

数学作业纸

(科目: 随机)

班级: 自93

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编号: 2019010702 第 1 页

28. $D = \{(x, y) | 0 < x < y < 1\}$

$$f(x, y) = f_{X|Y}(x|y) f_Y(y) \uparrow$$

$$= \begin{cases} 15x^2y & 0 < x < y < 1 \\ 0 & \text{其他} \end{cases}$$

$$P(X > \frac{1}{2}) = \iint_{\{(x,y) | x > \frac{1}{2}\} \cap D} 15x^2y \, dx \, dy$$

$$= \int_{\frac{1}{2}}^1 dy \int_{\frac{1}{2}}^y 15x^2y \, dx$$

$$= \int_{\frac{1}{2}}^1 5y^5 - \frac{5}{8}y \, dy$$

$$= y^5 - \frac{5}{16}y^2 \Big|_{\frac{1}{2}}^1 = \frac{47}{64}$$

$$EX = \iint_D 15x^3y \, dx \, dy$$

$$= \int_0^1 dy \int_0^y 15x^3y \, dx$$

$$= \int_0^1 \frac{15}{4}y^5 \, dy$$

$$= \frac{5}{8}$$

$$EXY = \iint_D xy \cdot 15x^2y \, dx \, dy$$

$$= \iint_D 15x^3y^2 \, dx \, dy$$

$$= \int_0^1 dy \int_0^y 15x^3y^2 \, dx$$

$$= \int_0^1 \frac{15}{4}y^6 \, dy = \frac{15}{28}$$

33. $X \sim N(\mu, \sigma^2)$

$$E(X | a < X < b) = \frac{E(X \mathbb{I}_{a < X < b})}{P(a < X < b)}$$

$$P(a < X < b) = \Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)$$

$$E(X \mathbb{I}_{a < X < b}) = \int_a^b x f(x) \, dx$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \int_a^b x e^{-\frac{(x-\mu)^2}{2\sigma^2}} \, dx$$

$$= \frac{1}{\sqrt{2\pi}} \int_{\frac{a-\mu}{\sigma}}^{\frac{b-\mu}{\sigma}} (\sigma t + \mu) e^{-\frac{1}{2}t^2} \, dt$$

$$= \frac{1}{\sqrt{2\pi}} \left(\sigma \int_{\frac{a-\mu}{\sigma}}^{\frac{b-\mu}{\sigma}} t e^{-\frac{1}{2}t^2} \, dt + \mu \int_{\frac{a-\mu}{\sigma}}^{\frac{b-\mu}{\sigma}} e^{-\frac{1}{2}t^2} \, dt \right)$$

$$= -\frac{\sigma}{\sqrt{2\pi}} \left(e^{-\frac{(b-\mu)^2}{2\sigma^2}} - e^{-\frac{(a-\mu)^2}{2\sigma^2}} \right) + \mu \int_{\frac{a-\mu}{\sigma}}^{\frac{b-\mu}{\sigma}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2} \, dt$$

$$= \frac{\sigma(e^{-\frac{(a-\mu)^2}{2\sigma^2}} - e^{-\frac{(b-\mu)^2}{2\sigma^2}})}{\sqrt{2\pi}} + \mu [\Phi(\frac{b-\mu}{\sigma}) - \Phi(\frac{a-\mu}{\sigma})]$$

$$\therefore E(X | a < X < b)$$

$$= \frac{\sigma(e^{-\frac{(a-\mu)^2}{2\sigma^2}} - e^{-\frac{(b-\mu)^2}{2\sigma^2}})}{\sqrt{2\pi}[\Phi(\frac{b-\mu}{\sigma}) - \Phi(\frac{a-\mu}{\sigma})]} + \mu$$

34. $f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) \, dx = \int_0^{+\infty} \lambda^2 e^{-\lambda x} \, dx$

$$= -\lambda \int_0^{+\infty} e^{-\lambda x} d(1-\lambda x) = -\lambda e^{-\lambda x} \Big|_0^{+\infty} = \lambda e^{-\lambda y} \quad y > 0$$

$$f_Y(y) = \begin{cases} \lambda e^{-\lambda y} & y > 0 \\ 0 & \text{其他} \end{cases}$$



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$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

$$= \frac{\lambda^2 e^{-\lambda x}}{\lambda e^{-\lambda y}} = \lambda e^{-\lambda(x-y)} \quad 0 < y < \infty$$

$$\therefore f_{X|Y}(x|y) = \begin{cases} \lambda e^{-\lambda(x-y)} & 0 < y < x \\ 0 & y > x \end{cases}$$

$$\therefore E(X|Y=y) = \int_{-\infty}^{+\infty} x f_{X|Y}(x|y) dx$$

$$= \int_y^{+\infty} \lambda x e^{-\lambda x} \cdot e^{\lambda y} dx$$

$$= e^{\lambda y} \int_y^{+\infty} \lambda x e^{-\lambda x} dx$$

$$= \frac{1}{\lambda} e^{\lambda y} \int_y^{+\infty} \lambda x e^{-\lambda x} d(\lambda x)$$

$$= \frac{1}{\lambda} e^{\lambda y} \int_{\lambda y}^{+\infty} t e^{-t} dt$$

$$= \frac{1}{\lambda} e^{\lambda y} \left(-(t+1)e^{-t} \Big|_{\lambda y}^{+\infty} \right)$$

$$= \frac{1}{\lambda} e^{\lambda y} (\lambda y + 1) e^{-\lambda y}$$

$$= \frac{\lambda y + 1}{\lambda} = y + \frac{1}{\lambda}$$

$$E(X^2|Y=y) = \lambda \int_y^{+\infty} x^2 e^{-\lambda(x-y)} dx$$

$$= \lambda e^{\lambda y} \int_y^{+\infty} x^2 e^{-\lambda x} dx$$

$$= \frac{1}{\lambda^2} e^{\lambda y} \int_y^{+\infty} (\lambda x)^2 e^{-\lambda x} d(\lambda x)$$

$$= \frac{1}{\lambda^2} e^{\lambda y} \int_{\lambda y}^{+\infty} t^2 e^{-t} dt$$

$$= \left(-t^2 e^{-t} - 2(t+1)e^{-t} \Big|_{\lambda y}^{+\infty} \right) \frac{1}{\lambda^2} e^{\lambda y}$$

$$= \frac{\lambda^2 y^2 + 2(\lambda y + 1)}{\lambda^2}$$

$$\therefore D(X|Y=y)$$

$$= E(X^2|Y=y) - (E(X|Y=y))^2$$

$$= \frac{\lambda^2 y^2 + 2(\lambda y + 1)}{\lambda^2} - \frac{(\lambda y + 1)^2}{\lambda^2}$$

$$= \frac{\lambda^2 y^2 + 2\lambda y + 1 + 1 - (\lambda y + 1)^2}{\lambda^2}$$

$$= \frac{(\lambda y + 1)^2 + 1 - (\lambda y + 1)^2}{\lambda^2}$$

$$= \frac{1}{\lambda^2}$$

$$\text{as } \eta \sim U[0, a], X \sim U[\eta, a]$$

$$F_{X, \eta}(x, \eta) = P(X \leq x, \eta \leq \eta)$$

$$= \begin{cases} \frac{\eta}{a} \cdot \frac{x - \eta}{a - \eta} & 0 \leq \eta \leq x < a \\ 0 & \text{else} \end{cases}$$

$$f_{X, \eta}(x, \eta) = \frac{\partial^2 F}{\partial x \partial \eta} = \begin{cases} \frac{1}{(a - \eta)^2} & 0 \leq \eta \leq x < a \\ 0 & \text{else} \end{cases}$$

$$f_{X|\eta}(x|\eta) = \frac{f_{X, \eta}(x, \eta)}{f_{\eta}(\eta)} = \frac{\frac{1}{(a - \eta)^2}}{\frac{1}{a}} = \frac{a}{(a - \eta)^2}, \quad 0 \leq \eta \leq x < a$$

$$E(X|\eta = \eta) = \int_{\eta}^a \frac{a}{(a - \eta)^2} x dx = \frac{a + \eta}{2(a - \eta)} \therefore E(X|\eta) = \frac{a + \eta}{2(a - \eta)}$$



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$$35. X | \eta = \eta_0 \sim U[\eta_0, a]$$

$$\therefore E(X | \eta = \eta_0) = \frac{a + \eta_0}{2}$$

$$\therefore E(X | \eta) = \frac{a + \eta}{2}$$

$$\therefore \eta \sim U[0, a]$$

$$\therefore \frac{a + \eta}{2} \sim U[\frac{a+0}{2}, \frac{a+a}{2}]$$

$$\therefore E(X | \eta) \sim U[\frac{a}{2}, a]$$

$$41. X_1 \sim E(\lambda_1)$$

$$F_{X_1 | X_1 < X_2}(x) = P(X_1 \leq x | X_1 < X_2)$$

$$= \frac{P(X_1 \leq x, X_1 < X_2)}{P(X_1 < X_2)}$$

$$P(X_1 < X_2) = \int_{-\infty}^{+\infty} P(X_1 < X_2 | X_2 = x) f_{X_2}(x) dx$$

$$= \int_{-\infty}^{+\infty} P(X_1 < x) f_{X_2}(x) dx$$

$$= \int_0^{+\infty} (1 - e^{-\lambda_1 x}) \lambda_2 e^{-\lambda_2 x} dx$$

$$= \frac{\lambda_1}{\lambda_1 + \lambda_2}$$

$$P(X_1 \leq x, X_1 < X_2)$$

$$= \int_0^x dx_1 \int_{x_1}^{+\infty} \lambda_1 \lambda_2 e^{-\lambda_1 x_1} e^{-\lambda_2 x_2} dx_2$$

$$= \int_0^x \lambda_1 e^{-\lambda_1 x_1} (e^{-\lambda_2 x_1}) dx_1$$

$$= \int_0^x \lambda_1 e^{-\lambda_1 x_1} e^{-\lambda_2 x_1} dx_1$$

$$= \int_0^x \lambda_1 e^{-(\lambda_1 + \lambda_2) x_1} dx_1 \quad x > 0$$

$$\therefore \bar{F}_{X_1 | X_1 < X_2}(x) = \int_0^x \frac{\lambda_1 e^{-(\lambda_1 + \lambda_2) x_1}}{\frac{\lambda_1}{\lambda_1 + \lambda_2}} dx_1$$

$$= \int_0^x (\lambda_1 + \lambda_2) e^{-(\lambda_1 + \lambda_2) x_1} dx_1 \quad x > 0$$

$$x \leq 0$$

$$\therefore f_{X_1 | X_1 < X_2}(x) = (\lambda_1 + \lambda_2) e^{-(\lambda_1 + \lambda_2) x}$$

$$\therefore X_1 | X_1 < X_2 \sim E(\lambda_1 + \lambda_2)$$

$$\therefore E(X_1 | X_1 < X_2) = \frac{1}{\lambda_1 + \lambda_2}$$

$$49. \begin{cases} U = X + Y \\ V = \frac{X}{Y} \end{cases} \Leftrightarrow \begin{cases} X = \frac{UV}{1+V} \\ Y = \frac{U}{1+V} \end{cases}$$

$$f_{X,Y}(x,y) = e^{-x} e^{-y} = e^{-x+y}$$

$$f_{U,V}(u,v) = f_{X,Y}(x(u,v), y(u,v)) \left| \frac{\partial(x,y)}{\partial(u,v)} \right|$$

$$= e^{-u} \frac{u}{(1+v)^2} = \frac{ue^{-u}}{(1+v)^2} \quad u, v > 0$$

$$\therefore f_{U,V}(u,v) = \begin{cases} \frac{ue^{-u}}{(1+v)^2} & u, v > 0 \\ 0 & \text{其他} \end{cases}$$

$$f_U(u) = \int_{-\infty}^{+\infty} f_{U,V}(u,v) dv = \int_0^{+\infty} ue^{-u} \left(-\frac{1}{1+v} \right) \Big|_0^{+\infty} = ue^{-u} \quad u > 0$$

$$f_V(v) = \int_{-\infty}^{+\infty} f_{U,V}(u,v) du = \int_0^{+\infty} \frac{1}{(1+v)^2} (-e^{-u}) \Big|_0^{+\infty} = \frac{1}{(1+v)^2} \quad v > 0$$

$$\therefore f_{U,V}(u,v) = f_U(u) f_V(v)$$

$$\therefore U, V \text{ 独立}$$



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$$\begin{cases} U+V=X+Y \\ V-U=|X-Y| \end{cases}$$

$$X > Y \text{ 时, } \begin{cases} V=X \\ U=Y \end{cases} \Rightarrow U < V \text{ 且 } \left| \frac{\partial(x,y)}{\partial(u,v)} \right| = 1$$

$$X \leq Y \text{ 时, } \begin{cases} V=Y \\ U=X \end{cases} \Rightarrow U \leq V \text{ 且 } \left| \frac{\partial(x,y)}{\partial(u,v)} \right| = 1$$

$$\begin{aligned} \therefore f_{U,V}(u,v) &= f_{X,Y}(v,u) \times 1 + f_{X,Y}(u,v) \times 1 \\ &= \frac{12}{7} v(u+v) + \frac{12}{7} u(u+v) \\ &= \frac{12}{7} (u+v)^2 \quad u < v \end{aligned}$$

$$\therefore f_{U,V}(u,v) = \begin{cases} \frac{12}{7} (u+v)^2 & u < v \\ 0 & \text{其他} \end{cases}$$

$$\begin{aligned} 51. \begin{cases} Y_1 = (-2 \ln X_1)^{\frac{1}{2}} \cos 2\pi X_2 \\ Y_2 = (-2 \ln X_2)^{\frac{1}{2}} \sin 2\pi X_2 \end{cases} &\Rightarrow \begin{cases} X_1 = e^{-\frac{1}{2}(Y_1^2 + Y_2^2)} \\ X_2 = \frac{1}{2\pi} \arctan \frac{Y_2}{Y_1} \end{cases} \end{aligned}$$

$$f_{X_1, X_2}(x_1, x_2) = 1 \quad 0 \leq x_1, x_2 \leq 1$$

$$\begin{aligned} \therefore f_{Y_1, Y_2}(y_1, y_2) &= \left| \frac{\partial(x_1, x_2)}{\partial(y_1, y_2)} \right| = \left| \begin{vmatrix} -\frac{y_1}{2} e^{-\frac{1}{2}(y_1^2 + y_2^2)} & -\frac{y_2}{2} e^{-\frac{1}{2}(y_1^2 + y_2^2)} \\ \frac{1}{2\pi} \frac{-y_2}{y_1^2 + y_2^2} & \frac{1}{2\pi} \frac{y_1}{y_1^2 + y_2^2} \end{vmatrix} \right| \\ &= \frac{1}{2\pi} e^{-\frac{1}{2}(y_1^2 + y_2^2)} \end{aligned}$$

$$\therefore Y_1, Y_2 \sim N(0, 0, 1, 1)$$

$$\therefore Y_1 \sim N(0, 1)$$

$$Y_2 \sim N(0, 1)$$

$$\therefore \rho = 0$$

$$\therefore Y_1, Y_2 \text{ 独立}$$

$$53. \text{ 设 } Y = \max(X_1, \dots, X_n)$$

$$\begin{aligned} F_Y(y) &= P(\max(X_1, \dots, X_n) \leq y) \\ &= P(X_1 \leq y, X_2 \leq y, \dots, X_n \leq y) \\ &= \prod_{i=1}^n P(X_i \leq y) \end{aligned}$$

$$\begin{aligned} F_X(x) &= \int_{-\infty}^x f(x) dx \\ &= \begin{cases} 0 & x < 0 \\ x^2 & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases} \end{aligned}$$

$$\begin{aligned} \therefore F_Y(y) &= \prod_{i=1}^n F_X(x) \\ &= (F_X(x))^n \end{aligned}$$

$$= \begin{cases} 0 & x < 0 \\ x^{2n} & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$

$$f_Y(y) = \begin{cases} 2n x^{2n-1} & 0 < x < 1 \\ 0 & \text{其他} \end{cases}$$

$$G_n(x) = P(Z_n \leq x)$$

$$= P(n(1-Y) \leq x) = P(Y \geq 1 - \frac{x}{n})$$

$$= 1 - P(Y < 1 - \frac{x}{n}) = \int_0^{1 - (1 - \frac{x}{n})^n} 1 - (1 - \frac{x}{n})^n \quad \begin{matrix} x > n \\ 0 < x \leq n \\ x \leq 0 \end{matrix}$$

$$\begin{aligned} n \rightarrow +\infty \text{ 时} \\ G_n(x) &\rightarrow \begin{cases} 0 & x \leq 0 \\ 1 - (1 - \frac{x}{n})^n = 1 - (1 - \frac{x}{n})^{\frac{n}{x} \cdot x} \rightarrow 1 - e^{-x} \end{cases} \end{aligned}$$



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$$\therefore \lim_{n \rightarrow \infty} G_n(x) = \begin{cases} 0 & x \leq 0 \\ 1 - e^{-2x} & x > 0 \end{cases} = \begin{cases} 0 & x < 0 \\ 1 - e^{-2x} & x \geq 0 \end{cases}$$

$$\therefore G(x) \sim E(2)$$

$$\therefore G(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-2x} & x \geq 0 \end{cases}$$

$$\therefore \lim_{n \rightarrow \infty} G_n(x) = G(x)$$

补充题: $(X, Y) \sim N(0, 0, \rho, 1)$ 求 $D(XY)$

$$D(XY) = E(X^2Y^2) - (E(XY))^2$$

由二维正态分布性质

$$Y|X=x \sim N(\rho x, 1-\rho^2)$$

$$\therefore E(Y|X) = \rho X \quad D(Y|X) = 1-\rho^2 \quad E(Y^2|X) = D(Y|X) + (E(Y|X))^2 = 1-\rho^2 + \rho^2 X^2$$

$$E(XY) = E(E(XY|X)) = E(XE(Y|X)) = E(X\rho X) = \rho EX^2 = \rho(DX + (EX)^2) = \rho DX = \rho$$

$$E(X^2Y^2) = E(E(X^2Y^2|X)) = E(X^2E(Y^2|X)) = E(X^2(1-\rho^2 + \rho^2 X^2)) = E((1-\rho^2)X^2 + \rho^2 X^4)$$

$$= (1-\rho^2)E(X^2) + \rho^2 E(X^4) = 1-\rho^2 + 3\rho^2 = 1+2\rho^2$$

