

## Markov Inequality

For RV  $X$  s.t.  $\Pr(X \geq 0) = 1$ ,

$$\Pr(X \geq t) \leq \mathbb{E}(X) / t$$

## Chebyshev Inequality

If  $\text{Var}(X) < \infty$ ,

$$\Pr(|X - \mathbb{E}(X)| \geq t) \leq \frac{\text{Var}(X)}{t^2}$$

$$\text{Proof: } \Pr((X - \mathbb{E}(X))^2 \geq t^2) \leq \mathbb{E}((X - \mathbb{E}(X))^2) / t^2 = \frac{\text{Var}(X)}{t^2}$$

## Sample Mean

For  $n$  RV  $X_1, \dots, X_n$ , sample mean  $\bar{X}_n = \frac{1}{n} \sum X_i$

If  $X_1, \dots, X_n$  random sample from distr w/

mean  $\mu$  and variance  $\sigma^2$ ,  $\bar{X}_n$  is sample mean,

$$\mathbb{E}(\bar{X}_n) = \mu, \quad \text{Var}(\bar{X}_n) = \sigma^2 / n$$

$$(\text{By Chebyshev, } \Pr(|\bar{X}_n - \mu| \geq t) \leq \sigma^2 / nt^2)$$

# Convergence of RV

## - Convergence in distribution

Sequence of RV  $Z_1, Z_2, \dots$  w/ CDF  $F_1, F_2, \dots$

converge to RV  $Z$  w/ CDF  $F$  if:

$$\lim_{n \rightarrow \infty} F_n(x) = F(x), \text{ for every } x \text{ s.t. } F \text{ continuous}$$

## - Convergence in Probability

Sequence of RV  $Z_1, Z_2, \dots$  converges to  $b$  in prob if:

$$\text{for every } \varepsilon > 0, \lim_{n \rightarrow \infty} \Pr(|Z_n - b| < \varepsilon) = 1$$

( $Z_n \xrightarrow{P} b$ )  $\downarrow$  only need  $Z_n$  be close to  $b$

## - Convergence with Probability 1 (almost surely)

$$\dots \text{ converges to } b \text{ if: } \Pr(\lim_{n \rightarrow \infty} Z_n = b) = 1$$

( $Z_n \xrightarrow{a.s.} b$ )

$\rightarrow$  need any sequence of  $\{Z_n\}$  converge to  $b$

Almost sure convergence  $\Rightarrow$  Convergence in prob  $\Rightarrow$  Convergence in distr

Example of converge in distr but not w/ prob 1:

$$\text{Let } Z_n = \begin{cases} 1 & \text{with } \Pr \frac{1}{n} \\ 0 & 1 - \frac{1}{n} \end{cases}$$

Then  $Z_n \rightarrow 0$  in prob.

However, since  $\sum_{n=1}^{\infty} \Pr(Z_i = 1) = +\infty$ , events  $\{Z_i = 1\}$  independent,

However, since  $\sum_{n=1}^{\infty} \Pr(Z_i = 1) = +\infty$ , and events  $\{Z_i = 1\}$

by Second Borel-Cantelli lemma,  $\Pr(\lim_{n \rightarrow \infty} Z_n = b) \neq 1$ .

## Weak Law of Large Numbers

Suppose  $X_1, \dots, X_n$  form a random sample from a distr w/ mean  $\mu$ , and finite variance. Let  $\bar{X}_n$  be sample mean,

then,  $\bar{X}_n \xrightarrow{P} \mu$ .

- If  $Z_n \xrightarrow{P} b$ ,  $g(z)$  is a function continuous at  $b$ ,  
 $g(Z_n) \xrightarrow{P} g(b)$ .

## Strong Law of Large Numbers

$$\bar{X}_n \xrightarrow{\text{a.s.}} \mu$$

## Chernoff Bounds

Idea: Want to generate a family of bounds for  $\Pr(X)$ .

Consider in Chebyshev, we applied Markov on  $[X - \mathbb{E}(X)]^2$

We can repeat that with a functional family of  $X$  with parameter  $\lambda$ . This will give a family of bounds parameterized by  $\lambda$ . Then optimize over  $\lambda$  to get the best bound.

Derivation: for  $t \in \mathbb{R}$ ,

$$\Pr(X \geq t) = \Pr(e^{\lambda X} \geq e^{\lambda t}) \leq \frac{\mathbb{E}(e^{\lambda X})}{e^{\lambda t}} = e^{-\lambda t} \overset{\text{MGF of } X}{\psi(\lambda)}, \quad \forall \lambda > 0.$$

$$\text{Take union bound: } \Pr(X \geq t) \leq \min_{\lambda > 0} e^{-\lambda t} \psi(\lambda).$$

## Central Limit Theorem (Lindberg & Lévy)

If RV  $X_1, \dots, X_n$  form a random sample from a given distr w/ mean  $\mu$ , var  $\sigma^2 < \infty$ , then for each fixed number  $x$ ,

$$\lim_{n \rightarrow \infty} \Pr\left(\frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \leq x\right) = \Phi(x)$$

CDF of  $N(0,1)$

(Standardized sample mean  $\sim N(0,1)$ )

Interpretation: The distribution of standardized sample mean converges in distribution to normal.

Usage: Estimate  $\Pr(|\bar{X} - \mathbb{E}X| \leq e) \approx \Phi\left(\frac{e}{\sigma/\sqrt{n}}\right) - \Phi\left(-\frac{e}{\sigma/\sqrt{n}}\right)$

# The Delta Method

(Computing approx distr for function of RV, esp. function of sample mean)

Let  $Y_1, Y_2, \dots$  be sequence of RV,  $F^*$  be a cdf.

Let  $\theta \in \mathbb{R}$ ,  $\{a_i\}$  s.t.  $a_i > 0$ ,  $a_n \rightarrow \infty$ .

Suppose  $a_n(Y_n - \theta)$  converges in distr to  $F^*$ ,

let  $\alpha(\cdot)$  be a function w/ continuous derivative s.t.  $\alpha'(\theta) \neq 0$ ,

then,  $a_n[\alpha(Y_n) - \alpha(\theta)] / \alpha'(\theta)$

converges in distr to  $F^*$ .

Proof outline:

$a_n \rightarrow \infty$ ,  $a_n(Y_n - \theta) \xrightarrow{d} F^*$ , then  $Y_n$  close to  $\theta$  as  $n \rightarrow \infty$ .

(Because for any CDF,  $F^*(N) \rightarrow 0$  as  $N \rightarrow \infty$ . Thus if  $Y_n$  not close to  $\theta$ ,  $F^*$  will have non-negligible mass at  $\infty$ .)

Use Taylor Exp of  $\alpha(Y_n)$  around  $\theta$ :

$$\alpha(Y_n) \approx \alpha(\theta) + \alpha'(\theta)(Y_n - \theta)$$

$$\Rightarrow \alpha(Y_n) - \alpha(\theta) \approx \alpha'(\theta)(Y_n - \theta)$$

$$\Rightarrow \frac{a_n}{\alpha'(\theta)} (\alpha(Y_n) - \alpha(\theta)) \approx a_n(Y_n - \theta)$$

Thus, CDF of LHS  $\approx$  CDF of RHS  $= F^*$

## Distribution of Function of Sample Mean

Consider  $\alpha(\bar{X}_n)$ .

By CLT,  $\frac{\sqrt{n}}{\sigma}(\bar{X}_n - \mu) \xrightarrow{d} \Phi$ .

Thus,  $\frac{\sqrt{n}}{\sigma}(\alpha(\bar{X}_n) - \alpha(\mu)) / \alpha'(\mu) \rightarrow \Phi$ .

(By Delta Method,  $Y_n = \bar{X}_n$ ,  $\theta = \mu$ ,  $a_n = \frac{\sqrt{n}}{\sigma}$ ,  $F^* = \Phi$ )