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Stat HW1

8.4a. If $B(\hat{\theta}) = 0$, then $MSE(\hat{\theta}) = V(\hat{\theta})$

b. $MSE(\hat{\theta}) > V(\hat{\theta})$ because

$$MSE = V + B$$

8.6a $\hat{\theta}_3 = a\hat{\theta}_1 + (1-a)\hat{\theta}_2$

$$E(\hat{\theta}_3) = a E(\hat{\theta}_1) + (1-a) E(\hat{\theta}_2)$$

$$= a\theta + (1-a)\theta$$

$$= a\theta + \theta - a\theta$$

$$= \theta$$

So, $\hat{\theta}_3$ is unbiased estimator of θ

8.8a $\hat{\theta}_1 = Y_1, E(\hat{\theta}_1) = E(Y_1) = \theta \rightarrow$ unbiased

$\hat{\theta}_2 = \frac{Y_1 + Y_2}{2}, E(\hat{\theta}_2) = 2\theta/2 = \theta \rightarrow$ unbiased

$E(\hat{\theta}_3) = 3\theta/3 \rightarrow$ unbiased

$E(\hat{\theta}_4) = \min(\theta, \theta, \theta) = \theta \rightarrow$ unbiased

$E(\hat{\theta}_5) = E((Y_1 + Y_2 + Y_3)/3) = 3\theta/3 = \theta \rightarrow$ unbiased

b. $V(\text{exp}) = \theta^2$

$V(\hat{\theta}_1) = V(Y_1) = \theta^2$

$\dots = V\left(\frac{Y_1 + Y_2}{2}\right) = \frac{1}{4}(V(Y_1) + V(Y_2)) = \frac{\theta^2}{2}$

$\dots = V\left(\frac{Y_1 + 2Y_2}{3}\right) = \frac{1}{9}\theta^2 + \frac{4}{9}\theta^2 = \frac{5\theta^2}{9}$

$\dots = \min(\theta^2, \theta^2, \theta^2) = \theta^2$
 $= \frac{1}{4}V(Y_1) + \frac{1}{4}\theta^2 + \frac{1}{4}\theta^2 = \boxed{\frac{1}{3}\theta^2}$

8.12a $f(y) = 1$ $\theta \leq y \leq \theta + 1$

$$E(\bar{Y}) = \theta + \frac{1}{2}$$

$$E(\bar{Y}) = E\left(\frac{Y_1 + Y_2 + \dots + Y_n}{n}\right)$$

$$E(\bar{Y}) \neq \theta, \text{ so biased.}$$

$$= \frac{1}{n} E(Y_1) + \dots$$

$$b(\theta) = \theta + \frac{1}{2} - \theta = \frac{1}{2}$$

$$= \sum_{i=1}^n \frac{1}{n} (\theta + \frac{1}{2}) = \theta + \frac{1}{2}$$

$$b. \frac{1}{n} E(\bar{Y} - \frac{1}{2}) = E(\bar{Y}) - E(\frac{1}{2}) = \theta$$

$$c. V(\bar{Y}) = \sum_{i=1}^n \frac{1}{n^2} V(Y_i) = \frac{1}{n^2} V(Y_1) + \frac{1}{n^2} V(Y_2) + \dots = \frac{1}{2n}$$

$$V(\bar{Y}) = 1/12 \quad \text{MSE}(\bar{Y}) = \frac{1}{12n} + \left(\frac{1}{2}\right)^2 = \frac{1+3n}{12n}$$

8.18 $f(y) = \frac{1}{\theta} \quad 0 \leq y \leq \theta. \quad F(y) = \frac{y}{\theta}$

$$Y_{(n)} \approx f_{Y_{(n)}}(y) = n \left(1 - \frac{y}{\theta}\right)^{n-1}$$

~~$$E(Y_{(n)}) = \int_0^\theta \frac{n}{\theta} \left(1 - \frac{y}{\theta}\right)^{n-1} dy = \frac{n}{\theta} \left[\frac{(1 - \frac{y}{\theta})^n}{-1} - \frac{1}{-1} \right]_0^\theta$$~~

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8.18 continued

$$\begin{aligned}
 E(Y_{(n)}) &= \frac{n}{\theta} \int_0^{\theta} y \left(1 - \frac{y}{\theta}\right)^{n-1} dy + \frac{u}{\theta} \int_0^{\theta} y \left(1 - \frac{y}{\theta}\right)^{n-1} dy \\
 &= \left(-\frac{y \theta \left(1 - \frac{y}{\theta}\right)^n}{n} - \frac{\theta^2 \left(1 - \frac{y}{\theta}\right)^{n+1}}{n(n+1)} \right) \Big|_0^{\theta} - \theta \left(1 - \frac{y}{\theta}\right)^n \Big|_0^{\theta} \\
 &= \left(-\frac{y \left(1 - \frac{y}{\theta}\right)^n}{n} - \frac{\theta \left(1 - \frac{y}{\theta}\right)^{n+1}}{n(n+1)} \right) \Big|_0^{\theta} - \frac{\theta^2 \left(1 - \frac{y}{\theta}\right)^{n+1}}{n(n+1)} \Big|_0^{\theta} \\
 &= (0 - 0) - \left(0 - \frac{\theta}{n+1}\right) = \boxed{\frac{\theta}{n+1}}
 \end{aligned}$$

unbiased = $\frac{\theta}{n+1} \cdot (n+1)$, so $Y_{(n)}(n+1)$ is unbiased

9.2a. $E(\hat{\mu}_1) = \frac{1}{2}(u + u) = u \quad \checkmark$

$E(\hat{\mu}_2) = \frac{1}{4}u + \frac{(n-2)u}{n-2} + \frac{1}{4}u = u \quad \checkmark$

$E(\hat{\mu}_3) = \frac{n \cdot u}{n} = u \quad \checkmark$

b. $V(\hat{\mu}_1) = \frac{1}{4}(\sigma^2 + \sigma^2) = \frac{\sigma^2}{2}$

$V(\hat{\mu}_2) = \frac{1}{16}\sigma^2 + \frac{(n-2)\sigma^2}{4(n-2)^2} + \frac{1}{16}\sigma^2 = \frac{2\sigma^2}{8(n-2)} = \frac{n \cdot \sigma^2}{8(n-2)}$

$V(\hat{\mu}_3) = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$

$Eff(\hat{\mu}_3, \hat{\mu}_1) = \frac{\sigma^2/2}{\sigma^2/n} = n/2$

$Eff(\hat{\mu}_3, \hat{\mu}_2) = \frac{n \cdot \sigma^2}{8(n-2)} / \frac{\sigma^2}{n} = \frac{n^2}{8(n-2)}$

9.3a from 8.12: $E(\bar{Y}) = \theta + \frac{1}{2}$

$$E(\hat{\theta}_1) = \theta + \frac{1}{2} - \frac{1}{2} = \theta \quad \checkmark$$

$$f_{Y_n}^*(y) = n(y-\theta)^{n-1}$$

$$E(Y_{(n)}) = \int_0^{\theta+1} n y (y-\theta)^{n-1} dy \quad \begin{array}{l} u = y - \theta \quad y = u + \theta \\ du = dy \end{array}$$

$$= n \int_0^1 (u+\theta) u^{n-1} du = n \int_0^1 u^n + \theta u^{n-1} du$$

$$= n \left(\frac{u^{n+1}}{n+1} + \frac{\theta u^n}{n} \right) \Big|_0^1 = \frac{n}{n+1} + \theta$$

So,

$$E(\hat{\theta}_2) = E(Y_{(n)}) - \frac{n}{n+1} = \theta \quad \checkmark$$

b. From 8.12: $V(\bar{Y}) = \frac{1}{12n}$

$$E(Y_{(n)}^2) = n \int_0^{\theta+1} y^2 (y-\theta)^{n-1} dy = n \int_0^1 (u+\theta)^2 u^{n-1} du$$

$$= n \int_0^1 (u^2 + 2u\theta + \theta^2) u^{n-1} du = n \int_0^1 u^{n+1} + 2u^n \theta + \theta^2 u^{n-1} du$$

$$= n \left(\frac{u^{n+2}}{n+2} + \frac{2u^{n+1}}{n+1} \theta + \frac{\theta^2 u^n}{n} \right) \Big|_0^1 = \frac{n}{n+2} + \theta^2 + \frac{2n\theta}{n+1}$$

$$V(Y_{(n)}) = \frac{n}{n+2} + \theta^2 + \frac{2n\theta}{n+1} - \left(\left(\frac{n}{n+1} \right)^2 + \theta^2 + \frac{2n\theta}{n+1} \right)$$

$$= \frac{n}{n+2} - \frac{n^2}{(n+1)^2} = \frac{n}{(n+2)(n+1)^2}$$

$$V(\hat{\theta}_2) = \frac{n}{(n+2)(n+1)^2} + 0 \quad \text{Eff}(\hat{\theta}_1, \hat{\theta}_2) = \frac{12n^2}{(n+2)(n+1)^2}$$

9.7 $f(y) = \frac{1}{\theta} e^{-y/\theta}$, $E(Y) = \theta$, $V(Y) = \theta^2$

$\hat{\theta}_1 = nY_{(1)}$ $MSE(\hat{\theta}_1) = \theta^2$

$\hat{\theta}_2 = \bar{Y}$ $V(\hat{\theta}_2) = V\left(\frac{Y_1 + Y_2 + \dots + Y_n}{n}\right)$
 $= \frac{1}{n^2} V(Y_1 + Y_2 + \dots + Y_n) = \frac{1}{n^2} n \theta^2 = \frac{\theta^2}{n}$

Since $\hat{\theta}_1$ is unbiased, $b(\hat{\theta}_1) = 0$

$MSE(\hat{\theta}_1) = V(\hat{\theta}_1)$

$E(Y) = E(Y_1 + Y_2 + \dots + Y_n) \frac{1}{n} = \frac{n\theta}{n} = \theta$

So, $\hat{\theta}_2$ is unbiased, $b(\hat{\theta}_2) = 0$ oops, this is not necessary

~~MSE~~ $MSE(\hat{\theta}_2) = V(\hat{\theta}_2)$

$Eff(\hat{\theta}_1, \hat{\theta}_2) = \frac{V(\hat{\theta}_2)}{V(\hat{\theta}_1)} = \frac{\theta^2}{n} \cdot \frac{1}{MSE(\hat{\theta}_1)} = \frac{\theta^2}{n} \cdot \frac{1}{\theta^2} = \boxed{\frac{1}{n}}$