## THU-70250043, Pattern Recognition (Spring 2018)

Homework: 3

Gaussian Mixture Model and Expectation Maximization Algorithm

Lecturer: Changshui Zhang zcs@mail.tsinghua.edu.cn

Student: XXX xxx@mails.tsinghua.edu.cn

## EM and Gradient Descent

In this problem you will investigate connections between the EM algorithm and gradient descent. Consider a GMM where  $\Sigma_k = \sigma_k^2 I$ , i.e., the covariances are spherical but of different spread. Moreover, suppose the mixture weight  $\pi_k$  is known. The log likelihood then is

$$l(\{\mu_k, \sigma_k^2\}_{k=1}^K) = \sum_{i=1}^n \log \left( \sum_{k=1}^K \pi_k \ N(x_i | \mu_k, \sigma_k^2 I) \right).$$

A maximization algorithm based on gradient descent is as follows:

- Initialize  $\mu_k$  and  $\sigma_k^2$ ,  $k \in \{1, \dots, K\}$ . Set the iteration counter t=1.
- Repeat the following until convergence:

- For 
$$k = 1, \dots, K$$
,

$$\boldsymbol{\mu}_k^{(t+1)} \leftarrow \boldsymbol{\mu}_k^{(t)} + \eta_k^{(t)} \nabla_{\mu_k} l\left(\{\boldsymbol{\mu}_k^{(t)}, (\sigma_k^2)^{(t)}\}_{k=1}^K\right)$$

- For 
$$k = 1, \dots, K$$
,

$$(\sigma_k^2)^{(t+1)} \leftarrow (\sigma_k^2)^{(t)} + s_k^{(t)} \nabla_{\sigma_k^2} l \left( \{ \mu_k^{(t+1)}, (\sigma_k^2)^{(t)} \}_{k=1}^K \right)$$

– Increase the iteration counter  $t \leftarrow t+1$ 

Show that with properly chosen step size  $\eta_k^{(t)}$  and  $s_k^{(t)}$ , the above gradient descent algorithm is equivalent to the following modified EM algorithm:

- Initialize  $\mu_k$  and  $\sigma_k^2$ ,  $k \in \{1, \dots, K\}$ . Set the iteration counter t=1.
- Repeat the following until convergence:

$$\tilde{z}_{ik}^{(t+0.5)} \leftarrow Prob\left(x_i \in cluster_k | \{(\mu_j^{(t)}, (\sigma_j^2)^{(t)})\}_{j=1}^K, x_i\right),$$

- M-step:

$$\{\mu_k^{(t+1)}\}_{k=1}^K \leftarrow \arg\max_{\{\mu_k\}_{k=1}^K} \sum_{i=1}^n \sum_{k=1}^K \tilde{z}_{ik}^{(t+0.5)} \left(\log N(x_i|\mu_k, (\sigma_k^2)^{(t)}I) + \log \pi_k\right)$$

- E-step:

$$\tilde{z}_{ik}^{(t+1)} \leftarrow Prob\left(x_i \in cluster_k | \{(\mu_j^{(t+1)}, (\sigma_j^2)^{(t)})\}_{j=1}^K, x_i\right),$$

- M-step:

$$\{(\sigma_k^2)^{(t+1)}\}_{k=1}^K \leftarrow \arg\max_{\{\sigma_k\}_{k=1}^K} \sum_{i=1}^n \sum_{k=1}^K \tilde{z}_{ik}^{(t+1)} \left(\log N(x_i | \mu_k^{(t+1)}, \sigma_k^2 I) + \log \pi_k\right)$$

- Increase the iteration counter  $t \leftarrow t+1$ 

The main modification is inserting an extra E-step between the M-step for  $\mu_k$ 's and the M-step for  $\sigma_k^2$ 's.

## EM for MAP Estimation

The EM algorithm that we talked about in class was for solving a maximum likelihood estimation problem in which we wished to maximize

$$\prod_{i=1}^{m} p(x^{(i)}; \theta) = \prod_{i=1}^{m} \sum_{z^{(i)}} p(x^{(i)}, z^{(i)}; \theta)$$
(1)

where the  $z^{(i)}$ 's were latent random variables. Suppose we are working in a Bayesian framework, and wanted to find the MAP estimate of the parameters  $\theta$  by maximizing

$$(\prod_{i=1}^{m} p(x^{(i)}; \theta)) p(\theta) = (\prod_{i=1}^{m} \sum_{z^{(i)}} p(x^{(i)}, z^{(i)} | \theta)) p(\theta)$$
(2)

Here,  $p(\theta)$  is our prior on the parameters. Generalize the EM algorithm to work for MAP estimation. You may assume that  $\log p(x, z|\theta)$  and  $\log p(\theta)$  are both concave in  $\theta$ , so that the M-step is tractable if it requires only maximizing a linear combination of these quantities. (This roughly corresponds to assuming that MAP estimation is tractable when x, z is fully observed, just like in the frequentist case where we considered examples in which maximum likelihood estimation was easy if x, z was fully observed.)

Make sure your M-step is tractable, and also prove that  $(\prod_{i=1}^m p(x^{(i)};\theta))p(\theta)$  (viewed as a function of  $\theta$ ) monotonically increases with each iteration of your algorithm.

## Programming

| Points | $\omega_1$ |        |       | $\omega_2$ |       |        |
|--------|------------|--------|-------|------------|-------|--------|
|        | $x_1$      | $x_2$  | $x_3$ | $x_1$      | $x_2$ | $x_3$  |
| 1      | 0.42       | -0.087 | 0.58  | -0.4       | 0.58  | 0.089  |
| 2      | -0.2       | -3.3   | -3.4  | -0.31      | 0.27  | -0.04  |
| 3      | 1.3        | -0.32  | 1.7   | 0.38       | 0.055 | -0.035 |
| 4      | 0.39       | 0.71   | 0.23  | -0.15      | 0.53  | 0.011  |
| 5      | -1.6       | -5.3   | -0.15 | -0.35      | 0.47  | 0.034  |
| 6      | -0.029     | 0.89   | -4.7  | 0.17       | 0.69  | 0.1    |
| 7      | -0.23      | 1.9    | 2.2   | -0.011     | 0.55  | -0.18  |
| 8      | 0.27       | -0.3   | -0.87 | -0.27      | 0.61  | 0.12   |
| 9      | -1.9       | 0.76   | -2.1  | -0.065     | 0.49  | 0.0012 |
| 10     | 0.87       | -1.0   | -2.6  | -0.12      | 0.054 | -0.063 |

Table 1: Data for Programming

Suppose we know that the ten data points in category  $\omega_1$  in the table above come from a three-dimensional Gaussian. Suppose, however, that we do not have access to the  $x_3$  components for the even-numbered data points.

- Write an EM program to estimate the mean and covariance of the distribution. Start your estimate with  $\mu_0 = 0$  and  $\Sigma_0 = I$ , the three-dimensional identity matrix.
- Compare your final estimate with that for the case when there is no missing data

Suppose we know that the ten data points in category  $\omega_2$  in the table above come from a three-dimensional uniform distribution  $p(x|\omega_2) \sim U(x_l, x_u)$ . Suppose, however, that we do not have access to the  $x_3$  components for the even-numbered data points.

- Write an EM program to estimate the six scalars comprising  $x_l$  and  $x_u$  of the distribution. Start your estimate with  $x_l = (-2, -2, -2)^t$  and  $x_u = (+2, +2, +2)^t$ .
- Compare your final estimate with that for the case when there is no missing data.