exercise

December 3, 2020

1 EXERCISE 1 - ML - Grundverfahren

1.1 Submission Instructions

Please follow the steps before submission: 1. Make sure that every cell is executed and the output is printed. 2. Create a PdF of the Jupyter notebook via $File \rightarrow ... \rightarrow PDF$ via LaTeX (.pdf) or $File \rightarrow Print$ $Preview \rightarrow Use$ your favorit PDF printing program 3. Zip your created PdF file and your original notebook, i.e. the .ipynb file, as well as your separate pen and paper solution if existent together. 4. Rename your zip file with the following naming convention: group_y_uxxxx_uxxxx_uxxxx where y is your group number, uxxxx is the kit user from each group member / 5. Upload the zip file to Ilias. Please make sure that every group member uploads the group submission.

1.2 1.) Linear Regression

1.2.1 1.1) Matrix Vector Calculus (1 Point)

Rewrite the following expression as a matrix-vector product

$$g = \alpha \sum_{i} q_i \sum_{j} x_{ij} \sum_{k} y_{jk} z_k$$

 $\#\#\#\#\Rightarrow$ Rewriting the expression as a matrix-vector product yields:

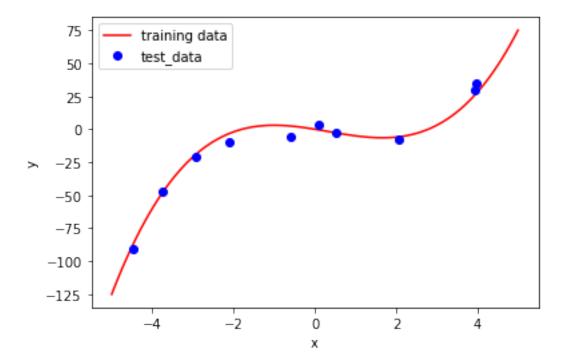
$$g = \alpha \begin{pmatrix} q_0 \\ q_1 \\ \dots \\ q_{m-1} \end{pmatrix}^{\mathsf{T}} \begin{bmatrix} x_{0,0} & x_{0,1} & \dots & x_{0,n-1} \\ x_{1,0} & x_{1,1} & \dots & x_{1,n-1} \\ \dots & \dots & \dots & \dots \\ x_{m-1,0} & x_{m-1,1} & \dots & x_{m-1,n-1} \end{bmatrix} \begin{bmatrix} x_{0,0} & x_{0,1} & \dots & x_{0,o-1} \\ x_{1,0} & x_{1,1} & \dots & x_{1,o-1} \\ \dots & \dots & \dots & \dots \\ x_{n-1,0} & x_{n-1,1} & \dots & x_{n-1,o-1} \end{bmatrix} \begin{pmatrix} z_0 \\ z_1 \\ \dots \\ q_{o-1} \end{pmatrix}$$

1.2.2 Ridge Regression

Let's first get the data

```
[5]: import numpy as np
     import matplotlib.pyplot as plt
     from typing import Tuple
     # Load data
     training_data = np.load('training_data.npy')
     test_data = np.load('test_data.npy')
     test_data_x = test_data[:, 0]
     test_data_y = test_data[:, 1]
     training_data_x = training_data[:, 0]
     training_data_y = training_data[:, 1]
     # Visualize data
     plt.plot(test_data_x, test_data_y, 'r')
     plt.plot(training_data_x, training_data_y, 'ob')
     plt.xlabel('x')
     plt.ylabel('y')
     plt.legend(["training data", "test_data"])
```

[5]: <matplotlib.legend.Legend at 0x11b23f970>



As in the lecture notebook, we will use polynomial-features here again. The following functions

will be used for calculating polynomial features, evaluating the model and calculating the Mean Squarred Error for assigning a performance to each model. Note that we have a different function called 'get_mean_std_trainset_features' here. This function will return the mean and the standard deviation of the training feature matrix. We will use the mean and the standard deviation to normalize our features according to:

$$ilde{oldsymbol{\Phi}} = rac{oldsymbol{\Phi}(oldsymbol{x}) - oldsymbol{\mu}_\Phi}{oldsymbol{\sigma}_\Phi},$$

where $\tilde{\Phi}$ are the (approximately) normalized features to any input x (not necessarily the training data), μ_{Φ} is the mean of the features applied to the training data and σ_{Φ} is the standard deviation of the features applied to the training data for each dimension. Normalization is a standard technique used in Regression to avoid numerical problems and to obtain better fits for the weight vectors w. Especially when the features transform the inputs to a very high value range, normalization is very useful. In this homework we will use features of degree 10. Since the input range of the data is roughly from -4 to 4 this will lead to very high values for higher order degrees. By normalizing each dimension of the feature matrix, we will map each dimension of the feature matrix applied to the training data to a zero mean unit variance distribution.

```
[6]: # Function for calculating the mean and standard deviation of the training
      \rightarrow feature set
     def get mean std trainset features(data: np.ndarray, degree:float) -> Tuple[np.
      →ndarray, np.ndarray]:
         :param data: training data points, shape: [n_samples] (we have 1-dim data)
         :param degree: degree of your polynomial, shape: scalar
         :return mean_feat: mean vector of the features applied to the training\Box
      \rightarrow data, shape: [1 x (degrees+1)]
         :return std feat: standard deviation vector(standard deviation for each_
      ⇒ dimension in feature matrix),
                            shape: [1 \ x \ (degrees+1)]
         11 11 11
         unnormalized_features = get_polynomial_features(data, degree, None, None)
         mean_feat = np.mean(unnormalized_features, axis=0, keepdims=True)
         mean_feat[:, 0] = 0.0 # we don't want to normalize the bias
         std_feat = np.std(unnormalized_features, axis=0, keepdims=True)
         std_feat[:, 0] = 1.0 # we don't want to normalize the bias
         return mean_feat, std_feat
     # Function to create Feature Matrix
     def get_polynomial_features(data: np.ndarray, degree:float, mean_train_features:
      → np.ndarray, std_train_features: np.ndarray) ->np.ndarray:
         11 11 11
         :param data: data points you want to evaluate the polynomials, shape: ⊔
      \rightarrow [n_samples] (we have 1-dim data)
         :param degree: degree of your polynomial, shape: scalar
         :param mean_train_features: mean of the feature matrix for the training_
      \rightarrowset, shape:[1 x (degrees+1)]
```

```
:param std train features: standard deviation of the feature matrix for the
\hookrightarrow training set, shape: [1 x (degrees+1)]
    :return features: feature matrix, shape: [n_data x (degree+1)]
    Extends the feature matrix according to the matrix form discussed in the
\hookrightarrow lectures.
    features = np.ones(data.shape)
    for i in range(degree):
        features = np.column_stack((features, (data)**(i+1)))
    if mean_train_features is not None: # if mean_test_features is None, do not_
 \rightarrownormalize
        # note: features: (n_samplesxn_dims), mean_train_features: (1xn_dims), u
 →std_train_features: (1xn_dims)
                due to these dimensionalities we can do element-wise operations.
 → By this we normalize each
                dimension independantly
        norm feat = (features-mean train features)/(std train features)
        return norm feat
    else:
        return features
# Evaluate the models
def eval(Phi:np.ndarray, w:np.ndarray)->np.ndarray:
    : param Phi: Feature matrix, shape: [n_data x (degree+1)]
    : param w: weight vector, shape: [degree + 1]
    : return : predictions, shape [n_data] (we have 1-dim data)
    Evaluates your model
    11 11 11
    return np.dot(Phi, w)
def mse(y_target:np.ndarray, y_pred:np.ndarray)->np.ndarray:
    :param y_target: the target outputs, which we want to have, shape: [n_data]__
\hookrightarrow (here 1-dim data)
    :param y_pred: the predicted outputs, shape: [n_data] (we have 1-dim data)
    :return : The Mean Squared Error, shape: scalar
    11 11 11
    dif = y_target - y_pred
    return np.sum(dif ** 2, axis=0) / y_pred.shape[0]
```

1.2.3 1.2) Ridge Regression Weights (4 Points)

Derive the weight updates for ridge regressin in matrix form. Hint: You will need derivatives for vectors/matrices. Start from the matrix objective for ridge regression as stated here

$$L = (\boldsymbol{y} - \boldsymbol{\Phi} \boldsymbol{w})^T (\boldsymbol{y} - \boldsymbol{\Phi} \boldsymbol{w}) + \lambda \boldsymbol{w}^T \boldsymbol{I} \boldsymbol{w}.$$

 $\#\#\#\#\Rightarrow$ Rewriting the equation above gives:

$$L = \boldsymbol{w}^T \boldsymbol{\Phi}^T \boldsymbol{\Phi} \boldsymbol{w} - \boldsymbol{y}^T \boldsymbol{\Phi} \boldsymbol{w} - \boldsymbol{w}^T \boldsymbol{\Phi}^T \boldsymbol{y} + \boldsymbol{y}^T \boldsymbol{y} + \lambda \boldsymbol{w}^T \boldsymbol{I} \boldsymbol{w}$$
$$= \boldsymbol{w}^T \boldsymbol{\Phi}^T \boldsymbol{\Phi} \boldsymbol{w} - 2 \boldsymbol{y}^T \boldsymbol{\Phi} \boldsymbol{w} + \boldsymbol{y}^T \boldsymbol{y} + \lambda \boldsymbol{w}^T \boldsymbol{I} \boldsymbol{w}$$

 $\#\#\#\#\Rightarrow$ Taking the derivative w.r.t. wyields:

$$\nabla_{w}L = \frac{\partial}{\partial w} \left\{ \boldsymbol{w}^{T} \boldsymbol{\Phi}^{T} \boldsymbol{\Phi} \boldsymbol{w} - 2 \boldsymbol{y}^{T} \boldsymbol{\Phi} \boldsymbol{w} + \boldsymbol{y}^{T} \boldsymbol{y} + \lambda \boldsymbol{w}^{T} \boldsymbol{I} \boldsymbol{w} \right\}$$
$$= \frac{\partial}{\partial w} \left\{ \boldsymbol{w}^{T} \boldsymbol{\Phi}^{T} \boldsymbol{\Phi} \boldsymbol{w} - 2 \boldsymbol{y}^{T} \boldsymbol{\Phi} \boldsymbol{w} + \lambda \boldsymbol{w}^{T} \boldsymbol{I} \boldsymbol{w} \right\}$$
$$= 2 \boldsymbol{\Phi}^{T} \boldsymbol{\Phi} \boldsymbol{w} - 2 \boldsymbol{\Phi}^{T} \boldsymbol{y} + 2 \lambda \boldsymbol{I} \boldsymbol{w} \stackrel{!}{=} 0$$

 $\#\#\#\#\Rightarrow$ Setting the gradient to 0 yields:

$$2\mathbf{\Phi}^{T}\mathbf{\Phi}\boldsymbol{w} - 2\mathbf{\Phi}^{T}\boldsymbol{y} + 2\lambda\boldsymbol{I}\boldsymbol{w} = 0$$

$$\mathbf{\Phi}^{T}\mathbf{\Phi}\boldsymbol{w} - \mathbf{\Phi}^{T}\boldsymbol{y} + \lambda\boldsymbol{I}\boldsymbol{w} = 0$$

$$\mathbf{\Phi}^{T}\mathbf{\Phi}\boldsymbol{w} + \lambda\boldsymbol{I}\boldsymbol{w} = \mathbf{\Phi}^{T}\boldsymbol{y}$$

$$(\mathbf{\Phi}^{T}\mathbf{\Phi} + \lambda\boldsymbol{I})\boldsymbol{w} = \mathbf{\Phi}^{T}\boldsymbol{y}$$

$$\boldsymbol{w} = (\mathbf{\Phi}^{T}\mathbf{\Phi} + \lambda\boldsymbol{I})^{-1}\mathbf{\Phi}^{T}\boldsymbol{y}$$

1.2.4 1.3) Implement Ridge Regression Weights (2 Point)

The following function will calculate the weights for ridge regression. Fill in the missing code according to the formula for calculating the weight updates for ridge regression. Recall that the formula is given by

$$\boldsymbol{w} = (\boldsymbol{\Phi}^T \boldsymbol{\Phi} + \lambda \boldsymbol{I})^{-1} \boldsymbol{\Phi}^T \boldsymbol{y},$$

where Φ is the feature matrix (the matrix storing the data points applied to the polynomial features). Hint: use np.linalg.solve for solving for the linear equation. If you got confused because of the normalization described before, don't worry, you do not need to consider it here:)

[7]:

```
def calc weights ridge(Phi:np.ndarray, y:np.ndarray, ridge_factor:float)->np.
→ndarray:
    11 11 11
    :param Phi: Feature Matrix, shape: [n_data x (degree+1)]
    :param y: Output Values, [n_data] (we have 1-dim data)
    :param ridge factor: lambda value, shape: scalar
    return : The weight vector, calculated according to the equation shown □
⇒before, shape: [degrees +1]
    11 11 11
    ## DONE
    dot = np.dot(Phi.T, Phi)
    summed = dot + ridge_factor * np.identity(len(Phi)+1)
    solved = np.linalg.solve(summed, Phi.T) # np.identity(len(Phi)+1))
    # is it possible to solve for Phi to make next step redundant?
    # dotted = np.dot(solved, Phi.T)
    return np.dot(solved, y) # dotted, y)
    # DONE
```

For demonstrating ridge regression we will pick the polynomial degree of 10. In the lecture notebook we have seen that this model is highly overfitting to the data. We will investigate the role of the ridge factor λ . For that purpose we first need to calculate the weights for different λ values. We will pick $\lambda = [1e-6, 1e-3, 1, 3, 5, 10, 20, 30, 40, 50, 1e2, 1e3, 1e5]$ to see the differences of the values. Practical note. We use here very high values for λ for demonstration purposes here. In practice we would not choose a model where we know from beginning that it is highly overfitting. When choosing an appropriate model, the value needed for λ automatically will be small (often in the range of $1e^{-6}$ or smaller).

```
# We further have to perform the predictions based on the models we have

calculated

y_training_ridge = []

y_test_ridge = []

for w in weights_ridge:
    y_training_ridge.append(eval(poly_10_train, w))
    y_test_ridge.append(eval(poly_10_test, w))
```

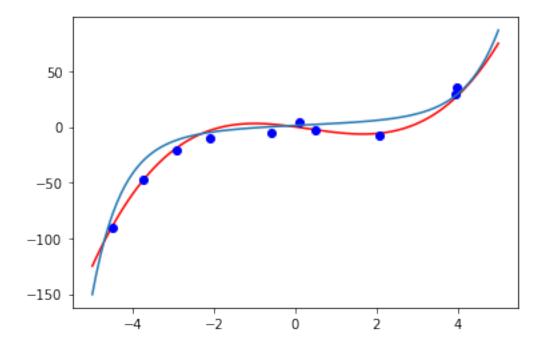
We are interested in the mean squarred error on the test and the training data. For that purpose we calculate them here and plot the errors for different λ values in log space.

```
[9]: training_error_ridge = []
     test_error_ridge = []
     for i in range(len(y_training_ridge)):
         training_error_ridge.append(mse(training_data_y, y_training_ridge[i]))
         test_error_ridge append(mse(test_data_y, y_test_ridge[i]))
     error_fig_ridge = plt.figure()
     plt.figure(error_fig_ridge.number)
     plt.title("Error Plot Ridge Regression")
     plt.xlabel("$\lambda$")
     plt.ylabel("MSE")
     x_axis = ["$1e-{6}$", "$1e-{3}$", "$1$", "$3$", \[ \]
     → "$5$", "$10$", "$20$", "$30$", "$40$", "$50$",
               "$1e2$", "$1e3$", "$1e5$"]
     plt.yscale('log')
     plt.plot(x_axis, training_error_ridge, 'b')
     plt.plot(x_axis, test_error_ridge, 'r')
     # let's find the index with the minimum training error
     min_error_idx = np.argmin(test_error_ridge)
     plt.plot(x_axis[min_error_idx], test_error_ridge[min_error_idx], 'xg')
     plt.legend(['Training Error', 'Test Error', 'Min Test Error'])
```

[9]: <matplotlib.legend.Legend at 0x11b318850>



[10]: [<matplotlib.lines.Line2D at 0x11b5aaf10>]



1.2.5 1.4) Error Plot (1 Point)

In the lecture we have seen the error plot for polynomial degrees (slide 44). Draw a connection to the conclusions regarding over- and underfitting learned in the lecture to the different values for λ here. What is characteristic for overfitting and what is characteristic for underfitting with respect to the λ values? Hint: Do not forget that we are in log space. Small changes in the y-axis mean high differences in the error values.

 $\#\#\#\#\Rightarrow$ Evaluating the graph w.r.t. the Lambda's yields:

Small Lambdas yield to overfitting, as Training Data Error is very small and Test Data Error is huge (\rightarrow Model too complex).

Big Lambda Values yield to underfitting as both errors - for Training and Test Data - rise (\rightarrow Model too simple).

2 Probability Basics and Linear Classification

2.1 First Example (Two Moons)

Let us start by loading a very simple toy dataset, the "two moons".

```
[11]: %matplotlib inline
import numpy as np
import matplotlib.pyplot as plt
```

```
from typing import Tuple, Callable

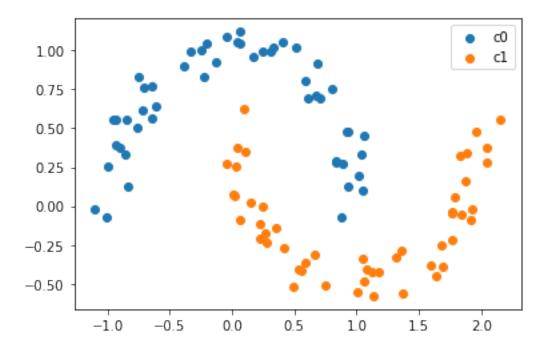
data = dict(np.load("two_moons.npz", allow_pickle=True))
samples = data["samples"]

labels = data["labels"]

c0_samples = samples[labels == 0]
c1_samples = samples[labels == 1]

plt.figure("Data")
plt.scatter(x=c0_samples[:, 0], y=c0_samples[:, 1])
plt.scatter(x=c1_samples[:, 0], y=c1_samples[:, 1])
plt.legend(["c0", "c1"])
```

[11]: <matplotlib.legend.Legend at 0x11b4e9a30>



Let us also define some plotting utility

```
b = plt_std * np.sqrt(smallest_eigval)

ellipse_x_r = a * np.cos(np.linspace(0, 2 * np.pi, num=200))
    ellipse_y_r = b * np.sin(np.linspace(0, 2 * np.pi, num=200))

R = np.array([[np.cos(phi), np.sin(phi)], [-np.sin(phi), np.cos(phi)]])
    r_ellipse = np.array([ellipse_x_r, ellipse_y_r]).T @ R
    plt.plot(mu[0] + r_ellipse[:, 0], mu[1] + r_ellipse[:, 1], *args, **kwargs)

# plot grid for contour plots
plt_range = np.arange(-1.5, 2.5, 0.01)
plt_grid = np.stack(np.meshgrid(plt_range, plt_range), axis=-1)
flat_plt_grid = np.reshape(plt_grid, [-1, 2])
plt_grid_shape = plt_grid.shape[:2]
```

2.2 2): Classification using Generative Models (Naive Bayes Classifier)

We first try a generative approach, the Naive Bayes Classifier. We model the class conditional distributions $p(\boldsymbol{x}|c)$ as Gaussians, the class prior p(c) as Bernoulli and apply bayes rule to compute the class posterior $p(c|\boldsymbol{x})$.

2.2.1 2.1): Implementing Generative Classifier (3 Points):

Fill in the missing code snippets below such that code runs and a prediction is made by the classifier. The final accuracy should be 87%.

Recall that the density of the Multivariate Normal Distribution is given by

$$p(\boldsymbol{x}) = \mathcal{N}\left(\boldsymbol{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}\right) = \frac{1}{\sqrt{\det\left(2\pi\boldsymbol{\Sigma}\right)}} \exp\left(-\frac{(\boldsymbol{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\boldsymbol{x} - \boldsymbol{\mu})}{2}\right)$$

```
[13]: def mvn_pdf(x: np.ndarray, mu: np.ndarray, sigma: np.ndarray) -> np.ndarray:
    """
    Density of the Multivariate Normal Distribution
    :param x: samples, shape: [N x dimension]
    :param mu: mean, shape: [dimension]
    :param sigma: covariance, shape: [dimension x dimension]
    :return p(x) with p(x) = N(mu, sigma), shape: [N]
    """
    norm_term = 1 / np.sqrt(np.linalg.det(2 * np.pi * sigma))
    diff = x - np.atleast_2d(mu)
    exp_term = np.sum(np.linalg.solve(sigma, diff.T).T * diff, axis=-1)
    return norm_term * np.exp(-0.5 * exp_term)
```

Practical Aspect: In praxis you would never implement it like that, but stay in the log-domain. Also for numerically stable implementations of the multivariate normal density the symmetry and

positive definitness of the covariance should be exploited by working with it's Cholesky decomposition.

The maximum likelihood estimator for a Multivariate Normal Distribution is given by

$$oldsymbol{\mu} = rac{1}{N} \sum_{i}^{N} oldsymbol{x}_{i} \qquad oldsymbol{\Sigma} = rac{1}{N} \sum_{i}^{N} (oldsymbol{x}_{i} - oldsymbol{\mu}) (oldsymbol{x}_{i} - oldsymbol{\mu})^{T}.$$

2.2): Derivation of Maximum Likelihood Estimator (4 Points): Derive the maximum likelihood estimator for Multivariate Normal distributions, given above. This derivations involves some matrix calculus. Matrix calculus is a bit like programming, you google the stuff you need and then plug it together in the right order. Good resources for such rules are the "matrix cookbook" (https://www.math.uwaterloo.ca/~hwolkowi/matrixcookbook.pdf) and the Wikipdia article about matrix calculus (https://en.wikipedia.org/wiki/Matrix_calculus). State all rules you use explicitly (except the ones given in the hints below).

Remark There are different conventions of how to define a gradient (as column-vector or row-vector). This results in different ways to write the Jacobian and thus different, usually transposed, matrix calculus rules: - In the lecture we define the gradient as column-vector - In the Wikipedia article this convention is referred to as "Denominator Layout". It also contains a nice explanaiton of the different conventions for the gourmets among you;) - The Matrix Cookbook uses the same convention (gradient as column vector) - Please also use it here

Hint Here are two of those rules that might come in handy

(1)
$$\frac{\partial \log \det(\boldsymbol{X})}{\partial \boldsymbol{X}} = \boldsymbol{X}^{-1}$$

(2)
$$\frac{\partial \mathbf{x}^T \mathbf{A} \mathbf{x}}{\partial \mathbf{x}} = 2\mathbf{A} \mathbf{x}$$
 for symmetric matrices \mathbf{A} (Hint: covariance matrices are always symmetric)

There is one missing to solve the exercise. You need to find it yourself. (Hint: Look in the matrix cookbook, chapter 2.2)

 $\#\#\#\#\Rightarrow$ The following matrix calculus rule is used in addition to the two given above:

(3)
$$\frac{\partial \boldsymbol{a}^T \boldsymbol{X}^{-1} \boldsymbol{b}}{\partial \boldsymbol{X}} = -\boldsymbol{X}^{-T} \boldsymbol{a} \boldsymbol{b}^T \boldsymbol{X}^{-T}$$

$$(3) \ \frac{\partial a^T X b}{X} = a b^T$$

 $\#\#\#\#\Rightarrow$ Deriving mean μ :

$$p(\boldsymbol{x}) = \mathcal{N} (\boldsymbol{x} | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$

$$= \frac{1}{\sqrt{\det(2\pi\boldsymbol{\Sigma})}} \exp\left(-\frac{(\boldsymbol{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\boldsymbol{x} - \boldsymbol{\mu})}{2}\right)$$

####⇒log-likelihood:

$$\sum_{i}^{N} \log \left(\frac{1}{\sqrt{\det(2\pi\Sigma)}} \exp\left(-\frac{(\boldsymbol{x}_{i} - \boldsymbol{\mu})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x}_{i} - \boldsymbol{\mu})}{2} \right) \right) \rightarrow \log(a \cdot b) = \log(a) + \log(b)$$

$$= \sum_{i}^{N} \log \left(\frac{1}{\sqrt{\det(2\pi\Sigma)}} \right) + \log \left(\exp\left(-\frac{(\boldsymbol{x}_{i} - \boldsymbol{\mu})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x}_{i} - \boldsymbol{\mu})}{2} \right) \right)$$

$$= -N \log \left(\frac{1}{\sqrt{\det(2\pi\Sigma)}} \right) - \sum_{i}^{N} \left(\frac{(\boldsymbol{x}_{i} - \boldsymbol{\mu})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x}_{i} - \boldsymbol{\mu})}{2} \right)$$

\Rightarrow MLE Solution for μ : $\frac{\partial \log likelihood(\mu; D)}{\partial \mu} \stackrel{!}{=} 0$

$$0 \stackrel{!}{=} \frac{\partial}{\partial \mu} \left\{ -N \log \left(\frac{1}{\sqrt{\det(2\pi \Sigma)}} \right) - \sum_{i}^{N} \left(\frac{(x_{i} - \mu)^{T} \Sigma^{-1}(x_{i} - \mu)}{2} \right) \right\}$$

$$= \frac{\partial}{\partial \mu} \left\{ -\sum_{i}^{N} \frac{(x_{i} - \mu)^{T} \Sigma^{-1}(x_{i} - \mu)}{2} \right\}$$

$$= -\sum_{i}^{N} \frac{2\Sigma^{-1}(x_{i} - \mu)}{2}$$

$$= -N\Sigma^{-1} \sum_{i}^{N} (x_{i} - \mu)$$

$$= N^{2} \Sigma^{-1} \mu - N\Sigma^{-1} \sum_{i}^{N} x_{i}$$

$$N^{2} \Sigma^{-1} \mu = N\Sigma^{-1} \sum_{i}^{N} x_{i}$$

$$\rightarrow \Sigma \cdot \Sigma^{-1} = I$$

$$N\mu = \sum_{i}^{N} x_{i}$$

$$\mu = \frac{1}{N} \sum_{i}^{N} x_{i}$$

 $\#\#\#\#\Rightarrow$ Deriving covariance Σ :

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

$$= \frac{1}{\sqrt{\det(2\pi\boldsymbol{\Sigma})}} \exp\left(-\frac{(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})}{2}\right)$$

####⇒log-likelihood:

$$\sum_{i}^{N} \log \left(\frac{1}{\sqrt{\det(2\pi\Sigma)}} \exp\left(-\frac{(\mathbf{x} - \boldsymbol{\mu})^{T} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})}{2} \right) \right) \rightarrow \log(a \cdot b) = \log(a) + \log(b)$$

$$= \sum_{i}^{N} \left(\log \left(\frac{1}{\sqrt{\det(2\pi\Sigma)}} \right) + \log \left(\exp\left(-\frac{(\mathbf{x} - \boldsymbol{\mu})^{T} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})}{2} \right) \right) \right) \rightarrow \log \left(\frac{1}{x} \right) = -\log(x)$$

$$\rightarrow \log(\sqrt{x}) = \frac{1}{2} \cdot \log(x)$$

$$= -\frac{N}{2} \log \left(\det(2\pi\Sigma) \right) - \sum_{i}^{N} \frac{(\mathbf{x}_{i} - \boldsymbol{\mu})^{T} \mathbf{\Sigma}^{-1} (\mathbf{x}_{i} - \boldsymbol{\mu})}{2} \rightarrow \operatorname{Da} \mathbf{x} \text{ d-Dimensional } \rightarrow \det(2\pi\Sigma) =$$

$$= -\frac{Nd}{2} \log(2\pi) - \frac{N}{2} \log \left(\det(\Sigma) \right) - \sum_{i}^{N} \frac{(\mathbf{x}_{i} - \boldsymbol{\mu})^{T} \mathbf{\Sigma}^{-1} (\mathbf{x}_{i} - \boldsymbol{\mu})}{2}$$

\Rightarrow MLE Solution for Σ : $\frac{\partial \log likelihood(\Sigma; D)}{\partial \Sigma} \stackrel{!}{=} 0$

$$0 \stackrel{!}{=} \frac{\partial}{\partial \Sigma} \left\{ -\frac{Nd}{2} \log(2\pi) - \frac{N}{2} \log(\det(\Sigma)) - \sum_{i}^{N} \frac{(\boldsymbol{x}_{i} - \boldsymbol{\mu})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x}_{i} - \boldsymbol{\mu})}{2} \right\} \rightarrow -\det(\boldsymbol{X}) = \det(\boldsymbol{X}^{-1}) \text{ For sym}$$

$$= \frac{\partial}{\partial \Sigma} \left\{ -\frac{Nd}{2} \log(2\pi) + \frac{N}{2} \log\left(\det(\Sigma^{-1})\right) - \sum_{i}^{N} \frac{(\boldsymbol{x}_{i} - \boldsymbol{\mu})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x}_{i} - \boldsymbol{\mu})}{2} \right\} \rightarrow \text{Equation (1) and (3)}$$

$$= \frac{N}{2} \boldsymbol{\Sigma} + \frac{1}{2} \sum_{i}^{N} \boldsymbol{\Sigma}^{-T} (\boldsymbol{x}_{i} - \boldsymbol{\mu}) (\boldsymbol{x}_{i} - \boldsymbol{\mu})^{T} \boldsymbol{\Sigma}^{-T}$$

$$\boldsymbol{\Sigma} = -\frac{1}{N} \sum_{i}^{N} \boldsymbol{\Sigma}^{-T} (\boldsymbol{x}_{i} - \boldsymbol{\mu}) (\boldsymbol{x}_{i} - \boldsymbol{\mu})^{T} \boldsymbol{\Sigma}^{-T}$$

$$????$$

Transform the log-likelihood function

$$\begin{split} \ell(x|\mu,\Sigma) &= \sum_{i}^{N} \log(\frac{1}{|2\pi\Sigma|} * e^{-\frac{(x_{i}-\mu)^{T}\Sigma^{-1}(x_{i}-\mu)}{2}}) \\ &= \sum_{i}^{N} \log(\frac{1}{|2\pi\Sigma|}) + \log(e^{-\frac{(x_{i}-\mu)^{T}\Sigma^{-1}(x_{i}-\mu)}{2}}) \\ &= \dots & \text{As always in math, the next steps} \\ &= \frac{-N}{2} * \frac{\log(|2\pi\Sigma|)}{2} - \frac{1}{2} * \sum_{i}^{N} (x_{i}-\mu)^{T}\Sigma^{-1}(x_{i}-\mu) & \text{Since } 2\pi \text{ is a scaler, we can } \text{do}|2\pi\Sigma| \\ &= \frac{-N}{2} * (\log(|2\pi\|) + \frac{\log(|\Sigma|)}{2}) - \frac{1}{2} * \sum_{i}^{N} (x_{i}-\mu)^{T}\Sigma^{-1}(x_{i}-\mu) & \log(x) = -\log x^{-1}, \text{i.e. multiplying} \\ &= \frac{-N}{2} * \log(|2\pi\|) + \frac{N}{2} * \log(|\Sigma^{-1}|) - \frac{1}{2} * \sum_{i}^{N} (x_{i}-\mu)^{T}\Sigma^{-1}(x_{i}-\mu) & \text{We can now derive this after } \Sigma^{-1} \end{split}$$

Calculate derivative

$$\begin{split} \frac{\partial l}{\partial \Sigma^{-1}} &= \frac{\partial}{\partial \Sigma^{-1}} [\frac{N}{2} * \log(|2\pi|) + \frac{-N}{2} * \log(|\Sigma^{-1}|) - \frac{1}{2} * \sum_{i}^{N} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)] \\ &= \frac{\partial}{\partial \Sigma^{-1}} [\frac{N}{2} \log(|2\pi|)] + \frac{\partial}{\partial \Sigma^{-1}} [\frac{-N}{2} \log(|\Sigma^{-1}|)] + \frac{\partial}{\partial \Sigma^{-1}} [-\frac{1}{2} \sum_{i}^{N} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)] \\ &= 0 + \frac{\partial}{\partial \Sigma^{-1}} [\frac{-N}{2} \log(|\Sigma^{-1}|)] + \frac{\partial}{\partial \Sigma^{-1}} [-\frac{1}{2} * \sum_{i}^{N} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)] \\ &= 0 + \frac{-N}{2} * \Sigma + \frac{\partial}{\partial \Sigma^{-1}} [-\frac{1}{2} * \sum_{i}^{N} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)] \\ &= 0 + \frac{-N}{2} * \Sigma + -\frac{1}{2} * \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T = 0 \end{split}$$

$$\text{Now, let's set this}$$

$$0 = \frac{N}{2} * \Sigma + -\frac{1}{2} * \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \Sigma - \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \Sigma - \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \Sigma - \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \sum_{i}^{N} \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \sum_{i}^{N} \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \sum_{i}^{N} \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \sum_{i}^{N} \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \sum_{i}^{N} \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \\ &= N * \sum_{i}^{N} \sum_{i}^{N} (x_i - \mu) (x_i - \mu)^T \end{aligned}$$

2.2.2 Implementation

```
[14]: def mvn_mle(x: np.ndarray) → Tuple[np.ndarray, np.ndarray]:

"""

Maximum Likelihood Estimation of parameters for Multivariate Normal

⇒Distribution

:param x: samples shape: [N x dimension]

:return mean (shape: [dimension]) und covariance (shape: [dimension x

⇒dimension]) that maximize likelihood of data.

"""

## DONE

return (x.mean(axis=0), np.cov(x[:,0],x[:,1]))

## DONE
```

We can now use this maximum likelihood estimator to fit generative models to the samples of both classes. Using those models and some basic rules of probability we can obtain the class posterior distribution p(c|x)

```
[15]: # Fit Gaussian Distributions using the maximum likelihood estimator to samples
       → from both classes
      mu_c0, sigma_c0 = mvn_mle(c0_samples)
      mu_c1, sigma_c1 = mvn_mle(c1_samples)
      # Prior obtained by "counting" samples in each class
      p_c0 = c0_samples.shape[0] / samples.shape[0] # 0.5
      p_c1 = c1_samples.shape[0] / samples.shape[0] # DONE
      def compute_posterior(
               samples: np.ndarray,
              p_c0: float, mu_c0: np.ndarray, sigma_c0: np.ndarray,
              p_c1: float, mu_c1: np.ndarray, sigma_c1: np.ndarray) \
               -> Tuple[np.ndarray, np.ndarray]:
          computes the posteroir distribution p(c|x) given samples x, the prior p(c)_{\sqcup}
       \hookrightarrow and the
          class conditional likelihood p(x/c)
          :param samples: samples x to classify, shape: [N x dimension]
          :param p_c0: prior probability of class 0, p(c=0)
          :param mu_c0: mean of class conditional likelihood of class 0, p(x/c=0)_{\sqcup}
       \hookrightarrow shape: [dimension]
           :param sigma_c0: covariance of class conditional likelihood of class 0, _
       \rightarrow p(x/c=0) shape: [dimension x dimension]
           :param p_c1: prior probability of class 1 p(c=1)
           :param mu_c1: mean of class conditional likelihood of class 1 p(x/c=1)_{\sqcup}
       ⇔shape: [dimension]
           :param sigma_c1: covariance of class conditional likelihood of class 1, ⊔
       \rightarrow p(x/c=1) shape: [dimension x dimension]
```

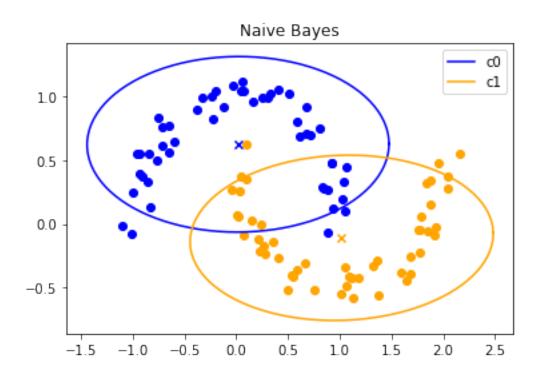
```
:return two arrays, p(c=0/x) and p(c=1/x), both shape [N]
    ## DONE
    p_x = mvn_pdf(samples, *mvn_mle(samples))
    # check: p(x/c) = mvn with c0 mu and sigma
    p_x_c0 = mvn_pdf(samples, mu_c0, sigma_c0)
    p_x_c1 = mvn_pdf(samples, mu_c1, sigma_c1)
    p_c0_given_x = (p_x_c0 * p_c0)/p_x
    p_c1_given_x = (p_x_c1 * p_c1)/p_x
    return p_c0_given_x, p_c1_given_x
    ## DONE
p_c0_given_x, p_c1_given_x = compute_posterior(samples, p_c0, mu_c0, sigma_c0, __
\rightarrowp_c1, mu_c1, sigma_c1)
# Prediction
predicted_labels = np.zeros(labels.shape)
# break at 0.5 arbitrary
predicted_labels[p_c0_given_x >= 0.5] = 0.0 # is not strictly necessary since_
→whole array already zero.
predicted_labels[p_c1_given_x > 0.5] = 1.0
# Evaluate
acc = (np.count_nonzero(predicted_labels == labels)) / labels.shape[0]
print("Accuracy:", acc)
```

Accuracy: 0.89

Lets look at the class likelihoods

```
plt.title("Naive Bayes")
plt.scatter(x=samples[labels == 0, 0], y=samples[labels == 0, 1], c="blue")
draw_2d_gaussian(mu_c0, sigma_c0, c="blue")
plt.scatter(x=samples[labels == 1, 0], y=samples[labels == 1, 1], c="orange")
draw_2d_gaussian(mu_c1, sigma_c1, c="orange")
plt.legend(["c0", "c1"])
```

[16]: <matplotlib.legend.Legend at 0x11b3d06d0>



And the final posterior distribution for the case $p(c = 1|\mathbf{x})$

```
pred_grid = np.reshape(compute_posterior(flat_plt_grid, p_c0, mu_c0, sigma_c0, p_c1, mu_c1, sigma_c1)[1], plt_grid_shape)

plt_contourf(plt_grid[..., 0], plt_grid[..., 1], pred_grid, levels=10)

plt.colorbar()

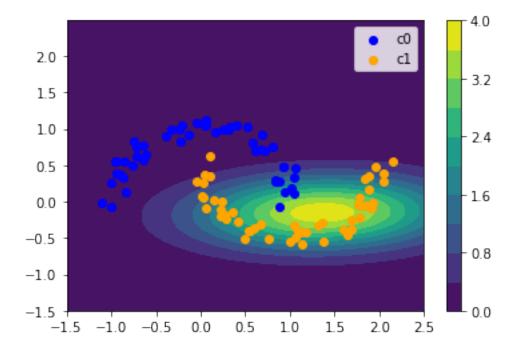
s0 = plt.scatter(c0_samples[..., 0], c0_samples[..., 1], color="blue")

s1 = plt.scatter(c1_samples[..., 0], c1_samples[..., 1], color="orange")

plt.legend([s0, s1], ["c0", "c1"])

plt.xlim(-1.5, 2.5)
```

[17]: (-1.5, 2.5)



We achieve a train accuracy of 87%. For such a simple task that is clearly not great, but it nicely illustrates a problem with generative approaches: They usually depend on quite a lot of assumptions.

2.2.3 2.3) Wrong Assumptions? (1 Point):

Which untrue assumption did we make?

We assume that our given attributes are iid. (identically and independently distributed).

$$P(C_0 \cap C_1|X) = P(C_0|X) * P(C_1|X)$$

2.2.4 Discriminative Classifier using Logistic Regression

This part of the Notebook was already presented in the Lecture and is only here for reference. There are no tasks in this part.

We start by implementing a few helper functions for affine mappings, the sigmoid function and the negative bernoulli log-likelihood.

```
[18]: def affine_features(x: np.ndarray) -> np.ndarray:
    """
    implements affine feature function
    :param x: inputs, shape: [N x sample_dim]
    :return inputs with additional bias dimension, shape: [N x feature_dim]
```

```
return np.concatenate([x, np.ones((x.shape[0], 1))], axis=-1)
def quad_features(x: np.ndarray) -> np.ndarray:
    implements quadratic feature function
    :param x: inputs, shape: [N x sample_dim]
    :return squared features of x, shape: [N x feature_dim]
    sq = np.stack([x[:, 0] ** 2, x[:, 1] ** 2, x[:, 0] * x[:, 1]], axis=-1)
    return np.concatenate([sq, affine features(x)], axis=-1)
def cubic features(x: np.ndarray) -> np.ndarray:
    implements cubic feature function
    :param x: inputs, shape: [N x sample_dim]
    :return cubic features of x, shape: [N x feature_dim]
    cubic = np.stack([x[:, 0]**3, x[:, 0]**2 * x[:, 1], x[:, 0] * x[:, 1]**2,\Box
\rightarrow x[:, 1]**3], axis=-1)
    return np.concatenate([cubic, quad features(x)], axis=-1)
def sigmoid(x: np.ndarray) -> np.ndarray:
    the sigmoid function
    :param x: inputs
    :return sigma(x)
    HHHH
    return 1 / (1 + np.exp(-x))
def bernoulli_nll(predictions: np.ndarray, labels: np.ndarray, epsilon: float = 0
\rightarrow1e-12) -> np.ndarray:
    11 11 11
    :param predictions: output of the classifier, shape: [N]
    :param labels: true labels of the samples, shape: [N]
    :param epsilon: small offset to avoid numerical instabilities (i.e loq(0))
    :return negative log-likelihood of the labels given the predictions
    return - (labels * np.log(predictions + epsilon) + (1 - labels) * np.log(1<sub>U</sub>
 →- predictions + epsilon))
```

2.2.5 Optimization by Gradient Descent

First, we implement a very simple gradient descent optimizer. It iteratively applies the gradient descent rule introduced in the lecture

$$\boldsymbol{x}_{t+1} = \boldsymbol{x}_t - \eta \nabla f(\boldsymbol{x}_t).$$

We also add some simple stopping criteria which terminate the minimization if the algorithm has converged.

```
[19]: def minimize(f: Callable, df: Callable, x0: np.ndarray, lr: float, num_iters:
       →int) -> \
              Tuple[np.ndarray, float, np.ndarray, np.ndarray]:
          :param f: objective function
          :param df: gradient of objective function
          :param x0: start point, shape [dimension]
          :param lr: learning rate
          :param num_iters: maximum number of iterations
          :return argmin, min, values of x for all interations, value of f(x) for all
       \rightarrow iterations
          11 11 11
          # initialize
          x = np.zeros([num_iters + 1] + list(x0.shape))
          f x = np.zeros(num iters + 1)
          x[0] = x0
          f x[0] = f(x0)
          for i in range(num_iters):
              # update using gradient descent rule
              grad = df(x[i])
              x[i + 1] = x[i] - lr * grad
              f_x[i + 1] = f(x[i + 1])
          return x[i+1], f_x[i+1], x[:i+1], f_x[:i+1] # logging info for visualization
```

Practical Aspect: While such a simple gradient descent optimizer works for the task we are considering and is simple enough to implement, in practice you should always use more sophisticated optimizers (e.g. L-BFGS) and use existing implementations. Such efficient and well-tested implementations are provided by software packages such as NlOpt or scipy.optimize.

Next, we need to define the cost function and its derivative. Maximizing the likelihood is equivalent to minimizing the negative log-likelihood, which we are using here. The derivation of the gradient is given in the lecture.

Note that we do not sum the losses as in the lecutre but take the mean. This is just a multiplication with a constant, thus the optimal parameters do not change. Yet, working with the mean makes the loss, and more importantly, the length of the gradient independent of the number of samples.

```
[20]: def objective_bern(weights: np.ndarray, features: np.ndarray, labels: np.

→ndarray) -> float:

"""

bernoulli log-likelihood objective

:param weights: current weights to evaluate, shape: [feature_dim]

:param features: train samples, shape: [N x feature_dim]

:param labels: class labels corresponding to train samples, shape: [N]

:return average negative log-likelihood

"""
```

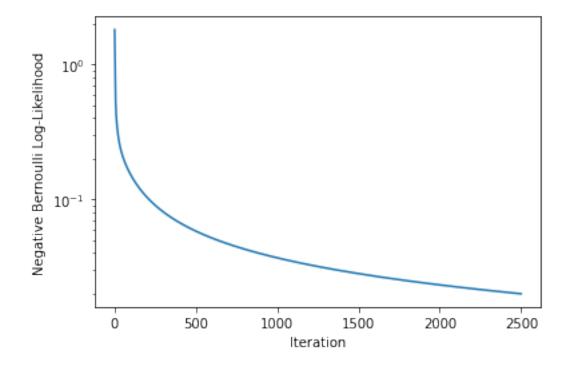
Finally, we can tie everything together and get our probabilistic classifier

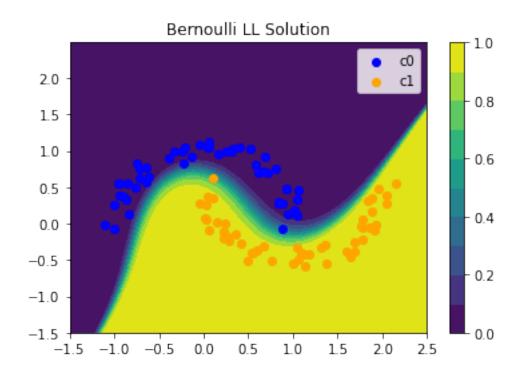
```
[21]: # Generate Features from Data
      # change this to play arround with feature functions
      #feature_fn = affine_features
      #feature_fn = quad_features
      feature_fn = cubic_features
      features = feature_fn(samples)
      # Optimize Loss
      w_bce, loss, x_history, f_x_history = \
          minimize(lambda w: objective_bern(w, features, labels),
                   lambda w: d_objective_bern(w, features, labels),
                   np.ones(features.shape[1]), 1, 2500)
      print("Final loss:", loss)
      # Plot
      plt.figure()
      plt.semilogy(f_x_history)
      plt.xlabel("Iteration")
      plt.ylabel("Negative Bernoulli Log-Likelihood")
      plt.figure()
      plt.title("Bernoulli LL Solution")
      pred_grid = np.reshape(sigmoid(feature_fn(flat_plt_grid) @ w_bce),_
      →plt_grid_shape)
      plt.contourf(plt_grid[..., 0], plt_grid[..., 1], pred_grid, levels=10)
      plt.colorbar()
      s0 = plt.scatter(c0_samples[..., 0], c0_samples[..., 1], color="blue")
```

```
s1 = plt.scatter(c1_samples[..., 0], c1_samples[..., 1], color="orange")
plt.legend([s0, s1], ["c0", "c1"])
plt.xlim(-1.5, 2.5)
```

Final loss: 0.019875249179541393

[21]: (-1.5, 2.5)





2.3 3) Stochastic and Batch Gradients (4 Points)

Note You need to run the cells above first to load the data

Usually it is cheaper to approximate the gradients on a small subset of the data, i.e. a batch. We implement a single function for this

Fill in the todos in the function below.

```
num_iterations: int, batch_size: int):
   11 II II
   :param features: all samples, shape: [N x feature_dim]
   :param labels: all labels, shape: [N]
   :param initial_weights: initial weights of the classifier, shape:\sqcup
\hookrightarrow [feature_dim * K]
   :param schedule: learning rate schedule (a callable function returning the \sqcup
\rightarrow learning rate, given the iteration
   :param num_iterations: number of times to loop over the whole dataset
   :param batch size: size of each batch, should be between 1 and size of data
   return "argmin", "min", logging info
   assert 1 <= batch_size <= features.shape[0]</pre>
   # This is a somewhat simplifying assumption but for the exercise its ok
   assert features.shape[0] % batch_size == 0, "Batch Size does not evenly"
{\scriptstyle \hookrightarrow} \texttt{divide number of samples"}
   batches_per_iter = int(features.shape[0] / batch_size)
   # setup
   # yta: weights is large enough to store all weights after every batch
   # batch = e.g. 10 datasamples that were used
   # goal = track down weight reduction over iterations/batches
   # n = 100, iterations = 10, batch = 20 -> 5
   weights = np.zeros([batches_per_iter * num_iterations + 1, initial_weights.
\rightarrowshape [0]])
   # the current loss
   loss = np.zeros(batches_per_iter * num_iterations + 1)
   # get initial weights and loss
   weights[0] = initial_weights
   loss[0] = objective bern(weights[0], features, labels)
   for iteration in range(num_iterations):
       # DONE: shuffle data
       features, labels = unison_shuffled_copies(features, labels)
       local_idx = 0
       #-----
       for batch in range(batches_per_iter):
            # this is the index for weights
```

```
# Iterations = 10, N = 100, Batch-Size = 20 -> Batches per Iteration:

→ 5

           global_idx = iteration * batches_per_iter + batch
           # DONE: do stochastic gradient descent update!
           # DONE: pick the data you want to minimize
           batch_range = range(local_idx, local_idx + batch_size)
           cur_features, cur_labels = features[batch_range, :],__
→labels[batch_range]
           cur_weights = weights[global_idx]
           w_bce, _, _, _ = minimize(
                       lambda w: objective_bern(cur_weights, cur_features,_
lambda w: d_objective_bern(cur_weights, cur_features,_

    cur_labels),
                       cur_weights,
                       schedule(global idx),
                       10
           )
           # DONE: store new and minimized weights
           weights[global_idx+1] = w_bce
           local_idx = local_idx + batch_size
           \# log loss (on all samples, usually you should not use all samples \sqcup
→ to evaluate after each stochastic
           # update step)
           loss[global_idx + 1] = objective_bern(weights[global_idx + 1],
→features, labels)
   return weights[-1], loss[-1], (weights, loss)
```

The loss curve is expected to look a bit jerky due to the stochastic nature of stochastic gradient descent. If it goes down asymptotically its fine. Also, feel free to play around a bit with the schedule, num iterations and batch size to see how they affect the behaviour

```
[23]: _, l, l_info = minimize_with_sgd(features, labels, np.zeros(features.shape[1]), schedule=(lambda t: 0.25), num_iterations=25,
```

```
batch_size=5)
print("Final loss", 1)

plt.figure()
plt.title("Cross Entropy Loss")
plt.semilogy(l_info[1])
```

Final loss 0.029276563746251802

[23]: [<matplotlib.lines.Line2D at 0x11b9c4220>]

