

Q1:Bayesian Inference

(a)

F represents Fair coin; H represents Head.

$$P(F) = 0.5; P(\neg F) = 0.5; P(H|F) = 0.5; P(H|\neg F) = 0.5$$

Therefore

$$\begin{aligned} P(F|H) &= \frac{P(F \cdot H)}{P(H)} = \frac{P(H|F)P(F)}{P(H|F)P(F) + P(\neg F)P(H|\neg F)} \\ &= \frac{(0.5 \times 0.5)}{(0.5 \times 0.5 + 0.5 \times 0.8)} \approx 0.38 \end{aligned}$$

(b)

Now, the probability of the coin being fair is updated to 0.38. Then, recompute the posterior probability of the coin being fair.

$$\begin{aligned} P(F|H) &= \frac{P(F \cdot H)}{P(H)} = \frac{P(H|F)P(F)}{P(H|F)P(F) + P(\neg F)P(H|\neg F)} \\ &= \frac{(0.5 \times 0.38)}{(0.5 \times 0.38 + (1 - 0.38) \times 0.8)} \approx 0.28 \end{aligned}$$

(c)

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In [13]: Pr_F=0.28
Pr_Not_F=1-Pr_F
Pr_H_given_F=0.5
Pr_H_given_not_F=0.8
n=0
while Pr_F>=0.05:
    Pr_F=(Pr_H_given_F*Pr_F)/(Pr_H_given_F*Pr_F+Pr_Not_F*Pr_H_given_not_F)
    Pr_Not_F=1-Pr_F
    n=n+1
print(n)
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Therefore, after 5 more flips, my belief about the probability of the coin being fair drops to below 0.05.

Q2: Fun with Linear Regression

(a)(1)

Because

$$\epsilon \sim N(0, \sigma)$$

Then, we can figure out

$$y^{(i)} | x^{(i)} \sim N(\omega^T x^{(i)}, \sigma)$$

Compute

$$\begin{aligned} L(\omega) &:= \prod_{i=1}^n P(y^{(i)} | x^{(i)}, \sigma, \omega) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{-(y^{(i)} - \omega^T x^{(i)})^2}{2\sigma^2}} \end{aligned}$$

(a)(2)

$$\begin{aligned}
LL(\omega) &:= \ln L(\omega) \\
&= \ln \left(\prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y^{(i)} - \omega^T x^{(i)})^2}{2\sigma^2}} \right) \\
&= \sum_{i=1}^n \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y^{(i)} - \omega^T x^{(i)})^2}{2\sigma^2}} \right) \\
&= \sum_{i=1}^n \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right) + \sum_{i=1}^n \ln \left(e^{-\frac{(y^{(i)} - \omega^T x^{(i)})^2}{2\sigma^2}} \right) \\
&= n \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y^{(i)} - \omega^T x^{(i)})^2
\end{aligned}$$

Remove constant and terms that don't have ω

Therefore,

$$\begin{aligned}
\hat{\omega} &= \underset{\omega}{\operatorname{argmax}} - \sum_{i=1}^n (y^{(i)} - \omega^T x^{(i)})^2 \\
&= \underset{\omega}{\operatorname{argmin}} \sum_{i=1}^n (\omega^T x^{(i)} - y^{(i)})^2
\end{aligned}$$

When X represents a $(n \times p)$ matrix which has n data samples and p features, Y represents a $(n \times 1)$ matrix which has n results, $\min_{\omega} \|X\omega - Y\|_2^2$ yields the maximizer of the posterior.

(b)(1)

Because

$$\epsilon \sim N(0, \sigma)$$

$$\omega \sim N(0, \tau)$$

Then, we can compute posterior by following steps:

$$\begin{aligned}
 M(\omega) &:= \prod_{i=1}^n P(\omega|y^{(i)}, x^{(i)}) \\
 &= \prod_{i=1}^n \frac{P(y^{(i)}, x^{(i)}|\omega)g(\omega)}{h(y^{(i)}, x^{(i)})} \\
 &= \prod_{i=1}^n \frac{P(y^{(i)}|x^{(i)}, \omega)P(x^{(i)})g(\omega)}{h(y^{(i)}, x^{(i)})}
 \end{aligned}$$

Remove constant multipliers that don't have ω

$$\begin{aligned}
 M(\omega) &= \prod_{i=1}^n P(y^{(i)}|x^{(i)}, \omega)g(\omega) \\
 &= \prod_{i=1}^n L(\omega)g(\omega) \\
 &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y^{(i)} - \omega^T x^{(i)})^2}{2\sigma^2}} \cdot \frac{1}{\sqrt{2\pi\tau^2}} e^{-\frac{\omega^2}{2\tau^2}}
 \end{aligned}$$

(b)(2)

Compute the log-likelihood of the posterier:

$$\begin{aligned}
 m(\omega) &:= \ln M(\omega) \\
 &= \ln \left(\prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y^{(i)} - \omega^T x^{(i)})^2}{2\sigma^2}} \cdot \frac{1}{\sqrt{2\pi\tau^2}} e^{-\frac{\omega^2}{2\tau^2}} \right) \\
 &= \sum_{i=1}^n \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y^{(i)} - \omega^T x^{(i)})^2}{2\sigma^2}} \cdot \frac{1}{\sqrt{2\pi\tau^2}} e^{-\frac{\omega^2}{2\tau^2}} \right) \\
 &= \sum_{i=1}^n \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right) + \sum_{i=1}^n \ln \left(e^{-\frac{(y^{(i)} - \omega^T x^{(i)})^2}{2\sigma^2}} \right) + \sum_{i=1}^n \ln \left(\frac{1}{\sqrt{2\pi\tau^2}} \right) + \sum_{i=1}^n \ln \left(e^{-\frac{\omega^2}{2\tau^2}} \right) \\
 &= n \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right) + n \ln \left(\frac{1}{\sqrt{2\pi\tau^2}} \right) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y^{(i)} - \omega^T x^{(i)})^2 - \frac{n}{2\tau^2} \omega^2
 \end{aligned}$$

Remove constant and terms that don't have ω

Therefore,

$$\begin{aligned}
\omega_{MAP} &= \underset{\omega}{\operatorname{argmax}} -\frac{1}{2\sigma^2} \sum_{i=1}^n (y^{(i)} - \omega^T x^{(i)})^2 - \frac{n}{2\tau^2} \omega^2 \\
&= \underset{\omega}{\operatorname{argmin}} \frac{1}{2\sigma^2} \sum_{i=1}^n (y^{(i)} - \omega^T x^{(i)})^2 + \frac{n}{2\tau^2} \omega^2 \\
&= \underset{\omega}{\operatorname{argmin}} \sum_{i=1}^n (y^{(i)} - \omega^T x^{(i)})^2 + \frac{n\sigma^2}{\tau^2} \omega^2
\end{aligned}$$

When X represents a $(n \times p)$ matrix which has n data samples and p features, Y represents a $(n \times 1)$ matrix which has n results, $\lambda = \frac{n\sigma^2}{\tau^2}$, $\min_{\omega} ||X\omega - Y||_2^2 + \lambda ||\omega||_2^2$ yields the maximizer of the posterior.