Homework 2

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Due date: Thursday, October 11

1. Show that (algebraically) in the absence of censoring $\hat{S}_{\text{KM}}(t) = \hat{S}_{\mathbf{e}}(t)$.

Without loss of generality, we assume $t_1 < t_2 < \ldots < t_n$ and $\Delta_i = 1$ for $i = 1, \ldots, n$. Under the assumption of no ties, we have $d_1 = \ldots = d_n = 1$ and $n_1 = n, n_2 = n - 1, \ldots, n_i = n - i - 1$.

$$\hat{S}_{km}(t) = \prod_{i:t_i \le t} \left(1 - \frac{d_i}{n_i} \right) = \prod_{i:t_i \le t} \left(\frac{n_i - d_i}{n_i} \right)$$
$$= \frac{n-1}{n} \times \frac{n-2}{n-1} \times \dots \times \frac{n-i}{n-i-1}$$
$$= \frac{n-i}{n},$$

which is the expression for the empirical survival estimator, $\hat{S}_{e}(t)$.

2. In the absence of censoring, show that the Greenwood Formula (page 30 on note 2) can be reduced to

$$\frac{\hat{S}_{\mathbf{KM}}(t) \times \{1 - \hat{S}_{\mathbf{KM}}(t)\}}{n}.$$

You might assume there are no ties among the observations.

Continue with the assumption outlined in #1, we can simplify the Greenwood formula as follows.

$$\begin{split} \hat{S}_{km}^{2}(t) \cdot \sum_{i:t_{i} \leq t} \frac{d_{i}}{n_{i}(n_{i} - d_{i})} &= \left(\frac{n - i}{n}\right)^{2} \cdot \left(\frac{1}{n(n - 1)} - \frac{1}{(n - 1)(n - 2)} - \dots - \frac{1}{(n - i - 1)(n - i)}\right) \\ &= \left(\frac{n - i}{n}\right)^{2} \cdot \left(\frac{1}{n - i} - \frac{1}{n}\right) \\ &= \left(\frac{n - i}{n}\right)^{2} \cdot \frac{i}{n(n - i)} = \frac{i(n - i)}{n^{3}} = \frac{\hat{S}_{\text{KM}}(t) \times \{1 - \hat{S}_{\text{KM}}(t)\}}{n}. \end{split}$$

3. Consider the Leukemia data from the survival package:

```
> library(survival)
> head(aml)
```

```
time status
1
     9
            1 Maintained
2
    13
            1 Maintained
3
    13
            0 Maintained
4
    18
            1 Maintained
5
    23
            1 Maintained
6
    28
            0 Maintained
```

In here, each row represent one patient. aml is the observed survival time, status is the censoring indicator (1 = event, 0 = censored), and x is the treatment indicator. We will ignore the treatment indicator for now.

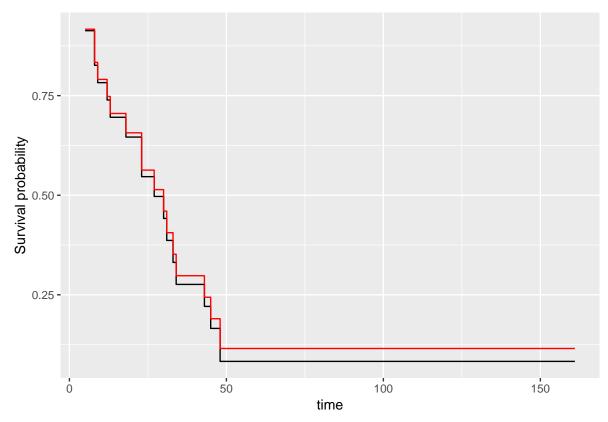
- a. Plot the Kaplan-Meier survival curve for the data.
- b. Add the Nelson-Aalen survival curve to the Kaplan-Meier plot from (3a).

Load libraries:

```
> library(tidyverse)
> library(survival)
```

The following gives both the Kaplan-Meier survival curve (black) and the Nelson-Aalen survival curve (red).

```
> dat <- aml %>% arrange(time) %>% select(-x) %>% group_by(time) %>%
+ summarize(di = sum(status), ni = length(status)) %>%
+ mutate(ni = rev(cumsum(rev(ni))), KM = cumprod(1 - di / ni), Na = cumsum(di / ni))
> ggplot(data = dat, aes(x = time)) +
+ geom_step(aes(y = KM), show.legend = TRUE) +
+ geom_step(aes(y = exp(-Na)), col = 2) + ylab("Survival probability")
```



4. The expected survival time for the Leukemia data in #(3) does not exist because the last observation is a censored event. An alternative is to lookInstead of looking at the expected survival time, an alternative is to look at the restricted mean survival time. Compute E(T|T<161) based on the survival curve in (3a).

```
> sum(diff(c(0, dat$time)) * c(1, dat$KM[-18]))
```

[1] 36.36439

- 5. Let $N_i(t)$ be the number of events over time interval (0,t] for the ith patient in #(3). Let $N(t) = \sum_{i=1}^{n} N_i(t)$ be the aggregated counting process.
 - a. Plot N(t).
 - b. Plot M(t), where $M(t) = N(t) \hat{H}(t)$ and $\hat{H}(t)$ is the Nelson-Aalen estimator for the cumulative hazard function.

Omit.