \text{by substituting } (*) \\ &rV_t^{\varphi} + \sigma^{(1)} \left[\frac{\alpha^{(1)} - r}{\sigma^{(2)}} - \frac{\alpha^{(2)} - r}{\sigma^{(2)}} \right] \varphi_t^{(1)}S_t^{(1)} > 0 \\ &\Rightarrow \frac{\alpha^{(1)} - r}{\sigma^{(2)}} - \frac{\alpha^{(2)} - r}{\sigma^{(2)}} > 0 \\ &V_0 = 0 = \left(-\varphi_t^{(1)}S_0^{(1)} + \frac{\sigma^{(1)}}{\sigma^{(2)}}\varphi_0^{(1)}S_0^{(1)} \right) + \varphi_t^{(1)}S_0^{(1)} - \frac{\sigma^{(1)}}{\sigma^{(2)}}\varphi_0^{(1)}S_0^{(1)} \end{split}$$

Theorem 7.2.2. (Girsanov's Change of Measure (1-dim)).

$$\mathbb{P} \to \mathbb{O}, \mathbb{O} \sim \mathbb{P}.$$

Let $\{B_t\}_{t\geq 0}$ be a B.M. on $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$

Focusing on [0,T] and assuming for adaptive $\{\Theta_t\}_{0 \leq t \leq T}$ with $\mathbb{E}[e^{\frac{1}{2}\int_0^T \Theta_u^2 \ du}] < +\infty$, define

$$\tilde{B}_t = B_t + \int_0^t \Theta_u \, du$$

$$Z_t = e^{-\int_0^t \Theta_u \, dB_u - \frac{1}{2} \int_0^t \Theta_u^2 \, du}$$

Then for $\tilde{\mathbb{P}}(\mathbb{Q})$ defined by $\mathbb{P}[A] := \mathbb{E}[Z_T \mathbb{1}_A], \forall A \in \mathcal{F}$,

we have $\{\tilde{B}_t\}_{t\geq 0}$ is a B.M. under $\tilde{\mathbb{P}}$.

Exercise 7.2.3. How are the Brownian Motions related between risky asset and risk-free asset. *Solution.*

$$dS_{t} = \alpha_{t} S_{t} dt + \sigma_{t} S_{t} dB_{t} = r S_{t} dt + \sigma_{t} S_{t} d\tilde{B}_{t}$$

$$\Rightarrow d\tilde{B}_{t} - dB_{t} = \frac{\alpha_{t} - r}{\sigma_{t}} dt$$

$$\Rightarrow d\tilde{B}_{t} = dB_{t} + \frac{\alpha_{t} - r}{\sigma_{t}} dt$$

$$\Rightarrow \tilde{B}_{t} = B_{t} + \int_{0}^{t} \frac{\alpha_{u} - r}{\sigma_{u}} du$$

7.3 Girsanov's Change of Measure (d-dim)

Example 7.3.1. m different risky assets, interest rate r.

Under \mathbb{P} ,

$$\begin{split} dS_t^{(i)} &= \alpha_t^{(i)} S_t^{(i)} dt + \sigma_t^{(i)} S_t^{(i)} dB_t^{(i)} \quad i = 1, \dots, m \\ \Theta_t^{(i)} &= \frac{\alpha_t^{(i)} - r}{\sigma_t^{(i)}} \\ dB_t^{(i)} \cdot dB_t^{(j)} &= \rho_{ij} dt \quad i \neq j, 0 \leq |\rho_{ij}| \leq 1 \\ &\Rightarrow dS_t^{(i)} = a_t^{(i)} S_t^{(i)} dt + S_t^{(i)} \sum_{j=1}^d \sigma_t^{ij} dB_t^j, \quad \{B_t^j\}_{t \geq 0} \text{ are independent B.M.} \\ \left(\sigma_t^{(i)}\right)^2 &= \sum_{j=1}^d \left(\sigma_t^{(ij)}\right)^2 \\ V_t^{\varphi} &= \left(V_t^{\varphi} - \varphi_t^{(1)} S_t^{(1)} - \varphi_t^{(2)} S_t^{(2)} - \varphi_t^{(m)} S_t^{(m)}\right) + \varphi_t^{(1)} S_t^{(1)} + \varphi_t^{(2)} S_t^{(2)} + \dots + \varphi_t^{(m)} S_t^{(m)} \\ \varphi_t &= \begin{pmatrix} \varphi_t^{(1)} \\ \vdots \\ \varphi_t^{(m)} \end{pmatrix}, \quad S_t &= \begin{pmatrix} S_t^{(1)} \\ \vdots \\ S_t^{(m)} \end{pmatrix}, \quad \sum_{k=1}^m \varphi_t^{(k)} S_t^{(k)} = \varphi_t^\top S_t, \quad dB_t &= \begin{pmatrix} dB_t^1 \\ \vdots \\ dB_t^d \end{pmatrix} \\ dV_t^{\varphi} &= r(V_t^{\varphi} - \varphi_t^\top S_t) dt + \varphi_t^\top dS_t \\ dS_t &= \begin{pmatrix} \alpha_t^{(1)} S_t^{(1)} \\ \vdots \\ \alpha_t^{(m)} S_t^{(m)} \end{pmatrix} dt + \begin{pmatrix} S_t^{(1)} \sigma_t^{11} & S_t^{(1)} \sigma_t^{12} & \dots & S_t^{(1)} \sigma_t^{1d} \\ \vdots & \vdots & \vdots \\ S_t^{(m)} \sigma_t^{m1} & S_t^{(m)} \sigma_t^{m2} & \dots & S_t^{(m)} \sigma_t^{md} \end{pmatrix} dB_t \\ dV_t^{\varphi} &= r(V_t^{\varphi} - \varphi_t^\top S_t) dt + \varphi_t^\top \begin{pmatrix} dS_t^{(1)} \\ \vdots \\ dS_t^{(m)} \end{pmatrix} dt + \varphi_t^\top \begin{pmatrix} S_t^{(1)} \sigma_t^{11} & S_t^{(1)} \sigma_t^{12} & \dots & S_t^{(1)} \sigma_t^{1d} \\ \vdots & \vdots & \vdots & \vdots \\ S_t^{(m)} \sigma_t^{m1} & S_t^{(m)} \sigma_t^{m2} & \dots & S_t^{(m)} \sigma_t^{md} \end{pmatrix} dB_t \\ dV_t^{\varphi} &= r(V_t^{\varphi} - \varphi_t^\top S_t) dt + \varphi_t^\top \begin{pmatrix} dS_t^{(1)} \\ \vdots \\ dS_t^{(m)} \end{pmatrix} dt + \varphi_t^\top \begin{pmatrix} S_t^{(1)} \sigma_t^{11} & S_t^{(1)} \sigma_t^{12} & \dots & S_t^{(1)} \sigma_t^{1d} \\ \vdots & \vdots & \vdots & \vdots \\ S_t^{(m)} \sigma_t^{m1} & S_t^{(m)} \sigma_t^{m1} & S_t^{(m)} \sigma_t^{m1} & S_t^{(m)} \sigma_t^{m1} & S_t^{(m)} \sigma_t^{m1} \end{pmatrix} dB_t \\ \end{pmatrix}$$

Theorem 7.3.2. (Girsanov's Change of Measure (d-dim)).

 $\mathbb{P} \to \mathbb{Q}, \mathbb{Q} \sim \mathbb{P}.$

Let $\{B_t\}_{t\geq 0}$ be a d-dim independent B.M. on $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$

Focusing on [0,T] and assuming for adaptive d-dim process $\{\Theta_t\}_{t\geq 0}$ with $\|\Theta_t\|^2 = \sum_{j=1}^d \left(\Theta_t^j\right)^2$ and $\mathbb{E}[e^{\frac{1}{2}\int_0^T \|\Theta_u\|^2 \, du}] < +\infty$, define

$$\tilde{B}_t = B_t + \int_0^t \Theta_u \, du$$

$$Z_t = e^{-\int_0^t \Theta_u^T \, dB_u - \frac{1}{2} \int_0^t \|\Theta_u\|^2 \, du}$$

Then for $\widetilde{\mathbb{P}}(\mathbb{Q})$ defined by $\mathbb{P}[A]:=\mathbb{E}[Z_T\mathbb{1}_A], \forall A\in\mathcal{F},$

we have $\{\tilde{B}_t\}_{0 \le t \le T}$ is a B.M. under $\tilde{\mathbb{P}}$.

Example 7.3.3. How to find risk-neutral measure?

Since Under
$$\mathbb{P}: B_t = \begin{pmatrix} B_t^1 \\ \vdots \\ B_t^d \end{pmatrix}$$
, Under $\tilde{\mathbb{P}}: \tilde{B}_t = \begin{pmatrix} \tilde{B}_t^1 \\ \vdots \\ \tilde{B}_t^d \end{pmatrix}$, then we have
$$dS_t = \begin{pmatrix} \alpha_t^{(1)} S_t^{(1)} \\ \vdots \\ \alpha_t^{(m)} S_t^{(m)} \end{pmatrix} dt + \begin{pmatrix} S_t^{(1)} \sigma_t^{11} & \cdots & S_t^{(1)} \sigma_t^{1d} \\ \vdots & \ddots & \vdots \\ S_t^{(m)} \sigma_t^{m1} & \cdots & S_t^{(m)} \sigma_t^{md} \end{pmatrix} dB_t$$

$$\Rightarrow \frac{dS_t}{S_t} = \begin{pmatrix} \frac{dS_t^{(1)}}{S_t^{(1)}} \\ \vdots \\ \frac{dS_t^{(m)}}{S_t^{(m)}} \end{pmatrix} = \begin{pmatrix} \alpha_t^{(1)} \\ \vdots \\ \alpha_t^{(m)} \end{pmatrix} dt + \begin{pmatrix} \sigma_t^{11} & \cdots & \sigma_t^{1d} \\ \vdots & \ddots & \vdots \\ \sigma_t^{m1} & \cdots & \sigma_t^{md} \end{pmatrix} dB_t$$

$$\Rightarrow \mathbb{P}: \frac{dS_t}{S_t} = \alpha_t dt + \sigma_t dB_t \quad (1) \quad \text{where } \alpha_t = \begin{pmatrix} \alpha_t^{(1)} \\ \vdots \\ \alpha_t^{(m)} \end{pmatrix}, \sigma_t = (\sigma_t^{ij}), i = 1, \dots, m, j = 1, \dots, d$$

$$\Rightarrow \mathbb{P}: dS_t^{(i)} = rS_t^{(i)} dt + S_t^{(i)} \sum_{j=1}^d \sigma_t^{ij} d\tilde{B}_t^j$$

$$\frac{dS_t}{S_t} = r\mathbb{1}_m dt + \sigma_t dB_t \quad (2)$$

From equation (1) and (2), we get

$$\alpha_t dt + \sigma_t dB_t = \frac{dS_t}{S_t} = r \mathbb{1}_m dt + \sigma_t d\tilde{B}_t$$

Assume $\tilde{B}_t = B_t + \int_0^t \Theta_u du \Rightarrow d\tilde{B}_t = dB_t + \Theta_t dt$, then we have

$$\sigma_t d\tilde{B}_t = \sigma_t (dB_t + \Theta_t dt)$$

$$\Rightarrow \sigma_t dB_t + (\alpha_t - r \mathbb{1}_m) dt = \sigma_t d\tilde{B}_t + \sigma_t \Theta_t dt$$

$$(\alpha_t - r \mathbb{1}_m) dt = \sigma_t \Theta_t dt$$

$$\Rightarrow \alpha_t^{(i)} - r = \sum_{j=1}^d \sigma_t^{ij} \Theta_t^j$$

There are d unknowns Θ_u^j and m equations $\alpha_t^{(i)} - r = \sum_{j=1}^d \sigma_t^{ij} \Theta_t^j$.

Linear equations:

- 0 solutions for Θ_t
- Exist 1 solution for Θ_t
- Exist ∞ solutions for Θ_t

7.4 Review of FTAP I and FTAP II

Theorem 7.4.1. (First fundamental theorem of asset pricing from Steve Shreve, Stochastic Calculus for Finance Volume II Theorem 5.4.7). If a market model has a risk-neutral probability measure, then it does not admit arbitrage.

Theorem 7.4.2. (Second fundamental theorem of asset pricing from Steve Shreve, Stochastic Calculus for Finance Volume II Theorem 5.4.9). Consider a market model that has a risk-neutral probability measure. The model is complete if and only if the risk-neutral probability measure is unique.

Definition 7.4.3. (Uniqueness of risk-neutral measure from Steve Shreve, Stochastic Calculus for Finance Volume II Definition 5.4.4). A market model is complete if every derivative security can be hedged.

Example 7.4.3. Solve the equation for 2-dim and 3-dim.

$$\begin{split} dS_t^{(1)} &= \alpha^{(1)} S_t^{(1)} dt + \sigma^{(1)} S_t^{(1)} dB_t^1 \\ dS_t^{(2)} &= \alpha^{(2)} S_t^{(2)} dt + \sigma^{(2)} S_t^{(2)} \rho dB_t^1 + \sigma^{(2)} S_t^{(2)} c dB_t^2 \quad \text{where } B_t^1 \perp B_t^2 \end{split}$$

The Black-Scholes volatility of $S_t^{(2)}$ is $\sigma^{(2)}$, thus

$$\begin{split} \left(\sigma^{(2)}\right)^2 \rho^2 + \left(\sigma_t^{(2)}\right)^2 c^2 &= \left(\sigma^{(2)}\right)^2 \Rightarrow c = \sqrt{1 - \rho^2} \\ \frac{d\langle S_t^{(1)}, S_t^{(2)} \rangle}{S_t^{(1)}, S_t^{(2)}} &= \frac{dS_t^{(1)}, S_t^{(2)}}{S_t^{(1)}, S_t^{(2)}} = \frac{\rho \sigma^{(1)} \sigma^{(2)} S_t^{(1)} S_t^{(2)}}{S_t^{(1)} S_t^{(2)}} (dB_t^1)^2 = \rho \sigma^{(1)} \sigma^{(2)} dt \end{split}$$

Assume $\frac{\alpha^{(1)}-r}{\sigma^{(1)}}>\frac{\alpha^{(2)}-r}{\sigma^{(2)}},$ then from Example 7.3.3 we have

$$\begin{cases} \alpha^{(1)} - r = \sigma^{(1)}\Theta_1 \\ \alpha^{(2)} - r = \sigma^{(2)}\Theta_2 = \sigma^{(2)}\rho\Theta_1 + \sigma^{(2)}\sqrt{1 - \rho^2}\Theta_2 \end{cases}$$

$$\Rightarrow \begin{cases} \Theta_1 = \frac{\alpha^{(1)} - r}{\sigma^{(1)}} \\ \Theta_2 = \frac{\alpha^{(2)} - r - \sigma^{(2)}\rho\Theta_t^1}{\sigma^{(2)}\sqrt{1 - \rho^2}} = \frac{1}{\sqrt{1 - \rho^2}} \left(\frac{\alpha^{(2)} - r}{\sigma^{(2)}} - \rho\frac{\alpha^{(1)} - r}{\sigma^{(1)}}\right) \end{cases}$$

For a 3-dim case, we have

$$dS_t^{(1)} = \alpha^{(1)} S_t^{(1)} dt + \sigma^{(1)} S_t^{(1)} dB_t^1$$

$$dS_t^{(2)} = \alpha^{(2)} S_t^{(2)} dt + \sigma^{(2)} S_t^{(2)} \left(\rho dB_t^1 + \tilde{\rho} dB_t^2 + \sqrt{1 - \rho - \tilde{\rho}} dB_t^3 \right)$$

The unknowns are $\Theta_1, \Theta_2, \Theta_3$.

The equations are

$$\begin{cases} \alpha^{(1)} - r = \sigma^{(1)}\Theta_1 \\ \alpha^{(2)} - r = \sigma^{(2)}\rho\Theta_1 + \sigma^{(2)}\tilde{\rho}^2\Theta_2 + \sigma^{(2)}\sqrt{1 - \rho^2 - \tilde{\rho}^2}\Theta_3 \end{cases}$$

Thus, we have infinite solutions of (Θ^2, Θ^3) , given by $\Theta_1 = \frac{\alpha^{(1)} - r}{\sigma^{(1)}}$.

Therefore, not all risk/derivatives are hedgable by FTAP II/Definition 7.4.3.

Example 7.4.4. Assume that the interest rate is not constant, and for the adaptive process $\{R_t\}_{t\geq 0}, R_t\geq 0$.

The discount factor is $D(t,T) := e^{-\int_t^T R_u du}$, or $D_t = e^{-\int_0^t R_u du}$.

Let
$$X_t = \int_0^t R_u du \Rightarrow dX_t = R_t dt$$
. Let $f(x) = e^{-x}$. Then $dD_t = f'(X_t) dX_t = -e^{-X_t} dX_t = -D_t dX_t = -D_t R_t dt$.

The discounted stock price is $D_t S_t$.

In the 1-dim case, $dS_t = R_t S_t dt + \sigma_t S_t d\tilde{B}_t$, where \tilde{B} is risk-neutral. Then we have

$$dD_t S_t = S_t dD_t + D_t dS_t + dD_t dS_t$$

= $-S_t D_t R_t dt + D_t (R_t S_t dt + \sigma_t S_t d\tilde{B}_t)$
= $D_t \sigma_t S_t d\tilde{B}_t$

In the d-dim case, $\{B_t\}_{t\geq 0}$ is independent for $B_t^j, j=1,\ldots,d$. For m risky assets, $\frac{dS_t}{S_t}=\alpha_t dt+\sigma_t dB_t, \alpha_t\in\mathbb{R}^{m\times 1}, \sigma_t\in\mathbb{R}^{m\times d}$.

$$d$$
 unknowns are $\begin{pmatrix} \Theta_t^{(1)} \\ \vdots \\ \Theta_t^d \end{pmatrix}$.

m equations are $\alpha_t - R_t \mathbb{1}_m = \sigma_t \Theta_t$ for each $t \in [0, T]$.

8 Lecture 8

 R_t is the interest rate, which can be constant, deterministic, or stochastic.

In the market where

- 1 stock S_t with interest rate R_t , there exists a risk-neutral $\tilde{\mathbb{P}} \sim \mathbb{P}$
- Zero-coupon bond with all maturities can be traded.

then by FTAP I, there is no arbitrage.

Discount factor : $D_t = e^{-\int_0^t R_u du}$

- is known at time t
- $D_t X_t$ is the discounted value of X_t

To prove D_t is a martingale, for $0 \le s \le t$, we compute

$$\tilde{\mathbb{E}}[D_t \mid \mathcal{F}_s] = \tilde{\mathbb{E}}\left[e^{-\int_0^t R_u du} \middle| \mathcal{F}_s\right] \\
= \tilde{\mathbb{E}}\left[e^{-\int_0^s R_u du} \cdot e^{-\int_s^t R_u du} \middle| \mathcal{F}_s\right] \\
= D_s \tilde{\mathbb{E}}\left[e^{-\int_s^t R_u du} \middle| \mathcal{F}_s\right] \\
= D_s$$

since expectation of discounted value equal to current value under risk-neutral measure, i.e., $\tilde{\mathbb{E}}\left[e^{-\int_s^t R_u du}\Big|\mathcal{F}_s\right]=1$. By definition of a risk-neutral measure, D_tS_t is a martingale since dD_tS_t has no drift term from Example 7.4.4.

8.1 Zero-Coupon Bond

Definition 8.2.1. Zero-coupon bond is represented as

$$B(t,T) = \tilde{\mathbb{E}}\left[e^{-\int_t^T R_u du}\middle|\mathcal{F}_t\right] = \tilde{\mathbb{E}}\left[\frac{D_T}{D_t}\middle|\mathcal{F}_t\right] = \frac{1}{D_t}\tilde{\mathbb{E}}\left[D_T \mid \mathcal{F}_t\right]$$

Exercise 8.1.2 Prove that $D_tB(t,T)$ is also a martingale.

By definition of zero-coupon bond, we have

$$B(t,T) = \frac{1}{D_t} \tilde{\mathbb{E}}[D_T \mid \mathcal{F}_t] \Rightarrow D_t B(t,T) = \tilde{\mathbb{E}}[D_T \mid \mathcal{F}_t] \quad (*)$$

Since $D_T \leq 1$, then we have

$$\tilde{\mathbb{E}}[|D_t B(t,T)|] = \tilde{\mathbb{E}}[|\tilde{\mathbb{E}}[D_T \mid \mathcal{F}_t]|] \le 1 < +\infty$$

For $0 \le s \le t \le T$, we compute

$$\widetilde{\mathbb{E}}\left[D_t B(t,T) \mid \mathcal{F}_s\right] = \widetilde{\mathbb{E}}\left[\widetilde{\mathbb{E}}\left[D_T \mid \mathcal{F}_t\right] \middle| \mathcal{F}_s\right] \quad \text{by tower property}$$

$$= \widetilde{\mathbb{E}}\left[D_T \mid \mathcal{F}_s\right] \quad \text{since } D_t \text{ is a martingale}$$

$$= D_s B(s,T) \quad \text{from equation (*)}$$

8.2 Forward Contract

Definition 8.2.1. (Forward contract from Steve Shreve, Stochastic Calculus for Finance Volume II Definition 5.6.1). A forward contract is an agreement to pay a specified price K at a delivery date T, where $0 \le T \le \overline{T}$, for the asset whose price at time t is S_t .

- have to pay K at time T
- · over-the-counter

Definition 8.2.2. (Forward contract from Steve Shreve, Stochastic Calculus for Finance Volume II Definition 5.6.1). The T-forward price $\operatorname{For}_S(t,S)$ of this asset at time t, where $0 \le t \le T \le \overline{T}$, is the value of K that makes the forward contract have no-arbitrage price zero at time t.

Theorem 8.2.3. (Forward contract price formula from Steve Shreve, Stochastic Calculus for Finance Volume II Theorem 5.6.2). Assume the zero-coupon bonds of all maturities can be traded. Then

$$\operatorname{For}_{S}(t,T) = \frac{S_{t}}{B(t,T)}, \quad 0 \le t \le \bar{T}$$

Remark 8.2.4. Only need risk-neutral measure $\tilde{\mathbb{P}}$

Remark 8.2.5. No assumptions on dS_t

Remark 8.2.6. No assumptions on R_t

Example 8.2.7. Let $X_T = S_T - K$ and $X_t = 0$.

Since the discounted portfolio is a martingale, i.e., $D_t X_t = \tilde{\mathbb{E}}[D_T X_T \mid \mathcal{F}_t]$, then we have

$$0 = X_t = \frac{1}{D_t} \tilde{\mathbb{E}} \left[D_T(S_T - K) \mid \mathcal{F}_t \right] \quad \text{by multiplying the discounted factor to the portfolio}$$

$$= \frac{1}{D_t} \left(D_t S_t - K \tilde{\mathbb{E}} \left[D_T \mid \mathcal{F}_t \right] \right) \quad \text{taking out known}$$

$$= S_t - K \tilde{\mathbb{E}} \left[\frac{D_T}{D_t} \middle| \mathcal{F}_t \right]$$

$$\Rightarrow K = \frac{S_t}{B(t,T)} \quad \text{by definition of zero-coupon bond}$$

Example 8.2.8. At time u > t, $\operatorname{For}_S(u,T) = \frac{S_u}{B(t,T)}$ is the time u value of constant $\operatorname{For}_S(t,T)$ to delivery at time T. Long forward position pays $S_T - \operatorname{For}_S(t,T)$ at time T. To check if a long forward position is a martingale, we compute

$$\tilde{\mathbb{E}}\left[S_{T} - \operatorname{For}_{S}(t,T)|\mathcal{F}_{u}\right] = \tilde{\mathbb{E}}\left[\frac{D_{T}}{D_{u}}S_{T} - \frac{D_{T}}{D_{u}}\operatorname{For}_{S}(t,T)\Big|\mathcal{F}_{u}\right] \\
= S_{u} - \tilde{\mathbb{E}}\left[\frac{D_{T}}{D_{u}}\frac{S_{t}}{B(t,T)}\Big|\mathcal{F}_{u}\right] \\
= S_{u} - S_{t}\tilde{\mathbb{E}}\left[\frac{D_{T}}{D_{u}}\frac{D_{t}}{D_{T}}\Big|\mathcal{F}_{u}\right] \\
= S_{u} - \frac{D_{t}}{D_{u}}S_{t} \quad \Rightarrow \tilde{\mathbb{E}}\left[D_{u}S_{u}|\mathcal{F}_{t}\right] = D_{t}S_{t} \\
\neq S_{u} - \frac{S_{u}}{B(u,T)}$$

Forward price $For_S(t, T)$ is not the price of a forward contract.

8.3 Futures Contract

Futures prices: $\operatorname{Fut}_S(t,T)$, $\{\operatorname{Fut}_S(u,T)\}_{t\leq u\leq T}$, $\operatorname{Fut}_{S_u}(t_{k-1},t_k)=\tilde{\mathbb{E}}[S_T\mid \mathcal{F}_{k-1}]$, $\operatorname{Fut}_S(T,T)=S_T$.

Long position in futures (t_k, t_{k+1}) : receive $\operatorname{Fut}_S(t_{k+1}, T) - \operatorname{Fut}_S(t_k, T)$.

From
$$(0,T)$$
: $\sum_{k=0}^{K} \text{Fut}_{S}(t_{k+1},T) - \text{Fut}_{S}(t_{k},T) = S_{T} - \text{Fut}_{S}(0,T)$

Definition 8.3.1. (Futures contract from Steve Shreve, Stochastic Calculus for Finance Volume II Definition 5.6.4). The futures price of an asset whose value at time T is S_T is given by the formula

$$\operatorname{Fut}_{S}(t,T) = \tilde{\mathbb{E}}[S_{T} \mid \mathcal{F}_{t}], \quad 0 \le t \le T$$

A *long position in the futures contract* is an agreement to receive as a cash flow the changes in the futures price (which may be negative as well as positive) during the time the position is held. A *short position in the futures contract* received the opposite cash flow.

Theorem 8.3.2. (Futures price is a martingale from Steve Shreve, Stochastic Calculus for Finance Volume II Theorem 5.6.5). The futures price is a martingale under the risk-neutral measure $\tilde{\mathbb{P}}$, it satisfies $\operatorname{Fut}_S(T,T)=S_T$, and the value of a long (or a short) futures position to be held over an interval of time is always zero.

Proof. For $0 \le t \le u \le T$, assume $\tilde{\mathbb{E}}[|S_T|] < +\infty$, and we evaluate

$$\begin{split} \tilde{\mathbb{E}}\left[|\operatorname{Fut}_S(t,T)|\right] &= \tilde{\mathbb{E}}\left[\left|\tilde{\mathbb{E}}[S_T \mid \mathcal{F}_t]\right|\right] \\ &\leq \tilde{\mathbb{E}}\left[\tilde{\mathbb{E}}[|S_T| \mid \mathcal{F}_t]\right] \quad \text{by Jensen's Inequality of Expectation} \\ &= \tilde{E}\left[|S_T|\right] \quad \text{by tower property} \\ &< +\infty \\ \tilde{\mathbb{E}}\left[\operatorname{Fut}_S(u,T) \mid \mathcal{F}_t\right] &= \tilde{\mathbb{E}}\left[\tilde{\mathbb{E}}[S_T \mid \mathcal{F}_u]\middle|\mathcal{F}_t\right] \quad \text{by definition of futures price} \\ &= \tilde{\mathbb{E}}\left[S_T \mid \mathcal{F}_t\right] \quad \text{by tower property} \\ &= \operatorname{Fut}_S(t,T) \end{split}$$

Therefore, futures price is a martingale under risk-neutral measure.

8.4 Forward-Futures Spread

$$\operatorname{For}_{S}(t,T) = \frac{S_{t}}{B(t,T)}, \operatorname{Fut}_{S}(t,T) = \tilde{\mathbb{E}}[S_{T} \mid \mathcal{F}_{t}]$$

1. Interest rate is constant r:

$$B(t,T) = e^{-r(T-t)} \Rightarrow \operatorname{For}_{S}(t,T) = e^{r(T-t)} S_{t}$$

$$\tilde{\mathbb{E}}[e^{-rT} S_{T} \mid \mathcal{F}_{t}] = e^{-rt} S_{t} \Rightarrow \tilde{\mathbb{E}}[S_{T} \mid \mathcal{F}_{t}] = e^{r(T-t)} S_{t}$$

$$\Rightarrow \operatorname{Fut}_{S}(t,T) = e^{r(T-t)} S_{t}$$

2. Interest rate R_t is deterministic, R(t):

Then $D_T = D(T) = e^{-\int_0^T R(u)du}$ is deterministic, thus

$$\tilde{\mathbb{E}}[D(T)S_T \mid \mathcal{F}_t] = D(T)\tilde{\mathbb{E}}[S_T \mid \mathcal{F}_t]
For_S(t,T) = Fut_S(t,T)
40$$

3. Interest rate R_t is stochastic: Note that the following does not hold in general

$$\tilde{\mathbb{E}}[D_T S_T \mid \mathcal{F}_t] \neq D_T \tilde{\mathbb{E}}[S_T \mid \mathcal{F}_t]
\tilde{\mathbb{E}}[D_T S_T \mid \mathcal{F}_t] \neq \tilde{\mathbb{E}}[D_T \mid \mathcal{F}_t] \tilde{\mathbb{E}}[S_T \mid \mathcal{F}_t]$$

Considering [0, T], we have

$$\begin{aligned} \operatorname{For}_S(0,T) &= \frac{S_0}{B(0,T)} = \frac{S_0}{\tilde{\mathbb{E}}[D_T]}, B[0,T] = \tilde{\mathbb{E}}[D_T], S_0 = \tilde{\mathbb{E}}[D_TS_T] \\ \operatorname{Fut}_S(0,T) &= \tilde{\mathbb{E}}[S_T] \\ \Rightarrow \operatorname{For}_S(0,T) - \operatorname{Fut}_S(0,T) &= \frac{\tilde{\mathbb{E}}[D_TS_T]}{\tilde{\mathbb{E}}[D_T]} - \tilde{\mathbb{E}}[S_T] \\ &= \frac{\tilde{\mathbb{E}}[D_TS_T] - \tilde{\mathbb{E}}[D_T]\tilde{\mathbb{E}}[S_T]}{\tilde{\mathbb{E}}[D_T]} \\ &= \frac{Cov(D_T,S_T)}{\tilde{\mathbb{E}}[D_T]} \end{aligned}$$

Conclusion:

- (1) If $Cov(D_T, S_T) \neq 0$, then $For_S(0, T) = Fut_S(0, T)$.
- (2) $\tilde{\text{Cov}}(D_T, S_T) < 0$
- (3) $\tilde{\text{Cov}}(D_T, S_T) > 0$

8.5 Dynamics of Futures Price

1. Suppose $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$ under risk-neutral measure.

$$\begin{split} \operatorname{Fut}_S(t,T) &= e^{r(T-t)} S_t = g(t,S_t) \\ g(t,x) &= e^{r(T-t)} x, \, g_t(t,x) = -rg(t,x), \, g_x(t,x) = e^{r(T-t)}, \, g_{xx} = 0 \\ d\operatorname{Fut}_S(t,T) &= dg(t,S_t) \\ &= g_t dt + g_x dS_t + \frac{1}{2} g_{xx} (dS_t)^2 \\ &= -rg(t,S_t) dt + e^{r(T-t)} dS_t \\ &= -re^{r(T-t)} S_t dt + e^{r(T-t)} (rS_t dt + \sigma S_t d\tilde{B}_t) \\ &= \sigma e^{r(T-t)} S_t d\tilde{B}_t \end{split}$$

$$\mathbb{E}[\int_0^t \sigma^2 e^{r(T-t)} S_t^2 dt] < +\infty$$

2. Replicating portfolio:

Suppose $dS_t = rS_t dt + \sigma S_t d\tilde{B}_T$ under risk-neutral measure

$$\begin{split} X_t &= X_t - \Delta S_t + \Delta_t S_t \\ dX_t &= r(X_t - \Delta_t S_t) dt + \Delta_t dS_t \\ &= r(X_t - \Delta_t S_t) dt + \Delta_t (rS_t dt + \sigma S_t d\tilde{B}_t) \\ &= rX_t dt + \Delta_t \sigma S_t d\tilde{B}_t \\ \Leftrightarrow d\text{Fut}_S(t,T) &= \sigma e^{r(T-t)} S_t d\tilde{B}_t \end{split}$$

$$\Rightarrow \begin{cases} \Delta_t = e^{r(T-t)}, & X_0 = 0\\ X_0 = 0 = -\Delta_0 S_0 + \Delta_0 S_0, & \Delta_0 = e^{rT} \end{cases}$$

$$\begin{split} \Delta_0 &= e^{rT} \to \Delta_t = e^{r(T-t)}, \quad \text{selling stock at rate} \ -r\Delta_t \\ &d\Delta_t = -r\Delta_t dt \\ &de^{-rt}X_t = e^{-rt}\Delta_t \sigma_t S_t d\tilde{B}_t \quad (\tilde{\mathbb{E}}[e^{-rt}X_t] = 0 \text{ since } X_0 = 0) \end{split}$$

3. Selling 1 futures contract at time 0, Δ_t shares of stock, constant interest r, $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$ under risk-neutral measure, we have

$$\begin{split} X_t &= X_t - \Delta_t S_t + \Delta_t S_t - \operatorname{Fut}_S(t,T) \\ dX_t &= r(X_t - \Delta_t S_t) dt + \Delta_t dS_t - d\operatorname{Fut}_S(t,T) \\ &= r(X_t - \Delta_t S_t) dt + \Delta_t (rS_t dt + \sigma S_t d\tilde{B}_t) - \sigma e^{r(T-t)} S_t d\tilde{B}_t \\ &= rX_t dt + (\Delta \sigma S_t - \sigma e^{r(T-t)} S_t) d\tilde{B}_t \end{split}$$

9 Lecture 9

9.1 Interest Rate Model

- 1. Interest rate is R_t , and discount factor is $D_t = e^{-\int_0^t R_u du}$. Under risk-neutral measure, $B(t,T) = \frac{1}{D_t} \tilde{\mathbb{E}}[D_T \mid \mathcal{F}_t]$.
- 2. Bond Price = Face Value $\times e^{-\text{Yield} \times \text{Time to Maturity}}$
- 3. Modeling of R_t short rate

$$dR_t = \beta(t, R_t)dt + \gamma(t, R_t)d\tilde{B}_t, \quad (*)$$

which is the dynamic of R_t under risk-neutral measure.

- 4. 1-factor model ⇔ 1-dim Brownian Motion.
- 5. We are interested in the risk-neutral measure that

$$B(t,T) = \frac{1}{D_t} \tilde{\mathbb{E}}[D_T \mid \mathcal{F}_t]$$
$$= e^{\int_0^t R_u du} \tilde{\mathbb{E}}[e^{-\int_0^T R_u du} \mid \mathcal{F}_t],$$

which can use 2 methods to solve:

- (1) Risk-neural pricing method: $\tilde{\mathbb{E}}[e^{-\int_0^T R_u du} \mid \mathcal{F}_t]$
 - (a) Need $\int_0^T R_t dt$ distributional information
 - (b) Very hard for general dynamic
- (2) PDE method: 4-step approach

9.2 PDE Approach

Step 1: Assume there exists a martingale:

$$g(t, R_t) = B(t, T) = \frac{1}{D_t} \tilde{\mathbb{E}}[D_T \mid \mathcal{F}_t], \text{ where } g(T, R_T) = 1.$$

Since $D_t B(t,T) = \tilde{\mathbb{E}}[D_T \mid \mathcal{F}_t]$, then $\{D_t B(t,T)\}_{0 \le t \le T}$ is a martingale under risk-neutral measure

Step 2: Apply Itô's formula for discounted option price:

$$dg(t, R_{t}) = g_{t}(t, R_{t})dt + g_{r}(t, R_{t})dR_{t} + \frac{1}{2}g_{rr}(t, R_{t})(dR_{t})^{2}$$

$$= \left(g_{t}(t, R_{t}) + g_{r}(t, R_{t})\beta(t, R_{t}) + \frac{1}{2}\gamma^{2}(t, R_{t})g_{rr}(t, R_{t})\right)dt + g_{r}(t, R_{t})\gamma(t, R_{t})d\tilde{B}_{t} \quad \text{from equation (*)}$$

$$dD_{t}g(t, R_{t}) = D_{t}dg(t, R_{t}) + g(t, R_{t})dD_{t} + dg(t, R_{t})dD_{t}$$

$$= D_{t}\left[\left(g_{t}(t, R_{t}) + g_{r}(t, R_{t})\beta(t, R_{t}) + \frac{1}{2}\gamma^{2}(t, R_{t})g_{rr}(t, R_{t})\right)dt + g_{r}(t, R_{t})\gamma(t, R_{t})d\tilde{B}_{t}\right] + g(t, R_{t})(-R_{t}D_{t}dt)$$

$$= D_{t}\left[\left(-R_{t}g(t, R_{t}) + g_{t}(t, R_{t}) + g_{r}(t, R_{t})\beta(t, R_{t}) + \frac{1}{2}\gamma^{2}(t, R_{t})g_{rr}(t, R_{t})\right)dt + D_{t}g_{r}(t, R_{t})\gamma(t, R_{t})d\tilde{B}_{t}\right]$$

Step 3: Set dt term equal to 0:

$$-R_t g(t, R_t) + g_t(t, R_t) + g_r(t, R_t) \beta(t, R_t) + \frac{1}{2} \gamma^2(t, R_t) g_{rr}(t, R_t) = 0$$

Step 4: Replace (t, R_t) to (t, r):

$$\begin{cases} -rg(t,r) + g_t(t,r) + g_r(t,r)\beta(t,r) + \frac{1}{2}\gamma^2(t,r)g_{rr}(t,r) = 0\\ g(T,r) = 1 \end{cases}$$

for all $0 \le t \le T$, for all possible values of r.

9.3 Mean-Reverting Processes

9.3.1 Ornstein-Uhlenbeck (OU) Process

$$dR_t = \kappa(\theta - R_t)dt + \sigma d\tilde{B}_t$$

where

- $\beta(t,r) = \kappa(\theta r)$
- $\gamma(t,r) = \sigma$
- $\kappa > 0$: mean-reverting speed
- θ : mean-reverting level

Let
$$g(t, R_t) = e^{\kappa t} R_t$$
, then

$$g_t(t, R_t) = \kappa e^{\kappa t} R_t = \kappa g(t, R_t)$$

$$q_r(t, R_t) = e^{\kappa t}$$

$$g_{rr}(t,R_t)=0$$

We calculate the dynamic of $e^{\kappa t}R_t$:

$$de^{\kappa t}R_t = dg(t, R_t) = g_t(t, R_t)dt + g_r(t, R_t)dR_t + \frac{1}{2}g_{rr}(t, R_t)(dR_t)^2$$

$$= \kappa e^{\kappa t}R_t dt + e^{\kappa t}(\kappa(\theta - R_t)dt + \sigma d\tilde{B}_t)$$

$$= \kappa \theta e^{\kappa t} dt + \sigma e^{\kappa t} d\tilde{B}_t$$

$$\Rightarrow e^{\kappa t}R_t - R_0 = \int_0^t \kappa \theta e^{\kappa u} du + \sigma \int_0^t e^{\kappa u} d\tilde{B}_u$$

The expectation R_t is:

$$\begin{split} &\tilde{\mathbb{E}}[e^{\kappa t}R_t] = R_0 + \int_0^t \kappa \theta e^{\kappa u} du \quad \text{since } \sigma^2 \int_0^t e^{2\kappa u} du < +\infty \\ &\Rightarrow e^{\kappa t} \tilde{\mathbb{E}}[R_t] = R_0 + \theta(e^{\kappa t} - 1) \\ &\Rightarrow \tilde{\mathbb{E}}[R_t] = e^{-\kappa t} R_0 + \theta(1 - e^{-\kappa t}) \stackrel{t \to +\infty}{\longrightarrow} \theta \quad \text{which is the long-term mean} \end{split}$$

The variance of R_t is

$$\begin{split} \widetilde{\mathrm{Var}}(e^{\kappa t}R_t) &= \tilde{\mathbb{E}}\left[\left(\sigma\int_0^t e^{\kappa u}d\tilde{B}_t\right)^2\right] \\ &= \sigma^2\tilde{\mathbb{E}}\left[\int_0^t e^{2\kappa u}du\right] \quad \text{since } \widetilde{\mathrm{Var}}(\tilde{B}_t) = t \\ &= \frac{\sigma^2}{2\kappa}(e^{2\kappa t}-1) \\ \Rightarrow \widetilde{\mathrm{Var}}(R_t) &= e^{-2\kappa t}\frac{\sigma^2}{2\kappa}(e^{2\kappa t}-1) \\ &= \frac{\sigma^2}{2\kappa}(1-e^{-2\kappa t}) \end{split}$$

Therefore,
$$R_t \sim \left(e^{-\kappa t}R_0 + \theta(1 - e^{-\kappa t}), \frac{\sigma^2}{2\kappa}(1 - e^{-2\kappa t})\right)$$
.

9.3.2 Cox-Ingersoll-Ross (CIR) Process

$$dR_t = \kappa(\theta - R_t)dt + \sigma\sqrt{R_t}d\tilde{B}_t$$

- $\kappa > 0, \theta > 0, R_t \ge 0$
- Continuous process, $R_t = 0 \Rightarrow$ volatility = 0, hit 0 but never goes down below 0
- Not Gaussian
- $R_t \ge 0$
- Still mean reverting

9.4 Diffusion Process

$$dX_t = \mu_t dt + \sigma_t dB_t$$
, $X_0 = X$, B_t is \mathbb{P} -martingale

- 1. Martingale for Diffusion Process: $\{X_t\}_{0 \le t \le T}$ is a \mathbb{P} -martingale \iff
 - (1) $\mu_t = 0$ for all $t \ge 0$
 - $(2) \ \mathbb{E}\left[\int_0^T \sigma_u^2 du\right] < +\infty$
- 2. Remark: Suppose $\mathbb{E}\left[\int_0^T \sigma_u^2 du\right] < +\infty$

Then
$$Y_t = \int_0^t \sigma_u dB_u, Y_0 = 0 \iff dY_t = \sigma_t dB_t, Y_0 = 0.$$

By definition of martingale for diffusion process, $\{Y_t\}_{0 \le t \le T}$ is a \mathbb{P} -martingale.

Therefore, $\mathbb{E}[Y_t \mid \mathcal{F}_s] = Y_s, \forall 0 \le t \le T.$

In particular, $Y_0 = 0 = \mathbb{E}[Y_t \mid \mathcal{F}_0] = \mathbb{E}[Y_t].$

$$\begin{split} dX_t &= \mu_t dt + \sigma_t dB_t = \mu_t dt + dY_t \\ \Rightarrow X_t &= X_0 + \int_0^t \mu_u du + Y_t \\ \Rightarrow \mathbb{E}[X_t] &= X_0 + \mathbb{E}\left[\int_0^t \mu_u du\right] \quad \text{since } \mathbb{E}[Y_t] = 0 \end{split}$$

9.5 Famous Interest Rate Models

- 1. Ho-Lee Model: $dR_t = \mu dt + \sigma d\tilde{B}_t$
- 2. Vasicek Model: $dR_t = \kappa(\theta R_t)dt + \sigma d\tilde{B}_t$
- 3. Hull-White Model: $dR_t = \kappa(\theta(t) R_t)dt + \sigma d\tilde{B}_t$
- 4. Cox-Ingersoll-Ross (CIR) Model: $dR_t = \kappa(\theta-R_t)dt + \sigma\sqrt{R_t}d\tilde{B}_t$

All leads to affine (linear) yield curve: $B(t,T)=e^{-\mathrm{yield}\cdot T}\Rightarrow \mathrm{yield}=-\frac{\partial}{\partial T}\log B(t,T)$

Example 9.5.1 Hull-White Model

Recall

$$\begin{cases} -rg(t,r) + g_t(t,r) + \beta(t,r)g_r(t,r) + \frac{1}{2}\gamma^2(t,r)g_{rr}(t,r) = 0\\ g(T,r) = 1, \forall r \end{cases}$$

For a Hull-White model, $\beta(t,r) = \kappa(\theta(t) - r)$, $\gamma(t,r) = \sigma$. Therefore the equations for the Hull-White model are

$$\begin{cases} -rg(t,r) + g_t(t,r) + \kappa(\theta(t) - r)g_r(t,r) + \frac{1}{2}\sigma^2 g_{rr}(t,r) = 0\\ g(T,r) = 1, \forall r \end{cases}$$

Guess $g(t,r) = e^{-C(t;T)r - A(t;T)}$ since the yield curve is affine. Then we have

$$g_t(t,r) = \left(-\frac{\partial C(t;T)}{\partial t}r - \frac{\partial A(t;T)}{\partial t}\right)e^{-C(t;T)r - A(t;T)} = -\left(\frac{\partial C(t;T)}{\partial t}r + \frac{\partial A(t;T)}{\partial t}\right)g(t,r)$$

$$g_r(t,r) = -C(t;T)e^{-C(t;T)r - A(t;T)} = -C(t;T)g(t,r)$$

$$g_{rr}(t,r) = C^2(t;T)g(t,r)$$

To solve the PDE equations, we first calculate

$$-rg(t,r)+g_t(t,r)+\kappa(\theta(t)-r)g_r(t,r)+\frac{1}{2}\sigma^2g_{rr}(t,r)=0$$

$$-rg(t,r)-\left(\frac{\partial C(t;T)}{\partial t}r+\frac{\partial A(t;T)}{\partial t}\right)g(t,r)-\kappa(\theta(t)-r)C(t;T)g(t,r)+\frac{1}{2}\sigma^2C^2(t;T)g(t,r)=0$$

$$-r-\frac{\partial C(t;T)}{\partial t}r-\frac{\partial A(t;T)}{\partial t}-\kappa(\theta(t)-r)C(t;T)+\frac{1}{2}\sigma^2C^2(t;T)=0 \quad \text{since } g(t,r)>0$$

$$-r\left(1+\frac{\partial C(t;T)}{\partial t}-\kappa C(t;T)\right)-\frac{\partial A(t;T)}{\partial t}-\kappa\theta(t)C(t;T)+\frac{1}{2}\sigma^2C^2(t;T)=0$$

$$\Rightarrow\begin{cases} 1+\frac{\partial C(t;T)}{\partial t}-\kappa C(t;T)=0\\ -\frac{\partial A(t;T)}{\partial t}-\kappa\theta(t)C(t;T)+\frac{1}{2}\sigma^2C^2(t;T)=0 \end{cases}$$
 since PDE holds for all r

From $g(T,r)=1, \forall r$, we know C(T;T)=0 and A(T;T)=0. Then we can transfer PDE to ODE:

$$\begin{cases} 1 + \frac{\partial C(T;T)}{\partial t} - \kappa C(t;T) = 0 \\ C(T;T) = 0 \end{cases} \begin{cases} -\frac{\partial A(t;T)}{\partial t} - \kappa \theta(t)C(t;T) + \frac{1}{2}\sigma^2 C^2(t;T) = 0 \\ A(T;T) = 0 \end{cases}$$

For C(t;T), we have

$$\begin{split} \frac{\partial}{\partial t}C(t;T) &= \kappa C(t;T) - 1 \\ \Rightarrow \frac{\partial}{\partial t}e^{-\kappa t}C(t;T) &= e^{-\kappa t}\frac{\partial}{\partial t}C(t;T) - \kappa e^{-\kappa t}C(t;T) \\ &= e^{-\kappa t}\kappa C(t;T) - e^{-\kappa t} - \kappa e^{-\kappa t}C(t;T) \\ &= e^{-\kappa t} \end{split}$$

$$\Rightarrow \int_t^T \frac{\partial}{\partial t}e^{-\kappa t}C(t;T) &= \int_t^T e^{-\kappa t} \\ e^{-\kappa T}C(T;T) - e^{-\kappa t}C(t;T) &= -\int_t^T e^{-\kappa u}du \\ -e^{-\kappa t}C(t;T) &= \frac{1}{\kappa}(e^{-\kappa T} - e^{-\kappa t}) \\ \Rightarrow C(t;T) &= \frac{1}{\kappa}(1 - e^{-\kappa(T-t)}), \quad \frac{\partial}{\partial T}C(t;T) = e^{-\kappa(T-t)} \end{split}$$

For A(t;T), we have

$$\begin{split} -\frac{\partial}{\partial t}A(t;T) &= \kappa\theta(t)C(t;T) - \frac{1}{2}\sigma^2C^2(t;T) \\ A(t;T) - A(T;T) &= \int_t^T \kappa\theta(u)C(t;T)du - \frac{1}{2}\sigma^2\int_t^T C^2(u;T)du \quad \text{by taking integrals on both sides} \\ A(t;T) &= \int_t^T \kappa\theta(u)C(t;T)du - \frac{1}{2}\sigma^2\int_t^T C^2(u;T)du \quad \text{since } A(T;T) = 0 \\ \frac{\partial}{\partial T}A(t;T) &= \kappa\theta(u)C(T;T) - \int_t^T \kappa\theta(u)\frac{\partial}{\partial T}C(u;T)du - \frac{1}{2}\sigma^2C(T;T) + \frac{1}{2}\sigma^2\int_t^T \frac{\partial}{\partial T}C^2(u;T)du \\ &= -\int_t^T \kappa\theta(u)\frac{\partial}{\partial T}C(u;T)du + \frac{1}{2}\sigma^2\int_t^T 2C(u;T)\frac{\partial}{\partial T}C(u;T)du \\ &= -\int_t^T \kappa\theta(u)e^{-\kappa(T-u)}du + \sigma^2\int_t^T C(u;T)e^{-\kappa(T-u)}du \\ &= -\int_t^T \kappa\theta(u)e^{-\kappa(T-u)}du \\ &= -\int_t^T \kappa\theta(u)e^{-\kappa(T-u)}du + \sigma^2\int_t^T C(u;T)e^{-\kappa(T-u)}du \\ &= -\int_t^T \kappa\theta(u)e^{-\kappa(T-u)}du \\ &= -\int_t^T \kappa\theta(u$$

With both C(t;T) and A(t;T), we are able to give the bond price as

$$B(t,T) = g(t,R_t) = e^{-C(t;T)R_t - A(t;T)}$$

$$B(0,T) = e^{-C(0;T)R_0 - A(0;T)}$$

$$yield = -\frac{\partial}{\partial T} \log B(0,T)$$

$$= \frac{\partial}{\partial T} (C(0;T)R_0 + A(0;T))$$

$$= \underbrace{e^{-\kappa T}}_{\text{slope}} R_0 + \underbrace{\frac{\partial}{\partial T} A(0;T)}_{\text{intercent}}$$

9.6 Derivatives on Interest Rate

9.6.1 Forward Interest Rate

Consider a contract that pays R_T at time T, P_t for price of the contract at time t, $0 \le t \le T$.

$$\begin{split} D_t P_t &= \tilde{\mathbb{E}}[D_T P_T \mid \mathcal{F}_t] \\ &= \tilde{\mathbb{E}}[D_T R_T \mid \mathcal{F}_t] \quad \text{since } P_T = R_T \text{ at maturity} \\ P_t &= \tilde{\mathbb{E}}\left[R_T e^{-\int_t^T R_u du} \middle| \mathcal{F}_t\right] \quad \text{by dividing } D_t \text{ on both sides and by definition of } B(t,T) \\ &= -\frac{\partial}{\partial T} \tilde{\mathbb{E}}\left[e^{-\int_t^T R_u du} \middle| \mathcal{F}_t\right] \quad \text{since } \frac{\partial}{\partial T} e^{-\int_t^T R_u du} = -R_T e^{-\int_t^T R_u du} \\ &= -\frac{\partial}{\partial T} B(t,T) \quad \text{by definition of } B(t,T) \end{split}$$

9.6.2 Forward Interest Rate

$$\operatorname{Fut}_R(t,T) = \tilde{\mathbb{E}}[R_T \mid \mathcal{F}_t]$$

Under Hull-White model, we have

$$dR_{t} = \kappa(\theta(t) - R_{t})dt + \sigma d\tilde{B}_{t}$$

$$\Rightarrow R_{T} = R_{0} + \int_{0}^{T} \kappa(\theta(u) - R_{u})du + \sigma \tilde{B}_{T}$$

$$de^{\kappa t}R_{t} = e^{\kappa t}dR_{t} + \kappa e^{\kappa t}R_{t}dt$$

$$= e^{\kappa t} \left[\kappa(\theta(t) - R_{t})dt + \sigma d\tilde{B}_{t}\right] + \kappa e^{\kappa t}R_{t}dt$$

$$= e^{\kappa t}\kappa\theta(t)dt + e^{\kappa t}\sigma d\tilde{B}_{t}$$

$$\Rightarrow e^{\kappa T}R_{T} = R_{0} + \int_{0}^{T} e^{\kappa t}\kappa\theta(t)dt + \sigma \int_{0}^{T} e^{\kappa t}d\tilde{B}_{t}$$

Since $\mathbb{E}[\int_0^T \sigma^2 e^{2\kappa t} dt] \leq \frac{\sigma^2}{2\kappa} e^{2\kappa T} < +\infty$, we have

$$\tilde{\mathbb{E}}[e^{\kappa T}R_T \mid \mathcal{F}_t] = R_t + \int_t^T e^{\kappa t} \kappa \theta(t) dt$$

$$\operatorname{Fut}_R(t, T) = \tilde{\mathbb{E}}[R_T \mid \mathcal{F}_t] = e^{-\kappa T} R_t + \kappa \int_t^T e^{-\kappa (T - t)} \theta(t) dt$$

10 Lecture 10

10.1 Girsanov's Change of Measure (1-dim)

Theorem 10.1.1. (Girsanov's Change of Measure (1-dim)).

On $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \mathbb{P})$, let $\{B_t\}_{t\geq 0}$ be a Brownian Motion and $\{\Theta_t\}_{t\geq 0}$ be an adaptive process.

Focusing [0,T] and assuming $\mathbb{E}[e^{\frac{1}{2}\int_0^T \Theta_t^2 dt}] < +\infty$, define

$$\begin{split} \tilde{B}_t &= B_t + \int_0^t \Theta_u \, du \Longleftrightarrow d\tilde{B}_t = dB_t + \Theta_t dt \\ Z_t &= e^{-\int_0^t \Theta_u \, dB_u - \frac{1}{2} \int_0^t \Theta_u^2 \, du} \Longleftrightarrow dZ_t = -Z_t \Theta_t dB_t \end{split}$$

Then for $\tilde{\mathbb{P}}$ defined by $\tilde{\mathbb{E}}[X] = \mathbb{E}[XZ_T]$, for all random variables X,

we have that $\{\tilde{B}_t\}_{t\geq 0}$ is a Brownian Motion under $\tilde{\mathbb{P}}$ on On $(\Omega, \mathcal{F}, \{\mathcal{F}\}_{t\geq 0}, \tilde{\mathbb{P}})$.

Example 10.1.2. Show that $dZ_t = -Z_t\Theta_t dB_t$.

Let $g(t) = -\int_0^t \Theta_u dB_u - \frac{1}{2} \int_0^t \Theta_u^2$. Then we have $dg(t) = -\frac{1}{2} \Theta_t^2 dt - \Theta_t dB_t$.

Let $f(t,g(t)) = e^{g(t)} = Z_t$. Then we have $f_g(t,g(t)) = e^{g(t)} = Z_t$ and $f_{gg}(t,g(t)) = e^{g(t)} = Z_t$.

Applying Itô's formula, we have

$$dZ_t = df(t, g(t)) = \left[f_g(t, g(t))\mu_t + \frac{1}{2} f_{gg}(t, g(t))\sigma_t^2 \right] dt + f_g(t, g(t))\sigma_t dB_t$$
$$= \left(-Z_t \frac{1}{2} \Theta_t^2 + \frac{1}{2} Z_t (-\Theta_t)^2 \right) dt - Z_t \Theta_t dB_t$$
$$= -Z_t \Theta_t dB_t$$

Exercise 10.1.3 In a Black-Scholes market model, let $S = \{S_t\}_{t\geq 0}$ denote the price process satisfying

$$dS_t = S_t(\mu + \sigma dW_t),$$

under the physical measure \mathbb{P} with interest rate r=0. Using the FTAP and Girsanov's Theorem, determine the Radon-Nikodym derivative from \mathbb{P} to a risk-neutral measure $\tilde{\mathbb{P}}$. Further, write down the stock dynamics under the risk-neutral measure $\tilde{\mathbb{P}}$ and determine the relationship between the Brownian motions under measure \mathbb{P} and $\tilde{\mathbb{P}}$.

Solution.

Under physical measure \mathbb{P} , $dS_t = \mu S_t dt + \sigma S_t dB_t$, where $\{B_t\}_{t\geq 0}$ is a Brownian Motion under \mathbb{P} (usually $\mu > r$).

Under risk-neutral measure $\tilde{\mathbb{P}}$, $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$, where $\{\tilde{B}_t\}_{t>}$ is a Brownian Motion under $\tilde{\mathbb{P}}$.

Equating $dS_t = \mu S_t dt + \sigma S_t dB_t$ and $dS_t = rS_t dt + \sigma S_t d\tilde{B}_t$, we have

$$\mu S_t dt + \sigma S_t dB_t = r S_t dt + \sigma S_t d\tilde{B}_t$$
$$d\tilde{B}_t - dB_t = \frac{\mu - r}{\sigma} dt$$
$$\Rightarrow d\tilde{B}_t = dB_t + \frac{\mu - r}{\sigma} dt$$

where $\frac{\mu-r}{\sigma} > 0$ and $\sigma > 0$.

By Girsanov's Change of Measure, define $\Theta_t = \frac{\mu - r}{\sigma}$, then the Radon-Nikodym derivative from physical measure \mathbb{P} to risk-neutral measure \mathbb{P} is

$$\begin{split} Z_t &= \frac{d\tilde{\mathbb{P}}}{d\mathbb{P}} = e^{-\int_0^t \frac{\mu - r}{\sigma} dB_u - \frac{1}{2} \int_0^t \left(\frac{\mu - r}{\sigma}\right)^2 du} \\ &= e^{-\frac{\mu - r}{\sigma} B_t - \frac{1}{2} \left(\frac{\mu - r}{\sigma}\right)^2 t} \quad \text{since } \frac{\mu - r}{\sigma} \text{ is constant} \end{split}$$

Note that when $r=0, Z_t=\frac{d\tilde{\mathbb{P}}}{d\mathbb{P}}=e^{-\frac{\mu}{\sigma}B_t-\frac{1}{2}\left(\frac{\mu}{\sigma}\right)^2t}.$

Aside:

Note that MGF is $\mathbb{E}[e^{\lambda Z}] = e^{\lambda \mu + \frac{1}{2}\lambda^2 \sigma^2}$, where $Z \sim N(\mu, \sigma^2)$.

The expectations of strike price and final stock price by change of measure (from physical measure \mathbb{P} to risk-neutral measure $\tilde{\mathbb{P}}$) are

$$\begin{split} \tilde{\mathbb{E}}[K] &= \mathbb{E}[KZ_T] = \mathbb{E}\left[Ke^{-\frac{\mu-r}{\sigma}B_T - \frac{1}{2}\left(\frac{\mu-r}{\sigma}\right)^2T}\right] = K \quad \text{since } K \text{ is constant and } \mathbb{E}[Z_T] = 1 \\ \tilde{\mathbb{E}}[S_T] &= \mathbb{E}[S_TZ_T] = \mathbb{E}\left[S_0e^{\left(\mu - \frac{1}{2}\sigma^2\right)T + \sigma B_T}e^{-\frac{\mu-r}{\sigma}B_T - \frac{1}{2}\left(\frac{\mu-r}{\sigma}\right)^2T}\right] \quad \text{where } \mathbb{E}[S_T] = \mathbb{E}\left[S_0e^{\left(\mu - \frac{1}{2}\sigma^2\right)T + \sigma B_T}\right] = e^{\mu T} \text{ by MGF} \\ &= S_0\mathbb{E}\left[e^{\left(\sigma - \frac{\mu-r}{\sigma}\right)B_T - \frac{1}{2}\left(\sigma^2 + \left(\frac{\mu-r}{\sigma}\right)^2 - 2(\mu-r)\right)T - (\mu-r)T + \mu T}\right] \\ &= S_0\mathbb{E}\left[e^{\left(\sigma - \frac{\mu-r}{\sigma}\right)B_T - \frac{1}{2}\left(\sigma^2 + \left(\frac{\mu-r}{\sigma}\right)^2 - 2(\mu-r)\right)T + rT}\right] \\ &= S_0e^{rT}\mathbb{E}\left[e^{\left(\sigma - \frac{\mu-r}{\sigma}\right)B_T - \frac{1}{2}\left(\sigma^2 + \left(\frac{\mu-r}{\sigma}\right)^2 - 2(\mu-r)\right)T + rT}\right] \\ &= S_0e^{rT}\mathbb{E}\left[e^{\left(\sigma - \frac{\mu-r}{\sigma}\right)B_T - \frac{1}{2}\left(\sigma^2 - \frac{\mu-r}{\sigma}\right)^2T}\right] \\ &= S_0e^{rT}\mathbb{E}\left[e^{\left(\sigma - \frac{\mu-r}{\sigma}\right)B_T}\right]e^{-\frac{1}{2}\left(\sigma - \frac{\mu-r}{\sigma}\right)^2T} \\ &= S_0e^{rT}e^{\frac{1}{2}\left(\sigma - \frac{\mu-r}{\sigma}\right)^2T}e^{-\frac{1}{2}\left(\sigma - \frac{\mu-r}{\sigma}\right)^2T} \\ &= S_0e^{rT} \quad \text{which aligns with the definition of risk-neutral measure} \end{split}$$

10.2 Black-Scholes Equation for European Call Option by Change of Measure

$$\operatorname{Recall} \tilde{\mathbb{E}} \left[e^{-rT} (S_T - K)^+ \right] = \tilde{\mathbb{E}} \left[e^{-rT} S_T \mathbb{1}_{[S_T \geq K]} \right] - e^{-rT} K \tilde{\mathbb{E}} \left[\mathbb{1}_{[S_T \geq K]} \right]$$

(1) For
$$\tilde{\mathbb{E}}\left[\mathbb{1}_{[S_T\geq K]}\right]=\tilde{\mathbb{P}}[S_T\geq K]$$
, since $S_T=S_0e^{\left(r-\frac{1}{2}\sigma^2\right)T+\sigma \tilde{B}_T}$, we calculate

$$S_{T} \ge K$$

$$S_{0}e^{\left(r-\frac{1}{2}\sigma^{2}\right)T+\sigma\tilde{B}_{T}} \ge K$$

$$\left(r-\frac{1}{2}\sigma^{2}\right)T+\sigma\tilde{B}_{T} \ge -\ln\frac{S_{0}}{K}$$

$$\sigma\tilde{B}_{T} \ge -\left(\ln\frac{S_{0}}{K}+\left(r-\frac{1}{2}\sigma^{2}\right)T\right)$$

$$-\frac{\tilde{B}_{T}}{\sqrt{T}} \le \frac{1}{\sigma\sqrt{T}}\left(\ln\frac{S_{0}}{K}+\left(r-\frac{1}{2}\sigma^{2}\right)T\right)$$

where $-\frac{\tilde{B}_T}{\sqrt{T}}$ is a standard normal variable. Therefore, $\tilde{\mathbb{P}}[S_T \geq K] = N(d_-(T, S_0))$.

(2) For $\tilde{\mathbb{E}}\left[e^{-rT}S_T\mathbbm{1}_{[S_T\geq K]}\right]=\tilde{\mathbb{E}}\left[e^{-rT}S_0e^{\left(r-\frac{1}{2}\sigma^2\right)T+\sigma\tilde{B}_T}\mathbbm{1}_{[S_T\geq K]}\right]=\tilde{\mathbb{E}}\left[S_0e^{\sigma\tilde{B}_T-\frac{1}{2}\sigma T}\mathbbm{1}_{[S_T\geq K]}\right]$, we define $\Theta_t=-\sigma$ since $e^{\sigma\tilde{B}_T-\frac{1}{2}\sigma T}$ looks like a Radon-Nikodym derivative.

By Girsanov's Change of Measure, the Radon-Nikodym derivative from risk-neutral measure $\tilde{\mathbb{P}}$ to the measure \mathbb{S} and the Brownian Motion under \mathbb{S} -measure are

$$\begin{split} Z_t &= \frac{d\mathbb{S}}{d\tilde{\mathbb{P}}} = e^{-\int_0^t \Theta_u \, d\tilde{B}_u - \frac{1}{2} \int_0^t \Theta_u^2 \, du} = e^{\sigma \tilde{B}_t - \frac{1}{2} \sigma^2 t} \\ B_t^{\mathbb{S}} &= \tilde{B}_t + \int_0^t \Theta_u du = \tilde{B}_t - \sigma t \end{split}$$

since σ is constant. Then we can write

$$\begin{split} &\tilde{\mathbb{E}}\left[e^{-rT}S_{T}\mathbb{1}_{[S_{T}\geq K]}\right] = \tilde{\mathbb{E}}\left[S_{0}e^{\sigma\tilde{B}_{T} - \frac{1}{2}\sigma T}\mathbb{1}_{[S_{T}\geq K]}\right] = \mathbb{E}^{\mathbb{S}}\left[S_{0}\mathbb{1}_{[S_{T}\geq K]}\right] = S_{0}\mathbb{P}^{\mathbb{S}}[S_{T}\geq K] \quad \text{and} \quad S_{T} = S_{0}e^{\left(r - \frac{1}{2}\sigma^{2}\right)T + \sigma\tilde{B}_{T}} = S_{0}e^{\left(r - \frac{1}{2}\sigma^{2}\right)T + \sigma(B_{T}^{\mathbb{S}} + \sigma T)} = S_{0}e^{\left(r + \frac{1}{2}\sigma^{2}\right)T + \sigma B_{T}^{\mathbb{S}}} \end{split}$$

by change of measure. Then we have

$$S_T \ge K$$

$$S_0 e^{\left(r + \frac{1}{2}\sigma^2\right)T + \sigma B_T^{\mathbb{S}}} \ge K$$

$$\sigma B_T^{\mathbb{S}} \ge -\left(\ln\frac{S_0}{K} + \left(r + \frac{1}{2}\sigma^2\right)T\right)$$

$$-\frac{B_T^{\mathbb{S}}}{\sqrt{T}} \le \frac{1}{\sigma\sqrt{T}}\left(\ln\frac{S_0}{K} + \left(r + \frac{1}{2}\sigma^2\right)T\right)$$

Therefore, $\mathbb{P}^{\mathbb{S}}[S_T \geq K] = N(d_+(T, S_0)).$

Therefore, we have arrived at $\tilde{\mathbb{E}}\left[e^{-rT}\left(S_T-K\right)^+\right]=S_0N(d_+(T,S_0))-e^{-rT}KN(d_-(T,S_0))$, which is the Black-Scholes equation for European call.

Example 10.2.1 Use the change of measure $\Theta_t = -\sigma$ to solve $\mathbb{E}\left[e^{\sigma B_T - \frac{1}{2}\sigma^2 T}B_T^4\right]$.

Solution. Since $\Theta_t = -\sigma$, we have $Z_t = e^{\sigma B_T - \frac{1}{2}\sigma^T}$ and $\hat{B}_t = B_t - \sigma t \Rightarrow B_t = \hat{B}_t + \sigma t$. Then we can calculate

$$B_T^4 = (\hat{B}_t + \sigma t)^4$$

$$= {4 \choose 0} (\hat{B}_T)^4 (\sigma T)^0 + {4 \choose 1} (\hat{B}_T)^3 (\sigma T)^1 + {4 \choose 2} (\hat{B}_T)^2 (\sigma T)^2 + {4 \choose 3} (\hat{B}_T)^1 (\sigma T)^3 + {4 \choose 4} (\hat{B}_T)^0 (\sigma T)^4$$

$$= \hat{B}_T^4 + 4\sigma T \hat{B}_T^3 + 6\sigma^2 T^2 \hat{B}_T^2 + 4\sigma^3 T^3 \hat{B}_T + \sigma^4 T^4$$

$$\begin{split} \mathbb{E}\left[e^{\sigma B_T - \frac{1}{2}\sigma^2 T}B_T^4\right] &= \hat{\mathbb{E}}\left[B_T^4\right] \quad \text{by change of measure} \\ &= \hat{\mathbb{E}}\left[\hat{B}_T^4 + 4\sigma T\hat{B}_T^3 + 6\sigma^2 T^2\hat{B}_T^2 + 4\sigma^3 T^3\hat{B}_T + \sigma^4 T^4\right] \\ &= \hat{\mathbb{E}}\left[\hat{B}_T^4\right] + 6\sigma^2 T^2\hat{\mathbb{E}}\left[\hat{B}_T^2\right] + \sigma^4 T^4 \quad \text{since } \hat{\mathbb{E}}\left[\hat{B}_T\right] = 0 \text{ and } \hat{\mathbb{E}}\left[\hat{B}_T^3\right] = 0 \\ &= 3T^2 + 6\sigma^2 T^3 + \sigma^4 T^4 \quad \text{since } \hat{\mathbb{E}}\left[\hat{B}_T^2\right] = T \text{ and } \hat{\mathbb{E}}\left[\hat{B}_T^4\right] = 3T^2 \end{split}$$

Remark 10.2.2 Moments of Brownian Motion B_t :

- (1) Mean: $\mathbb{E}[B_t] = 0$.
- (2) Second Moment: $\mathbb{E}[B_t^2] = t$.
- (3) Third Moment: $\mathbb{E}[B_t^3] = 0$.
- (4) Fourth Moment: $\mathbb{E}[B_t^4] = 3t^2$

10.3 Girsanov's Change of Measure (d-dim) in Currencies

CAD\$/USD\$

EUR€/USD\$

Consider the Brownian Motions

- $(\tilde{B}_t^{(1)}, \tilde{B}_t^{(2)}, \dots, \tilde{B}_t^{(d)})$ under risk-neutral measure $\tilde{\mathbb{P}}$ in the American market (\$), and
- $(\hat{B}_t^{(1)}, \hat{B}_t^{(2)}, \dots, \hat{B}_t^{(d)})$ under risk-neutral measure $\tilde{\mathbb{P}}$ in the European market (\mathfrak{C}) .

Changing from \$ to €, we have

$$d\hat{B}_{t} = \begin{pmatrix} d\hat{B}_{t}^{(1)} \\ \vdots \\ d\hat{B}_{t}^{(d)} \end{pmatrix}, \quad d\hat{B}_{t} = \begin{pmatrix} \hat{B}_{t}^{(1)} \\ \vdots \\ \hat{B}_{t}^{(d)} \end{pmatrix}, \quad \Theta_{t} = \begin{pmatrix} \Theta_{t}^{(1)} \\ \vdots \\ \Theta_{t}^{(d)} \end{pmatrix}$$
$$d\hat{B}_{t} = d\tilde{B}_{t} + \Theta_{t}dt$$
$$Z_{t} = e^{-\int_{0}^{t} \Theta_{u}^{T} dB_{u} - \frac{1}{2} \int_{0}^{t} \|\Theta_{u}\|^{2} du}$$

Let $\{S_t\}_{t\geq 0}$ and $\{N_t\}_{t\geq 0}$ be the prices of two assets determined in the same currency,

$$dS_t = R_t S_t dt + \sigma_t^{(1)} S_t d\tilde{B}_t^{(1)} + \sigma_t^{(2)} S_t d\tilde{B}_t^{(2)} + \dots + \sigma_t^{(d)} S_t d\tilde{B}_t^{(d)} , \text{ and } dN_t = R_t N_t dt + \nu_t^{(1)} N_t d\tilde{B}_t^{(1)} + \nu_t^{(2)} N_t d\tilde{B}_t^{(2)} + \dots + \nu_t^{(d)} N_t d\tilde{B}_t^{(d)} ,$$

where the discount factor is $D_t = e^{-\int_0^t R_u du}$. For $\sigma_t = (\sigma_t^{(1)}, \dots, \sigma_t^{(d)})$ and $\nu_t = (\nu_t^{(1)}, \dots, \nu_t^{(d)})$, we define

$$\sigma_t d\tilde{B}_t = \sum_{i=1}^d \sigma_t^{(i)} d\tilde{B}_t \quad \text{and} \quad \nu_t d\tilde{B}_t = \sum_{i=1}^d \nu_t^{(i)} d\tilde{B}_t$$
$$dD_t S_t = D_t S_t \sigma_t d\tilde{B}_t \quad \text{and} \quad dD_t N_t = D_t N_t \nu_t d\tilde{B}_t$$

Here, we take N_t as the numéraire.

To verify, we write $D_t S_t = D_0 S_0 e^{\int_0^t \sigma_u d\tilde{B}_u - \frac{1}{2} \int_0^t \sum_{i=1}^d (\sigma_u^{(i)})^2 du} = D_0 S_0 e^{\sum_{i=1}^d \int_0^t \sigma_u^{(i)} d\tilde{B}_u^{(i)} - \frac{1}{2} \int_0^t \sum_{i=1}^d (\sigma_u^{(i)})^2 du}$.

By letting $f(t) = \frac{1}{2} \int_0^t \sum_{i=1}^d (\sigma_u^{(i)})^2 du$ and $M_t = \sum_{i=1}^d \int_0^t \sigma_u^{(i)} d\tilde{B}_u^{(i)} - f(t)$, we calculate

$$df(t) = \frac{1}{2} \sum_{i=1}^{d} (\sigma_t^{(i)})^2 dt$$

$$dM_t = \sum_{i=1}^{d} \sigma_t^{(i)} d\tilde{B}_t^{(i)} - \frac{1}{2} \sum_{i=1}^{d} (\sigma_t^{(i)})^2 dt$$

$$dD_t S_t = D_t S_t \left(dM_t + \frac{1}{2} d[M]_t \right)$$

$$= D_t S_t \left(\sum_{i=1}^{d} \sigma_t^{(i)} d\tilde{B}_t^{(i)} - \frac{1}{2} \sum_{i=1}^{d} (\sigma_t^{(i)})^2 dt + \frac{1}{2} \sum_{i=1}^{d} (\sigma_t^{(i)})^2 dt \right)$$

$$= D_t S_t \sum_{i=1}^{d} \sigma_t^{(i)} d\tilde{B}_t^{(i)}$$

To continue the previous steps, we have

$$\begin{cases} D_t S_t = D_0 S_0 e^{\sum_{i=1}^d \int_0^t \sigma_u^{(i)} d\tilde{B}_u^{(i)} - \frac{1}{2} \int_0^t \sum_{i=1}^d (\sigma_u^{(i)})^2 du} \\ D_t S_t = D_0 S_0 e^{\sum_{i=1}^d \int_0^t \nu_u^{(i)} d\tilde{B}_u^{(i)} - \frac{1}{2} \int_0^t \sum_{i=1}^d (\nu_u^{(i)})^2 du} \end{cases}$$

$$\begin{split} \frac{D_t S_t}{D_t N_t} &= \frac{S_t}{N_t} = \frac{S_0}{N_0} e^{\sum_{i=1}^d \int_0^t (\sigma_u^{(i)} - \nu_u^{(i)}) d\tilde{B}_u^{(i)} - \frac{1}{2} \int_0^t \sum_{i=1}^d ((\sigma_u^{(i)})^2 du - (\nu_u^{(i)})^2) du} \\ &= \frac{S_0}{N_0} e^{\sum_{i=1}^d \int_0^t (\sigma_u^{(i)} - \nu_u^{(i)}) d\tilde{B}_u^{(i)} - \frac{1}{2} \int_0^t \sum_{i=1}^d (\sigma_u^{(i)} - \nu_u^{(i)})^2 du - \int_0^t \sum_{i=1}^d (\sigma_u^{(i)} - \nu_u^{(i)}) \nu_u^{(i)} du} \\ &= \frac{S_0}{N_0} e^{\sum_{i=1}^d \int_0^t (\sigma_u^{(i)} - \nu_u^{(i)}) (d\tilde{B}_u^{(i)} - \nu_u^{(i)} du) - \frac{1}{2} \int_0^t \sum_{i=1}^d (\sigma_u^{(i)} - \nu_u^{(i)})^2 du} \end{split}$$

since
$$(\sigma_u^{(i)} - \nu_u^{(i)})(\sigma_u^{(i)} + \nu_u^{(i)}) - (\sigma_u^{(i)} - \nu_u^{(i)})^2 = 2\nu_u^{(i)}(\sigma_u^{(i)} - \nu_u^{(i)}).$$

By d-dim Girsanov's Theorem, the change of measure from \$ to €is

$$\begin{split} d\hat{B}_{t}^{(i)} &= d\tilde{B}_{t}^{(i)} - \nu_{t}^{(i)} dt \\ \frac{S_{t}}{N_{t}} &= \frac{S_{0}}{N_{0}} e^{\sum_{i=1}^{d} \int_{0}^{t} (\sigma_{u}^{(i)} - \nu_{u}^{(i)}) d\hat{B}_{u}^{(i)} - \frac{1}{2} \int_{0}^{t} \sum_{i=1}^{d} (\sigma_{u}^{(i)} - \nu_{u}^{(i)})^{2} du} \\ d\frac{S_{t}}{N_{t}} &= \frac{S_{t}}{N_{t}} \sum_{i=1}^{d} (\sigma_{t}^{(i)} - \nu_{t}^{(i)}) d\hat{B}_{t}^{(i)} \end{split}$$

Example 10.3.1 Suppose

- $N_0 = 1$ in the bank account at time 0, and
- N_t in the bank account at time t.

Then we have $N_t = e^{\int_0^t R_u du} = \frac{1}{D_t} \Rightarrow D_t N_t = 1$.

Using bank account as numéraire, we have that $\frac{S_t}{N_t} = D_t S_t$ is a martingale under risk-neutral measure.

Remark 10.3.2 Quotient of martingale is a martingale with respect to the measure you change to using the denominator and is normalized to start at 1 as the Radon-Nikodym derivative.

In our case, $\{D_tS_t\}_{t\geq 0}$ and $\{D_tN_t\}_{t\geq 0}$ are martingales under risk-neutral measure.

Remark 10.3.3 Volatility subtracts component-by-component for each $\tilde{B}_t^{(i)}$.

In our case, rrom $\tilde{\mathbb{P}}$ to $\hat{\mathbb{P}}$, we have

$$\begin{split} d\hat{B}_t^{(i)} &= d\tilde{B}_t^{(i)} - \nu_t^{(i)} dt \\ \hat{\mathbb{E}}[X] &= \tilde{\mathbb{E}}\left[\frac{D_T N_T}{N_0} X\right] \quad \text{focusing on } [0, T] \\ d\frac{S_t}{N_t} &= \frac{S_t}{N_t} \sum_{i=1}^d (\sigma_t^{(i)} - \nu_t^{(i)}) d\hat{B}_t^{(i)} \end{split}$$

10.4 Foreign Exchange Rate

Define variables:

- S_t : Stock Price in USD\$.
- X_t : Price of $\mathfrak{C}1$ in USD\$.
- R_t : Interest rate of the US market.
- $R_t^{\mathfrak{C}}$: Interest rate of the Euro market.

€1 in Euro bank $\longrightarrow \$ \frac{X_t}{D_t^{\mathfrak{C}}}$ asset in US market, where $\mathfrak{C} \frac{1}{D_t^{\mathfrak{C}}} = \mathfrak{C} e^{\int_0^t R_u^{\mathfrak{C}}} du$.

By definition of risk-neutral measure, in the US market, the discounted asset price process (using $D_t = e^{-\int_0^t R_u du}$) is a martingale, that is, $\{\frac{D_t}{D_s^c}X_t\}_{0 \le t \le T}$ is a martingale.

Suppose we have

$$\begin{cases} dX_t = \gamma_t X_t dt + X_t \sigma_t^X \left(\rho_t d\tilde{B}_t^{(1)} + \sqrt{1 - \rho_t^2} d\tilde{B}_t^{(2)} \right) \\ dS_t = R_t S_t dt + S_t \sigma_t^S d\tilde{B}_t^{(1)} \end{cases}$$
(*)

1. Focusing on X_t , what is γ_t ?

Since $\{\frac{D_t}{D_t^c}X_t\}_{0 \le t \le T}$ is a martingale, $d\frac{D_t}{D_t^c}X_t$ will only contain $d\tilde{B}_t^{(1)}$ and $d\tilde{B}_t^{(2)}$. We have

$$\frac{D_t}{D_t^{\mathfrak{C}}} X_t = X_t e^{\int_0^t (R_u^{\mathfrak{C}} - R_u) du}
d \frac{D_t}{D_t^{\mathfrak{C}}} X_t = e^{\int_0^t (R_u^{\mathfrak{C}} - R_u) du} dX_t + X_t e^{\int_0^t (R_u^{\mathfrak{C}} - R_u) du} (R_t^{\mathfrak{C}} - R_t) dt + 0
= X_t e^{\int_0^t (R_u^{\mathfrak{C}} - R_u) du} \left[\left(\gamma_t + R_t^{\mathfrak{C}} - R_t \right) dt + \sigma_t^X \left(\rho_t d\tilde{B}_t^{(1)} + \sqrt{1 - \rho_t^2} d\tilde{B}_t^{(2)} \right) \right]$$

Setting the drift term to zero, we get $\gamma_t = R_t - R_t^{\mathfrak{C}}$.

2. Change to \mathfrak{C} risk-neutral measure, that is, $\frac{S_t}{X_t}$.

Consider martingales $D_t S_t$ and $\frac{D_t}{D_t^C} X_t$ under \$ risk-neutral measure, taking quotient, we have that $\frac{D_t S_t}{D_t^C} X_t = \frac{D_t^C S_t}{X_t}$ is a martingale under $\mathbb C$ risk-neutral measure. Then we define

$$\begin{split} d\hat{B}_{t}^{(1)} &= d\tilde{B}_{t}^{(1)} - \rho_{t}\sigma_{t}^{X}dt, \quad d\hat{B}_{t}^{(2)} = d\tilde{B}_{t}^{(2)} - \sqrt{1 - \rho_{t}^{2}}\sigma_{t}^{X}dt \\ d\left(\frac{S_{t}}{X_{t}}\right) &= \frac{S_{t}}{X_{t}}\left(\frac{dS_{t}}{S_{t}} - \frac{dX_{t}}{X_{t}} + \frac{d[X,X]_{t}}{X_{t}^{2}} - \frac{d[S,X]_{t}}{S_{t}X_{t}}\right) \quad \text{by Itô's Quotient Rule} \\ &= \frac{S_{t}}{X_{t}}\left[R_{t}dt + \sigma_{t}^{S}d\tilde{B}_{t}^{(1)} - \gamma_{t}dt - \sigma_{t}^{X}\left(\rho_{t}d\tilde{B}_{t}^{(1)} + \sqrt{1 - \rho_{t}^{2}}d\tilde{B}_{t}^{(2)}\right) + \frac{X_{t}^{2}(\sigma_{t}^{X})^{2}dt}{X_{t}^{2}} - \frac{S_{t}X_{t}\sigma_{t}^{S}\sigma_{t}^{X}\rho_{t}dt}{S_{t}X_{t}}\right] \\ &= \frac{S_{t}}{X_{t}}\left[R_{t}^{C}dt + \left(\sigma_{t}^{S} - \sigma_{t}^{X}\rho_{t}\right)d\tilde{B}_{t}^{(1)} - \sigma_{t}^{X}\sqrt{1 - \rho_{t}^{2}}d\tilde{B}_{t}^{(2)} + \left(\sigma_{t}^{X}\right)^{2}dt - \sigma_{t}^{S}\sigma_{t}^{X}\rho_{t}dt\right] \\ &= \frac{S_{t}}{X_{t}}\left[R_{t}^{C}dt + \left(\sigma_{t}^{S} - \sigma_{t}^{X}\rho_{t}\right)\left(d\hat{B}_{t}^{(1)} + \rho_{t}\sigma_{t}^{X}dt\right) - \sigma_{t}^{X}\sqrt{1 - \rho_{t}^{2}}\left(d\hat{B}_{t}^{(2)} + \sqrt{1 - \rho_{t}^{2}}\sigma_{t}^{X}dt\right) + \left(\sigma_{t}^{X}\right)^{2}dt - \sigma_{t}^{S}\sigma_{t}^{X}\rho_{t}dt\right] \\ &= \frac{S_{t}}{X_{t}}\left[R_{t}^{C} + \sigma_{t}^{S}\sigma_{t}^{X}\rho_{t} - \left(\sigma_{t}^{X}\right)^{2}\rho_{t}^{2} - \left(\sigma_{t}^{X}\right)^{2}\left(1 - \rho_{t}^{2}\right) + \left(\sigma_{t}^{X}\right)^{2} - \sigma_{t}^{S}\sigma_{t}^{X}\rho_{t}\right]dt \\ &+ \frac{S_{t}}{X_{t}}\left[\left(\sigma_{t}^{S} - \sigma_{t}^{X}\rho_{t}\right)d\hat{B}_{t}^{(1)} - \sigma_{t}^{X}\sqrt{1 - \rho_{t}^{2}}d\hat{B}_{t}^{(2)}\right] \\ &= R_{t}^{C}\frac{S_{t}}{X_{t}}dt + \frac{S_{t}}{X_{t}}\left[\left(\sigma_{t}^{S} - \sigma_{t}^{X}\rho_{t}\right)d\hat{B}_{t}^{(1)} - \sigma_{t}^{X}\sqrt{1 - \rho_{t}^{2}}d\hat{B}_{t}^{(2)}\right]. \end{split}$$

3. $f(x) = \frac{1}{x}, f'(x) = -\frac{1}{x^2}, f''(x) = \frac{2}{x^3}$. Under $\hat{\mathbb{P}}^{\mathbb{C}}$ -measure, we have

$$\begin{split} d\left(\frac{1}{X_{t}}\right) &= -\frac{1}{X_{t}^{2}}dX_{t} + \frac{1}{2}\frac{2}{X_{t}^{3}}\left(dX_{t}\right)^{2} \\ &= -\frac{1}{X_{t}^{2}}\left(\gamma_{t}X_{t}dt + X_{t}\sigma_{t}^{X}\left(\rho_{t}d\tilde{B}_{t}^{(1)} + \sqrt{1 - \rho_{t}^{2}}d\tilde{B}_{t}^{(2)}\right)\right) + \frac{1}{X_{t}^{3}}X_{t}^{2}\left(\sigma_{t}^{X}\right)^{2}dt \\ &= \frac{1}{X_{t}}\left(R_{t}^{\mathfrak{C}} - R_{t}\right)dt + \frac{\sigma_{t}^{X}}{X_{t}}\left(\rho_{t}\left(d\hat{B}_{t}^{(1)} + \rho_{t}\sigma_{t}^{X}dt\right) + \sqrt{1 - \rho_{t}^{2}}\left(d\hat{B}_{t}^{(2)} + \sqrt{1 - \rho_{t}^{2}}\sigma_{t}^{X}dt\right) + \left(\sigma_{t}^{X}\right)^{2}dt\right) \\ &= \frac{1}{X_{t}}\left(R_{t}^{\mathfrak{C}} - R_{t}\right)dt + \frac{\sigma_{t}^{X}}{X_{t}}\left(\rho_{t}d\hat{B}_{t}^{(1)} + \sqrt{1 - \rho_{t}^{2}}d\hat{B}_{t}^{(2)}\right) \end{split}$$

Note that under $\tilde{\mathbb{P}}^{\$}$ -measure, we have $dX_t = X_t(R_t - R_t^{\mathbf{C}})dt + X_t\sigma_t^X\left(\rho_t d\tilde{B}_t^{(1)} + \sqrt{1 - \rho_t^2}d\tilde{B}_t^{(2)}\right)$ from (*).

4. Useful table:

Currency	US Bank	Stock	Euro Bank
$\$$ risk-neutral measure, \tilde{B}_t	Time 0: \$1	Time 0: S_0	Time 0: X_0
	Time t : $\frac{1}{D_t}$	Time t : S_t	Time $t: \frac{X_t}{D_t^{\mathbf{C}}}$
Martingale, D_t	$\{\$1\}_{0 \le t \le T}$ is a martingale	$\{D_tS_t\}_{0\leq t\leq T}$ is a martingale	$\{\frac{D_t}{D_t^{\mathfrak{C}}}X_t\}_{0\leq t\leq T}$ is a martingale
$lacktriangleright$ risk-neutral measure, \hat{B}_t	Time 0: $\frac{1}{X_0}$	Time 0: $\frac{S_0}{X_0}$	Time 0: €1
	Time t : $\frac{1}{X_t D_t}$	Time t : $\frac{S_t^{\circ}}{X_t}$	Time $t: \frac{\mathfrak{C}_1}{D_t^{\mathfrak{C}}}$
Martingale, $D_t^{\mathfrak{C}}$	$\left\{\frac{D_t^{\mathfrak{C}}}{D_t}\frac{1}{X_t}\right\}_{0\leq t\leq T}$ is a martingale	$\{\frac{D_t^{\epsilon}S_t}{X_t}\}_{0 \le t \le T}$ is a martingale	$\{\mathfrak{C}1\}_{0\leq t\leq T}$ is a martingale

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