

VI. Probability Distribution

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Special Distribution Recap



Special Discrete Distributions



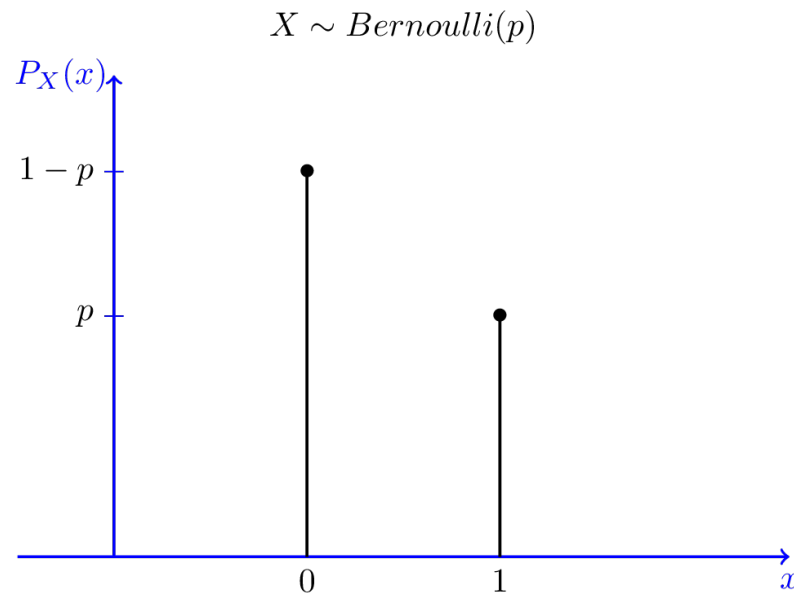
Bernoulli Distribution

Definition 3.4

A random variable X is said to be a *Bernoulli* random variable with *parameter* p , shown as $X \sim \text{Bernoulli}(p)$, if its PMF is given by

$$P_X(x) = \begin{cases} p & \text{for } x = 1 \\ 1 - p & \text{for } x = 0 \\ 0 & \text{otherwise} \end{cases}$$

where $0 < p < 1$.



Special Discrete Distributions

❖ PMFs

Discrete Random Variables

Distribution	PMF / Formula	Mean (μ)	Variance (σ^2)
Discrete Uniform	$P(X = x) = \frac{1}{n}$ for $x = 1, 2, \dots, n$	$\frac{n+1}{2}$	$\frac{(n^2-1)}{12}$
Binomial	$P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$	np	$np(1-p)$
Geometric	$P(X = k) = (1-p)^{k-1} p$	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Negative Binomial	$P(X = k) = \binom{k-1}{r-1} p^r (1-p)^{k-r}$	$\frac{r}{p}$	$\frac{r(1-p)}{p^2}$
Hypergeometric	$P(X = k) = \frac{\binom{K}{k} \binom{N-K}{n-k}}{\binom{N}{n}}$	$\frac{nK}{N}$	$\frac{nK(N-K)(N-n)}{N^2(N-1)}$
Poisson	$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}$	λ	λ



Special Continuous Distributions



Special Continuous Distributions

❖ PDFs

Continuous Random Variables

Distribution	PDF / Formula	Mean (μ)	Variance (σ^2)
Continuous Uniform	$f(x) = \frac{1}{b-a}$ for $a \leq x \leq b$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Exponential	$f(x) = \lambda e^{-\lambda x}$ for $x \geq 0$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$
Normal	$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	μ	σ^2

❖ CDFs:

$$X \sim U(a, b)$$

The CDF of X , $F(x)$, is defined as:

$$F(x) = P(X \leq x) = \begin{cases} 0 & \text{for } x < a \\ \frac{x-a}{b-a} & \text{for } a \leq x \leq b \\ 1 & \text{for } x > b \end{cases}$$

$$X \sim \text{Exponential}(\lambda)$$

The CDF of X , $F(x)$, is defined as:

$$F(x) = P(X \leq x) = \begin{cases} 0 & \text{for } x < 0 \\ 1 - e^{-\lambda x} & \text{for } x \geq 0 \end{cases}$$



Special Distributions – Supplementary Contents



Poisson Distribution



❖ Standard Poisson Distribution for a fixed interval

❖ **Formula:**

$$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}$$

• **Explanation:**

- λ is the average number of events per unit time or space.
- k is the number of events.
- This formula gives the probability of exactly k events occurring in a fixed interval when the average rate of events is λ .



- ❖ **A Poisson process** models a sequence of events occurring randomly over time or space.
- ❖ **Key Properties:**
 - **Independent Events:** Number of events in disjoint intervals are independent.
 - **Stationary Increments:** Probability of events depends only on the length of the interval.
 - **No Simultaneous Events:** Probability of multiple events occurring at the same time is zero.
 - **Homogeneity:** In a homogeneous Poisson process, the rate (λ) is constant.



Poisson Process Distribution (1 unit time $\rightarrow t$)

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❖ Poisson distribution within the context of a Poisson process over time t

❖ **Formula:**

$$P(N(t) = k) = \frac{(\lambda t)^k e^{-\lambda t}}{k!}$$

❖ **Explanation:**

- λ is the average number of events per unit time.
- t is the length of the time interval.
- k is the number of events.
- This formula gives the probability of exactly k events occurring in a time interval of length t when the average rate of events is λ .



Exponential Distribution

- ❖ Purpose: Time between events in a Poisson process with rate parameter λ .
- ❖ PDF: $f(x) = \lambda e^{-\lambda x}, x \geq 0$
- ❖ CDF: $F(x) = 1 - e^{-\lambda x}, x \geq 0$



Poisson Dis. and Exponential Dis.

- ❖ **Poisson Distribution:** Describes the number of events in a fixed interval of time or space.
- ❖ **Exponential Distribution:** Describes the time between successive events in a Poisson process.

❖ Summary:

- **Link:** The parameter λ is common to both distributions, representing the rate of events in a Poisson process.
- **Difference:**
 - In the Poisson distribution, λ describes the average number of events in a fixed interval.
 - In the exponential distribution, λ describes the rate of events per unit time, which is the reciprocal of the average time between events.



Normal Distribution



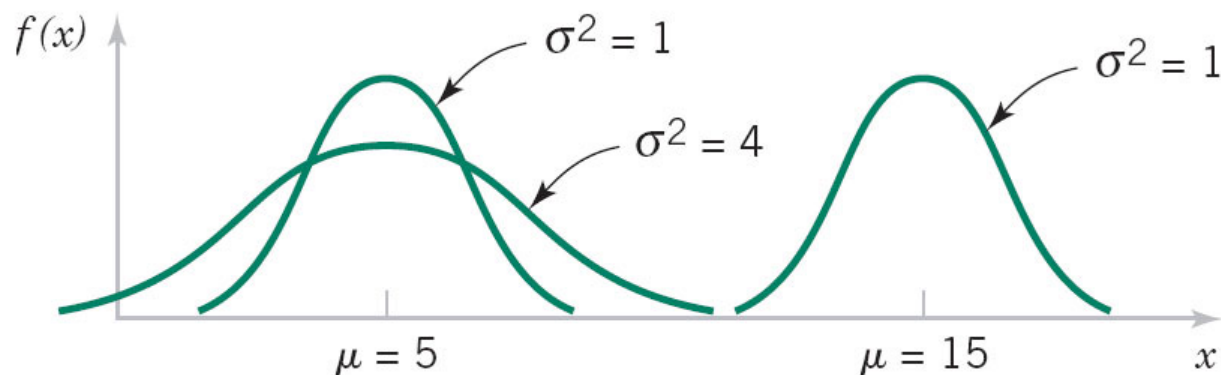
❖ The PDF of a normal RV is

$$f(X) = N(\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

❖ with

$$E(X) = \mu$$

$$V(X) = \sigma^2$$



Standard Normal Distribution

- ❖ Special case with mean 0 and standard deviation 1
- ❖ PDF:

$$f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

- ❖ A standard normal RV is often denoted as Z
- ❖ The CDF of a normal RV is denoted as

$$\Phi(z) = P(Z \leq z)$$

- ❖ Values for $\Phi(z)$ can be found in many Normal tables



Standard Normal Distribution

❖ Special case with mean 0 and standard deviation 1

❖ PDF:

$$f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

❖ **Properties:**

- **68-95-99.7 Rule:**

- 68% within one standard deviation
- 95% within two standard deviations
- 99.7% within three standard deviations

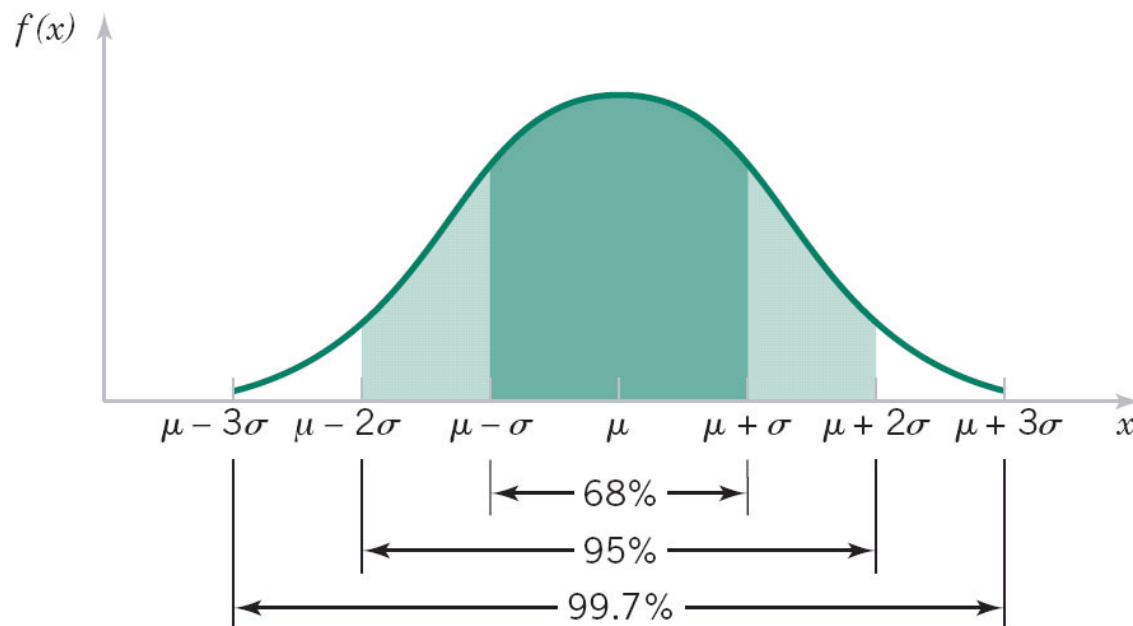


❖ For any normal RV

❖ $P(\mu - \sigma < X < \mu + \sigma) = 0.6827$

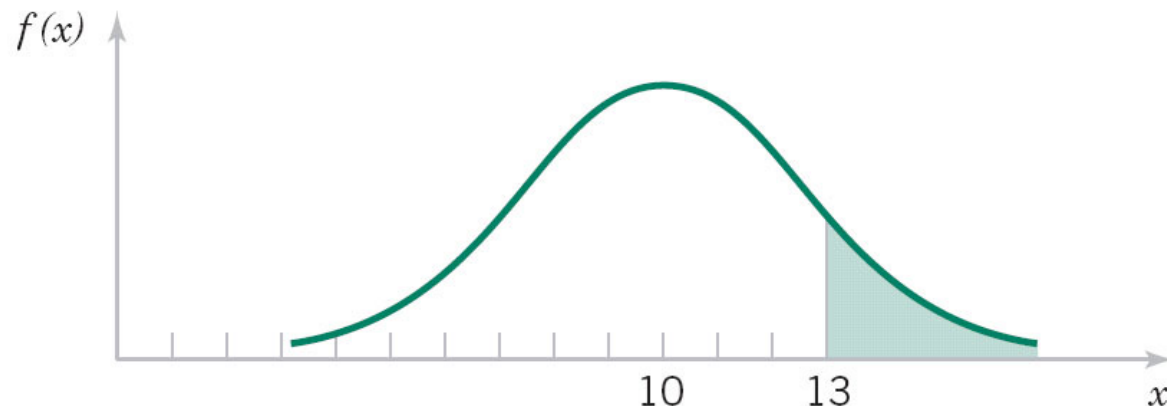
❖ $P(\mu - 2\sigma < X < \mu + 2\sigma) = 0.9545$

❖ $P(\mu - 3\sigma < X < \mu + 3\sigma) = 0.9973$



Example [$X = N(\mu, \sigma^2) \rightarrow Z = N(0, 1)$]

- ❖ A concrete strength has mean 10MPa/variance 4MPa²
- ❖ What is the probability that a test result exceeds 13 MPa?

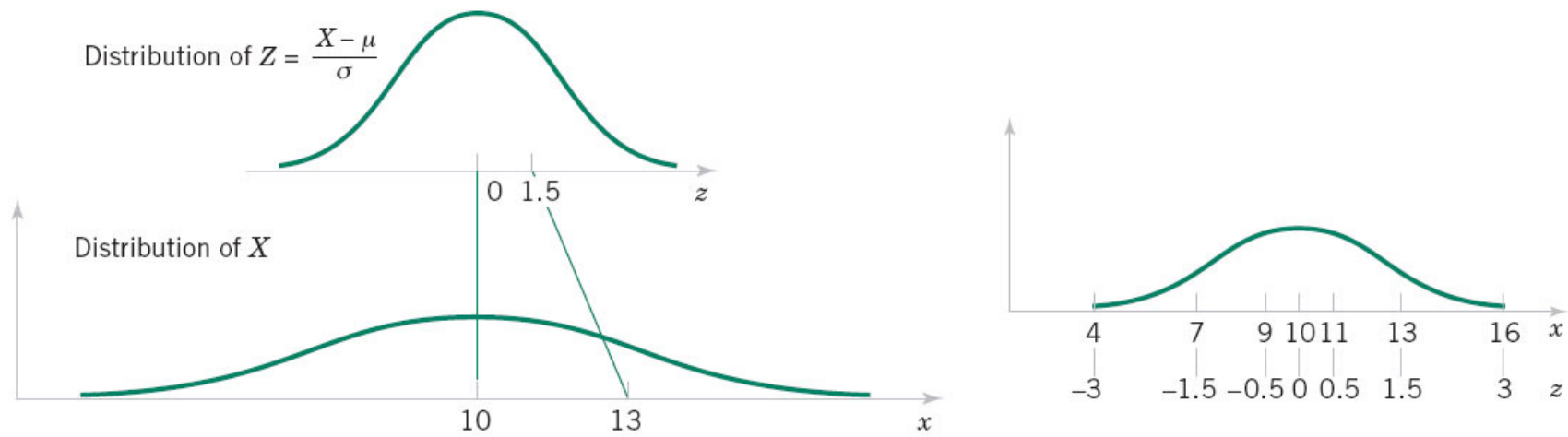


- ❖ Unfortunately, there is no closed form for the normal CDF
- ❖ To this purpose, tables are used



Example [$X = N(\mu, \sigma^2) \rightarrow Z = N(0, 1)$]

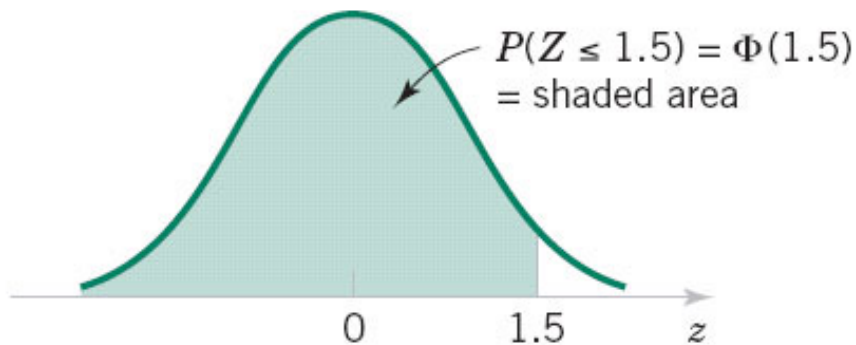
- ❖ A concrete strength has mean 10MPa/variance 4MPa²
- ❖ What is the probability that a test result exceeds 13 MPa?
 - ❖ Standardize: $Z = (X - 10)/2$
 - ❖ Then: $X > 13 \rightarrow Z > 1.5$
 - ❖ From Table III: $P(Z > 1.5) = 1 - P(Z \leq 1.5) = 1 - 0.93319 = 0.06681$



Example [$X = N(\mu, \sigma^2) \rightarrow Z = N(0, 1)$]

- ❖ Find the probability that a standard NRV is lower than 1.5

$$P(Z \leq 1.5) = \Phi(1.5)$$



z	0.00	0.01	0.02	0.03
0	0.50000	0.50399	0.50398	0.51197
\vdots		\vdots		
1.5	0.93319	0.93448	0.93574	0.93699

If you have to find $P(Z \leq 1.53)$, go here



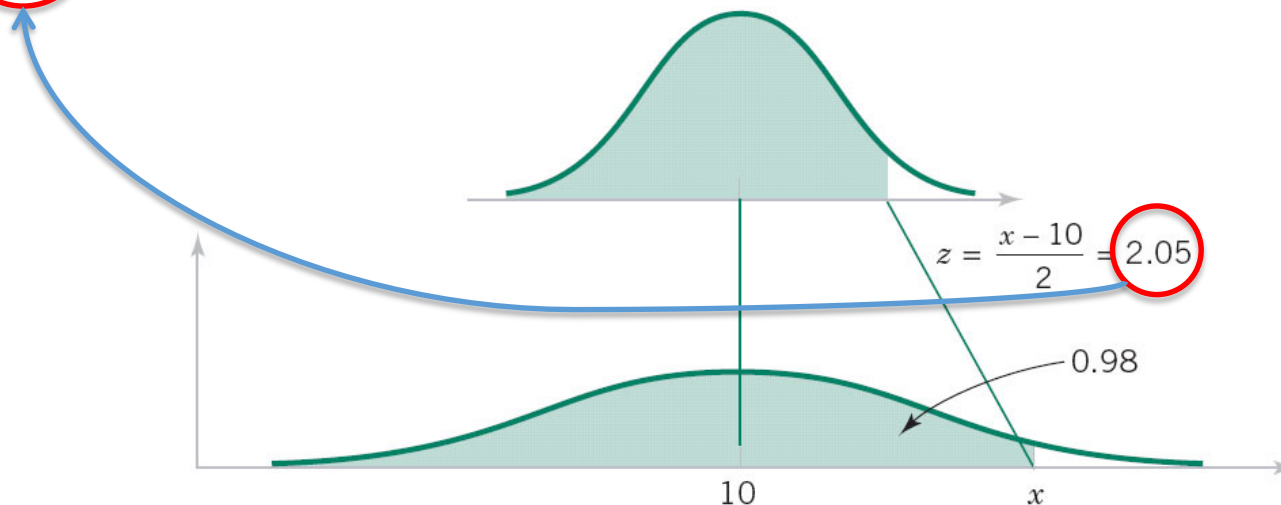
Example [$X = N(\mu, \sigma^2) \rightarrow Z = N(0, 1)$]

❖ What is $P(9 < X < 11)$?

$$\begin{aligned} \text{❖ } P(9 < X < 11) &= P(-0.5 < Z < 0.5) = P(Z < 0.5) - P(Z < -0.5) \\ &= 0.69146 - 0.30854 = 0.38292 \end{aligned}$$

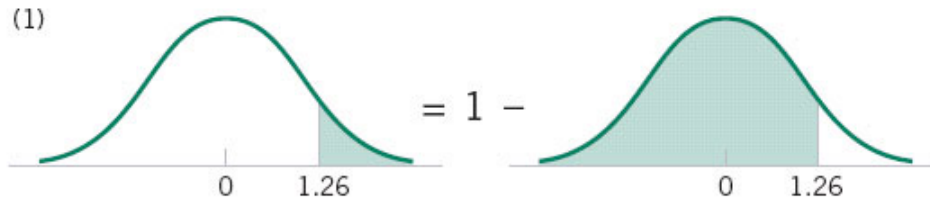
❖ Inverse problem: What is x so that $P(X < x) = 0.98$?

$$\text{❖ } x = 2(2.05) + 10 = 14.1 \text{ MPa}$$

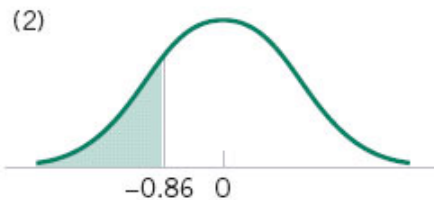


Examples (using basic probability and symmetry)²⁴

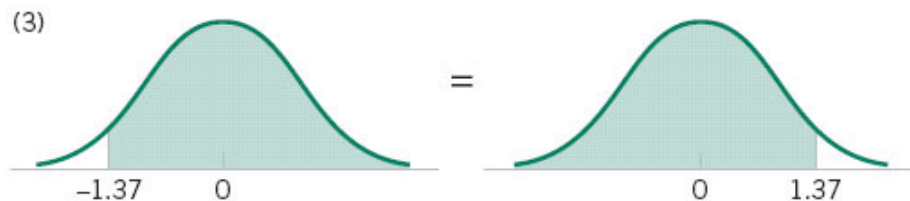
$$P(Z > 1.26)$$



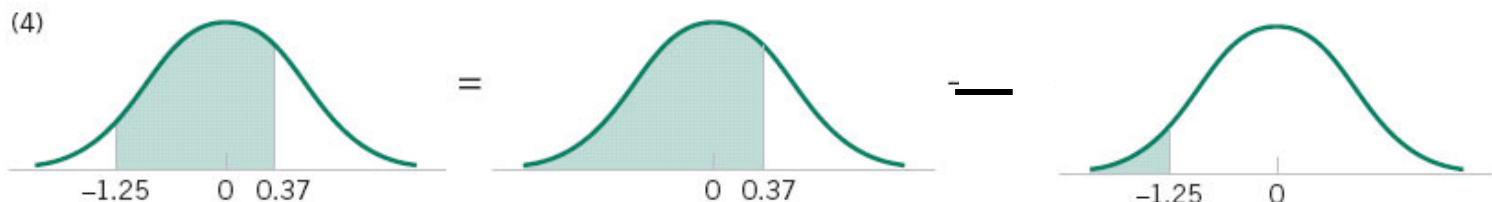
$$P(Z < -0.86)$$



$$P(Z > -1.37)$$

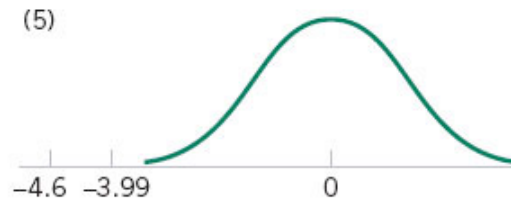


$$P(-1.25 < Z < 0.37)$$



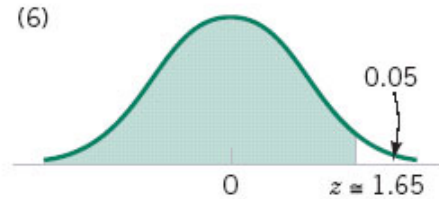
Examples (using basic probability and symmetry)²⁵

$$P(Z \leq -4.6)$$

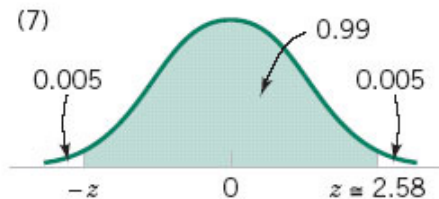


$$P(Z > z) = 0.05$$

$$P(Z \leq z) = 0.95$$



$$P(-z < Z < z) = 0.99$$



Example [using basic probability and symmetry]²⁶

Assume X is normally distributed with a mean of 10 and a standard deviation of 2. Determine:

$$\begin{aligned}\text{a) } P(2 < X < 4) &= P\left(\frac{2-10}{2} < Z < \frac{4-10}{2}\right) \\ &= P(-4 < Z < -3) \\ &= P(Z < -3) - P(Z < -4) \\ &= 0.00132\end{aligned}$$

$$\begin{aligned}\text{b) } P(X > 9) &= 1 - P(X < 9) \\ &= 1 - P(Z < (9-10)/2) \\ &= 1 - P(Z < -0.5) \\ &= 0.69146\end{aligned}$$

$$\begin{aligned}\text{c) } P(-2 < X < 8) &= P(X < 8) - P(X < -2) \\ &= P\left(Z < \frac{8-10}{2}\right) - P\left(Z < \frac{-2-10}{2}\right) \\ &= P(Z < -1) - P(Z < -6) \\ &= 0.15866\end{aligned}$$

$$\begin{aligned}\text{d) } P(6 < X < 14) &= P\left(\frac{6-10}{2} < Z < \frac{14-10}{2}\right) \\ &= P(-2 < Z < 2) \\ &= P(Z < 2) - P(Z < -2) \\ &= 0.9545\end{aligned}$$

Review Examples – Special Discrete Distributions



PMF – General Exampe

❖ I toss a fair coin twice, and let X be defined as the number of heads I observe. Find the range of X , R_X , as well as its probability mass function P_X .

❖ Sol:

$$❖ S = \{HH, HT, TH, TT\}$$

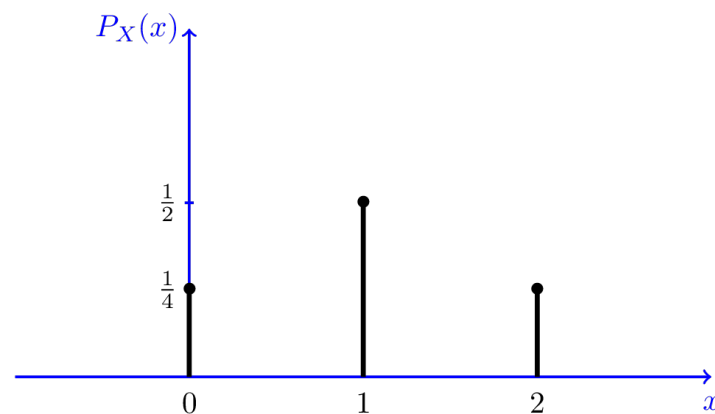
$$❖ R_X = \{0, 1, 2\}, P_X(k) = P(X = k) \text{ for } k = \{0, 1, 2\}$$

$$P_X(0) = P(X = 0) = P(TT) = \frac{1}{4},$$

$$P_X(1) = P(X = 1) = P(\{HT, TH\}) = \frac{1}{4} + \frac{1}{4} = \frac{1}{2},$$

$$P_X(2) = P(X = 2) = P(HH) = \frac{1}{4}.$$

$$P_X(x) = \begin{cases} P(X = x) & \text{if } x \text{ is in } R_X \\ 0 & \text{otherwise} \end{cases}$$



Binomial vs Hypergeometric Dist.



Binomial vs Hypergeometric Dist. – Example 1

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- ❖ We draw 10 balls at random from an urn containing 30 red and 20 blue balls, and let $X = \# \text{red balls obtained}$.
- ❖ Derive the PMFs for the following two scenarios:
 - ❖ Drawing balls with replacement
 - ❖ Drawing balls without replacement



❖ **Scenario 1: Drawing with Replacement**

❖ When drawing with replacement, each draw is independent of the previous ones. Therefore, the number of red balls obtained follows a **Binomial distribution**.

❖ Let's define:

❖ **Binomial RV**: $X = \text{\#red balls obtained}$.

❖ $n=10$ as the number of draws.

❖ $p = 30/50 = 0.6$ as the probability of drawing a red ball on each draw.

❖ PMF: $X \sim B(10, 0.6)$ $P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}$

❖ Substituting the values:

$$P(X = k) = \binom{10}{k} (0.6)^k (0.4)^{10-k}$$



❖ Scenario 2: Drawing without Replacement

❖ When drawing without replacement, the draws are dependent on each other. Therefore, the number of red balls obtained follows a **Hypergeometric distribution**.

❖ Let's define:

- **Hypergeometric RV:** $X = \# \text{red balls obtained}$.
- $N=50$ as the total number of balls.
- $K=30$ as the number of red balls.
- $n=10$ as the number of draws.

❖ PMF: $X \sim \text{HyperGem}(N, k, n) = P(X = k) = \frac{\binom{K}{k} \binom{N-K}{n-k}}{\binom{N}{n}}$

❖ Substituting the values:

$$P(X = k) = \frac{\binom{30}{k} \binom{50-30}{10-k}}{\binom{50}{10}}$$



Geometric Distribution



Geometric Distribution – Example 1

- ❖ Suppose Max owns a lightbulb manufacturing company and determines that 3 out of every 75 bulbs are defective.
- ❖ Q:
 - ❖ What is the probability that Max will find the first faulty lightbulb on the 6th one that he tested?
 - ❖ What is the number lightbulbs we would expect Max to inspect until he finds his first defective, as well as the standard deviation?



Geometric Distribution – Example 1

- ❖ Q1 Sol. $P(\{\text{the first faulty lightbulb on the 6th one}\})$
- ❖ Define Geometric RV: $X = \# \text{tests to find the first faulty lightbulb}$
- ❖ $X \sim \text{Geom}(p)$

$$p = \frac{3}{75} = 0.04$$

$$k = 6$$

$$P(X = k) = p(1 - p)^{k-1}$$

$$P(X = 6) = 0.04(1 - 0.04)^{6-1}$$

$$P(X = 6) = 0.04(0.96)^5 = 0.0326$$



Geometric Distribution – Example 1

- ❖ Q2 Sol. $P(\{\text{Expected \#tests until the first defective}\})$
- ❖ Define Geometric RV: $X = \text{\#tests to find the first faulty lightbulb}$
- ❖ $X \sim \text{Geom}(p)$

Mean:	$\mu = E(X) = \frac{1}{p}$ $E(X) = \frac{1}{0.04} = 25$	
Variance:	$\sigma^2 = V(X) = \frac{(1-p)}{p^2}$ $V(X) = \frac{(1-0.04)}{(0.04)^2} = 600$	$p = \text{probability of success}$ $p = 0.04$
Standard Deviation	$\sigma = \sqrt{\sigma^2} = \sqrt{600} = 24.49$	



Poisson Distribution



Poisson Distribution – Example 1

- ❖ The number of emails that one gets in a weekday can be modeled by a Poisson distribution with an average of 0.2 emails per minute.
 - What is the probability that one gets no emails in an interval of length 5 minutes?
 - What is the probability that one gets more than 3 emails in an interval of length 10 minutes?



Poisson Distribution – Example 1

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❖ Q1

- ❖ Define Poisson RV: **X** = #emails obtained in the 5-minute interval
- ❖ Then, by the assumption **X** is a Poisson RV with parameter $\lambda = 5 \cdot 0.2 = 1$
- ❖ $X \sim \text{Pois}(\lambda = 1)$

$$P(X = 0) = P_X(0) = e^{-\lambda} \frac{\lambda^0}{0!} = e^{-1} \cdot \frac{1^0}{1} = e^{-1} \approx 0.3679.$$

❖ Q2

- ❖ Define Poisson RV: **Y** = #emails obtained in the 10-minute interval.
- ❖ Then, by the assumption **Y** is a Poisson RV with $\lambda = 10 \cdot 0.2 = 2$
- ❖ $Y \sim \text{Pois}(\lambda = 2)$

$$\begin{aligned} P(Y > 3) &= 1 - P(Y \leq 3) = 1 - (P_Y(0) + P_Y(1) + P_Y(2) + P_Y(3)) \\ &= 1 - (e^{-2} + 2e^{-2} + 2e^{-2} + \frac{8}{6}e^{-2}) \\ &= 1 - e^{-2}(1 + 2 + 2 + \frac{8}{6}) \\ &= 1 - \frac{19}{3}e^{-2} \approx 0.1429. \end{aligned}$$



Review Examples – Special Continuous Distributions



Continuous Uniform Distribution



Continuous Uniform Distribution – Example 1

- ❖ The amount of time, in minutes, that a person must wait for a bus is **uniformly distributed** between zero and 15 minutes, inclusive.
- ❖ Problem
 - ❖ a. What is the probability that a person waits **fewer than 12.5 minutes**?
 - ❖ b. On the average, how long must a person wait? Find the mean, μ , and the standard deviation, σ .
 - ❖ c. To what value does the wait time fall below 90% of the time?



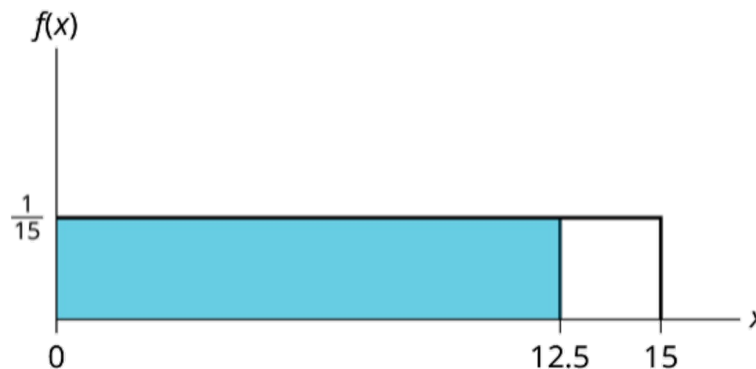
Continuous Uniform Distribution – Example 1

❖ Sol. P1a. What is the probability that a person waits **fewer than 12.5 minutes**?

❖ Let X = #minutes a person must wait for a bus.

❖ $a = 0$ and $b = 15$. $X \sim U(0, 15)$.

❖ PDF: $f(x) = \frac{1}{15-0} = \frac{1}{15}$, for $0 \leq x \leq 15$.



$$P(x < k) = (\text{base})(\text{height}) = (12.5 - 0) \left(\frac{1}{15} \right) = 0.8333$$



Continuous Uniform Distribution – Example 1

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❖ Sol. P2

❖ PDF: $f(x) = \frac{1}{15}$, for $0 \leq x \leq 15$.

b. $\mu = \frac{a+b}{2} = \frac{15+0}{2} = 7.5$. On the average, a person must wait 7.5 minutes.

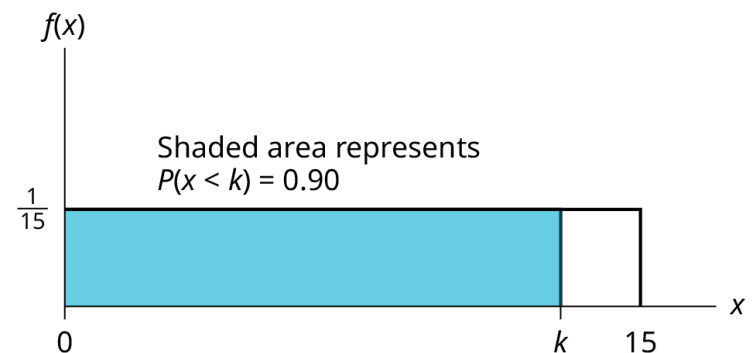
$\sigma = \sqrt{\frac{(b-a)^2}{12}} = \sqrt{\frac{(15-0)^2}{12}} = 4.3$. The Standard deviation is 4.3 minutes.

❖ Sol. P3 To what value does the wait time fall below 90% of the time?

❖ $P(X < k) = F(x = k) = \frac{k-a}{b-a} = (k - 0)\left(\frac{1}{15}\right)$

❖ $0.90 = (k) \left(\frac{1}{15}\right) 0.90 = (k) \left(\frac{1}{15}\right)$

❖ $k = (0.90)(15) = 13.5$



❖ The 90th percentile is 13.5 minutes. Ninety percent of the time, a person must wait at most 13.5 minutes.



Continuous Uniform Distribution – Exercise

- ❖ The thickness x of a protective coating applied to a conductor designed to work in corrosive conditions follows a uniform distribution over the interval $[20, 40]$ microns.
- ❖ Problems
 - ❖ Find the mean, standard deviation and cumulative distribution function of the thickness of the protective coating.
 - ❖ Find also the probability that the coating is less than 35 microns thick.





Exponential Distribution



Exponential Dist. vs Poisson Dist. – Example 1

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- ❖ Suppose, in a call center, calls arrive at an average rate of 3 calls per hour.
- ❖ Problems:
 - ❖ A. What is the probability that the next call will arrive **within** the next 10 minutes?
 - ❖ B. What is the probability that **exactly** 5 calls will arrive in the next hour?



Poisson Dist. vs Exponential Dist. – Example 1

- ❖ Sol. $Q(A)$ $P(\{\text{the next call will arrive **within** the next 10 minutes}\})$
- ❖ Define:
 - **Exponential RV**: X = The time until the next call arrives at the call center
 - **Parameter**: Average rate $\lambda=3$ calls per hour.
 - Time $x = \frac{10}{60} = \frac{1}{6}$ hours (since 10 minutes is $\frac{1}{6}$ of an hour).
- ❖ PDF: $X \sim \text{Exp}(\lambda = 3), f(x) = \lambda e^{-\lambda x}, x \geq 0$
- ❖ CDF: $F(x) = 1 - e^{-\lambda x}$
- ❖ Substitute the values:

$$P(X \leq \frac{1}{6}; 3) = 1 - e^{-3 \times \frac{1}{6}} = 1 - e^{-0.5} \approx 0.3935$$



Poisson Dist. vs Exponential Dist. – Example 1

- ❖ Sol. Q(B) $P(\{\textbf{exactly}$ 5 calls will arrive in the next hour $\})$
- ❖ Given:
 - **Poisson RV:** $X = \# \text{calls arriving at the call center in an hour}$
 - **Parameter:** Average rate $\lambda=3$ calls per hour.
 - Number of calls $k=5$.
- $X \sim \text{Pois}(\lambda = 3)$
- Using the Poisson distribution formula:

$$P(X = k; \lambda) = \frac{\lambda^k e^{-\lambda}}{k!}$$

- ❖ Substitute the values:

$$P(X = 5; 3) = \frac{3^5 e^{-3}}{5!} = \frac{243 e^{-3}}{120} \approx 0.1008$$



- ❖ Example Scenario: Hospital ER
- ❖ In a hospital ER, patients arrive at an average rate of 6 patients per hour.
- ❖ Questions:
 - ❖ What is the probability that exactly 8 patients will arrive at the ER in the next hour?
 - ❖ What is the probability that the next patient will arrive within the next 5 minutes?

(You may leave the answer expressed as a fraction involving e and factorials.)





Normal Distribution



❖ **Example Scenario: Heights of Adults**

- ❖ Suppose the heights of adult men in a certain country are normally distributed with a mean $\mu=175$ cm and a standard deviation $\sigma=10$ cm.

❖ **Question:**

- ❖ What is the probability that a randomly selected adult man is taller than 185 cm?



Normal Distribution – Example 1

❖ Sol:

❖ 1. Define the Random Variable:

❖ NRV: X be the height of an adult man

❖ 2. Given the parameters

❖ $X \sim N(\mu, \sigma^2)$ $f(x) = \frac{1}{10\sqrt{2\pi}} e^{-\frac{(x-175)^2}{2 \cdot 100}}$

❖ 3. Calculate the Z-Score: $Z = \frac{X - \mu}{\sigma}$

❖ For $x=185$ cm $Z = \frac{185 - 175}{10} = \frac{10}{10} = 1$

❖ $P(X > 185) \rightarrow P(Z > 1)$ in the standard $N(0, 1)$

$$P(Z > 1) = 1 - P(Z \leq 1) \quad P(Z \leq 1) \approx 0.8413.$$

$$P(Z > 1) = 1 - 0.8413 = 0.1587$$



❖ Question:

- ❖ Find the probability that a randomly selected adult man is between 165 cm and 185 cm
- ❖ (Hint: Draw the PDF and use the symmetry of Normal dist.)

$$P(Z \leq 1) \approx 0.8413.$$

$$P(Z > 1) = 1 - 0.8413 = 0.1587$$



