

第四章

1. 一箱产品中有 3 件正品和 2 件次品，不放回任取两件，X 表示得到的次品数，求平均次品数 $E(X)$

X	0	1	2
p_i	$\frac{C_3^2}{C_5^2} = \frac{3}{10}$	$\frac{C_3^1 C_2^1}{C_5^2} = \frac{3}{5}$	$\frac{C_2^2}{C_5^2} = \frac{1}{10}$

$$E(X) = 0 \cdot \frac{3}{10} + 1 \cdot \frac{3}{5} + 2 \cdot \frac{1}{10} = \frac{4}{5}$$

2. 已知随机变量 $X \sim P(\lambda)$ ，试求 $E(\frac{1}{1+X})$

$$\begin{aligned} E(\frac{1}{1+X}) &= \sum_{k=0}^{+\infty} \frac{1}{1+k} \cdot \frac{\lambda^k}{k!} e^{-\lambda} = \frac{1}{\lambda} \sum_{k=0}^{+\infty} \frac{\lambda^{k+1}}{(k+1)!} e^{-\lambda} \\ &= \frac{1}{\lambda} \left(\sum_{k=0}^{+\infty} \frac{\lambda^k}{k!} e^{-\lambda} - e^{-\lambda} \right) = \frac{1}{\lambda} (1 - e^{-\lambda}) \end{aligned}$$

3. 设随机变量 X 的概率密度为 $f(x) = \begin{cases} x, & 0 \leq x \leq 1 \\ 2-x, & 1 < x \leq 2 \\ 0, & \text{其他} \end{cases}$ ，试求 $E(X)$ 和 $D(X)$.

$$\begin{aligned} E(X) &= \int_{-\infty}^{+\infty} x f(x) dx = \int_0^1 x^2 dx + \int_1^2 x(2-x) dx \\ &= \frac{x^3}{3} \Big|_0^1 + [x^2 - \frac{x^3}{3}]_1^2 = 1 \end{aligned}$$

$$\begin{aligned} E(X^2) &= \int_{-\infty}^{+\infty} x^2 f(x) dx = \int_0^1 x^3 dx + \int_1^2 x^2(2-x) dx \\ &= \frac{x^4}{4} \Big|_0^1 + [\frac{2}{3}x^3 - \frac{x^4}{4}]_1^2 = \frac{7}{6} \end{aligned}$$

$$D(X) = E(X^2) - E^2(X) = \frac{7}{6} - 1 = \frac{1}{6}$$

4. 地面雷达搜索飞机，在时间 $(0, t)$ 内发现飞机的概率是 $P(t) = 1 - e^{-\lambda t}$, $(\lambda > 0)$ ，试求发现飞机所需的平均搜索时间。

5. 设发现飞机所需时间为 T，则

$$\begin{aligned} F(t) = P\{T \leq t\} &= \begin{cases} P\{0 < T \leq t\} = P(t) = 1 - e^{-\lambda t} & t > 0 \\ 0 & t \leq 0 \end{cases} \\ f(t) = F'(t) &= \begin{cases} \lambda e^{-\lambda t} & t > 0 \\ 0 & t \leq 0 \end{cases} \end{aligned}$$

可见 T 服从参数为 λ 的指数分布，平均搜索时间为

$$\begin{aligned} E(T) &= \int_{-\infty}^{+\infty} t f(t) dx = \int_0^{+\infty} t \lambda e^{-\lambda t} dx \\ &\quad \underline{\text{分部积分}} - e^{-\lambda t} [t + 1/\lambda]_0^{+\infty} = 1/\lambda \end{aligned}$$

5. 设随机变量 X 服从几何分布: $P(X=k)=pq^k (q=1-p), k=0,1,2,\dots$, 求 $E(X), D(X)$.

$$\text{解: } E(X) = \sum_{k=0}^{+\infty} kP\{X=k\} = \sum_{k=1}^{+\infty} kpq^k = pq \sum_{k=1}^{+\infty} kq^{k-1} = pq \left[\left(\sum_{k=1}^{+\infty} x^k \right)' \right]_{x=q} = \frac{q}{p}$$

$$\begin{aligned} E(X^2) &= \sum_{k=0}^{+\infty} k^2 P\{\xi=k\} = \sum_{k=1}^{+\infty} k^2 pq^k = pq \sum_{k=1}^{+\infty} k^2 q^{k-1} \\ &= pq \sum_{k=1}^{+\infty} k(k+1-1)q^{k-1} = pq \sum_{k=1}^{+\infty} k(k+1)q^{k-1} - pq \sum_{k=1}^{+\infty} kq^{k-1} \\ &= pq \left[\left(\sum_{k=1}^{+\infty} x^{k+1} \right)'' \right]_{x=q} - \frac{q}{p} = pq \frac{2}{p^3} - \frac{q}{p} = \frac{2q-pq}{p^2} = \frac{(2-p)q}{p^2} \end{aligned}$$

$$D(X) = E(X^2) - E^2(X) = \frac{(2-p)q}{p^2} - \frac{q^2}{p^2} = \frac{q}{p^2}$$

6. 随机变量 X 的概率密度为 $f(x) = \begin{cases} e^{-x}, & x > 0 \\ 0, & x \leq 0 \end{cases}$, 试求 $Y=2X$ 和 $Z=e^{-2X}$ 的数学期望。

$$E(Y) = \int_{-\infty}^{+\infty} 2xf(x)dx = \int_0^{+\infty} 2xe^{-x}dx = -2e^{-x}[x+1]_0^{+\infty} = 2$$

$$E(Z) = \int_{-\infty}^{+\infty} e^{-2x} f(x)dx = \int_0^{+\infty} e^{-3x} dx = [-\frac{1}{3}e^{-3x}]_0^{+\infty} = \frac{1}{3}$$

7. 设某种产品每周的需求量 $X \sim U(10, 30)$, 而经销商进货数量为区间 $[10, 30]$ 中的某一整数。商店每销售一件商品可获利 500 元; 若供大于求则削价处理, 每处理一件商品亏损 100 元; 若供不应求可从外部调货, 但此时每件商品仅获利 300 元。为使该商店每周所获平均利润至少为 9280 元, 试确定最少进货量。

解: 设商店的进货量为 n ($10 \leq n \leq 30$), 则商店每周所获利润为

$$g(X) = \begin{cases} 500X - 100(n-X) = 600X - 100n, & X \leq n, \\ 500n + 300(X-n) = 300X + 200n, & X > n \end{cases}$$

$$\begin{aligned} E[g(X)] &= \int_{-\infty}^{+\infty} g(x)f_X(x)dx = \frac{1}{20} \int_{10}^{30} g(x)dx \\ &= \frac{1}{20} \left[\int_{10}^n (600x - 100n)dx + \int_n^{30} (300x + 200n)dx \right] \\ &= \frac{1}{20} [-150n^2 + 7000n + 105000] \geq 9280 \\ &\Rightarrow \frac{62}{3} \leq n \leq 26 \Rightarrow n = 21 \end{aligned}$$

8. 设 (X, Y) 的联合概率密度为 $f(x, y) = \begin{cases} 12y^2, & 0 \leq y \leq x \leq 1 \\ 0, & \text{其它} \end{cases}$, 求 $E(X), E(Y), E(XY), E(X^2+Y^2)$.

$$E(X) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xf(x, y) dx dy = \int_0^1 dx \int_0^x 12xy^2 dy = \frac{4}{5}$$

$$E(Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} yf(x, y) dx dy = \int_0^1 dx \int_0^x 12y^3 dy = \frac{3}{5}$$

也可以先算X,Y的边缘密度再求期望,但较麻烦

$$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xyf(x, y) dx dy = \int_0^1 dx \int_0^x 12xy^3 dy = \frac{1}{2}$$

$$\begin{aligned} E(X^2 + Y^2) &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (x^2 + y^2) f(x, y) dx dy \\ &= \int_0^1 dx \int_0^x (x^2 + y^2) \cdot 12y^2 dy = \frac{16}{15} \end{aligned}$$

9. 设随机变量 X_1, X_2, \dots, X_n 相互独立, 都服从区间 $(0, \theta)$ 上的均匀分布, 求 $Y = \max\{X_1, X_2, \dots, X_n\}$ 的数学期望和方差。

$$\text{解: 若 } X \sim U(0, \theta), \text{ 则 } F_X(x) = \begin{cases} 0, & x < 0 \\ x/\theta, & 0 \leq x \leq \theta \\ 1, & x \geq \theta \end{cases}, \quad f_X(x) = \begin{cases} 1/\theta, & 0 \leq x \leq \theta \\ 0, & \text{其它} \end{cases}$$

$$F_Y(y) = P\{Y \leq y\} = P\{\max\{X_1, X_2, \dots, X_n\} \leq y\}$$

$$= P\{X_1 \leq y, X_2 \leq y, \dots, X_n \leq y\}$$

$$= P\{X_1 \leq y\} P\{X_2 \leq y\} \cdots P\{X_n \leq y\}$$

$$= [F_X(y)]^n$$

$$f_Y(y) = F'_Y(y) = n[F_X(y)]^{n-1} f_X(y) = \begin{cases} 0, & \text{其它} \\ n\left(\frac{y}{\theta}\right)^{n-1}, & 0 \leq y \leq \theta \end{cases}$$

$$E(Y) = \int_{-\infty}^{+\infty} y f_Y(y) dy = \int_0^\theta y \frac{n}{\theta} \left(\frac{y}{\theta}\right)^{n-1} dy = \frac{n}{\theta^n} \int_0^\theta y^n dy = \frac{n}{n+1} \theta$$

$$E(Y^2) = \int_0^\theta y^2 \frac{n}{\theta} \left(\frac{y}{\theta}\right)^{n-1} dy = \frac{n}{\theta^n} \int_0^\theta y^{n+1} dy = \frac{n}{n+2} \theta^2$$

$$D(Y) = E(Y^2) - E^2(Y) = \frac{n(n+1)^2 - n^2(n+2)}{(n+2)(n+1)^2} \theta^2 = \frac{n^3 + n^2 + n}{(n+2)(n+1)^2} \theta^2$$

10. 民航机场的送客汽车载有 20 名乘客, 从机场开出, 乘客可以在 10 个车站下车, 如果到达某一车站时无顾客下车, 则在该站不停车。设随机变量 X 表示停车次数, 假定每个乘客在各个车站下车是等可能的, 求平均停车次数。

$$\text{解: 设 } X_i = \begin{cases} 0, & \text{第 } i \text{ 站有人下车} \\ 1, & \text{第 } i \text{ 站无人下车} \end{cases}, \text{ 于是停车次数 } X = \sum_{i=1}^{10} X_i$$

$$P\{X_i = 0\} = \frac{9^{20}}{10^{20}}, \text{ 于是 } E(X) = \sum_{i=1}^{10} E(X_i) = \sum_{i=1}^{10} P(X_i = 1) = 10(1 - (\frac{9}{10})^{20})$$

11. 证明: 对取值于区间 (a, b) 内的随机变量 X, 恒成立不等式:

$$a \leq E(X) \leq b, \quad D(X) \leq \frac{(b-a)^2}{4}$$

$$\text{证: } a = \int_a^b a dF_X(x) \leq E(X) = \int_{-\infty}^{+\infty} x dF_X(x) = \int_a^b x dF_X(x) \leq \int_a^b b dF_X(x) = b$$

令 $g(x) = E[(X-x)^2]$, 知 $g(x)$ 在 $x = E(X)$ 处取得最小值 $D(X) = E[(X - E(X))^2]$,

$$\text{取 } x = \frac{a+b}{2},$$

$$\begin{aligned} D(X) &\leq g\left(\frac{a+b}{2}\right) = E\left[\left(X - \frac{a+b}{2}\right)^2\right] \\ &= \int_{-\infty}^{+\infty} \left(x - \frac{a+b}{2}\right)^2 dF_X(x) = \int_a^b \left(x - \frac{a+b}{2}\right)^2 dF_X(x) \leq \int_a^b \left(b - \frac{a+b}{2}\right)^2 dF_X(x) \\ &= \int_a^b \left(\frac{b-a}{2}\right)^2 dF_X(x) = \frac{(b-a)^2}{4} \end{aligned}$$

12. 证明: 如果随机变量 X, Y 相互独立, 则

$$D(XY) = D(X)D(Y) + E^2(X)D(Y) + E^2(Y)D(X)$$

证明: $E(XY) = E(X)E(Y)$

$$E(X^2Y^2) = E(X^2)E(Y^2) = [D(X) + E^2(X)][D(Y) + E^2(Y)]$$

$$D(XY) = E(X^2Y^2) - E^2(XY)$$

$$\begin{aligned} &= [D(X) + E^2(X)][D(Y) + E^2(Y)] - E^2(X)E^2(Y) \\ &= D(X)D(Y) + E^2(X)D(Y) + E^2(Y)D(X) \end{aligned}$$

13. 设 (X, Y) 的联合概率密度为 $f(x, y) = \begin{cases} 1, & 0 \leq x \leq 1, |y| < x \\ 0, & \text{其他} \end{cases}$, 判断 X 与 Y 的相关性和独立性。

$$E(X) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xf(x, y) dx dy = \int_0^1 dx \int_{-x}^x x dy = \frac{2}{3}$$

$$E(Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} yf(x, y) dx dy = \int_0^1 dx \int_{-x}^x y dy = 0$$

$$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xyf(x, y) dx dy = \int_0^1 dx \int_{-x}^x xy dy = 0$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 0 \Rightarrow \rho_{XY} = 0$$

所以 X, Y 不相关

$$\text{又 } f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} 2x, & 0 < x < 1 \\ 0, & \text{其它} \end{cases}, f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} 1 - |y|, & -1 < y < 1 \\ 0, & \text{其它} \end{cases}$$

可见, 在区域 $D = \{0 \leq x \leq 1, |y| < x\}$ 上, $f(x, y) \neq f_X(x)f_Y(y)$, X, Y 不相互独立。

14. 设随机变量 (X, Y) 的联合分布律为

	Y	0	1
X			
0		0.1	0.15
1		0.25	0.2
2		0.15	0.15

(1) 求 X, Y 的协方差矩阵。

(2) 求 $Z = \sin\left(\frac{\pi}{2}(X+Y)\right)$ 的数学期望。

解: (1) $E(X) = 1 \times (0.25 + 0.2) + 2 \times (0.15 + 0.15) = 1.05$

$$E(Y) = 1 \times (0.15 + 0.2 + 0.15) = 0.5$$

$$E(XY) = 0 \times 0 \times 0.1 + 0 \times 1 \times 0.15 + 1 \times 0 \times 0.25 + 1 \times 1 \times 0.2 + 2 \times 0 \times 0.15 + 2 \times 1 \times 0.15 = 0.5$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = -0.025$$

$$E(X^2) = 1^2 \times (0.25 + 0.2) + 2^2 \times (0.15 + 0.15) = 1.65$$

$$E(Y^2) = 1^2 \times (0.15 + 0.2 + 0.15) = 0.5$$

$$D(X) = E(X^2) - E^2(X) = 0.5475, \quad D(Y) = E(Y^2) - E^2(Y) = 0.25$$

$$\text{协方差矩阵 } C = \begin{pmatrix} 0.5475 & -0.025 \\ -0.025 & 0.25 \end{pmatrix}$$

(2)

$$\begin{aligned} E(Z) &= E\left[\sin\left(\frac{\pi}{2}(X+Y)\right)\right] \\ &= \sin\left(\frac{\pi}{2}(0+0)\right) \times 0.1 + \sin\left(\frac{\pi}{2}(0+1)\right) \times 0.15 + \sin\left(\frac{\pi}{2}(1+0)\right) \times 0.25 \\ &\quad + \sin\left(\frac{\pi}{2}(1+1)\right) \times 0.2 + \sin\left(\frac{\pi}{2}(2+0)\right) \times 0.15 + \sin\left(\frac{\pi}{2}(2+1)\right) \times 0.15 \\ &= 1 \times 0.15 + 1 \times 0.25 - 1 \times 0.15 = 0.25 \end{aligned}$$

15. 设 $D(X)=25$, $D(Y)=36$, 相关系数 $\rho_{XY} = 0.4$, 试求: $D(X+Y)$ 和 $D(X-Y)$.

$$\text{Cov}(X, Y) = \rho_{XY} \sqrt{D(X)} \sqrt{D(Y)} = 0.4 \times 5 \times 6 = 12$$

$$D(X+Y) = D(X) + D(Y) + 2\text{Cov}(X, Y) = 25 + 36 + 2 \times 12 = 85$$

$$D(X-Y) = D(X) + D(Y) - 2\text{Cov}(X, Y) = 25 + 36 - 2 \times 12 = 37$$

16. 已知 (X, Y) 的联合密度函数为 $f(x, y) = \begin{cases} 3x, & 0 < y < x, 0 < x < 1, \\ 0, & \text{其他} \end{cases}$,

求 $E(X|Y)$ 和 $D(X|Y = \frac{1}{2})$.

$$\text{解: } f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_y^1 3x dx = \frac{3}{2}(1-y^2), & 0 < y < 1 \\ 0, & \text{其他} \end{cases}$$

$$\text{当 } 0 < y < 1 \text{ 时, } f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{2x}{1-y^2}, & y < x < 1 \\ 0, & \text{其他} \end{cases}$$

$$E(X|Y=y) = \int_{-\infty}^{+\infty} x f_{X|Y}(x|y) dx = \int_y^1 x \frac{2x}{1-y^2} dx = \frac{2(1+y+y^2)}{3(1+y)}$$

$$E(X|Y) = \frac{2(1+Y+Y^2)}{3(1+Y)}$$

$$E(X|Y=1/2) = \frac{7}{9},$$

$$D(X|Y=\frac{1}{2}) = \int_{-\infty}^{+\infty} (x - \frac{7}{9})^2 f_{X|Y}(x|y) dx = \int_{\frac{1}{2}}^1 (x - \frac{7}{9})^2 \frac{2x}{1-\frac{1}{4}} dx = \frac{13}{648}$$

(写出积分式子即可)

17. 设每天到站的货物件数 N 的分布律为:

N	10	11	12	13	14	15
P	0.05	0.1	0.1	0.2	0.35	0.2

若每天到达的货物次品率均为 0.1, 用 X 表示每天到达的货物中次品的件数, 求 $E(X)$ 。

$$\begin{aligned}\text{解: } E(X) &= E[E(X|N)] = \sum_{n=10}^{15} E(X|N=n)P(N=n) \\ &= 10 \times 0.1 \times 0.05 + 11 \times 0.1 \times 0.1 + 12 \times 0.1 \times 0.1 \\ &\quad + 13 \times 0.1 \times 0.2 + 14 \times 0.1 \times 0.35 + 15 \times 0.1 \times 0.2 = 1.33\end{aligned}$$

18. 小猫走进一个山洞, 其中有三个门洞。第一个走 2 小时回到地面, 第二个走 3 小时重回山洞, 第三个走 5 小时重回山洞。若小猫随机选择一个门洞, 求它回到地面的平均时间。

解 设小猫回到地面的时间为 X , 小猫选定的通道为 Y , Y 服从三点均匀分布。小猫回到地面的平均时间为

$$\begin{aligned}E(X) &= E[E(X|Y)] = \sum_{n=1}^3 E(X|Y=n)P(Y=n) \\ &= E(X|Y=1)P(Y=1) + E(X|Y=2)P(Y=2) + E(X|Y=3)P(Y=3) \\ &= 2 \times \frac{1}{3} + [3 + E(X)] \times \frac{1}{3} + (5 + E(X)) \times \frac{1}{3} \Rightarrow E(X) = 10\end{aligned}$$

19. 设 X, Y 相互独立且都服从正态分布 $N(\mu, \sigma^2)$, 令 $Z_1 = \alpha X + \beta Y, Z_2 = \alpha X - \beta Y$,

(1) 求 ρ_{Z_1, Z_2} ; (2) 确定 (Z_1, Z_2) 的联合分布; (3) 讨论 Z_1 与 Z_2 的独立性。

解一 已知

$$\begin{bmatrix} X \\ Y \end{bmatrix} \sim N(\mu, C) = N\left(\begin{pmatrix} \mu \\ \mu \end{pmatrix}, \begin{bmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{bmatrix}\right)$$

由正态分布的线性变换不变性, 其线性变换也服从联合正态分布。

$$Z = \begin{bmatrix} Z_1 \\ Z_2 \end{bmatrix} = \begin{bmatrix} \alpha & \beta \\ \alpha & -\beta \end{bmatrix} \begin{bmatrix} X \\ Y \end{bmatrix} = B \begin{bmatrix} X \\ Y \end{bmatrix},$$

向量 Z 的均值向量和协方差矩阵分别为

$$\begin{aligned}E \begin{bmatrix} Z_1 \\ Z_2 \end{bmatrix} &= B\mu = \begin{bmatrix} \alpha & \beta \\ \alpha & -\beta \end{bmatrix} \begin{bmatrix} \mu \\ \mu \end{bmatrix} = \begin{bmatrix} (\alpha + \beta)\mu \\ (\alpha - \beta)\mu \end{bmatrix} \\ C_Z &= BCB^T = \begin{pmatrix} \alpha & \beta \\ \alpha & -\beta \end{pmatrix} \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{pmatrix} \begin{pmatrix} \alpha & \beta \\ \alpha & -\beta \end{pmatrix}^T = \sigma^2 \begin{pmatrix} \alpha^2 + \beta^2 & \alpha^2 - \beta^2 \\ \alpha^2 - \beta^2 & \alpha^2 + \beta^2 \end{pmatrix}\end{aligned}$$

有 $D(Z_1) = (\alpha^2 + \beta^2)\sigma^2$ $D(Z_2) = (\alpha^2 + \beta^2)\sigma^2$ $Cov(Z_1, Z_2) = (\alpha^2 - \beta^2)\sigma^2$

$$(1) \quad \rho_{Z_1, Z_2} = \frac{Cov(Z_1, Z_2)}{\sqrt{D(Z_1)D(Z_2)}} = \frac{(\alpha^2 - \beta^2)\sigma^2}{(\alpha^2 + \beta^2)\sigma^2} = \frac{\alpha^2 - \beta^2}{\alpha^2 + \beta^2}$$

(2) 代入参数得 Z 的联合概率密度

$$\varphi(x, y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)}\left[\frac{(x-\mu_1)^2}{\sigma_1^2} - 2\rho\frac{(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2}\right]\right\}, \quad (x, y) \in R^2,$$

(3) 因变换矩阵的行列式 $|\mathbf{B}| = -2\alpha\beta \neq 0 \Rightarrow \alpha \neq 0, \beta \neq 0$ 时, Z 为非奇异 (非退化) 二维正态随机向量的满秩线性变换, 仍服从非退化二维正态分布, 二者相互独立等价于不相关, 故 Z_1 与 Z_2 相互独立的充要条件是

$$\frac{\alpha^2 - \beta^2}{\alpha^2 + \beta^2} = 0 \quad \text{即 } \alpha^2 = \beta^2, \quad \alpha \neq 0, \beta \neq 0$$

或 直接验证 Z 的协方差矩阵 C_Z 非奇异的成立条件.

$$\begin{aligned} \text{解二 (1)} \quad \text{Cov}(Z_1, Z_2) &= \text{Cov}(\alpha X + \beta Y, \alpha X - \beta Y) = \alpha^2 D(X) - \beta^2 D(Y) = (\alpha^2 - \beta^2)\sigma^2 \\ D(Z_1) &= D(\alpha X + \beta Y) = \alpha^2 D(X) + \beta^2 D(Y) = (\alpha^2 + \beta^2)\sigma^2 \\ D(Z_2) &= D(\alpha X - \beta Y) = \alpha^2 D(X) + \beta^2 D(Y) = (\alpha^2 + \beta^2)\sigma^2 \\ \rho_{Z_1 Z_2} &= \frac{\text{Cov}(Z_1, Z_2)}{\sqrt{D(Z_1)D(Z_2)}} = \frac{(\alpha^2 - \beta^2)\sigma^2}{(\alpha^2 + \beta^2)\sigma^2} = \frac{\alpha^2 - \beta^2}{\alpha^2 + \beta^2} \end{aligned}$$

(2) 根据题意知 $(X, Y) \sim N(\mu, \sigma^2; \mu, \sigma^2; 0)$, 由正态分布的线性变换不变性知 (Z_1, Z_2) 仍服从二维正态分布.

$$E(Z_1) = E(\alpha X + \beta Y) = \alpha E(X) + \beta E(Y) = (\alpha + \beta)\mu$$

$$E(Z_2) = E(\alpha X - \beta Y) = \alpha E(X) - \beta E(Y) = (\alpha - \beta)\mu$$

结合第 (1) 问的结果得

$$(Z_1, Z_2) \sim N((\alpha + \beta)\mu, (\alpha^2 + \beta^2)\sigma^2; (\alpha - \beta)\mu, (\alpha^2 + \beta^2)\sigma^2; \frac{\alpha^2 - \beta^2}{\alpha^2 + \beta^2})$$

由此可写出 (Z_1, Z_2) 的联合概率密度.

(3) 因 (Z_1, Z_2) 服从二维正态分布, 故 Z_1 与 Z_2 相互独立的充要条件是 $\frac{\alpha^2 - \beta^2}{\alpha^2 + \beta^2} = 0$

即 $\alpha^2 = \beta^2$

20. 设二维正态随机变量, $(X, Y) \sim N(1, 3^2; 0, 4^2; -\frac{1}{2})$, 设 $Z = \frac{X}{3} + \frac{Y}{2}$, 试求:

(1) Z 的数学期望和方差; (2) ρ_{XZ} ; (3) 判断 X 与 Z 的独立性.

. 1)由正态分布的边缘分布还是正态分布可知

$$(X,Y) \sim N(1,3^2, 0,4^2, -\frac{1}{2}) \Rightarrow X \sim N(1,3^2), Y \sim N(0,4^2)$$

$$\therefore E(X)=1, D(X)=3^2=9, E(Y)=0, D(Y)=4^2=16$$

$$E(Z) = E(\frac{X}{3} + \frac{Y}{2}) = \frac{1}{3}E(X) + \frac{1}{2}E(Y) = \frac{1}{3} \times 1 + \frac{1}{2} \times 0 = \frac{1}{3}$$

$$\rho_{XY} = -\frac{1}{2} \Rightarrow Cov(X,Y) = \rho_{XY} \sqrt{D(X)} \sqrt{D(Y)} = -\frac{1}{2} \times 3 \times 4 = -6$$

$$D(Z) = D(\frac{X}{3} + \frac{Y}{2}) = D(\frac{X}{3}) + D(\frac{Y}{2}) + 2Cov(\frac{X}{3}, \frac{Y}{2})$$

$$= \frac{1}{9}D(X) + \frac{1}{4}D(Y) + \frac{1}{3}Cov(X,Y) = 3$$

$$(2) Cov(X, Z) = Cov(X, \frac{X}{3} + \frac{Y}{2}) = \frac{1}{3}Cov(X, X) + \frac{1}{2}Cov(X, Y)$$

$$= \frac{1}{3}D(X) + \frac{1}{2}Cov(X, Y) = \frac{1}{3} \times 9 + \frac{1}{2} \times (-6) = 0$$

$$\therefore \rho_{XZ} = 0$$

(3) $\because (X,Y) \sim N(1,3^2, 0,4^2, -\frac{1}{2})$, 而 $Z = \frac{X}{3} + \frac{Y}{2}$ 是 X 和 Y 的非零线性

组合, $\therefore (X, Z)$ 也服从正态分布 (书上159页性质3). 又由(2)

的结果知 $\rho_{XZ} = 0$. 故 X 和 Z 相互独立。
