

P1.

(a). $\sum_{i=1}^n (X_i - \bar{X})^2 = \sum_{i=1}^n [(X_i - \bar{X}) + \bar{X}]^2$
 $= \sum_{i=1}^n (X_i - \bar{X})^2 + 2\bar{X} \sum_{i=1}^n (X_i - \bar{X}) + \sum_{i=1}^n \bar{X}^2$
As $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, $\sum_{i=1}^n (X_i - \bar{X}) = \sum_{i=1}^n X_i - n\bar{X} = 0$, we can get
 $\sum_{i=1}^n X_i^2 = \sum_{i=1}^n (X_i - \bar{X})^2 + n\bar{X}^2$
 $(n-1)S^2 + n\bar{X}^2 = (n-1) \cdot \frac{1}{n-1} \cdot \sum_{i=1}^n (X_i - \bar{X})^2 + n\bar{X}^2 = \sum_{i=1}^n (X_i - \bar{X})^2 + n\bar{X}^2$.
So the equation is correct.

(b) As $\frac{n-1}{\sigma^2} S^2 \sim \chi^2_{n-1}$, with $n-1$ degrees of freedom

$$E\left(\frac{n-1}{\sigma^2} S^2\right) = E(\chi^2_{n-1}) = n-1$$

$$\begin{aligned} \text{Thus } E(S^2) &= E\left(\frac{\sigma^2}{n-1} \cdot \frac{n-1}{\sigma^2} S^2\right) \\ &= \frac{\sigma^2}{n-1} \cdot E\left(\frac{n-1}{\sigma^2} S^2\right) \\ &= \frac{\sigma^2}{n-1} \cdot n-1 = \sigma^2. \end{aligned}$$

So S^2 is an unbiased estimator of ~~σ^2~~ σ^2 .(c) To prove \bar{X} is independent of $X_i - \bar{X}$, we can prove $\text{Cov}(\bar{X}, X_i - \bar{X}) = 0$.

$$\begin{aligned} \text{Cov}(\bar{X}, X_i - \bar{X}) &= \text{Cov}(\bar{X}, X_i) - \text{Cov}(\bar{X}, \bar{X}) \\ &= \text{Cov}\left(\frac{1}{n} \sum_{i=1}^n X_i, X_i\right) - \text{Var}(\bar{X}) \\ &= \frac{1}{n} \text{Cov}\left(\sum_{i=1}^n X_i, X_i\right) - \frac{\sigma^2}{n} \\ &= \frac{1}{n} [\text{Cov}(X_1, X_1) + \text{Cov}(X_1, X_2) + \dots + \text{Cov}(X_1, X_n)] - \frac{\sigma^2}{n}. \end{aligned}$$

As X_1, X_2, \dots, X_n are i.i.d., $\text{Cov}(X_i, X_j)$ ($i \neq j$) = 0.
So $\text{Cov}(\bar{X}, X_i - \bar{X}) = \frac{1}{n} \text{Cov}(X_i, X_i) - \frac{\sigma^2}{n}$
 $= \frac{6\sigma^2}{n} - \frac{6\sigma^2}{n} = 0.$

So \bar{X} is independent of $X_i - \bar{X}$.(d). As $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$, as proved in (c), \bar{X} is independent of $X_i - \bar{X}$,
thus \bar{X} is independent of S^2 .

As $\bar{X} = \bar{Y} = 0$.

$$\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n y_i x_i}{\sum x_i^2}, \quad \hat{\beta}_0 = 0, \text{ i.e. } \hat{Y} = \hat{\beta}_1 X$$

To proof $R^2 = r^2$:

$$\frac{\sum_{i=1}^n \hat{y}_i^2}{\sum_{i=1}^n y_i^2} = \frac{\left(\sum_{i=1}^n x_i y_i\right)^2}{\sum_{i=1}^n x_i^2 \sum_{i=1}^n y_i^2}$$

$$\sum_{i=1}^n \hat{y}_i^2 = \left(\frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2}\right) \cdot \left(\sum_{i=1}^n x_i y_i\right)$$

$$\sum_{i=1}^n \hat{y}_i^2 = \hat{\beta}_1 \cdot \sum_{i=1}^n x_i y_i$$

$$\sum_{i=1}^n \hat{\beta}_1^2 x_i^2 = \hat{\beta}_1 \sum_{i=1}^n x_i y_i$$

$$\sum_{i=1}^n \hat{\beta}_1 x_i \cdot x_i = \hat{\beta}_1 \sum_{i=1}^n x_i y_i$$

$$\sum_{i=1}^n x_i y_i$$

$$\sum_{i=1}^n x_i (\hat{\beta}_1 x_i - y_i) = 0.$$

$$\sum_{i=1}^n x_i \left(\frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2} x_i - y_i \right) = 0.$$

$$\sum_{i=1}^n x_i y_i \left(\frac{\sum_{i=1}^n x_i}{\sum_{i=1}^n x_i^2} x_i - 1 \right) = 0.$$

$$\sum_{i=1}^n x_i y_i \cdot 0 = 0.$$

As the equation is correct, it's proved that $R^2 = r^2$.

P6

$$\bar{x} = \frac{\sum_{i=1}^n x_i}{20} = 0.428, \bar{y} = \frac{\sum_{i=1}^n y_i}{20} = 19.91.$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n x_i y_i - \bar{x} \sum_{i=1}^n y_i - \bar{y} \sum_{i=1}^n x_i + \bar{x} \bar{y}}{\sum_{i=1}^n x_i^2 - 2 \bar{x} \sum_{i=1}^n x_i + \bar{x}^2} = \frac{30.101}{34.84} = 0.86$$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x} = 22.772 - 0.86 \cdot 0.428 = 22.039$$

$$\begin{aligned} S^2 &= \frac{1}{n-2} \sum_{i=1}^n (y_i - \hat{y}_i)^2 = \frac{1}{n-2} \sum_{i=1}^n y_i^2 - 2\bar{y} \sum_{i=1}^n y_i + \bar{y}^2 \\ &= \frac{1}{n-2} \sum_{i=1}^n (y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i))^2 = \frac{1}{n-2} \left(\sum_{i=1}^n (y_i^2 - 2(\hat{\beta}_0 + \hat{\beta}_1 x_i) y_i + (\hat{\beta}_0 + \hat{\beta}_1 x_i)^2) \right) = 1.36 \end{aligned}$$

$$\text{When } x=0.5, y = 22.772 - 0.86 \cdot 0.5 = 22.487$$

$$R^2 = \frac{ESS}{TSS} = \frac{\sum (\hat{y}_i - \bar{y})^2}{\sum (y_i - \bar{y})^2} = \frac{\sum (\hat{\beta}_0 + \hat{\beta}_1 x_i)^2 - 2\bar{y} \sum (\hat{\beta}_0 + \hat{\beta}_1 x_i) y_i + \bar{y}^2}{\sum y_i^2 - 2\bar{y} \sum y_i + \bar{y}^2} = 0.977$$

P7

$$MSR = \frac{RSS}{p-1} \sim \lambda^2$$

$$MSE = \frac{ESS}{n-p} \sim \lambda^2$$

ESS has $n-p$ degrees of freedom associated with it since p parameters need to be estimated in the regression function.

RSS has $p-1$ degrees of freedom associated with it, representing the number of X variables X_1, \dots, X_{p-1} . here $p-1 = 6$. $p=7$.

To test whether there's a regression relation between the response Y and the set of X variables X_1, \dots, X_{p-1} , i.e. to choose between alternatives

$$H_0: \beta_1 = \beta_2 = \dots = \beta_{p-1} = 0.$$

$$H_1: \text{not all } \beta_k (k=1, \dots, p-1) \text{ equal zero.}$$

$$F^* = \frac{MSR}{MSE} = \frac{\frac{RSS}{6}}{\frac{ESS}{45-7}} = \frac{\frac{RSS}{6}}{\frac{TSS-RSS}{38}} = \frac{1.89}{1.89} = 1.89$$

The p-value can be calculated by F-test \rightarrow P-value calculator.

p-value is 0.108 for $F(6, 38)$ with $F^* = 1.89$