6 Inference Based On Normal Distribution

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[1]Thanks to my family, my friend and freedom.

1 Introduction

Finding probability distributions to describe—and, ultimately, to predict—empirical data is one of the most important contributions a statistician can make to the research scientist. Already we have seen a number of functions playing that role.

2 Estimatior of variance σ^2

Suppose that a random sample of n measurements, $Y_1, Y_2, \cdots, Y_n$, is to be taken on a trait that is thought to be normally distributed, the objective being to draw an inference about the underlying pdf' s true mean, μ. If the variance σ^2 is known, we already know how to proceed.

In practice, though, the parameter σ^2 is seldom known, so the ratio $\frac{\overline{Y}-\mu}{\sigma/\sqrt{n}}$ can not be calculated. The usual estimator for the population variance,

$$S^2 = \frac{1}{n-1}\sum_{i=1}^{n}(Y_i - \overline{Y})^2$$

whose deduction is at [2].

Historically, many early practitioners of statistics felt that replacing σ with S had, in fact, no effect on the distribution of the Z ratio. Sometimes they were right. If the sample size is very large (which was not an unusual state of affairs in many of the early applications of statistics), the estimator S is essentially a constant and for all intents and purposes equal to the true σ.

The ratio $\frac{\overline{Y}-\mu}{\sigma/\sqrt{n}}$ is called the **Student t distribution**.

3 Chi square distribution

To introduce the Student t distribution, we first introduce the **chi square distribution**.

Theorem 3.1. *Let $U = \sum_{j=1}^{m} Z_j^2$, where $\{Z_j\}$ are independent standard normal random variables. Then U has a gamma distribution with $r = \frac{m}{2}$ and $\lambda = \frac{1}{2}$. That is,*

$$f_U(u) = \frac{1}{2^{m/2}\Gamma\left(\frac{m}{2}\right)} u^{(m/2)-1} e^{-u/2}, \quad u \geq 0$$

which we called the chi square distribution with m degrees of freedom.

Theorem 3.2. *Let $Y_1, \cdots, Y_n$ be a random sample from a normal distribution with mean μ and variance σ^2. Then*

1. *S^2 and $\overline{Y}$ are independent.*
2. *$\frac{(n-1)S^2}{\sigma^2} = \frac{1}{\sigma}\sum_{i=1}^{n}(Y_i - \overline{Y})^2$ has a chi square distribution with $n-1$ degrees of freedom.*

Proof. The proof is at [1]. □

Suppose that U and V are independent chi square random variables with n and m degrees of freedom, respectively. A random variable of the form $\frac{V/m}{U/n}$ is said to have an F distribution with m and n degrees of freedom.

Theorem 3.3. *Suppose $F_{m,n} = \frac{V/m}{U/n}$ denotes an F random variable with m and n degrees of freedom. The pdf of $F_{m,n}$ has the form*

$$f_F(w) = \frac{\Gamma\left(\frac{m+n}{2}\right) m^{m/2} n^{n/2} w^{(m/2)-1}}{\Gamma\left(\frac{m}{2}\right)\Gamma\left(\frac{n}{2}\right)(n+mw)^{(m+n)/2}}$$

4 Student t distribution

Now, we introduce the formal definition of **Student t distribution** derived from *chi square distribution*.

Theorem 4.1. *Let Z be a standard normal random variable and let U be a chi square random variable independent of Z with n degrees of freedom. The Student t ratio with n degrees of freedom is denoted T_n, where*

$$T_n = \frac{Z}{\sqrt{U/n}}$$

We review the Student t distribution $\frac{\overline{X}-\mu}{\sqrt{S/n}}$, and we easily get that it is a Student t distribution with $n-1$ degrees of freedom.

The pdf of T_n is given by the theorem.

Theorem 4.2. *The pdf of a Student t random variable with n degrees of freedom is given by*

$$f_{T_n}(t) = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{n\pi}\,\Gamma\left(\frac{n}{2}\right)\left(1+\frac{t^2}{n}\right)^{(n+1)/2}}$$

BTW, the estimation of μ of a Student t distribution is the same as the case of normal distribution. The only differencs is the table.

In the case that X is not normally distributed, we may use computer to calculate the integral.

4.1 Estimation of σ^2

Since we have know that

$$\frac{(n-1)S^2}{\sigma^2} = \frac{1}{\sigma^2}\sum_{k=1}^{n}(Y_k - \overline{Y})^2$$

which satisfies *chi square distribution*, the confidence interval of σ is derived from the equation

$$P\left[\mathcal{X}^2_{\alpha/2,n-1} \leq \frac{(n-1)S^2}{\sigma^2} \leq \mathcal{X}^2_{1-\alpha/2,n-1}\right] = 1-\alpha$$

where $\mathcal{X}^2_{p,n}$ denotes that in the chi square distribution with n degrees of freedom, $P(x \leq \mathcal{X}^2_{p,n}) = p$.

Theorem 4.3. *Let s^2 denote the sample variance calculated from a random sample of n observations drawn from a normal distribution with mean μ and variance σ^2. Let $\mathcal{X}_2 = (n-1)s^2/\sigma_0^2$.*

1. *To test $H_0 : \sigma^2 = \sigma_0^2$ versus $H_1 : \sigma^2 > \sigma_0^2$ at the α level of significance, reject H_0 if $\mathcal{X}_2 \geq \mathcal{X}^2_{1-\alpha,n-1}$.*
2. *To test $H_0 : \sigma^2 = \sigma_0^2$ versus $H_1 : \sigma^2 < \sigma_0^2$ at the α level of significance, reject H_0 if $\mathcal{X}_2 \leq \mathcal{X}^2_{\alpha,n-1}$.*
3. *To test $H_0 : \sigma^2 = \sigma_0^2$ versus $H_1 : \sigma^2 \neq \sigma_0^2$ at the α level of significance, reject H_0 if $\mathcal{X}^2 \leq \mathcal{X}^2_{\alpha/2,n-1}$ or $\mathcal{X}_2 \geq \mathcal{X}^2_{1-\alpha/2,n-1}$.*

References

[1] jld. *Independence of meean and estimator*. `https://stats.stackexchange.com/questions/344960/showing-s2-and-overliney-are-independent-seeking-a-solution-to-this-tex`. May 2018.

[2] Wikipedia. *Bias of an estimator*. `https://en.wikipedia.org/wiki/Bias_of_an_estimator`. June 2020.