



$$(a) f(w) = \frac{1}{n} \sum_{i=1}^n \{ -y_i w^T x_i + \log(1 + \exp(w^T x_i)) \} + \frac{\lambda}{2} \|w\|_2^2$$

assume  $\|x_i\| \leq R$

$$\begin{aligned} \nabla_w f(w) &= \frac{1}{n} \sum_{i=1}^n \left( -y_i x_i + \frac{x_i \exp(w^T x_i)}{1 + \exp(w^T x_i)} \right) + \lambda \|w\| \\ &= \frac{1}{n} \sum_{i=1}^n (b(w^T x_i) - y_i) x_i + \lambda \|w\|, \text{ where } b(w^T x_i) = \frac{\exp(w^T x_i)}{1 + \exp(w^T x_i)} \end{aligned}$$

Steps: ① Initialize  $w_i$  values as random values ( $i=0, 1, \dots, d$ )

where  $d$  is the number of features,  $w = \begin{bmatrix} w_1 \\ \vdots \\ w_d \end{bmatrix}$

② Compute  $\nabla_w f(w)$  above

③ compute  $y_{k+1} = w_k - \eta \nabla_w f(w)$ , until it converges

④ if  $\|y_{k+1}\| \leq R$ ,  $w_{k+1} = y_{k+1}$ , if  $\|y_{k+1}\| > R$ ,  $w_{k+1} = \frac{R}{\|y_{k+1}\|} y_{k+1}$

(b) Yes, for every  $w_1 \succ w_2$ , to prove

$$f(w_1) \geq f(w_2) + (w_1 - w_2)^T \nabla f(w_2) + \frac{\lambda}{2} \|w_1 - w_2\|^2$$

$$\begin{aligned} \text{We need to prove } & \frac{1}{n} \sum_{i=1}^n \{ -y_i w_1^T x_i + \log(1 + \exp(w_1^T x_i)) \} + \frac{\lambda}{2} \|w_1\|^2 \geq \\ & \frac{1}{n} \sum_{i=1}^n \{ -y_i w_2^T x_i + \log(1 + \exp(w_2^T x_i)) \} + \frac{\lambda}{2} \|w_2\|^2 + (w_1 - w_2)^T \left\{ \frac{1}{n} \sum_{i=1}^n (b(w_2^T x_i) - y_i) x_i \right. \\ & \left. + \lambda \|w_2\| \right\} + \frac{\lambda}{2} \|w_1 - w_2\|^2. \end{aligned}$$

To prove  $f(w)$  is strongly convex, we can prove  $f(w) - \frac{\lambda}{2} \|w\|^2$  is convex

$$\nabla^2 (f(w) - \frac{\lambda}{2} \|w\|^2) = \frac{1}{n} \sum_{i=1}^n \left\{ \frac{\exp(w^T x_i) \cdot [(1 + \exp(w^T x_i) + \exp(w^T x_i)) x_i]}{(1 + \exp(w^T x_i))^2} \right\} x_i \geq 0$$

Since  $\exp(w^T x_i) > 0$  and  $x_i \geq 0$  ( $i=0, 1, \dots, n$ )

The subjective function  $f(w)$  is strongly convex.

教师点评

评后反思

(c) Yes. Since for  $\beta > \lambda$ , since  $\log\left(\frac{1+\exp(x_i)}{1+\exp(x_j)}\right) + (x_i - x_j)^T \frac{\exp(x_i)}{1+\exp(x_i)} + \frac{\beta}{2} \|x_i - x_j\|^2 \geq 0$  if  $\beta = 2(x_i - x_j)^T \frac{\exp(x_i)}{1+\exp(x_i)}$  because  $\frac{\exp(x_i)}{1+\exp(x_i)} > 0$  and  $\log\left(\frac{1+\exp(x_i)}{1+\exp(x_j)}\right)$  is growing less than gradient function, and  $-y^T x$  is  $\beta$ -smooth since  $-y(w_1^T - w_2^T)x + (w_1 - w_2)^T(-yx) + \frac{\beta}{2} \|w_1 - w_2\|^2 \geq 0$ . Therefore

for  $\beta > \lambda$ , there exists  $w_1, w_2$  where  $f(w_1) \leq f(w_2) + (w_1 - w_2)^T \nabla f(w_2) + \frac{\beta}{2} \|w_1 - w_2\|^2$ .

If  $\beta = 2 \cdot y^T x$ , obviously  $\log\left(\frac{1+\exp(x_i)}{1+\exp(x_j)}\right) + (w_1 - w_2)^T \lambda \|w_1 - w_2\| + \frac{\beta}{2} \|w_1 - w_2\|^2 \geq 0$

(d) For  $\beta$ -smooth and  $\alpha$ -strongly convex function.

Strong-convexity:  $f(w_1) - f(w^*) \leq (w_1 - w^*)^T \nabla f(w^*) + \frac{\beta}{2} \|w_1 - w^*\|^2$   
since  $\nabla f(w^*) \neq 0$

$$\begin{aligned} w_{t+1} &= w_t - \eta_T \nabla f(w_t) \\ \|w_{t+1} - w^*\|^2 &= \|w_t - w^*\|^2 + \eta_T^2 \|\nabla f(w_t)\|^2 - 2\eta_T \nabla f(w_t)^T (w_t - w^*) \\ &\leq \|w_t - w^*\|^2 + \eta_T^2 \|\nabla f(w_t)\|^2 - 2\eta_T \left( \frac{\partial \beta}{\partial \beta} \|w_t - w^*\|^2 + \frac{1}{\partial \beta} \|\nabla f(w_t)\|^2 \right) \\ &= (1 - 2\eta_T \frac{\partial \beta}{\partial \beta}) \|w_t - w^*\|^2 + \eta_T \left( \eta_T - \frac{2}{\partial \beta} \right) \|\nabla f(w_t)\|^2 \end{aligned}$$

Suppose  $\eta_T \leq \frac{2}{\partial \beta}$ ,  $\|w_{t+1} - w^*\|^2 \leq (1 - 2\eta_T \frac{\partial \beta}{\partial \beta}) \|w_t - w^*\|^2$   
 $\leq \|w_0 - w^*\|^2 \prod_{n=1}^k (1 - 2\eta_n \frac{\partial \beta}{\partial \beta})$  and  $1 - x \leq e^{-x}$

Therefore,  $f(w_t) - f(w^*) \leq \frac{\beta}{2} \|w_t - w^*\|^2 \leq \frac{\beta}{2} \prod_{n=1}^k (1 - 2\eta_n \frac{\partial \beta}{\partial \beta}) \|w_0 - w^*\|^2$   
 $\leq \frac{\beta}{2} \|w_0 - w^*\|^2$

2.



2.

(a)

E-step: Denote the current parameter values as  $\Phi$ . Compute  $W_{ik}$  for all data points  $x_i$  given the old set of parameters  $\Phi^b$  and set the label  $G_i$  with the largest likelihood  $P(G_i | x_i, \Phi^b) = h_i^b$  to the datapoint.

M-step: recompute the prior probability  $P(G_i)$ , mean vector and the covariance matrix with the posterior  $h_i^b$ , the data points  $X^b$  and the total number of class  $N$ , update the new set of parameters to be  $\Phi^{(t)}$ , maximum likelihood is what to get new parameters.

Keep doing E-step & M-step until  $h_i^b$  converges.

$$(b) \quad \pi_i^{(t+1)} = \frac{\sum_{j=1}^N h_j^b}{N}, \quad M_i^{(t+1)} = \frac{\sum_{j=1}^N h_j^b x_j^b}{\sum_{j=1}^N h_j^b}$$

$$\bar{\Sigma}_i^{(t+1)} = \frac{\sum_{j=1}^N h_j^b (x_j^b - M_i^{(t+1)}) (x_j^b - M_i^{(t+1)})^T}{\sum_{j=1}^N h_j^b} \quad \text{where } l \text{ is the step number}$$

$$(c) \quad h_i^t = P(\pi_i | x^t, M_i, \bar{\Sigma}_i) = \frac{P(x^t | \pi_i, M_i, \bar{\Sigma}_i) P(\pi_i)}{\sum_{j=1}^K P(x^t | \pi_j, M_j, \bar{\Sigma}_j) P(\pi_j)}$$

$$P(x^t | \pi_i, M_i, \bar{\Sigma}_i) = \pi_i \frac{1}{(2\pi)^{p/2} |\bar{\Sigma}_i|^{p/2}} \exp\left(-\frac{1}{2} (x^t - M_i)^T \bar{\Sigma}_i^{-1} (x^t - M_i)\right)$$

3.

Error rates for MyLogisticReg2 with Boston50						
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5	Mean	SD
0.1980198 01980197 97	0.1188118 811881188	0.1089108 910891089	0.2475247 524752475 2	0.1188118 811881188	0.1584158 415841584	0.0549482 650515311 16

Error rates for MyLogisticReg2 with Boston25						
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5	Mean	SD
0.1386138 61386138 63	0.1881188 118811880 6	0.0396039 603960396 4	0.0594059 405940594 6	0.2178217 821782178	0.1287128 712871287 2	0.0697299 677689465 4

Error rates for LogisticRegression with Boston50						
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5	Mean	SD
0.1386138 61386138 63	0.1386138 613861386 3	0.0990099 009900989 9	0.2772277 227722772 5	0.0990099 009900989 9	0.1504950 495049505	0.0657950 40498361

Error rates for LogisticRegression with Boston25						
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5	Mean	SD
0.1584158 41584158 45	0.2178217 821782178	0.0396039 603960396 4	0.1287128 712871287 2	0.1386138 613861386 3	0.1366336 633663366 4	0.0575280 754310255 26