

# GESIS Spring Seminar 23

## Comparative Social Research with Multi-Group SEM

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Day 2 - 28.02.2023

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## Packages

```
# basic tools
library(tidyverse)
library(sjmisc)

# checking data
library(MVN)
library(correlation)
library(performance)
library(parameters)
library(insight)

# building the models
library(lavaan)

# cheking the models
library(semTools)

# model visualization
library(semPlot)
library(semptools)
```

# Data

```
ESS07 <- read.csv("../data/ESS07.csv")
```

## Confirmatory measurement analysis (CFA)

Many variables in the social and behavioral sciences represent hypothetical constructs (e.g., traits, attitudes, values)

Direct access to hypothetical constructs is rare. Many constructs can only be captured by (multiple) observable indicators (e.g., in a questionnaire)

It is assumed that the such indicators are manifestations of hypothetical constructs - Manifest indicators: directly observable - Latent variables: not directly observable

We need a model that is able to conceptualize the relationship between manifest indicators and latent variables

The classic test theory Novick (1966)

## Procedure

Step 1: Translation of the theoretical model into a measurement model

- We need a model that is able to conceptualize the relationship between manifest indicators and latent variables

Step 2: Checking whether it is a reflexive or formative construct

- Reflective measurement: indicators are “effects” of the latent variable
- Formative measurement: latent variable is “caused” or composed by the manifest indicators

Step 3: Specification of the model (Choose a method to assign a scale to the latent variable)

- Reference indicator method: factor loading of one indicator is fixed to 1.0
- Fixed factor method: the factor variance is fixed to 1.0

Step 4: Checking the model assumptions (ML estimation requirements)

- Data are continuous and multivariate normal (case of non-normal data standard errors and chi-square-statistics should be adjusted - e.g., robust ML = MLR)
- Sample size is sufficiently large (ratio of number of cases and number of parameters - N:q rule => 20:1 or 10:1)

Step 5: Checking whether the model is identified

- Complexity of a CFA models is limited by the total number of observations (p) and the number of unknown model parameters (q)
- Degrees of freedom must be at least zero:  $df = p - q \geq 0$
- $j$  = number of indicators
- $p = j * (j + 1) / 2$
- $q$  = counting of the parameters to be calculated (factor loading, residual error, factor variance)
- $df = 0$  => model is identified and saturated /  $df > 0$  => model is overidentified

Step 6: Checking the fit measurements

Index	Typ	Theoretical range	Cut-off	N sensitive	Penalty for complexity
X2/df	badness	$\geq 0$	$< 5$	yes	no
CFI	godness	0.0 – 1.0	$\geq 0.95$ ( $\geq 0.90$ )	no	yes
TLI	godness	0.0 – 1.0*	/	no	yes

Index	Typ	Theoretical range	Cut-off	N sensitive	Penalty for complexity
SRMR	badness	$\geq 0$	$< 0.08$	yes	no
RMSEA	badness	$\geq 0$	$\leq 0.05$ ( $\leq 0.08$ )	yes to small N	yes
PCLOSE	badness	0.0 – 1.0	$\geq 0.95$	yes	/

\*negative values indicate extremely misspecified model; when exceeds 1, model is extremely well-fitting

Step 7: Checking for misspecification

- What causes model misfit?  $\Rightarrow$  Indicator choice, Factor choice, Violations of assumptions (e.g., multivariate normality) and Causal structure (e.g., restrictions)
- How to diagnose model misfit?  $\Rightarrow$  Parameter estimates (Heywood cases?), Residual matrices (i.e., differences between observed and estimated covariances), Modification indices (approximation of the reduction of chi-square if a single constrained parameter is freely estimated)

Step 8: Identifying the mean structure (optional)

- Fixing the intercept of the reference indicator to zero
- Fixing the latent mean to zero

## CFA model of Universalism

### Step 1

Our example measurement model Universalism comes from the Theory of Basic Human Values according to Schwartz (1992) and comprises three facets - Concern, Nature, Tolerance.

Concern: Commitment to equality, justice and protection for all people

Nature: Preservation of the natural environment

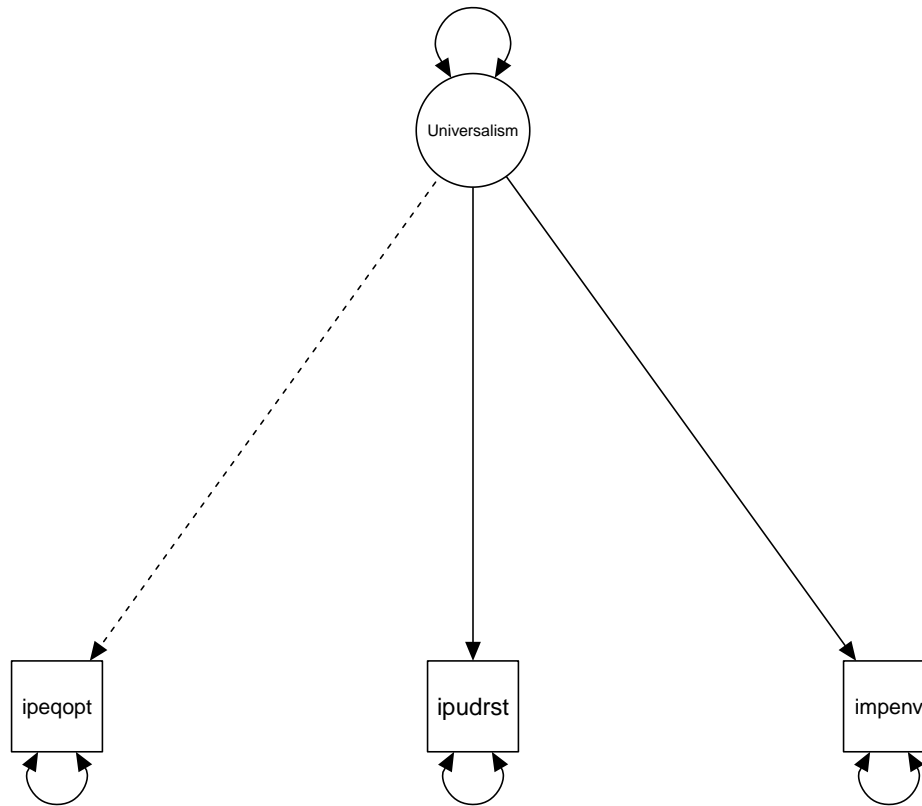
Tolerance: Acceptance and understanding of those who are different from oneself

In order to measure this attitude construct, three questions (items) were created in ESS 7 (2014) and queried in various countries (e.g. Czech Republic, Germany, Great Britain).

Name	Label	Construct
ipeqopt	Important that people are treated equally and have equal opportunities	Universalism
ipudrst	Important to understand different people	Universalism
impenv	Important to care for nature and environment	Universalism

### Step 2

Indicators are “effects” of the latent variable and therefore we specify a reflective measurement model



### Step 3

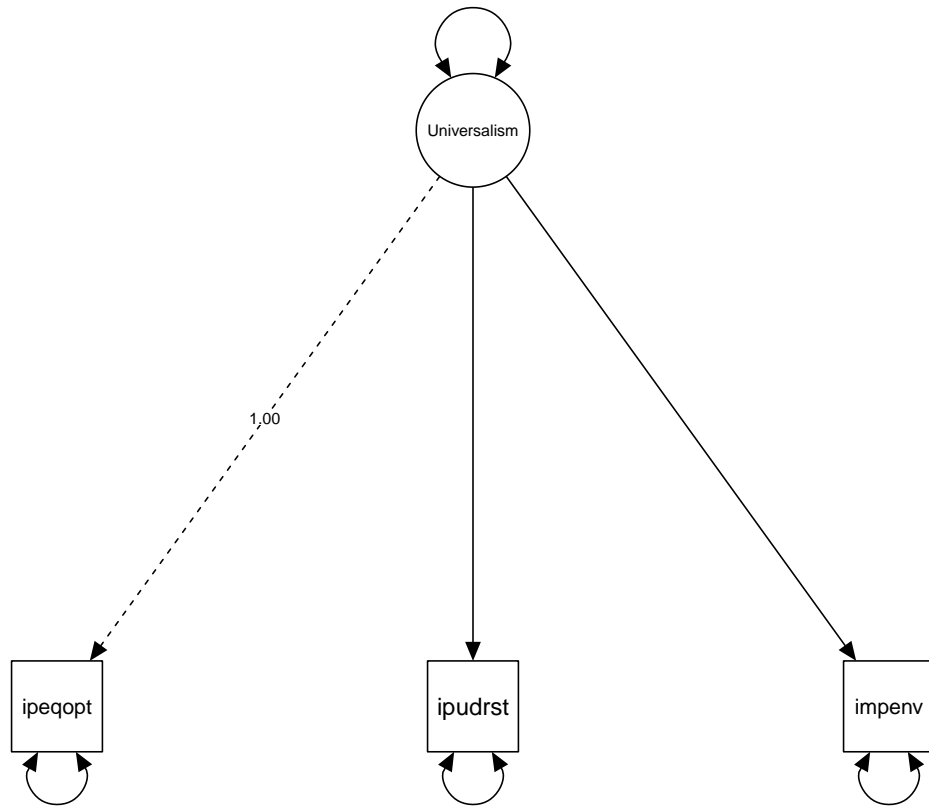
The syntax commands for specifying a lavaan model:

<https://lavaan.ugent.be/tutorial/syntax1.html>

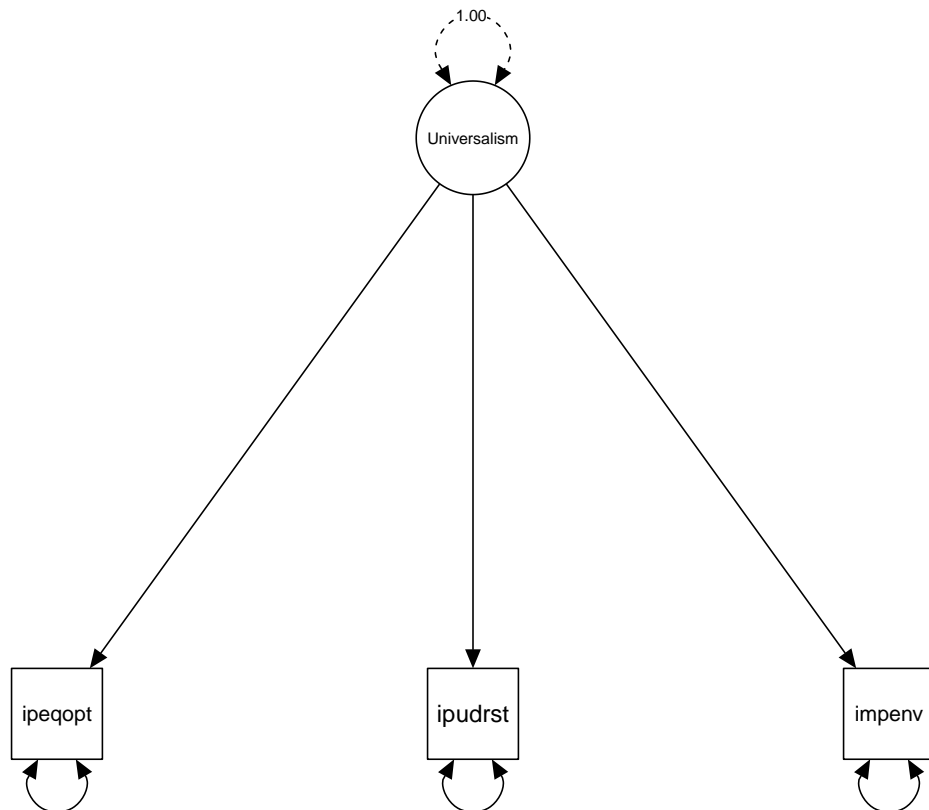
Formula type	Operator	Mnemonic
latent variable definition	<code>=~</code>	is measured by
regression	<code>~</code>	is regressed on
(residual) (co)variance	<code>~~</code>	is correlated with
intercept	<code>~1</code>	intercept

```
universalism <- "uni =~ ipeqopt + ipudrst + impenv"
```

Lavaan automatically uses the reference indicator method, so one indicator's factual loading is fixed at 1. Accordingly, the variance of the latent factor is estimated.



In Lavaan, however, it is also very easy to use the fixed factor method, i.e. to fix the variance of the latent factor to 1. This allows all factor loadings to be freely estimated.



This requires the use of labels, i.e. the labeling of individual parameters in order to be able to control them directly. The syntax commands for labeling a lavaan model can be found here: <https://lavaan.ugent.be/tutorial/syntax2.html>. For this step, we only need to understand how individual parameters can be fixed and freely calculated. As soon as it comes to the calculation of measurement invariance, we will work with labels again. But more on that later.

Reference indicator method:

```
universalism <- "uni =~ 1*ipeqopt + ipudrst + impenv"
```

Fixed factor method:

```
universalism <- "
# specification of the measurement model
uni =~ NA*ipeqopt + ipudrst + impenv

# specification of the variance
uni ~~ 1*uni
"
```

Freely estimate the factor loading of the first indicator by NA\*ipeqopt and fix the variance of the latent factor by 1\*uni. The variance of a variable is specified as covariance with itself, hence uni ~~ uni. In the end we get uni ~~ 1\*uni

#### Step 4

Now let's get an overview of the properties of our variables. First, let's look at the descriptive statistics of our items to gain knowledge of missing values, central tendency properties, variability, and distribution.

```
ESS07 %>%
  select(ipeqopt, ipudrst, impenv) %>%
  mvn() %>%
  # optional for better display
  export_table(table_width = 1, digits = 3)
```

Test	HZ	p value	MVN
Henze-Zirkler	133.042	0	NO

Test	Variable	Statistic	p value	Normality
Anderson-Darling	ipeqopt	362.6880	<0.001	NO
Anderson-Darling	ipudrst	332.9049	<0.001	NO
Anderson-Darling	impenv	322.7889	<0.001	NO

n	Mean	Std.Dev	Median	Min	Max	25th	75th	Skew	Kurtosis
5844	4.945	1.040	5	1	6	4	6	-1.142	1.295
5844	4.717	1.027	5	1	6	4	5	-0.937	0.892
5844	4.878	1.025	5	1	6	4	6	-0.933	0.732

This shows that the individual variables are not normally distributed and that there is no multivariate normal distribution (Henze-Zirkler value). So, we don't have any missing values in our example dataset and a more than sufficient sample size. In addition, the variables show a right skewed distribution and a moderate variance.

Due to the missing multivariate normal distribution, we use the MLR estimator. If we were dealing with missing values, we would use Full Information Maximum Likelihood (FIML) as the estimation strategy.

## Step 5

Now let's estimate the model using Lavaan. For this we use the `cfa()` function.

```
universalism_fit <- cfa(  
  model = "Universalism =~ ipeqopt + ipudrst + impenv",  
  data = ESS07,  
  estimator = "MLR",  
  # optional if there are missing values  
  missing = "fiml.x"  
)
```

Now we can look at the estimated model. To do this, we use the `summary()` function.

`Fit.measures = T` shows us the most important fit measurements of our model and thus the fit to the data. `Standardized = T` shows us standardized values in the output and `rsquare` shows us the explained variance of the dependent variable.

```
summary(  
  object = universalism_fit,  
  fit.measures = T,  
  standardized = T,  
  rsquare = T  
)
```

lavaan 0.6-12 ended normally after 24 iterations

Estimator	ML
Optimization method	NLMINB
Number of model parameters	9
Number of observations	5844
Number of missing patterns	1

Model Test User Model:

	Standard	Robust
Test Statistic	0.000	0.000
Degrees of freedom	0	0

Model Test Baseline Model:

Test statistic	1532.717	1023.100
Degrees of freedom	3	3
P-value	0.000	0.000
Scaling correction factor		1.498

User Model versus Baseline Model:

Comparative Fit Index (CFI)	1.000	1.000
Tucker-Lewis Index (TLI)	1.000	1.000
Robust Comparative Fit Index (CFI)		NA
Robust Tucker-Lewis Index (TLI)		NA

Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-24640.784	-24640.784
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Loglikelihood unrestricted model (H1)	-24640.784	-24640.784
Akaike (AIC)	49299.567	49299.567
Bayesian (BIC)	49359.626	49359.626
Sample-size adjusted Bayesian (BIC)	49331.026	49331.026

Root Mean Square Error of Approximation:

RMSEA	0.000	0.000
90 Percent confidence interval - lower	0.000	0.000
90 Percent confidence interval - upper	0.000	0.000
P-value RMSEA <= 0.05	NA	NA
Robust RMSEA		0.000
90 Percent confidence interval - lower		0.000
90 Percent confidence interval - upper		0.000

Standardized Root Mean Square Residual:

SRMR	0.000	0.000
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Parameter Estimates:

Standard errors	Sandwich
Information bread	Observed
Observed information based on	Hessian

Latent Variables:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
Universalism =~						
ipeqopt	1.000				0.595	0.572
ipudrst	1.059	0.063	16.852	0.000	0.630	0.613
impenv	0.837	0.046	18.010	0.000	0.498	0.486

Intercepts:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.ipeqopt	4.945	0.014	363.381	0.000	4.945	4.753
.ipudrst	4.717	0.013	351.024	0.000	4.717	4.592
.impenv	4.878	0.013	363.981	0.000	4.878	4.761
Universalism	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.ipeqopt	0.729	0.030	24.070	0.000	0.729	0.673
.ipudrst	0.659	0.029	22.371	0.000	0.659	0.624
.impenv	0.802	0.023	34.294	0.000	0.802	0.764
Universalism	0.354	0.027	13.341	0.000	1.000	1.000

R-Square:

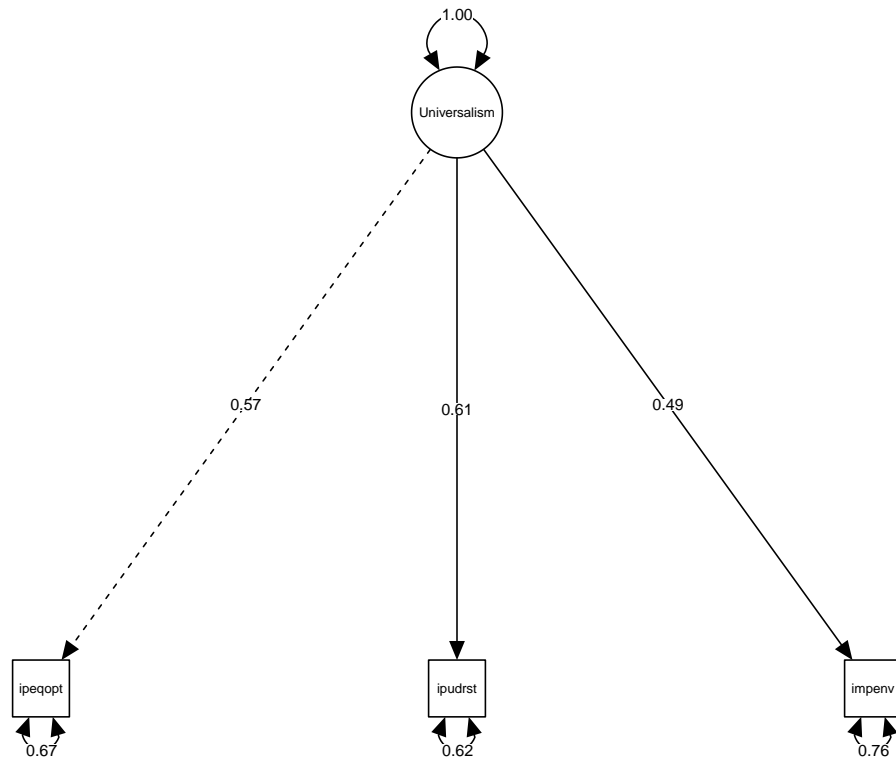
	Estimate
ipeqopt	0.327
ipudrst	0.376
impenv	0.236



Since we estimated a model with three indicators, it is a saturated model, i.e. a model identified with 0 degrees of freedom.

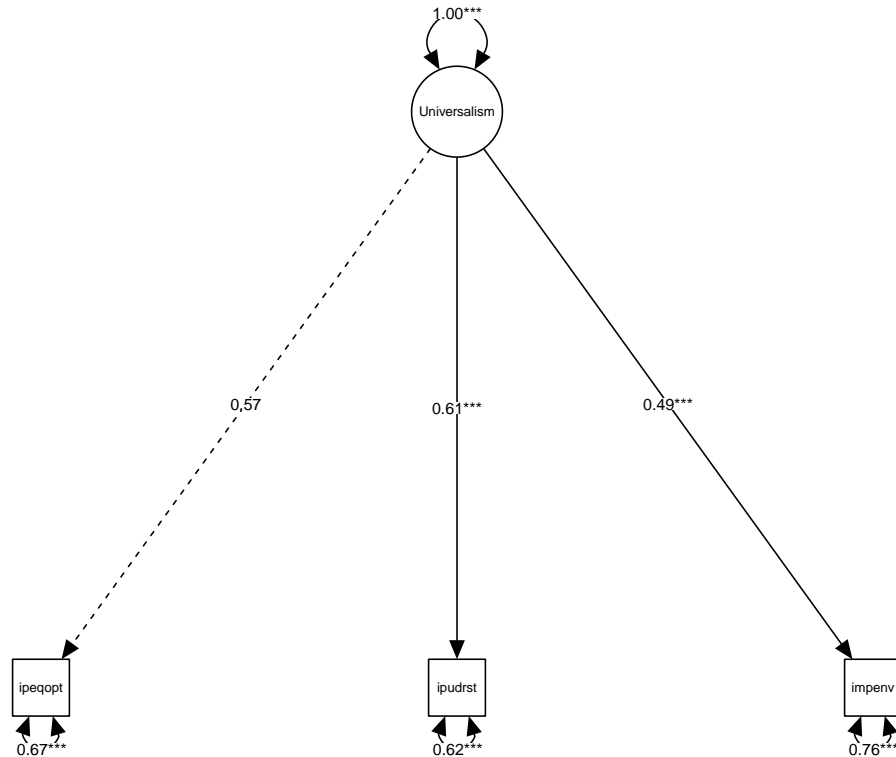
This allows the following graphic to be created with standardized factor loadings.

```
universalism_plot <- universalism_fit %>%
  semPaths(
    whatLabels = "std",
    style = "mx",
    layout = "tree",
    edge.color = "black",
    intercepts = F,
    nCharNodes = 50
  )
```



The semptools package also allows us to easily add additional information to our graphics ( <https://cran.r-project.org/web/packages/semptools/vignettes/semptools.html>). Here I have also included the degree of statistical significance.

```
universalism_plot %>%
  mark_sig(universalism_fit) %>%
  plot()
```



### Step 6:

A look at the fit measures shows us that the model is saturated due to the 0 degrees of freedom and thus achieves a perfect data fit.

### Step 7:

In the following we want to check whether the selected items represent a factor. It can be seen that the factor loadings all reach the cut off of  $\geq 0.4$  and thus, in combination with the fit indices, there is an appropriate measurement model.

If there are recognizable misfits, we can access the modification indices.

Modification indices (MIs) = Approximation of the reduction of  $\chi^2$  if a single constrained parameter is freely estimated

Constrained parameters in CFA: - Factor cross-loadings - Residual correlations

-> Test of local fit -> One parameter at a time

```
universalism_fit %>%
  modificationindices(standardized = T, sort. = T)
```

lhs	op	rhs	mi	epc	sepc.lv	sepc.all	sepc.nox
-----	----	-----	----	-----	---------	----------	----------

Accordingly, there are no modification options, since the model has just been identified and no degrees of freedom are available.

So let's look at two more models of Theory of Basic Human Values from the ESS07.

## CFA model of Tradition/Conformity

### Step 1:

Name	Label	Construct
ipmodst	Important to be humble and modest, not draw attention	Tradition
imptrad	Important to follow traditions and customs	Tradition
ipfrule	Important to care for nature and environment	Conformity
ipbhprp	Important to behave properly	Conformity

### Step 2:

Indicators are “effects” of the latent variable and therefore we specify a reflective measurement model.

### Step 3:

We use the reference indicator method again to calculate the model.

### Step 4:

```
ESS07 %>%
  select(ipmodst, imptrad, ipfrule, ipbhprp) %>%
  mvn() %>%
  # optional
  export_table(table_width = 1, digits = 3)
```

Test		HZ	p value	MVN
Henze-Zirkler	20.393		0	NO

Test	Variable	Statistic	p value	Normality
Anderson-Darling	ipmodst	242.6587	<0.001	NO
Anderson-Darling	imptrad	202.8157	<0.001	NO
Anderson-Darling	ipfrule	173.6286	<0.001	NO
Anderson-Darling	ipbhprp	227.9644	<0.001	NO

n	Mean	Std.Dev	Median	Min	Max	25th	75th	Skew	Kurtosis
5844	4.344	1.208	5	1	6	4	5	-0.643	-0.191
5844	4.155	1.380	4	1	6	3	5	-0.529	-0.601
5844	3.700	1.393	4	1	6	2	5	-0.161	-0.984
5844	4.258	1.281	5	1	6	3	5	-0.554	-0.514

The variables are scaled metrically on a scale from 0 to 6 and distributed skewed to the right. In addition, there is neither a univariate nor a multivariate normal distribution.

### Step 5:

```
tradition_conformity_fit <- cfa(
  model = "TraditionConformity =~ ipmodst + imptrad + ipfrule + ipbhprp",
  data = ESS07,
  estimator = "MLR",
  # optional if there are missing values
```

```

missing = "fiml.x"
)

summary(
  object = tradition_conformity_fit,
  fit.measures = T,
  standardized = T,
  rsquare = T
)

```

lavaan 0.6-12 ended normally after 31 iterations

Estimator	ML
Optimization method	NLMINB
Number of model parameters	12
Number of observations	5844
Number of missing patterns	1

Model Test User Model:

	Standard	Robust
Test Statistic	17.439	14.378
Degrees of freedom	2	2
P-value (Chi-square)	0.000	0.001
Scaling correction factor		1.213
Yuan-Bentler correction (Mplus variant)		

Model Test Baseline Model:

Test statistic	2339.880	1848.011
Degrees of freedom	6	6
P-value	0.000	0.000
Scaling correction factor		1.266

User Model versus Baseline Model:

Comparative Fit Index (CFI)	0.993	0.993
Tucker-Lewis Index (TLI)	0.980	0.980
Robust Comparative Fit Index (CFI)		0.994
Robust Tucker-Lewis Index (TLI)		0.981

Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-38377.968	-38377.968
Scaling correction factor		1.002
for the MLR correction		
Loglikelihood unrestricted model (H1)	-38369.248	-38369.248
Scaling correction factor		1.032
for the MLR correction		
Akaike (AIC)	76779.936	76779.936
Bayesian (BIC)	76860.014	76860.014
Sample-size adjusted Bayesian (BIC)	76821.881	76821.881

Root Mean Square Error of Approximation:

RMSEA	0.036	0.033
90 Percent confidence interval - lower	0.022	0.019
90 Percent confidence interval - upper	0.053	0.048
P-value RMSEA <= 0.05	0.909	0.973
Robust RMSEA		0.036
90 Percent confidence interval - lower		0.020
90 Percent confidence interval - upper		0.054

Standardized Root Mean Square Residual:

SRMR	0.011	0.011
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Parameter Estimates:

Standard errors	Sandwich
Information bread	Observed
Observed information based on	Hessian

Latent Variables:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
TraditionConformity =~						
ipmodst	1.000				0.426	0.353
imptrad	1.546	0.095	16.327	0.000	0.659	0.477
ipfrule	1.875	0.112	16.689	0.000	0.799	0.574
ipbhprp	2.084	0.110	18.900	0.000	0.888	0.693

Intercepts:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.ipmodst	4.344	0.016	274.875	0.000	4.344	3.596
.imptrad	4.155	0.018	230.189	0.000	4.155	3.011
.ipfrule	3.700	0.018	203.091	0.000	3.700	2.657
.ipbhprp	4.258	0.017	254.054	0.000	4.258	3.323
TraditnCnfrmt	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.ipmodst	1.278	0.027	48.093	0.000	1.278	0.876
.imptrad	1.470	0.036	41.297	0.000	1.470	0.772
.ipfrule	1.301	0.038	34.612	0.000	1.301	0.671
.ipbhprp	0.853	0.039	21.855	0.000	0.853	0.519
TraditnCnfrmt	0.182	0.018	10.217	0.000	1.000	1.000

R-Square:

	Estimate
ipmodst	0.124
imptrad	0.228
ipfrule	0.329
ipbhprp	0.481

The model is identified with two degrees of freedom.

### Step 6:

The fit indices show a good fit to the data. All fit indices are below their cut offs.

### Step 7:

In this model, the factor loading of the item ipmodst is clearly misfit, since the factor loading cut-off of 0.4 is not reached.

Here it has to be considered whether the item has a value for the construct or whether the item has to be removed from the measurement model. We decide to take it out.

Further modifications are not necessary since the model shows a very good fit to the data.

```
tradition_conformity_fit2 <- cfa(  
  model = "TraditionConformity =~ imptrad + ipfrule + ipbhprp",  
  data = ESS07,  
  estimator = "MLR",  
  # optional if there are missing values  
  missing = "fiml.x"  
)
```

```
summary(  
  object = tradition_conformity_fit2,  
  fit.measures = T,  
  standardized = T,  
  rsquare = T  
)
```

lavaan 0.6-12 ended normally after 28 iterations

Estimator	ML
Optimization method	NLMINB
Number of model parameters	9
Number of observations	5844
Number of missing patterns	1

Model Test User Model:

	Standard	Robust
Test Statistic	0.000	0.000
Degrees of freedom	0	0

Model Test Baseline Model:

Test statistic	1849.694	1427.618
Degrees of freedom	3	3
P-value	0.000	0.000
Scaling correction factor		1.296

User Model versus Baseline Model:

Comparative Fit Index (CFI)	1.000	1.000
Tucker-Lewis Index (TLI)	1.000	1.000
Robust Comparative Fit Index (CFI)		1.000
Robust Tucker-Lewis Index (TLI)		1.000

Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-29217.708	-29217.708
Loglikelihood unrestricted model (H1)	-29217.708	-29217.708
Akaike (AIC)	58453.417	58453.417
Bayesian (BIC)	58513.475	58513.475
Sample-size adjusted Bayesian (BIC)	58484.876	58484.876

Root Mean Square Error of Approximation:

RMSEA	0.000	0.000
90 Percent confidence interval - lower	0.000	0.000
90 Percent confidence interval - upper	0.000	0.000
P-value RMSEA <= 0.05	NA	NA
Robust RMSEA		0.000
90 Percent confidence interval - lower		0.000
90 Percent confidence interval - upper		0.000

Standardized Root Mean Square Residual:

SRMR	0.000	0.000
------	-------	-------

Parameter Estimates:

Standard errors	Sandwich
Information bread	Observed
Observed information based on	Hessian

Latent Variables:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
TraditionConformity =~						
imptrad	1.000				0.673	0.488
ipfrule	1.247	0.059	21.170	0.000	0.839	0.602
ipbhprp	1.252	0.061	20.475	0.000	0.842	0.658

Intercepts:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.imptrad	4.155	0.018	230.189	0.000	4.155	3.011
.ipfrule	3.700	0.018	203.091	0.000	3.700	2.657
.ipbhprp	4.258	0.017	254.054	0.000	4.258	3.323
TraditnCnfrmty	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.imptrad	1.452	0.036	39.882	0.000	1.452	0.762
.ipfrule	1.236	0.043	29.051	0.000	1.236	0.637
.ipbhprp	0.932	0.040	23.156	0.000	0.932	0.568
TraditnCnfrmty	0.453	0.032	14.009	0.000	1.000	1.000

R-Square:

Estimate

imptrad	0.238
ipfrule	0.363
ipbhprp	0.432

This makes the measurement model suitable for further use.

## CFA model Perceived Threat

### Step 1:

Name	Label	Construct
imbgeco	Immigration bad or good for country's economy	Perceived threat
imueclt	Country's cultural life undermined or enriched by immigrants	Perceived threat
imtcjob	Immigrants take jobs away in country or create new jobs	Perceived threat
rlgueim	Religious beliefs and practices undermined or enriched by immigrants	Perceived threat

### Step 2:

Indicators are “effects” of the latent variable and therefore we specify a reflective measurement model

### Step 3:

We use the reference indicator method again to calculate the model.

### Step 4:

```
ESS07 %>%
  select(imbgeco, imueclt, imtcjob, rlgueim) %>%
  mvn() %>%
  # optional for better display
  export_table(table_width = 1, digits = 3)
```

Test		HZ	p value	MVN
Henze-Zirkler		17.085		0   NO

Test		Variable	Statistic	p value	Normality
Anderson-Darling		imbgeco	61.6271	<0.001	NO
Anderson-Darling		imueclt	53.8325	<0.001	NO
Anderson-Darling		imtcjob	103.8354	<0.001	NO
Anderson-Darling		rlgueim	108.6560	<0.001	NO

n	Mean	Std.Dev	Median	Min	Max	25th	75th	Skew	Kurtosis
5844	4.979	2.466	5	0	10	3	7	0.186	-0.521
5844	4.696	2.577	5	0	10	3	6	0.174	-0.623
5844	5.187	2.301	5	0	10	4	7	0.228	-0.165
5844	5.318	2.188	5	0	10	4	7	0.041	0.053

The variables are scaled metrically on a scale from 0 to 10. Nevertheless, there is neither a univariate nor a multivariate normal distribution.



## Step 5:

```
perceived_threat_fit <- cfa(  
  model = "PerceivedThreat =~ imbgeco + imueclt + imtcjob + rlgueim",  
  data = ESS07,  
  estimator = "MLR",  
  # optional if there are missing values  
  missing = "fiml.x"  
)  
  
summary(  
  object = perceived_threat_fit,  
  fit.measures = T,  
  standardized = T,  
  rsquare = T  
)
```

lavaan 0.6-12 ended normally after 27 iterations

Estimator	ML
Optimization method	NLMINB
Number of model parameters	12
Number of observations	5844
Number of missing patterns	1

### Model Test User Model:

	Standard	Robust
Test Statistic	279.766	184.727
Degrees of freedom	2	2
P-value (Chi-square)	0.000	0.000
Scaling correction factor		1.514
Yuan-Bentler correction (Mplus variant)		

### Model Test Baseline Model:

Test statistic	8797.948	5161.721
Degrees of freedom	6	6
P-value	0.000	0.000
Scaling correction factor		1.704

### User Model versus Baseline Model:

Comparative Fit Index (CFI)	0.968	0.965
Tucker-Lewis Index (TLI)	0.905	0.894
Robust Comparative Fit Index (CFI)		0.969
Robust Tucker-Lewis Index (TLI)		0.906

### Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-49161.345	-49161.345
Scaling correction factor		1.214
for the MLR correction		
Loglikelihood unrestricted model (H1)	-49021.462	-49021.462

Scaling correction factor 1.257  
for the MLR correction

Akaike (AIC)	98346.690	98346.690
Bayesian (BIC)	98426.768	98426.768
Sample-size adjusted Bayesian (BIC)	98388.636	98388.636

Root Mean Square Error of Approximation:

RMSEA	0.154	0.125
90 Percent confidence interval - lower	0.139	0.113
90 Percent confidence interval - upper	0.170	0.138
P-value RMSEA <= 0.05	0.000	0.000

Robust RMSEA	0.154
90 Percent confidence interval - lower	0.136
90 Percent confidence interval - upper	0.173

Standardized Root Mean Square Residual:

SRMR	0.026	0.026
------	-------	-------

Parameter Estimates:

Standard errors	Sandwich
Information bread	Observed
Observed information based on	Hessian

Latent Variables:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
PerceivedThreat =~						
imbgeco	1.000				1.967	0.798
imueclt	1.105	0.022	51.221	0.000	2.174	0.844
imtcjob	0.792	0.015	52.565	0.000	1.559	0.677
rlgueim	0.692	0.020	34.946	0.000	1.362	0.622

Intercepts:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.imbgeco	4.979	0.032	154.352	0.000	4.979	2.019
.imueclt	4.696	0.034	139.327	0.000	4.696	1.823
.imtcjob	5.187	0.030	172.337	0.000	5.187	2.254
.rlgueim	5.318	0.029	185.807	0.000	5.318	2.431
PerceivedThret	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.imbgeco	2.211	0.085	26.148	0.000	2.211	0.364
.imueclt	1.913	0.094	20.355	0.000	1.913	0.288
.imtcjob	2.865	0.087	33.077	0.000	2.865	0.541
.rlgueim	2.933	0.080	36.474	0.000	2.933	0.613
PerceivedThret	3.870	0.119	32.528	0.000	1.000	1.000

R-Square:

Estimate

imbgeco	0.636
imueclt	0.712
imtcjob	0.459
rlgueim	0.387

The model is identified with two degrees of freedom.

#### Step 6:

The model shows a clear misspecification of the model for the RMSEA indice.

#### Step 7:

The factor loadings all show sufficient loading. So we can immediately devote ourselves to the modification indices in order to understand the misspecification evident in the RMSEA.

```
perceived_threat_fit %>%
  modificationindices(standardized = T, sort. = T)
```

	lhs	op	rhs	mi	epc	sepc.lv	sepc.all	sepc.nox
19	imueclt	~~	rlgueim	238.993104	0.8771423	0.8771423	0.3702775	0.3702775
16	imbgeco	~~	imtcjob	238.993083	0.9085458	0.9085458	0.3609752	0.3609752
17	imbgeco	~~	rlgueim	207.424644	-0.7466824	-0.7466824	-0.2931758	-0.2931758
18	imueclt	~~	imtcjob	207.424451	-0.9445400	-0.9445400	-0.4034743	-0.4034743
20	imtcjob	~~	rlgueim	1.848284	-0.0628233	-0.0628233	-0.0216717	-0.0216717
15	imbgeco	~~	imueclt	1.848267	-0.1265846	-0.1265846	-0.0615456	-0.0615456

The modification indices tell us that adjusting the residual covariance between items imueclt and rlgueim would have the largest effect on model fitting (in terms of the chi-square test statistic).

This is because some of the covariation of imueclt and rlgueim is due to sources other than the common factor.

Reasons: method effects, item wording, acquiescence, social desirability, unaccounted theoretical common causes.

We then modify our initial model and remove the residual covariance constraint between the two items, which is fixed at 0, and let it calculate freely.

```
perceived_threat_fit2 <- cfa(
  model = "
    PerceivedThreat =~ imbgeco + imueclt + imtcjob + rlgueim

    imueclt    ~~    rlgueim
  ",
  data = ESS07,
  estimator = "MLR",
  # optional if there are missing values
  missing = "fiml.x"
)

summary(
  object = perceived_threat_fit2,
  fit.measures = T,
  standardized = T,
  rsquare = T
)
```

lavaan 0.6-12 ended normally after 31 iterations

Estimator	ML
Optimization method	NLMINB
Number of model parameters	13
Number of observations	5844
Number of missing patterns	1

Model Test User Model:

	Standard	Robust
Test Statistic	47.224	31.440
Degrees of freedom	1	1
P-value (Chi-square)	0.000	0.000
Scaling correction factor		1.502
Yuan-Bentler correction (Mplus variant)		

Model Test Baseline Model:

Test statistic	8797.948	5161.721
Degrees of freedom	6	6
P-value	0.000	0.000
Scaling correction factor		1.704

User Model versus Baseline Model:

Comparative Fit Index (CFI)	0.995	0.994
Tucker-Lewis Index (TLI)	0.968	0.965
Robust Comparative Fit Index (CFI)		0.995
Robust Tucker-Lewis Index (TLI)		0.969

Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-49045.074	-49045.074
Scaling correction factor		1.238
for the MLR correction		
Loglikelihood unrestricted model (H1)	-49021.462	-49021.462
Scaling correction factor		1.257
for the MLR correction		
Akaike (AIC)	98116.148	98116.148
Bayesian (BIC)	98202.900	98202.900
Sample-size adjusted Bayesian (BIC)	98161.589	98161.589

Root Mean Square Error of Approximation:

RMSEA	0.089	0.072
90 Percent confidence interval - lower	0.068	0.055
90 Percent confidence interval - upper	0.111	0.091
P-value RMSEA <= 0.05	0.001	0.016
Robust RMSEA		0.088
90 Percent confidence interval - lower		0.064

90 Percent confidence interval - upper 0.116

Standardized Root Mean Square Residual:

SRMR 0.012 0.012

Parameter Estimates:

Standard errors Sandwich  
Information bread Observed  
Observed information based on Hessian

Latent Variables:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
PerceivedThreat =~						
imbgeco	1.000				2.100	0.851
imueclt	0.963	0.019	50.389	0.000	2.022	0.785
imtcjob	0.760	0.017	46.031	0.000	1.596	0.694
rlgueim	0.559	0.018	30.307	0.000	1.174	0.537

Covariances:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.imueclt ~~						
.rlgueim	0.848	0.072	11.715	0.000	0.848	0.287

Intercepts:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.imbgeco	4.979	0.032	154.352	0.000	4.979	2.019
.imueclt	4.696	0.034	139.327	0.000	4.696	1.823
.imtcjob	5.187	0.030	172.337	0.000	5.187	2.254
.rlgueim	5.318	0.029	185.807	0.000	5.318	2.431
PerceivedThret	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
.imbgeco	1.673	0.091	18.481	0.000	1.673	0.275
.imueclt	2.551	0.097	26.176	0.000	2.551	0.384
.imtcjob	2.747	0.084	32.790	0.000	2.747	0.519
.rlgueim	3.409	0.086	39.542	0.000	3.409	0.712
PerceivedThret	4.408	0.126	34.979	0.000	1.000	1.000

R-Square:

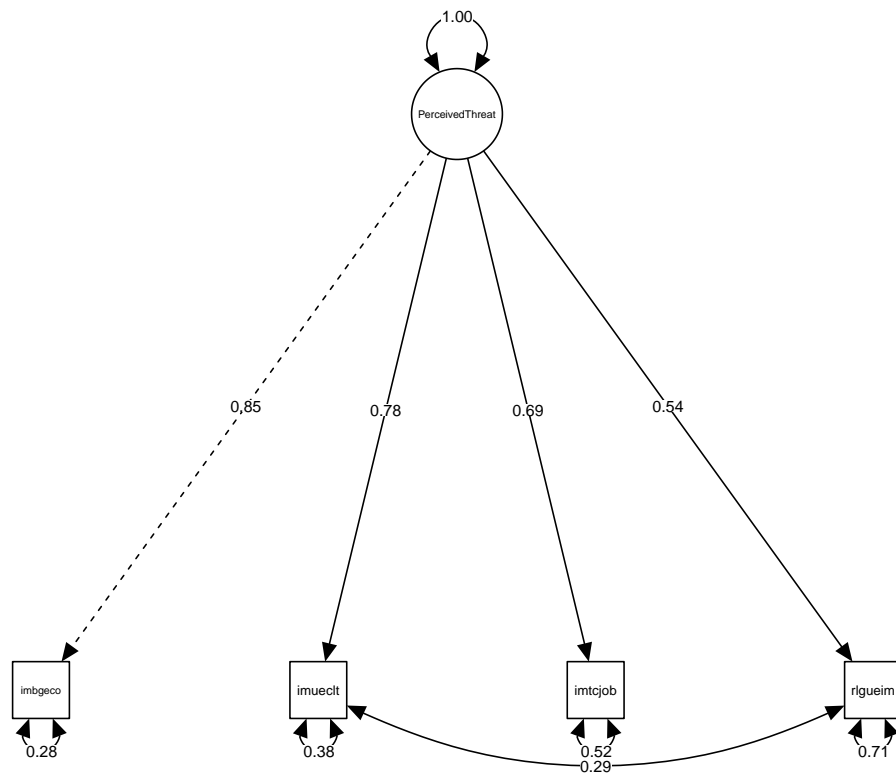
	Estimate
imbgeco	0.725
imueclt	0.616
imtcjob	0.481
rlgueim	0.288

```
perceived_threat_fit2 %>%
  semPaths(
    whatLabels = "std",
    style = "mx",
    layout = "tree",
    edge.color = "black",
```

```

intercepts = F,
nCharNodes = 50
)

```



The RMSEA is then back in our anticipated cut-off area and we can continue working with the measurement model.

### Step 8:

Finally, we want to display the mean structure of our model.

In practice, the only reason why a user would add intercept-formulas in the model syntax is because some constraints must be specified on them.

```

perceived_threat_mean <- cfa(
  model = "
    PerceivedThreat =~ imbgeco + imueclt + imtcjob + rlgueim

    imueclt    ~~    rlgueim
  ",
  data = ESS07,
  estimator = "MLR",
  # optional if there are missing values
  missing = "fiml.x",
  meanstructure = T
)

```

```

summary(
  object = perceived_threat_mean,
  fit.measures = T,
  standardized = T,

```

```
rsquare = T
)
```

lavaan 0.6-12 ended normally after 31 iterations

Estimator	ML
Optimization method	NLMINB
Number of model parameters	13
Number of observations	5844
Number of missing patterns	1

Model Test User Model:

	Standard	Robust
Test Statistic	47.224	31.440
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Degrees of freedom	6	6
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Bayesian (BIC)	98202.900	98202.900
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90 Percent confidence interval - upper	0.111	0.091
P-value RMSEA <= 0.05	0.001	0.016

Robust RMSEA	0.088
90 Percent confidence interval - lower	0.064
90 Percent confidence interval - upper	0.116

Standardized Root Mean Square Residual:

SRMR	0.012	0.012
------	-------	-------

Parameter Estimates:

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Information bread	Observed
Observed information based on	Hessian

Latent Variables:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
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Covariances:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
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.rlgueim	0.848	0.072	11.715	0.000	0.848	0.287

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	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
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.rlgueim	5.318	0.029	185.807	0.000	5.318	2.431
PerceivedThret	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z )	Std.lv	Std.all
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.imueclt	2.551	0.097	26.176	0.000	2.551	0.384
.imtcjob	2.747	0.084	32.790	0.000	2.747	0.519
.rlgueim	3.409	0.086	39.542	0.000	3.409	0.712
PerceivedThret	4.408	0.126	34.979	0.000	1.000	1.000

R-Square:

	Estimate
imbgeco	0.725
imueclt	0.616
imtcjob	0.481
rlgueim	0.288



# Measurement Invariance

## Procedure

Are measurements actually comparable across groups/time?

Definition: Measurement invariance is a property of a measurement instrument, implying that the instrument measures the same concept in the same way across subgroups of respondents or time points of data assessment.

Measurement invariance (MI) can be assessed with multiple-group CFA (MGCFA). Impose equality constraints on measurement parameters across groups/time (“fix them to be equal”).

Use a sequential testing strategy to test different levels of measurement invariance: - Configural MI - Metric (weak) MI - Scalar (strong) MI - Strict MI - ...

### Configural measurement invariance

Equality constraints: Equal structure: pattern of factors and relationships of indicators and factors is equal across groups

Implications: Same construct exists across groups; no comparisons of estimates!

### Metric measurement invariance

Equality constraints: Equal structure + equal factor loadings across groups

$$\lambda^1 = \lambda^2 = \lambda^G$$

### Scalar measurement invariance

Equality constraints: Equal structure + equal factor loadings + equal intercepts across groups

$$\lambda^1 = \lambda^2 = \lambda^G$$

$$\tau^1 = \tau^2 = \tau^G$$

Implications: Respondents from different groups “use” the scale in a similar manner; comparison of latent means are valid

## Result

- Configural invariance: same structure across groups
- Metric invariance: factor loadings set equal across groups
- Scalar invariance: factor loadings and intercepts set equal across groups

How decide which level of measurement invariance is supported by the data?

1. Consecutively estimate the configural, metric, and scalar invariance models
2. Inspect model fit as usual:  $\chi^2$  (df), CFI, RMSEA
3. Inspect model fit differences between models:
  - $\chi^2/\Delta\chi^2(df/\Delta df)$
  - $CFI/\Delta CFI; RMSEA/\Delta RMSEA$

## References

Novick, M. R. (1966). The axioms and principal results of classical test theory. *Journal of Mathematical Psychology*, 3(1), 1–18. Retrieved February 20, 2023, from <https://linkinghub.elsevier.com/retrieve/pii/0022249666900022>