Probability for Computer Science

Spring 2021

Lecture 12



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Today

- Multiple discrete random variables
 - Conditional Expectation (continued)
 - Indep. of multiple r.v.s
- The Sample Mean
- If time: Continuous Random Variables



Conditional Expectation

- The conditional expectation of r.v. X, given an event A with P(A)>0 is defined as: $E[X|A] = \sum_{x} x \cdot p_{X|A}(x)$
- For a function g(X), $E[g(X)|A] = \sum_{x} g(x) \cdot p_{X|A}(x)$
- Given r.v.s X and Y associated with the same experiment, the conditional expectation of X given a value y of Y is:

$$E[X|Y = y] = \sum_{x} x \cdot p_{X|Y}(x|y)$$



Total Expectation Theorem

• Given disjoint events A_1 , ..., A_n that partition the sample space, with $P(A_i) > 0$ for all i,

$$E[X] = \sum_{i=1}^{n} P(A_i)E[X|A_i]$$

Example

Messages are sent from a computer in Boston, over the internet to the following destinations, with the following probabilities:

- NYC with probability 0.5
- DC with probability 0.3
- Denver with probability 0.2

The transit time is a random variable, T. Its expectation, conditioned on each city, is:

- 0.05 if message destination is NYC
- 0.1 if message destination is DC
- 0.3 if message destination is Denver

Q: What is E[T]?

Independence of a r.v. from an event

A random variable, X, is independent of an event, A, if, for all x,

$$P(X = x \text{ and } A) = P(X = x)P(A)$$

= $p_X(x)P(A)$

i.e., X is independent of A if the events {X=x} and A are independent, for every value of x.

To prove or disprove that r.v.s X and Y are independent, it is enough to prove or disprove any of the following statements (as they are equivalent):

•
$$p_{X,Y}(a,b) = p_X(a)p_Y(b)$$

for all a and b

•
$$p_X(a) = p_{X|Y}(a | b)$$

for all a and b s.t. $p_y(b) > 0$

•
$$p_{Y}(b) = p_{Y|X}(b | a)$$

for all a and b s.t. p_X (a) > 0

Are X and Y independent?

Y

	y 1	y ₂	y ₃	Y 4
x ₁	0.05	0.15	0	0.2
x ₂	0.025	0.075	0	0.1
X ₃	0.05	0.15	0	0.2

Are X and Y independent? Yes.

How can we tell?

X

Y

	y ₁	y ₂	y ₃	y ₄
X ₁	0.05	0.15	0	0.2
X ₂	0.025	0.075	0	0.1
X ₃	0.05	0.15	0	0.2

How can we tell that X and Y are independent?

The columns are multiples of each other. Therefore $p_{X|Y}(x | y)$ is the same for every value of y, and therefore does not depend on y, so $p_{X|Y}(x | y) = p_X(x)$.

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	y 1	y ₂	y ₃	Y 4
X ₁	0.05	0.15	0	0.2
X ₂	0.025	0.075	0	0.1
X ₃	0.05	0.15	0	0.2

Conditional independence of r.v.s

Random variables X and Y are conditionally independent, given event A if, for all x, y,

$$P(X = x, Y = y \mid A) = P(X = x \mid A)P(Y = y \mid A)$$

$$= p_{X|A}(x)p_{Y|A}(y)$$

Equivalently, for all x and y s.t. $p_{Y|A}(y) > 0$,

$$p_{X|Y,A}(x|y) = p_{X|A}(x)$$

Properties of independent r.v.s

If X and Y are independent random variables, then:

- E[XY] = E[X]E[Y]
- E[g(X)h(Y)] = E[g(X)]E[h(Y)]

• var(X + Y) = var(X) + var(Y)

NOTE: Not true in general for arbitrary r.v.s!

Independence of multiple r.v.s

Random variables X, Y, and Z are independent if:
 For all x, y, z:

$$p_{X,Y,Z}(x,y,z) = p_X(x)p_Y(y)p_Z(z)$$

Let X₁ ,..., X_n be independent random variables.
 Then:

$$\operatorname{\mathtt{var}}\left(\sum_{i=1}^n X_i
ight) = \sum_{i=1}^n \operatorname{\mathtt{var}}(X_i)$$

Note: These properties need not hold in general – need independence.

The Sample Mean

- Suppose we want to estimate the approval rating of a public figure, B.
- We ask n people drawn uniformly at random from the population.
- Define X_i as an indicator random variable for whether the i-th person approves of B.
- We model X_1 , X_2 , ..., X_n as independent Bernoulli random variables, with common mean, p, and variance p(1-p).
- That is, we assume p is the true approval rating of B. It is unknown, so we try to estimate it.
- We compute the Sample Mean from the *n* responses, i.e. the average approval rating in the *n*-person sample:

$$S_n = \frac{X_1 + X_2 + \dots + X_n}{n}$$

Mean and Variance of the Sample Mean

The Sample Mean is:

$$S_n = \frac{X_1 + X_2 + \dots + X_n}{n}$$

What are E[S_n] and var(S_n)?

Continuous Random Variables

- Probability Density Function (PDF)
- Expectation, Variance
- Random variables: Uniform, Exponential
- Conditional Distribution Function (CDF)

Continuous random variables

A random variable, X, is continous if there is a nonnegative function f_X , called the probability density function (PDF), such that for every subset of the real line, $B \subseteq \mathbb{R}$,

$$P(X \in B) = \int_{B} f_{X}(x)dx$$

The probability that X falls in interval [a, b] is:

$$P(a \le X \le b) = \int_a^b f_X(x) dx$$

Properties of PDF

• Nonnegativity: $f_X(x) \ge 0 \quad \forall x \in \mathbb{R}$

Normalization:

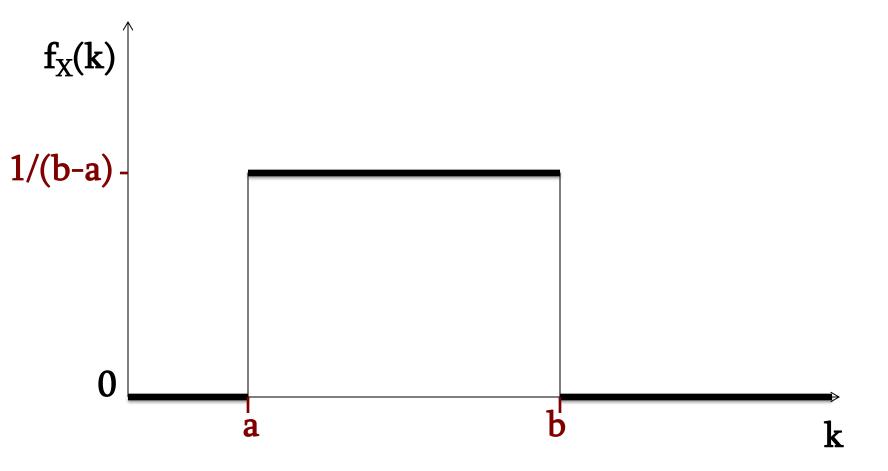
$$P(-\infty < X < \infty) = \int_{-\infty}^{\infty} f_X(x) dx = 1$$

Uniform random variable

A continuous r.v., X, is uniform (or uniformly distributed) if it takes values in an interval [a, b], and any two sub-intervals of the same length have the same probability. Its PDF is:

$$f_X(k) = \begin{cases} \frac{1}{b-a} & a \le k \le b \\ 0 & \text{otherwise} \end{cases}$$

Uniform PDF



Expectation of a continuous r.v.

If X is a continuous r.v., its expectation (or mean, or expected value) is defined as:

$$\left(E[X] = \int_{-\infty}^{\infty} x f_X(x) dx\right)$$

Note: this is exactly the center of gravity of the PDF. That is, it is the exact tipping point of the area under the PDF.

[True in the discrete case as well].

Functions of a continuous r.v.

If X is a continuous random variable, and Y is a real-valued function of X, g(X), then Y is a random variable.

Y can be continuous, e.g. Y = 3 X.

• Y can be discrete, e.g.
$$Y = \begin{cases} 1 & X > 0 \\ 0 & \text{otherwise} \end{cases}$$

Expected value of a function of a continuous r.v.:

$$E[Y] = \left[E[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx \right]$$

Variance of continuous r.v.

Variance of a continuous r.v., X, is defined:

$$var(X) = E[(X - E[X])^{2}] = \int_{-\infty}^{\infty} (x - E[X])^{2} f_{X}(x) dx$$

Note: same definition (LHS) as in discrete case, but for the expectation of a function of a continuous r.v., need to integrate.

Similar to discrete case, n-th moment is E[Xⁿ], and:

$$0 \le \operatorname{var}(X) = E[X^2] - (E[X])^2$$

Linear functions of a continuous r.v.

If X is a continuous r.v. and Y is a linear function of X, i.e. Y = a X + b for some constants a, b, then:

1.
$$E[Y] = E[a X + b] = a E[X] + b$$

2.
$$var(Y) = var(a X + b) = a^2 var(X)$$

Note: just as in discrete case.

Exponential random variable

An exponential random variable, X, is a continuous r.v. with the following PDF, for a constant $\lambda > 0$.

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x \ge 0\\ 0, & \text{otherwise} \end{cases}$$

Exponential PDF

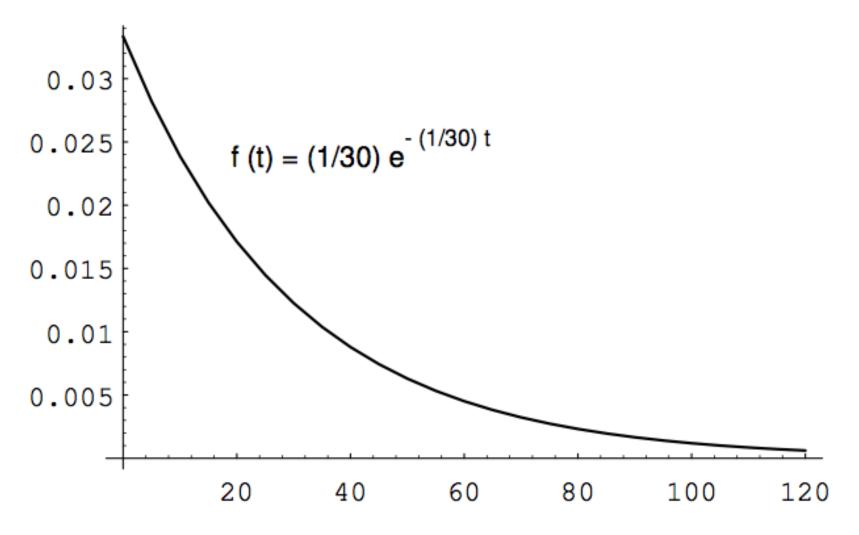


Figure 2.20: Exponential density with $\lambda = 1/30$.

Credit: Grinstead and Snell, 1997.

Cumulative distribution function (CDF)

For any r.v., X, discrete or continuous, the cumulative distribution function, CDF is:

$$F_X(x) = P(X \le x)$$

which is computed as follows:

$$F_X(x) = \begin{cases} \sum_{k \le x} p_X(k), & \text{if X is a discrete r.v.} \\ \int_{-\infty}^x f_X(t) dt, & \text{if X is a continuous r.v.} \end{cases}$$

Compute PMF from CDF, for discrete r.v.

If X is a discrete r.v., then CDF of X is:

$$F_X(x) = \sum_{k \le x} p_X(k)$$

Can therefore compute PMF from CDF as follows:

$$p_X(k) = P(X = k) = P(X \le k) - P(X \le k - 1)$$
$$= F_X(k) - F_X(k - 1)$$

Compute PDF from CDF, for continuous r.v.

If X is a continuous r.v., then CDF of X is:

$$F_X(x) = \int_{-\infty}^x f_X(t)dt$$

In other words, the value of the CDF at x, $F_X(x)$, is equal to the area under PDF, f_X , from $-\infty$ to x.

Can therefore compute PMF by differentiating the CDF:

$$f_X(x) = \frac{dF_X}{dx}(x)$$

Properties of the CDF

F_X is monotonically non-decreasing:
 if j ≤ k then F_X (j) ≤ F_X(k)

- $F_X(k)$ tends to 0 as $k \rightarrow -\infty$
- $F_x(k)$ tends to 1 as $k \rightarrow \infty$

- If X is a discrete r.v. then its CDF, $F_X(k)$, is a piecewise constant function of k.
- If X is a continuous r.v. then its CDF, $F_X(k)$, is a continuous function of k.

