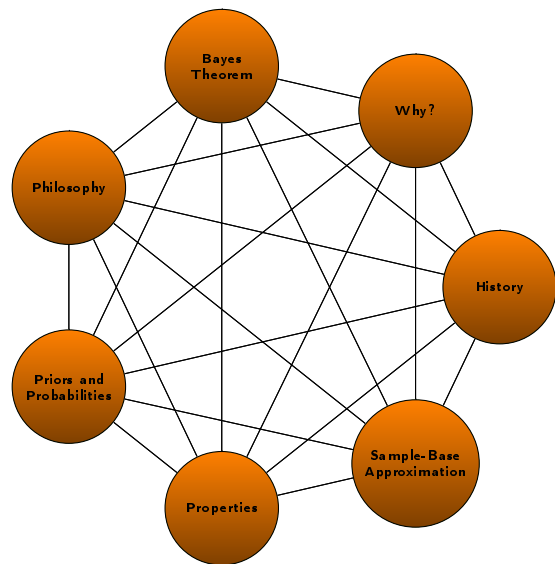


# Introduction to Bayesian Inference

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# Why?

Why are you interested in Bayesianism?

# Why? – some advantages

## Things Ecologist say...

- small sample sizes: exact inference
- missing data: easy to impute
- integrate other information
- derived quantities
- complex, hierarchical process models

multiple sources of variation (space,time)  
"honest" epistemology

## Theoretical

- conditional on the observed data
  - probabilistic statements
  - evidential
  - coherence, decision making
  - good frequentist properties
- shrinkage, decision-theory
- model selection

- what is probability (basing inference on something that doesn't exist!!!)?
- objective basis for science?
- misalignment: probability theory and human psychology
- biased (towards the prior)<sup>1</sup>
- language dependence

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<sup>1</sup>A Bayesian would claim that given a prior, it would be irrational to believe in anything other than the posterior.

## Neo Bayesian Revival (>1992)

Gelfand and Smith 1992 - Sampled based approximations of Bayesian posteriors  
2

## Revival (~1920s - )

### Subjective Bayesian, Decision Theory

- Ramsey (1926)
- De Finetti (1937)
- Savage (1954)
- (Wald, 1939, 1954)

### Hypothesis Testing, Logical/Objective Bayesism

- Jeffreys (1939)
- Jaynes (2003)

### Hierarchical Bayesian

- Good (1953,65)

### Relationship to Compact Coding Theory

- Rissanen (1978), Wallace (1968)

### Prediction

## Frequentist "lethal blow"<sup>3</sup>

Rallied against use of prior probabilities in statistical inference

- Sir Ronald A. Fisher (1925,1935,...)

Maximum likelihood, significance testing, ANOVA, sufficiency, randomized experiments

*Inductive inference*

- Jerzy Neyman & Egon Pearson (1933)

Hypothesis testing, confidence intervals, Type-I/II error rates

*Inductive Behaviour*

## Philosophical developments

- Karl Popper (1959,1963) and **Falsificationism**

Anti-Induction: scientific progress is by **falsifying** theories, only

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<sup>3</sup>S. Zabell 1989

frequentists:

- reject probabilistic confirmation of models <sup>4</sup>
- reject Bayesian notions of probability
- frequentists care about *good frequency properties*

Estimation: unbiased, efficiency, obtain minimum variance

Hypothesis testing: Type-I error rates, most powerful tests<sup>5</sup>

## Frequencies as probabilities

- probabilities only meaningful as long-run frequencies of events

$$\mathbf{Y}: \{H, T, H, H, H, T, H, T, T, H, T, H, T, \dots\} \quad (1)$$

- The probability of flipping a coin and getting a head is...

$$p(y = \text{Head}) \equiv \lim_{n \rightarrow \infty} \frac{1}{n} \sum_i^n \mathbb{I}[y_i = \text{Head}] \quad (2)$$

the mean counts of heads in the long-run

<sup>4</sup>Some use the AICc for model confirmation

<sup>5</sup>Neyman-Pearson-type testing, not Fisherians



## Frequentism

### Fisher's p-value

- continuous index of evidence **against** a hypothesis  $H$
- **NEVER** prove a hypothesis  $H$ , only disprove

### Neyman-Pearson $\alpha$ and $\beta$

- long-run error rates of Type-I and Type-II
- **bound** Type-I at  $\alpha \leq 0.05$  and hopefully maximize power  $(1 - \beta)$  with high  $n$  and most powerful tests
- never confirm a hypothesis: only **act** so as to "not be wrong too often"

## Bayesians

### Model probabilities

probabilistic confirmation of hypotheses

- $p(H_k|Y)$  what is the probability of Hypothesis  $H_k$  given the data?

### Bayes Factors

evidence in favour of one hypothesis over another

- $BF = \frac{p(Y|H_1)}{p(Y|H_2)}$ .

find hypothesis that is more likely to be true

# History - Inverse Probability

from late 1700's to ~1920's: Method of Inverse Probability, the bread and butter of applied analyses

- Rev. Thomas Bayes (1778)
- Simone-Pierre Laplace (1774)

## Bayes

"PROBLEM: Given the number of times in which an unknown event ( $y \in [0, 1]$ ) has happened and failed: Required the **chance** that the probability of its happening in a single trial lies between any two degrees of probability that can be named. . . By *chance* I mean the same as *probability*."

- *probability of a probability*:  $p(\theta|y)$
- two types of probabilities

## Laplace

developed Bayes Theorem close to its modern form:

$$\underbrace{p(\theta|Y)}_{\text{posterior}} = \frac{\overbrace{p(\theta) \mathcal{L}(Y|\theta)}^{\text{prior likelihood}}}{\underbrace{\int_{\theta} p(\theta) \mathcal{L}(Y|\theta) d\theta}_{\text{marginal likelihood}}} \quad (3)$$

Bayes great innovation: two types of probability

Probability of an *observable* event  $Y$ :

$$p(y=\text{Head}) \equiv \theta$$

- $\theta$  is like a *parameter in a model*
- $p(y=\text{Head}|\theta) = \text{Bern}(y; \theta) \rightarrow$  What we now call a likelihood
- how the data was generated:  $y \sim \text{Bern}(\theta)$

Probability distribution for the *parameter*  $p(\theta)$

for inference...

- **Before data:**  $p(\theta)$  (the *prior* probability distribution)
- **After data:**  $p(\theta|Y)$  (the *posterior* probability distribution)

- conditional probability: want a probability distribution  $p(\theta|y)$  conditional on observed data  $y$  -> need a likelihood  $f(Y|\theta)$  and prior probability distribution  $p(\theta)$ .

$$p(\theta|Y) = \frac{p(\theta)\mathcal{L}(Y|\theta)}{\int_{\theta} p(\theta)\mathcal{L}(Y|\theta)d\theta} \text{ where } \dots$$

$p(\theta) \equiv$  prior information (before the data)

$\mathcal{L}(Y|\theta) \equiv$  likelihood (information from the data)

$p(\theta|Y) \equiv$  distribuion of  $\theta$  after the data 6 (4)

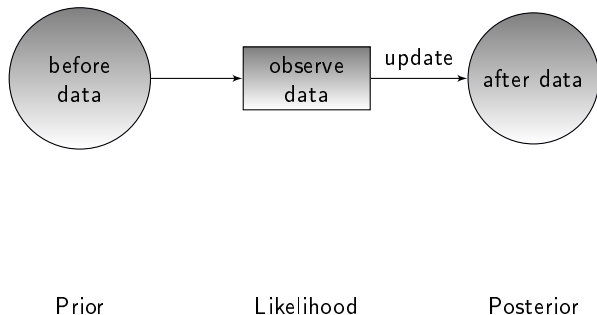
denominator : ~~marginal likelihood~~ (often ignore)

more common form...

$$p(\theta|Y) \propto \mathcal{L}(Y|\theta)p(\theta)$$

- posterior is a mixture of information in **prior** and **likelihood**

- posterior is a mixture of information in **prior** and **likelihood**



## Mixture of information

- $\mathcal{L}(Y|\theta)$ : Likelihood, specified by model. Similar between Bayesian and non-Bayesian analyses<sup>7</sup>
- $p(\theta)$  ... where do they come from?

## How to specify priors (HUGE topic)

- a previous posterior distribution
- elicitation from experts, previous studies
- Priors as degrees-of-beliefs: **Subjectivist/personalist** Bayesians
- Default prior and reference priors: **Objective/logical** Bayesians
- adhoc

---

<sup>7</sup>Frequentists reserve the term likelihood for a function of  $\theta$  for fixed  $y$ , whereas Bayesians consider "joint probability density of the data" given  $\theta$ .

# Posterior Inference (for estimation $\theta$ )

- Probabilistic statements about abstract quantity ( $\theta$ ) (*only* Bayesians can do)
- Posterior probability necessarily depends on a *prior*

“to make an Omelette, you must crack a few eggs” (Savage)

## The joy of Posterior Inference

can make statements like...

- what is the probability that  $\theta > 0$ ?
- what is the most probable value of  $\theta$ ? (**MAP**)
- what is the expected value of  $\theta$ ? (**posterior mean**)
- what is a *high probability region* of  $\theta$  (**Q% credibility interval**)

## Example 1:

- men's height,  $n=20$  observations.

- $y_i \sim \mathcal{N}(175, 10^2)$

```
y <- c(183.46, 182.32, 178.31, 181.36, 165.12, 185.68, 170.47,  
178.11, 174.86, 182.03, 180.09, 172.88, 177.94, 177.26, 182.58,  
171, 173.74, 177.78, 180.02, 163.05)
```

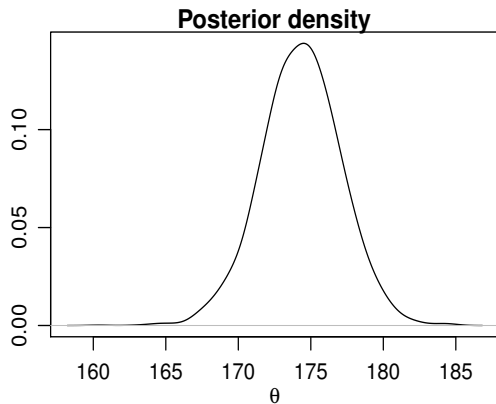
- estimate  $\theta = [\mu, \sigma^2]$ : mean population height and variance
- priors:  $p(\mu) = \mathcal{N}(0, 90^2)$ ,  $p(\sigma^2) = \mathcal{IG}(0.1, 0.1)$
- specify a likelihood:  $\mathcal{L}(\mathbf{y}|\mu, \sigma^2) = \prod_i^n \mathcal{N}(y_i; \mu, \sigma^2)$

*now run a Gibbs sampler to approximate the posterior  $p(\mu, \sigma^2|\mathbf{y}) \dots$*



# Why are Posteriors so Great!

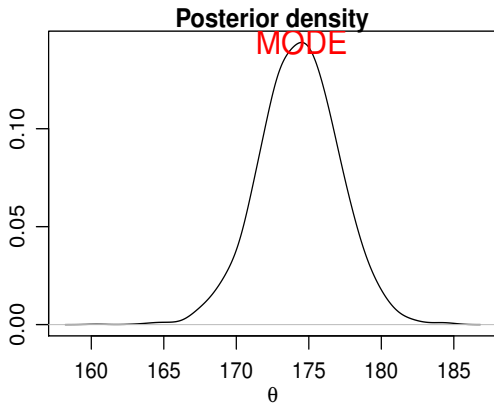
- IS a probability distribution



- easy to interpret

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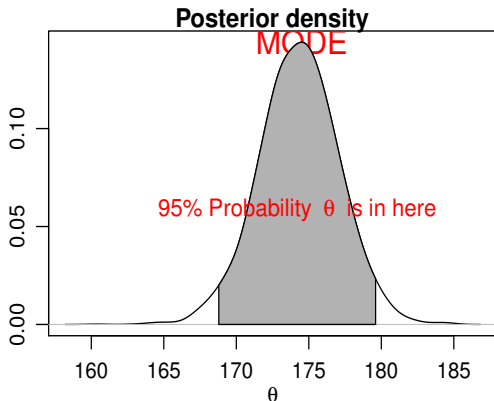
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- Posterior mode: most probable value
- Posterior mean  $\mathbb{E}[\theta] = \int p(\theta|Y)\theta d\theta$ : expected value

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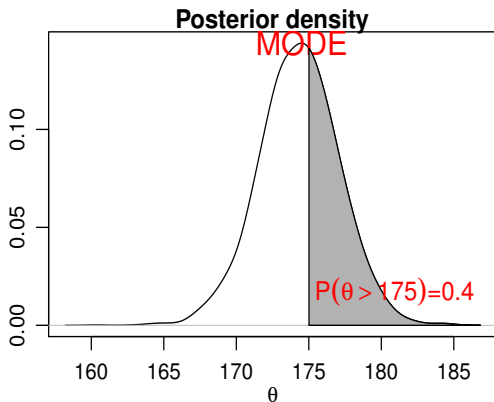
- IS a probability distribution



- Posterior mode: most probable value
- Posterior mean  $\mathbb{E}[\theta] = \int p(\theta|Y)\theta d\theta$ : expected value
- 95%CI of  $\theta$

# Why are Posteriors so Great!

- IS a probability distribution



- Posterior mode: most probable value
- Posterior mean  $\mathbb{E}[\theta] = \int p(\theta|Y)\theta d\theta$ : expected value
- What is the probability that  $\theta > X$ ? Area of  $p(\theta|Y) > X$

Bayesian vs. frequentist estimates: compare posteriors to maximum-likelihood method

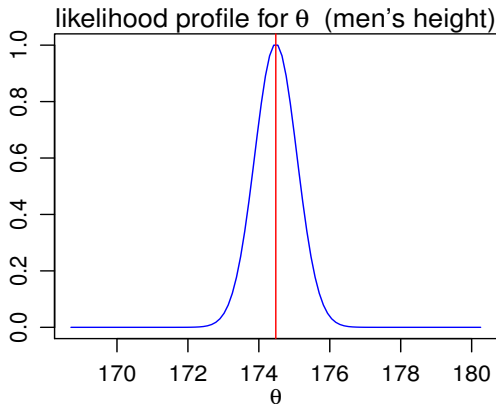
## method of maximum likelihood

- Choose  $\theta$  such that we *maximize* the likelihood ( $\mathcal{L}$ ) of seeing  $y$
- **interpretation** "It would be very (un)likely to see the data that I saw, if the value of  $\theta$  were  $X$ "
- Most common method among Frequentists (single model estimation)
- $\hat{\theta}_{\text{MLE}}$  is **NOT** the "most probability value of  $\theta$ "
- **optimality**: unbiased, efficient <sup>8</sup>

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<sup>8</sup>but see shrinkage estimators for high-dimensional problems

# Posterior Estimation vs Maximum Likelihood Estimation



- frequentist point estimates: `glm(y~1)`

MLE	se	95CI-low	95CI-hi
172.3	1.48	169.4	175.17

- compare to (approximate)<sup>9</sup> Posterior descriptive statistics

$E[\theta]$	SD	95CI-low	95CI-hi
172.21	1.51	169.21	175.16

nearly the same

## Example 2:

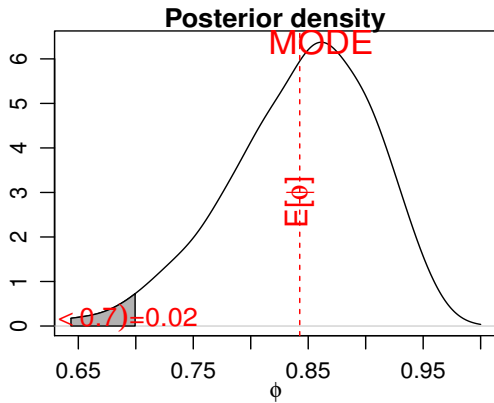
- survival [0 = died, 1 = survived],  $n = 30$  observations.
- $s_i \sim \text{Bern}(0.9)$

`s <-`

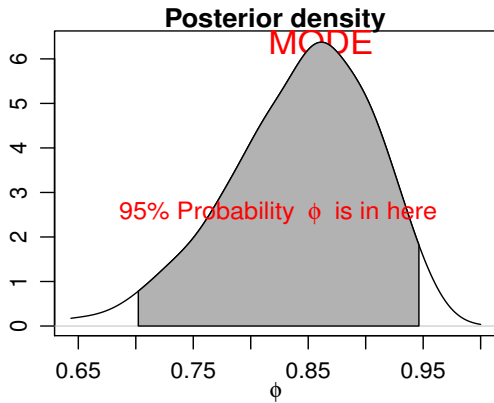
`c(1,1,1,1,0,1,0,1,1,1,1,1,1,1,1,0,1,0,1,1,1,1,1,1,1,1,1,1,1,1)`

- estimate  $\theta = [\phi]$ : mean population survival
- priors:  $p(\phi) = \text{Beta}(1, 1)$
- specify a likelihood:  $\mathcal{L}(\mathbf{s}|\phi, n_s) = \prod_s^n \text{Bern}(s_i; \phi, n_s)$

now run a Gibbs sampler to approximate the posterior  $p(\phi|\mathbf{s})$

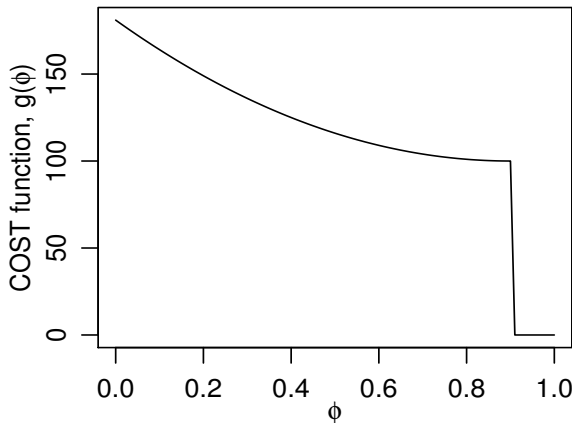




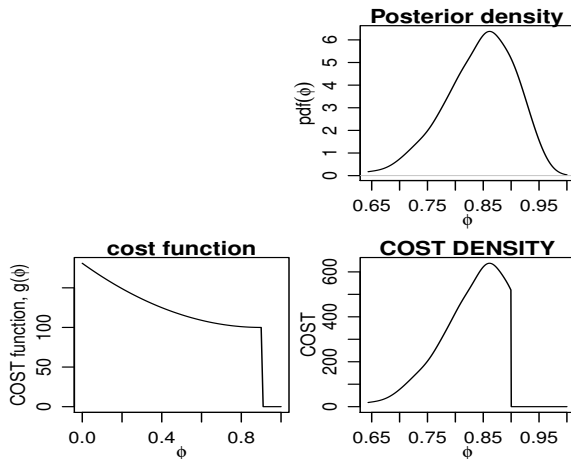


## Posterior inference: cost functions

What if you have a “cost function”  $g(\phi)$ ? e.g., cost of conservation action conditional on the estimated values of  $\phi$ ?



# Posterior inference: cost functions



$$\mathbb{E}[\text{COST}] = \int_0^1 g(\phi) p(\phi|s) d\phi$$

- full cost including all uncertainty in  $\phi$

How do Bayesian posterior estimates compared to more-familiar (frequentist) point-estimates based on maximum likelihood?

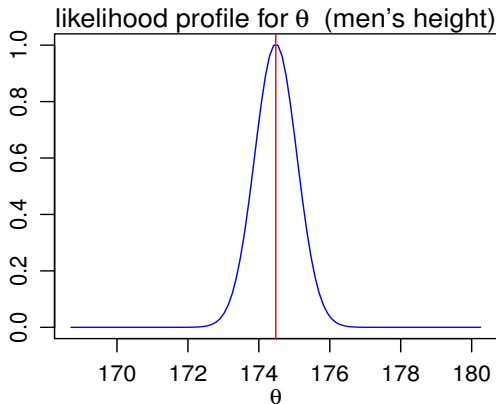
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# Posterior Estimation vs Maximum Likelihood Estimation

for  $n$  getting **LARGE**, and for **WEAK** priors

- Posterior Mode  $\theta_{\text{MAP}} \rightarrow \hat{\theta}_{\text{MLE}}$
- Posterior Confidence Intervals  $\rightarrow$  Confidence Intervals

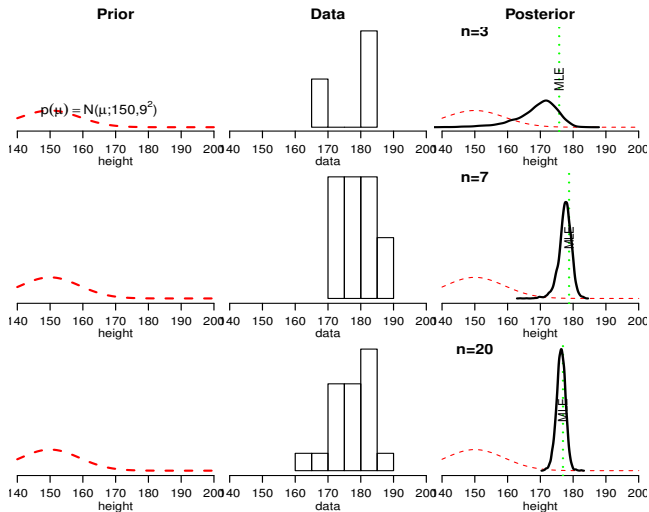
for **low**  $n$  and/or for **STRONG** priors

- shrinkage:  $\theta \rightarrow$  Prior expectation.
- Posterior mean  $\bar{\theta}$  is “biased” towards the priors

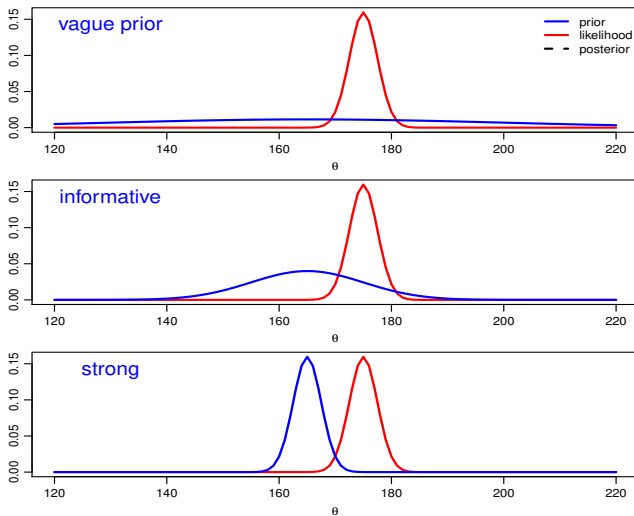
role of priors (from an estimation perspective)

- Priors retard/accelerate rate of convergence of  $\bar{\theta} \rightarrow$  truth
- At low samples-sizes, “sensible” priors induce *shrinkage* and have better estimation properties than MLEs
- **Key POINTS**: you must be a master of prior distributions.

# Posteriors and Sample Size

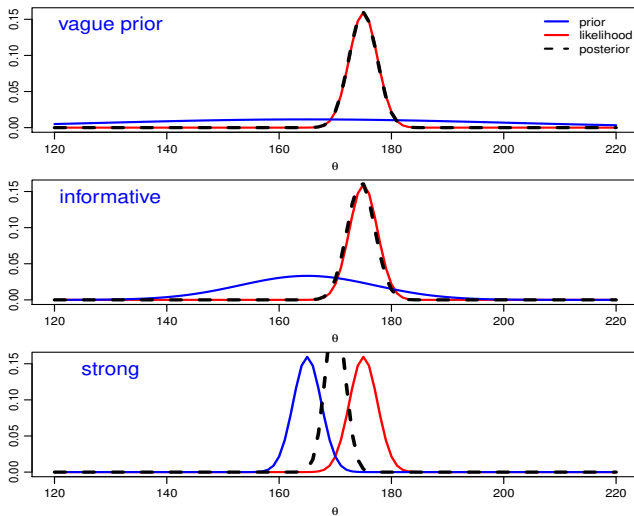


# Posteriors and Prior information





# Posteriors and Prior information



## Most Biologists are reluctant Bayesians

- **Frequentist** vs. **Bayesian**: often desire that point estimates are identical between posteriors and MLEs
- **but**, only for: i) weak priors, and ii) large-samples sizes
- key point: **Be a Master of Priors!**

## Subjective Personalist Bayesians

Probabilities are your “degree of belief”

- priors: prior beliefs
- posteriors: bring your beliefs into alignment with posterior
- **decision making**

## Objective Logical Bayesians

Probabilities are continuous extension of Aristotean logic, deductive

- Probabilities capture “degree of truth”
  - Priors: non-informative, set by **default** (Jeffrey’s Priors, reference priors, language-invariant priors)
- e.g.,  $p(\phi) = \text{Beta}(0.5, 0.5)$  (Jeffrey’s Prior)

- Elicit priors from previous studies (posterior becomes new prior)

## Instrumentalist

priors useful for good estimation properties

- shrinkage, efficiency

## Frequentist

Principal principle: "probabilities ( $p(\text{Event})$ ) should align with long-run frequencies of Event"

- probabilities do not exist in reality

## Other

- Quantum mechanics
- Propensities (Karl Popper)

# Probability Distributions in JAGS/BUGS

- you must express your prior information **probabilitistically**

Know the distributions and their parameters (JAGS Manual)

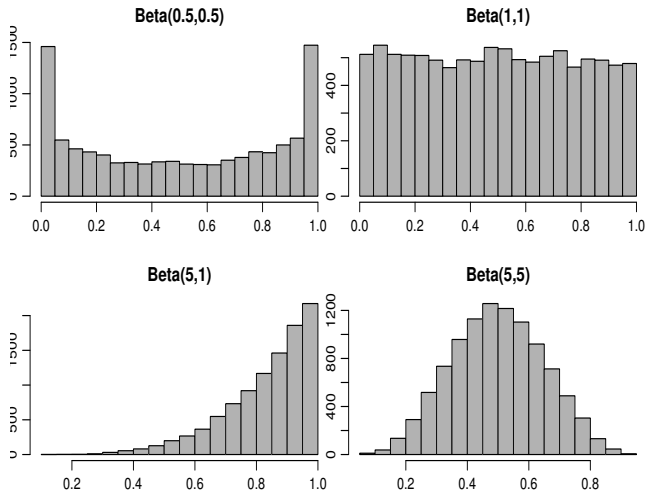
Name	Usage	Density	Lower	Upper
Beta	<code>dbeta(a,b)</code> $a > 0, b > 0$	$\frac{x^{a-1}(1-x)^{b-1}}{\beta(a,b)}$	0	1
Chi-square	<code>dchisqr(k)</code> $k > 0$	$\frac{x^{\frac{k}{2}-1} \exp(-x/2)}{2^{\frac{k}{2}} \Gamma(\frac{k}{2})}$	0	
Double exponential	<code>ddexp(mu,tau)</code> $\tau > 0$	$\tau \exp(-\tau x - \mu )/2$		
Exponential	<code>dexp(lambda)</code> $\lambda > 0$	$\lambda \exp(-\lambda x)$	0	
F	<code>df(n,m)</code> $n > 0, m > 0$	$\frac{\Gamma(\frac{n+m}{2})}{\Gamma(\frac{n}{2})\Gamma(\frac{m}{2})} \left(\frac{n}{m}\right)^{\frac{n}{2}} x^{\frac{n}{2}-1} \left\{1 + \frac{nx}{m}\right\}^{-\frac{(n+m)}{2}}$	0	
Gamma	<code>dgamma(r, lambda)</code> $\lambda > 0, r > 0$	$\frac{\lambda^r x^{r-1} \exp(-\lambda x)}{\Gamma(r)}$	0	
Generalized gamma	<code>dgen.gamma(r, lambda, b)</code> $\lambda > 0, b > 0, r > 0$	$\frac{b\lambda^{br} x^{br-1} \exp\{-(\lambda x)^b\}}{\Gamma(r)}$	0	
Logistic	<code>dlogis(mu, tau)</code> $\tau > 0$	$\frac{\tau \exp\{(x - \mu)\tau\}}{[1 + \exp\{(x - \mu)\tau\}]^2}$		
Log-normal	<code>dlnorm(mu,tau)</code> $\tau > 0$	$\left(\frac{\tau}{2\pi}\right)^{\frac{1}{2}} x^{-1} \exp\left\{-\tau(\log(x) - \mu)^2/2\right\}$	0	
Noncentral Chi-squre	<code>dnchisqr(k, delta)</code> $k > 0, \delta \geq 0$	$\sum_{r=0}^{\infty} \frac{\exp(-\frac{\delta}{2})(\frac{\delta}{2})^r}{r!} \frac{x^{(k/2+r-1)} \exp(-\frac{x}{2})}{2^{(k/2+r)} \Gamma(\frac{k}{2}+r)}$	0	
Normal	<code>dnorm(mu,tau)</code> $\tau > 0$	$\left(\frac{\tau}{2\pi}\right)^{\frac{1}{2}} \exp\{-\tau(x - \mu)^2/2\}$		
Pareto	<code>dpar(alpha, c)</code> $\alpha > 0, c > 0$	$\alpha c^\alpha x^{-(\alpha+1)}$	c	

# Intuiting Probability Distributions

- easy to learn in R

e.g.,  $r \sim \text{Beta}(a, b)$

```
r <- rbeta(10000, 0.5, 0.5)
```



## Posteriors

often no 'analytical' solution to  $(\theta|Y)$

## Solution: Sampling

- it is a Probability Distribution!!!
- find a way to sample from posterior
- with enough samples:  $\text{mean}(\text{samples}) = \text{Posterior Expectation}$

assuming  $\theta_j \sim p(\theta|y)$  for  $j = 1, \dots, J$

Expected Value	$= \int \theta p(\theta y) d\theta$	$\approx \frac{1}{J} \sum_j^J \theta_j$
Standard Error( $\theta$ )	$= SE(\theta)$	$\approx SD(\theta_j)$
Probability $\theta > 0$	$= \int \mathbb{I}[\theta > 0] p(\theta y) d\theta$	$\approx \frac{1}{J} \sum_j^J \mathbb{I}[\theta_j > 0]$

## Sampling Algorithms

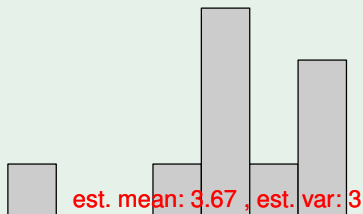
MCMC; Gibbs Sampling; Metropolis-Hastings; Slice-Sampling; Importance Sampling; "Conjugate Priors";

# Approximate the joint-posterior distribution"

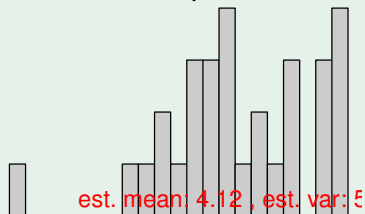
example: estimate mean and variance of  $\theta$

$$\theta_{\text{true}} = 3.44; \text{Var}(\theta)_{\text{true}} = 4.89$$

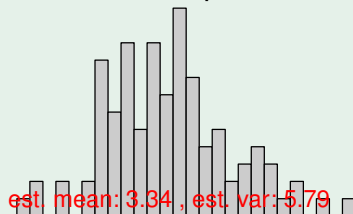
10 samples from  $N(\mu=3.44, \sigma=4.89)$



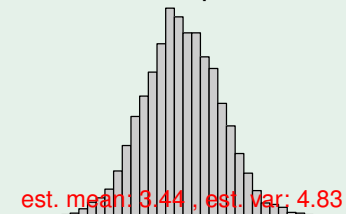
30 samples



100 samples



3000 samples





break-down joint posterior into (simpler) conditional distributions

- difficult: sampling  $P(\beta_0, \beta_1, \beta_2, \sigma^2 | Y)$
- easy: sampling  $P(\beta_0, \beta_1, \beta_2, | \sigma^2, Y)$  then  $P(\sigma^2 | \beta_0, \beta_1, \beta_2, Y)$  then repeat

approximates the joint posterior

## algorithm

- initialize:  $\beta_0^{(0)}, \beta_1^{(0)}, \beta_2^{(0)}, \sigma^{2(0)}$

$$\begin{aligned}\{\beta_0^{(1)}, \beta_1^{(1)}, \beta_2^{(1)}\} &\sim P(\beta | \sigma^{2(0)}, Y) \\ \sigma^{2(1)} &\sim P(\sigma^2 | \beta_0^{(1)}, \beta_1^{(1)}, \beta_2^{(1)}, Y) \\ \{\beta_0^{(2)}, \beta_1^{(2)}, \beta_2^{(2)}\} &\sim P(\beta | \sigma^{2(1)}, Y) \\ \sigma^{2(2)} &\sim P(\sigma^2 | \beta_0^{(2)}, \beta_1^{(2)}, \beta_2^{(2)}, Y)\end{aligned}\tag{5}$$

- repeat 1000's or 1000000 's times

Previously, Bayesian analysis demanded custom-coding MCMC algorithms

WinBUGS & OpenBUGS & JAGS

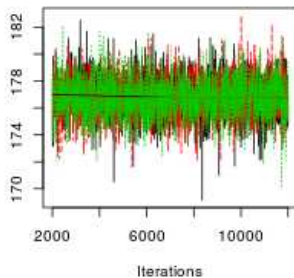
automatically use appropriate sampling techniques; so we don't have to worry

BUT you must: Monitor the MCMC!

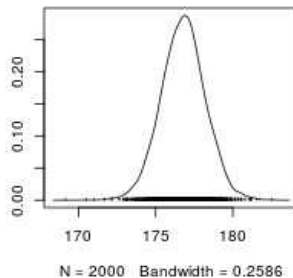
- give reasonable **initial values**
- ensure **convergence**: no trend; independent chains give same answer
- ensure adequate **mixing**: independent samples

# MCMC: Good mixing

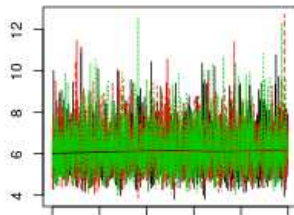
Trace of  $\mu$



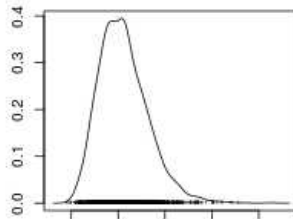
Density of  $\mu$



Trace of  $\sigma$

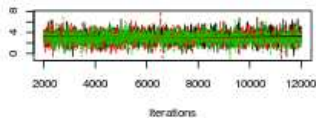


Density of  $\sigma$

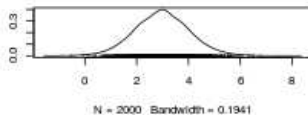


# MCMC: Poor convergence

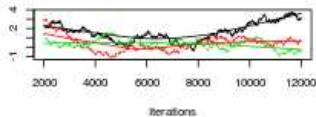
Trace of beta1



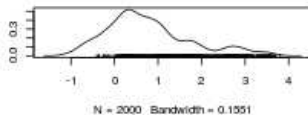
Density of beta1



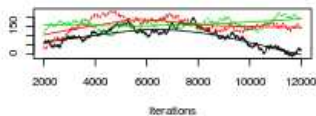
Trace of beta2



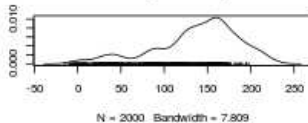
Density of beta2



Trace of intercept



Density of intercept



Trace of sigma



Density of sigma



## MCMC parameters in JAGS

- `n.chains`: num. of MCMC chains; more is better
- `n.adapt`: discard first samples; let algorithm 'adapt'
- `n.burn`: discard extra samples; allow algorithm to reach stationary distribution
- `n.iter`: total number of sample; more is better
- `thin`: take every  $k^{th}$  iteration for a sample; decorrelates one sample from the next; higher is better
- `total samples`: number of samples to approximate your Posterior; target at least 2000 to 5000

# MCMC: what to do with bad mixing

- run longer chains
- ensure long enough adaption phase
- misspecified priors
- bad initial values?

Time to open up R and JAGS

## 'JAGS: Just Another Gibbs Sampler'

Uses BUGS-like syntax (similar to OpenBUGS, WinBUGS)

- `rjags` Package: R friendly JAGS interface
- easy easy **easy** Bayesian *estimation*
- not so easy for *model selection*

Don't worry about 'samplers': JAGS does the hard work

- specify **likelihood** (how the data arose) and the **priors**

example model: counts of survival of 30 animals

```
s <-  
c(1,1,1,1,0,1,0,1,1,1,1,1,1,1,1,0,1,0,1,1,1,1,1,1,1,1,1,1,1)
```

Same data: three models

- Exercise 1: Bernoulli model: estimate mean survival with `dbeta` priors
- Exercise 2: Bernoulli model with `logit-normal` priors
- Exercise 3: logistic regression model



- open up R and rjags
- go to the file `BayesCMR_workshop/PART3_introJAGS/R_jags_intro.R`

In every analysis, we will precede as follows:

- 1 Data pre-processing in R
- 2 Write the JAGS syntax -> save to a JAGS file
- 3 Assemble the JAGS data
- 4 Initialize random variables (aka stochastic nodes)
- 5 Compile the jags model
- 6 Burn-in phase
- 7 Sample from the posteriors
- 8 Do what you want with posteriors!

## 1) Data pre-processing in R

duh

## 2) Write the JAGS syntax -> save to a JAGS file

all syntax looks like this basic structure...

---

```
model{  
  # SPECIFY PRIORS  
  phi ~ dbeta(pr.phi[1],pr.phi[2])  
  # SPECIFY LIKELIHOOD | parameter  
  for(i in 1:n){  
    y[i] ~ dbern(phi)  
  }  
}
```

---

Always need section for:

- priors
- likelihood | parameter

## 3) Assemble the JAGS data

- Jags wants a **NAMED LIST** of data.
- The names in the JAGS syntax must match the names in the list

---

```
jags.data<-list(  
  y = c(1,0,0,0,0,1,0,0,2), # response variable  
  n = length(n), # length of data (sample size) for loop  
  pr.phi = c(4,1) # prior parameters  
)
```

---

## 4) Initialize random variables (aka stochastic nodes)

- make a function that will *seed* random values for all RANDOM VARIABLES.
- the function must return a **NAMED LIST** with names equal to the names of random variables in the jags syntax

---

```
jags.inits.f <- function(){ # no arguments
  ret<-list(
    phi=runif(1,0,1)
  )
  return(ret)}
```

---

What is a random variable?

Every with a tilde  $\sim$  after it (except the data  $y$ )

## 5) Compile the jags model

```
m <- jags.model(file = "my.first.model", data=jags.data,  
  inits = jags.inits.f, n.chains=3, n.adapt=1000)
```

## 6) Burn-in phase

```
update(m, 1000)
```

## 7) Sample from the posteriors

```
post <- coda.samples(model=m, variable.names=c("phi"),  
  n.iter=30000, thin=300)
```

## 8) Do what you want with posteriors!

```
summary(post)
```

There are three exercises

- Exercise 1: A simple **coin-flip survival** model with Beta priors
- Exercise 2: A simple **coin-flip survival** model with logit-Normal priors
- Exercise 3: A **logistic regression** model