

Series 7, May 20th, 2019 (Mixture Models, EM Algorithm)

Solutions will be published on Monday, May 27th.

Problem 1 (Mixture Models and Expectation-Maximization Algorithm):

Consider a one-dimensional Gaussian Mixture Model with 2 clusters and parameters $(\mu_1, \sigma_1^2, \mu_2, \sigma_2^2, w_1, w_2)$. Here (w_1, w_2) are the mixing weights, and (μ_1, σ_1^2) , (μ_2, σ_2^2) , are the centers and variances of the clusters. We are given a dataset $\mathcal{D} = \{\mathbf{x}_1, \mathbf{x}_2, \mathbf{x}_3\} \subset \mathbb{R}$, and apply the EM-algorithm to find the parameters of the Gaussian mixture model.

1. Write down the complete log-likelihood that is being optimized, *for this problem*.

Assume that the dataset \mathcal{D} consists of the following three points, $\mathbf{x}_1 = 1, \mathbf{x}_2 = 10, \mathbf{x}_3 = 20$. At some step in the EM-algorithm, we compute the expectation step which results in the following matrix:

$$R = \begin{bmatrix} 1 & 0 \\ 0.4 & 0.6 \\ 0 & 1 \end{bmatrix}$$

where r_{ic} denotes the probability of \mathbf{x}_i belonging to cluster c .

In the next questions, leave all results unsimplified, i.e. in fractional form.

2. Given the above R for the expectation step, write the result of the maximization step for the mixing weights w_1, w_2 . You can use the equations for maximum likelihood updates without proof.
3. Do the same for μ_1, μ_2 . Given the above R for the expectation step, write the result of the maximization step for the centers μ_1, μ_2 . You can use the equations for maximum likelihood updates without proof.
4. Do the same for σ_1^2, σ_2^2 . Given the above R for the expectation step, write the result of the maximization step for the variance values σ_1^2, σ_2^2 . You can use the equations for maximum likelihood updates without proof.
5. The previous two questions are doing soft-EM. Calculate the maximization step of μ_1, μ_2 for hard-EM.

Problem 2 (Mixture Models and MAP estimation):

Consider a mixture of K multivariate Bernoulli distributions with parameters $\boldsymbol{\mu} = \{\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_K\}$, where $\boldsymbol{\mu}_i = \{\mu_{i1}, \dots, \mu_{iD}\}$. You will use EM algorithm to compute MLE and MAP estimates.

1. What is the M step for μ_{ij} using MLE?
2. Now, suppose you want to do MAP estimation. What is the E step?
3. What is the M step for μ_{ij} using MAP? You can assume a Beta(α, β) prior.

Problem 3 (A Different Perspective on EM):

In this question you will show that EM can be seen as an iterative algorithm which maximizes a lower bound on the log-likelihood. We will treat any general model $P(X, Z)$ with observed variables X and latent variables Z . For the sake of simplicity, we will assume that Z is discrete and takes values in $\{1, 2, \dots, m\}$. If we observe X , the goal is to maximize the log-likelihood

$$\ell(\theta) = \log P(\mathbf{x}; \theta) = \log \sum_{z=1}^m P(\mathbf{x}, z; \theta)$$

with respect to the parameter vector θ . $Q(Z)$ denotes *any* distribution over the latent variables.

- Show that if $Q(z) > 0$ when $P(\mathbf{x}, z) > 0$, then it holds that

$$\ell(\theta) \geq \mathbb{E}_Q[\log P(X, Z)] - \sum_{z=1}^m Q(z) \log Q(z).$$

Hence, we have a bound on the log-likelihood parametrized by a distribution $Q(Z)$ over the latent variables.
(Hint: Consider using Jensen's inequality)

- Show that for a fixed θ , the lower bound is maximized for $Q^*(Z) = P(Z | X; \theta)$. Moreover, show that the bound is exact (holds with equality) for this specific distribution $Q^*(Z)$.
(Hint: Do not forget to add Lagrange multipliers to make sure that Q^* is a valid distribution.)
- Show that if we optimize with respect to Q and θ in an alternating manner, this corresponds to the EM procedure. Discuss what this implies for the convergence properties of EM.