

Chapter 1: Probability Theory

1.1 Probability Inequalities

Theorem 1.1 (The Gaussian Tail Inequality) Let $X \sim N(0, 1)$. Then

$$\mathbb{P}(|X| > \epsilon) \leq \frac{2e^{-\epsilon^2/2}}{\epsilon}.$$

If X_1, \dots, X_n i.i.d $\sim N(0, 1)$, then

$$\mathbb{P}(|\bar{X}_n| > \epsilon) \leq \frac{2}{\sqrt{n}\epsilon} e^{-n\epsilon^2/2} \stackrel{\text{large } n}{\leq} e^{-n\epsilon^2/2}. \quad (1.1)$$

Lemma 1.2 Supposed that $a \leq X \leq b$. Then

$$\mathbb{E}(e^{tX}) \leq e^{t\mu} e^{\frac{t^2(b-a)^2}{8}}. \quad (1.2)$$

Theorem 1.3 (Hoeffding's Inequality) Let Y_1, \dots, Y_n be i.i.d observations such that $\mathbb{E}(Y_i) = \mu$ and $a \leq Y_i \leq b$. Then for any $\epsilon > 0$,

$$\mathbb{P}(|\bar{Y}_n - \mu| \geq \epsilon) \leq 2e^{-2n\epsilon^2/(b-a)^2}. \quad (1.3)$$

Theorem 1.4 (Bernstein Inequality) If X_1, \dots, X_n are independent and $\mathbb{E}(X_i) = 0$. if $|X_i| \leq b$, then we have

$$P\left(\left|\sum_{i=1}^n X_i\right| \geq \epsilon\right) \leq \exp\left[-\frac{\epsilon^2}{2\sum_{i=1}^n \mathbb{E}(X_i)^2 + 2\epsilon/3b}\right]. \quad (1.4)$$

Remark: If $\sum_{i=1}^n \mathbb{E}(X_i^2) = O(n)$, then let $\epsilon = C \frac{\sqrt{\log n}}{\sqrt{n}}$ we have

$$P\left(\left|\frac{1}{n} \sum_{i=1}^n X_i\right| \geq C \frac{\sqrt{\log n}}{\sqrt{n}}\right) \leq n^{-C'}. \quad (1.5)$$

Note that for i.i.d. random variables $X_i, i = 1, 2, \dots, n$ (satisfy CLT conditions), we have $\frac{1}{n} \sum_{i=1}^n (X_i - \mathbb{E}(X_i)) = O_p(1/\sqrt{n})$, but it's not a uniform result.

For example, we want to obtain the convergence rate of $\max_{k \leq N} \frac{1}{n} \sum_{i \in H_k} (X_i - \mathbb{E}(X_i))$,

$$P\left(\max_{k \leq N} \frac{1}{n} \sum_{i \in H_k} (X_i - \mathbb{E}(X_i)) \geq ca_n\right) \leq \sum_{k=1}^N P\left(\frac{1}{n} \sum_{i \in H_k} (X_i - \mathbb{E}(X_i)) \geq ca_n\right),$$

then we need Bernstein inequality (1.4) rather than CLT. Combine (1.4) and (1.5), we have

$$\max_{k \leq N} \frac{1}{n} \sum_{i \in H_k} (X_i - \mathbb{E}(X_i)) = O_p\left(\frac{\sqrt{\log n}}{\sqrt{n}}\right)$$

Theorem 1.5 (McDiarmid) Let X_1, \dots, X_n be independent r.v. Supposed that

$$\sup_{x_1, \dots, x_n, x'_i} |g(x_1, \dots, x_{i-1}, x_i, x_{i+1}, \dots, x_n) - g(x_1, \dots, x_{i-1}, x'_i, x_{i+1}, \dots, x_n)| \leq c_i$$

for $i = 1, 2, \dots, n$. Then

$$\mathbb{P}(g(X_1, \dots, X_n) - \mathbb{E}(g(X_1, \dots, X_n)) \geq \epsilon) \leq \exp \left\{ -\frac{2\epsilon^2}{\sum_{i=1}^n c_i^2} \right\}. \quad (1.6)$$

1.2 Moment Inequalities

1. Hölder's inequality: $E|XY| \leq (E|X|^p)^{1/p} (E|Y|^q)^{1/q}$
2. Liapounov's inequality: $(E|X|^r)^{1/r} \leq (E|X|^s)^{1/s}$, where r and s are constants satisfying $1 \leq r \leq s$
3. Minkowski's inequality: $(E|X + Y|^p)^{1/p} \leq (E|X|^p)^{1/p} + (E|Y|^p)^{1/p}$, where X and Y are random variables and $p \geq 1$ is a constant.
4. Esseen and von Bahr (1965): Let X_1, \dots, X_n be independent random variables with mean 0 and $E|X_i|^p < \infty, i = 1, 2, \dots, n$, where p is a constant in $[1, 2]$, then

$$E \left| \sum_{i=1}^n X_i \right|^p \leq C_p \sum_{i=1}^n E|X_i|^p \quad (1.7)$$

where C_p is a constant depending only on p .

5. C_r inequality: $E|X + Y|^r \leq C_r (E|X|^r + E|Y|^r)$, where $C_r = 0$ if $0 < r \leq 1$ and $C_r = 2^{r-1}$ if $r > 1$.
6. Marcinkiewicz and Zygmund: For $p \geq 2$,

$$E \left| \sum_{i=1}^n X_i \right|^p \leq \frac{C_p}{n^{1-p/2}} \sum_{i=1}^n E|X_i|^p. \quad (1.8)$$

where C_p is a constant depending only on p .

7. Hájek and Rényi: Let Y_1, \dots, Y_n be independent random variables having finite variances, then we have

$$P \left(\max_{1 \leq l \leq n} c_l \left| \sum_{i=1}^l (Y_i - EY_i) \right| > t \right) \leq \frac{1}{t^2} \sum_{i=1}^n c_i^2 \text{Var}(Y_i), \quad t > 0 \quad (1.9)$$

where c_i 's are positive constants satisfying $c_1 \leq c_2 \leq \dots \leq c_n$. If $c_i = 1$, then inequality reduces to famous Kolmogorov's inequality.

Corollary 1.6 For independent random variables X_1, \dots, X_n with $\sup_n \mathbb{E}(X_n) < \infty$, we have

$$P \left(\left| \frac{1}{n} \sum_{i=1}^n (X_i - EX_i) \right| > t \right) \leq \begin{cases} O(n^{1-p}) & \text{if } 1 < p < 2 \\ O(n^{-p/2}) & \text{if } p \geq 2 \end{cases}$$

1.3 Asymptotic theory

1.3.1 Convergence Mode

Theorem 1.7 For random k -vectors X, X_1, X_2, \dots on a probability space, $X_n \rightarrow \text{a.s.} X$ if and only if for every $\epsilon > 0$

$$\lim_{n \rightarrow \infty} P \left(\bigcup_{m=n}^{\infty} \{\|X_m - X\| > \epsilon\} \right) = 0.$$

Theorem 1.8 Let $\{X_n\}$ be a sequence of **independent** random variables, then $X_n \rightarrow \text{a.s.} 0$ if and only if for every $\epsilon > 0$

$$\sum_{i=1}^n P(|X_n| \geq \epsilon) < \infty.$$

Remark:

1. $X_n \rightsquigarrow X$ and $Y_n \rightsquigarrow Y$ does not imply that $X_n + Y_n \rightsquigarrow X + Y$.
2. $X_n \xrightarrow{P} b$ does not imply that $\mathbb{E}(X_n) \rightarrow b$.
3. If $X_n \xrightarrow{P} X$ and $|X_n|^r$ is uniformly integrable, then $X_n \xrightarrow{r} X$

Theorem 1.9 Let X_1, X_2, \dots be a sequence of identically distributed random variables with a finite $E|X_1|$ and let $Y_n = n^{-1} \max_{i \leq n} |X_i|$, then (1) $Y_n \rightarrow_{L_1} 0$ (2) $Y_n \rightarrow_{\text{a.s.}} 0$

1.3.2 Stochastic convergence rate

Theorem 1.10 Let $X, X_1, X_2, \dots, Y_1, Y_2, \dots, Z_1, Z_2, \dots$ be random variables, then the following statements hold:

- (a) If $X_n \rightarrow_d X$, then $X_n = O_p(1)$.
- (b) If $X_n = O_p(Z_n)$ and $P(Y_n = 0) = 0$, then $X_n Y_n = O_p(Y_n Z_n)$.
- (c) If $X_n = O_p(Z_n)$ and $Y_n = O_p(Z_n)$, then $X_n + Y_n = O_p(Z_n)$.
- (d) If $E|X_n| = O(a_n)$, then $X_n = O_p(a_n)$, where $a_n \in (0, \infty)$.
- (e) If $X_n \rightarrow_{\text{a.s.}} X$, then $\sup_n |X_n| = O_p(1)$.

1.4 The law of large numbers

Theorem 1.11 (Jun Shao p62) Let X_1, X_2, \dots be a sequence of independent identically distributed random variables,

(i) (The WLLN) A necessary and sufficient condition for the existence of a sequence of real numbers $\{a_n\}$ for which

$$\frac{1}{n} \sum_{i=1}^n X_i - a_n \rightarrow_p 0$$

is that $nP(|X_1| > n) \rightarrow 0$, in which case take $a_n = E(X_1 I_{\{|X_1| \leq n\}})$.

(ii) (The SLLN) A necessary and sufficient condition for the existence of a constant c for which

$$\frac{1}{n} \sum_{i=1}^n X_i \rightarrow_{a.s.} c$$

is that $\mathbb{E}(|X_1|) < \infty$, in which case take $c = E(X_1)$ and

$$\frac{1}{n} \sum_{i=1}^n c_i (X_i - EX_1) \rightarrow_{a.s.} 0$$

for any bounded sequence of real numbers $\{c_i\}$.

Theorem 1.12 (Jun Shao p65) Let X_1, X_2, \dots be a sequence of independent random variables with finite expectations.

(i) (The SLLN) If there is a constant $p \in [1, 2]$ such that

$$\sum_{i=1}^{\infty} \frac{E|X_i|^p}{i^p} < \infty$$

then

$$\frac{1}{n} \sum_{i=1}^n (X_i - EX_i) \rightarrow_{a.s.} 0.$$

(ii) (The WLLN) If there is a constant $p \in [1, 2]$ such that

$$\lim_{n \rightarrow \infty} \frac{1}{n^p} \sum_{i=1}^n E|X_i|^p = 0$$

then

$$\frac{1}{n} \sum_{i=1}^n (X_i - EX_i) \rightarrow_p 0.$$