

## Chapter 4 Branching process

"Population" dynamics are often modelled by branching process.

Assume that each unit in the population at time  $n$  produces offspring according to some distribution.

Each member of the population produces offspring

- ① independently
- ② according the same distribution

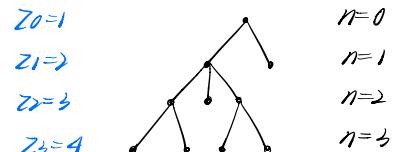
$$\begin{aligned}\underline{\alpha} &= (\alpha_0, \alpha_1, \dots) \\ &= (\Pr(X_1=0), \Pr(X_1=1), \dots)\end{aligned}$$

where  $X_i = \# \text{offspring}$

Let  $Z_n$  be the number of population units at time  $n$  (or generation  $n$ ). Then  $\{Z_n\}$  is a branching process.

$\{Z_n\}$  can be represented as a Markov chain

$$Z_n \in N = \{0, 1, 2, \dots\}$$



$$\Pr(Z_{n+1}=i_{n+1} | Z_n=i_n, Z_{n-1}=i_{n-1}, \dots, Z_0=i_0) = \Pr(Z_{n+1}=i_{n+1} | Z_n=i_n)$$

$$\begin{cases} \text{If } \alpha_0=0, \text{ then } Z_{n+1} \geq Z_n \quad (\text{Every pair produces at least one offspring}) \\ \text{If } \alpha_0=1, \text{ then } Z_1=0 \text{ and the population is extinct at generation } n+1 \end{cases}$$

For the rest of chapter 4, we will assume  $0 < \alpha_0 < 1$

Fact: 0 is an absorbing state

Lemma: All non-zero states are transient

$$\Pr(Z_{n+1}=0 | Z_n=i) = \alpha_0^i \quad \text{Probability that I return to } i \text{ at all.}$$

$$\Leftrightarrow \Pr(Z_1>0 | Z_0=i)$$

If  $n=1$  and

$Z_1=0$ , you cannot get back to  $Z_2=i$

$$\downarrow$$

$$1 - \Pr(Z_1=0 | Z_0=i) = 1 - \alpha_0^i < 1$$

## 4.2 Possibilities for our process.

$$\begin{cases} E(X_i) = \mu \\ \text{Var}(X_i) = \sigma^2. \end{cases}$$

- ① Get absorbed
- ② Process grows without bound.

### Mean generation size

Let  $Z_n = \sum_{i=1}^{Z_{n-1}} X_i$  where  $X_i$  is the number of offspring for the  $i$ th member of  $(n-1)^{\text{st}}$  generation

- $Z_n$  is a sum of iid random variable  $X_1, \dots, X_{Z_{n-1}}$
- $Z_n$  is a random sum of iid random variables

$E(X_i) = \mu = \sum_{k=0}^{\infty} k \Pr(X_i=k)$  mean of the offspring distribution

$$\begin{aligned} E_{Z_{n-1}}(Z_n) &= E_{Z_{n-1}} E_{Z_{n-1}|Z_{n-1}}[Z_n | Z_{n-1}] \\ &= \sum_{k=0}^{\infty} \left\{ E_{Z_{n-1}|Z_{n-1}=k} \left[ \sum_{i=1}^k X_i | Z_{n-1}=k \right] \right\} \times \Pr(Z_{n-1}=k) \\ &= \sum_{k=0}^{\infty} E\left(\sum_{i=1}^k X_i\right) \Pr(Z_{n-1}=k) \\ &= \sum_{k=0}^{\infty} k \mu \Pr(Z_{n-1}=k) \\ &= \mu \sum_{k=0}^{\infty} k \Pr(Z_{n-1}=k) E_{Z_{n-1}}(Z_{n-1}) \\ &= \mu E_{Z_{n-1}}(Z_{n-1}) \\ &= \mu [ \mu E_{Z_{n-2}}(Z_{n-2}) ] \\ &\quad \cdots \\ &= \mu^n E_{Z_0}(Z_0) \end{aligned}$$

If  $Z_0=1$  with probability 1, then  $E_{Z_0}(Z_0) = \mu^n$ .

$$\text{As } n \rightarrow \infty : \lim_{n \rightarrow \infty} E(Z_n) = \lim_{n \rightarrow \infty} \mu^n = \begin{cases} 0 & \text{if } \mu < 1 \\ 1 & \text{if } \mu = 1 \\ \infty & \text{if } \mu > 1 \end{cases}$$

The probability that the population is extinct by generation  $n$  is

$$\begin{aligned} P(E) = P(Z_n=0) &= 1 - P(Z_n > 0) = 1 - \sum_{k=1}^{\infty} \Pr(Z_n=k) \\ &\geq 1 - \sum_{k=1}^{\infty} k \Pr(Z_n=k) = 1 - E(Z_n) = 1 - \mu^n. \end{aligned}$$

For subcritical process ( $\mu < 1$ ),  $\lim_{n \rightarrow \infty} P(Z_n=0) \geq \lim_{n \rightarrow \infty} 1 - \mu^n = 1$   
 $\Rightarrow P(E)=1$  with probability 1, a subcritical process eventually goes extinct.

## Variance

$$\begin{aligned}
 \text{Var}_{Zn}(Z_n) &= \text{Var}_{Z_{n-1}}[E_{Z_{n-1}}(Z_n)] + E_{Z_{n-1}}[\text{Var}_{Z_{n-1}}(Z_n)] \\
 &\quad \downarrow \\
 &\text{Var}_{Z_{n-1}}(\mu_{Z_{n-1}}) \\
 &= \alpha^2 \text{Var}_{Z_{n-1}}(Z_{n-1}) \\
 &\quad \downarrow \\
 &E_{Z_{n-1}}[\text{Var}_{Z_{n-1}Z_{n-1}}(\sum_{i=1}^{Z_{n-1}} X_i | Z_{n-1})] \\
 &= E_{Z_{n-1}}[Z_{n-1} \alpha^2] \\
 &= \alpha^2 E_{Z_{n-1}}(Z_{n-1}) \\
 &= \alpha^2 \mu^{n-1}
 \end{aligned}$$

$$\Rightarrow \text{Var}(Z_n) = \mu^2 \text{Var}(Z_{n-1}) + \alpha^2 \mu^{n-1}.$$

If  $Z_0=1$  with probability 1.

$$\text{Var}(Z_0) = 0$$

$$\text{Var}(Z_1) = \mu^2 \cdot 0 + \alpha^2 \mu^0 = \alpha^2$$

$$\text{Var}(Z_2) = \mu^2 \cdot \alpha^2 + \alpha^2 \mu = \mu(1+\mu)\alpha^2.$$

$$\text{Var}(Z_3) = \mu^2 (\mu(1+\mu)\alpha^2) + \alpha^2 \mu^2 = \alpha^2 \mu^2 (1+\mu+\mu^2)$$

...

$$\text{Var}(Z_n) = \alpha^2 \mu^{n-1} \sum_{k=0}^{n-1} \mu^k.$$

	$E(Z_n)$	$\text{Var}(Z_n)$	
$\mu=1$	1	$n\alpha^2$	Critical process
$\mu<1$	$\mu^n \xrightarrow{n \rightarrow \infty} 0$	$\frac{\alpha^2 \mu^{n-1} (\mu^{n-1})}{n-1} \xrightarrow{n \rightarrow \infty} 0$	Subcritical process.
$\mu>1$	$\mu^n \xrightarrow{n \rightarrow \infty} \infty$	$\frac{\alpha^2 \mu^{n-1} (\mu^{n-1})}{n-1} \xrightarrow{n \rightarrow \infty} \infty$	Supercritical process

## 4.5 Probability generating functions

For a discrete random variable  $X$  taking values in  $\{0, 1, \dots\}$ , the probability generating function is

$$G(s) = E(s^X) = \sum_{k=0}^{\infty} s^k P(X=k) \quad \text{A power series with coefficients summing to 1.}$$

$$= P(X=0) + P(X=1)s + P(X=2)s^2 + \dots$$

①  $G_X(1) = \sum_{k=0}^{\infty} \Pr(X=k) = 1$

② Series converges absolutely for  $|s| \leq 1$ .

Note: Similar to MGF (moment generating function)

$$M_X(t) = E(e^{xt}) = E((e^t)^X)$$

e.g. geometric distribution  $X \sim \text{Geom}(p)$  failure

$$G(s) = E(s^X) = \sum_{k=1}^{\infty} s^k p(1-p)^{k-1} = sp \sum_{k=1}^{\infty} (s(1-p))^{k-1} = \frac{sp}{1-s(1-p)} \text{ for } |s| < 1$$

### ① Find Probability

$$G(0) = \Pr(X=0)$$

$$G'(0) = \sum_{k=1}^{\infty} k s^{k-1} \Pr(X=k) = 0 + \sum_{k=1}^{\infty} k s^{k-1} \Pr(X=k) = \Pr(X=1)$$

$$G''(0) = \sum_{k=2}^{\infty} k(k-1) s^{k-2} \Pr(X=k) = 2 \Pr(X=2)$$

$$G^{(j)}(0) = \sum_{k=j}^{\infty} k(k-1) \dots (k-j+1) s^{k-j} \Pr(X=k) = j! \Pr(X=j)$$

$$\Rightarrow \Pr(X=j) = \frac{G^{(j)}(0)}{j!}$$

### ② Sums of independent random variables

Let  $Y = X_1 + X_2 + \dots + X_n$

$$G_Y(s) = E(s^Y) = E(s^{X_1+X_2+\dots+X_n}) = E(\prod_{k=1}^n s^{X_k}) = \prod_{k=1}^n E(s^{X_k}) = \prod_{k=1}^n G_{X_k}(s)$$

$\Rightarrow$  If  $X_1, X_2, \dots, X_n$  are iid, then

$$G_Y(s) = [G_X(s)]^n \text{ where } X_1, X_2, \dots, X_n \sim F.$$

### ③ Moments

$$\begin{aligned} a_X^{(0)}(s) &= \sum_{k=0}^{\infty} k s^{k-1} \Pr(X=k) \\ \Rightarrow a_X^{(1)}(1) &= \sum_{k=0}^{\infty} k \Pr(X=k) = E(X). \\ a_X^{(2)}(1) &= \sum_{k=0}^{\infty} k(k-1) \Pr(X=k) = E(X^2) - E(X)^2 \end{aligned}$$

### Properties of PAF

- ① (a)  $a(1)=1$
- (b)  $P(X=k) = \frac{a^{(k)}(0)}{k!}$  for  $k \geq 0$
- (c)  $E(X) = a'(1)$
- (d)  $\text{Var}(X) = a''(1) + a'(1) - a'(1)^2$
- ② If  $X$  and  $Y$  are random variables s.t.  $a_X(s) = a_Y(s)$  for all  $s$ , then  $X$  and  $Y$  have the same distribution.
- ③ If  $X$  and  $Y$  are independent, then  $a_{X+Y}(s) = a_X(s) a_Y(s)$ .

44 Extinction is forever!

PAF for  $Z_n$  : # of population units at generation  $n$ .

$$a_{Z_n}(s) = \sum_{k=0}^{\infty} s^k \Pr(Z_n=k)$$

Let  $a_X(s) = \sum_{k=0}^{\infty} s^k a_k$  be PAF of the offspring distribution.

$$\begin{aligned} \text{so } a_{Z_n}(s) &= E(s \sum_{i=1}^{Z_{n-1}} X_i) \quad \text{by } Z_n = \sum_{i=1}^{Z_{n-1}} X_i \\ &= E_{Z_{n-1}} [E(s \sum_{i=1}^{Z_{n-1}} X_i) | Z_{n-1}] \end{aligned}$$

$$\text{Then, } E(s \sum_{i=1}^{Z_{n-1}} X_i) | Z_{n-1}=z = E(s \sum_{i=1}^z X_i) = \prod_{i=1}^z a_X(s) = [a_X(s)]^z.$$

$$\Rightarrow \text{An}(s) = E_{Z=1} [(\text{Ax}(s))^{Z+1}] \quad \text{If } |s| \leq 1, \quad \text{Ax}(s) \leq 1.$$

$$= E_{Z=1} [u^{Z+1}]$$

$$= \text{An}_1(u)$$

$$= \text{An}_1(\text{Ax}(s))$$

$$\text{An}(s) = \sum_{k=0}^{\infty} s^k \Pr(Z_0=k) = s \text{ if } Z_0=1 \text{ with probability 1.}$$

$$\text{A}_1(s) = \text{A}_0(\text{Ax}(s)) = \text{Ax}(s)$$

$$\text{A}_2(s) = \text{A}_1(\text{Ax}(s)) = \text{Ax}(\text{Ax}(s)) = \text{Ax}(\text{A}_1(s))$$

$$\text{A}_3(s) = \text{A}_2(\text{Ax}(s)) = \text{A}_1(\text{Ax}(\text{Ax}(s))) = \text{Ax}(\text{Ax}(\text{Ax}(s))) = \text{A}_1(\text{A}_2(s))$$

$$\dots$$

$$\text{A}_n(s) = \text{A}_{n-1}(\text{Ax}(s)) = \underbrace{\text{Ax}(\text{Ax}(\text{Ax}(\dots)))}_{n \text{ Ax's}} = \text{Ax}(\text{A}_{n-1}(s))$$

Theorem For a branching process, let  $\text{Ax}(s)$  be the PAF for the off spring distribution. The probability of extinction ever is the smallest positive root of

$$s = \text{Ax}(s).$$

Important

If  $\mu \leq 1$  (subcritical / critical), the extinction probability is equal to 1.

e.g. Geometric (# of failures)

$$a_k = (1-p)^k p$$

$$Ax(s) = \sum_{k=0}^{\infty} s^k (1-p)^k p = p \sum_{k=0}^{\infty} s^k (1-p)^k = \frac{p}{1-s(1-p)}$$

$$M = Ax'(s) = \frac{p(1-p)}{(1-s(1-p))}, \quad |s=1| = \frac{p(1-p)}{p} = \frac{1-p}{p}$$

$\begin{cases} \text{if } p \geq \frac{1}{2}, \quad M \leq 1 \\ \text{if } p < \frac{1}{2}, \quad M > 1 \end{cases}$  super critical process

$$s = \frac{p}{1-s(1-p)}$$

$$\Rightarrow \frac{p}{1-s(1-p)} - s = 0$$

$$\Rightarrow s^2(1-p) - s + p = 0$$

$$\Rightarrow s = \frac{1 \pm \sqrt{1+4p^2-4p}}{2(1-p)} = \frac{1 \pm (1-p)}{2(1-p)}$$

$$\Rightarrow \text{smaller positive root} \quad \begin{cases} \frac{p}{1-p} & p < \frac{1}{2} \\ 1 & p \geq \frac{1}{2} \end{cases}$$

Let  $\ell_n = P(Z_n=0)$  denote the probability that the population goes extinct by generation  $n$ . We have

$$\ell_n = P(Z_n=0) = \ell_n(0) = \alpha_x(\ell_{n+1}(0)) = \alpha_x(P(Z_{n+1}=0)) = \alpha_x(\ell_{n+1})$$

$$\Rightarrow \lim_{n \rightarrow \infty} \ell_n = \lim_{n \rightarrow \infty} \alpha_x(\ell_{n+1}) \quad \text{Note: } \ell = \lim_{n \rightarrow \infty} \ell_n$$

$$\Rightarrow \ell = \alpha_x(\lim_{n \rightarrow \infty} \ell_{n+1}) = \alpha_x(\ell)$$

$\Rightarrow \ell$  is a root of  $s = \alpha_x(s)$ .

Let  $x$  be a positive solution of  $s = \alpha_x(s)$ . We need to show that  $\ell \leq x$ .

Since  $\alpha_x(s) = \sum_{k=0}^{\infty} s^k P(X=k)$  is an increasing function on  $(0, 1]$  and  $0 < x$

$$\ell_1 = P(Z_1=0) = \ell_1(0) = \alpha_x(0) \quad \text{H} Z_0=1$$

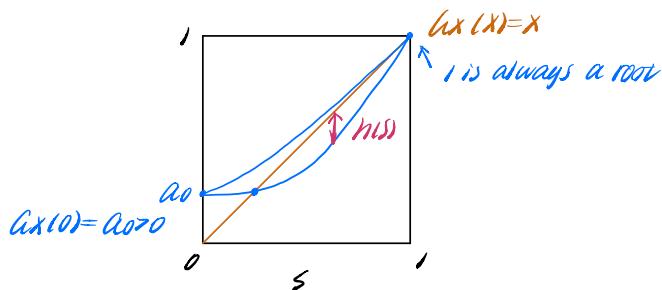
$$\leq \alpha_x(x) = x$$

$\alpha_x(s)$  is nondecreasing.

By induction, assuming  $\ell_k \leq x$ , for all  $k < n$

$$\begin{aligned} \ell_n &= P(Z_n=0) = \ell_n(0) = \alpha_x(\ell_{n+1}(0)) = \alpha_x(\ell_{n+1}) \\ &\leq \alpha_x(x) = x \end{aligned}$$

$\Rightarrow$  Take limits  $\lim_{n \rightarrow \infty} \ell_n \leq \lim_{n \rightarrow \infty} x \Rightarrow \ell \leq x$ .



$$\alpha''_x(s) = \sum_{k=0}^{\infty} k(k-1)s^{k-2} P(X=k) \geq 0 \Rightarrow \alpha_x(s) \text{ is convex}$$

$$\alpha'_x(s) = \sum_{k=0}^{\infty} ks^{k-1} P(X=k) \text{ increasing function of } s$$

$$\Rightarrow \alpha'_x(s) \leq \alpha'_x(1) = \mu$$

$$\text{Let } h(s) = s - \alpha_x(s) \quad \text{At } s, \text{ the distance between } s \text{ and } \alpha_x(s)$$

$$h'(s) = 1 - \alpha'_x(s)$$

positive below  
 0 on the line  
 negative above

$$\underline{\mu \leq 1} \quad \begin{cases} h(1) = 0 \\ h(s) < 0 \text{ for } s < 1 \end{cases} \Rightarrow \alpha_X(s) > s \Rightarrow \ell = 1 \\ \downarrow \\ \mu = \alpha'_X(s) \leq 1 \Rightarrow h'(s) \geq 0 \Rightarrow h(s) \text{ increasing}$$

$$\underline{\mu > 1} \quad \begin{cases} h(0) = 0 - \mu < 0 \\ h(s) = s - \alpha_X(s) \\ h(1) = 0 \end{cases} \Rightarrow \exists \text{ some } t \in (0, 1) \text{ s.t. } h(t) = 0 \\ \mu = \alpha'_X(s) > 1 \Rightarrow h'(s) < 0 \Rightarrow h(s) \text{ decreasing}$$

e.g. Let  $\alpha = (\frac{0}{4}, \frac{1}{4}, \frac{1}{2})$ ,  $\mu_X = 0 \cdot \frac{1}{4} + 1 \cdot \frac{1}{4} + 2 \cdot \frac{1}{2} = \frac{5}{4} > 1$

$$\alpha_X(s) = \sum_{k=0}^{\infty} s^k \Pr(X=k) = \frac{1}{4} + \frac{1}{4}s + \frac{1}{2}s^2$$

$$\alpha_X(s) - s = \frac{1}{4} - \frac{3}{4}s + \frac{1}{2}s^2 = 0$$

$$\Rightarrow (2s-1)(s-1) = 0 \quad s = \left\{ \frac{1}{2}, 1 \right\}$$

Not the probability  $X_1 = 0$ , which is  $\frac{1}{4}$ .

Probability of extinction  $\ell = \frac{1}{2} = \sum_{n=1}^{\infty} \Pr(T_n = 0)$  where  
 $T_n = \min \{ n : Z_n = 0 \}$