1 The marginal likelihood

We will perform Type-II Maximum-Likelihood estimation so to calculate the Type-II MLE we wish to find the equation for the following marginal:

$$\mathbf{W}' = \arg\max_{\mathbf{W}} \int p(\mathbf{Y} \mid \mathbf{X}, \mathbf{W}) p(\mathbf{X}) d\mathbf{X}$$

We have the equation for the linear model as:

$$p(\mathbf{Y}, \mathbf{X}, \mathbf{W}) = p(\mathbf{Y} \mid \mathbf{X}, \mathbf{W}) p(\mathbf{X}) p(\mathbf{W})$$

We assume additive Gaussian noise for a linear non-parametric Gaussian process and a spherical prior so our likelihood and prior are as follows:

$$p(y_i|x_i, \mathbf{W}) = \mathcal{N}(y_i|\mathbf{W}x_i, \sigma^2 \mathbf{I})$$

 $p(x) = \mathcal{N}(0, \mathbf{I})$

Our marginalisation is the product of two Gaussians which is itself another Gaussian:

$$p(y_i|\mathbf{W}) = \int p(y_i|x_i, \mathbf{W}) p(x_i) dx$$
$$= \int \mathcal{N}(y_i|\mathbf{W}x_i, \sigma^2 \mathbf{I}) \mathcal{N}(0, \mathbf{I}) dx$$

By replacing these Gaussians with their probability density function and combining we can form a single exponential, by then completing the square with regard to x we can integrate x out to form a new exponent for the marginal which by re-arranging we can find the mean 0 and covariance $\mathbf{W}\mathbf{W}^T + \sigma^2\mathbf{I}$ thus making our marginal likelihood:

$$p(y_i|\mathbf{W}) = \mathcal{N}(y_i|0, \mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})$$
$$p(\mathbf{Y}|\mathbf{W}) = \prod_{i=1}^{N} \mathcal{N}(y_i|0, \mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})$$

2 The objective function

We have our marginal likelihood from above:

$$p(\mathbf{Y}|\mathbf{W}) = \prod_{i=1}^{N} \mathcal{N}(y_i|0, \mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})$$

From this we can evaluate, using log rules, our negative log likelihood objective function as follows:

$$\mathcal{L}(\mathbf{W}) = -\log\left(\prod_{i=1}^{N} \mathcal{N}(y_{i}|0, \mathbf{W}\mathbf{W}^{T} + \sigma^{2}\mathbf{I})\right)$$

$$= -\sum_{i=1}^{N} \log\left(\mathcal{N}(y_{i}|0, \mathbf{W}\mathbf{W}^{T} + \sigma^{2}\mathbf{I})\right)$$

$$= -\sum_{i=1}^{N} \log\left(\frac{1}{\sqrt{2\pi|\Sigma|}}\right) + \log\left(\exp\left(-\frac{1}{2}y_{i}^{T}\Sigma^{-1}y_{i}\right)\right)$$

$$= -\sum_{i=1}^{N} \log\frac{1}{(2\pi)^{\frac{D}{2}}|\mathbf{W}\mathbf{W}^{T} + \sigma^{2}\mathbf{I}|^{\frac{1}{2}}} + \log\left(\exp\left(-\frac{1}{2}y_{i}^{T}(\mathbf{W}\mathbf{W}^{T} + \sigma^{2}\mathbf{I})^{-1}y_{i}\right)\right)$$

$$= -\log\left((2\pi)^{D}|\mathbf{W}\mathbf{W}^{T} + \sigma^{2}\mathbf{I}|\right)^{-\frac{N}{2}} - \sum_{i=1}^{N} -\frac{1}{2}y_{i}^{T}(\mathbf{W}\mathbf{W}^{T} + \sigma^{2}\mathbf{I})^{-1}y_{i}$$

$$= \frac{N}{2}\left(D\log 2\pi + \log(|\mathbf{W}\mathbf{W}^{T} + \sigma^{2}\mathbf{I}|\right) + \frac{1}{2}\sum_{i=1}^{N} y_{i}^{T}(\mathbf{W}\mathbf{W}^{T} + \sigma^{2}\mathbf{I})^{-1}y_{i}$$

From the fact that $Tr(\mathbf{AB}) = Tr(\mathbf{BA})$ and $Tr(\mathbf{C} = \mathbf{c})$ for the dimensions in our sum we can evaluate it as follows:

$$\mathcal{L}(\mathbf{W}) = \frac{N}{2} \left(D \log 2\pi + \log(|\mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I}|) + \text{Tr}\left[\mathbf{Y}(\mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})^{-1}\mathbf{Y}^T\right] \right)$$

3 The derivative of the objective function

To find the gradient of \mathcal{L} we will look at each term in turn, we will make extensive use of [1] to calculate the derivatives of matrices.

$$\mathcal{L}(\mathbf{W}) = \frac{N}{2} \left(\underbrace{\operatorname{Tr} \left[(\mathbf{W} \mathbf{W}^T + \sigma^2 \mathbf{I})^{-1} \mathbf{Y} \mathbf{Y}^T \right]}_{A} + \underbrace{\log(|\mathbf{W} \mathbf{W}^T + \sigma^2 \mathbf{I}|)}_{B} + \underbrace{D \log 2\pi}_{C} \right)$$

C is a constant term so we shall discard that. To find the derivative of B we use the rules that $\partial(\log(\det(\mathbf{X}))) = \text{Tr}(\mathbf{X}^{-1}\partial\mathbf{X})$ and that $\sigma^2\mathbf{I}$ is constant with respect to \mathbf{W} and so evaluate B to:

$$\frac{\partial B}{\partial \mathbf{W}} = \text{Tr}\left((\mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})^{-1} \partial (\mathbf{W}\mathbf{W}^T) \right)$$

We also have the rules that $\partial(\mathbf{XY}) = (\partial \mathbf{X})\mathbf{Y} + \mathbf{X}(\partial \mathbf{Y})$ and that $\frac{\partial \mathbf{X}}{\partial \mathbf{X}_{ij}} = \mathbf{J}^{ij}$ where \mathbf{J} is the single-entry matrix, having 1 at (i,j) and 0 elsewhere and so can reduce B to:

$$\frac{\partial B}{\partial \mathbf{W}_{ij}} = \text{Tr}\left((\mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})^{-1} (\mathbf{J}^{ij}\mathbf{W}^T + \mathbf{W}\mathbf{J}^{ijT}) \right)$$

Using the rule that $\partial(\text{Tr}(\mathbf{X})) = \text{Tr}(\partial \mathbf{X})$, the chain rule and that $\frac{\partial \mathbf{XCX}^T}{\partial \mathbf{C}} = \mathbf{X}^T \mathbf{X}$ we can find the derivative of A:

$$\frac{\partial A}{\partial \mathbf{W}} = \text{Tr}\left(\partial \left(\mathbf{Y}(\mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})^{-1}\mathbf{Y}^T\right)\right)$$
$$= \text{Tr}\left(\mathbf{Y}^T\mathbf{Y}\partial \left((\mathbf{W}\mathbf{W}^T + \sigma^2 \mathbf{I})^{-1}\right)\right)$$

As $\partial(\mathbf{X}^{-1}) = -\mathbf{X}^{-1}(\partial\mathbf{X})\mathbf{X}^{-1}$ we can evaluate A to:

$$\frac{\partial A}{\partial \mathbf{W}_{ij}} = \text{Tr}\left(\mathbf{Y}^T \mathbf{Y} \left(-(\mathbf{W} \mathbf{W}^T + \sigma^2 \mathbf{I})^{-1} (\mathbf{J}^{ij} \mathbf{W}^T + \mathbf{W} \mathbf{J}^{ijT}) (\mathbf{W} \mathbf{W}^T + \sigma^2 \mathbf{I}) \right) \right)$$

And so by combining our terms A and B we have the gradient for our log likelihood \mathcal{L} :

$$\begin{split} \frac{\partial \mathcal{L}}{\partial \mathbf{W}_{ij}} &= \frac{N}{2} \Bigg(\operatorname{Tr} \left[\mathbf{Y}^T \mathbf{Y} \Big(- (\mathbf{W} \mathbf{W}^T + \sigma^2 \mathbf{I})^{-1} (\mathbf{J}^{ij} \mathbf{W}^T + \mathbf{W} \mathbf{J}^{ijT}) (\mathbf{W} \mathbf{W}^T + \sigma^2 \mathbf{I}) \Big) \right] \\ &+ \operatorname{Tr} \left[(\mathbf{W} \mathbf{W}^T + \sigma^2 \mathbf{I})^{-1} (\mathbf{J}^{ij} \mathbf{W}^T + \mathbf{W} \mathbf{J}^{ijT}) \right] \Bigg) \end{split}$$

References

[1] K. B. Petersen and M. S. Petersen. The Matrix Cookbook. November 2012.