

Therapy

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April 2, 2019

Load data

```
# The data
Y<- cbind(c(28,4,41,12),c(45,12,44,7),c(29,5,20,3),c(26,2,20,1))
colnames(Y)<-c("ProgressiveDisease","NoChange","PartialRemission","CompleteRemission")
therapy <- c("S","S","A","A")
gender <- rep(c("M","F"),2)
```

Baseline Categorical Logit Model

Fit

```
library(nnet)
#Progressive Disease as baseline
b1<-multinom(Y~as.factor(therapy))

## # weights:  12 (6 variable)
## initial  value 414.502014
## iter   10 value 395.641305
## final   value 395.641300
## converged

beta1<-coefficients(b1)
summary(b1)

## Call:
## multinom(formula = Y ~ as.factor(therapy))
##
## Coefficients:
##                (Intercept) as.factor(therapy)S
## NoChange          -0.03846754          0.6157805
## PartialRemission  -0.83479810          0.8954216
## CompleteRemission -0.92576929          0.7922358
##
## Std. Errors:
##                (Intercept) as.factor(therapy)S
## NoChange           0.1961524          0.2954143
## PartialRemission   0.2496922          0.3507248
## CompleteRemission  0.2578506          0.3653097
##
## Residual Deviance: 791.2826
## AIC: 803.2826
```

Results

```
result=predict(b1,type ="prob")
result=cbind(therapy,result)
#predicted probability
result

##   therapy ProgressiveDisease NoChange          PartialRemission
## 1 "S"      "0.211920853535378" "0.37748312784509" "0.225165653963416"
## 2 "S"      "0.211920853535378" "0.37748312784509" "0.225165653963416"
## 3 "A"      "0.358108276712193" "0.344594321659436" "0.155405416074088"
## 4 "A"      "0.358108276712193" "0.344594321659436" "0.155405416074088"
##   CompleteRemission
## 1 "0.185430364656116"
## 2 "0.185430364656116"
## 3 "0.141891985554283"
## 4 "0.141891985554283"

# proportional Therapy S vs A
S <-c(28,4, 45,12,29,5,26,2)
A <-c(41,12,44,7,20,3,20,1)
PS=sum(S)/(sum(S)+sum(A))
PA=sum(A)/(sum(S)+sum(A))
cbind(PS,PA)

##           PS           PA
## [1,] 0.5050167 0.4949833
```

Interpretation

In our model with Progressive Disease as baseline:

Under the alternating therapy treatment, we have $\log(\frac{\pi_N}{\pi_{Prog}}) = \beta_{11} = -0.03846754$; Under the sequential therapy treatment, we have $\log(\frac{\pi_N}{\pi_{Prog}}) = \beta_{11} + \beta_{12} = -0.03846754 + 0.6157805 = 0.577313$.

The log odds ratio: $\log\frac{odd_S}{odd_A} = \log(\frac{\pi_N}{\pi_{Prog}}|S) - \log(\frac{\pi_N}{\pi_{Prog}}|A) = \beta_{12} = 0.6157805$, which suggests the odds of “Progressive Disease” vs “No Disease” with alternating sequential and alternating therapy are $\frac{odd_S}{odd_A} = e^{0.6157805} = 1.851101$, meaning sequential therapy are better than alternating therapy.

Confidence Interval

```
SE <- summary(b1)$standard.errors
#95% CI
CI.u <- exp(beta1[1,2]+ 1.96*SE[1,2])
CI.l <- exp(beta1[1,2]- 1.96*SE[1,2])
cbind(CI.l,CI.u)

##           CI.l           CI.u
## [1,] 1.037453 3.302873
```

From the confidence interval, we can also conclude that we are 95% confident that sequential therapy has a better treatment effect than alternating treatment effect since 1 is not trapped by our confidence interval.

Cummulative Logit Model

Fit

```
library(VGAM)

## Warning: package 'VGAM' was built under R version 3.4.4
## Loading required package: stats4
## Loading required package: splines
c1=vglm(Y~therapy,propodds(reverse=F))
summary(c1)

##
## Call:
## vglm(formula = Y ~ therapy, family = propodds(reverse = F))
##
## Pearson residuals:
##   logitlink(P[Y<=1]) logitlink(P[Y<=2]) logitlink(P[Y<=3])
## 1          -0.2129          -0.2992          0.04569
## 2          -1.0015           0.9976          1.25878
## 3          -0.1322          -0.4179         -1.12409
## 4           1.6003           0.5881           0.91423
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept):1  -0.6468    0.1620  -3.993 6.52e-05 ***
## (Intercept):2   0.9081    0.1667   5.449 5.06e-08 ***
## (Intercept):3   1.9502    0.1984   9.831 < 2e-16 ***
## therapyS        -0.5699    0.2116  -2.694 0.00706 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Names of linear predictors: logitlink(P[Y<=1]), logitlink(P[Y<=2]),
## logitlink(P[Y<=3])
##
## Residual deviance: 9.1642 on 8 degrees of freedom
##
## Log-likelihood: -27.3399 on 8 degrees of freedom
##
## Number of Fisher scoring iterations: 4
##
## No Hauck-Donner effect found in any of the estimates
##
## Exponentiated coefficients:
##   therapyS
## 0.5655741
```

Coefficients and interpretation

Exponentiated coefficients: $e^\beta = e^{-0.5699} = 0.5655741$. Interpretation: $\log \frac{\frac{Pr(Y \leq j | Therapy=S)}{Pr(Y > j | Therapy=S)}}{\frac{Pr(Y \leq j | Therapy=A)}{Pr(Y > j | Therapy=A)}} = \beta \implies \frac{Pr(Y \leq j | Therapy=S)}{Pr(Y > j | Therapy=S)} = e^\beta = e^{-0.5699} = 0.5655741$. Since 0.5655741 is smaller than 1, we can conclude that sequential therapy has a better effect than alternating therapy.

Confidence Interval

```
e=exp(1)
e^confint(c1)

##                2.5 %      97.5 %
## (Intercept):1 0.3812571 0.7193911
## (Intercept):2 1.7887255 3.4375737
## (Intercept):3 4.7655893 10.3708532
## therapyS      0.3735997 0.8561948
```

95% confidence interval for exponentiated coefficients of therapy:[0.3735997, 0.8561948]. Interpretation: we are 95% confident that the true odds ratio for the effect of therapy fall within [0.3735997, 0.8561948].

Should Gender Be Included in the Models Above?

Baseline Categorical Model

```
library(nnet)
b2<-multinom(Y~therapy+gender)

## # weights:  16 (9 variable)
## initial value 414.502014
## iter  10 value 393.034226
## final value 393.017781
## converged

beta2<-coefficients(b2)
summary(b2)

## Call:
## multinom(formula = Y ~ therapy + gender)
##
## Coefficients:
##              (Intercept)  therapyS    genderM
## NoChange              -0.112579  0.6164483  0.09023528
## PartialRemission      -1.138719  0.8978800  0.36184990
## CompleteRemission     -2.060487  0.7986133  1.27828139
##
## Std. Errors:
##              (Intercept)  therapyS    genderM
## NoChange              0.3676934  0.2954664  0.3786301
## PartialRemission      0.4759568  0.3511357  0.4771676
## CompleteRemission      0.6614644  0.3677732  0.6605390
```

```
##
## Residual Deviance: 786.0356
## AIC: 804.0356
```

```
anova(b1,b2,test="Chi")
```

```
##           Model Resid. df Resid. Dev   Test    Df LR stat.    Pr(Chi)
## 1 as.factor(therapy)         6   791.2826          NA      NA      NA
## 2  therapy + gender          3   786.0356 1 vs 2     3 5.247038 0.1545762
```

According to the anova table, the difference between baseline category logit model with therapy and the one with therapy+gender is not statistically significant. Therefore, we would not include gender as a predictor.

Cummulative Logit Model

```
c2<-vglm(Y~therapy+gender,propodds(reverse=F))
summary(c2)
```

```
##
## Call:
## vglm(formula = Y ~ therapy + gender, family = propodds(reverse = F))
##
## Pearson residuals:
##   logitlink(P[Y<=1]) logitlink(P[Y<=2]) logitlink(P[Y<=3])
## 1           0.1720           0.06056           0.2809
## 2          -1.6543           0.16312           0.8425
## 3           0.2655          -0.13909          -0.9386
## 4           0.6174          -0.01519           0.6444
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept):1  -0.1960    0.2947  -0.665  0.50605
## (Intercept):2   1.3713    0.3059   4.482 7.38e-06 ***
## (Intercept):3   2.4221    0.3276   7.393 1.43e-13 ***
## therapyS        -0.5807    0.2119  -2.741  0.00613 **
## genderM         -0.5414    0.2953  -1.834  0.06671 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Names of linear predictors: logitlink(P[Y<=1]), logitlink(P[Y<=2]),
## logitlink(P[Y<=3])
##
## Residual deviance: 5.5677 on 7 degrees of freedom
##
## Log-likelihood: -25.5417 on 7 degrees of freedom
##
## Number of Fisher scoring iterations: 5
##
## No Hauck-Donner effect found in any of the estimates
##
## Exponentiated coefficients:
##   therapyS   genderM
## 0.5595152 0.5819358
```

```
p_val <- pchisq(2*(logLik(c2)-logLik(c1)),df=df.residual(c1)-df.residual(c2), lower.tail=FALSE)
p_val
```

```
## [1] 0.05790126
```

The p value from the test is $0.057 < 0.05$, from which we can conclude that gender should not be included in the cumulative logit model.

Conclusion

Therefore, we conclude that gender should not be included in the model(s).