# Density Estimation and Expectation Maximization

Maschinelles Lernen 1 - Grundverfahren WS20/21

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# **Learning Outcomes**

#### What will we learn today?

- Understand the density estimation problem
- Recap parametric density estimation and its Bayesian variant
- Understand non-parametric density estimation and its advantages
- Understand mixture models and how to train it using EM
- Analysis of the EM algorithm and why it converges

# Agenda for today

#### **Parametric Density Estimation**

- Maximum Likelihood
- Bayesian Estimation (see Bayesian Learning lecture...)

#### **Non-Parametric Density Estimation**

- Histograms
- Kernel-density estimates
- Nearest neighbour density estimates

#### **Mixture Models**

Gaussian Mixture Models (GMM)

#### The Expectation Maximization Algorithm

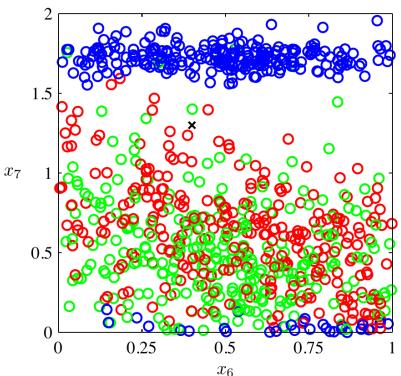
- EM decomposition
- E- and M-step
- Convergence analysis
- EM for GMMs

# **Density estimation**

# How do we get the probability distributions from this?

#### **Applications:**

- Classify (generative approaches)
- Outlier / unseen event detection
- Generate new data



# **Probability Density Estimation**

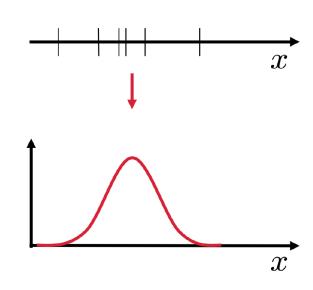
#### Training data

$$\mathcal{D} = \{oldsymbol{x}_1, oldsymbol{x}_2, \dots, oldsymbol{x}_N\}$$

Estimation

$$p(\boldsymbol{x})$$

- Methods
  - Parametric model
  - Non-parametric model
  - Mixture models



# Recap: Parametric models

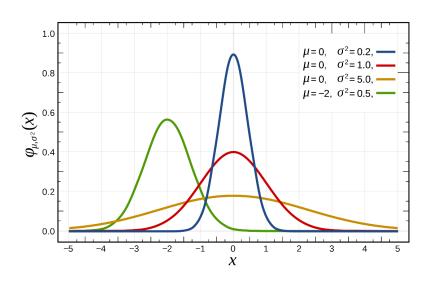
#### Most commonly used: Gaussian distribution

#### Parametric model:

$$p_{\boldsymbol{\theta}}(x) = \mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\} \stackrel{\mathfrak{F}}{\overset{\circ}{\Rightarrow}_{0.4}}$$

### **2** Parameters: $\theta = \{\mu, \sigma\}$

- Mean  $\mu$
- Variance  $\sigma^2$



# Recap: Maximum Likelihood Estimation (MLE)

#### Maximize the Log-likelihood:

$$loglik(\boldsymbol{\theta}; D) = \sum_{i} log p_{\boldsymbol{\theta}}(x_i)$$

- Assumption: Independently identically distributed (iid.) dataset
- Sums are "nicer" to optimize than products
- Log cancels exponential form (most distributions are in the exponential family)

#### The MLE solution is given by:

$$\boldsymbol{\theta}_{\mathrm{ML}} = \operatorname{argmax}_{\boldsymbol{\theta}} \operatorname{loglik}(\boldsymbol{\theta}; D)$$

### Maximum Likelihood for a Gaussian

ML estimation for a Gaussian

$$\mu, \Sigma = \operatorname{argmax}_{\theta} \operatorname{loglik}(\theta; D) = \sum_{i=1}^{N} \log \mathcal{N}(x_i | \mu, \Sigma)$$

Take the partial derivatives and set it to 0

$$\frac{\partial \text{loglike}(\boldsymbol{\theta}; \mathcal{D})}{\partial \boldsymbol{\mu}} = \mathbf{0} \qquad \qquad \frac{\partial \text{loglike}(\boldsymbol{\theta}; \mathcal{D})}{\partial \boldsymbol{\Sigma}} = \mathbf{0}$$

Which leads to the closed form solution (also see homework 1)

$$oldsymbol{\mu} = rac{1}{N} \sum_{i=1}^N oldsymbol{x}_i oldsymbol{\Sigma} = rac{1}{N} \sum_{i=1}^N (oldsymbol{x}_i - oldsymbol{\mu}) (oldsymbol{x}_i - oldsymbol{\mu})^T$$

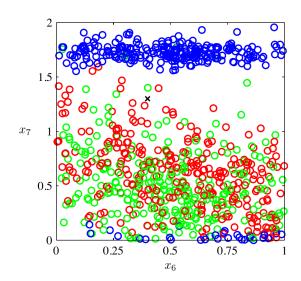
# Non-parametric Models

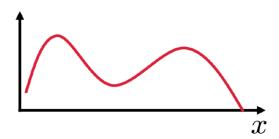
#### Does this look Gaussian?

- No! Indeed most data-sets are not Gaussian distributed. Typically we have:
  - Multi-Modality
  - Non-symmetric
  - No infinite support

#### We need more complex representations:

- Non-parametric
- Mixture models





# Non-parametric models

#### Why use Non-parametric representations?

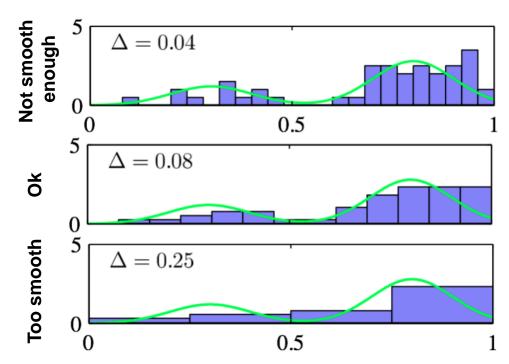
 Often we do not know what functional form the class-conditional density takes (or we do not know what class of function we need)

# Probability density is estimated directly from the data (i.e. without an explicit parametric model)

- Histograms
- Kernel density estimation (Parzen windows)
- K-nearest neighbors

# Histograms

- Discretize the input space into bins
- Count the samples per bin



# Histograms

#### **Properties**

- They are very general, because in the infinite data limit any probability density can be approximated arbitrarily well
- At the same time it is a Brute-force method

#### **Problems**

- High-dimensional feature spaces
- Exponential increase in the number of bins
- Hence requires exponentially much data
- Commonly known as the curse of dimensionality
- How to choose the size of the bins?
  - This is again a model-selection problem!

### More formal definition

- Data point x is sampled from probability density p(x)
- Probability that x falls in region R

$$p(\boldsymbol{x} \in R) = \int_{R} p(\boldsymbol{x}) d\boldsymbol{x}$$

• If R is is sufficiently small, with volume V, then p(x) is almost constant

$$p(\boldsymbol{x} \in R) = \int_{R} p(\boldsymbol{x}) d\boldsymbol{x} \approx p(\boldsymbol{x}) V$$

• We can also compute  $p(x \in R)$  from samples (If we have sufficiently large dataset)

$$p(x \in R) \approx \frac{K}{N} \Rightarrow p(x) \approx \frac{K}{NV}$$

where N is the number of total points and K is the number of points falling in the region R

# Regions

$$p(\boldsymbol{x}) pprox rac{K}{NV}$$

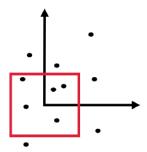
- For histograms, the regions are of equal size and span across the whole input space.
- Can we find a more adaptive representation of regions?
  - Yes, make the region always centred on the input x!

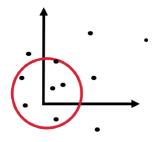
#### Kernel density estimation

- Fix V and determine K
- Example: determine the number of data points K in a fixed hypercube

#### K-nearest neighbor

- Fix K and determine V
- Example: increase the size of a sphere until K data points fall into the sphere





# Kernel density estimation

### A kernel k(x,y) "compares" two samples x and y

- Required properties for density estimation:
  - Non-negative:  $k({m x},{m y}) \geq 0$  Distance-dependent:  $k({m x},{m y}) = g(\underbrace{{m x} {m y}}_{ ext{difference }{m u}})$
- Volume:  $V = \int g({m u}) d{m u}$
- Summed kernel activation:  $K(m{x}_*) = \sum_{i=1}^{m} g(m{x}_* m{x}_i)$
- Estimated density:  $p(m{x}_*) pprox rac{K(m{x}_*)}{NV} = rac{1}{NV} \sum_{i=1}^N g(m{x}_* m{x}_i)$
- A more formal definition can be found in the kernel lecture

### Parzen Window

Kernel function: Hypercubes in d dimensions with edge length h

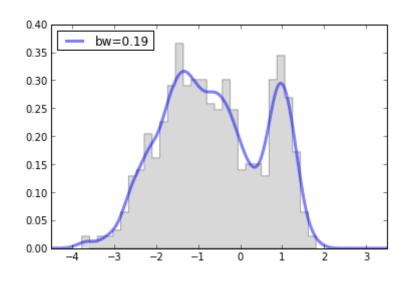
$$g(\boldsymbol{u}) = \left\{ egin{array}{ll} 1, & |u_j| \leq h/2, \ j=1\ldots d \\ 0, & \mathrm{else} \end{array} 
ight.$$
 dimensionality

Volume:

$$V = \int g(\boldsymbol{u})d\boldsymbol{u} = h^d$$

Estimated Density:

$$p(oldsymbol{x}_*) pprox rac{1}{Nh^d} \sum_{i=1}^N g(oldsymbol{x}_* - oldsymbol{x}_i)$$



- ✓ Simple to compute
- x Not very smooth

### Gaussian kernel...

#### Kernel function:

$$g(\boldsymbol{u}) = \exp\left(-\frac{\|\boldsymbol{u}\|^2}{2h}\right)$$
 bandwidth

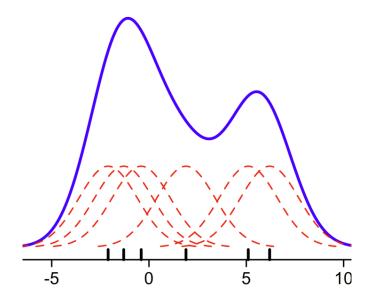
Volume:

$$V = \int g(\boldsymbol{u})d\boldsymbol{u} = \sqrt{2\pi h^d}$$
 dimensionality

#### Estimated Density:

$$p(\boldsymbol{x}_*) pprox rac{1}{NV} \sum_{i=1}^{N} g(\boldsymbol{x}_* - \boldsymbol{x}_i)$$

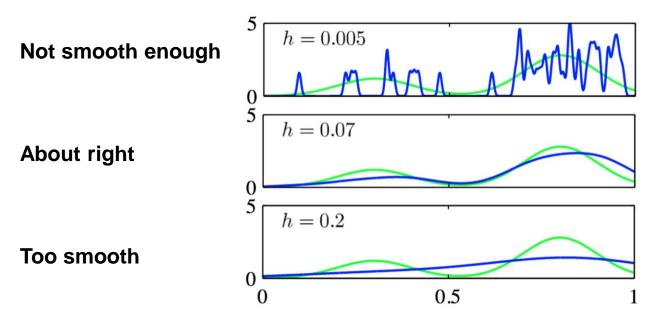
$$= \frac{1}{N\sqrt{2\pi h^d}} \sum_{\substack{i=1 \ \text{Cerhard Neumann | Machine Tearning 1 | KIT | WS 2020/2021}}^N \exp\left(-\frac{\|\boldsymbol{x}_* - \boldsymbol{x}_i\|^2}{2h}\right)$$



- ✓ Smooth
- × Infinite support
- × Requires a lot of computation

# Gaussian KDE Example

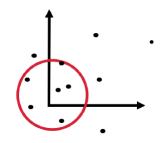
Problem with kernel methods: We have to select the kernel bandwidth h
appropriately

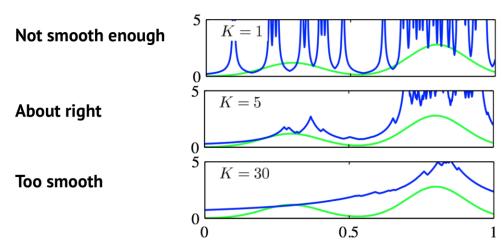


# K-nearest neighbour density estimation

#### K-nearest neighbour: Fix K and determine V

 Example: increase the size of a sphere until K data points fall into the sphere





### Model-Selection

#### Nonparametric probability density estimation

- Histograms: Size of the bins?
  - too large: too smooth
  - too small: not smooth enough
- Kernel density estimation: Kernel bandwidth?
  - h too large: too smooth
  - h too small: not smooth enough
- K-nearest neighbor: Number of neighbors?
  - K too large: too smooth
  - K too small: not smooth enough

#### A general problem of many density estimation approaches

Select via cross-validation: Select model with highest likelihood on test-set

# Agenda for today

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- Nearest neighbour density estimates

#### **Mixture Models**

Gaussian Mixture Models (GMM)

#### The Expectation Maximization Algorithm

- EM decomposition
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# Mixture Models

### Mixture Models

#### Parametric models

- Gaussian, Neural Networks, ...
- ✓ Good analytic properties
- ✓ Simple
- ✓ Small memory requirements
- ✓ Fast
- Limited representation power (most parametric distributions have only one mode)

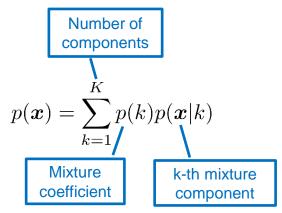
#### Non-Parametric models

- Kernel-density estimation, k-NN
- General (can represent any distribution)
- × Curse of dimensionality
- × Large memory requirements
- × Slow

- Mixture models combine the advantages of both worlds
- Key idea: Create a complex distribution by combining simple ones (e.g. Gaussians)

### Mixture model

# A mixture distribution is the sum of individual distributions:



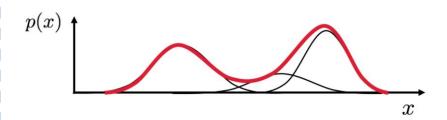
• In the limit with many / infinite components, this can approximate any smooth density

### **Example: Mixture of Gaussians (MoG)**

Individual Gaussians



Sum of Gaussians



### Mixture of Gaussians

#### Mixture coefficient:

$$p(k) = \pi_k$$
, with  $0 \le \pi_k \le 1$ ,  $\sum_k \pi_k = 1$ 

Mixture component:

$$p(\boldsymbol{x}|k) = \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

Mixture distribution:

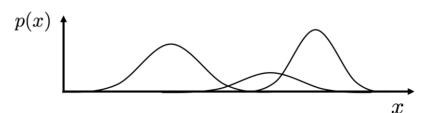
$$p(\boldsymbol{x}) = \sum_{k=1}^{K} \pi_k \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Always integrates to 1
- Parameters of the mixture

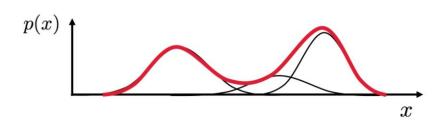
$$\boldsymbol{\theta} = \{\pi_1, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1, \dots, \pi_K, \boldsymbol{\mu}_K, \boldsymbol{\Sigma}_K\}$$

### **Example: Mixture of Gaussians (MoG)**

Individual Gaussians



Sum of Gaussians



### Maximum Likelihood of a mixture

(Marginal-)Log-Likelihood with N iid. points

$$\mathcal{L} = \log L(\boldsymbol{\theta}) = \sum_{i=1}^{N} \log \underbrace{p(\boldsymbol{x}_i | \boldsymbol{\theta})}_{\text{marginal}} = \sum_{i=1}^{N} \log \underbrace{\left(\sum_{k=1}^{K} \pi_k \mathcal{N}(\boldsymbol{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)\right)}_{\text{non-exponential family}}$$

Q: Can we do gradient descent?

$$\begin{split} \frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}_j} &= \sum_{i=1}^N \frac{\pi_j \mathcal{N}(\boldsymbol{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}{\sum_{k=1}^K \pi_k \mathcal{N}(\boldsymbol{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)} \boldsymbol{\Sigma}_j^{-1} (\boldsymbol{x}_i - \boldsymbol{\mu}_j) \\ &= \sum_{i=1}^N \frac{p(j) p(\boldsymbol{x}_i | j)}{p(\boldsymbol{x}_i)} \boldsymbol{\Sigma}_j^{-1} (\boldsymbol{x}_i - \boldsymbol{\mu}_j) \\ &= \sum_{i=1}^N \boldsymbol{\Sigma}_j^{-1} (\boldsymbol{x}_i - \boldsymbol{\mu}_j) p(j | \boldsymbol{x}_i) \end{split}$$

- x Gradient depends on all other components (cyclic dependency)
- × No closed form solution
- x Typically very slow convergence
- A: Yes, but the sum (marginalization) does not go well with the log

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# Basics: Kullback-Leibler Divergences

#### The KL-divergence is a important similarity measure for distributions

$$\mathrm{KL}(q(oldsymbol{x})||p(oldsymbol{x})) = \sum_{oldsymbol{x}} q(oldsymbol{x}) \log rac{q(oldsymbol{x})}{p(oldsymbol{x})}$$

- Its always non-negative
- If its zero, both distributions are the same:
- It is non-symmetric (hence, its not a distance metric):

$$\mathrm{KL}(q||p) \ge 0$$

$$KL(q||p) = 0 \iff q = p$$

$$KL(q||p) \neq KL(p||q)$$

- Can be used to find different approximations of distributions
- Used a lot in Variational Inference, Reinforcement Learning, Information theory...

### Latent Variable Models

#### Mixture models are an instance of latent variable models

- Examples: Missing data, latent factors, mixtures, ...
- Observed variables: x, Latent variables: z (e.g., index of mixture component)
- Parametric model:  $p(\boldsymbol{x}, \boldsymbol{z} | \boldsymbol{\theta})$

• Marginal distribution: 
$$p(\boldsymbol{x}|\boldsymbol{\theta}) = \sum_{z} p(\boldsymbol{x}, \boldsymbol{z}|\boldsymbol{\theta}), \quad p(\boldsymbol{x}|\boldsymbol{\theta}) = \int_{z} p(\boldsymbol{x}, \boldsymbol{z}|\boldsymbol{\theta}) d\boldsymbol{z}$$

#### (Marginal) Log-Likelihood:

$$\mathcal{L} = \log L(oldsymbol{ heta}) = \sum_{i=1}^N \log p(oldsymbol{x}_i | oldsymbol{ heta}) = \sum_{i=1}^N \log \left( \sum_{oldsymbol{z}} p(oldsymbol{x}_i, oldsymbol{z}) 
ight)$$

... which is hard to optimize for all latent variable models (due to log of a sum)

# Expectation-Maximization (EM)

# Expectation-Maximization (EM) is a general algorithm for estimating latent variable models

- Most common application: Gaussian Mixture models
- ... but many other (deep) models as well
- Its extension is called Variational Bayes, which is underlying variational auto-encoder and other variational inference techniques
- Very hot research topic... pays off to look into the math of it

#### EM can be derived in 2 ways:

- Jensen's inequality (not covered)
- Decomposition in lower-bound and KL-term

# **Expectation-Maximization (EM)**

#### EM uses a lower bound of the marginal log-likelihood for the optimization

For simplicity, lets consider only a single data-point first

$$\underbrace{\log p(\boldsymbol{x}|\boldsymbol{\theta})}_{\text{marginal log-like}} = \underbrace{\sum_{z} q(z) \log \frac{p(\boldsymbol{x}, z|\boldsymbol{\theta})}{q(z)}}_{\text{Lower Bound } \mathcal{L}(q, \boldsymbol{\theta})} + \underbrace{\sum_{z} q(z) \log \frac{q(z)}{p(z|\boldsymbol{x})}}_{\text{KL Divergence: KL}(q(z)||p(z|\boldsymbol{x}))}$$

- Where q(z) is called the variational / auxiliary distribution
  - This decomposition holds for any q(z)
  - By introducing q(z), the optimization will become much simpler
- Why is that the same?
  - We can use Bayes rule for  $p(z|x) = \frac{p(x,z|\theta)}{p(x|\theta)}$  and all terms except  $p(x|\theta)$  cancel

# **EM-Decomposition**

#### **Derivation:**

$$\begin{split} \log p(\boldsymbol{x}) &= \sum_{z} q(\boldsymbol{z}) \log p(\boldsymbol{x}) \\ &= \sum_{z} q(\boldsymbol{z}) \left( \log p(\boldsymbol{x}, \boldsymbol{z}) - \log p(\boldsymbol{z} | \boldsymbol{x}) \right) \\ &= \sum_{z} q(z) \left( \log p(\boldsymbol{x}, z) - \log q(z) \right. \\ &\quad + \log q(z) - \log p(\boldsymbol{z} | \boldsymbol{x}) \right) \\ &= \underbrace{\sum_{z} q(z) \log \frac{p(\boldsymbol{x}, z)}{q(z)}}_{\text{Lower Bound } \mathcal{L}(q)} + \underbrace{\sum_{z} q(z) \log \frac{q(z)}{p(z | \boldsymbol{x})}}_{\text{ELovergence: KL}(q(z) || p(z | \boldsymbol{x}))} \end{split}$$

1. Introduce variational distribution q(z)

2. Use Bayesian theorem

$$p(\boldsymbol{x}) = \frac{p(\boldsymbol{x}, z)}{p(\boldsymbol{z}|x)}$$

3. Add and substract  $\log q(z)$ 

4. Write as 2 sums

# **EM** Decomposition

### Marginal Likelihood decomposes in 2 terms: $\log p(\boldsymbol{x}|\boldsymbol{\theta}) = \mathcal{L}(q,\boldsymbol{\theta}) + \mathrm{KL}(q(z)||p(z|\boldsymbol{x}))$

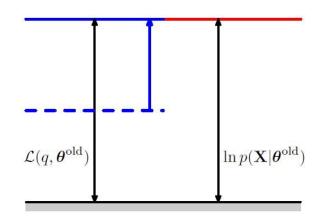
- Lower bound  $\mathcal{L}(q, m{ heta}) = \sum_z q(z) \log p(m{x}, z | m{ heta}) \sum_z q(z) \log q(z)$ 
  - Contains  $\log p({m x},z|{m heta})$  instead of  $\log p({m x}|{m heta}) = \log \sum p({m x},z|{m heta})$
  - ... which is much easier to optimize (convex for most distributions)
  - Each  $\log p(\boldsymbol{x}, z|\boldsymbol{\theta})$  is weighted by q(z)
- Why is it a lower bound?
  - Since  $\mathrm{KL}(q||p) \geq 0$  it follow that  $\mathcal{L}(q, \boldsymbol{\theta}) \leq \log p(\boldsymbol{x}|\boldsymbol{\theta})$

# **Expectation-Maximization Steps**

#### EM iteratively applies 2 steps:

- (E)xpectation-step:  $q(z) = \underset{q}{\operatorname{arg \, min}} \operatorname{KL}(q(z)||p(z|\boldsymbol{x}))$ 
  - Find q(z) that minimizes KL
  - Can be done in closed form for discrete z:

$$q(z) = p(z|\boldsymbol{x}, \boldsymbol{\theta}_{\mathrm{old}}) = \frac{p(\boldsymbol{x}, z|\boldsymbol{\theta}_{\mathrm{old}})}{\sum_{z} p(\boldsymbol{x}, z|\boldsymbol{\theta}_{\mathrm{old}})}$$



#### Observations:

- The marginal log-likelihood  $\log p(\boldsymbol{x}|\boldsymbol{\theta})$  is unaffected by the E-step
- As KL is minimized, lower bound has to go up
- After the E-step  $\mathrm{KL}(q(z)||p(z|\boldsymbol{x})) = 0$  and therefore, the lower bound is tight, i.e.:

$$\log p(\boldsymbol{x}|\boldsymbol{\theta}) = \mathcal{L}(q,\boldsymbol{\theta})$$

# **Expectation-Maximization Steps**

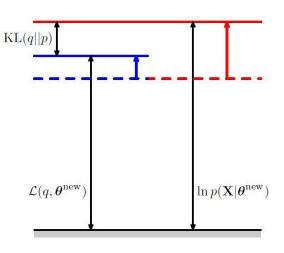
#### **EM** iteratively applies 2 steps:

(M)aximization-step:

$$\boldsymbol{\theta} = \underset{\boldsymbol{\theta}}{\operatorname{arg\,max}} \mathcal{L}(q, \boldsymbol{\theta}) = \underset{\boldsymbol{\theta}}{\operatorname{arg\,max}} \sum_{z} q(z) \log p(\boldsymbol{x}, z | \boldsymbol{\theta}) + \operatorname{const}$$

- Maximize lower bound with respect to θ
- Also called the complete-data likelihood
- Each possible value of the missing data is weighted by

$$q(z) = p(z|\boldsymbol{x}, \boldsymbol{\theta}_{\mathrm{old}})$$



# **EM Convergence Properties**

#### EM improves the lower bound

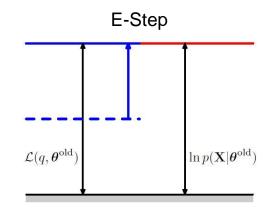
$$\mathcal{L}(q_{\text{new}}, \boldsymbol{\theta}_{\text{new}}) \geq \mathcal{L}(q_{\text{old}}, \boldsymbol{\theta}_{\text{old}})$$

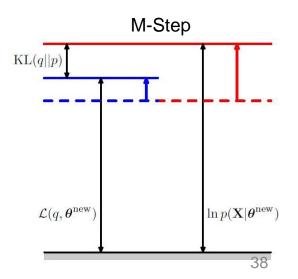
- M-step: Lower bound is maximized
- E-step: KL is set to 0, lower bound has to go up

### EM improves the marginal likelihood

$$\log p(\boldsymbol{x}|\boldsymbol{\theta}_{\text{new}}) \geq \log p(\boldsymbol{x}|\boldsymbol{\theta}_{\text{old}})$$

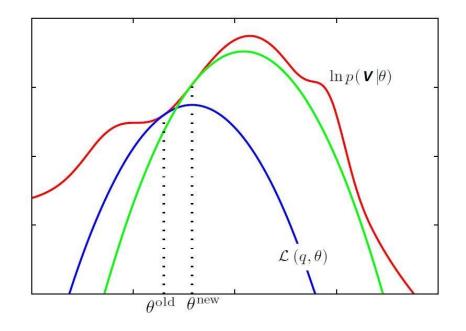
- M-step: Lower bound increases and KL increases (can't get smaller than 0)
- E-step: Marginal likelihood is unaffected





### Illustration of EM

- Lower bound (blue curve) is a convex approximation of the marginal likelihood (red curve)
  - Maximum of lower bound can be easily obtained ( $oldsymbol{ heta}^{ ext{new}}$ )
  - Closed form solutions available, no gradient descent required
- Compute new lower bound for  $oldsymbol{ heta}^{ ext{new}}$  (green curve)
- Due to the local approximation of the lowerbound, EM can only find local optima



### EM for full dataset

#### For all data-points, the lower bound is given by:

$$\mathcal{L}(q, \boldsymbol{\theta}) = \sum_{i} \sum_{z} q_i(z) \log p(\boldsymbol{x}_i, z | \boldsymbol{\theta}) - \sum_{z} q_i(z) \log q_i(z)$$

- One latent variable  $z_i$  per data-point
- If z is discrete with K different values, than

$$q_i(z=k) = p(z=k|\boldsymbol{x}_i,\boldsymbol{\theta}_{\mathrm{old}})$$

can be represented as a N x K matrix

• We will write  $q_{ik} = q_i(z=k)$ 

### **EM for Mixture of Gaussians**

#### Mixture coefficient:

$$p(k) = \pi_k$$
, with  $0 \le \pi_k \le 1$ ,  $\sum_k \pi_k = 1$ 

#### Mixture component:

$$p(\boldsymbol{x}|k) = \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

#### Mixture distribution:

$$p(\boldsymbol{x}) = \sum_{k=1}^K \pi_k \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Always integrates to 1
- Parameters of the mixture

$$\boldsymbol{\theta} = \{\pi_1, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1, \dots, \pi_K, \boldsymbol{\mu}_K, \boldsymbol{\Sigma}_K\}$$

### E-Step

Compute "responsibilities"

$$q_{ik} = \frac{\pi_k \mathcal{N}(\boldsymbol{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\boldsymbol{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

 How much component k contributes to generation of x<sub>i</sub> according to current mixture model

### **EM for Mixture of Gaussians**

#### Mixture coefficient:

$$p(k) = \lambda_k$$
, with  $0 \le \lambda_k \le 1$ ,  $\sum_k \lambda_k = 1$ 

#### Mixture component:

$$p(\boldsymbol{x}|k) = \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

#### Mixture distribution:

$$p(\boldsymbol{x}) = \sum_{k=1}^K \pi_k \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Always integrates to 1
- Parameters of the mixture

$$\boldsymbol{\theta} = \{\pi_1, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1, \dots, \pi_K, \boldsymbol{\mu}_K, \boldsymbol{\Sigma}_K\}$$

**M-Step:** 
$$\theta = \underset{\theta}{\operatorname{arg max}} \sum_{i} \sum_{k} q_{ik} \log p(k) p(\boldsymbol{x}_i|k)$$

- We can separate updates of single components and coefficients
  - just additive objectives in lower bound

#### Update coefficients:

$$\pi = \operatorname*{arg\,max}_{\pi} \sum_{i} \sum_{k} q_{ik} \log \pi_{k}$$

Update components:

$$oldsymbol{\mu}_k, oldsymbol{\Sigma}_k = rg \max_{oldsymbol{\mu}_k, oldsymbol{\Sigma}_k} \sum_i q_{ik} \log \mathcal{N}(oldsymbol{x}_i | oldsymbol{\mu}_k, oldsymbol{\Sigma}_k)$$

- Each data-point is weighted by q<sub>ik</sub>
- Weighted maximum likelihood estimate

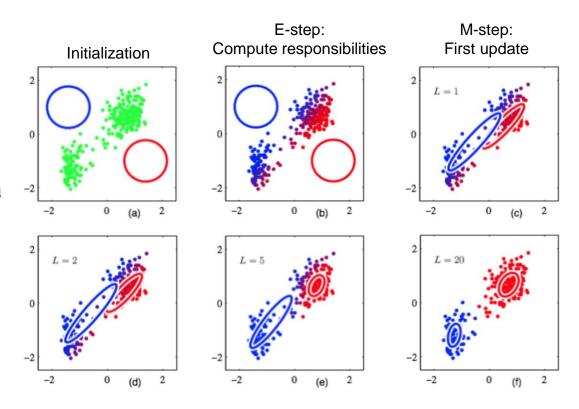
### **EM for Mixture of Gaussians**

#### **Weighted Maximum Likelihood updates:**

- Update coefficients:  $\pi = \argmax_{\boldsymbol{\pi}} \sum_{i} \sum_{k} q_{ik} \log \pi_k$ 
  - Result:  $\pi_k = \frac{\sum_i q_{ik}}{\sum_k \sum_i q_{ik}} = \frac{\sum_i q_{ik}}{N}$
- Update components:  $\mu_k, \Sigma_k = rg \max_{\mu_k, \Sigma_k} \sum_i q_{ik} \log \mathcal{N}(\boldsymbol{x}_i | \boldsymbol{\mu}_k, \Sigma_k)$ 
  - Mean:  $oldsymbol{\mu}_k = rac{\sum_i q_{ik} oldsymbol{x}_i}{\sum_i q_{ik}}$
  - Covariance:  $m{\Sigma}_k = rac{\sum_i q_{ik} (m{x}_i m{\mu}_k) (m{x}_i m{\mu}_k)^T}{\sum_i q_{ik}}$

### Illustration

- Each component represents a cluster in the data set
- EM is very sensitive to the initialization



### EM versus k-means

#### K-means can be seen as special case of EM with:

- Co-Variances are always set to 0 (in the limit)
- E-Step / Assignment Step:
  - responsibilities  $q_{ik}$  of nearest cluster k are set to 1, all other values are 0
- M-Step / Adjustment Step:
  - Update for the mean is the same
  - Co-Variances are ignored (set to close to 0)
- EM is harder to learn than k-means but it also gives you variances and densities
- Often k-means is used to initialize the means for EM

### Practical considerations...

#### How many mixture components do we need?

- More components will typically lead to a better likelihood
- But are more components necessarily better? Not always, because of overfitting!
- It's again a model-selection problem (cross-validate on a validation-set)
- Bayesian methods can be used to integrate out number of components (tricky to get them to work)

#### How do we initialize:

- EM can give very poor results with wrong initialization
- Most common approach:
  - Use k-means (simple clustering algorithm) to initialize the centers
  - Use a fixed value for the covariance

### **Additional Notes**

#### EM assumes that E-step can set the KL to zero:

- Lower bound is tight
- Marginal likelihood always improves

#### For more complex latent variable models (e.g. continuous z), this is not possible:

• In this case, we also have to use an approximate q(z) that minimizes

$$q(oldsymbol{z}) = rg \min_q \mathrm{KL}(q(oldsymbol{z}) || p(oldsymbol{z} || oldsymbol{x}))$$

- Hence,  $\mathrm{KL}(q(\boldsymbol{z})||p(\boldsymbol{z}|\boldsymbol{x}))>0$  after the E-step
  - No guaranteed improvement of the marginal likelihood
  - Still improvement of the lower bound
- Algorithms that do that are called Variational Bayes / Variational Inference
  - Very active research, underlying algorithm of many deep learning architectures (e.g. variational autoencoder)

# Takeaway messages

#### You know now:

- The difference between parametric and non-parametric models
- Different non-parametric models (histogram, kernel density estimation and k-nearest neighbors)
- What mixture models and latent variable models are
- What the Expectation-Maximization idea and algorithm are
- Why does EM converge
- How to apply EM to GMMs



# Self-test questions

- What are parametric methods and how to obtain their parameters?
- How many parameters have non-parametric methods?
- What are mixture models?
- Should gradient methods be used for training mixture models?
- How does the EM algorithm work?
- What is the biggest problem of mixture models?
- How does EM decomposes the marginal likelihood?
- Why does EM always improve the lower bound?
- Why does EM always improve the marginal likelihood