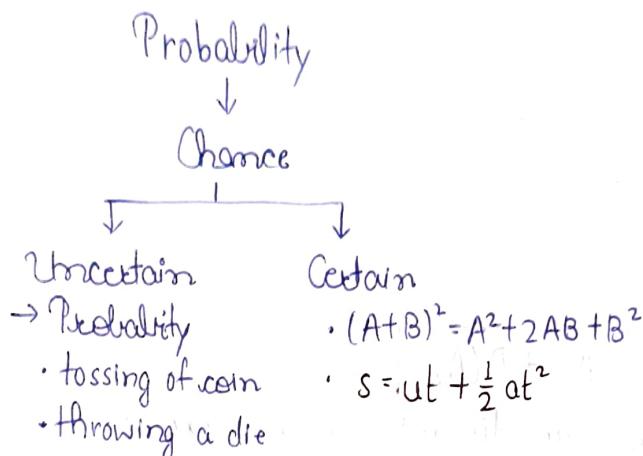
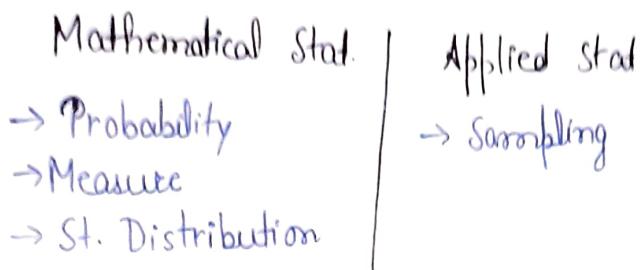


Engineering Mathematics II

(Probability and Statistics)



Apriori Probability

let us suppose a random experiment 'E' in its sample space 'S' contain a finite number

• $n(S) \rightarrow$ Sample points

and event 'A' which contains

• $n(A)$

$$P(A) = \frac{n(A)}{n(S)}$$

Statistical Regularity

$$f_n(A) = \frac{n(A)}{A}$$

event A is found to occur on event A

Statistical Regularity Probability/Apriori

$$P(A) = \lim_{n \rightarrow \infty} f_n = \lim_{n \rightarrow \infty} \frac{n(A)}{A}$$

e.g.- a company manufactures blades

Let $A = \text{blade is defective}$

$n = \text{no. of blades taken}$

$n(A) = \text{no. of defective blades}$

$$n(A) = 3 \quad 2 \quad 5 \quad 8 \quad 9 \quad 12 \quad 12 \quad 15 \quad \dots$$

$$n = 10 \quad 15 \quad 20 \quad 25 \quad 30 \quad 35 \quad 40 \quad 45 \quad \dots$$

$$P(A) = \frac{n(A)}{n} = \frac{1}{3}$$

Theorem

$$0 \leq P(A) \leq 1$$

$P(A) = 0 \rightarrow \text{impossible case}$

$P(A) = 1 \rightarrow \text{certain case}$

$$0 < P(A) < 1$$

$$P(A_1 \cup A_2) = P(A_1) + P(A_2)$$

• A_1 and A_2 are disjoint (mutually exclusive)

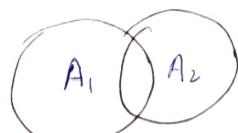


$$P\left(\bigcup_{i=1}^n A_i\right) = \sum_{i=1}^n P(A_i)$$

Total Probability Theorem

$$P(A_1 \cup A_2) = P(A_1) + P(A_2) - P(A_1 \cap A_2)$$

Conditional Probability



• not disjoint

- Let, A and B are two events with random experiment E .
- Make the hypothesis that the event A has occurred $n(A)$ times and B occurs simultaneously with A , $n(A \cap B)$ times in n separation.
- Then ration $\frac{n(A \cap B)}{n(A)}$ is called "conditional probability".

Conditional Frequency Ratio

$$f_n(B/A) = \frac{n(A \cap B)}{n(A)}$$

If $f_n(B/A)$ tends to finite as $n \rightarrow \infty$,

then conditional frequency ratio is called conditional probability.

- $P(B/A) = \lim_{n \rightarrow \infty} f_n(B/A) = \lim_{n \rightarrow \infty} \frac{n(A \cap B)}{n(A)}$
- $= \frac{P(A \cap B)}{P(A)}, P(A) \neq 0$
- $P(A/B) = \frac{P(A \cap B)}{P(B)}, P(B) \neq 0$

e.g.- Let one card be drawn from a full pack, A = spade, B = king
find the probability of king supposing spade occurred

$$\therefore \text{Probability } P(B/A) = \frac{P(A \cap B)}{P(A)} = \frac{1}{13}$$

Multiplication Rule of Probability

$$P(A \cap B) = P(A) \cdot P\left(\frac{B}{A}\right) = P(B) \cdot P\left(\frac{A}{B}\right)$$

Similarly, for n events A_1, A_2, \dots, A_n ,

$$P(A_1 \cap A_2 \cap \dots \cap A_n) = P(A_1) \cdot P\left(\frac{A_2}{A_1}\right) \cdot P\left(\frac{A_3}{A_1 \cap A_2}\right) \cdots P\left(\frac{A_n}{\bigcap_{i=1}^{n-1} A_i}\right)$$

For independent events,

$$P(A \cap B) = P(A) \cdot P(B)$$

Bay's Theorem

If A_1, A_2, \dots, A_n be a given set of pairwise mutually exclusive and exhaustive events, then for any event 'A',

$$\text{I} P(A) = \sum_{i=1}^n P(A_i) \cdot P\left(A/A_i\right)$$

$$\text{II } P\left(\frac{A_i}{A}\right) = \frac{P(A_i) \cdot P(A/A_i)}{P(A)}$$

e.g. There are three urns

first urn contains \rightarrow 3 red + 4 black balls

second " " \rightarrow 2 red + 6 black "

third " " \rightarrow 3 red + 3 black "

one urn is chosen and a ball is drawn from the urn.

let $A_1 = \text{Ist urn chosen}$

$A_2 = \text{IInd } ", "$

$A_3 = \text{IIIrd } ", "$

$A \rightarrow \text{ball is red}$

$$\begin{aligned}\therefore P(A) &= P(A_1) \cdot P\left(\frac{A}{A_1}\right) + P(A_2) \cdot P\left(\frac{A}{A_2}\right) + P(A_3) \cdot P\left(\frac{A}{A_3}\right) \\ &= \frac{1}{3} \left(\frac{3}{7} + \frac{2}{8} + \frac{5}{8} \right) \\ &= \frac{73}{168}\end{aligned}$$

Q) In a belt factory, machine 'A', 'B', 'C' manufacture respectively 25%, 35% and 40% of the total. Their outputs are 5%, 2% respectively defective belts. A belt is drawn at random from the product and is found to be defective. What are the probabilities that it was manufactured by machine (i) A (ii) B (iii) C?

Let: $X_A \rightarrow \text{event that belt is manufactured by A}$

$X_B \rightarrow \text{event that belt is manufactured by B}$

$X_C \rightarrow \text{event that belt is manufactured by C}$

$X \rightarrow \text{event that belt is defective}$

$$P(X_A) = \frac{1}{4}$$

$$P(X_B) = \frac{7}{20}$$

$$P(X_C) = \frac{2}{5}$$

$$\begin{aligned}P(X) &= P(X_A) \cdot P\left(\frac{X}{X_A}\right) + P(X_B) \cdot P\left(\frac{X}{X_B}\right) \\ &\quad + P(X_C) \cdot P\left(\frac{X}{X_C}\right) \\ &= \left(\frac{1}{4}\right)\left(\frac{1}{20}\right) + \left(\frac{7}{20}\right)\left(\frac{1}{25}\right) + \left(\frac{2}{5}\right)\left(\frac{1}{50}\right) \\ &= \frac{1}{80} + \frac{7}{500} + \frac{1}{125} \\ &= \frac{69}{2000}\end{aligned}$$

$$\textcircled{i} \quad P\left(\frac{X_A}{X}\right) = \frac{P(X_A \cap X)}{P(X)}$$

$$= \frac{P(X_A) \cdot P\left(\frac{X}{X_A}\right)}{P(X)}$$

$$= \frac{\frac{1}{4} \times \frac{1}{20}}{\frac{69}{2000}}$$

$$= \frac{25}{69}$$

$$\textcircled{ii} \quad P\left(\frac{X_B}{X}\right) = \frac{28}{69}$$

$$\textcircled{iii} \quad P\left(\frac{X_C}{X}\right) = \frac{16}{69}$$

Q.) The probability of a college student being a male is $\frac{1}{8}$ and that of being a female is $\frac{7}{8}$. The probability that a male student completes the course is $\frac{2}{3}$ and that of a female student is $\frac{1}{3}$. A student is selected at random and is found to have completed the course. What is the probability that the student is (i) male (ii) female?

Random Variables (R.V.)

"STOCHASTIC / CHANCE Variables" "Spectrum"

e.g.- three electronic components are tested and result may be either defective or non-defective

$$S = \{NNN, NND, NDN, \dots, DDD\}$$

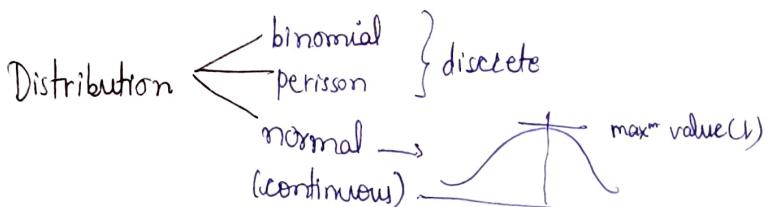
$$R.V(D) = \{0, 1, 2, 3\}$$

- A r.v. is called a "discrete" variable if the set of possible outcomes is countable.
- A r.v. is called a "continuous" variable if it can take values on a continuous scale.

PMF \rightarrow Probability Mass Function \rightarrow discrete

PDF \rightarrow Probability Density Function \rightarrow continuous

$$F(X=x) = P(X \leq x), -\infty < x < \infty \rightarrow \text{Cumulative distribution function}$$



e.g.- a consignment of 8 similar micro-computer to a retail outlet contain 3 defectives.

if a firm makes a random purchase of 2 of these computer.

Find a prob. distribution for no. of objects defectives

$X \rightarrow$ r.v. for two defectives

So, here x takes the value 0, 1, 2

$$P(X=0) = F(X=0) = \frac{{}^3C_0 \cdot {}^5C_2}{{}^8C_2} = \frac{10}{28}$$

$$F(X=1) = \frac{{}^3C_1 \cdot {}^5C_1}{{}^8C_2} = \frac{15}{28}$$

$$F(X=2) = \frac{{}^3C_2 \cdot {}^5C_0}{{}^8C_2} = \frac{2}{28}$$

x	0	1	2
$P(X=x)$	0.36	0.63	0.11

$$\sum F(x=x) = 1$$

$$\sum F(x=x) \geq 0$$

$$0 < F(x=x) < 1$$

Continuous Distribution Function

A function $F(x=x)$ is a probability density function (PDF) for continuous r.v defined:

$$\textcircled{1} \quad F(x=x) \geq 0 \quad \forall x \in \mathbb{R}$$

$$\textcircled{2} \quad \int_{-\infty}^{\infty} F(x=x) dx = 1$$

$$\textcircled{3} \quad P(a \leq x \leq b) = \int_a^b F(x=x) dx$$

e.g.- suppose that the error in the rx temp. (${}^{\circ}\text{C}$) for a controlled lab experiment is continuous r.v 'X' having probability density function

$$F(x=x) = \begin{cases} \frac{x^2}{3}, & -1 < x < 2 \\ 0, & \text{otherwise} \end{cases}$$

\textcircled{a} Verify that the total probability is unity

\textcircled{b} Find $P(0 \leq x \leq 1)$.

$$\textcircled{a} \quad \int_{-\infty}^{\infty} F(x=x) dx = \int_{-1}^2 \frac{x^2}{3} dx = \left[\frac{x^3}{9} \right]_{-1}^2 = \frac{8}{9} - \left(\frac{-1}{9} \right) = \frac{9}{9} = 1$$

$$\textcircled{b} \quad P(0 \leq x \leq 1) = \int_0^1 \frac{x^2}{3} dx = \left[\frac{x^3}{9} \right]_0^1 = \frac{1}{9}$$

Expectation or Mean of R.V.

① Let, X be a discrete R.V. whose distribution is given:

$$X : x_0 \ x_1 \ x_2 \dots$$

$$P(X=x) : f_0 \ f_1 \ f_2 \dots$$

Mean = $E(X) = x_0 f_0 + x_1 f_1 + x_2 f_2 + \dots$

$$(M.m) \quad \boxed{E(X) = \sum_{i=0}^n x_i f_i}$$

② Let, the PDF of a continuous R.V

$$\boxed{E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx}$$

e.g. - $f(x) = \begin{cases} \frac{1}{2}, & -1 < x < 1 \\ 0, & \text{otherwise} \end{cases}$

$$E(X) = \int_{-1}^1 x \cdot \frac{1}{2} dx$$

$$E(2x^3) = \int_{-\infty}^{\infty} 2x^3 \cdot f(x) dx$$

$$= \left. \frac{x^4}{4} \right|_{-1}^1$$

$$= 0$$

Properties of Expectation

(i) $E(a) = a$, where $a \rightarrow \text{constant}$

(ii) $E(ax) = aE(x)$

(iii) $E(x \pm y) = E(x) \pm E(y)$

(iv) $E(xy) = E(x) \cdot E(y)$

if independent

Q.) A number is chosen at random from the set $\{1, 2, \dots, 100\}$ and another number is chosen at random from the set $\{1, 2, \dots, 50\}$. Find the expected value of the product.

Let, $x \rightarrow$ first no.

and, $y \rightarrow$ second no.

for

$$X: 1 \quad 2 \quad 3 \quad \dots \quad 100$$

$$f_i: \frac{1}{100} \quad \frac{1}{100} \quad \frac{1}{100} \quad \dots \quad \frac{1}{100}$$

$$y: 1 \quad 2 \quad \dots \quad 50$$

$$g_i: \frac{1}{50} \quad \frac{1}{50} \quad \dots \quad \frac{1}{50}$$

$$\begin{aligned} E(X) &= \sum_{i=1}^{100} x_i f_i \\ &= \frac{1}{100} (1+2+3+\dots+100) \\ &= \frac{100 \times 101}{100 \times 2} \\ &= \frac{101}{2} \\ &= 50.5 \end{aligned}$$

$$\begin{aligned} E(y) &= \frac{1}{50} (1+2+\dots+50) \\ &= \frac{50 \times 51}{2 \times 50} \\ &= \frac{51}{2} \\ &= 25.5 \end{aligned}$$

$$\therefore E(XY) = E(X) \cdot E(Y)$$

$$= \left(\frac{101}{2}\right) \left(\frac{51}{2}\right)$$

Variance and Standard Deviation

The variance of a r.v. denoted by $\text{Var}(X)$ and,

$$\text{Var}(X) = E[(x - m)^2]$$

$$\boxed{\text{Var}(X) = E[(x - E(x))^2]}$$

$$\text{S.D} = \sqrt{\text{Var}(X)}$$

$$\boxed{\sigma_x = \sqrt{E(x - E(x))^2}}$$

$$\boxed{\text{Var}(X) = E(x^2) - m^2}$$

$$\boxed{\text{Var}(X) = E(x^2) - [E(x)]^2}$$

$$\text{Var}(\text{constant}) = 0$$

$$\textcircled{I} \quad \text{Var}(X) = E(x^2) - [E(x)]^2$$

(iv)

$$\textcircled{II} \quad \text{Var}(ax+b) = a^2 \text{Var}(X)$$

$$M = m = E(x) = \text{mean}$$

$$\text{Var}(X) = E\{x - m\}^2$$

$$= E(x^2 + m^2 - 2xm)$$

$$= E(x^2) + E(m^2) - E(2xm)$$

$$= E(x^2) - 2mE(x) + m^2$$

$$= E(x^2) - 2m^2 + m^2$$

$$= E(x^2) - m^2$$

$$\text{e.g. - } f(x) = \begin{cases} \frac{1}{2}x & , 0 \leq x \leq 2 \\ 0 & , \text{elsewhere} \end{cases}$$

$$E(x) = \int_{-\infty}^{\infty} x \cdot f(x) dx$$

$$= \int_0^2 x \cdot \frac{1}{2} dx$$

$$= \left[\frac{x^2}{4} \right]_0^2$$

$$= \frac{1}{4} (4)$$

$$= 1$$

$$E(x^2) = \int_0^2 x^2 \cdot \frac{1}{2} dx$$

$$= \left[\frac{x^3}{2 \cdot 3} \right]_0^2$$

$$= \frac{2^3}{2 \cdot 3}$$

$$= \frac{4}{3}$$

$$\text{Var}(x) = E(x^2) - [E(x)]^2$$

$$= \frac{4}{3} - 1$$

$$= \frac{1}{3}$$

$$S.D(x) = \frac{1}{\sqrt{3}}$$

Q.) A random variable x has following probability mass function

$$X = 0 \quad 1 \quad 2 \quad 3 \quad 4 \quad 5 \quad 6 \quad 7$$

$$P(X=x) : 0 \quad k \quad 2k \quad 2k \quad 3k \quad k^2 \quad 2k^2 \quad 7k^2 + k$$

i) Determine the constant 'k'

ii) Determine the min^m value of x , so that $P(X \leq x) > \frac{1}{2}$

iii) Obtain the distribution $F(x)$. iv) Evaluate $P(X < 6)$, $P(X \geq 6)$, $P(6 < X \leq 6)$, $P(X \leq 3 / X \leq 6)$

$$\sum_{x=0}^6 P(X=x) = 1$$

$$\Rightarrow 0 + k + 2k + 2k + 3k + k^2 + 2k^2 + 7k^2 + k = 1$$

$$\Rightarrow 10k^2 + 9k - 1 = 0$$

$$\Rightarrow 10k^2 + 10k - k - 1 = 0$$

$$\Rightarrow 10k(k+1) - 1(k+1) = 0$$

$$\Rightarrow k = \frac{1}{10}, -1$$

$$\boxed{k = \frac{1}{10}}$$

$$\textcircled{iv} \quad P(X < 6) = 1 - P(X \geq 6)$$

$$= 1 - \{P(X=6) + P(X=7)\}$$

$$= 1 - \left\{ \frac{2}{100} + \frac{17}{100} \right\}$$

$$= \frac{81}{100}$$

$$P(X \geq 6) = \frac{19}{100}$$

$$P(3 < X \leq 6) = P(X=4) + P(X=5) + P(X=6)$$

$$= \frac{30}{100} + \frac{1}{100} + \frac{2}{100}$$

$$= \frac{33}{100}$$

$$P(X \geq 3 | X \leq 6) = \frac{P\{(X \geq 3) \cap (X \leq 6)\}}{P(X \leq 6)}$$

$$= \frac{\frac{33}{100}}{\frac{83}{100}}$$

$$= \frac{33}{83}$$

$$\textcircled{i} \quad P(X \leq m) > \frac{1}{2}$$

$$P(X \leq 1) = P(X=0) + P(X=1)$$

$$= 0 + \frac{1}{10}$$

$$= \frac{1}{10}$$

$$P(X \leq 2) = 0 + \frac{1}{10} + \frac{2}{10}$$

$$= \frac{3}{10}$$

$$P(X \leq 3) = 0 + \frac{1}{10} + \frac{2}{10} + \frac{2}{10}$$

$$= \frac{5}{10}$$

$$= \frac{1}{2}$$

$$P(X \leq 4) = \frac{1}{2} + \frac{3}{10} = \frac{8}{10} > \frac{1}{2}$$

$$\therefore \boxed{n=4}$$

(11) Distribution function:

$$F(x) = \begin{cases} 0 & , -\infty \leq x < 1 \\ \frac{1}{10} & , 1 \leq x < 2 \\ \frac{3}{10} & , 2 \leq x < 3 \\ \frac{1}{2} & , 3 \leq x < 4 \\ \frac{4}{5} & , 4 \leq x < 5 \\ \frac{81}{100} & , 5 \leq x < 6 \\ \frac{83}{100} & , 6 \leq x < 7 \\ 1 & , 7 \leq x < \infty \end{cases}$$

↑
cumulative
frequency

Q) Find the mean variance of continuous variable having pdf

$$f(x) = \begin{cases} 1 - |1-x| & , 0 < x < 2 \\ 0 & , \text{elsewhere} \end{cases}$$

Also find $E(x^3)$.

$$\begin{aligned} E(x) &= \int_0^2 [1 - |1-x|] dx \\ &= \int_0^1 [1 - (1-x)] dx + \int_1^2 [1 - (-1+x)] dx \\ &= \int_0^1 x dx + \int_1^2 (2-x) dx \\ &= \left[\frac{x^2}{2} \right]_0^1 + 2[x]_1^2 - \left[\frac{x^2}{2} \right]_1^2 \\ &= \frac{1}{2} + 2 - \frac{3}{2} \\ &= 1 \end{aligned}$$

$$\begin{aligned} E(x^3) &= \int_0^2 [1 - |1-x^2|] dx \\ &= \int_0^1 [1 - (1+x^2)] dx + \int_1^2 [1 - (-1+x^2)] dx \\ &= \int_0^1 x^2 dx + \int_1^2 (2-x^2) dx \\ &= \frac{1}{3} + 2.3 - \frac{1}{3} \times 6 \\ &= \frac{1}{3} + \frac{18}{3} - \frac{63}{3} = -\frac{44}{3} \end{aligned}$$

$$\begin{aligned} \text{Var}(X) &= E(X^2) - [E(X)]^2 \\ &= -\frac{44}{3} - 1 \\ &= -\frac{47}{3} \end{aligned}$$

Theoretical Discrete and Continuous Distribution

(Special Types of Distribution)

↳ 7 types

- ① Binomial Distribution ? discrete
- ② Poisson Distribution
- ③ Normal Distribution → continuous
↳ sociology, market research, etc.

Binomial Distribution (Bernoulli's dist.)

A discrete r.v. is said to be binomial dist. with parameters P ($0 < P < 1$) and n (a positive integer)

$$X: 0 \quad 1 \quad 2 \quad 3 \quad \dots \quad n$$

$$f_i: f_0 \quad f_1 \quad f_2 \quad f_3 \quad \dots \quad f_n$$

$$P(X=i) = {}^n C_r \frac{p^i}{\uparrow} \frac{(1-p)^{n-i}}{\uparrow} ; i=0, 1, 2, \dots$$

$\boxed{B(p, n)}$

$$X \sim B(p, n)$$

↑
parameters

Properties:

- ① The experiment consists of 'n' repeated times (trials)
- ② Each trial result is an outcome classified as successful or unsuccessful
 (p) (q)
- ③ The repeated trials are independent

Limitations:

- ① $n \rightarrow$ finite
- ② The value of p and q are around $\frac{1}{2}$ (never very small)

$$P(X=x) = \begin{cases} {}^n C_i p^i q^{n-i}, & i \in \mathbb{Z}^+ \cup \{0\} \\ 0 & \text{otherwise} \end{cases}$$

Tossing of coin

$n=1, H \quad T$

$$P = p(H) + p(T) = p(H) \cup p(T) = P(p \cup q)$$

$$= (p+q)^1$$

$n=2, HH \quad HT \quad TH \quad TT$

$$\begin{aligned} & P(HH) + P(HT) + P(TH) + P(TT) \\ &= P(H) \cdot P(H) + 2 \cdot P(H) \cdot P(T) + P(T) \cdot P(T) \\ &= p^2 + 2pq + q^2 \\ &= (p+q)^2 \end{aligned}$$

$n=3, HHH, HHT, \dots, TTT$

$$(p+q)^3$$

\vdots
 $n=n, (p+q)^n$

$$= {}^nC_0 p^n q^0 + {}^nC_1 p^1 q^{n-1} + \dots + {}^nC_i p^i q^{n-i} + \dots + {}^nC_n p^n q^0$$

Q) The probability that a certain kind of component will survive a given shock test is $\frac{3}{4}$. Find the probability that exactly 2 of the next 4 components will survive.

Soln: $n=4, p=\frac{3}{4}, q=\frac{1}{4}$

$\therefore X \sim B(n, p)$

$$P(X=2) = {}^4C_2 \left(\frac{3}{4}\right)^2 \left(\frac{1}{4}\right)^2$$

$$= \frac{4 \times 3}{2} \times \frac{3 \times 3}{4 \times 4} \times \frac{1}{16}$$

$$= 0.21$$

Mean and Variance of Binomial Distribution

$$\begin{aligned}
 \text{Mean} = E(X) &= \sum_{x=0}^n x \cdot P(X=x) \\
 &= \sum_{x=0}^n x \cdot {}^n C_x p^x q^{n-x} \\
 &= \sum_{x=0}^n x \cdot \left(\frac{n}{x}\right) {}^{n-1} C_{x-1} p^{x-1} q^{n-x} \\
 &= np \sum_{x=0}^n {}^{n-1} C_{x-1} p^{x-1} q^{n-x} \\
 &= np (p+q)^{n-1} \\
 &= np \cdot (1)^{n-1} \\
 &= np
 \end{aligned}$$

$\therefore \boxed{\text{Mean} = np}$

$$\begin{aligned}
 \text{Variance} = \text{Var}(X) &= E(X^2) - [E(X)]^2 \\
 &= npq
 \end{aligned}$$

$$\begin{aligned}
 E(X^2) &= \sum_{x=0}^n x^2 \cdot {}^n C_x p^x q^{n-x} \\
 &= \sum_{x=0}^n \{x(x-1)+x\} {}^n C_x p^x q^{n-x} \\
 &= \sum_{x=0}^n x(x-1) \cdot {}^n C_x p^x q^{n-x} + \sum_{x=0}^n x {}^n C_x p^x q^{n-x} \\
 &= \sum_{x=0}^n x(x-1) \cdot \frac{n(n-1)}{x(x-1)} \cdot {}^{n-2} C_{x-2} p^{x-2} q^{n-x} + \underbrace{\sum_{x=0}^n x {}^n C_{x-1} p^n q^{n-x}}_{E(X)} \\
 &= n(n-1) \sum_{x=0}^n {}^{n-2} C_{x-2} p^{x-2} q^{n-x} + \underbrace{\sum_{x=0}^n {}^n C_{x-1} p^n q^{n-x}}_{E(X)} \\
 &= n(n-1)p^2 \sum_{x=0}^n {}^{n-2} C_{x-2} p^{x-2} q^{n-x} + np \\
 &= n(n-1)p^2 + np
 \end{aligned}$$

$$\therefore \text{Var}(X) = \frac{n(n-1)p^2 + np - (np)^2}{\boxed{\text{Var}(X) = npq}}$$

e.g.- Let x be the no. of 1s obtained in 15 throws of a die. Find its mean and variance.

Defn: X is a binomial distribution

$$P(\text{getting } 1) = \frac{1}{6}$$

$$q = \frac{5}{6}$$

$$n = 15$$

$$\text{Mean} = np \\ = \frac{15}{6} = 2.5$$

$$\text{Var}(X) = npq \\ = 15 \left(\frac{1}{6} \right) \left(\frac{5}{6} \right) \\ = \frac{75}{36} \\ = 2.09$$

Q.) A defective die is thrown TEN times independently. The prob. that an even will appear 5 times is twice the probability that an even no. appears 4 times. What is the prob. that odd face appears in each of ten times?

Defn: Let success(p) = even face

X = no. of even faces among 10 throws

$$P(X=x) = {}^{10}C_x p^x (1-p)^{10-x}$$

A/qv,

$$P(X=5) = 2 P(X=4)$$

$${}^{10}C_5 p^5 (1-p)^5 = 2 \cdot {}^{10}C_4 p^4 (1-p)^6$$

$$p(1-p) = \frac{2 \cdot {}^{10}C_4}{{}^{10}C_5} = \frac{2 \times 10! \times 5! \times 5!}{4! \times 6! \times 10!} = \frac{2 \times 5 \times 5! \times 5!}{4! \times 6 \times 5!} = \frac{5}{3}$$

$$3p - 3p^2 = 5$$

$$3p^2 - 3p + 5 = 0$$

$$p = \frac{5}{8}$$

$$\therefore P(X=0) = {}^{10}C_0 p^0 (1-p)^{10} = 10 \cdot \frac{5}{8} \left(\frac{3}{8} \right)^{10}$$

Poisson Distribution

Def: A r.v 'X' is said to follow poisson distribution iff it assumes non-negative value and its prob mass function is

$$P(x, \lambda) = \begin{cases} P(X=x) = \frac{e^{-\lambda} \lambda^x}{x!}, & x=0, 1, 2, \dots \\ 0, & \text{elsewhere} \end{cases}$$

• λ is a parameter

$$\boxed{\lambda > 0}$$

$$\boxed{\lambda = np}$$

$$X \sim P(X)$$

• Used when prob. of success is very less
(small)

Derivation:

Binomial dist. tends to poisson dist. under the following conditions —

- ① Probability of success is very small ($p \rightarrow 0$)
- ② Number of trials is very large ($n \rightarrow \infty$)

$$\boxed{\lambda = np}$$

$$P_n = \begin{cases} {}^n C_x \cdot p^x \cdot q^{n-x}, & x=0, 1, 2, \dots \\ 0, & \text{otherwise} \end{cases}$$

$$= \frac{n!}{x!(n-x)!} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x}$$

$$= \frac{n(n-1)(n-2) \dots (n-(x-1))(n-x)!}{x!(n-x)!} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x}$$

$$= \lambda^n \cdot \left(1 - \frac{1}{n}\right) \left(1 - \frac{2}{n}\right) \cdots \left(1 - \frac{(n-1)}{n}\right) \cdot \frac{\lambda^x}{x!} \cdot \left(1 - \frac{1}{n}\right)^{n-x}$$

We know that,

$$\lim_{n \rightarrow \infty} \left(1 - \frac{1}{n}\right)^n = e^{-1}$$

$$\text{and, } \lim_{n \rightarrow \infty} \left(1 - \frac{1}{n}\right)^{n-x} = 1$$

Then,

$$P_x = \boxed{\frac{e^{-\lambda} \cdot \lambda^x}{x!}}$$

Mean and Variance for Poisson Distribution

$$\text{Mean} = E(X)$$

$$= \sum_{x=0}^{\infty} x \cdot \frac{e^{-\lambda} \cdot \lambda^x}{x!}$$

$$= \sum_{x=0}^{\infty} \frac{e^{-\lambda} \cdot \lambda^x}{(x-1)!}$$

$$= \cancel{\sum_{x=2}^{\infty}} e^{-\lambda} \lambda \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!}$$

$$= e^{-\lambda} \cdot \lambda \cdot e^{\lambda}$$

$$= \lambda$$

$$\text{Mean} = E(X)$$

$$= \lambda$$

~~$$= np$$~~

$$= np$$

Variance:

$$\text{Var}(X) = E(X^2) - [E(X)]^2$$

$$E(X^2) = \sum_{x=0}^{\infty} x^2 \cdot \frac{e^{-\lambda} \lambda^x}{x!}$$

$$= \sum_{x=0}^{\infty} [x(x-1)+x] \frac{e^{-\lambda} \lambda^x}{x!}$$

$$= \underbrace{\sum_{x=0}^{\infty} \frac{e^{-\lambda} \cdot \lambda^x}{(x-1)!}}_{\lambda} + \sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{(x-2)!}$$

$$= \lambda + \sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{(x-2)!}$$

$$= \lambda^2 e^{-\lambda} \sum_{x=2}^{\infty} \frac{\lambda^{x-2} \lambda^x}{(x-2)!} + \lambda$$

$$= \lambda^2 e^{-\lambda} e^{\lambda} + \lambda$$

$$= \lambda^2 + \lambda$$

$$\therefore \text{Var}(X) = (\lambda^2 + \lambda) - \lambda^2$$

$$= \lambda$$

$$= np$$

Q) The average no. of radioactive particles through a counter during 1 ms in a laboratory experiment is 4. What is the probability that 6 particles enter the counter in a given millisecond?

Sol": $n=6$ and $\lambda=4$

$$P(X=6, \lambda=4) = \frac{e^{-\lambda} \cdot \lambda^x}{x!}$$

$$= \frac{e^{-4} \cdot 4^6}{6!}$$

$$= \frac{0.0183 \times 4^6}{720}$$

$$= 0.1041$$

Q) In a certain factory turning razor blades, there is a small chance $\frac{1}{5000}$ for any blade to be defective. The blades are in packets of 10. Use Poisson Distribution to calculate the approx. no. of packets containing (i) No defective (ii) 2 defective blades respectively in one consignment of 10000 packets ($e^{-0.02} = 0.9802$)

Sol": 1 defective blade in 5000 blades

$$\text{So, avg. no. of defective blades (in each packet)} = 10 \times \frac{\frac{1}{5000}}{500} \\ = \frac{1}{500} \\ = 0.002$$

Let, 'X' be no. of defective blades in packet.

$$P_x = P(X=x) = \frac{e^{-\lambda} \cdot \lambda^x}{x!}$$

① $x=0$

$$P_0 = \frac{e^{-\lambda} \cdot \lambda^0}{0!} = e^{-\lambda} \\ = e^{-0.02}$$

$$= 0.9802$$

\therefore No. of pkt. in consignment containing no defectives = 10000×0.9802

$$\textcircled{1} \quad x=2$$

$$P_x = \frac{e^{-\lambda} \cdot \lambda^x}{x!}$$

$$= \frac{e^{-0.02} \times (0.02)^2}{2!}$$

$$= 0.00019604$$

\therefore No. of p.t. in consignments containing two defectives

$$= 0.00019604 \times 10000$$

$$\approx 1.9604$$

$$\approx 2$$

Q) 2% of the items made by a machine are defective. Find the prob. that 3 or more items are defective in a sample of 100 items.

(Given : $e^{-1} = 0.368$, $e^{-2} = 0.135$, $e^{-3} = 0.0498$)

Sol": $n = 100$

$$\phi = 2\% \cancel{\times 100}$$

$$= \frac{2}{100}$$

$$= 0.02$$

$$P_x = \frac{e^{-\lambda} \cdot \lambda^x}{x!} = \frac{e^{-2} \cdot 2^x}{x!}$$

\therefore Required Probability = $P(X \geq 3)$

$$= 1 - P(X < 3)$$

$$= 1 - \left\{ P(X=0) + P(X=1) + P(X=2) \right\}$$

$$= 1 - \left\{ \frac{e^{-2} \cdot 2^0}{0!} + \frac{e^{-2} \cdot 2^1}{1!} + \frac{e^{-2} \cdot 2^2}{2!} \right\}$$

$$= 1 - e^{-2} \{ 1 + 2 + 2 \}$$

$$= 1 - 5(0.135)$$

$$= 0.325$$

Q.) Red blood cell deficiency may be determined by examining a specimen of blood under microscope. Suppose a certain small fixed volume contained on the average 20 red cells for normal people. Find the prob. that a specimen from a normal person will contain less than 15 red cells.

Sol: $\lambda = 8.20$ (given)

Probability that a normal person contain x red cells

$$= \frac{e^{-\lambda} \cdot \lambda^x}{x!}$$

$$= \frac{e^{-20} \cdot 20^x}{x!}$$

$$P(X < 15) = \sum_{x=0}^{14} \frac{e^{-20} \cdot 20^x}{x!}$$

$$= e^{-20} \left[1 + \frac{20}{1!} + \frac{20^2}{2!} + \dots + \frac{20^{14}}{14!} \right]$$

• Measures of Central Tendency

Mean, Median, Mode, Geometrical Mean } Grouped data
Raw data

- ① 7, 8, 9, 10, 11 }
 ② 3, 6, 9, 12, 15 } Mean = 9
 ③ 1, 5, 9, 13, 17 }

- Homogeneous data → data points lie closer to the mean
- Heterogeneous data

• Measures of Dispersion

↳ Skewness, Kurtosis

① Range : A B
(greatest) (smallest)

$$\boxed{\text{Range} = A - B}$$

② Quartile

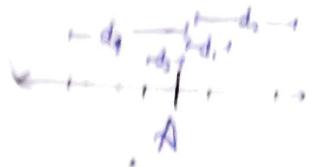
Q_1 Q_2 Q_3 Q_4

$$\boxed{Q = \frac{Q_1 - Q_3}{2}}$$

III Mean Deviation

$$MD = \frac{1}{N} \sum f_i |x_i - A|$$

$$MD = \frac{1}{N} \sum f_i |x_i - \bar{x}|$$



$$MD = \frac{1}{N} \sum f_i |x_i - \bar{x}|$$

$$\sum f_i = N$$

IV R.M.S - standard deviation

$$S^2 = \frac{1}{N} \sum f_i (x_i - A)^2$$

$$\sigma^2 = \frac{1}{N} \sum f_i (x_i - A)^2$$

V Moment

$$\mu_1 = \frac{1}{N} \sum f_i (x_i - A)^1$$

$$\mu_1 = \frac{1}{N} \sum f_i d_i$$

Here, $d_i = x_i - A$

$$\mu_r = \frac{1}{N} \sum (x_i - \bar{x})^r$$

$$\mu_0 = \frac{1}{N} \sum f_i = 1$$

$$\mu_1 = \frac{1}{N} \sum f_i (x_i - \bar{x})^1$$

$$= \frac{1}{N} \sum f_i x_i - \frac{1}{N} \bar{x} \sum f_i$$

$$= 1 - 1$$

$$\mu_1 = 0$$

$$\mu_2 = \frac{1}{N} \sum f_i (x_i - \bar{x})^2$$

$$\mu_2 = \sigma^2$$

Pearson's β and γ coefficient

$$\beta = \frac{\mu_3^2}{\mu_2^2}$$

$$\gamma = \beta - 3$$

$$\left[\begin{array}{c} \beta \\ \gamma \end{array} \right] = \left[\begin{array}{c} \mu_3 \\ \mu_2 \end{array} \right]$$

$$\left[\begin{array}{c} \gamma \\ \beta \end{array} \right] = \left[\begin{array}{c} \mu_2 - 3 \\ \mu_3^2 / \mu_2^2 \end{array} \right]$$

convexity of curve

Skewness:

↳ "Lack of symmetry"



Mean
= Median
= Mode

"Symmetric Curve"

"Normal Curve"

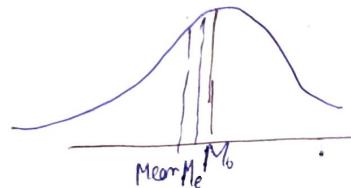
↳ $N(0, 1)$

mean = 0

$\sigma = 1$

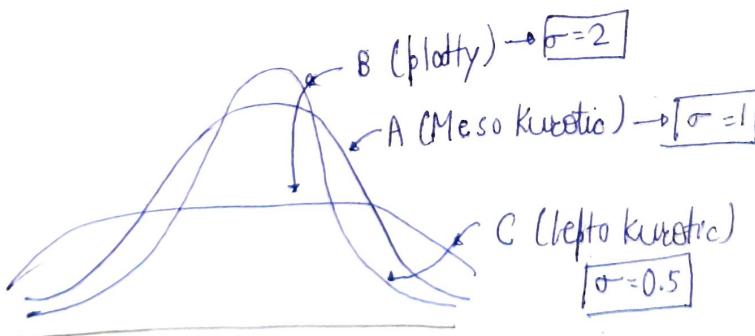


$M_o > M_e > M_m$



$M_m > M_e > M_o$

Kurtosis



A \rightarrow normal curve \rightarrow meso kurtotic

B \rightarrow platyk \rightarrow more flatter

C \rightarrow highly peaked \rightarrow lepto kurtotic

$$\beta_2 = \frac{\mu_2}{\mu_1^2}$$

$$\gamma_2 = \beta_2 - 3$$

(A) $\rightarrow \beta_2 = 3 ; \gamma_2 = 0$ (normal / mesokurtic curve)

Normal Distribution

• For continuous data (r.v.)

- Continuous r.v. 'X' is said to have a normal distribution if its probability density function is given by:

$$N(X) \sim f(x) = \begin{cases} \frac{1}{\sigma\sqrt{2\pi}} \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}} & , -\infty < x < \infty \\ 0 & , \text{otherwise} \end{cases}$$

where μ, σ are two parameters of the distribution

$$X \sim N(\mu, \sigma)$$

• $f(x) \geq 0$

$$\text{Moreover, } \int_{-\infty}^{\infty} f(x) dx = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$= \frac{1}{\sqrt{\pi}} \int_{-\infty}^{\infty} e^{-z^2} dz \quad (\text{by } z = \frac{x-\mu}{\sqrt{2}\sigma})$$

$$= \frac{2}{\sqrt{\pi}} \int_{0}^{\infty} e^{-z^2} dz$$

$$= \frac{1}{\sqrt{\pi}} \int_{0}^{\infty} e^{-u} \cdot u^{\frac{1}{2}} du \quad (z^2 = u)$$

$$= \frac{1}{\sqrt{\pi}} \Gamma\left(\frac{1}{2}\right)$$

$$= \frac{1}{\sqrt{\pi}} \cdot \sqrt{\pi}$$

$$= 1$$

$$\boxed{\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}}$$

(Gamma function)

$$\boxed{\Gamma(m+1) = m\Gamma(m)}$$

$$\boxed{F(x) = \int_{-\infty}^{\infty} f(x) dx = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx}$$

Mean for Normal Distribution

$$\begin{aligned}
 \text{Mean} \cdot E(X) &= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} x \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \\
 &= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} (x-\mu) \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx + \frac{\mu}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \\
 &= \frac{1}{\sqrt{\pi}} \int_{-\infty}^{\infty} u \cdot e^{-u^2} du + \frac{\mu}{\sqrt{\pi}} \int_{-\infty}^{\infty} e^{-u^2} du \quad \left(\frac{x-\mu}{\sigma\sqrt{2}} = u \right) \\
 &= 0 + \frac{\mu}{\sqrt{\pi}} \cdot 2 \int_{-\infty}^{\infty} e^{-u^2} du \\
 &= \frac{\mu}{\sqrt{\pi}} \int_0^{\infty} z^{\frac{1}{2}-1} \cdot e^{-z} dz \quad (u^2 = z) \\
 &= \frac{\mu}{\sqrt{\pi}} \Gamma\left(\frac{1}{2}\right) \\
 &= \frac{\mu}{\sqrt{\pi}} \cdot \sqrt{\pi} \\
 &= \mu
 \end{aligned}$$

$$\boxed{\text{Mean} = \mu}$$

Standard Deviation for Normal Distribution

$$\begin{aligned}
 \text{Var}(X) &= E\{(X-\mu)^2\} \\
 &= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} (x-\mu)^2 \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \quad E(X^2) = \int x^2 P_x dx \\
 &= \frac{2\sigma^2}{\sqrt{\pi}} \int_{-\infty}^{\infty} z^2 \cdot e^{-z^2} dz \quad \left[z = \frac{x-\mu}{\sqrt{2}\sigma} \right] \\
 &= \frac{4\sigma^2}{\sqrt{\pi}} \int_{-\infty}^{\infty} z^2 \cdot e^{-z^2} dz \\
 &= \frac{2\sigma^2}{\sqrt{\pi}} \int_0^{\infty} u^2 \cdot e^{-u} du \quad \left[z^2 = u \right]
 \end{aligned}$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \Gamma(\frac{3}{2})$$

$$\Gamma(3) = 2 \cdot 1 = 2$$

$$\frac{2\sigma^2}{\sqrt{\pi}} \approx \Gamma(1)$$

$$\frac{\sigma^2}{\sqrt{\pi}}$$

$$\therefore \sigma^2$$

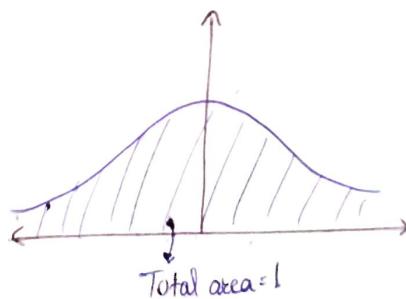
$\text{Var}(X) = \sigma^2$	For normal distribution
$\text{S.D.} = \sigma$	
$\text{Mean} = \mu$	

- $n \rightarrow \text{large}$
- Prob. is moderate $\approx \frac{1}{2}$
(PDF \rightarrow continuous)

Important Result

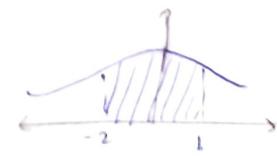
- In continuous random variable X for normal distribution with parameters ' μ ' and ' σ ', then

$$Z = \frac{X - \mu}{\sigma}$$



- Q) To find the no. of days on which power supplied will lie b/w 280 to 310 MW. Suppose $m=300$, $\sigma=10$. Find $P(280 < X < 310)$.

$$Z = \frac{X - \mu}{\sigma} = \frac{X - 300}{10}$$



$$P(280 < X < 310) = P\left(\frac{280-300}{10} < Z < \frac{310-300}{10}\right)$$

$$= P(-2 < Z < 1)$$

$$= \int \phi(z) dz$$

$$\phi(z) = \frac{1}{\sqrt{2\pi}} \int e^{-\frac{z^2}{2}} dz$$

= area enclosed by std. Normal curve and $z = -2$ and $z = 1$

$$= 0.4772 + 0.3413$$

$$= 0.8185 = 81.85\%$$

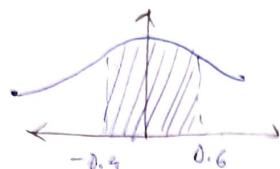
Q) The length of belt produced by a machine is normally distributed with mean = 4 and S.D. = 0.5. A belt is defective if its length doesn't lie in the interval (3.8, 4.3). Find the % of defective belts produced by machine. $\left[\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{-t_2} e^{-t^2/2} dt = 0.7257 \text{ and } \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+t_2} e^{-t^2/2} dt = 0.6564 \right]$

Let $x \rightarrow \text{length of belt}$

$$m = 4, SD = 0.5$$

$$\begin{aligned} P(3.8 < x < 4.3) &= P\left(\frac{3.8-4}{0.5} < \frac{x-4}{0.5} < \frac{4.3-4}{0.5}\right) \\ &= P(-0.4 < z < 0.6) \end{aligned}$$

= area enclosed by std. normal curve and $z = -0.4$ to $z = 0.6$



$$= \int_{-0.4}^{\infty} \phi(t) dt$$

$$= \int_{-\infty}^{0.6} \phi(t) dt - \int_{-\infty}^{-0.4} \phi(t) dt$$

$$= 0.7257 - \int_{0.4}^{\infty} \phi(t) dt \quad (\text{Graph is symmetrical})$$

$$= 0.7257 - \left[1 - \int_{-\infty}^{0.4} \phi(t) dt \right]$$

$$= 0.7257 - [1 - 0.6564]$$

$$= 0.3811$$

\therefore Probability that the length of the whole belt lies b/w 3.8 and 4.3

$$= 0.3811$$

$$= 38.11 \%$$

Probability belt is NOT defective = 0.6189

$$= 61.89\%$$

Q) X is normal distribution and mean of X is 12 and $SD = 4$.

Q) Find out the prob. of the following

(i) $X \geq 20$ (ii) $X \leq 20$ (iii) $0 \leq X \leq 12$

(i) find ' x ' when $P(X > x) = 0.24$

(ii) find x_0 and x_1 when $P(x_0 < X < x_1) = 0.5$ and $P(X > x_1) = 0.25$

soln. $m = 12$, $SD(\sigma) = 4$

(i) (i) $P(X \geq 20)$

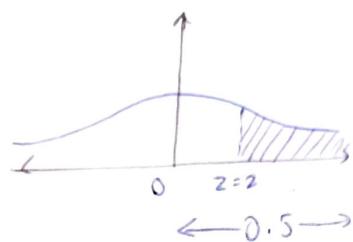
$$Z = \frac{X-12}{4} = \frac{20-12}{4} = 2$$

$$\therefore P(X \geq 20) = P(Z \geq 2)$$

$$= 0.5 - P(0 \leq Z \leq 2)$$

$$= 0.5 - 0.4772$$

$$= 0.0228$$



(ii) $P(X \leq 20) = 1 - P(X \geq 20)$

$$= 1 - 0.0228$$

$$= 0.9772$$

(iii) $P(0 \leq X \leq 12) = P\left(\frac{0-12}{4} \leq \frac{X-12}{4} \leq \frac{12-12}{4}\right)$

$$= P(-3 \leq Z \leq 0)$$

$$= 0.49865$$

(b) $P(X > x) = 0.24$

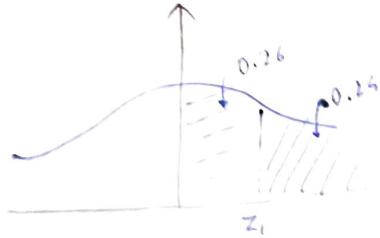
$$\Rightarrow P\left(\frac{X-12}{4} > \frac{x-12}{4}\right) = 0.24$$

$$\Rightarrow P(Z > z_1) = 0.24$$

$$\Rightarrow P(0 < Z < z_1) = \frac{1}{2} - 0.24 = 0.26$$

$$z_1 = \frac{x-12}{4} = 0.71$$

$$\Rightarrow x = 14.84$$



$$\textcircled{c} \quad P(x_0 < X < x_1) = 0.5$$

$$P(X > x_1) = 0.25$$

$$x = x_1 \Rightarrow z = \frac{x_1 - \mu}{\sigma} = z_1 \text{ (say)}$$

$$x = x_0 \Rightarrow z = \frac{x_0 - \mu}{\sigma} = -z_1 \text{ (say)}$$

$$P(z > z_1) = 0.25$$

$$P(0 < z < z_1) = 0.5 - 0.25 = 0.25$$

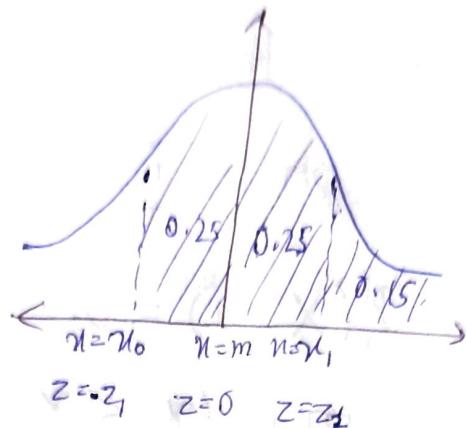
$$\Rightarrow z_1 = 0.67 \text{ (from Table)}$$

$$\frac{x_1 - \mu}{\sigma} = 0.67$$

$$\frac{x_0 - \mu}{\sigma} = -0.67$$

$$\therefore x_1 = 14.68$$

$$x_0 = 9.32$$



Q) In a distribution exactly normal, 7% of the items are under 35 and 89% are under 63. What is mean and SD of the distribution?

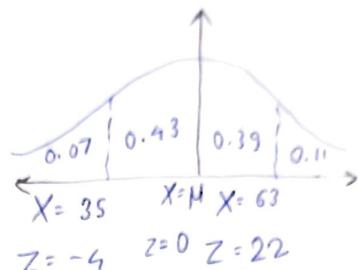
Let: If $X \sim N(\mu, \sigma^2)$, then we are given

$$P(X < 63) = 0.89 \quad P(X < 35) = 0.07$$

$$P(X > 63) = 0.11 \quad P(X > 35) = 0.93$$

$$\text{when, } X=35 \quad Z = \frac{35-\mu}{\sigma} = -Z_1 \text{ (say)}$$

$$X=63 \quad Z = \frac{63-\mu}{\sigma} = Z_2 \text{ (say)}$$



$$P(0 < Z < Z_2) = 0.39 \Rightarrow \frac{63-\mu}{\sigma} = 1.23 \quad \text{--- (i)}$$

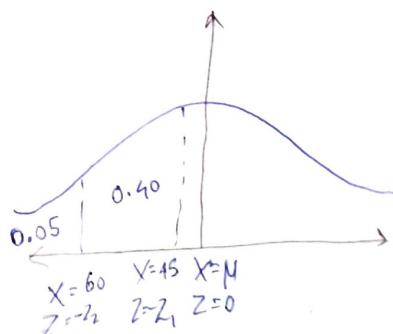
$$P(0 < Z < Z_1) = 0.43 \Rightarrow \frac{35-\mu}{\sigma} = -1.48 \quad \text{--- (ii)}$$

(i) - (ii), we get:

$$\frac{28}{\sigma} = 2.71 \Rightarrow \sigma = 10.33$$

$$\begin{aligned} \text{From (i), } \mu &= 35 + 1.4(10.33) \\ &= 50.3 \end{aligned}$$

Q) If a large group of men, 5% are under 60 inches high and 40% b/w 60 and 45. Assuming normal dist., find mean and SD.



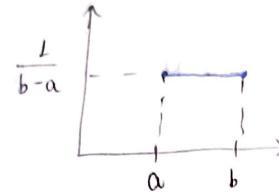
Another Distribution

I. Uniform or Rectangular Distribution

A rv 'X' is said to be a uniform dist. on interval $[a, b]$, $-\infty < a < b$

$$f(x) = \begin{cases} \frac{1}{b-a}, & a \leq x \leq b \\ 0, & \text{otherwise} \end{cases}$$

$$X \sim U[a, b]$$



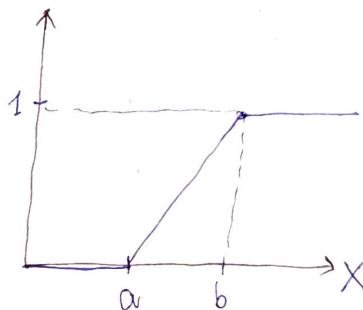
e.g.-

$$F(x) = \int_{-\infty}^x f(x) dx$$

is given by:

$$F(x) = \begin{cases} 0, & -\infty < x < a \\ \frac{x-a}{b-a}, & a \leq x < b \\ 1, & b \leq x < \infty \end{cases}$$

$$F(x)$$



$$\textcircled{1} \text{ Mean} = \frac{1}{2}(a+b)$$

$$\textcircled{2} \text{ Variance} = \frac{(a-b)^2}{12}$$

$$\text{Mean} = E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_{-\infty}^a x \cdot 0 dx + \int_a^b x \cdot \frac{1}{b-a} dx + \int_b^{\infty} x \cdot 0 dx$$

$$= \frac{b^2 - a^2}{2(b-a)}$$

$$= \frac{1}{2}(a+b)$$

$$\text{Var}(X) = E(X^2) - [E(X)]^2$$

I Exponential Distribution

$$f(x) = \begin{cases} \frac{1}{b} e^{-\frac{(x-a)}{b}}, & x \geq a \\ 0 & \text{otherwise} \end{cases}$$

where $a, b (b > 0)$ are two parameters

$$\int_{-\infty}^{\infty} f(x) dx = \int_a^{\infty} \frac{1}{b} e^{-\frac{(x-a)}{b}} dx = 1$$

① Mean = μ

② Variance = $\frac{1}{\lambda^2}$ ($\lambda = \frac{1}{\mu}$)

II Erlang Distribution (General Gamma Distribution)

$$f(x) = \begin{cases} \frac{\lambda^K \cdot x^{K-1} \cdot e^{-\lambda x}}{\Gamma(K)}, & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

III Weibull Distribution

$$\begin{aligned} f_x(x) &= e^{-r} \alpha \beta \cdot x^{\beta-1} \\ &= \alpha \beta x^{\beta-1} \cdot e^{-\alpha x \beta}, \quad x > 0 \end{aligned}$$

IV Normal distribution (Gaussian dist.)

Next Episode:

① Moment Generating function

② Central Limit Theorem

Moment Generating Function (MGF)

Def: The MGF of a r.v. 'X' is a function in a real variable given by $M_x(t) = E(e^{tx})$

① Discrete

$$M(t) = \sum_i e^{tx_i} P(X=x_i)$$

② Continuous

$$M(t) = \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

Theorem ①: If a r.v. has MGF $M_x(t)$, then about a

$$M_x^{(r)}(t) = E((x-a)^r) = r^{\text{th}} \text{ moment about } a, r=1, 2, 3, \dots$$

Theorem ②: Let x_1, x_2, \dots, x_n be the independent r.v., then M.G.F of their sum is,

$$\begin{aligned} M_{x_1+x_2+x_3+\dots+x_n}(t) &= E(e^{t(x_1+x_2+\dots+x_n)}) \\ &= E(e^{tx_1}, e^{tx_2}, \dots, e^{tx_n}) \\ &= E(e^{tx_1}), E(e^{tx_2}), \dots, E(e^{tx_n}) \\ &= M_{x_1}(t), M_{x_2}(t), \dots, M_{x_n}(t) \end{aligned}$$

MGF of Bernoulli's distribution:

$$\text{P.M.F.} \Rightarrow f_i = P(X=i) = {}^n C_i \cdot p^i (1-p)^{n-i}, i=0, 1, 2, \dots$$

$$\begin{aligned} M(t) &= E(e^{tx}) \\ &= \sum_{i=0}^n e^{ti} \cdot f_i \\ &= \sum_{i=0}^n {}^n C_i (pe^t)^i (1-p)^{n-i} \\ &= (1-p + pe^t)^n \\ &= (q + pe^t)^n \quad \text{where, } q = (1-p) \end{aligned}$$

Mean and Variance

$$M^{(1)}(t) = \frac{d}{dt} (M(t)) = q + p e^t$$

$$= n(q + p e^t)^{n-1} \cdot p e^t$$

$$= np e^t (q + p e^t)^{n-1}$$

$$M^{(2)}(t) = \frac{d^2}{dt^2} (M(t)) = np \left[e^t \cdot (q + p e^t)^{n-1} + e^t \cdot (n-1)(q + p e^t)^{n-2} \cdot p e^t \right]$$

$$= np \left[e^t (q + p e^t)^{n-1} + np e^{2t} (q + p e^t)^{n-2} - np e^{2t} (q + p e^t)^{n-2} \right]$$

$$= np e^t (q + p e^t)^{n-2} \{ (q + p e^t) + (n-1)p e^t \}$$

$$\text{Mean} = M^{(1)}(0)$$

$$= np (q + p)^{n-1}$$

$$= np \cdot 1$$

$$= np$$

$$\text{Var}(X) = M^{(2)}(0) - \{M^{(1)}(0)\}^2$$

$$= np (p+q)^{n-2} \{ q + p + (n-1)p \} - (np)^2$$

$$= np (1-p)$$

$$= npq$$

Limit Theorem

Tchebycheff's Inequality

Theorem: Let 'X' be a r.v. having mean ' μ ' and variance ' σ^2 ', then for any $\epsilon > 0$,

$$\boxed{P(|X-\mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon}}$$

Q) Show by Chebychev's Inequality that in 2000 throws with a coin the probability that the no. of heads lie b/w 900 and 1100 is at least $\frac{19}{20}$

Sol: Let 'X' denote the no. of heads

This is the case of binomial $B(n, p)$.

$$n = 2000, p = \frac{1}{2}$$

$$\mu = E(X) = np = 1000$$

$$\text{and } \text{Var}(X) = \sigma^2 = npq = 2000 \times \frac{1}{2} \times \left(1 - \frac{1}{2}\right) \\ = 500$$

$$\text{Next, } P(900 < X < 1100) = P(-100 < X - 1000 < 100)$$

$$= P(|X - 1000| < 100)$$

$$= 1 - P(|X - 1000| \geq 100) \quad \text{--- (1)}$$

Now by T-theorem,

$$P(|X - 1000| \geq 100) \leq \frac{\sigma^2}{100}$$

$$P(|X - 1000| \geq 100) \leq \underbrace{\frac{500}{(100)^2}}_{\geq} = \frac{1}{20}$$

$$\text{From eq (1), } P(900 < X < 1100) \geq 1 - \frac{1}{20} = \frac{19}{20}$$

Central Limit Theorem (C.L.T)

Let X_1, X_2, \dots, X_n be 'n' independent and identically r.v. with common μ and variance (σ^2) respectively.

$$\text{Also, } \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

Then, the distribution of \bar{X} converges to the normal distribution with μ and variance $\frac{\sigma^2}{n}$,

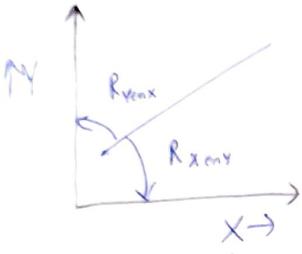
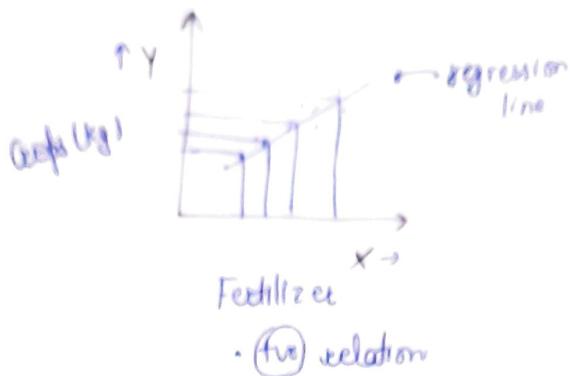
as $n \rightarrow \infty$
(large n)

$$\boxed{Z = \frac{\bar{X} - \mu}{\left(\frac{\sigma}{\sqrt{n}}\right)}}$$

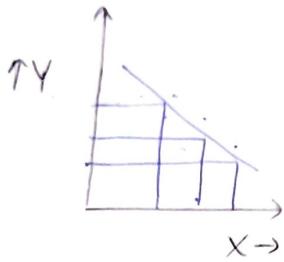
Co-relation

Relation (X, Y)

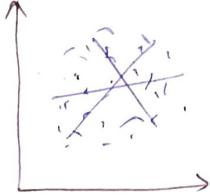
$R(X, Y)$



- regression line
- best fit
- average line
- greatest of fit

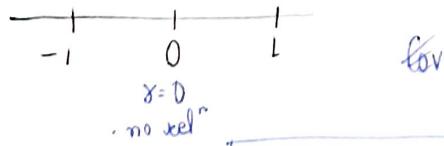
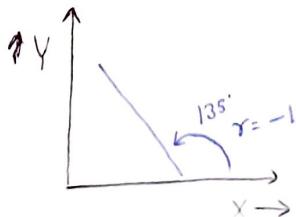
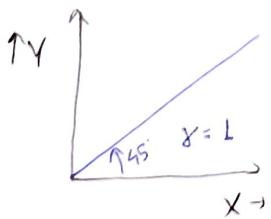


• (+ve) relation



• No relation (most points can't be covered with any line)

Coefficient of correlation

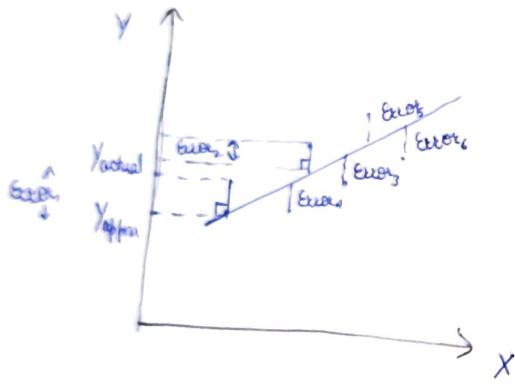


$$r = \frac{\text{Cov}(X, Y)}{\sigma_x \sigma_y}, |r| \leq 1$$

$$\text{Cov}(X, Y) = \frac{1}{n} \sum_{i=1}^n [(x_i - \bar{x})(y_i - \bar{y})]$$

$$\sigma_x = \text{Var}(X) = \frac{1}{n} (\sum x_i - \bar{x}) ; \sigma_y = \text{Var}(Y) = \frac{1}{n} (\sum y_i - \bar{y})$$

$$\begin{aligned}
 \text{Cov}(x, y) &= \frac{1}{n} \sum_{i=1}^n [(x_i - \bar{x})(y_i - \bar{y})] \\
 &= \frac{1}{n} \sum_{i=1}^n [x_i y_i - \bar{x} \bar{y} - \bar{x} y_i + \bar{x} \bar{y}] \\
 &= \frac{1}{n} \sum_{i=1}^n x_i y_i - \bar{Y} \left(\frac{\sum x_i}{n} \right) - \bar{x} \left(\frac{\sum y_i}{n} \right) + \bar{x} \bar{Y} - \frac{\sum x_i \bar{y}}{n} \\
 &= \frac{1}{n} \sum_{i=1}^n x_i y_i - \bar{X} \bar{Y} - \cancel{\bar{x} \bar{Y}} + \cancel{\bar{x} \bar{Y}} \\
 &= \frac{1}{n} \sum_{i=1}^n x_i y_i - \bar{X} \bar{Y}
 \end{aligned}$$



Let $Y = ax + b$ covers most points

$$Y_1 = a x_1 + b + e_1$$

$$Y_2 = a x_2 + b + e_2$$

{

$$Y_n = a x_n + b + e_n$$

$$e_i = Y_i - (a x_i + b)$$

$$e_i^2 = (Y_i - a x_i - b)^2$$

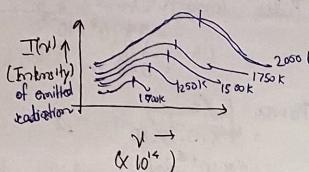
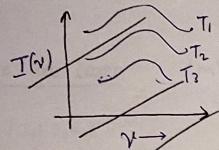
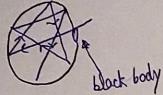
$$\mathcal{J} = \sum (e_1^2 + e_2^2 + \dots + e_n^2)$$

$$= \sum (Y_i - (a x_i + b))^2$$

$$\frac{\partial \mathcal{J}}{\partial a} = 0 \quad \frac{\partial \mathcal{J}}{\partial b} = 0$$

Quantum Mechanics

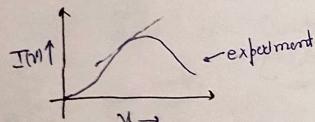
Black Body Radiation



- If temp. T_1 , total amount of emitted radiation $\propto T_1^4$
- The position of maxima shifts towards higher frequency region with temperature.

A/c Rayleigh-Jeans law,

$$I(v) = \frac{8\pi v^2 k T}{c^3} dV$$



- Satisfied at lower frequency region, not at higher frequencies
- "UV catastrophe"

A/c Planck's modification,

$$I(v) = \frac{8\pi h v^3}{c^3} \left(\frac{1}{e^{hv/kT} - 1} \right)$$

$$\text{Case I: } hv \gg kT$$

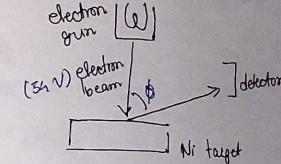
$$\text{Case II: } hv = kT$$

$$\text{Case III: } hv < kT$$

- It explained black body radiation (partially)

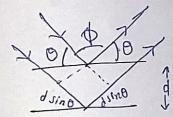
Davission-Germer Experiment

Demonstration of matter wave



Max^m intensity at $\theta = 50^\circ$

Bragg's Law



$$2\theta + \phi = 180^\circ$$

$$\Rightarrow 2\theta + 50^\circ = 180^\circ$$

$$\Rightarrow \boxed{\theta = 65^\circ}$$

$$\boxed{2d \sin \theta = m\lambda}$$

$$\boxed{d_{m1} = 0.91 \text{ \AA}}$$

Let, $m=1$

$$2d \sin \theta = \lambda$$

$$\Rightarrow \lambda = 2 \times 0.91 \times \sin 65^\circ \text{ \AA} \\ = 1.67 \text{ \AA} \quad (\text{when considered wave})$$

Chemically,

$$\boxed{\frac{1}{2}mv^2 = eV}$$

$$\text{Also, } \boxed{\lambda = \frac{h}{p}}$$

$$\boxed{\lambda = \sqrt{\frac{150}{V}}}$$

$$\Rightarrow \lambda = 1.67 \text{ \AA} \quad (\text{when considered wave particle})$$

Wave \rightarrow interference, diffraction, frequency, wavelength, polarization

Particle \rightarrow mass, velocity, momentum, blackbody radiation, PE effect

De-Broglie Hypothesis

• Wave - Particle duality

$$\lambda = \frac{h}{mv} = \frac{h}{p}$$

wave property particle property

$$E = h\nu \quad E = mc^2$$

$$\text{Kinetic Energy} = \frac{1}{2}mv^2 = \frac{p^2}{2m}$$

$$p = \sqrt{2mE}$$

$$\lambda = \frac{h}{\sqrt{2mE}}$$

Velocity of De-Broglie Wave

$$\lambda = \sqrt{\frac{150}{V}}$$

potential

$$eV = \frac{1}{2}mv^2$$

$$\lambda = \frac{h}{mv}$$

$$\lambda = \frac{12.25}{\sqrt{V}}$$

$$\Rightarrow \lambda = \frac{h}{\sqrt{2meV}} = \sqrt{\frac{150}{V}} = 12.25$$

$$h = 6.63 \times 10^{-34} \text{ Js}$$

$$c = 3 \times 10^{10} \text{ cm/s}$$

$$m_e = 9 \times 10^{-31} \text{ kg}$$

$$V = 100 \text{ V}$$

$$\lambda_0 = 1.226 \text{ Å}$$

$$m_p = 1.67 \times 10^{-27} \text{ kg}$$

$$\lambda_p = 0.28 \text{ Å}$$

• Particle heavier $\Rightarrow \lambda$ is less

Phase and Group Velocities

$$y = a \sin(\omega t - kx)$$

Here, $\omega t - kx \rightarrow \text{constant}$

$$\frac{dx}{dt} = \frac{\omega}{k} = v_r$$

$$\lambda = \frac{h}{p} \quad E = h\nu \quad v = \lambda\nu$$

$$v_p = \frac{E}{h} \times \frac{h}{p} = \frac{E}{mv} = \frac{mc^2}{mv}$$

$$\boxed{\nu_p \propto c^2}$$

$c > v_r$ $v_p > c$

The phase vel. of a de-broglie wave associated with particle moving with vel. v is greater than vel. of light c .

Group Velocities

$$y_1 = a \sin(\omega_1 t - k_1 x)$$

$$y_2 = a \sin(\omega_2 t - k_2 x)$$

After superposition,

$$y = y_1 + y_2$$

$$= a \left[2 \sin\left(\frac{(\omega_1 + \omega_2)t - (k_1 + k_2)x}{2}\right) \cos\left(\frac{(\omega_1 - \omega_2)t - (k_1 - k_2)x}{2}\right) \right]$$

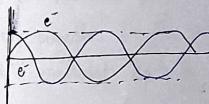
$$\text{Let, } (\omega = \frac{\omega_1 + \omega_2}{2}); k = \frac{(k_1 + k_2)}{2}$$

$$= 2a \cos(\omega t - kx) \cdot \sin\left(\frac{\Delta\omega}{2}t - \frac{\Delta k}{2}x\right)$$

$$v_g = \frac{\Delta\omega/2}{\Delta k/2} = \frac{\Delta\omega}{\Delta k}$$

$$v_g = \frac{\partial \omega}{\partial k} = \frac{\partial(2\pi\nu)}{\partial(2\pi/l)} = \lambda^2 \frac{\partial \nu}{\partial \lambda}$$

$$v_g = -\frac{\lambda^2}{2\pi} \frac{\partial \omega}{\partial \lambda}$$



maxima/minima related to v_g .

Relationship b/w v_g and v_p

$$v_g = \frac{\partial \omega}{\partial k} \quad v_p = \frac{\omega}{k}$$

$$= \frac{\partial}{\partial k} (v_p k) = v_p + k \frac{dv_p}{dk}$$

$$k = \frac{2\pi}{\lambda} \Rightarrow dk = -\frac{2\pi}{\lambda^2} d\lambda \Rightarrow \frac{k}{dk} = \frac{-\lambda}{d\lambda} \Rightarrow v_g << v_p$$

$$v_g = v_p - \lambda \frac{dv_p}{d\lambda}$$

Heisenberg's Uncertainty Principle

It is impossible to determine the exact position and momentum of a small moving particle (like e^-) simultaneously.

$$\boxed{\Delta x \cdot \Delta p \geq \frac{h}{4\pi}}$$

$$\boxed{\Delta x \cdot \Delta p \approx \frac{\hbar}{2\pi}} ; \quad \hbar = \frac{h}{2\pi}$$

$$\boxed{\Delta E \cdot \Delta t \approx \frac{\hbar}{2\pi}}$$

Applications:

① Non-existence of e^- in the nucleus.

$$\begin{aligned} \Delta p &= \frac{h}{2\pi \Delta x} \\ &= \frac{6.63 \times 10^{-34}}{2 \times 3.14 \times 10^{-14}} \\ &= 5.275 \times 10^{-21} \text{ kg m/s} \end{aligned}$$

$$E = \frac{1}{2}mv^2 = \frac{p^2}{2m} = \frac{(5.275 \times 10^{-21})^2}{2 \times 9.1 \times 10^{-31} \times 1.6 \times 10^{-19}} \text{ eV}$$

$$= 9.6 \text{ MeV}$$

The max energy that e^- can possess to exist in nucleus is 9 MeV. but it requires 9.6 MeV to reside inside the nucleus.

② Radius of Bohr's 1st orbit

$$\Delta x \cdot \Delta p \approx \hbar$$

$$KE = \frac{(\Delta p)^2}{2m} = \frac{1}{2m} \left(\frac{\hbar}{\Delta x} \right)^2$$

$$PE = \frac{1}{4\pi\epsilon_0} \cdot \frac{(Ze)(-e)}{\Delta x}$$

$$\begin{aligned} \text{Total energy } \Delta E &= KE + PE \\ &= \frac{1}{2m} \left(\frac{\hbar}{\Delta x} \right)^2 - \frac{Ze^2}{4\pi\epsilon_0 \Delta x} \end{aligned}$$

$$\text{For min/max: } \frac{\partial (\Delta E)}{\partial (\Delta x)} = 0$$

- (iii) Energy in a box of a particle in infinite potential well (of a particle)
- (iv) Ground State Energy of a linear harmonic oscillator.

③ Energy of a particle in a box (infinite potential well)

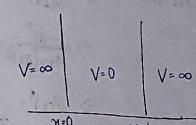
The particle is confined in a box of length 'L',

$$\therefore \Delta x \approx L$$

$$\text{Now, } \Delta x \cdot \Delta p \geq \frac{\hbar}{2} \quad (\text{Uncertainty Principle})$$

$$\Rightarrow \Delta p \geq \frac{\hbar}{2L}$$

$$\text{Also potential } V(x) = \begin{cases} 0, & 0 < x < L \\ \infty, & x < 0 \text{ or } x > L \end{cases}$$



i.e. PE=0 inside the box

$$KE = \frac{p^2}{2m}$$

Taking $p \approx \Delta p$

$$E_{min} = \frac{(\Delta p)^2}{2m}$$

$$\Rightarrow E_{min} \left(\frac{\hbar}{2L} \right)^2 \left(\frac{1}{2m} \right) = \frac{\hbar^2}{8mL}$$

* The existence of non-zero ground state energy $E = \frac{\hbar^2}{8mL}$ is a direct consequence of the uncertainty principle. If this energy were zero, then momentum will be exactly zero which'll contradict $\Delta p \geq \frac{\hbar}{2L}$.

(W) Ground State Energy of a linear harmonic oscillator

$$\Delta x \cdot \Delta p \approx \frac{\hbar}{2} \Rightarrow \Delta p \approx \frac{\hbar}{2\Delta x}$$

$$KE = \frac{(\Delta p)^2}{2m} = \left(\frac{\hbar}{2\Delta x}\right)^2 \left(\frac{1}{2m}\right) = \frac{\hbar^2}{8m(\Delta x)^2}$$

$$PE = \frac{1}{2}k(\Delta x)^2$$

$$E = KE + PE$$

$$= \frac{\hbar^2}{8m(\Delta x)^2} + \frac{1}{2}k(\Delta x)^2 \quad \text{--- (1)}$$

$$\text{For min energy, } \frac{\partial E}{\partial (\Delta x)} = 0$$

$$\Rightarrow \frac{-\hbar^2}{4m(\Delta x)^3} + k(\Delta x) = 0$$

$$\Rightarrow (\Delta x)^2 = \frac{\hbar^2}{4mk}$$

$$\Rightarrow \Delta x = \left(\frac{\hbar^2}{4mk}\right)^{1/2} \quad \text{--- (2)}$$

From (1) and (2),

$$E_{\min} = \frac{1}{2}\hbar\omega, \text{ where } \omega = \sqrt{\frac{k}{m}} \text{ (angular frequency of oscillator)}$$

↳ "Zero-point energy"

- The existence of the zero-point energy is a direct consequence of the uncertainty principle because if this ground state energy were zero, both Δx and Δp would be precisely zero, contradicting the uncertainty principle.

Wave Function

$\psi(x, y, z, t)$ → has both characters (particle & wave)
has no meaning by itself

$|\psi|^2$ → probability of finding a particle,

$$\int_{-\infty}^{\infty} |\psi|^2 dV = 1 \quad (\text{Normalization condition})$$

Properties of $\psi(x, y, z, t)$:

i) Continuous

- Finite

- Single-valued

ii) $\frac{\partial \psi}{\partial x}, \frac{\partial \psi}{\partial y}, \frac{\partial \psi}{\partial z}$ be continuous, finite, single-valued

iii) ψ should be normalized

* ψ is complex ($a+ib$) in nature.

Schrodinger's Equation

Time Independent:

By Newtonian mechanics,

$$\frac{\partial^2 \psi}{\partial x^2} = \frac{1}{v^2} \frac{\partial^2 \psi}{\partial t^2}$$

Same analogy,

$$\frac{\partial^2 \psi}{\partial x^2} + \frac{\partial^2 \psi}{\partial y^2} + \frac{\partial^2 \psi}{\partial z^2} = \frac{1}{v^2} \frac{\partial^2 \psi}{\partial t^2}$$

→ velocity of particle associated with wave

$$\psi(x, y, z, t) = \underbrace{\psi_0(x, y, z)}_{\text{time-independent amplitude}} e^{-i\omega t} = \psi_0(x, y, z) e^{-i\omega t}$$

$$\begin{aligned} \frac{\partial^2 \psi}{\partial t^2} &= (-i\omega)(-i\omega) \psi_0 e^{-i\omega t} \\ &= -\omega^2 \psi_0 e^{-i\omega t} \\ &= -\omega^2 \psi \end{aligned}$$

$$\psi = a+ib$$

$$\psi^* = a-ib, (\text{conjugate})$$

$$\text{Now, } \frac{\partial^2 \Psi}{\partial x^2} + \frac{\partial^2 \Psi}{\partial y^2} + \frac{\partial^2 \Psi}{\partial z^2} + \frac{\omega^2}{V^2} \Psi = 0, \quad \omega = 2\pi\nu$$

$$\nabla^2 \Psi + \frac{\omega^2}{V^2} \Psi = 0 \Rightarrow \frac{\omega}{V} = \frac{2\pi}{\lambda}$$

$$\Rightarrow \boxed{\nabla^2 \Psi + \frac{4\pi^2}{\lambda^2} \Psi = 0}$$

$$\text{Now } \lambda = \frac{h}{p} = \frac{h}{mv}$$

$$\Rightarrow \nabla^2 \Psi + \frac{4\pi^2 m^2 v^2}{h^2} \Psi = 0$$

$$\text{Also, Energy} = KE + PE \\ (E) \quad (V)$$

$$\Rightarrow (E - V) = \frac{1}{2} mv^2 \quad \text{from } \frac{\Psi G}{xG}, \frac{\Psi G}{yG}, \frac{\Psi G}{zG} \quad (i)$$

$$\Rightarrow \nabla^2 \Psi + \frac{8\pi^2 m}{h^2} (E - V) \Psi = 0$$

$$\Rightarrow \boxed{\nabla^2 \Psi + \frac{2m}{h^2} (E - V) \Psi = 0}$$

$$\boxed{\nabla^2 \Psi + \frac{2m E}{h^2} \Psi = 0} \quad (\text{For free particle } V=0)$$

Time-Dependent Schrödinger's eqn

$$\Psi = \Psi_0 e^{-i\omega t}$$

$$\frac{\partial \Psi}{\partial t} = -i\omega \Psi_0 e^{-i\omega t}$$

$$= -i2\pi\nu \Psi$$

$$= -i\frac{2\pi E}{\hbar} \Psi$$

$$= -\frac{iE}{\hbar} \Psi$$

$$= \frac{E}{i\hbar} \Psi$$

$$\frac{\partial \Psi}{\partial t} = \frac{E}{i\hbar} \Psi$$

$$\Rightarrow E\Psi = i\hbar \frac{\partial \Psi}{\partial t}$$

Time independent eqn:

$$\nabla^2 \Psi + \frac{2m}{\hbar^2} (E\Psi - V\Psi) = 0$$

$$\nabla^2 \Psi + \frac{2m}{\hbar^2} \left(i\hbar \frac{\partial \Psi}{\partial t} - V\Psi \right) = 0$$

$$\Rightarrow \boxed{-\frac{\hbar^2}{2m} \nabla^2 \Psi + V\Psi = i\hbar \frac{\partial \Psi}{\partial t}}$$

$$\Rightarrow \left(-\frac{\hbar^2}{2m} \nabla^2 + V \right) \Psi = i\hbar \frac{\partial \Psi}{\partial t}$$

$$\boxed{H\Psi = E\Psi}$$

$$\boxed{-\frac{\hbar^2}{2m} \nabla^2 + V = H \text{ (hamiltonian)}}$$

↑ eigenfunction
 $A\Psi = a\Psi$
 ↓ ↑
 operator eigen-value
 (energy value)

e.g. $A = \frac{\partial^2}{\partial x^2}$ and $\Psi = e^{2x}$, find eigen-values

$$A\Psi = \frac{\partial^2}{\partial x^2}(e^{2x})$$

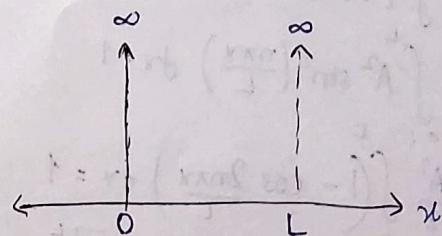
$$= 4e^{2x}$$

$$\therefore \text{Eigenvalue } (a) = 4$$

Particle in box (having infinite potential well)

• Particle is free ($V=0$)

$$V = \begin{cases} 0, & 0 \leq x \leq L \\ \infty, & x < 0 \text{ and } x > L \end{cases}$$



$$\nabla^2 \Psi + \frac{2m}{\hbar^2} (E - V) \Psi = 0$$

$$\therefore V=0$$

$$\therefore \nabla^2 \Psi + \frac{2mE}{\hbar^2} \Psi = 0$$

$$\Rightarrow \boxed{\nabla^2 \Psi + k^2 \Psi = 0} \quad \text{where, } k^2 = \frac{2mE}{\hbar^2}$$

$$\text{Let, } \Psi = A \sin(kx) + B \cos(kx)$$

$$\text{at } x=0, \Psi=0$$

$$0 = 0 + B \cos k$$

$$\Rightarrow \boxed{B=0}$$

$$\text{at } x=L, \Psi=0$$

$$0 = A \sin(kL)$$

$$\Rightarrow \boxed{kL = n\pi}, n \in \mathbb{Z}$$

$$\Rightarrow k^2 L^2 = n^2 \pi^2$$

$$\Rightarrow \frac{2mE}{\hbar^2} = \frac{n^2 \pi^2}{L^2}$$

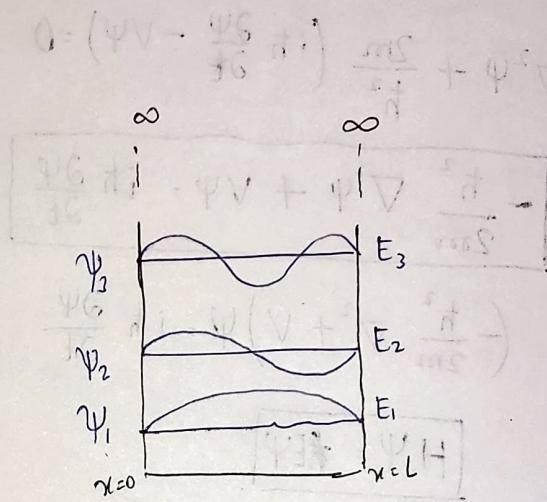
$$\Rightarrow \boxed{E_n = \frac{\hbar^2 n^2 \pi^2}{2mL^2}}$$

$$\text{at } x=0, \Psi=0$$

$$x=L, \Psi=0$$

(Particle won't exist at walls
as position will be exactly known.)

$$0 = (\Psi V - \frac{\hbar^2}{2m} \nabla^2) \frac{\sin kx}{k} + \Psi \nabla$$



$$E_1 = \frac{\hbar^2 \pi^2}{2mL^2}, E_2 = 4 \frac{\hbar^2 \pi^2}{2mL^2}, E_3 = 9 \frac{\hbar^2 \pi^2}{2mL^2}, \dots$$

\therefore Energy is quantized (discrete energy levels)

if $\Psi \rightarrow$ normalized

$$\text{then, } \int |\Psi|^2 dV = 1$$

In 1-D,

$$\int |\Psi|^2 dx = 1$$

$$\Rightarrow \int_0^L A^2 \sin^2 \left(\frac{n\pi x}{L} \right) dx = 1$$

$$\Rightarrow \frac{A^2}{2} \int_0^L \left(1 - \cos \frac{2n\pi x}{L} \right) dx = 1$$

$$\Rightarrow \frac{A^2}{2} \left[x - \frac{\sin \left(\frac{2n\pi x}{L} \right)}{\left(\frac{2n\pi}{L} \right)} \right]_0^L = 1$$

$$\Rightarrow \frac{A^2}{2} \left[L - \frac{L}{2n\pi} (0 - 0) \right] = 1$$

$$\Rightarrow A = \pm \sqrt{\frac{2}{L}}$$

$$\Psi = A \sin(kx)$$

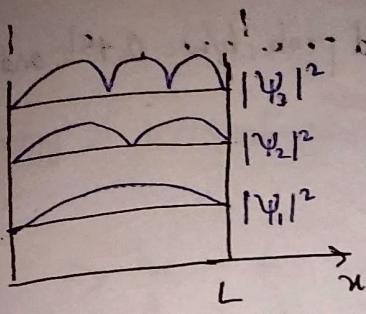
$$k = \frac{n\pi}{L}$$

$$\Psi = A \sin \left(\frac{n\pi x}{L} \right)$$

$$\therefore \boxed{\Psi_n = \sqrt{\frac{2}{L}} \sin \left(\frac{n\pi}{L} x \right)}$$

normalized wave function

$$\boxed{E_n = \frac{\hbar^2 n^2 \pi^2}{2mL^2}}$$



$|\psi_i|^2 \rightarrow$ probability of finding particle

if $E = E_1$, then max^m probability of finding particle = $\frac{1}{2}$

$$E = E_2, " " " " " = \frac{L}{\sqrt{3}}, \frac{3L}{\sqrt{3}}$$

Q) Find the probability that particle present in a box 'L' wide can be found b/w $0.45L$ and $0.55L$ for ground state and first excited state.

Ans: For ground state,

$$0.55L \quad \psi_1 = \sqrt{\frac{2}{L}} \sin\left(\frac{\pi}{L}x\right)$$

$$\int_{0.45L}^{0.55L} |\psi_1|^2 dx = \sqrt{\frac{2}{L}} \int_{0.45L}^{0.55L} \sin^2\left(\frac{\pi}{L}x\right) dx$$

$$= \frac{2}{L} \int_{0.45L}^{0.55L} \left(1 - \cos\left(\frac{2\pi x}{L}\right)\right) dx$$

$$= \frac{2}{L} \left[(0.10L) - \frac{L}{2\pi} \left(\sin\left(\frac{2\pi}{L}(0.55L)\right) - \sin\left(\frac{2\pi}{L}(0.45L)\right) \right) \right]$$

For $n=2$,



Q) If $\Psi = \alpha x$ and $x=0$ to $x=1$, then find prob. b/w 0.45 and 0.55.

$$\int |\Psi|^2 dx = 1$$

$$\Rightarrow \alpha^2 \int x^2 dx = 1$$

$$\Rightarrow \frac{\alpha^2}{3} = 1$$

$$\Rightarrow \alpha = \pm \sqrt{3}$$

Now,

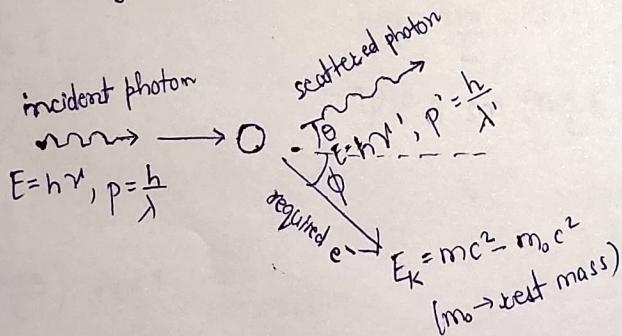
$$\int_{0.45}^{0.55} |\Psi|^2 dx = 3 \int_{0.45}^{0.55} x^2 dx$$

$$= \frac{3}{3} \left[x^3 \right]_{0.45}^{0.55}$$

=

Compton Scattering

- coherent scattering: energy absorbed = energy released (higher wavelengths)
- incoherent scattering: energy absorbed > energy released (lower wavelengths)



$$m = \frac{m_0}{\sqrt{1 - \frac{v^2}{c^2}}}$$

$p' \sin \theta$

$p' \cos \theta + p'' \cos \phi$

$p'' \sin \theta$

A/c conservation of energy,

Energy of incident photon = Energy of scattered photon +
Energy of required e-

$$\Rightarrow \frac{hc}{\lambda} = \frac{hc}{\lambda'} + (m - m_0)c^2$$

$$\Rightarrow hc \left(\frac{1}{\lambda} - \frac{1}{\lambda'} \right) = (m - m_0)c^2$$

$$\Rightarrow h \left(\frac{1}{\lambda} - \frac{1}{\lambda'} \right) = (m - m_0)c \quad \text{--- (1)}$$

A/c conservation of momentum,

$$\frac{h}{\lambda'} \sin \theta = \frac{m_0 v}{\sqrt{1 - \frac{v^2}{c^2}}} \quad \text{--- (II)}$$

$$\frac{h}{\lambda} = \frac{h}{\lambda'} \cos \theta + \frac{m_0 v}{\sqrt{1 - \frac{v^2}{c^2}}} \cos \phi \quad \text{--- (III)}$$

$$\boxed{\Delta \lambda = \lambda' - \lambda = \frac{h}{m_0 c} (1 - \cos \theta)}$$

$$\boxed{\frac{h}{m_0 c} \approx 0.0242 \text{ A}^\circ}$$

$$\theta = 0^\circ \Rightarrow \Delta \lambda = 0$$

$$\theta = 90^\circ \Rightarrow \Delta \lambda = \frac{h}{m_0 c}$$

$$\theta = 180^\circ \Rightarrow \Delta \lambda = \frac{2h}{m_0 c}$$

$\Delta \lambda \rightarrow$ compton wavelength shift

$$\% \text{ compton shift for } \lambda = 4000 \text{ A}^\circ \text{ (visible region)} = \frac{(\Delta \lambda)_{\max}}{\lambda} \times 100\%$$

$$= 0.001\% \text{ (Not observed)}$$

• Significant for X-rays ($\lambda \approx 1.5 \text{ A}^\circ$)