# Optimization and Simulation

Simulating events: the Poisson process

#### Michel Bierlaire

Transport and Mobility Laboratory
School of Architecture, Civil and Environmental Engineering
Ecole Polytechnique Fédérale de Lausanne







## Siméon Denis Poisson





Siméon-Denis Poisson French mathematician (1781–1840).

## Outline

Poisson random variable

Poisson process

3 Non homogeneous Poisson process

# Binomial random variable

#### Context

- n: number of independent trials
- p: probability of a success
- X: number of successes

## Probability of k successes

$$\Pr(X=k) = \binom{n}{k} p^k (1-p)^{n-k}.$$

# Poisson random variable

#### Context: continuous trials

- $n \to +\infty$ : large number of trials
- $p \rightarrow 0$ : low probability of a success
- $np \rightarrow \lambda$ : success rate
- X: number of successes

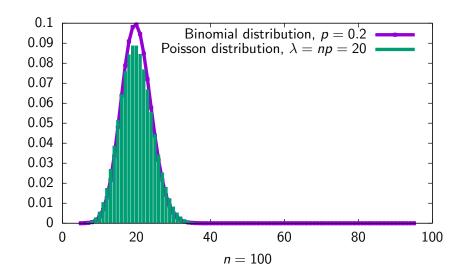
## Probability of k successes

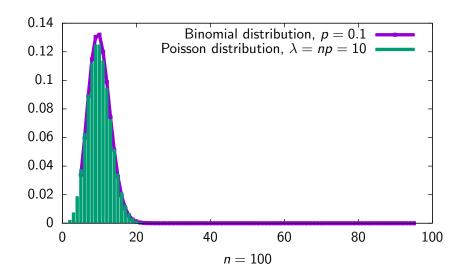
$$\Pr(X = k) = e^{-\lambda} \frac{\lambda^k}{k!}.$$

## Property

$$E[X] = Var(X) = \lambda.$$







# Outline

- 2 Poisson process
- 3 Non homogeneous Poisson process



## Events are occurring at random time points

N(t) is the number of events during [0, t]

# Poisson process with rate $\lambda > 0$ if

- N(0) = 0,
- # of events occurring in disjoint time intervals are independent,
- 3 distribution of N(t+s) N(t) depends on s, not on t,
- **1** probability of one event in a small interval is approx.  $\lambda h$ :

$$\lim_{h\to 0}\frac{\Pr(N(h)=1)}{h}=\lambda,$$

oprobability of two events in a small interval is approx. 0:

$$\lim_{h \to 0} \frac{\Pr(N(h) \ge 2)}{h} = 0.$$

## Property

$$N(t) \sim \mathsf{Poisson}(\lambda t), \quad \mathsf{Pr}(N(t) = k) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

#### Inter-arrival times

- $S_k$  is the time when the kth event occurs,
- $X_k = S_k S_{k-1}$  is the time elapsed between event k-1 and event k.
- $X_1 = S_1$
- Distribution of  $X_1$ :  $\Pr(X_1 > t) = \Pr(N(t) = 0) = e^{-\lambda t}$ .
- Distribution of  $X_k$ :

$$\Pr(X_k > t | S_{k-1} = s) = \Pr(0 \text{ events in } ]s, s+t] | S_{k-1} = s)$$

$$= \Pr(0 \text{ events in } ]s, s+t])$$

$$= e^{-\lambda t}.$$

## Inter-arrival times (ctd.)

Therefore, the CDF of  $X_k$  is, for any k,

$$F(t) = \Pr(X_k \le t) = 1 - \Pr(X_k > t) = 1 - e^{-\lambda t}.$$

The pdf is

$$f(t) = \frac{dF(t)}{dt} = \lambda e^{-\lambda t}.$$

#### Conclusion

The inter-arrival times  $X_1$ ,  $X_2$ ,... are independent and identically distributed exponential random variables with parameter  $\lambda$ , and mean  $1/\lambda$ .

#### Simulation

- Simulation of event times of a Poisson process with rate  $\lambda$  until time T:
  - **1** t = 0, k = 0.
  - ② Draw  $r \sim U(0,1)$ .

  - 4 If t > T, STOP.

  - 6 Go to step 2.



# Outline

- Poisson process
- 3 Non homogeneous Poisson process

#### Rate varies with time

 $\lambda(t)$ .

# Non homogeneous Poisson process with rate $\lambda(t)$ if

- N(0) = 0
- # of events occurring in disjoint time intervals are independent,
- **o** probability of one event in a small interval is approx.  $\lambda(t)h$ :

$$\lim_{h\to 0}\frac{\Pr\left(\left(\textit{N}(t+h)-\textit{N}(t)\right)=1\right)}{h}=\lambda(t),$$

probability of two events in a small interval is approx. 0:

$$\lim_{h\to 0} \frac{\Pr\left(\left(N(t+h)-N(t)\right)\geq 2\right)}{h}=0.$$

#### Mean value function

$$m(t) = \int_0^t \lambda(s) ds, \ t \geq 0.$$

#### Poisson distribution

$$N(t+s) - N(t) \sim \mathsf{Poisson}(m(t+s) - m(t))$$

### Link with homogeneous Poisson process

- Consider a Poisson process with rate  $\lambda$ .
- If an event occurs at time t, count it with probability p(t).
- The process of counted events is a non homogeneous Poisson process with rate  $\lambda(t) = \lambda p(t)$ .

#### Proof

- N(0) = 0 [OK]
- ullet of events occurring in disjoint time intervals are independent,  $[\mathsf{OK}]$
- **9** probability of one event in a small interval is approx.  $\lambda(t)h$ : [?]

$$\lim_{h\to 0}\frac{\Pr\left(\left(N(t+h)-N(t)\right)=1\right)}{h}=\lambda(t),$$

probability of two events in a small interval is approx. 0: [OK]

$$\lim_{h\to 0} \frac{\Pr\left(\left(N(t+h)-N(t)\right)\geq 2\right)}{h}=0.$$

# Proof (ctd.)

- N(t) number of events of the non homogeneous process
- N'(t) number of events of the underlying homogeneous process

$$\begin{split} \Pr \left( (N(t+h) - N(t)) = 1 \right) \\ &= \sum_{k=1}^{\infty} \Pr \left( (N'(t+h) - N'(t)) = k, 1 \text{ is counted} \right) \\ &= \Pr \left( (N'(t+h) - N'(t)) = 1, 1 \text{ is counted} \right) \\ &= \Pr \left( (N'(t+h) - N'(t)) = 1 \right) \Pr (1 \text{ is counted}) \\ &= \Pr \left( N'(h) = 1 \right) \Pr (1 \text{ is counted}) \end{split}$$

$$\lim_{h\to 0} \frac{\Pr((N(t+h)-N(t))=1)}{h} = \lim_{h\to 0} \frac{\Pr(N'(h)=1)}{h} \Pr(1 \text{ is counted})$$
$$= \lambda p(t).$$

# Simulation of event times of a non homogeneous Poisson process with rate $\lambda(t)$ until time T

- **①** Consider  $\lambda$  such that  $\lambda(t) \leq \lambda$ , for all  $t \leq T$ .
- ② t = 0, k = 0.
- **3** Draw  $r \sim U(0, 1)$ .
- $t = t \ln(r)/\lambda.$
- lacksquare If t > T, STOP.
- **o** Generate  $s \sim U(0,1)$ .
- If  $s \leq \lambda(t)/\lambda$ , then k = k + 1,  $S_k = t$ .
- Go to step 3.



# Summary

#### Outline

- Poisson random variable
- Poisson process
- Non homogeneous Poisson process

#### Comments

- Main assumption: events occur continuously and independently of one another
- Typical usage: arrivals of customers in a queue
- Easy to simulate