

Mathematical Analysis (New)

1 DC-ACs Model

→ DCs release biomarker cells & they are received by ACs

No of received cells in n^{th} time slot at m^{th} AC is

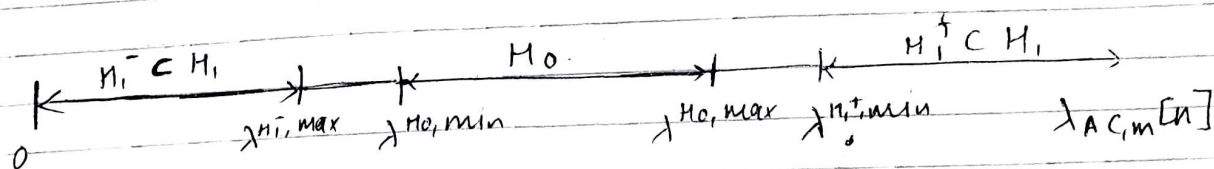
$$y_{AC,m}[n] = \text{Pois}(\text{pr} \cdot \lambda_{AC,m}[n] + \text{pr} \cdot \bar{\eta}_m)$$

\swarrow probability of reception
 \downarrow avg of no. of biomarker molecules received at m^{th} AC, n^{th} time slot
 \searrow mean of Poisson Noise

→ w.k.t,

$$P(y_{AC,m}; \lambda_{AC,m}) = \prod_{n=1}^N \frac{\exp(-\lambda_{AC,m}[n] - \text{pr} \cdot \bar{\eta}_m)}{(\lambda_{AC,m}[n] + \text{pr} \cdot \bar{\eta}_m)^{y_{AC,m}[n]}} \cdot \frac{y_{AC,m}[n]!}{y_{AC,m}[n]!}$$

→ we consider,



$\therefore \forall n,$

$$H_0: \lambda_{AC,m}^{H_0, \min} \leq \lambda_{AC,m}[n] \leq \lambda_{AC,m}^{H_0, \max}$$

$$H_1: \begin{cases} H_1^-: \lambda_{AC,m}[n] \leq \lambda_{AC,m}^{H_1^-, \max} \\ H_1^+: \lambda_{AC,m}[n] \geq \lambda_{AC,m}^{H_1^+, \min} \end{cases}$$

the likelihood ratio of hypothesis test based on GLRT can be given as,

$$\Lambda_{AC,m} = \frac{P(y_{AC,m}; H_1)}{P(y_{AC,m}; H_0)} \underset{H_1}{>} \gamma$$

$$= \frac{N}{\prod_{n=1}^N} \exp(-\lambda_{AC,m}^{H_1} [u] - p_{\beta 1} \bar{u}) (\lambda_{AC,m}^{H_1} [u] + p_{\beta 1} \bar{u})$$

$$\Lambda_{AC,m} = \frac{N}{\prod_{n=1}^N} \exp(-\lambda_{AC,m}^{H_1} [u] - p_{\beta 1} \bar{u}) (\lambda_{AC,m}^{H_1} [u] + p_{\beta 1} \bar{u}) \underset{H_1}{>} \gamma$$

$$\frac{N}{\prod_{n=1}^N} \exp(-\lambda_{AC,m}^{H_0} [u] - p_{\beta 1} \bar{u}) (\lambda_{AC,m}^{H_0} [u] + p_{\beta 1} \bar{u}) y_{AC,m} [u]$$

\therefore The simplified decision rule deduced by bounding likelihood ratio,

$$H_0: \gamma_L \leq \frac{1}{N} \sum_{n=1}^N y_{AC,m} [u] \leq \gamma_u$$

$$H_1: \begin{cases} \frac{1}{N} \sum_{n=1}^N y_{AC,m} [u] > \gamma_u & \text{or} \\ \frac{1}{N} \sum_{n=1}^N y_{AC,m} [u] < \gamma_L \end{cases}$$

where,

$$\gamma_L = \lambda_{AC,m}^{H_0, \min} + p_{\beta 1} \bar{u} - \gamma'$$

$$\gamma_u = \lambda_{AC,m}^{H_0, \max} + p_{\beta 1} \bar{u} + \gamma'$$

\therefore the simplified test statistic can be given as,

$$y_{AC,m} = \frac{1}{N} \sum_{n=1}^N y_{AC,m} [u]$$

Using this, we can define P_D & P_F

→ Probability of false alarm:

$$\begin{aligned} P_F^{AC, M} &= P_H \{ \gamma_L \leq Y_{AC, M} \leq \gamma_u ; H_0 \} \\ &= P_H \{ 0 \leq Y_{AC, M} < \gamma_u ; H_0 \} - \\ &\quad P_H \{ 0 \leq Y_{AC, M} < \gamma_L ; H_0 \} \end{aligned}$$

$$\begin{aligned} P_F^{AC, M} &= P_H \{ 0 \leq \text{Pois} (N \lambda_{AC, M}^{H_0, \max} + N P_H \bar{\eta}) \leq N \gamma_u \} - \\ &\quad P_H \{ 0 \leq \text{Pois} (N \lambda_{AC, M}^{H_0, \min} + N P_H \bar{\eta}) < N \gamma_u \} \\ &= \text{Pois}cdf (N \lambda_{AC, M}^{H_0, \max} + N P_H \bar{\eta}, N \gamma_u) - \\ &\quad \text{Pois}cdf (N \lambda_{AC, M}^{H_0, \min} + N P_H \bar{\eta}, N \gamma_L) \end{aligned}$$

→ Probability of Detection:

$$\begin{aligned} P_D^{AC, M} &= P_H \{ 0 < Y_{AC, M} < \gamma_L ; H_1^- \} \cdot P_H \{ H_1^- | H_1 \} + \\ &\quad P_H \{ Y_{AC, M} > \gamma_u ; H_1^+ \} \cdot P_H \{ H_1^+ | H_1 \} \end{aligned}$$

Assuming, symmetric distribution,

$$P_H \{ H_1^- | H_1 \} = P_H \{ H_1^+ | H_1 \} = 0.5$$

$$\begin{aligned} P_D^{AC, M} &= 0.5 \left[P_H \{ \text{Pois} (N \lambda_{AC, M}^{H_1^+, \min} + N P_H \bar{\eta}) > N \gamma_u \} + \right. \\ &\quad \left. P_H \{ \text{Pois} (N \lambda_{AC, M}^{H_1^-, \max} + N P_H \bar{\eta}) < N \gamma_L \} \right] \\ &= 1 - \frac{\text{Pois}cdf (N \lambda_{AC, M}^{H_1^+, \min} + N P_H \bar{\eta}, N \gamma_u)}{2} + \\ &\quad \frac{\text{Pois}cdf (N \lambda_{AC, M}^{H_1^-, \max} + N P_H \bar{\eta}, N \gamma_L)}{2} \end{aligned}$$

Now, w.k.t error can be generated in two cases i.e.,

1. we detect the signal, even though it's not sent (P_F - False Alarm)
2. we don't detect the signal, even though it's sent (P_{MD} - Miss Detection)

∴ Probability of error will be summation of both, with β (prior probability) as a controlling parameter.

$$P_e^{AC} = P_{MD}^{AC} \cdot \beta_{AC} + P_F^{AC} \cdot (1 - \beta_{AC})$$

$$= (1 - P_D^{AC}) \cdot \beta_{AC} + P_F^{AC} (1 - \beta_{AC})$$

Assuming, symmetric distribution, $\beta_{AC} = 0.5$

ACs-BC Model

→ Anomaly in ACs,

$$H_0: x^{H_0, \min} \leq x_{D_0} \leq x^{H_0, \max}$$

$$H_1: \begin{cases} H_1^-: x_{D_0} \leq x^{H_1^-, \max} \\ H_1^+: x_{D_0} \geq x^{H_1^+, \min} \end{cases}$$

→ Anomaly in BC using hypothesis test,

$$W_0: \sum_{m=1}^M \lambda_{BC, m} < K \bar{\lambda}_{BC}$$

$$W_1: \sum_{m=1}^M \lambda_{BC, m} \geq K \bar{\lambda}_{BC}$$

where, W_0 & W_1 are events corresponding to H_0 & H_1 in BC

k is min. no. of ACs which detect anomaly in BC-AC link.

$\bar{\lambda}_{BC}$ is avg. no. of received molecules at BC using,

$$Y_{AC,m} = \frac{1}{N} \sum_{n=1}^N Y_{AC,m}[n]$$

It can be given for $\forall m \in 1, \dots, M$.

→ If AC-BC link will alarm the presence of anomaly when atleast k no. of ACs alarm hypothesis H_1 by transmitting $\lambda_{AC,m}$ molecules.

The decision rule, for the BC with Poisson observation, where y_{BC}^{THR} is decision threshold at BC,

$$H_0: y_{BC} < y_{BC}^{THR}$$

$$H_1: y_{BC} \geq y_{BC}^{THR}$$

For perfect AC-BC link, we derive lower bound for prob. of detection & upper bound for prob. of false alarm.

$$Q_D = P_{H_1} \left\{ \sum_{m=1}^M \lambda_{BC,m} \geq k \bar{\lambda}_{BC}; H_1 \right\} = \sum_{m=k}^M P'_m$$

$$Q_F = P_{H_1} \left\{ \sum_{m=1}^M \lambda_{BC,m} \geq k \bar{\lambda}_{BC}; H_0 \right\} = \sum_{m=k}^M P''_m$$



where,

min. no. of ACs which detect the anomaly in order to enable BC

→ Probability of Detection:

$$P_D = P_{H1} \{ y_{BC} \geq y_{BC}^{THR}; H_1 \}$$

$$= P_{H1} \{ y_{BC} \geq y_{BC}^{THR} | w_0; H_1 \} P_{H1} \{ w_0; H_1 \} + \\ P_{H1} \{ y_{BC} \geq y_{BC}^{THR} | w_1; H_1 \} P_{H1} \{ w_1; H_1 \}$$

(using Total Prob. Theorem)

From definitions of Q_D , w_0 & w_1 ,

$$= P_{H1} \{ y_{BC} \geq y_{BC}^{THR} | \sum_{m=1}^{K-1} \lambda_{BC,m} < K \bar{\lambda}_{BC}; H_1 \} (1 - Q_D)$$

$$+ P_{H1} \{ y_{BC} \geq y_{BC}^{THR} | \sum_{m=1}^M \lambda_{BC,m} \geq K \bar{\lambda}_{BC}; H_1 \} Q_D$$

$$= P_{H1} \{ y_{BC} \geq y_{BC}^{THR} | \bigcup_{i=0}^{K-1} \{ \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC} \}; H_1 \} (1 - Q_D)$$

$$+ P_{H1} \{ y_{BC} \geq y_{BC}^{THR} | \bigcup_{i=K}^M \{ \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC} \}; H_1 \} Q_D$$

$$= \frac{P_{H1} \{ \bigcup_{i=0}^{K-1} \{ \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC} \} \cap \{ y_{BC} \geq y_{BC}^{THR} \}; H_1 \}}{P_{H1} \{ \bigcup_{i=0}^{K-1} \{ \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC} \}; H_1 \}} (1 - Q_D)$$

$$+ \frac{P_{H1} \{ \bigcup_{i=K}^M \{ \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC} \} \cap \{ y_{BC} \geq y_{BC}^{THR} \}; H_1 \}}{P_{H1} \{ \bigcup_{i=K}^M \{ \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC} \}; H_1 \}} (Q_D)$$

$$= \frac{\sum_{i=0}^{K-1} (P_{H1} \{ y_{BC} \geq y_{BC}^{THR} | \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC}; H_1 \})}{P_{H1} \{ \sum_{m=1}^M \lambda_{BC,m} < K \bar{\lambda}_{BC}; H_1 \}}$$

$$\begin{aligned}
 & \lambda P_{H1} \left\{ \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC}, H_1 \right\} (1 - Q_D) \\
 & + \sum_{i=k}^M \left(P_{H1} \left\{ y_{BC} \geq y_{BC}^{THR} \mid \sum_{m=1}^M \lambda_{BC,m} [n] = i \bar{\lambda}_{BC}, H_1 \right\} \right. \\
 & \quad \left. P_{H1} \left\{ \sum_{m=1}^M \lambda_{BC,m} \geq k \bar{\lambda}_{BC}, H_1 \right\} \right) \\
 & * P_{H1} \left\{ \sum_{m=1}^M \lambda_{BC,m} = i \bar{\lambda}_{BC}, H_1 \right\} Q_D
 \end{aligned}$$

Following the Poisson Distribution of y_{BC} ,

$$= \frac{1}{\sum_{i=0}^{K-1} P_i'} \sum_{i=0}^{K-1} P_i' \left(\frac{1 - \Gamma(\Gamma_{y_{BC}^{THR}}, i \bar{\lambda}_{BC}) + q_n \bar{E}}{\Gamma_{y_{BC}^{THR}} - 1!} \right) (1 - Q_D)$$

$$+ \frac{1}{\sum_{i=1}^M P_i'} \sum_{i=k}^M P_i' \left(\frac{1 - \Gamma(\Gamma_{y_{BC}^{THR}}, i \bar{\lambda}_{BC}) + q_n \bar{E}}{\Gamma_{y_{BC}^{THR}} - 1!} \right) Q_D$$

$$\therefore P_D = \sum_{i=0}^M P_i' \left(\frac{1 - \Gamma(\Gamma_{y_{BC}^{THR}}, i \bar{\lambda}_{BC}) + q_n \bar{E}}{\Gamma_{y_{BC}^{THR}} - 1!} \right)$$

FINALLY!

Probability of False Alarm:

$$\begin{aligned}
 P_F &= P_{H1} \{ y_{BC} \geq y_{BC}^{THR}; H_0 \} \\
 &= P_{H1} \{ y_{BC} \geq y_{BC}^{THR} \mid w_0; H_0 \} P_{H1} \{ w_0; H_0 \} \\
 &\quad + P_{H1} \{ y_{BC} \geq y_{BC}^{THR} \mid w_1; H_0 \} P_{H1} \{ w_1; H_0 \}
 \end{aligned}$$

Following similar steps as of P_D , we get,

$$= \frac{1}{\sum_{i=0}^{K-1} P_i''} \sum_{i=0}^{K-1} P_i'' \left(1 - \frac{\Gamma(\gamma_{BC}^{THR}, i \bar{\lambda}_{BC}) + q_{g1} \bar{E}}{\gamma_{BC}^{THR} - 1} \right) (1 - Q_F)$$

$$+ \frac{1}{\sum_{i=K}^M P_i''} \sum_{i=K}^M P_i'' \left(1 - \frac{\Gamma(\gamma_{BC}^{THR}, i \bar{\lambda}_{BC}) + q_{g1} \bar{E}}{\gamma_{BC}^{THR} - 1} \right) Q_F$$

$$= \sum_{i=0}^M P_i'' \left(1 - \frac{\Gamma(\gamma_{BC}^{THR}, i \bar{\lambda}_{BC}) + q_{g1} \bar{E}}{\gamma_{BC}^{THR} - 1} \right)$$

$$\therefore P_E = P_{MD} \beta + P_F (1 - \beta)$$

$$= (1 - P_D) \beta + P_F (1 - \beta)$$

where $\beta = 0.5$