

# Chapter 1.7 Binomial, Poisson and Geometric Distributions

1.7.1 Binomial Distribution

1.7.2 Geometric Distribution

1.7.3 Poisson Distribution

1.7.4 Summary

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## 1.7.1

# Binomial Distribution

The **Binomial distribution** with parameters  $n$  and  $p$  is the discrete probability distribution of the number of successes in a sequence of  $n$  independent trials, each of which yields success with probability  $p$ .

We write  $X \sim \text{Bin}(n, p)$  for a random variable that has Binomial distribution with  $n$  trials and probability of success  $p$ . The outcome for each trial is success/failure.  $X$  takes the value of  $0, 1, 2, \dots, n$ .

## 1.7.1 Binomial Distribution: example

Examples include tossing a fair coin  $n$  times. In each trial, we have 2 outcomes: success or failures. The probability of success does not change from one trial to another.

The probability of any combination of 2 heads is  $\left(\frac{1}{2}\right)^2 \left(1 - \frac{1}{2}\right)^{n-2}$

There are  $\binom{n}{2}$  possible combinations of 2 heads in  $n$  tosses, therefore

$$P(2H|n \text{ tosses}) = \binom{n}{2} \left(\frac{1}{2}\right)^2 \left(1 - \frac{1}{2}\right)^{n-2} = \binom{n}{2} \left(\frac{1}{2}\right)^n$$

## 1.7.1

# Binomial Distribution

In general if  $p$  is the probability of an event (we call this event a *success*) and we repeat  $n$  independent trials, the of getting exactly  $k$  successes in  $n$  trials is

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}; \quad k = 0, 1, \dots, n$$

where  $X$  denotes the number of successes and is called Binomial variable.

## 1.7.1

# Binomial Distribution

If  $X$  is Binomial  
its mean is

$$\mu = E(X) = np$$

and the variance is

$$\sigma^2 = Var(X) = np(1 - p) = (1 - p)E(X)$$

**Note** We usually denote  $q = 1 - p$ .

## 1.7.1

# Binomial Distribution

### Example 1

Compute the probability of obtaining at least two '6' in rolling a fair die 4 times. [Hint:  $p = \frac{1}{6}, n = 4, k = 2$ ]

### Solution

Let  $X \sim \text{Bin}\left(4, \frac{1}{6}\right)$ . Required probability is

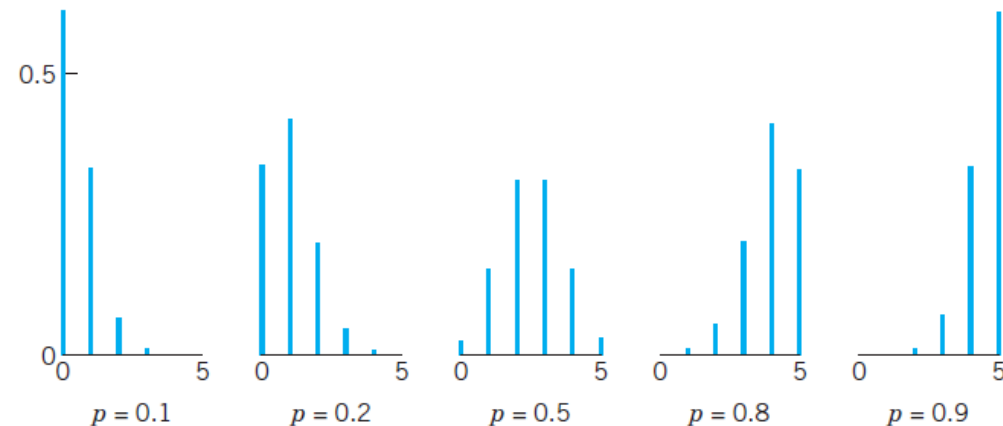
$$\begin{aligned} P(X \geq 2) &= \sum_{k=2}^4 \binom{4}{k} p^k (1-p)^{n-k} = P(X = 2) + P(X = 3) + P(X = 4) = \\ &= \binom{4}{2} \left(\frac{1}{6}\right)^2 \left(\frac{5}{6}\right)^2 + \binom{4}{3} \left(\frac{1}{6}\right)^3 \left(\frac{5}{6}\right)^1 + \binom{4}{4} \left(\frac{1}{6}\right)^4 \left(\frac{5}{6}\right)^0 = \frac{171}{1296} \\ &= 0.1320 \quad \blacksquare \end{aligned}$$

## 1.7.1

# Binomial Distribution

### Example 2

If  $X \sim \text{Bin}(n, p)$  then as  $k$  goes from 0 to  $n$ , the probability  $P(X = k)$  will first increase monotonically and then decrease monotonically, reaching its largest value when  $k$  is the largest integer less than or equal to  $(n + 1)p$ .



Probability function (2) of the binomial distribution for  $n = 5$  and various values of  $p$

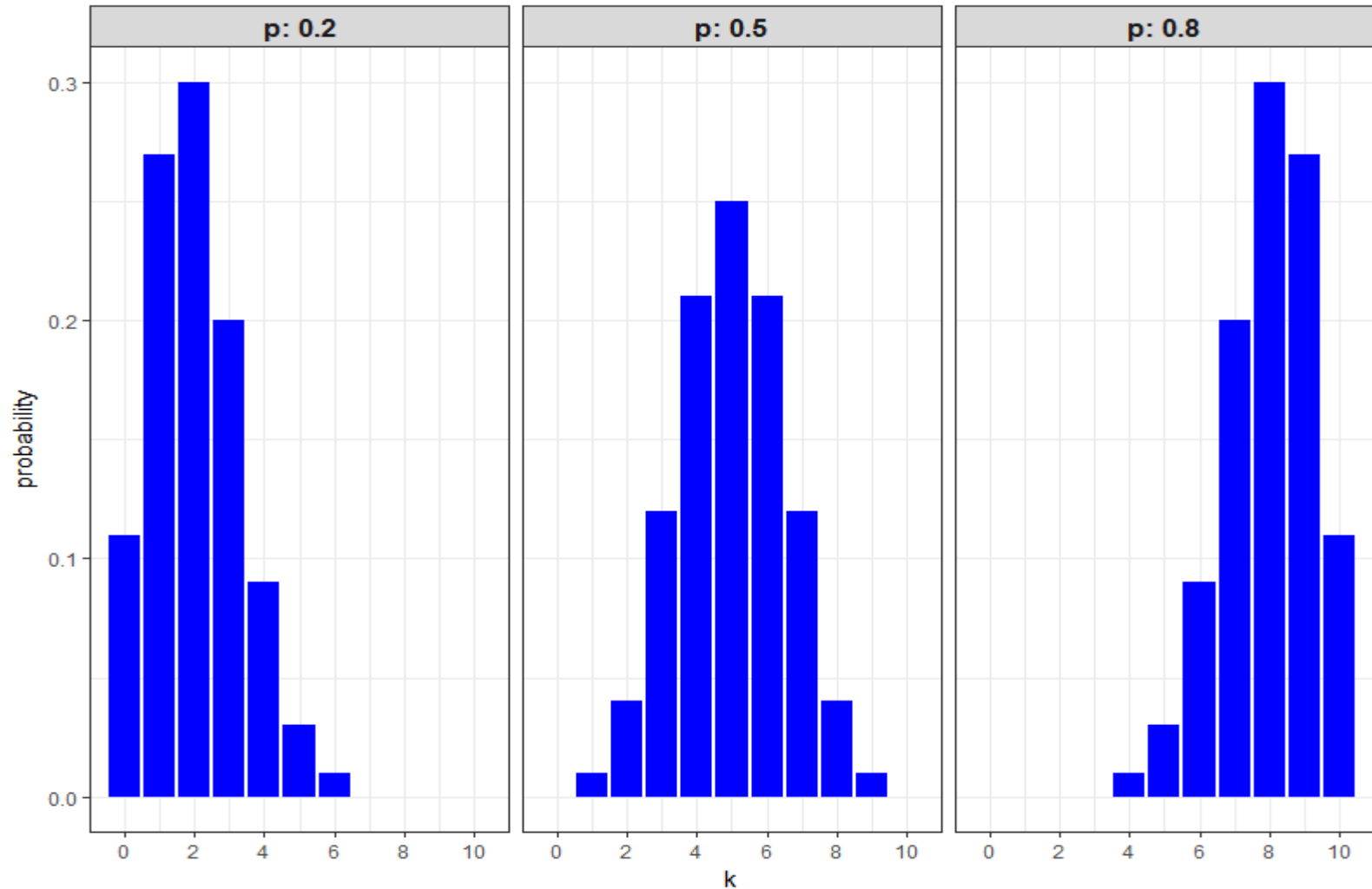
## 1.7.1

# Binomial Distribution

The Binomial distribution is symmetric for  $p = 0.5$  and skewed for other values.

Example  $n = 10$

$$n + 1 = 11$$





## 1.7.1

# Binomial Distribution

### Example 3

The random variable  $X \sim \text{Bin}(n, p)$  where  $0 < p < 1$ . Given that  $\text{Var}(X) = \frac{1}{5}E(X)$ , find the least value of  $n$  such that  $P(X \geq 1) > 0.95$ .

### Solution

Solving  $\text{Var}(X) = np(1 - p) = \frac{1}{5}E(X) = \frac{1}{5}np$ , we have  $p = \frac{4}{5}$ . So, we need  $n$  for which  $P(X \geq 1) = 1 - P(X = 0) > 0.95$ .

## 1.7.1

## Binomial Distribution

$$\therefore P(X = 0) = \left(\frac{1}{5}\right)^n < 0.05$$

Taking log we obtain  $n \ln\left(\frac{1}{5}\right) > \ln(0.05)$ ; rearranging, we have

$$n > \frac{\ln(0.05)}{\ln\left(\frac{1}{5}\right)} = 1.86$$

Hence minimum  $n$  is 2. ■

[for computing, recall that  $\ln(0.05) = \ln\left(\frac{1}{20}\right) = -\ln(20)$  and  $\ln\left(\frac{1}{5}\right) = -\ln(5)$ . So,  $\frac{\ln(0.05)}{\ln(1/5)} = \ln(20) / \ln(5)$ ]

## 1.7.1 Binomial Distribution: problem

- There are 20 racers. The probability that one finishes the race is 0.9. Find the probability that
  1. 12 racers finish the race
  2. Less than 19 finish the race

## 1.7.2

# Geometric Distribution

The **geometric distribution** is a discrete probability distribution of the number  $X$  of independent trials needed in order to get the **first success**. We assume that each trial has constant probability  $p \in (0,1)$  of a success.

So, let's denote  $S = \text{success}$  and  $F = \text{failure}$ , it is the probability of  $k - 1$  failures before a success.

For example, if  $X = \text{number of failures}$  and  $p = \text{prob. of success}$ ,  $q = 1 - p = \text{prob. of failure}$ ,

$$P(X = 4) = P(FFFS) = P(F)P(F)P(F)P(S) = q^3p$$

There is only one possible order: first 3 failures and then 1 success

## 1.7.2

## Geometric Distribution

We write  $X \sim \text{Geo}(p)$  for a random variable that has Geometric distribution which is the number of trials before the *1<sup>st</sup> success*.  $X$  takes the value of  $1, 2, \dots$ .

$X = 1$ , means success at the first try,  $X = 2$  at the second try, and so on.

## 1.7.2

## Geometric Distribution

To have the first success at the  $k^{th}$  trial, we need to have first  $(k - 1)$  failures and then a success. Therefore, letting  $q = 1 - p$ , the probability of getting the 1<sup>st</sup> success in  $k^{th}$  trials is

$$P(X = k) = p(1 - p)^{k-1} = pq^{k-1}; \quad k = 1, 2, \dots$$

The cdf is found as

$$P(X \leq k) = \sum_{i=1}^k (1 - p)^{i-1} p = p \sum_{i=0}^{k-1} q^i = p \frac{1 - q^k}{1 - q} = 1 - q^k$$

Recall that  $\sum_{i=0}^n r^i = \frac{1 - r^{n+1}}{1 - r}$

## 1.7.2

# Geometric Distribution

The expected number of trials for the first success is the mean

$$E(X) = \mu = \frac{1}{p}$$

The variance is  $\text{Var}(X) = \sigma^2 = \frac{1-p}{p^2}$ .

## 1.7.2 Geometric Distribution

The geometric distribution is *memoryless*:

$$P(X > k + i | X > i) = P(X > k)$$

**Proof.** We know that  $P(X \leq k) = 1 - q^k$ , so

$$P(X > k) = q^k$$

Hence,

$$P(X > k + i | X > i) = \frac{q^{k+1}}{q^i} = q^k = P(X > k)$$



## 1.7.2

# Geometric Distribution

### Example 7

A box contains  $N$  white and  $M$  black balls. Balls are selected randomly, one at a time and replaced, until a black one is obtained. If we assume that each ball is replaced before the next one is drawn, what is the probability that:

- i. exactly  $k$  draws are needed?
- ii. at least  $k$  draws are needed?

## 1.7.2

## Geometric Distribution

### Solution

The probability of a black ball is  $p = \frac{M}{M+N}$ , for white it is

$$q = 1 - p = \frac{N}{M+N}$$

i. Required probability is

$$P(X = k) = p(1-p)^{k-1} = \frac{MN^{k-1}}{(M+N)^k}.$$

ii. Required probability is

$$P(X \geq k) = 1 - P(X \leq k-1) = 1 - (1 - q^{k-1}) = q^{k-1}$$

## 1.7.3

# Poisson Distribution

The **Poisson distribution** is a discrete probability distribution that gives the probability of a given number of events occurring in a fixed interval of time and/or space if these events occur with a **known average rate** and **independently of the time since the last event**.

We write  $X \sim \text{Poi}(\lambda)$  for a random variable that has Poisson distribution with rate of occurring in a fixed interval of time and/or space.  $X$  takes the value of  $0, 1, 2, \dots$ .

## 1.7.3

# Poisson Distribution

### Example 4

The Poisson distribution may be useful to model events such as

1. The number of goals scored in a World Cup soccer match
2. The number of meteors greater than 1 meter diameter that strike earth per year
3. The number of occurrences of the DNA sequence "ACGT" in a gene
4. The number of patients arriving in an emergency room between 11 and 12 pm ■

## 1.7.3

# Poisson Distribution

### **When is the Poisson distribution an appropriate model?**

The Poisson distribution is appropriate if the following assumptions are true:

- $X$  is the number of times an event occurs in an interval and  $X$  can take values  $0, 1, 2, \dots$ .
- Events occur randomly, independently and singly
- The rate at which events occur is constant. The rate cannot be higher in some intervals and lower in other intervals.
- The probability of an event in an interval is proportional to the length of the interval.

If these conditions are true, then  $X$  is a Poisson random variable, and the distribution of  $X$  is a Poisson distribution.

## 1.7.3

# Poisson Distribution

### Example 5

**The Poisson assumptions are violated in the following cases:**

- The number of students who arrive at the lecture hall will likely not follow a Poisson distribution, because the **rate is not constant** (low rate during class time, high rate between class times) and the arrivals of individual students are not independent (students tend to come in groups).
- The number of magnitude 5 earthquakes per year in China may not follow a Poisson distribution if **one large earthquake increases the probability of aftershocks of similar magnitude**.
- Among patients admitted to the intensive care unit of a hospital, the number of days that the patients spend in the ICU is not Poisson distributed because **the number of days cannot be zero**. ■

### 1.7.3

## Poisson Distribution

The probability of getting exactly  $k$  events in a specified interval with rate  $\lambda$  is

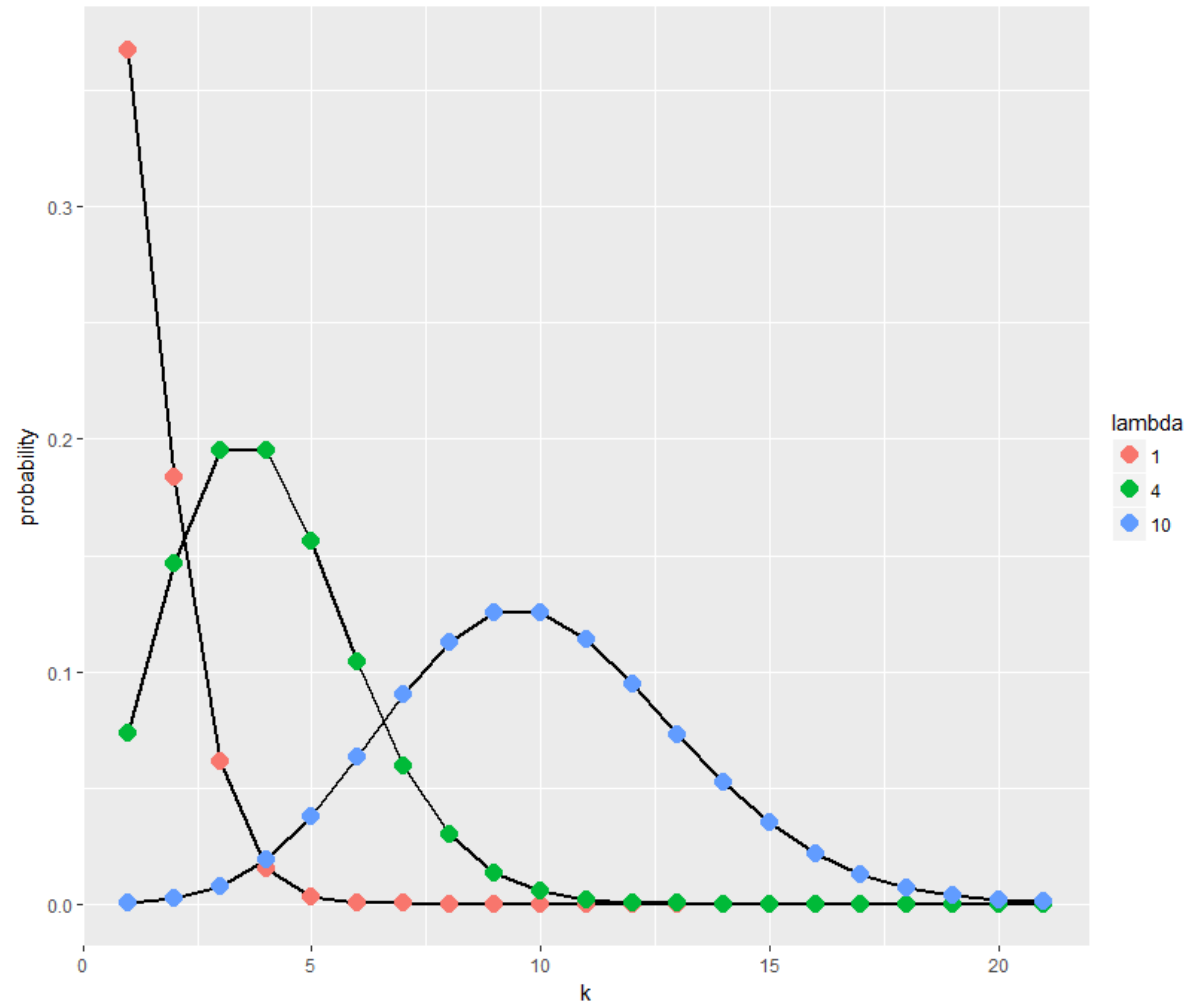
$$P(X = k) = \frac{e^{-\lambda} \lambda^k}{k!} \quad ; \quad k = 0, 1, 2, \dots$$

with mean  **$E(X) = \lambda$**  and variance  **$Var(X) = \lambda$** .

## 1.7.3

# Poisson Distribution

The shape of the Poisson pmf changes with  $\lambda$ , but converges for large  $\lambda$





### 1.7.3

## Poisson Distribution

The cdf of a Poisson is the sum of the probabilities

$$F(k) = P(X \leq k) = \sum_{i=0}^k \frac{e^{-\lambda} \lambda^k}{k!} ; \quad k = 0, 1, 2, \dots$$

$$F(0) = P(X = 0) = e^{-\lambda}, F(1) = e^{-\lambda} + \lambda e^{-\lambda} = e^{-\lambda}(1 + \lambda)$$

So, the probability of at least one event is

$$P(X > 0) = 1 - P(X = 0) = 1 - e^{-\lambda}$$

## 1.7.3

# Poisson Distribution

### Example 6

Suppose that earthquakes occur at the rate of 2 per week. By modeling this as a Poisson distribution, find:

- i. The probability of 3 earthquakes in the next week;
- ii. the probability that at least 2 earthquakes occur during the next week.

## 1.7.2 Poisson Distribution

### Solution

Since the rate is  $\lambda = 2$  earthquakes per week

i.  $X \sim \text{Poi}(2)$ , so  $P(X = 3) = \frac{e^{-2}2^3}{3!} = \frac{8e^{-2}}{6} = 0.18045$

ii. The required probability is

$$\begin{aligned} P(X \geq 2) &= 1 - P(X = 0) - P(X = 1) = \\ 1 - \frac{e^{-2}2^0}{0!} - \frac{e^{-2}2^1}{1!} &= 1 - e^{-2} - 2e^{-2} \\ &= 1 - 3e^{-2} \approx 0.594 \quad \blacksquare \end{aligned}$$

### 1.7.3

## Poisson Distribution: problem

Suppose that calls to a switchboard arrive as a Poisson variable with rate 2.5 per minute. Find the probability:

1. that in the next minute there won't be a call
2. That in the next minute there won't be more than 1 call.

Also find the average number of calls per minute.

## 1.7.3

# Poisson Distribution: problem

Suppose that calls to a switchboard arrive as a Poisson variable with rate 2.5 per minute. Find the probability:

1. that in the next minute there won't be a call

$$P(x = 0) = e^{-2.5} \approx 0.082$$

2. That in the next minute there won't be more than 1 call.

$$P(x \leq 1) = P(x = 0) + P(x = 1) = e^{-2.5} + 2.5e^{-2.5} \approx 0.288$$

Also find the average number of calls per minute.

The average number of calls is the mean of the variable which is equal to the rate. Hence  $E(X) = \lambda = 2.5$ .

## 1.7.4

## Summary

### 1.7.1 Binomial Distribution

$$P(X = k|n) = \binom{n}{k} p^k (1 - p)^{n-k}$$

### 1.7.2 Geometric Distribution

$$P(X = k) = p(1 - p)^{k-1}$$

### 1.7.3 Poisson Distribution

$$P(X = k|\lambda) = \frac{e^{-\lambda} \lambda^k}{k!}$$