

Exercise Sheet 1

Exercise 1: Estimating the Bayes Error (10 + 10 + 10 P)

The Bayes decision rule for the two classes classification problem results in the Bayes error

$$P(\text{error}) = \int P(\text{error} \mid \mathbf{x}) p(\mathbf{x}) d\mathbf{x},$$

where $P(\text{error} \mid \mathbf{x}) = \min [P(\omega_1 \mid \mathbf{x}), P(\omega_2 \mid \mathbf{x})]$ is the probability of error for a particular input \mathbf{x} . Interestingly, while class posteriors $P(\omega_1 \mid \mathbf{x})$ and $P(\omega_2 \mid \mathbf{x})$ can often be expressed analytically and are integrable, the error function has discontinuities that prevent its analytical integration, and therefore, direct computation of the Bayes error.

- (a) Show that the full error can be upper-bounded as follows:

$$P(\text{error}) \leq \int \frac{2}{\frac{1}{P(\omega_1 \mid \mathbf{x})} + \frac{1}{P(\omega_2 \mid \mathbf{x})}} p(\mathbf{x}) d\mathbf{x}.$$

Note that the integrand is now continuous and corresponds to the harmonic mean of class posteriors weighted by $p(\mathbf{x})$.

- (b) Show using this result that for the univariate probability distributions

$$p(x \mid \omega_1) = \frac{\pi^{-1}}{1 + (x - \mu)^2} \quad \text{and} \quad p(x \mid \omega_2) = \frac{\pi^{-1}}{1 + (x + \mu)^2},$$

the Bayes error can be upper-bounded by:

$$P(\text{error}) \leq \frac{2 P(\omega_1) P(\omega_2)}{\sqrt{1 + 4\mu^2 P(\omega_1) P(\omega_2)}}$$

(Hint: you can use the identity $\int \frac{1}{ax^2+bx+c} dx = \frac{2\pi}{\sqrt{4ac-b^2}}$ for $b^2 < 4ac$.)

- (c) Explain how you would estimate the error if there was no upper-bounds that are both tight and analytically integrable. Discuss following two cases: (1) the data is low-dimensional and (2) the data is high-dimensional.

Exercise 2: Bayes Decision Boundaries (15 + 15 P)

One might speculate that, in some cases, the generated data $p(x \mid \omega_1)$ and $p(x \mid \omega_2)$ is of no use to improve the accuracy of a classifier, in which case one should only rely on prior class probabilities $P(\omega_1)$ and $P(\omega_2)$ assumed here to be strictly positive.

For the first part of this exercise, we assume that the data for each class is generated by the univariate Laplacian probability distributions:

$$p(x \mid \omega_1) = \frac{1}{2\sigma} \exp\left(-\frac{|x - \mu|}{\sigma}\right) \quad \text{and} \quad p(x \mid \omega_2) = \frac{1}{2\sigma} \exp\left(-\frac{|x + \mu|}{\sigma}\right).$$

where $\mu, \sigma > 0$.

- (a) Determine for which values of $P(\omega_1), P(\omega_2), \mu, \sigma$ the optimal decision is to always predict the first class (i.e. under which conditions $P(\text{error} \mid x) = P(\omega_2 \mid x) \quad \forall x \in \mathbb{R}$).
- (b) Repeat the exercise for the case where the data for each class is generated by the univariate Gaussian probability distributions:

$$p(x \mid \omega_1) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) \quad \text{and} \quad p(x \mid \omega_2) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x + \mu)^2}{2\sigma^2}\right).$$

where $\mu, \sigma > 0$.

Exercise 3: Programming (40 P)

Download the programming files on ISIS and follow the instructions.

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where $P(\text{error} | \mathbf{x}) = \min [P(\omega_1 | \mathbf{x}), P(\omega_2 | \mathbf{x})]$ is the probability of error for a particular input \mathbf{x} . Interestingly, while class posteriors $P(\omega_1 | \mathbf{x})$ and $P(\omega_2 | \mathbf{x})$ can often be expressed analytically and are integrable, the error function has discontinuities that prevent its analytical integration, and therefore, direct computation of the Bayes error.

(a) Show that the full error can be upper-bounded as follows:

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Note that the integrand is now continuous and corresponds to the harmonic mean of class posteriors weighted by $p(\mathbf{x})$.

case 1. $P(\omega_1 | \mathbf{x}) \leq P(\omega_2 | \mathbf{x})$

$$\min [P(\omega_1 | \mathbf{x}), P(\omega_2 | \mathbf{x})] = P(\omega_1 | \mathbf{x}) = \frac{1}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}} \leq \frac{2}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}}$$

case 2. $P(\omega_1 | \mathbf{x}) \geq P(\omega_2 | \mathbf{x})$

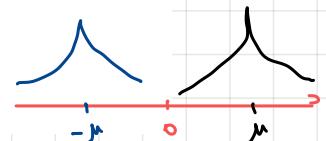
$$\min [P(\omega_1 | \mathbf{x}), P(\omega_2 | \mathbf{x})] = P(\omega_2 | \mathbf{x}) = \dots \leq \dots$$

(b) Show using this result that for the univariate probability distributions

$$p(x | \omega_1) = \frac{\pi^{-1}}{1 + (x - \mu)^2} \quad \text{and} \quad p(x | \omega_2) = \frac{\pi^{-1}}{1 + (x + \mu)^2},$$

the Bayes error can be upper-bounded by:

Cauchy distribution



(Hint: you can use the identity $\int \frac{1}{ax^2+bx+c} dx = \frac{2\pi}{\sqrt{4ac-b^2}}$ for $b^2 < 4ac$.)

$$P(\text{error}) \leq \int \frac{2}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}} p(\mathbf{x}) d\mathbf{x}$$

$P_1 = P(\omega_1)$
 $P_2 = P(\omega_2)$

$$= \int \frac{2}{\frac{P_1(x)}{P(x|\omega_1)P_1} + \frac{P_2(x)}{P(x|\omega_2)P_2}} P(x) dx$$

$$= \int \frac{2}{\frac{1}{P(x|\omega_1)P_1} + \frac{1}{P(x|\omega_2)P_2}} dx$$

$$= \int \frac{2}{\frac{1+(x-\mu)^2}{\pi^{-1}P_1} + \frac{1+(x+\mu)^2}{\pi^{-1}P_2}} dx$$

$$= \int \frac{2P_1P_2}{\pi [P_2(1+(x-\mu)^2) + P_1(1+(x+\mu)^2)]} dx$$

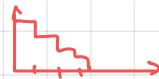
$$= 2\pi^{-1}P_1P_2 \int \frac{1}{\frac{(P_1+P_2)x^2+2\mu(P_1-P_2)x+1+\mu^2}{a} + b} dx \quad (P_1+P_2=1)$$

$$b^2 < 4ac \Rightarrow 4\mu^2(P_1-P_2)^2 < 4(P_1+P_2)(1+\mu^2) \quad \checkmark$$

$$\begin{aligned}
&= 2\pi^{-1} p_1 p_2 \cdot \frac{\sqrt{4(\underbrace{p_1+p_2}_{=1})(1+\mu^2) - 4\mu^2(p_1-p_2)^2}}{2\pi} \\
&= 2\pi^{-1} p_1 p_2 \cdot \frac{\sqrt{4(1+\mu^2) - 4\mu^2(p_1-p_2)^2}}{2\pi} \\
&\Rightarrow 2\pi^{-1} p_1 p_2 \cdot \frac{\sqrt{4 + 4\mu^2 [1 - (p_1-p_2)^2]}}{2\pi} \\
&= 2\pi^{-1} p_1 p_2 \cdot \frac{\sqrt{4 + 4\mu^2 [(p_1+p_2)^2 - (p_1-p_2)^2]}}{2\pi} \\
&= 2\pi^{-1} p_1 p_2 \cdot \frac{\sqrt{4 + 4\mu^2 [(p_1+p_2 - R_1+R_2)(p_1+p_2 + R_1-R_2)]}}{2\pi} \\
&\Rightarrow p_1 p_2 \cdot \frac{\sqrt{1 + 4\mu^2 p_1 p_2}}{2\pi}
\end{aligned}$$

- (c) Explain how you would estimate the error if there was no upper-bounds that are both tight and analytically integrable. Discuss following two cases: (1) the data is low-dimensional and (2) the data is high-dimensional.

(1) Riemann Sum $\int f(x)p(x)dx \approx \sum_i f(x_i)p(x_i)\Delta x$



(2) Monte Carlo

1. sample $x_i \sim p(x)$

2. $\int f(x)p(x)dx \approx \frac{1}{N} \sum f(x_i)$

Exercise 2: Bayes Decision Boundaries (15 + 15 P)

One might speculate that, in some cases, the generated data $p(x | \omega_1)$ and $p(x | \omega_2)$ is of no use to improve the accuracy of a classifier, in which case one should only rely on prior class probabilities $P(\omega_1)$ and $P(\omega_2)$ assumed here to be strictly positive.

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where $\mu, \sigma > 0$.

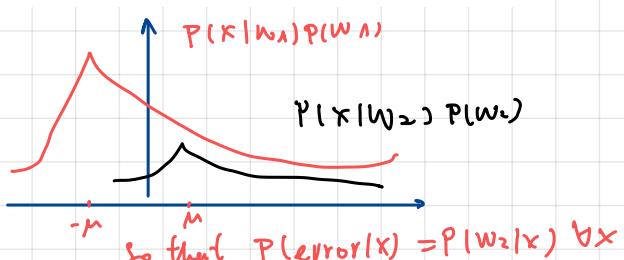
- (a) Determine for which values of $P(\omega_1), P(\omega_2), \mu, \sigma$ the optimal decision is to always predict the first class (i.e. under which conditions $P(\text{error} | x) = P(\omega_2 | x) \forall x \in \mathbb{R}$).

$$P(W_1|x) > P(W_2|x)$$

$$\Leftrightarrow P(x|\omega_1)P(\omega_1) > P(x|\omega_2)P(\omega_2)$$

$$\Leftrightarrow \frac{1}{2\sigma} e^{-\frac{|x-\mu|}{\sigma}} P(\omega_1) > \frac{1}{2\sigma} e^{-\frac{|x+\mu|}{\sigma}} P(\omega_2)$$

$$\Leftrightarrow -\frac{|x-\mu|}{\sigma} + \log P(\omega_1) > -\frac{|x+\mu|}{\sigma} + \log P(\omega_2)$$



$$\Leftrightarrow x \approx -\mu \Rightarrow -\frac{\mu-x}{\sigma} + \log P(\omega_1) > -\frac{-\mu-\mu}{\sigma} + \log P(\omega_2)$$

$$\Leftrightarrow 2\mu < \sigma(\log P(\omega_1) - \log P(\omega_2))$$

(the a. bound)

$$\textcircled{2} \quad -\mu < x < \mu \Rightarrow -\frac{\mu-x}{\delta} + \log P(\omega_1) > -\frac{x+\mu}{\delta} + \log P(\omega_2)$$

$$\Leftrightarrow -2x < \delta (\log P(\omega_1) - \log P(\omega_2))$$

$$\textcircled{3} \quad x > \mu \Rightarrow -2\mu < \delta (\log P(\omega_1) - \log P(\omega_2))$$

$$\left\{ (P(\omega_1), P(\omega_2), \mu, \delta) \mid \log \frac{P(\omega_1)}{P(\omega_2)} > \frac{2\mu}{\delta} \right\}$$

(b) Repeat the exercise for the case where the data for each class is generated by the univariate Gaussian probability distributions:

$$p(x | \omega_1) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \quad \text{and} \quad p(x | \omega_2) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x+\mu)^2}{2\sigma^2}\right).$$

where $\mu, \sigma > 0$.

$$P(\omega_1|x) > P(\omega_2|x)$$

$$P(x|\omega_1)P(\omega_1) > P(x|\omega_2)P(\omega_2)$$

$$-\frac{(x-\mu)^2}{2\sigma^2} + \log P(\omega_1) > -\frac{(x+\mu)^2}{2\sigma^2} + \log P(\omega_2)$$

$$-2x\mu < \delta^2 [\log P(\omega_1) - \log P(\omega_2)]$$

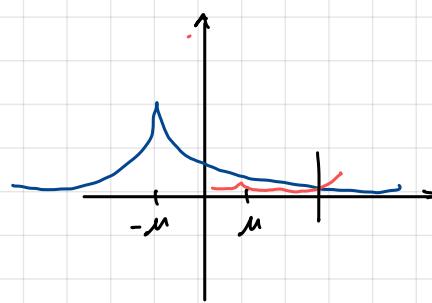
$$\Rightarrow \phi$$

$$\frac{1}{2}(-(x-\mu)^2 + (x+\mu)^2) > -\delta^2 \log \left(\frac{P(\omega_1)}{P(\omega_2)} \right)$$

$$\frac{1}{2}((x-\mu)^2 - (x+\mu)^2) < \delta^2 \log \left(\frac{P(\omega_1)}{P(\omega_2)} \right)$$

$$\frac{1}{2}(2x)(-2\mu) < \delta^2 \log \left(\frac{P(\omega_1)}{P(\omega_2)} \right)$$

$$-2x\mu < \delta^2 \log \left(\frac{P(\omega_1)}{P(\omega_2)} \right)$$



2. 右边的表达式：

$\sigma^2[\log P(\omega_1) - \log P(\omega_2)]$ 是一个固定的常数，记作 C 。

3. 是否存在解？

- 如果 $C > 0$, 即 $\log P(\omega_1) > \log P(\omega_2)$, 则 $\sigma^2[\log P(\omega_1) - \log P(\omega_2)]$ 是正数。但是, 由于 $-2x\mu$ 可正可负, 是否存在 x 使得它严格小于 C 取决于 x 的取值范围。

- 如果 $C < 0$, 即 $P(\omega_1) < P(\omega_2)$, 则 $\sigma^2[\log P(\omega_1) - \log P(\omega_2)]$ 是负数, 而左侧的 $-2x\mu$ 总是可以大于这个负数, 但题目要求的是 "总是" 预测 ω_1 , 即不等式应当对所有 x 成立。

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Note that the integrand is now continuous and corresponds to the harmonic mean of class posteriors weighted by $p(\mathbf{x})$.

$$\begin{aligned} \text{(a)} \quad P(\text{error}) &= \int P(\text{error} | \mathbf{x}) p(\mathbf{x}) d\mathbf{x} \\ &= \int \min [P(\omega_1 | \mathbf{x}), P(\omega_2 | \mathbf{x})] p(\mathbf{x}) d\mathbf{x} \quad \text{because of } \min[x_1, x_2, \dots, x_n] \leq \frac{n}{\sum \frac{1}{x_i}} \\ &= \int \frac{2}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}} p(\mathbf{x}) d\mathbf{x} \end{aligned}$$

- (b) Show using this result that for the univariate probability distributions

$$p(x | \omega_1) = \frac{\pi^{-1}}{1 + (x - \mu)^2} \quad \text{and} \quad p(x | \omega_2) = \frac{\pi^{-1}}{1 + (x + \mu)^2},$$

the Bayes error can be upper-bounded by:

$$P(\text{error}) \leq \frac{2 P(\omega_1) P(\omega_2)}{\sqrt{1 + 4\mu^2 P(\omega_1) P(\omega_2)}}$$

(Hint: you can use the identity $\int \frac{1}{ax^2 + bx + c} dx = \frac{2\pi}{\sqrt{4ac - b^2}}$ for $b^2 < 4ac$.)

$$\begin{aligned} P(\text{error}) &\leq \int \frac{2}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}} p(\mathbf{x}) d\mathbf{x} \\ &= \int \frac{2}{\frac{1}{P(\omega_1 | \mathbf{x})} + \frac{1}{P(\omega_2 | \mathbf{x})}} \frac{p(\mathbf{x})}{P(\mathbf{x})} d\mathbf{x} \\ &= \int \frac{\frac{1}{P(x | \omega_1) P(\omega_1)}}{\frac{P(x | \omega_1) P(\omega_1)}{P(\mathbf{x})}} + \frac{\frac{1}{P(x | \omega_2) P(\omega_2)}}{\frac{P(x | \omega_2) P(\omega_2)}{P(\mathbf{x})}} d\mathbf{x} \\ &= \int \frac{\frac{1}{P(x | \omega_1) P(\omega_1)}}{\frac{1}{P(x | \omega_1) P(\omega_1)}} + \frac{\frac{1}{P(x | \omega_2) P(\omega_2)}}{\frac{1}{P(x | \omega_2) P(\omega_2)}} d\mathbf{x} \\ &= \int \frac{\frac{\pi^{-1}}{(1 + (x - \mu)^2) P(\omega_1)}}{\frac{1}{P(x | \omega_1) P(\omega_1)}} + \frac{\frac{\pi^{-1}}{(1 + (x + \mu)^2) P(\omega_2)}}{\frac{1}{P(x | \omega_2) P(\omega_2)}} d\mathbf{x} \\ &= \int \frac{\frac{1}{\pi^{-1} P(\omega_1)}}{\frac{1 + (x - \mu)^2}{P(x | \omega_1) P(\omega_1)}} + \frac{\frac{1}{\pi^{-1} P(\omega_2)}}{\frac{1 + (x + \mu)^2}{P(x | \omega_2) P(\omega_2)}} d\mathbf{x} \end{aligned}$$

$$\begin{aligned}
&= \int \frac{\frac{2}{(\lambda + (x - \mu)^2) P(W_2) + (\lambda + (x + \mu)^2) P(W_1)}}{\pi^{-1} P(W_1) P(W_2)} dx \\
&= \int \frac{\pi^{-1} P(W_1) P(W_2)}{(\lambda + (x - \mu)^2) P(W_2) + (\lambda + (x + \mu)^2) P(W_1)} dx \\
&= \int \frac{2\pi^{-1} P(W_1) P(W_2)}{(\lambda + x^2 - 2x\mu + \mu^2) P(W_2) + (\lambda + x^2 + 2x\mu + \mu^2) P(W_1)} dx \\
&= \int \frac{2\pi^{-1} P(W_1) P(W_2)}{x^2(P(W_2) + P(W_1)) + \alpha(-2\mu P(W_2) + 2\mu P(W_1)) + (\lambda + \mu^2) P(W_2) + (\lambda + \mu^2) P(W_1)} dx
\end{aligned}$$

之后使用 $P(W_1) + P(W_2) = 1$
可以减少很多计算量

$$\begin{aligned}
4ac - b^2 &= 4(P(W_1) + P(W_2))[(\lambda + \mu^2)P(W_2) + (\lambda + \mu^2)P(W_1)] - (-2\mu P(W_2) + 2\mu P(W_1))^2 \\
&= 4(\lambda + \mu^2) \cdot (P(W_1) + P(W_2))^2 - 4\mu^2(P(W_1) - P(W_2))^2 \\
&= 4(\lambda + \mu^2)(P(W_1)^2 + 2P(W_1)P(W_2) + P(W_2)^2) - 4\mu^2(P(W_1)^2 - 2P(W_1)P(W_2) + P(W_2)^2) \\
&= 4P(W_1)^2 + 8P(W_1)P(W_2) + 4P(W_2)^2 + 4\mu^2 P(W_1)^2 + 8\mu^2 P(W_1)P(W_2) + 4\mu^2 P(W_2)^2 - 4\mu^2(P(W_1)^2 - 2P(W_1)P(W_2) + P(W_2)^2) \\
&= 4P^2(W_1) + (8 + 16\mu^2)P(W_1)P(W_2) + 4P^2(W_2) \geq 0
\end{aligned}$$

$$\begin{aligned}
P(\text{error}) &\leq \frac{(2\pi^{-1} P(W_1) P(W_2))}{[4(P(W_1) + P(W_2))[(\lambda + \mu^2)P(W_2) + (\lambda + \mu^2)P(W_1)] - (-2\mu P(W_2) + 2\mu P(W_1))^2]} \\
&= \frac{4P(W_1)P(W_2)}{\sqrt{4P^2(W_1) + (8 + 16\mu^2)P(W_1)P(W_2) + 4P^2(W_2)}} \\
&= \frac{2P(W_1)P(W_2)}{\sqrt{(P(W_1) + P(W_2))^2 + 4\mu^2 P(W_1)P(W_2)}} \\
&= \frac{2P(W_1)P(W_2)}{\sqrt{1 + 4\mu^2 P(W_1)P(W_2)}}
\end{aligned}$$

- (c) Explain how you would estimate the error if there was no upper-bounds that are both tight and analytically integrable. Discuss following two cases: (1) the data is low-dimensional and (2) the data is high-dimensional.

For low dimensional data, it's possible to estimate the posterior probability. And we could use the optimal decision function $\arg\min_j P(W_j|x)$ to calculate the error

For high dimensional data, the error computation would be very complex. To solve this, we could reduce the dimension and map the data from a high Dimension to a low Dimension.

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where $\mu, \sigma > 0$.

- (a) Determine for which values of $P(\omega_1), P(\omega_2), \mu, \sigma$ the optimal decision is to always predict the first class (i.e. under which conditions $P(\text{error} | x) = P(\omega_2 | x) \forall x \in \mathbb{R}$).

a) $p(W_1 | x) = P(W_2 | x)$

$$\frac{p(W_1 \cdot x)}{P(x)} = \frac{p(W_2 \cdot x)}{P(x)}$$

$$\frac{p(x | W_1) p(W_1)}{P(x)} = \frac{p(x | W_2) p(W_2)}{P(x)}$$

$$\frac{\frac{1}{2\sigma} e^{-\frac{|x-\mu|}{\sigma}}}{P(x)} p(W_1) = \frac{\frac{1}{2\sigma} e^{-\frac{|x+\mu|}{\sigma}}}{P(x)} p(W_2)$$

$$-\frac{|x-\mu|}{\sigma} + \ln p(W_1) = -\frac{|x+\mu|}{\sigma} + \ln p(W_2)$$

$$\delta \ln\left(\frac{p(W_1)}{p(W_2)}\right) = |x-\mu| - |x+\mu|$$

① $x < \mu$ and $x < -\mu$: $|x-\mu| - |x+\mu| = \mu - x + x + \mu = 2\mu > 0$

② $x < \mu$ and $x \geq -\mu$: $|x-\mu| - |x+\mu| = \mu - x - x - \mu = -2x \begin{cases} < 0 & \text{if } 0 < x < \mu \\ \geq 0 & \text{if } -\mu \leq x \leq 0 \end{cases}$

③ $x \geq \mu$: ($\mu > 0$) $|x-\mu| - |x+\mu| = x - \mu - x - \mu = -2\mu < 0$

Decision Boundary:

$$\min(p(W_1 | x), p(W_2 | x)) = \begin{cases} p(W_1 | x) & \text{if } x < -\mu < 0 \text{ and } -\mu < x \leq 0 \\ & (\text{x} < 0) \\ p(W_2 | x) & \text{if } x \geq \mu \text{ and } 0 < x < \mu \\ & (\text{x} > 0) \end{cases}$$

$$\ln\left(\frac{p(W_1)}{p(W_2)}\right) > 0 \Rightarrow p(W_1) > p(W_2) \Rightarrow p(W_1 | x) - p(W_2 | x) > 0$$

(b) Repeat the exercise for the case where the data for each class is generated by the univariate Gaussian probability distributions:

$$p(x | \omega_1) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) \quad \text{and} \quad p(x | \omega_2) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x + \mu)^2}{2\sigma^2}\right).$$

where $\mu, \sigma > 0$.