

Applying Urinary Biomarkers 8-isoprostane to Understand the Health Effects of NO₂ Exposure

https://github.com/Yang190809/data_repository_yang

Yang Wang

Abstract

Mounting evidence shows that exposure to NO₂ generates oxidative stress. Oxidative stress can cause lipid damage, which plays an important role in various respiratory and cardiovascular diseases. Urinary 8-isoprostane is a product of lipid peroxidation which can reflect systemic lipid damage. I built linear mix models to analyze the relationship between the level of urinary 8-isoprostane with the level of NO₂ exposure. I found that short term NO₂ exposure was associated with lipid peroxidation, reflected by increased concentrations of urinary 8-isoprostane associated with increasing exposure. One IQR (7.41 µg/m³) incremental change of 12-h NO₂ exposure was associated with an increase of 8-isoprostane level by 19.45% (95% CI: 14.37%, 24.55%, p-value <0.001). One IQR (8.64 µg/m³) incremental change of 24-h NO₂ exposure was associated an increase in 8-isoprostane level by 27.69% (95% CI: 20.50%, 34.95% , p-value <0.001). One IQR (4.47 µg/m³) incremental change of 1-week NO₂ exposure was associated with an increase in 8-isoprostane level by 25.15% (95% CI: 14.14%,36.31%, p-value <0.05). One IQR (2.74 µg/m³) incremental change of 2-week NO₂ exposure was associated with an increase in 8-isoprostane level by 15.68% (95% CI: 8.87%,22.65%, p-value <0.05).

Contents

1	Rationale and Research Questions	5
2	Dataset Information	5
2.1	Data Wrangling	7
3	Exploratory Analysis	22
4	Analysis	25
4.1	Result 1	26
4.2	Result 2	26
4.3	Result 3	28
4.4	Result 4	28
5	Summary and Conclusions	30
6	References	31

List of Tables

4	Summary of NO2 exposure	10
---	-----------------------------------	----

List of Figures

1	Plot of 8-isoprostane	23
2	Plot of log(8-isoprostane)	24
3	Relationship between log(8-isoprostane) and 12-hour NO2 exposure	26
4	Relationship between log(8-isoprostane) and 24-hour NO2 exposure	27
5	Relationship between log(8-isoprostane) and 1-week NO2 exposure	28
6	Relationship between log(8-isoprostane) and 2-week NO2 exposure	29

1 Rationale and Research Questions

Exposure to NO₂ associated with cardiovascular disease, lung function impairment and asthma (Mölter, A., et. al.,2013, Collart, P., et. al.,2018, Takenoue, Y., et. al., 2012). Mounting evidence shows that exposure to NO₂ generates oxidative stress (Hashemzadeh, B., et. al.,2019). Oxidative stress can cause lipid damage (Black, C. N., et. al., 2017), which plays an important role in various respiratory and cardiovascular diseases. (Zanolin, M. E., et. a.,2015, Lakshmi, S. V., et. al., 2009). I applied 8-isoprostane to investigate the health effects of short term NO₂ exposure.

8-isoprostane is the product of the oxidized cell membrane after being attacked by reactive oxygen species such as peroxides. Therefore, 8-isoprostane can reflect lipid damage (Danielsen et al., 2009). 8-isoprostane in urine does not have a diurnal variation and has been proven to be a stable biomarker showing systemic lipid damage. It has been applied in studying diseases such as type 2 diabetes mellitus, obesity, coronary heart disease, asthma, and acute respiratory distress syndrome (Nuernberg, A. M., et al., 2008). Nevertheless, the use of this biomarker in the research of air pollution and its health effects is scarce. One study in Iran discovered a significant positive relationship between short-term NO₂ exposure and 8-isoprostane in exhaled breath condensate (EBC) in healthy children aged 12-13 years old (Hashemzadeh, B., et. al., 2019). One study in New York City found that the increases in 1- to 5-day averages of nitrogen dioxide were significantly associated with increases in 8-isoprostane in EBC among 18-year old healthy and asthma affected individuals (Patel, M. M., et. al., 2013). My research question is: If the NO₂ exposure is positively associated with urinary 8-isoprostane? Among periods of 12-hour,24-hour,1-week and 2-weeks, which period has the most significant association with urinary 8-isoprostane?

2 Dataset Information

##Data Collection All the data are obtained from Jim Zhang’s lab. The urine samples were collected from participants of a previously described study conducted in Changsha, China by Dr. Zhang and other researchers (Day, D. B., et. al.,2018). The study recruited 89 healthy individuals (age>18years old) who were living and working at the Broad Company Campus and divided them into 2 groups, Group A (n = 36) and Group B (n = 53). The study periods include pre-intervention, intervention, and post-intervention, lasting 5 weeks. Urine samples were collected from each participant once during the pre-intervention, twice during the intervention, and once during the post-intervention. The study measured hourly indoor and outdoor concentrations of PM_{2.5}, O₃, NO₂, and SO₂, surveyed each participant’s time-activity, and calculated the participants’ exposure for each pollutant over 12h, 24h, 1 week and 2 week periods, which were counted backward from the visit points.

The file 8-iso is the data set of urinary biomarker 8-isoprostane.The file urine_info is the data set that has information about the subjects who provided the urine samples. The urine list is a data set with lists of information of sample ID, subject ID and visits. The meaning of the columns as well as units and class of the data in each folder is listed below. Column

names without descriptors are irrelevant to this study.

##Urine List The urine list is a data set with lists of information of sample ID, subject ID and visits.

Column Name	Meaning	units	class of the data
Sample ID	the identity number of the samples	NA	integer
Subject ID	the identity number of the subjects	NA	integer
visit	the number of the 4 visits (1,2,3 or 4)	NA	integer

kable(Table1,caption="Urine list raw data set")

##8-isoprostane Data Set The 8-isoprostane data set has information about 8-isoprostane concentration tested in the mass spectrum and the concentration of 8-isoprostane in original urine samples. The limit of detection for 8-isoprostane was 0.016ng/ml. Any value which is below 0.016 in the column of calculated Conc in 8-is data set should be excluded as an error.

Column Name	Meaning	Units	Class of the data
Sample ID	the identity number of the samples	NA	integer
Calculated Conc	the concentration tested by the machine	ng/ml	numeric
Sample Conc	the concentration in the original urine	ng/ml	numeric

kable(Table2,caption="8-isoprostane data set")

##Urine Info Data Set The urine info is the data set that has information about subjects characteristics and average pollutant exposure (NO2, SO2, O3, and PM2.5)over the periods of 12 hours, 24 hours, one week, and two weeks. It also includes average temperature and humidity over the periods of 12 hours, 24 hours, one week, and two weeks.

Column name	Meaning	Units	Class of the data
ample_ID	the identity number of the samples	NA	integer
SubjectID	the identity number of the subjects	NA	integer
COLD	cold (represent respiratory infection)	NA	category
MNST	menstration during visit	NA	category
last.ate	the hours to the last meal	hours	integer
wkday.start	the day that the subject start their work	NA	category
dt_smoke	second-hand smoke exposure in hours	hours	numeric
USG urine	specific gravity	g/ml	numeric
o3exp.12h	the exposure of ozone in 12h	ug/m3	numeric
pmexp.12h	the exposure of PM2.5 in 12h	ug/m3	numeric
no2exp.12h	the exposure of NO2 in 12h	ug/m3	numeric
so2exp.12h	the exposure of SO2 in 12h	ug/m3	numeric
Temp.12h	temperature in 12h	ug/m3	numeric
RHx.12h	humidity in 12h	ug/m3	numeric

```
kable(Table3,caption="Urine info data set")
```

2.1 Data Wrangling

My data wrangling started with loading the data set of 8-isoprostane and changing the column names. Then I merged this data set with the urine list to match the sample ID with subjects ID and visit. Then this merged data was merged with the urine info data set using the subject ID and visit as the matching keys. After the second merging, the rows with NAs were removed. Then I calculated the normalized 8-isoprostane using the specific urine gravity. This normalization can adjust the dilution in the urine. I obtained my final data by selecting columns that are relevant to my research question. Finally, the processed data was saved into a processed data folder.

1. Data wrangling

```
# Set working directory  
getwd()
```

```
## [1] "C:/Users/26059/OneDrive/Desktop/ENV 872 R/Yang_ENV872/directory_yang"
```

```
# Load packages  
library(tidyverse)  
library(lmerTest)  
library(MuMIn)  
library(car)  
library(tidyverse)  
library(cowplot)  
library("stargazer")  
library(knitr)
```

```
# Set ggplot theme  
mytheme <- theme_classic(base_size = 25) +  
  theme(axis.text = element_text(color = "black"),  
        legend.position = "top")  
theme_set(mytheme)
```

```
# Load dataset 1  
is <- read.csv("raw_data/8iso.csv")  
is <- select(is, Sample.ID, Calculated.Conc, Sample.Conc)  
colnames(is)[colnames(is) == "Calculated.Conc"] <- "is.origin"  
colnames(is)[colnames(is) == "Sample.Conc"] <- "is.conc"
```

```
# Load dataset 2  
list <- read.csv("raw_data/urine_list.csv")
```

```
#merge  
merge1 <- merge(x=is,y=list,by="Sample.ID",all=TRUE)  
  
urine.info <- read.csv("raw_data/urine_info.csv")  
merge3 <- merge(x=merge1,y=urine.info,by=c("Subject.ID","visit"),all=TRUE)  
dim(merge3)
```

```
## [1] 344 69
```



```
full.data <- na.omit(merge3)
dim(full.data)
```

```
## [1] 307 69
```

```
mean.cre<- mean(full.data$USG)
mean.cre
```

```
## [1] 1.016117
```

```
#calculate normalized biomarker concentration
full.data$is.ad <- (full.data$is.conc/(1-full.data$USG))*(1-mean.cre)

#select useful columns
final.dat<- select(full.data, Sample.ID, Subject.ID, is.ad, is.conc, is.origin, gender,

#save processed data set
write.csv(final.dat, row.names = FALSE, file = "processed_data/final.dat.csv")
```

```
##Explore the data
```

```
#load processed data set
dat <- read.csv("processed_data/final.dat.csv")
#remove outliers
summary(dat$is.origin)
```

```
##      Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
##  1.508   6.220   13.421   45.703   38.281  605.614
```

```
dat1<- dat[-c(4,7,89,147,167),]
```

```
#test normality
shapiro.test(dat1$is.ad)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  dat1$is.ad
## W = 0.58916, p-value < 2.2e-16
```

```
shapiro.test(log(dat1$is.ad))
```

```
##
##  Shapiro-Wilk normality test
##
## data:  log(dat1$is.ad)
## W = 0.98127, p-value = 0.0005511
```

```
#explore the data
```

```
##
## Summary of Statistics
## -----
## Statistic          N   Mean  St. Dev.   Min    Pctl(25) Pctl(75)   Max
## =====
## 12-hour NO2 exposure 302 24.114  8.158    12.124  19.959    27.436   41.964
## 24-hour NO2 exposure 302 23.720  6.827    12.373  18.405    27.049   39.351
## 1-week NO2 exposure  302 24.090  4.913    16.562  19.243    27.869   35.193
## 2-week NO2 exposure  302 25.489  2.542    20.805  23.999    26.755   32.173
## =====
```

Table 4: Summary of NO2 exposure

x

Summary of Statistics

Statistic	N	Mean	St. Dev.	Min	Pctl(25)	Pctl(75)	Max
12-hour NO2 exposure	302	24.114	8.158	12.124	19.959	27.436	41.964
24-hour NO2 exposure	302	23.720	6.827	12.373	18.405	27.049	39.351
1-week NO2 exposure	302	24.090	4.913	16.562	19.243	27.869	35.193
2-week NO2 exposure	302	25.489	2.542	20.805	23.999	26.755	32.173

```
#calculate IQR for each period
no2.12 <-no2.sum1[5]-no2.sum1[2]
no2.12
```

```
## 3rd Qu.
## 7.476984
```

```
no2.24 <-no2.sum2[5]-no2.sum2[2]
no2.24
```

```
## 3rd Qu.
## 8.643751
```

```
no2.1w <-no2.sum3[5]-no2.sum3[2]
no2.1w
```

```
## 3rd Qu.
## 8.626535
```

```
no2.2w <-no2.sum4[5]-no2.sum4[2]
no2.2w
```

```
## 3rd Qu.
## 2.755933
```

3.Build models

3.1 Build models for 12-hour NO2 exposure

```
mo1<- lmer(log(is.ad) ~no2exp.12h+o3exp.12h+ pmexp.12h+so2exp.12h+Temp.12h+RHx.12h + las
#find the best model with lowest AIC
step(mo1)
```

```
## Backward reduced random-effect table:
```

```
##
##               Eliminated npar  logLik      AIC      LRT Df Pr(>Chisq)
## <none>                16 -474.63   981.26
## (1 | Subject.ID)      0   15 -493.94 1017.88 38.616   1 5.159e-10 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
## Backward reduced fixed-effect table:
## Degrees of freedom method: Satterthwaite
```

```
##
##               Eliminated  Sum Sq Mean Sq NumDF  DenDF F value  Pr(>F)
## Temp.12h           1    0.0261  0.0261     1 276.301  0.0314 0.859411
## COLD                2    0.0677  0.0677     1 266.912  0.0819 0.774976
## pmexp.12h          3    0.0956  0.0956     1 271.141  0.1159 0.733779
## go.home            4    0.6585  0.3293     2  83.839  0.4012 0.670812
```

```
## wkday.start      5  0.3986  0.3986      1 291.709  0.4848 0.486792
## so2exp.12h       6  0.4401  0.4401      1 293.944  0.5343 0.465386
## MNST             7  0.8942  0.8942      1 272.520  1.0861 0.298261
## last.ate         8  0.8829  0.8829      1 276.178  1.0773 0.300216
## dt_smoke         9  0.9658  0.9658      1 293.061  1.1845 0.277332
## RHx.12h          10 2.0108  2.0108      1 251.786  2.4885 0.115939
## no2exp.12h       0 12.2057 12.2057      1 228.742 14.9910 0.000141 ***
## o3exp.12h        0  4.0334  4.0334      1 225.577  4.9538 0.027023 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Model found:
## log(is.ad) ~ no2exp.12h + o3exp.12h + (1 | Subject.ID)
```

```
#best model for 12-hour NO2 exposure
mo1.1<- lmer(log(is.ad) ~ no2exp.12h +o3exp.12h + (1 | Subject.ID),data=dat1)
summary(mo1.1)
```

```
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: log(is.ad) ~ no2exp.12h + o3exp.12h + (1 | Subject.ID)
## Data: dat1
##
## REML criterion at convergence: 910.8
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -1.99963 -0.62729 -0.03595  0.60429  2.58701
##
## Random effects:
## Groups      Name                Variance Std.Dev.
## Subject.ID (Intercept) 0.5187     0.7202
## Residual              0.8142     0.9023
## Number of obs: 302, groups: Subject.ID, 89
##
## Fixed effects:
##              Estimate Std. Error      df t value Pr(>|t|)
## (Intercept) -0.187878   0.212372 292.646414  -0.885 0.377064
## no2exp.12h   0.025912   0.006693 228.741989   3.872 0.000141 ***
## o3exp.12h    0.031066   0.013958 225.576673   2.226 0.027023 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
```

```
##          (Intr) n2x.12
## no2exp.12h -0.812
## o3exp.12h  -0.489  0.132
```

```
#check F value
anova(mo1.1)
```

```
## Type III Analysis of Variance Table with Satterthwaite's method
##          Sum Sq Mean Sq NumDF  DenDF F value    Pr(>F)
## no2exp.12h 12.2057 12.2057     1 228.74 14.9910 0.000141 ***
## o3exp.12h   4.0334  4.0334     1 225.58  4.9538 0.027023 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
#check collinearity
vif(mo1.1)
```

```
## no2exp.12h  o3exp.12h
##   1.017673   1.017673
```

```
#get R2
r.squaredGLMM (mo1.1)
```

```
##          R2m      R2c
## [1,] 0.03761012 0.412126
```

```
#identify outliers
res1.1 <- resid(mo1, type = "pearson")
dat[which(abs(res1.1) > 2.5),]
```

```
##  [1] Sample.ID   Subject.ID   is.ad       is.conc     is.origin   gender
##  [7] COLD         MNST        last.ate    wkday.start go.home     dt_smoke
## [13] o3exp.12h    pmexp.12h   no2exp.12h  so2exp.12h  Temp.12h    RHx.12h
## [19] o3exp.24h    pmexp.24h   no2exp.24h  so2exp.24h  Temp.24h    RHx.24h
## [25] o3exp.1w     pmexp.1w    no2exp.1w   so2exp.1w   Temp.1w     RHx.1w
## [31] o3exp.2w     pmexp.2w    no2exp.2w   so2exp.2w   Temp.2w     RHx.2w
## <0 rows> (or 0-length row.names)
```

```
#calculate IQR change
(exp(0.025912)-1)*no2.12
```

```
## 3rd Qu.
## 0.1962756
```

```
(exp(0.025912 +0.006693)-1)*no2.12
```

```
## 3rd Qu.  
## 0.2478049
```

```
(exp(0.025912 -0.006693)-1)*no2.12
```

```
## 3rd Qu.  
## 0.1450899
```

3.2 Build models for 24-hour NO2 exposure

```
mo2<- lmer(log(is.ad) ~ no2exp.24h+o3exp.24h+ pmexp.24h+so2exp.24h+Temp.24h+RHx.24h + last.ate, data=mo2, REML=FALSE)  
step(mo2)
```

```
## Backward reduced random-effect table:
```

```
##  
##           Eliminated npar logLik      AIC      LRT Df Pr(>Chisq)  
## <none>                16 -475.04  982.07  
## (1 | Subject.ID)      0  15 -493.92 1017.84 37.768  1 7.969e-10 ***  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##
```

```
## Backward reduced fixed-effect table:
```

```
## Degrees of freedom method: Satterthwaite
```

```
##  
##           Eliminated Sum Sq Mean Sq NumDF DenDF F value Pr(>F)  
## so2exp.24h          1  0.0016  0.0016    1 280.151  0.0019 0.9655516  
## Temp.24h            2  0.0161  0.0161    1 277.135  0.0193 0.8895889  
## wkday.start         3  0.0126  0.0126    1 279.050  0.0152 0.9021154  
## COLD                4  0.1555  0.1555    1 268.565  0.1873 0.6655198  
## go.home             5  0.9235  0.4617    2  84.443  0.5575 0.5747522  
## RHx.24h             6  0.4390  0.4390    1 258.415  0.5298 0.4673562  
## dt_smoke            7  0.5832  0.5832    1 291.289  0.7037 0.4022322  
## pmexp.24h           8  0.8868  0.8868    1 273.549  1.0799 0.2996466  
## MNST               9  0.8850  0.8850    1 274.024  1.0795 0.2997231  
## last.ate          10  1.0377  1.0377    1 281.175  1.2708 0.2605860  
## no2exp.24h          0 12.3002 12.3002    1 231.652 15.1608 0.0001292 ***  
## o3exp.24h           0  5.0095  5.0095    1 225.520  6.1746 0.0136869 *  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##
```

```
## Model found:
```

```
## log(is.ad) ~ no2exp.24h + o3exp.24h + (1 | Subject.ID)
```

```
mo2.1<- lmer(log(is.ad) ~ no2exp.24h + o3exp.24h + (1 | Subject.ID),data=dat1)
summary(mo2.1)
```

```
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: log(is.ad) ~ no2exp.24h + o3exp.24h + (1 | Subject.ID)
## Data: dat1
##
## REML criterion at convergence: 910.4
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -1.88882 -0.65529 -0.03642  0.58790  2.66141
##
## Random effects:
## Groups      Name      Variance Std.Dev.
## Subject.ID (Intercept) 0.5243   0.7241
## Residual              0.8113   0.9007
## Number of obs: 302, groups: Subject.ID, 89
##
## Fixed effects:
##              Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)  -0.339832   0.243838 284.149326  -1.394 0.164503
## no2exp.24h    0.031555   0.008104 231.652052   3.894 0.000129 ***
## o3exp.24h     0.030798   0.012394 225.520144   2.485 0.013687 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) n2x.24
## no2exp.24h  -0.857
## o3exp.24h   -0.505  0.196
```

```
anova(mo2.1)
```

```
## Type III Analysis of Variance Table with Satterthwaite's method
##              Sum Sq Mean Sq NumDF  DenDF F value    Pr(>F)
## no2exp.24h 12.3002 12.3002      1 231.65 15.1608 0.0001292 ***
## o3exp.24h   5.0095  5.0095      1 225.52  6.1746 0.0136869 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
vif(mo2.1)
```

```
## no2exp.24h o3exp.24h  
## 1.039908 1.039908
```

```
r.squaredGLMM(mo2.1)
```

```
## R2m R2c  
## [1,] 0.03854593 0.4159847
```

```
res2.1 <- resid(mo2.1, type = "pearson")  
dat[which(abs(res2.1) > 2.5),]
```

```
## [1] Sample.ID Subject.ID is.ad is.conc is.origin gender  
## [7] COLD MNST last.ate wkday.start go.home dt_smoke  
## [13] o3exp.12h pmexp.12h no2exp.12h so2exp.12h Temp.12h RHx.12h  
## [19] o3exp.24h pmexp.24h no2exp.24h so2exp.24h Temp.24h RHx.24h  
## [25] o3exp.1w pmexp.1w no2exp.1w so2exp.1w Temp.1w RHx.1w  
## [31] o3exp.2w pmexp.2w no2exp.2w so2exp.2w Temp.2w RHx.2w  
## <0 rows> (or 0-length row.names)
```

```
#calculate IQR change  
(exp(0.031555)-1)*no2.24
```

```
## 3rd Qu.  
## 0.2771025
```

```
(exp(0.031555 +0.008104)-1)*no2.24
```

```
## 3rd Qu.  
## 0.3496909
```

```
(exp(0.031555 -0.008104)-1)*no2.24
```

```
## 3rd Qu.  
## 0.2051001
```

3.3 Build models for 1-week NO2 exposure


```
mo3<- lmer(log(is.ad) ~ no2exp.1w+o3exp.1w+ pmexp.1w+so2exp.1w+Temp.1w+RHx.1w + last.ate
step(mo3)
```

```
## Backward reduced random-effect table:
```

```
##
##               Eliminated npar  logLik      AIC      LRT Df Pr(>Chisq)
## <none>                16 -475.45   982.89
## (1 | Subject.ID)      0   15 -493.45 1016.90 36.001   1  1.972e-09 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
## Backward reduced fixed-effect table:
```

```
## Degrees of freedom method: Satterthwaite
```

```
##
##               Eliminated Sum Sq Mean Sq NumDF   DenDF F value    Pr(>F)
## Temp.1w          1 0.0000  0.0000     1 256.639  0.0000 0.9986105
## RHx.1w           2 0.0271  0.0271     1 288.772  0.0317 0.8588147
## COLD             3 0.0560  0.0560     1 271.043  0.0657 0.7979384
## wkday.start      4 0.1240  0.1240     1 258.909  0.1459 0.7027697
## go.home          5 0.8717  0.4359     2  84.614  0.5134 0.6002845
## pmexp.1w         6 0.2371  0.2371     1 293.316  0.2793 0.5975635
## dt_smoke         7 0.4455  0.4455     1 291.942  0.5270 0.4684347
## last.ate         8 0.6026  0.6026     1 277.597  0.7195 0.3970452
## MNST             9 0.7572  0.7572     1 273.657  0.9102 0.3409085
## o3exp.1w        10 1.4655  1.4655     1 229.000  1.7684 0.1849006
## no2exp.1w        0 4.4044  4.4044     1 229.706  5.2961 0.0222690 *
## so2exp.1w        0 9.8969  9.8969     1 230.979 11.9006 0.0006669 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
## Model found:
```

```
## log(is.ad) ~ no2exp.1w + so2exp.1w + (1 | Subject.ID)
```

```
mo3.1<- lmer(log(is.ad) ~ no2exp.1w + so2exp.1w + (1 | Subject.ID),data=dat1)
summary(mo3.1)
```

```
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
```

```
## lmerModLmerTest]
```

```
## Formula: log(is.ad) ~ no2exp.1w + so2exp.1w + (1 | Subject.ID)
```

```
## Data: dat1
```

```
##
```

```
## REML criterion at convergence: 913.2
```

```
##
```

```
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -1.95288 -0.61932 -0.07413  0.62697  2.63863
##
## Random effects:
##   Groups      Name      Variance Std.Dev.
##   Subject.ID (Intercept) 0.5199   0.7211
##   Residual                0.8316   0.9119
## Number of obs: 302, groups: Subject.ID, 89
##
## Fixed effects:
##              Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)   0.76276    0.29669 261.12832   2.571 0.010698 *
## no2exp.1w     0.02877    0.01250 229.70566   2.301 0.022269 *
## so2exp.1w    -0.12342    0.03578 230.97927  -3.450 0.000667 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) n2xp.1
## no2exp.1w  -0.624
## so2exp.1w  -0.331 -0.476
```

```
anova(mo3.1)
```

```
## Type III Analysis of Variance Table with Satterthwaite's method
##              Sum Sq Mean Sq NumDF  DenDF F value    Pr(>F)
## no2exp.1w  4.4044  4.4044     1 229.71  5.2961 0.0222690 *
## so2exp.1w  9.8969  9.8969     1 230.98 11.9006 0.0006669 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
vif(mo3.1)
```

```
## no2exp.1w so2exp.1w
##  1.293508  1.293508
```

```
r.squaredGLMM (mo3.1)
```

```
##              R2m      R2c
## [1,] 0.02719595 0.4014215
```

```
#calculate IQR change
(exp( 0.02877)-1)*no2.1w
```

```
## 3rd Qu.
## 0.25179
```

```
(exp( 0.02877 +0.01250)-1)*no2.1w
```

```
## 3rd Qu.
## 0.3634656
```

```
(exp( 0.02877 -0.01250)-1)*no2.1w
```

```
## 3rd Qu.
## 0.1415017
```

3.4 Build models for 2-week NO2 exposure

```
mo4<- lmer(log(is.ad) ~ no2exp.2w+o3exp.2w+ pmexp.2w+so2exp.2w+Temp.2w+RHx.2w + last.ate
summary(mo4)
```

```
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: log(is.ad) ~ no2exp.2w + o3exp.2w + pmexp.2w + so2exp.2w + Temp.2w +
##      RHx.2w + last.ate + wkday.start + go.home + COLD + MNST +
##      dt_smoke + (1 | Subject.ID)
##      Data: dat1
##
## REML criterion at convergence: 946.3
##
## Scaled residuals:
##      Min      1Q   Median      3Q      Max
## -2.03734 -0.65274 -0.04204  0.61324  2.64864
##
## Random effects:
##      Groups      Name      Variance Std.Dev.
## Subject.ID (Intercept) 0.4961   0.7043
## Residual              0.8617   0.9283
## Number of obs: 302, groups: Subject.ID, 89
##
## Fixed effects:
```

	Estimate	Std. Error	df	t value	Pr(> t)
## (Intercept)	-0.293745	2.831683	283.544175	-0.104	0.917
## no2exp.2w	0.066813	0.056690	283.012735	1.179	0.240
## o3exp.2w	0.040363	0.074186	278.563880	0.544	0.587
## pmexp.2w	-0.002124	0.005735	284.434919	-0.370	0.711
## so2exp.2w	-0.085834	0.104764	286.681042	-0.819	0.413
## Temp.2w	0.023248	0.120432	279.830238	0.193	0.847
## RHx.2w	-0.009879	0.036280	285.735687	-0.272	0.786
## last.ate	0.010671	0.013349	280.291066	0.799	0.425
## wkday.start	0.042621	0.084409	190.300335	0.505	0.614
## go.homewed	-0.252430	0.302338	79.849719	-0.835	0.406
## go.homeweekend	-0.130340	0.343381	80.299526	-0.380	0.705
## COLDY	0.058713	0.204809	270.397059	0.287	0.775
## MNSTY	-0.482831	0.490378	266.959455	-0.985	0.326
## dt_smoke	-0.022624	0.032845	284.361134	-0.689	0.492

step(mo4)

Backward reduced random-effect table:

	Eliminated	npar	logLik	AIC	LRT	Df	Pr(>Chisq)				
## <none>		16	-473.16	978.33							
## (1 Subject.ID)	0	15	-490.95	1011.90	35.572	1	2.458e-09 ***				
## ---											
## Signif. codes:	0	'***'	0.001	'**'	0.01	'*'	0.05	'.'	0.1	' '	1

Backward reduced fixed-effect table:

Degrees of freedom method: Satterthwaite

	Eliminated	Sum Sq	Mean Sq	NumDF	DenDF	F value	Pr(>F)				
## Temp.2w	1	0.0321	0.0321	1	279.830	0.0373	0.8470709				
## COLD	2	0.0678	0.0678	1	271.212	0.0790	0.7788488				
## pmexp.2w	3	0.0718	0.0718	1	280.831	0.0839	0.7722626				
## RHx.2w	4	0.1283	0.1283	1	259.795	0.1506	0.6982381				
## go.home	5	0.5742	0.2871	2	83.645	0.3385	0.7138446				
## o3exp.2w	6	0.1074	0.1074	1	270.074	0.1266	0.7222560				
## dt_smoke	7	0.4478	0.4478	1	292.437	0.5298	0.4672622				
## wkday.start	8	0.5832	0.5832	1	263.734	0.6961	0.4048516				
## last.ate	9	0.7859	0.7859	1	278.303	0.9379	0.3336707				
## MNST	10	0.8190	0.8190	1	274.894	0.9847	0.3219216				
## no2exp.2w	0	4.5348	4.5348	1	221.626	5.4726	0.0202064 *				
## so2exp.2w	0	10.7883	10.7883	1	222.380	13.0193	0.0003808 ***				
## ---											
## Signif. codes:	0	'***'	0.001	'**'	0.01	'*'	0.05	'.'	0.1	' '	1

```
##
## Model found:
## log(is.ad) ~ no2exp.2w + so2exp.2w + (1 | Subject.ID)

mo4.1<- lmer(log(is.ad) ~ no2exp.2w + so2exp.2w + (1 | Subject.ID),data=dat1)
summary(mo4.1)

## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: log(is.ad) ~ no2exp.2w + so2exp.2w + (1 | Subject.ID)
## Data: dat1
##
## REML criterion at convergence: 910.7
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.07871 -0.63093 -0.03594  0.59131  2.66937
##
## Random effects:
## Groups      Name      Variance Std.Dev.
## Subject.ID (Intercept) 0.5184   0.7200
## Residual              0.8286   0.9103
## Number of obs: 302, groups: Subject.ID, 89
##
## Fixed effects:
##              Estimate Std. Error      df t value Pr(>|t|)
## (Intercept)   0.07802    0.54464 232.36291   0.143 0.886220
## no2exp.2w     0.05572    0.02382 221.62605   2.339 0.020206 *
## so2exp.2w    -0.13645    0.03782 222.37980  -3.608 0.000381 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              (Intr) n2xp.2
## no2exp.2w  -0.904
## so2exp.2w   0.079 -0.471

anova(mo4.1)

## Type III Analysis of Variance Table with Satterthwaite's method
##              Sum Sq Mean Sq NumDF  DenDF F value    Pr(>F)
## no2exp.2w  4.5348  4.5348     1 221.63  5.4726 0.0202064 *
## so2exp.2w 10.7883 10.7883     1 222.38 13.0193 0.0003808 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
vif(mo4.1)
```

```
## no2exp.2w so2exp.2w  
## 1.284372 1.284372
```

```
r.squaredGLMM (mo4.1)
```

```
## R2m R2c  
## [1,] 0.02811885 0.4021322
```

```
#calculate IQR change  
(exp(0.05572)-1)*no2.2w
```

```
## 3rd Qu.  
## 0.1579193
```

```
(exp(0.05572 +0.02382)-1)*no2.2w
```

```
## 3rd Qu.  
## 0.2281605
```

```
(exp(0.05572 -0.02382)-1)*no2.2w
```

```
## 3rd Qu.  
## 0.08933151
```

3 Exploratory Analysis

The Shapiro tests showed that concentrations of urinary 8-isoprostane are not normalized (p-value < 2.2e-16). The Shapiro tests showed that concentrations of log-scale urinary 8-isoprostane are still not normalized (p-value < 2.2e-16). To build linear models, I need to use the more normalized log-transformed concentrations of urinary 8-isoprostane. The distribution of data in log(8-isoprostane) (Figure 2) looks more normalized than the distribution of data of 8-isoprostane in (Figure 1).

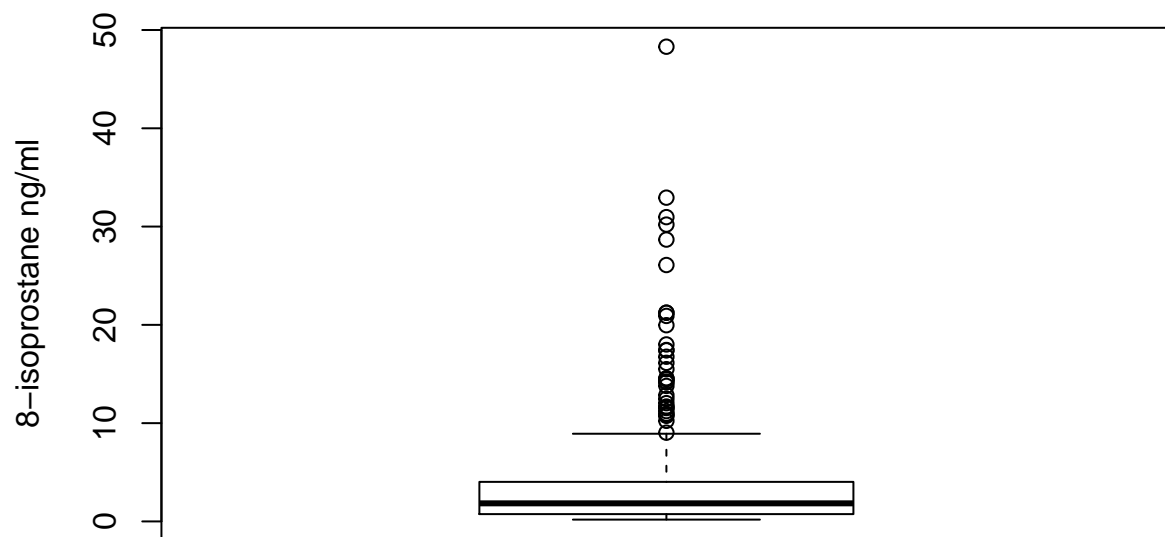


Figure 1: Plot of 8-isoprostane

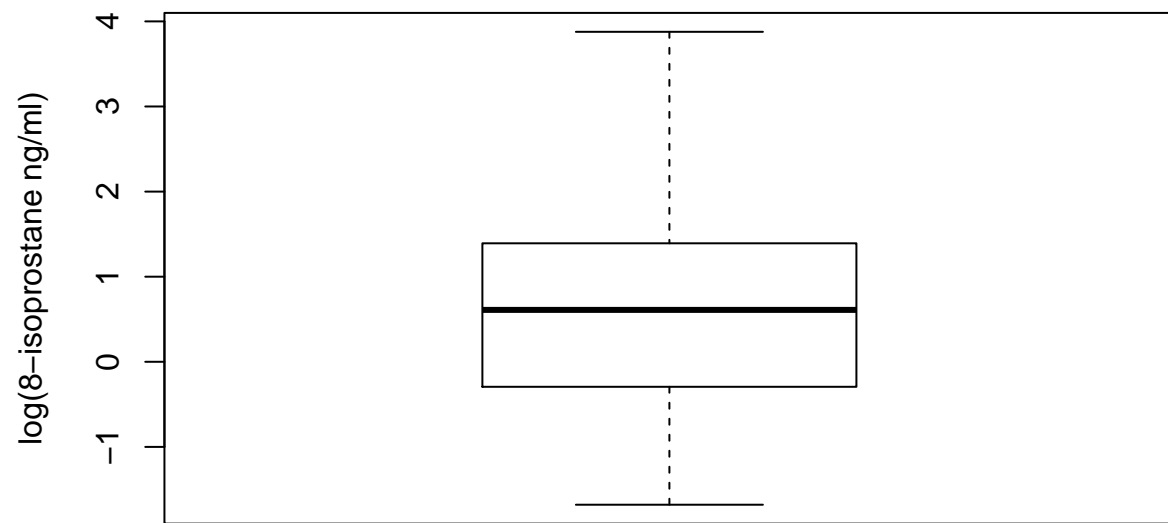


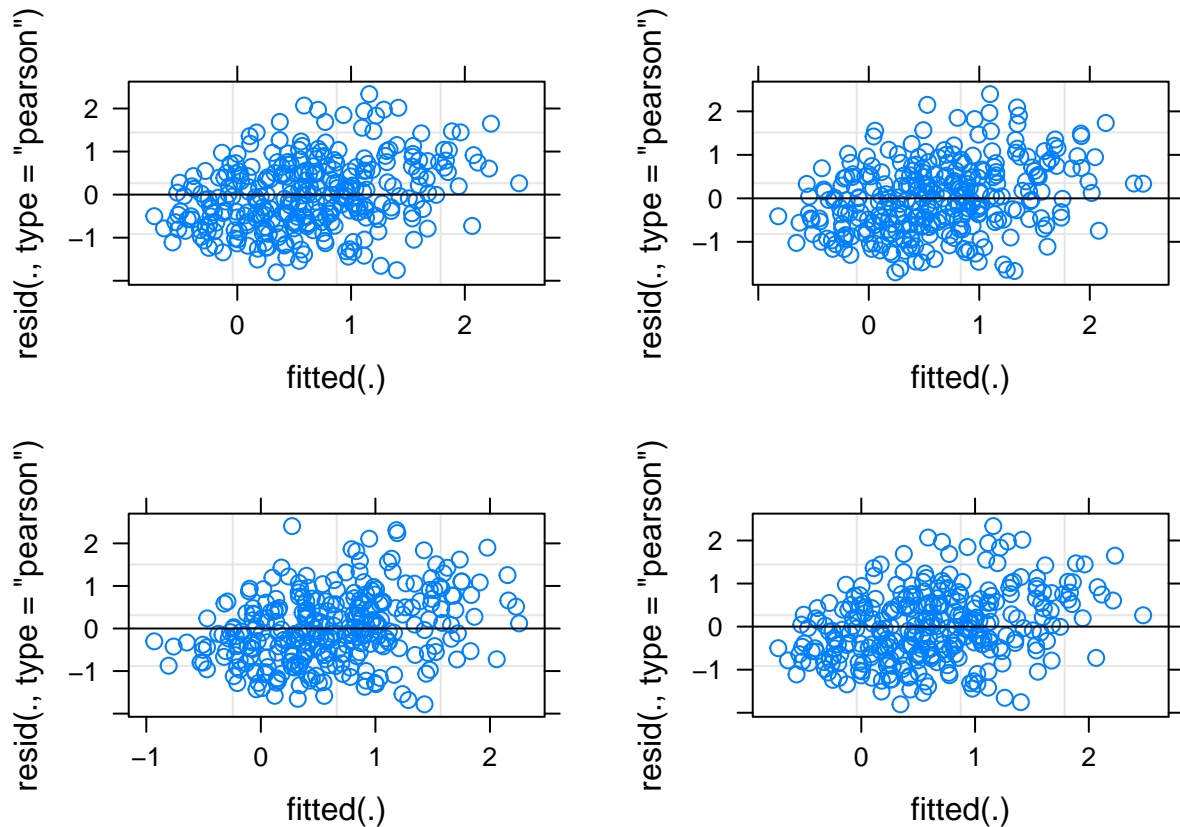
Figure 2: Plot of $\log(8\text{-isoprostane})$

4 Analysis

I used linear mixed models with participant-specific intercepts. The inclusion of participants as a variable in mixed-effects models account for the correlation of repeated measurements from the same individuals and precludes the need to control for participant characteristics (e.g., age, gender, BMI, the identity of smoker or non-smoker) that do not change across the four longitudinal measurements.

I used NO₂ exposure as predicting variables and log-transformed urinary biomarkers as dependent variables. For each biomarker, I built 4 models to the exposure over the periods of 12h, 24h, one week, and two weeks. I used a backward stepwise model selection method to select the co-variables for each model. The co-variables that I tested in the models were last meal, the start of the workday, respiratory infection status, menstruation, and the time of second-hand exposure. I also tested the SO₂ exposure, ozone exposure and PM_{2.5} exposure, the average temperature and humidity during a corresponding period.

With the stepAIC function, the best models were chosen with the lowest AIC. The vif tests show that there is no model with excess intercorrelation in the predicting variables. Figure 3 shows that all four diagnostic plots have the shape of a triangle. The higher the fitted value is, the higher the residue is. This residue is a result of skewed distribution of log scaled of the predicting variable. In general, the residue distribution is not bad. No value is bigger than an absolute value of 2.5.



Both 12-hour, 24-hour, 1-week and 2-week NO₂ exposure showed significant correlations

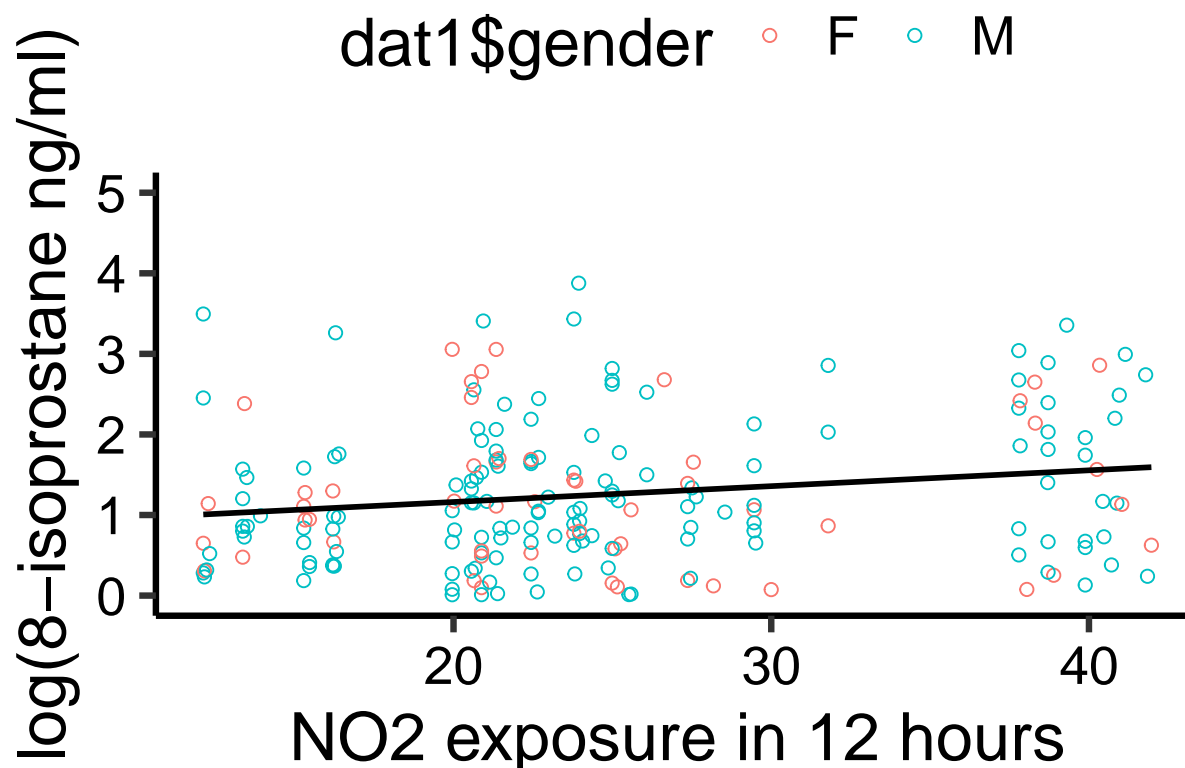


Figure 3: Relationship between log(8-isoprostane) and 12-hour NO2 exposure

with the level of urinary 8-isoprostane, especially the 2-week NO2 exposure.

4.1 Result 1

There is a significant positive relationship between 12-hour NO2 exposure and log(8-isoprostane)(Figure 7).One IQR (7.41 $\mu\text{g}/\text{m}^3$) incremental change of 12-h NO2 exposure was associated with an increase of 8-isoprostane level by 19.45% (95% CI: 14.37%, 24.55%, F value=14.99, p-value <0.001). The Adjusted R-squared =0.4121, which is the fraction of total variance explained by the model

4.2 Result 2

There is a significant positive relationship between 24-hour NO2 exposure and log(8-isoprostane)(Figure 8).One IQR (8.64 $\mu\text{g}/\text{m}^3$) incremental change of 24-h NO2 exposure was associated an increase in 8-isoprostane level by 27.69% (95% CI: 20.50%, 34.95% , F value=15.16, p-value <0.001). The Adjusted R-squared = 0.4160, which is the fraction of total variance explained by the model.

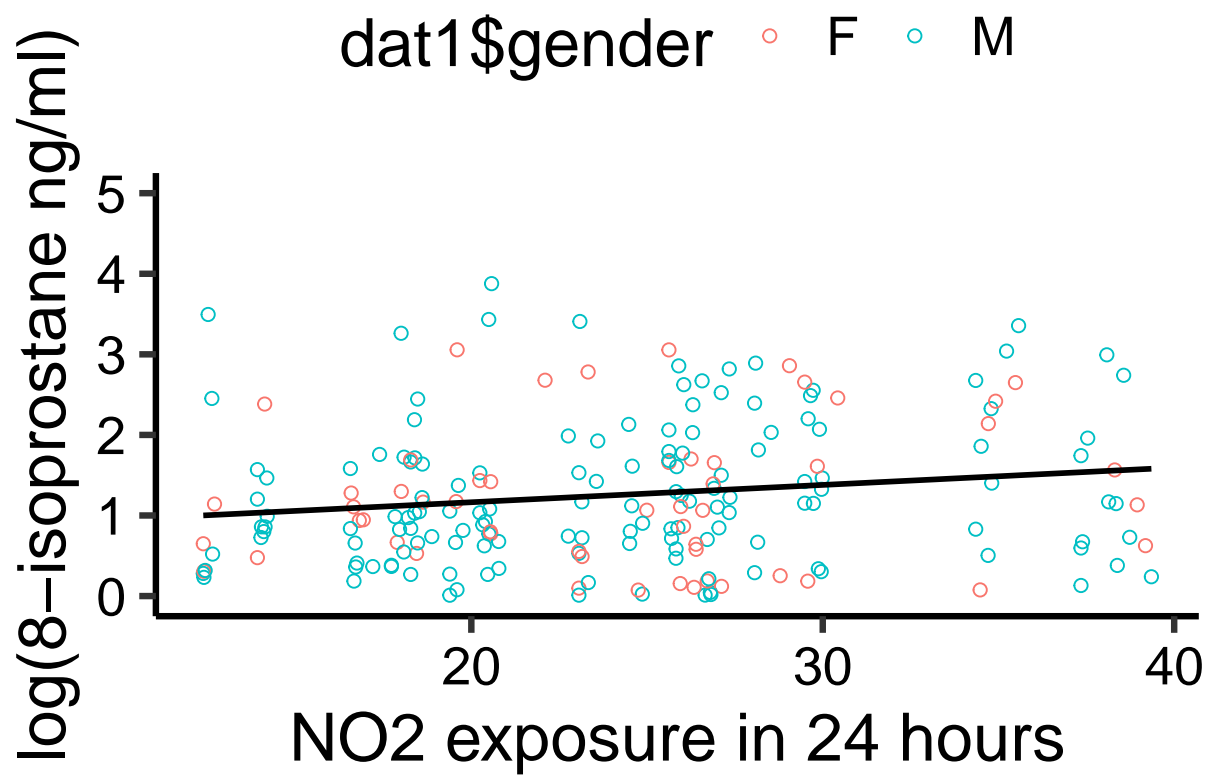


Figure 4: Relationship between log(8-isoprostane) and 24-hour NO2 exposure

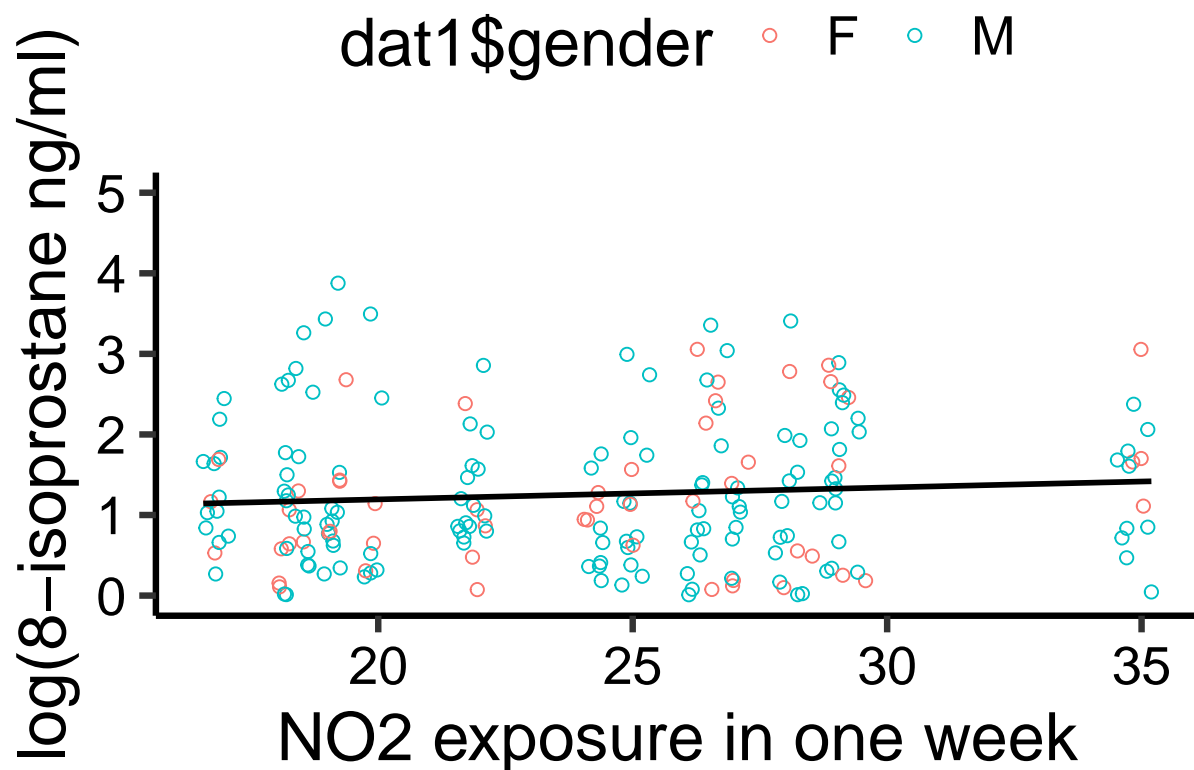


Figure 5: Relationship between log(8-isoprostane) and 1-week NO2 exposure

4.3 Result 3

There is a significant positive relationship between 1-week NO2 exposure and log(8-isoprostane)(Figure 9).One IQR (4.47 $\mu\text{g}/\text{m}^3$) incremental change of 1-week NO2 exposure was associated with an increase in 8-isoprostane level by 25.15% (95% CI: 14.14%,36.31%, F value=5.30, p-value <0.05). The Adjusted R-squared = 0.4014, which is the fraction of total variance explained by the model.

4.4 Result 4

There is a significant positive relationship between 2-week NO2 exposure and log(8-isoprostane)(Figure 10).One IQR (2.74 $\mu\text{g}/\text{m}^3$) incremental change of 2-week NO2 exposure was associated with an increase in 8-isoprostane level by 15.68% (95% CI: 8.87%,22.65%, F value=5.47, p-value <0.05). The Adjusted R-squared = 0.4021, which is the fraction of total variance explained by the model.

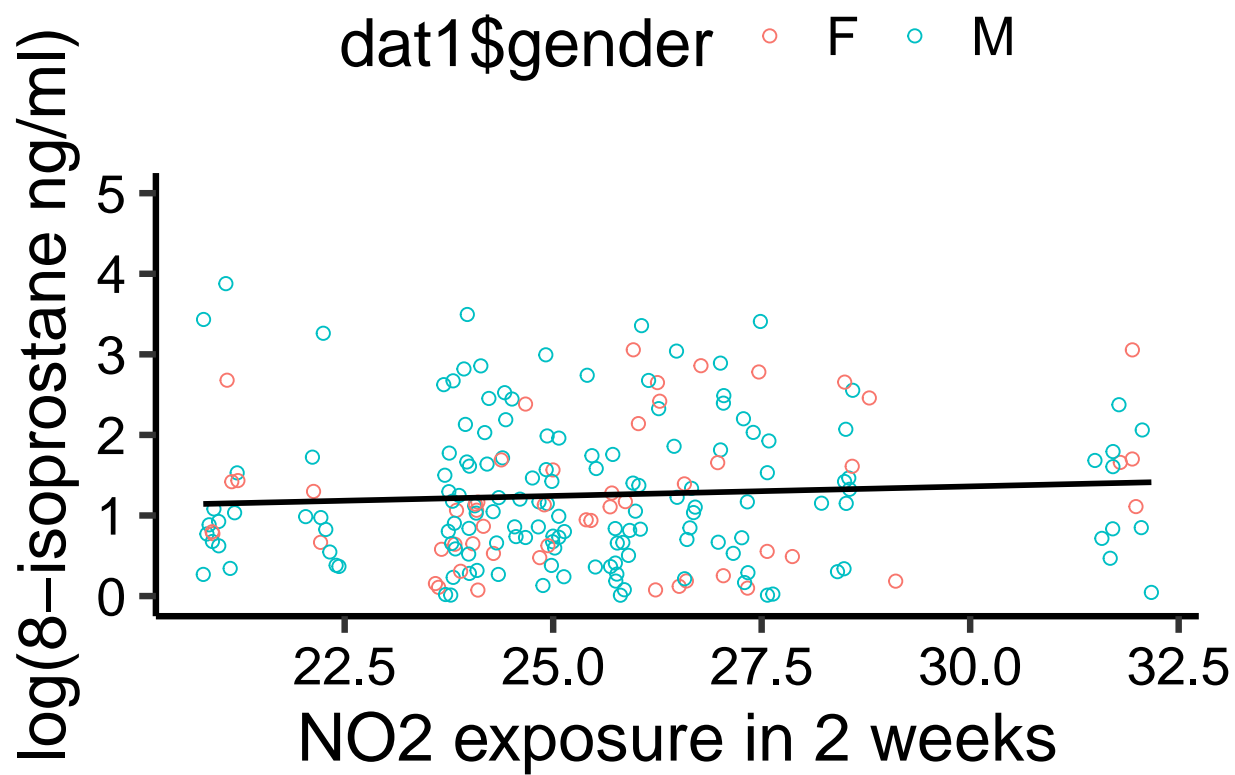


Figure 6: Relationship between log(8-isoprostane) and 2-week NO2 exposure

5 Summary and Conclusions

Short term NO₂ exposure was associated with urinary 8-isoprostane. This finding meet my hypothesis. Urinary 8-isoprostane indicates the lipid peroxidation from the whole body, reflecting the systemic oxidative stress. It indicates short term NO₂ exposure can cause signifiant higher systemic oxidative stress.

6 References

Mölter, A., Agius, R. M., de Vocht, F., Lindley, S., Gerrard, W., Lowe, L., ... & Simpson, A. (2013). Long-term exposure to PM10 and NO2 in association with lung volume and airway resistance in the MAAS birth cohort. *Environmental health perspectives*, 121(10), 1232-1238. Collart, P., Dubourg, D., Levêque, A., Sierra, N. B., & Coppieters, Y. (2018). Short-term effects of nitrogen dioxide on hospital admissions for cardiovascular disease in Wallonia, Belgium. *International journal of cardiology*, 255, 231-236. Takenoue, Y., Kaneko, T., Miyamae, T., Mori, M., & Yokota, S. (2012). Influence of outdoor NO2 exposure on asthma in childhood: Meta-analysis. *Pediatrics International*, 54(6), 762-769. Hashemzadeh, B., Idani, E., Goudarzi, G., Ankali, K. A., Sakhvidi, M. J. Z., Babaei, A. A., ... & Neisi, A. (2019). Effects of PM2. 5 and NO2 on the 8-isoprostane and lung function indices of FVC and FEV1 in students of Ahvaz city, Iran. *Saudi journal of biological sciences*, 26(3), 473-480. Black, C. N., Bot, M., Révész, D., Scheffer, P. G., & Penninx, B. (2017). The association between three major physiological stress systems and oxidative DNA and lipid damage. *Psychoneuroendocrinology*, 80, 56-66. Lakshmi, S. V., Padmaja, G., Kuppusamy, P., & Kutala, V. K. (2009). Oxidative stress in cardiovascular disease. Zanolin, M. E., Chamitava, L., Degan, P., Pasini, A., Fratta-Pasini, A., Nicolis, M., ... & De Marco, R. (2015). Biomarkers of oxidative stress in chronic respiratory diseases. Danielsen, P. H., Loft, S., Kocbach, A., Schwarze, P. E., & Møller, P. (2009). Oxidative damage to DNA and repair induced by Norwegian wood smoke particles in human A549 and THP-1 cell lines. *Mutation Research/Genetic Toxicology and Environmental Mutagenesis*, 674(1-2), 116-122. Nuernberg, A. M., Boyce, P. D., Cavallari, J. M., Fang, S. C., Eisen, E. A., & Christiani, D. C. (2008). Urinary 8-isoprostane and 8-OHdG concentrations in boilermakers with welding exposure. *Journal of occupational and environmental medicine*, 50(2), 182-189. Hashemzadeh, B., Idani, E., Goudarzi, G., Ankali, K. A., Sakhvidi, M. J. Z., Babaei, A. A., ... & Neisi, A. (2019). Effects of PM2. 5 and NO2 on the 8-isoprostane and lung function indices of FVC and FEV1 in students of Ahvaz city, Iran. *Saudi journal of biological sciences*, 26(3), 473-480. Patel, M. M., Chillrud, S. N., Deepti, K. C., Ross, J. M., & Kinney, P. L. (2013). Traffic-related air pollutants and exhaled markers of airway inflammation and oxidative stress in New York City adolescents. *Environmental research*, 121, 71-78.