



班级: 自93

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科目: 随机数学53第

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19. 记 $\begin{cases} X_1: \text{第一份报纸发出时的过路人数} \\ X_i: \text{第 } i-1 \text{ 份报纸发出到第 } i \text{ 份报纸发出时的过路人数} \end{cases}$

则由题意可知 $X_i \sim Ge(\frac{1}{3})$ $EX_i = \frac{1}{p} = 3$ $DX_i = \frac{q}{p^2} = 6$
令 $Y = \sum_{i=1}^{100} X_i$ $EY = 100EX_i = 300$, $DY = 600$

由 Levy-Lindeberg 中心极限定理

$$S_n^* = \frac{Y - EY}{\sqrt{DY}} \xrightarrow{D} N(0, 1)$$

$$\therefore P(|Y - 300| \leq 20) = P\left(\left|\frac{Y - 300}{\sqrt{600}}\right| \leq \frac{20}{\sqrt{600}}\right)$$

$$= P\left(\left|\frac{Y - EY}{\sqrt{DY}}\right| \leq \frac{2}{\sqrt{6}}\right) = 2\Phi\left(\frac{2}{\sqrt{6}}\right) - 1 \approx 0.588$$

20. (a) $Y_i \stackrel{iid}{\sim} \begin{pmatrix} 1 & -1 \\ \frac{1}{2} & \frac{1}{2} \end{pmatrix}$ $EY_i = 0$, $DY_i = EY_i^2 = 1$

$$X_n = \sum_{i=1}^n Y_i \quad EX_n = nEY_i = 0$$

$$DX_n = D\left(\sum_{i=1}^n Y_i\right) = nDY_i = n$$

(b) 由 Levy-Lindeberg 中心极限定理:

$$n \rightarrow \infty \text{ 时, } \frac{X_n - EX_n}{\sqrt{DX_n}} \rightarrow N(0, 1)$$

$$\therefore \lim_{n \rightarrow \infty} P\left(\frac{X_n}{\sqrt{n}} \leq x\right) = \lim_{n \rightarrow \infty} P\left(\frac{X_n - EX_n}{\sqrt{DX_n}} \leq x\right)$$

$$= \Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt$$

9. $X \sim U[-1, 1]$ $X_i \stackrel{iid}{\sim} U[-1, 1]$

$$E\bar{X} = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n EX_i = 0$$

$$D\bar{X} = D\left(\frac{1}{n} \sum_{i=1}^n X_i\right)$$

$$= \frac{DX_i}{n} = \frac{1}{n} \cdot \frac{(1+1)^2}{12}$$

$$= \frac{1}{3n}$$

$$10. S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

$$= \frac{1}{n-1} \sum_{i=1}^n (X_i^2 - 2\bar{X}X_i + \bar{X}^2)$$

$$= \frac{1}{n-1} \left(\sum_{i=1}^n X_i^2 - 2\bar{X} \sum_{i=1}^n X_i + \sum_{i=1}^n \bar{X}^2 \right)$$

$$= \frac{1}{n-1} \left(\sum_{i=1}^n X_i^2 - 2n\bar{X} \cdot \frac{1}{n} \sum_{i=1}^n X_i + n\bar{X}^2 \right)$$

$$= \frac{1}{n-1} \left(\sum_{i=1}^n X_i^2 - 2n\bar{X}^2 + n\bar{X}^2 \right)$$

$$= \frac{1}{n-1} \left(\sum_{i=1}^n X_i^2 - n\bar{X}^2 \right) \quad \textcircled{1}$$

$$\therefore \frac{1}{n(n-1)} \sum_{i < j} (X_i - X_j)^2$$

$$= \frac{1}{n(n-1)} \sum_{j < i} (X_j - X_i)^2$$

$$= \frac{1}{n(n-1)} \sum_{i < j} (X_i - X_j)^2$$

$$\text{且 } \sum_{i < j} (X_i - X_j)^2 = 0$$

$$\therefore \frac{1}{n(n-1)} \sum_{i < j} (X_i - X_j)^2$$

$$= \frac{1}{n(n-1)} \cdot \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n (X_i - X_j)^2$$

$$= \frac{1}{2n(n-1)} \sum_{i=1}^n \sum_{j=1}^n (X_i^2 - 2X_iX_j + X_j^2)$$





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$$\begin{aligned}
&= \frac{1}{2n(n-1)} \sum_{i=1}^n \left(n x_i^2 - 2 x_i \sum_{j=1}^n x_j + \sum_{j=1}^n x_j^2 \right) \\
&= \frac{1}{2n(n-1)} \left(n \sum_{i=1}^n x_i^2 - 2 \sum_{i=1}^n x_i \sum_{j=1}^n x_j + n \sum_{j=1}^n x_j^2 \right) \\
&= \frac{1}{2n(n-1)} \left(2n \sum_{i=1}^n x_i^2 - 2 \left(\sum_{i=1}^n x_i \right)^2 \right) \\
&= \frac{1}{n-1} \left(\sum_{i=1}^n x_i^2 - n \left(\frac{1}{n} \sum_{i=1}^n x_i \right)^2 \right) \\
&= \frac{1}{n-1} \left(\sum_{i=1}^n x_i^2 - n \bar{x}^2 \right) \quad (2)
\end{aligned}$$

由 (1) 得

$$S^2 = \frac{1}{n(n-1)} \sum_{i=1}^n (x_i - \bar{x})^2$$

24. (1) 令 $y_i = \frac{x_i - 8}{2}$ $i=1, 2, \dots, 16$

~~且 x_1, x_2, \dots, x_{16} 是来自 $N(0, 1)$ 的样本~~

$$\begin{aligned}
P(X_{(16)} > 10) &= 1 - P(X_{(16)} \leq 10) \\
&= 1 - P(X_1 \leq 10, X_2 \leq 10, \dots, X_{16} \leq 10) \\
&= 1 - \prod_{i=1}^{16} P(X_i \leq 10) \\
&= 1 - \left(\frac{16}{16} \right)^{16} P\left(\frac{X_1 - 8}{2} \leq 1\right) \\
&= 1 - (\Phi(1))^{16} \\
&= 1 - (0.8413)^{16} \\
&\approx 0.9370
\end{aligned}$$

$$(2) P(X_{(1)} > 5)$$

$$= P(X_1 > 5, X_2 > 5, \dots, X_{16} > 5)$$

$$\begin{aligned}
&= \prod_{i=1}^{16} P(X_i > 5) \\
&= \left(\frac{16}{16} \right)^{16} (1 - P(X_1 \leq 5)) \\
&= \left(\frac{16}{16} \right)^{16} (1 - P\left(\frac{X_1 - 8}{2} \leq -\frac{3}{2}\right)) \\
&= \left(\frac{16}{16} \right)^{16} (1 - \Phi(-\frac{3}{2})) \\
&= \left(\frac{16}{16} \right)^{16} \Phi\left(\frac{3}{2}\right) = \Phi\left(\frac{3}{2}\right) \\
&\approx 0.3308
\end{aligned}$$

29.

$$(1) f_{X_{(6)}}(x) = \frac{10!}{5!4!} [F(x)]^5 [1-F(x)]^4 f(x)$$

$$\text{记 } Y = F(X_{(6)}) \text{ 则 } X_{(6)} = F^{-1}(Y)$$

$$\begin{aligned}
f_Y(y) &= \frac{10!}{5!4!} [F(F^{-1}(y))]^5 [1-F(F^{-1}(y))]^4 f(F^{-1}(y)) \cdot \frac{1}{f(x)} \\
&= \frac{10!}{5!4!} y^5 (1-y)^4 \\
&= \frac{P(11)}{P(6)P(5)} y^{6-1} (1-y)^{5-1} \quad 0 \leq y < 1
\end{aligned}$$

$$\therefore Y = F(X_{(6)}) \sim \text{Beta}(6, 5)$$

$$\therefore EY = \frac{a}{a+b} = \frac{6}{11} \quad DY = \frac{ab}{(a+b)^2(a+b+1)} = \frac{30}{11^2 \times 12} = \frac{5}{242}$$

$$(2) F_Y(0.15) = \int_0^{0.15} \frac{10!}{5!4!} y^5 (1-y)^4 dy \approx 0.00138$$

$$32. f_{X_{(4)}}(x) = \frac{5!}{3!1!1!} (F(x))^3 (1-F(x)) f(x)$$

$$\begin{aligned}
f_{X_{(2)}, X_{(4)}}(x, y) &= 5! F(x) [F(y) - F(x)] [1 - F(y)] f(x) f(y) \quad x \leq y \\
\text{令 } \begin{cases} U = \frac{X_{(2)}}{X_{(4)}} \\ V = X_{(4)} \end{cases} &\Rightarrow \begin{cases} X_{(2)} = UV \\ X_{(4)} = V \end{cases} \quad \frac{\partial(x_{(2)}, x_{(4)})}{\partial(u, v)} = \begin{vmatrix} V & U \\ 0 & 1 \end{vmatrix} = V
\end{aligned}$$





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$$f_{U,V}(u,v) = 5! F(uv) [F(v) - F(uv)] [1 - F(v)] p(uv) p(v) v \quad uv \leq v$$

$$p(x) = \begin{cases} 3x^2 & 0 < x < 1 \\ 0 & \text{其他} \end{cases} \quad F(x) = \begin{cases} 0 & x < 0 \\ x^3 & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$

$$f_{X(4)}(x) = \begin{cases} 20x^9(1-x^3)3x^2 = 60x^{11}(1-x^3) & 0 \leq x < 1 \\ 0 & \text{其他} \end{cases}$$

$$\therefore f_{U,V}(u,v) = 120 (uv)^3 [v^3 - (uv)^3] [1 - v^3] 3u^2v^2 \cdot 3v^2v \\ = 1080 u^5(1-u^3) \cdot v''(1-v^3) \quad \begin{cases} 0 < v < 1 \\ 0 < u < 1 \end{cases}$$

$$f_U(u) = \int_0^1 1080 u^5(1-u^3) v''(1-v^3) dv$$

$$= 1080 u^5(1-u^3) \int_0^1 v''(1-v^3) dv$$

$$= 1080 u^5(1-u^3) \frac{1}{60}$$

$$= 18 u^5(1-u^3) \quad 0 < u < 1$$

$$\therefore \frac{X_{(2)}}{X_{(4)}} \sim 18 u^5(1-u^3) \quad 0 < u < 1$$

$$X_{(4)} \sim 60 v^{11}(1-v^3) \quad 0 < v < 1$$

$$\text{且联合分布 } \left(\frac{X_{(2)}}{X_{(4)}}, X_{(4)} \right) \sim 1080 u^5(1-u^3) v^{11}(1-v^3) \quad \begin{cases} 0 < v < 1 \\ 0 < u < 1 \end{cases}$$

$$\therefore f_{\left(\frac{X_{(2)}}{X_{(4)}}, X_{(4)} \right)}(u,v) = f_{\frac{X_{(2)}}{X_{(4)}}}(u) \cdot f_{X_{(4)}}(v)$$

$$\therefore \frac{X_{(2)}}{X_{(4)}} \text{ 与 } X_{(4)} \text{ 独立.}$$

$$8. \quad Z = \frac{n}{m} X / \left(1 + \frac{n}{m} X \right) = \frac{1}{1 + \frac{m}{n} \cdot \frac{1}{X}}$$

$$\text{令 } Y = \frac{1}{X} \text{ 则 } Y \sim F(m, n), \text{ 故 } Z = \frac{1}{1 + \frac{m}{n} Y} \quad Y = \frac{n}{m} \left(\frac{1}{Z} - 1 \right) \quad \frac{dy}{dz} = -\frac{n}{m} \cdot \frac{1}{z^2}$$

$$\text{由 } F \text{ 分布的概率密度 } f_Y(y) = \frac{\Gamma(\frac{m+n}{2}) (\frac{m}{n})^{\frac{m}{2}}}{\Gamma(\frac{m}{2}) \Gamma(\frac{n}{2})} y^{\frac{m}{2}-1} \left(1 + \frac{m}{n} y \right)^{-\frac{m+n}{2}}, \quad y > 0$$



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$$\therefore f_Z(z) = \frac{\Gamma(\frac{m+n}{2}) (\frac{m}{n})^{\frac{m}{2}}}{\Gamma(\frac{m}{2}) \Gamma(\frac{n}{2})} \left(\frac{n}{m}\right)^{\frac{m}{2}-1} \left(\frac{1}{z}-1\right)^{\frac{m}{2}-1} z^{\frac{m+n}{2}} \cdot \frac{n}{m} \cdot \frac{1}{z^2} \quad 0 < z < 1$$

$$= \frac{\Gamma(\frac{m+n}{2})}{\Gamma(\frac{m}{2}) \Gamma(\frac{n}{2})} \frac{m}{n} (1-z)^{\frac{m}{2}-1} z^{\frac{n}{2}+1} \cdot \frac{n}{m} \cdot \frac{1}{z^2} \quad 0 < z < 1$$

$$= \frac{\Gamma(\frac{m+n}{2})}{\Gamma(\frac{m}{2}) \Gamma(\frac{n}{2})} z^{\frac{n}{2}-1} (1-z)^{\frac{m}{2}-1} \quad 0 < z < 1$$

$$\therefore Z \sim \text{Beta}\left(\frac{n}{2}, \frac{m}{2}\right)$$

$$19. T = -2 \sum_{i=1}^n \ln F(x_i) = \sum_{i=1}^n \ln \frac{1}{F^2(x_i)}$$

$$\text{令 } Y = \ln \frac{1}{F^2(x)} \quad x_1$$

$$F_Y(y) = P(Y \leq y) = P\left(\ln \frac{1}{F^2(x)} \leq y\right) = P\left(\frac{1}{F^2(x)} \leq e^y\right) = P(F^2(x) \geq e^{-y})$$

$$= P(F(x) \geq e^{-\frac{y}{2}}) = 1 - P(F(x) \leq e^{-\frac{y}{2}}) = 1 - P(X < F^{-1}(e^{-\frac{y}{2}})) = 1 - F(F^{-1}(e^{-\frac{y}{2}}))$$

$$= 1 - e^{-\frac{y}{2}} \quad \therefore Y \sim E(1/2) \quad \text{即 } Y \sim G(1, 1/2)$$

由Gamma分布的可加性(可由特征函数证明) 且 X_1, \dots, X_n 相互独立.

$$T = \sum_{i=1}^n \ln \frac{1}{F^2(x_i)} \sim G(n, 1/2)$$

$$\therefore T \sim \chi^2(2n)$$

