

# Modeling and Forecasting with ARMA models

## 1 Method of Moments Estimator: Yule-Walker Estimation

AR 모형에서 많이 사용

### (1) AR models

- consider AR(p) process:

$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + \dots + \phi_p X_{t-p} + Z_t$$

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- Then, we know

$$\begin{aligned} \text{Cov}(X_t, X_t) \gamma(0) &= \phi_1 \gamma(1) + \phi_2 \gamma(2) + \dots + \phi_p \gamma(p) + \sigma_w^2 \quad // E X_t Z_t = r(0) \text{ 일때만 성립.} \\ \text{Cov}(X_t, X_{t+1}) \gamma(1) &= \phi_1 \gamma(0) + \phi_2 \gamma(-1) + \dots + \phi_p \gamma(1-p) \quad // E X_t Z_{t+1} = 0 \\ &\vdots \\ \text{Cov}(X_t, X_{t+p}) \gamma(p) &= \phi_1 \gamma(p-1) + \phi_2 \gamma(p-2) + \dots + \phi_p \gamma(-1) \end{aligned}$$

- Goal: estimation of  $\phi_1, \phi_2, \dots, \phi_p$ .
- Replace population ACF  $\gamma(h)$  to sample ACF  $\hat{\gamma}(h)$

$$r(k) = E X_t X_{t-k}$$

$$\hat{r}(k) = \frac{1}{n-k} \sum_{t=k+1}^n X_t X_{t-k}$$

$$\begin{aligned} \hat{\gamma}(0) &= \phi_1 \hat{\gamma}(1) + \phi_2 \hat{\gamma}(2) + \dots + \phi_p \hat{\gamma}(p) + \sigma_w^2 \\ \hat{\gamma}(1) &= \phi_1 \hat{\gamma}(0) + \phi_2 \hat{\gamma}(-1) + \dots + \phi_p \hat{\gamma}(1-p) \\ &\vdots \\ \hat{\gamma}(p) &= \phi_1 \hat{\gamma}(p-1) + \phi_2 \hat{\gamma}(p-2) + \dots + \phi_p \hat{\gamma}(-1) \end{aligned}$$

- Example with R code: AR (2)

$$\begin{aligned} \hat{r}(0) &= 0.434 \\ \hat{r}(1) &= 0.034 \\ \hat{r}(2) &= 0.476 \end{aligned} \quad \left. \begin{array}{l} \end{array} \right\} \text{이것들을 } \sigma^2 \text{ 추정}$$

$$\begin{pmatrix} 1 & \hat{r}(1) \\ \hat{r}(1) & 1 \end{pmatrix} \begin{pmatrix} \hat{\phi}_1 \\ \hat{\phi}_2 \end{pmatrix} = \begin{pmatrix} \hat{r}(1) \\ \hat{r}(2) \end{pmatrix}$$

$$\begin{pmatrix} \hat{\phi}_1 \\ \hat{\phi}_2 \end{pmatrix} = \frac{1}{1 - \hat{r}(1)^2} \begin{pmatrix} 1 & -\hat{r}(1) \\ -\hat{r}(1) & 1 \end{pmatrix} \begin{pmatrix} \hat{r}(1) \\ \hat{r}(2) \end{pmatrix} = \begin{pmatrix} 1.439 \\ -0.725 \end{pmatrix}$$

$$\begin{aligned} \therefore \hat{\sigma}_w^2 &= \hat{r}(0) - \hat{\phi}_1 \hat{r}(1) - \hat{\phi}_2 \hat{r}(2) \\ &= \hat{r}(0) (1 - \hat{\phi}_1 \hat{r}(1) - \hat{\phi}_2 \hat{r}(2)) = 1.215 \end{aligned}$$

$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + Z_t$$

$$r(1) = \phi_1 r(0) + \phi_2 r(0)$$

$$r(2) = \phi_1 r(1) + \phi_2 r(0)$$

$$\begin{pmatrix} r(0) & r(1) \\ r(1) & r(0) \end{pmatrix} \begin{pmatrix} \phi_1 \\ \phi_2 \end{pmatrix} = \begin{pmatrix} r(1) \\ r(2) \end{pmatrix}$$

$$= r(0) \begin{pmatrix} 1 & \hat{r}(1) \\ \hat{r}(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_1 \\ \phi_2 \end{pmatrix} = \begin{pmatrix} r(1) \\ r(2) \end{pmatrix}$$

$$\rightarrow \begin{pmatrix} 1 & \hat{r}(1) \\ \hat{r}(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_1 \\ \phi_2 \end{pmatrix} = \begin{pmatrix} \hat{r}(1) \\ \hat{r}(2) \end{pmatrix}$$

- Large Sample Distribution of Yule-Walker Estimator

For a large sample from AR (p) process,

$$\hat{\phi} \approx \text{Normal}\left(\phi, \frac{1}{n}\sigma^2\Gamma_p^{-1}\right)$$

Fact 2.  $\frac{1}{n}\sigma^2\Gamma_p^{-1}$

$$\hat{\phi} \sim N\left(\phi, \frac{1}{n}\sigma^2 \begin{pmatrix} \gamma(0) & \dots & \gamma(n-1) \\ \gamma(1) & \ddots & \gamma(n-2) \\ \vdots & \ddots & \vdots \\ \gamma(p) & \dots & \gamma(0) \end{pmatrix}^{-1}\right)$$

- (2) MA models

ex) AR(1)

$$\hat{\phi} \sim N\left(\phi, \frac{1}{n}\sigma^2 \gamma(0)^{-1}\right)$$

- Consider the MA(1) model  $X_t = \theta X_{t-1} + Z_t$  where  $|\theta| < 1$ .

- Yule-walker equations are

$$\gamma(0) = \sigma^2(1 + \theta^2)$$

$$\gamma(1) = \sigma^2\theta$$

- The method of moments is an effective procedure for fitting auto-regressive models, it does not perform as well for ARMA models with  $q > 0$ . From a computational point of view, it requires as much computing time as the more efficient estimators based on either the innovations algorithm or the Hannan-Rissanen procedure and is therefore rarely used except when  $q = 0$ .

• The Hannan-Rissanen Algorithm: MA 모형.

- Assume ARMA(p,q) model
- $X_t$  is regressed both on  $X_{t-1}, X_{t-2}, \dots, X_{t-p}$  and  $Z_{t-1}, Z_{t-2}, \dots, Z_{t-q}$ .
- Since  $\{Z_t\}$  are unobserved, we replace them with estimated residuals from

$$\hat{Z}_t = X_t - \hat{\phi}_{m1}X_{t-1} - \hat{\phi}_{m2}X_{t-2} - \dots - \hat{\phi}_{mm}X_{t-m}$$

for  $m$  large enough.

- One the estimated residuals  $\{\hat{Z}_t\}$  are computed, the vector of parameters  $(\beta', \phi')'$  is estimated by least squares linear regression of  $X_t$  onto  $(X_{t-1}, \dots, X_{t-p}, \hat{Z}_{t-1}, \dots, \hat{Z}_{t-q})$ ,  $t = m+1+q, \dots, n$  by minimizing the sum of squares

$$S(\beta) = \sum_{t=m+1+q}^n \left( X_t - \phi_1 X_{t-1} - \dots - \phi_p X_{t-p} - \theta_1 \hat{Z}_{t-1} - \dots - \theta_q \hat{Z}_{t-q} \right)^2$$

with respect to  $\beta$ . MA 모형

- Finally

$$X_t = \underbrace{\phi_1 X_{t-1} + \dots + \phi_p X_{t-p}}_{\text{관측 가능}} + \underbrace{\theta_1 Z_{t-1} + \dots + \theta_q Z_{t-q}}_{\text{관측 불가}} + Z_t$$

$$\hat{\sigma}^2 = \frac{S(\hat{\beta})}{n - m - q} \rightarrow \hat{\sigma}^2(0) = \phi_1 \hat{\sigma}^2(1) + \dots + \phi_p \hat{\sigma}^2(p) + \hat{\sigma}^2$$

가능

$$X_t \approx \phi_{m1} X_{t-1} + \dots + \phi_{mn} X_{t-n} + Z_t \text{ for sufficiently large } m.$$

$$\rightarrow \hat{\phi}_{m1}, \dots, \hat{\phi}_{mn} \text{ 추정 가능. (Yule walker or regression)}$$

$$\rightarrow \hat{Z}_t \approx X_t - \hat{\phi}_{m1} X_{t-1} - \dots - \hat{\phi}_{mn} X_{t-m} \rightarrow S(\beta) \text{ 적용}$$

이건 Yule walker 과정과 동일!!  
Yule walker로  $\phi$  추정한후  
과정만 추가된 것!!

ex) MA(1)

$$X_t = Z_t + \theta Z_{t-1}$$

$$Z_t = X_t - \theta Z_{t-1} \quad \rightarrow \theta \text{에 대한 편미분 (구분할 수 있음)}$$

$$S(\beta) = \sum_{t=m+1}^n (X_t - \theta Z_{t-1})^2$$

$$\hat{\sigma}^2 = \frac{-2Z_{t-1}(X_t - \theta Z_{t-1})}{n - m - 1}$$

원래 주어진

# Model Selection by Information Criteria

Information Criteria =  $-2 \log -\text{likelihood value}$  + penalty

Let  $k$  = number of parameters *변수 개수*

(a) Akaike's Information Criteria (AIC)

*와해계*

$$-2 \log -\text{likelihood value} + \{2k\}$$

(b) Bias corrected AIC

$$-2 \log -\text{likelihood value} + 2 \left\{ \frac{n(n+k)}{n-k-2} \right\}$$

(c) Bayesian Information Criteria (BIC)

*각원*

$$-2 \log -\text{likelihood value} + \{k \log n\} \quad n: \text{sample size}$$