## **UCI** Paul Merage School of Business

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# MFIN 290: Financial Econometrics

Lecture 5-1



#### **Last Time**

- Cointegration
- Error Correction Models
- Yule Walker Equations, Moment Restrictions
- AR(2), ARMA(1,1) examples



# **Today**

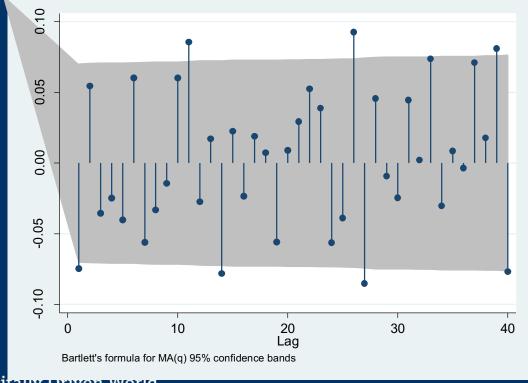
• Thus far in the course, when we have looked at time series effects and dependencies, we have focused on dependencies in the first moment.

• 
$$y = \beta y_{t-1} + u_t + \gamma u_{t-1}$$

How does the expected value of y depend on previous values of y or previous values of u?

- There is usually no autocorrelation in financial prices (why does this make sense)?
- But that doesn't mean there isn't anything we can say here...

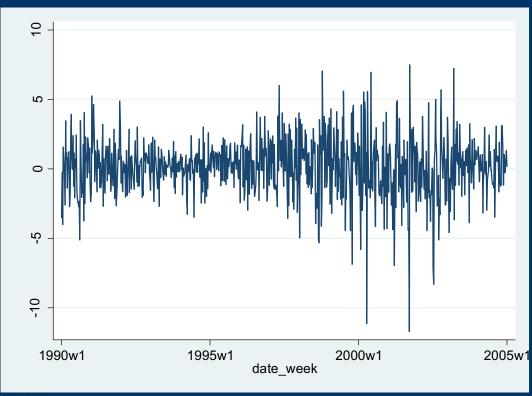
• Below: AC plot of weekly S&P 500 returns from Jan 1990 – Dec 2004. Nothing is a perfect textbook example, but this is pretty close...



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Below: Weekly S&P 500 returns from Jan 1990 – Dec 2004

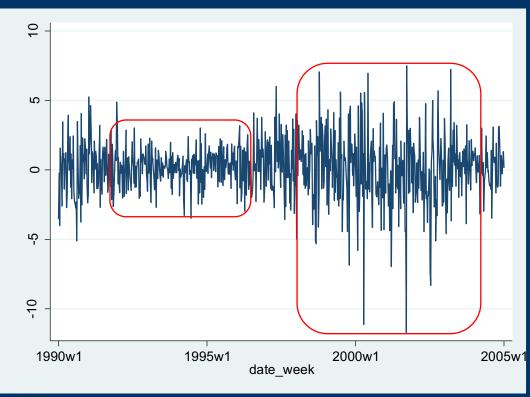


What do we see?

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Below: Weekly S&P 500 returns from Jan 1990 – Dec 2004



Clear periods of lower and higher volatility.

AND it looks like the periods of low and high volatility tend to follow one another.

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Consider a model given by

$$y_t = x_t \beta + e_t$$

$$\bullet \ e_t = u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}$$

• Where  $u_t$  is white noise, unit variance, uncorrelated, standard normal as usual

$$y_t = x_t \beta + e_t$$

$$\bullet \ e_t = u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}$$

- $E(u_t^2) = 1$
- $E(e_t|x_t, e_{t-1}) = 0$
- $E(e_t|x_t) = 0$
- $E(y_t|x_t) = x_t\beta$

- These are the classical regression assumptions. We can get  $\hat{eta}$  unbiased through OLS

$$y_t = x_t \beta + e_t$$

$$\bullet \ e_t = u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}$$

•  $var(e_t|e_{t-1}) = E(e_t^2|e_{t-1})$ 

$$= E(\left[u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}\right]^2 = \alpha_0 + \alpha_1 e_{t-1}^2$$

• Here, the **variance** of  $e_t$  is autoregressive. Called "Conditionally Heteroskedastic"

- "Conditionally Heteroskedastic"
- Not with respect to the x values, but with earlier values of  $e_t$  ...
- "Auto Regressive Conditionally Heteroskedastic"

- This has implications for pricing securities (such as options) whose value depends on future volatilities.
- If we model volatility as flat, when it is really conditionally heteroskedastic, our option prices will be systematically too high when times are quiet (assuming the option becomes more valuable when volatility increases) and too low when times are turbulent.
- You can see this in traded option prices relative to vanilla Black Scholes! Using a GARCH or ARCH model helps mitigate the implied vol smile for out of the money options.



• 
$$y_t = x_t \beta + e_t$$

$$\bullet \ e_t = u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}$$

$$\bullet \ e_t = u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}$$

• 
$$var(e_t|e_{t-1}) = E(e_t^2|e_{t-1})$$

$$= E(\left[u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}\right]^2 = \alpha_0 + \alpha_1 e_{t-1}^2$$

$$y_t = x_t \beta + e_t$$

$$\bullet \ e_t = u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}$$

• 
$$var(e_t) = \alpha_0 + \alpha_1 var(e_{t-1})$$

• If weakly stationary  $=> var(e_t) = var(e_{t-1}) \Rightarrow$ ?

• 
$$y_t = x_t \beta + e_t$$

$$\bullet \ e_t = u_t \sqrt{\alpha_0 + \alpha_1 e_{t-1}^2}$$

• 
$$var(e_t) = \frac{\alpha_0}{1-\alpha_1}$$

• Note:  $|\alpha_1|$  < 1 here for this to be finite and positive (like a variance has to be)

• But then...  $e_t$  is distributed normal with mean 0 and variance  $\frac{\alpha_0}{1-\alpha_1}$ 

• => we have our Gauss Markov Assumptions => OLS is the minimum variance linear unbiased estimator for  $\beta$ 

So can't we just run OLS here? What are we missing?

• There is a more efficient NON-LINEAR estimator!

We estimate these models through maximum likelihood (later in this course).

 Engle (1982) derived the function needed here (and won a Nobel prize for it!). The likelihood function is in his paper for those interested.

Higher Order Processes

• 
$$\sigma_t^2 = E(e_t^2 | e_{t-s} \text{ for all } S > 0)$$

• Given all of the relevant information known on period t... usually written compactly, say  $\psi_t$ :

$$\bullet \ \sigma_t^2 = E(e_t^2 | \psi_t)$$

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GARCH

$$y_t = x_t \beta + e_t$$

• 
$$e_t | \psi_t \sim N(0, \sigma_t^2)$$

GARCH

$$y_t = x_t \beta + e_t$$

•  $e_t | \psi_t \sim N(0, \sigma_t^2)$ 

- The conditional variance is defined as an ARMA process in  $e_t^2$
- GARCH(p,q) where p = order or AR component, q = order of MA component

- GARCH
- Some extra conditions are required to ensure that the specification is covariance stationary...
- Ensuring that  $\sigma_t^2$  is stable doesn't guarantee that all of the higher moments are.
- Bollerslev (1986) ("Generalized Autoregressive Conditional Heteroskedasticity" Journal of Econometrics) contains some of these conditions. Worth keeping this caveat in mind with these models and looking them up/checking as needed.
- Often the case that relatively simple GARCH models (such as (1,1)) perform as well characterizing real data as a longer (say order 8) ARCH.

ARCH in Mean

• What if the conditional mean of y depends in part on the variance?

ARCH in Mean

• What if the conditional mean of y depends in part on the variance?

- Intuition:
- The returns on a portfolio should be increasing with its risk, which may vary over time.

ARCH in Mean

$$y_t = \beta x_t + \delta \sigma_t^2 + e_t$$

•  $var(e_t|\psi_t) = ARCH(q)$ 

• This makes a lot of sense, but is pretty rare.... Why don't we see this more?

- ARCH in Mean
- $\boxed{\bullet \ y_t = \beta x_t + \delta \sigma_t^2 + e_t}$
- $var(e_t|\psi_t) = ARCH(q)$
- Mis-specification of the variance function will give you inconsistent estimates of  $\sigma_t^2$ , which can give you inconsistent estimates of  $\beta$ !
- Unlike in weighted least squares/GLS from earlier in the course here, the "right" weights are determined by (and thus correlated with) the disturbances themselves.
- ⇒ If there is mis-specification in our ARCH model, the coefficients we care about will be inconsistent.

ARCH in Mean

$$y_t = \beta x_t + \delta \sigma_t^2 + e_t$$

• 
$$var(e_t|\psi_t) = ARCH(q)$$

• Can have an efficient estimate in one case, but it might be inconsistent?

Sounds like a Hausman test!

- T-ARCH
- Idea is that the effect of the residual will be asymmetric. "Good" news may have less of an impact on volatility than "bad" news.
- Conditional Variance function is asymmetric:

• 
$$h_t = \alpha_0 + \alpha_1 e_{t-1}^2 + \gamma D_{t-1} e_{t-1}^2 + \beta h_{t-1} + u_t + others \dots$$

$$D_t = \begin{cases} 1 e_t < 0 \\ 0 e_t \ge 0 \end{cases}$$

```
Title
    [TS] arch — Autoregressive conditional heteroskedasticity (ARCH) family of estimators
Syntax
        arch depvar [indepvars] [if] [in] [weight] [, options]
    options
                                Description
    Model
      noconstant
                                suppress constant term
     arch(numlist)
                                ARCH terms
     garch(numlist)
                                GARCH terms
      saarch (numlist)
                                simple asymmetric ARCH terms
      tarch (numlist)
                                threshold ARCH terms
     aarch(numlist)
                                asymmetric ARCH terms
      narch (numlist)
                                nonlinear ARCH terms
     narchk(numlist)
                                nonlinear ARCH terms with single shift
      abarch(numlist)
                                absolute value ARCH terms
      atarch (numlist)
                                absolute threshold ARCH terms
      sdgarch (numlist)
                                lags of s t
      earch(numlist)
                                new terms in Nelson's EGARCH model
      egarch (numlist)
                                lags of ln(s t^2)
     parch(numlist)
                                power ARCH terms
     tparch (numlist)
                                threshold power ARCH terms
     aparch (numlist)
                                asymmetric power ARCH terms
      nparch (numlist)
                                nonlinear power ARCH terms
     nparchk (numlist)
                                nonlinear power ARCH terms with single shift
     pgarch (numlist)
                                power GARCH terms
      constraints (constraints) apply specified linear constraints
                                keep collinear variables
      collinear
```

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#### **Extensions to ARCH**

```
Model 2
  archm
                            include ARCH-in-mean term in the mean-equation specification
 archmlags(numlist)
                            include specified lags of conditional variance in mean equation
 archmexp(exp)
                            apply transformation in exp to any ARCH-in-mean terms
 arima(#p, #d, #q)
                            specify ARIMA(p,d,q) model for dependent variable
 ar(numlist)
                            autoregressive terms of the structural model disturbance
 ma(numlist)
                            moving-average terms of the structural model disturbances
Model 3
  distribution(dist [#])
                            use dist distribution for errors (may be qaussian, normal, t, or qed; default is qaussian)
 het(varlist)
                            include varlist in the specification of the conditional variance
  savespace
                            conserve memory during estimation
Priming
 arch0(xb)
                            compute priming values on the basis of the expected unconditional variance; the default
 arch0(xb0)
                            compute priming values on the basis of the estimated variance of the residuals from OLS
 arch0(xbwt)
                            compute priming values on the basis of the weighted sum of squares from OLS residuals
                            compute priming values on the basis of the weighted sum of squares from OLS residuals, with more weight at earlier times
 arch0(xb0wt)
 arch0(zero)
                            set priming values of ARCH terms to zero
 arch0(#)
                            set priming values of ARCH terms to #
 arma0(zero)
                            set all priming values of ARMA terms to zero; the default
 arma0(p)
                            begin estimation after observation p, where p is the maximum AR lag in model
                            begin estimation after observation q, where q is the maximum MA lag in model
 arma0(q)
                            begin estimation after observation (p + q)
 arma0 (pq)
                            set priming values of ARMA terms to #
 arma0(#)
  condobs (#)
                            set conditioning observations at the start of the sample to #
```

 Usually we aren't concerned with volatility for its own sake, we want it for something else (option pricing is one example).

 ARCH style models are usually estimated, and the easiest way (IMO) to demonstrate the ARCH impacts is through simulation.

 A good exercise would be to see the effect an ARCH or GARCH model would have on option prices for the S&P 500 and compare that to actual observed prices...

#### **ARCH Simulation**

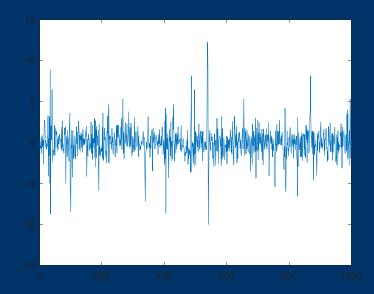
MATLAB!

```
clear
• alpha = [1 \ 0.8];
• eps = zeros(1000,1);
• u = mvnrnd(0, 1, 1000);
• T = [1:1:1000]';
• for i = 2:1000
      eps(i) = u(i)*(alpha(1) + alpha(2)*(eps(i-1))^2)^(0.5);
■ End
plot(T,eps)
• eps sq = eps.^2;
plot(T,eps sq)
autocorr (eps)
autocorr (eps sq)
parcorr(eps sq)
```

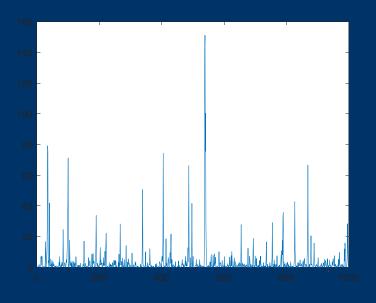
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## **ARCH Simulation**

- plot(T,eps)
- plot(T,eps\_sq)



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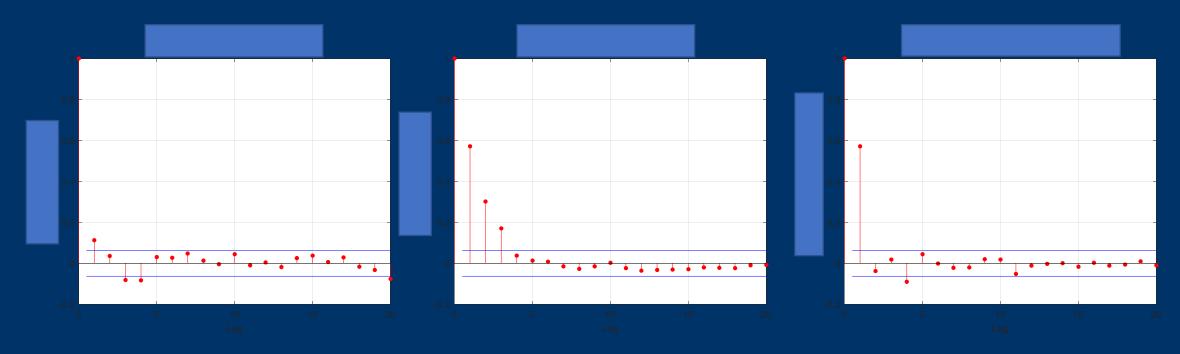


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# Which Graph is Which?

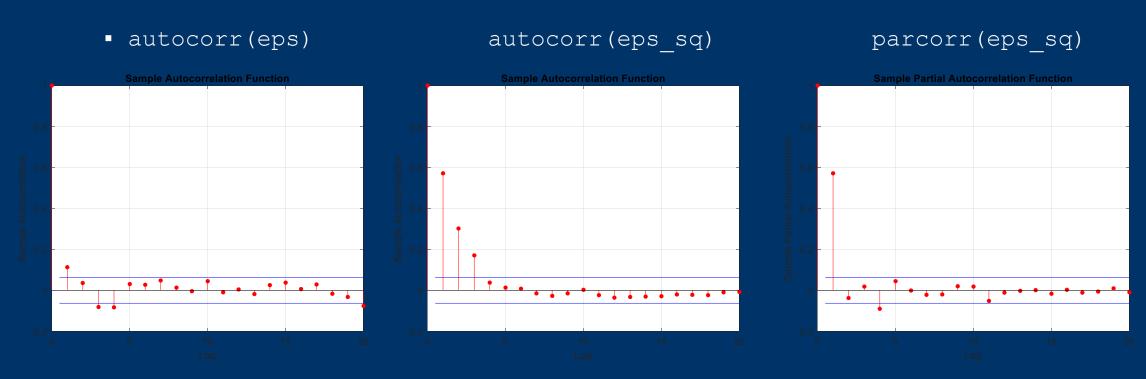
- autocorr(eps)
- autocorr(eps\_sq)
- parcorr(eps\_sq)



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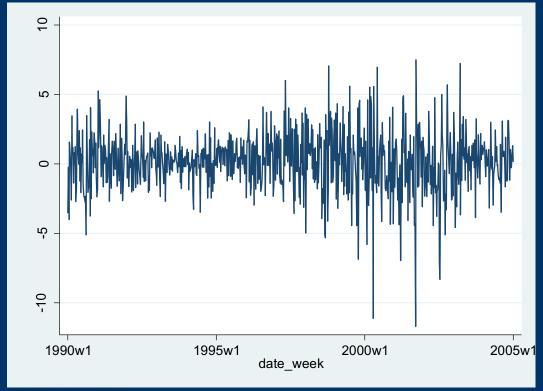


# Which Graph is Which?



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 You should always inspect the data first. Sometimes the eyeball test can be the most obvious!



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- Can we be more quantitative?
- Idea: We can get consistent estimates (exc. ARCH in mean) with plain vanilla OLS.
- Use the estimated error terms to see if there is any autocorrelation (or MA terms) in a regression of  $e_t^2$  (the fitted values) on a constant and the q lagged values.
- If there are no ARCH effects, a joint test of coefficient significance will show it.

$$\widehat{e_t^2} = \alpha_0 + \alpha_1 e_{t-1}^{2} + v_t$$

•  $H_0$ :  $\alpha_1 = 0$  (no ARCH effects)

- Test statistic =  $(T q)R^2$
- T = number of obs
- Q = number of lags in the regression
- $R^2 = R^2$  from the regression above. Distributed  $\chi_q^2$ .
- This form is generic for longer lags with one lag, a t-test gets you to the same place.

- Just like all heteroskedasticity, keep an eye on relevance.
- Depending on the application, OLS may yield an unbiased estimate, just not an efficient one.
- How much we care about characterizing the volatility series is usually a function of:
- 1) The question being asked
- 2) The consequences to mis-specifying the volatility relationship

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## **Testing for ARCH effects**

Example using S&P data

Source	SS	df	MS		mber of obs		781
Model	0	0			ob > F		0.00
Residual	3598.50032	780	4.6134619				0.0000
Total	3598.50032	780	4.6134619		j R-squared ot MSE		0.0000 2.1479
r	Coef.	Std. Err.	t	P> t	[95% Co	nf.	Interval]
_cons	.1582282 n, lags(1/5)	.0768578	2.06	0.040	.007355		
_cons	.1582282	.0768578	2.06	0.040	.007355	5	.3091008 chi2
_cons estat archlr test for a lags(p)	.1582282  a, lags(1/5) atoregressive chi	.0768578  conditiona	2.06  l heteroske	0.040	.007355	5 b >	.3091008
_cons estat archlr test for an lags(p)	.1582282  a, lags(1/5) atoregressive  chi 50.7	.0768578  conditiona 2 41	2.06  l heteroske  df	0.040	.007355	5 b >	.3091008 chi2
_cons estat archlr test for an lags(p)	.1582282  n, lags(1/5) itoregressive  chi  50.7 50.9	.0768578  conditiona 2 41 56	2.06  l heteroske  df  1 2	0.040	.007355 city (ARCH) Pro	b > .000	.3091008  chi2 0
_cons estat archlr test for an lags(p)	.1582282  a, lags(1/5) atoregressive  chi 50.7	.0768578  conditiona 2 41 56 78	2.06  l heteroske  df	0.040	.007355	5 b >	.3091008  chi2 0 0

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- Example using S&P data
- Two points to think about.
- Second seems MA-style, less
- plausible...

