Exam 1

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Instructions

- Due date: Tuesday, November 13.
- This is a open resource exam, but you are not allowed to ask post exam questions online.
- You are not allowed to collaborate with classmates and/or people outside of class.
- Please circle or highlight your final answer.
- The toal possible point is 80.

Violation of this agreement will result in an **F** on this exam and it will be averaged in as a 0%.

1. Use the complete WHAS100 dataset and gender as the group indicator, compute the log-rank statistic, Q, presented as Equation (9) on page 60 of note 2 with $\omega_i = 1$. Use Q to compute a p-value to test the null hypothesis of $H_o: S_0(t) = S_1(t)$. Do this without a software package (6 pts), and verify the calculation with survdiff (4 pts).

Solution:

We will first load the tidyverse package (but not the survival package):

```
> library(tidyverse)
```

The following codes calculates the necessary quantities to calculate Q:

Compute Q:

```
> with(dat, sum(di.1 - E)^2 / sum(V))
```

[1] 3.971377

- 2. There are many ways to form a basis for survival curve comparison. Here are some:
 - a. The numerator term in Q without the square:

$$D_1 = \sum_{i=1}^{D} \{d_{1i} - E(d_{1i})\}.$$

b. The largest distance between the two curves:

$$D_2 = \max |S_1(t) - S_0(t)|.$$

c. The difference between the median survival times:

$$D_3 = S_1^{-1}(0.5) - S_0^{-1}(0.5)$$

d. The difference between the mean survival times:

$$D_4 = \int_0^{t_{(n)}} \{ S_1(u) - S_0(u) \} du,$$

where $t_{(n)}$ is the maximum observed survival time. Compute each of the above statistic for the WHAS100 dataset (5 pts $\times 4$).

Solution:

Taking advantage of the dat created in #1, these quantity can be easily computed with the following codes:

- 3. The statistics computed in (2) do not provide meaningful interpretations when standing along. We will use a permutation approach to test for the null hypothesis of $H_o: S_0(t) = S_1(t)$ based on these statistics. The idea of a permutation test is simple. The general procedure can be summarized into the following steps:
 - i. Compute the desired statistic based on the observed data; we will call this the observed statistic.
 - ii. Permute the data under the null.
 - iii. Compute the statistics for each possible permutation in Step ii.; we will call these permutation statistics.
 - iv. Draw conclusion based on where the observed statistic stands among the permutation statistics.

The statistics we computed in (2) are the observed statistics in Step i. If the null hypothesis of $H_o: S_0(t) = S_1(t)$ is true, then one can randomly shuffle the group indicator to generate different permutations (Step ii) and the statistics for these permutations should be similar (Step iii).

a. (5 pts×4) Generate 5000 permutation and, for each of the permutation, compute the four statistics presented in (2). We will call the permutated statistics D_{1i}^* , D_{2i}^* , D_{3i}^* and D_{4i}^* for $i=1,\ldots,5000$. Create a histgram for these permutated statistics and print the summary.

b. (5 pts \times 4) Compute the *p*-value based on these statistics by

$$p = 2 \cdot \frac{\min(N_1, N_2)}{5000},$$

where $N_1 = \#\{D \ge D^*\}$, $N_2 = \#\{D \le D^*\}$, and # means the "number of", e.g., N_1 is the number of these permutated statistics less than or equal to the observed statistic.

Solution:

We first prepare a function that returns the 4 statistics in #2:

```
> #' @param obs is the observed survival time
> #' @param event is the censoring indicator
> #' @param x is the categorical covariate; taking values of 0 and 1
> getD <- function(obs, event, x) {
      dat0 <- tibble(Time = unique(sort(obs[event > 0]))) %>%
          mutate(ni = colSums(outer(obs, Time, ">=")),
                 di = colSums(outer(obs, Time, "==")),
                 ni.1 = colSums(outer(obs[x > 0], Time, ">=")),
                 di.1 = colSums(outer(obs[x > 0], Time, "==")),
                 ni.0 = ni - ni.1, di.0 = di - di.1,
                 E = ni.1 * di / ni, V = ni.1 * ni.0 * di * (ni - di) / ni / ni / (ni - 1),
                 km0 = pmax(0, cumprod((ni.0 - di.0) / ni.0), na.rm = TRUE),
                 km1 = pmax(0, cumprod((ni.1 - di.1) / ni.1), na.rm = TRUE))
      attach(dat0)
      d1 <- sum(di.1 - E)
      d2 \leftarrow \max(abs(km0 - km1))
      d3 <- Time[which.max(km1 <= .5)] - Time[which.max(km0 <= .5)]
      d4 \leftarrow sum(-diff(rev(Time)) * rev(km1 - km0)[-1])
      detach(dat0)
      c(d1 = d1, d2 = d2, d3 = d3, d4 = d4)
+ }
```

Check the function with whas 100 to confirm it returns values in #2:

```
> (d0 <- with(whas100, getD(lenfol, fstat, gender)))
```

```
d1 d2 d3 d4
6.6199997 0.3853606 -818.0000000 -428.7407041
```

a. The following codes perform the permutation statistics based on 5000 permutation:

```
> set.seed(123)
> system.time(
+ permd <- t(replicate(5000, with(whas100, getD(lenfol, fstat, sample(gender)))))
+ )</pre>
```

```
user system elapsed 20.632 0.020 20.674
```

> summary(permd)

```
d1
                            d2
                                              d3
                             :0.1439
                                                :-2195.00
Min.
        :-11.78196
                      Min.
                                        Min.
 1st Qu.: -2.25294
                      1st Qu.:0.2229
                                        1st Qu.: -395.00
Median : 0.05511
                      Median :0.3154
                                        Median : -136.00
       : 0.02653
                             :0.3189
                                                : -81.82
                      Mean
                                        Mean
3rd Qu.: 2.31044
                      3rd Qu.:0.3931
                                        3rd Qu.: 390.00
                             :0.7143
                                               : 2195.00
Max.
        : 11.97221
                      Max.
                                        Max.
       d4
Min.
        :-763.848
 1st Qu.:-154.421
Median : -4.830
       : -4.286
Mean
 3rd Qu.: 146.719
Max.
        : 743.986
> ggdat <- gather(as.tibble(permd), "d", factor_key = TRUE) %>% mutate(d0 = unlist(d0[d]))
> ggplot(ggdat, aes(value)) +
      geom_histogram(alpha = .5, bins = 20) + geom_vline(aes(xintercept = d0), lty = I(2)) +
      facet_wrap(d ~., scale = "free")
                                                                        d2
                          d1
                                                600 -
  600 -
                                                400 -
  400 -
                                                200 -
  200 -
    0 -
                                                  0 -
                 -5
                         0
                                                         0.2
                                        10
                                                                                 0.6
                                 5
                                                                     0.4
          -10
                          d3
                                                                        d4
                                                600 -
  900 -
                                                400 -
  600 -
  300 -
                                                200 -
                                                  0 -
                                                                         Ö
                                                                                   500
                          Ö
       -2000
                -1000
                                 1000
                                          2000
                                                            -500
                                               value
                                                                                              b.
The p-values are
> sapply(1:4, function(x) 2 * min(sum(d0[[x]] < permd[,x]), sum(d0[[x]] > permd[,x])) / 5000)
```

[1] 0.0460 0.5532 0.1748 0.0532

4. Another method to compare two survival curves is to consider a sign test. Suppose we have two groups of uncensored survival times:

```
Males: x_1, x_2, ..., x_{n_0}.
Females: y_1, y_2, ..., y_{n_1}.
```

The sign test looks at the statistic

$$U = \sum_{i=1}^{n_0} \sum_{j=1}^{n_1} \operatorname{sgn}(x_i - y_j),$$

where $sgn(\cdot)$ is the sign function. In the pretense of right censoring, survival times can not be compared directly and a modified version of $U = \sum_{i=1}^{n_0} \sum_{j=1}^{n_1} U_{ij}$ is considered, where

$$U_{ij} = \begin{cases} 1 & \text{if } x_i > y_j, y_j \text{ is uncersored.} \\ -1 & \text{if } x_i < y_j, x_i \text{ is uncersored.} \\ 0 & \text{otherwise.} \end{cases}$$

- a. (5 pts) Compute U for the WHAS100 dataset.
- b. (5 pts) Create a histgram for these permutated statistics and print the summary. Obtain a permutation *p*-value based on 5000 permutations.

Solution:

We will take a similar approach as in #3; first write a function to compute U then use it for permutation. The following function can be used to compute U:

```
> getU <- function(obs, event, x) {
+    t1 <- obs[x > 0]
+    d1 <- event[x > 0]
+    t0 <- obs[x == 0]
+    d0 <- event[x == 0]
+    sum(t(outer(t0, t1, ">")) * d1) - sum(outer(t0, t1, "<") * d0)
+ }</pre>
```

a. For the whas100 data:

```
> (U0 <- with(whas100, getU(lenfol, fstat, gender)))</pre>
```

[1] 459

b. Perform permutation test, calculate p-value and produce histogram:

```
> set.seed(123)
> permU <- replicate(5000, with(whas100, getU(lenfol, fstat, sample(gender))))
> summary(permU)
```

```
Min. 1st Qu. Median Mean 3rd Qu. Max.

-916.000 -166.000 2.000 3.146 177.000 896.000

> 2 * min(sum(U0 < permU), sum(U0 > permU)) / 5000
```

[1] 0.0644

```
> ggplot() + aes(permU) + geom_histogram(alpha = .5, bins = 20) +
    geom_vline(aes(xintercept = U0), lty = I(2))
```

