# Model Checking

Overview: to check the adequacy of the fit of the model to the data and the plausibility of the model for the purposes for which the model will be used.

We assume  $p(y|\theta)$  and  $p(\theta)$ , so it would be prudent to determine if these assumptions are reasonable.

- (Prior) sensitivity analysis
- Posterior predictive checks
  - Graphical checks
  - Posterior predictive pvalues

# Sensitivity analysis

- How much do different choices in model structure and priors affect the results?
  - test different models and priors
  - alternatively combine different models to one model
    - e.g. hierarchical model instead of separate and pooled
    - e.g. *t* distribution contains Gaussian as a special case
  - robust models are good for testing sensitivity to "outliers"
    - e.g. t instead of Gaussian
- Compare sensitivity of essential inference quantities
  - extreme quantiles are more sensitive than means and medians
  - extrapolation is more sensitive than interpolation

# Prior sensitivity analysis

Since a prior specifies our prior belief, we may want to check to determine whether our conclusions would change if we held different prior beliefs. Suppose a particular scientific question can be boiled down to

$$Y_i \stackrel{ind}{\sim} Ber(\theta)$$

and that there is wide disagreement about the value for  $\theta$  such that the following might reasonably characterize different individual beliefs before the experiment is run:

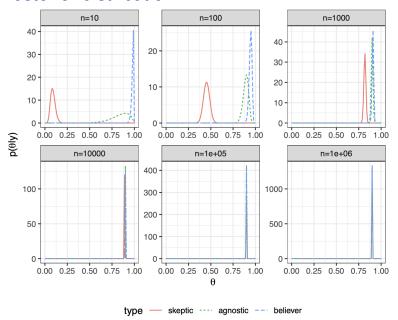
• Skeptic:  $\theta \sim \textit{Beta}(1, 100)$ 

Agnostic: θ ~ Beta(1, 1)

Believer: θ ~ Beta(100, 1)

An experiment is run and the posterior under these different priors are compared.

#### Posterior distribution



# Hierarchical variance prior (Review Gelman 2006)

Recall the normal hierarchical model

$$y_i \stackrel{ind}{\sim} N(\theta_i, s_i^2), \quad \theta_i \stackrel{ind}{\sim} N(\mu, \tau^2)$$

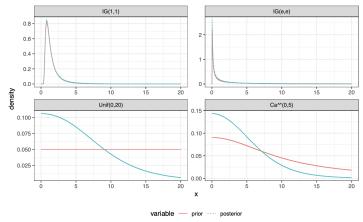
which results in the posterior distribution for  $\tau$  of

$$p( au|y) \propto p( au) V_{\mu}^{1/2} \prod_{j=1}^{J} (\sigma_{j}^{2} + au^{2})^{-1/2} \exp(-rac{(y_{j} - \hat{\mu})^{2}}{2(\sigma_{j}^{2} + au^{2})})$$

As an attempt to be non-informative, consider an  $IG(\epsilon, \epsilon)$  prior for  $\tau^2$ . As an alternative, consider  $\tau \sim Unif(0, C)$  or  $\tau \sim Ca^+(0, C)$  where C is problem specific, but is chosen to be relatively large for the particular problem.

# Posterior distribution - 8 schools example

#### Reproduction of Gelman 2006:



## Summary

For a default prior on a variance ( $\sigma^2$ ) or standard deviation ( $\sigma$ ), use

#### Easy

• Half-Cauchy on the standard deviation ( $\sigma \sim Ca^+(0, C)$ ).

#### 2. Complex

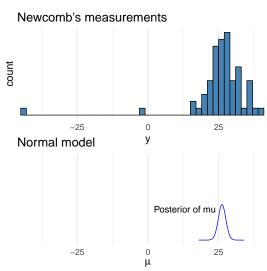
- Data-level variance
  - Use default prior  $(p(\sigma^2) \propto 1/\sigma^2)$
- · Hierarchical standard deviation
  - Use uniform prior (Unif(0, C)) if there are enough reps (5 or more) of that parameter.
  - Use half-Cauchy prior ( $Ca^+(0, C)$ ) otherwise.

#### When assigning the values for C

- For a uniform prior (Unif(0, C)) make sure C is large enough to capture any reasonable value for the standard deviation.
- For a half-Cauchy prior (Ca<sup>+</sup>(0, C)), a value of C that is too small will fail
  to capture the tail, implying that standard deviation needs to be larger
  whereas a value of C that is too large will put too much weight toward
  large values of the standard deviation and make the prior more
  informative.

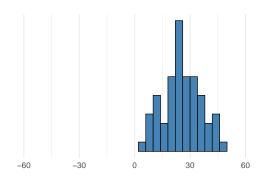
### Simon Newcomb's light of speed experiment in 1882

Newcomb measured (n = 66) the time required for light to travel from his laboratory on the Potomac River to a mirror at the base of the Washington Monument and back, a total distance of 7422 meters.



# Posterior predictive checking – example

- Newcomb's speed of light measurements
  - model  $y \sim N(\mu, \sigma)$  with prior  $(\mu, \log \sigma) \propto 1$
- Posterior predictive replicate y<sup>rep</sup>
  - draw  $\mu^{(s)}, \sigma^{(s)}$  from the posterior  $p(\mu, \sigma | y)$
  - draw  $y^{\text{rep }(s)}$  from  $N(\mu^{(s)}, \sigma^{(s)})$
  - repeat n times to get y<sup>rep</sup> with n replicates

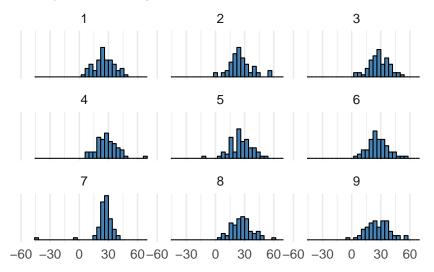


# Replicates vs. future observation

• Predictive  $\tilde{y}$  is the next not yet observed possible observation.  $y^{\text{rep}}$  refers to replicating the whole experiment (potentially with same values of x) and obtaining as many replicated observations as in the original data.

# Posterior predictive checking – example

- Generate several replicated datasets y<sup>rep</sup>
- Compare to the original dataset

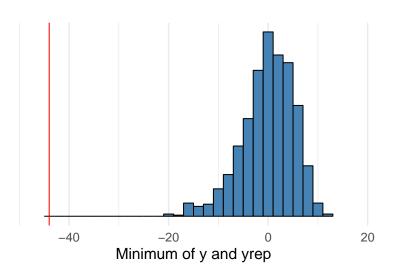


# Posterior predictive checking with test statistic

- Replicated data sets y<sup>rep</sup>
- Test quantity (or discrepancy measure)  $T(y, \theta)$ 
  - summary quantity for the observed data  $T(y, \theta)$
  - summary quantity for a replicated data  $T(y^{\text{rep}}, \theta)$
  - can be easier to compare summary quantities than data sets

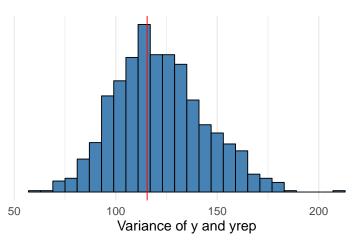
## Posterior predictive checking – example

- Compute test statistic for data  $T(y, \theta) = \min(y)$
- Compute test statistic  $min(y^{rep})$  for many replicated datasets



# Posterior predictive checking – example

- Good test statistic is ancillary (or almost)
  - ancillary if it depends only on observed data and if its distribution is independent of the parameters of the model
- Bad test statistic is highly dependent of the parameters
  - e.g. variance for normal model



# Posterior predictive checking

Posterior predictive p-value

$$\rho = \Pr(T(y^{\text{rep}}, \theta) \ge T(y, \theta)|y) 
= \int \int I_{T(y^{\text{rep}}, \theta) \ge T(y, \theta)} p(y^{\text{rep}}|\theta) p(\theta|y) dy^{\text{rep}} d\theta$$

where I is an indicator function

 having (y<sup>rep (s)</sup>, θ<sup>(s)</sup>) from the posterior predictive distribution, easy to compute

$$T(y^{\text{rep}(s)}, \theta^{(s)}) \ge T(y, \theta^{(s)}), \quad s = 1, \dots, S$$

- Posterior predictive p-value (ppp-value) estimated whether difference between the model and data could arise by chance
- Not commonly used, since the distribution of test statistic has more information

# Marginal predictive checking

- Consider marginal predictive distributions  $p(\tilde{y}_i|y)$  and each observation separately
  - · marginal posterior p-values

$$p_i = \mathsf{Pr}(T(y_i^{\mathrm{rep}}) \leq T(y_i)|y)$$
 if  $T(y_i) = y_i$   $p_i = \mathsf{Pr}(y_i^{\mathrm{rep}} \leq y_i|y)$ 

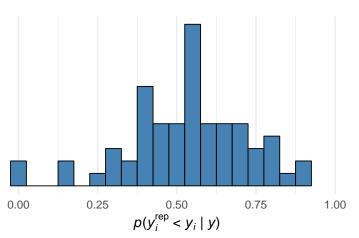
• if  $Pr(\tilde{y}_i|y)$  well calibrated, distribution of  $p_i$  would be uniform between 0 and 1

# Marginal predictive checking - Example

Marginal tail area or Probability integral transform (PIT)

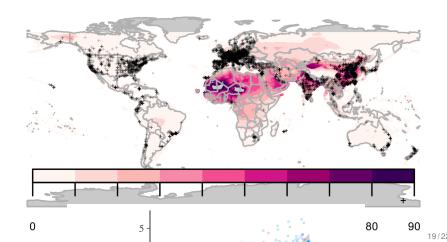
$$p_i = p(y_i^{\text{rep}} \leq y_i|y)$$

• if  $p(\tilde{y}_i|y)$  is well calibrated, distribution of  $p_i$ 's would be uniform between 0 and 1

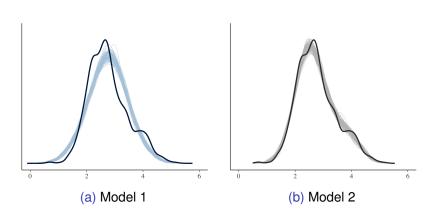


- Example from Jonah Gabry, Daniel Simpson, Aki Vehtari, Michael Betancourt, and Andrew Gelman (2019).
   Visualization in Bayesian workflow. https://doi.org/10.1111/rssa.12378
- Estimation of human exposure to air pollution from particulate matter measuring less than 2.5 microns in diameter (PM<sub>2.5</sub>)
  - Exposure to PM<sub>2.5</sub> is linked to a number of poor health outcomes and a recent report estimated that PM<sub>2.5</sub> is responsible for three million deaths worldwide each year (Shaddick et al., 2017)
  - In order to estimate the public health effect of ambient PM<sub>2.5</sub>, we need a good estimate of the PM<sub>2.5</sub> concentration at the same spatial resolution as our population estimates.

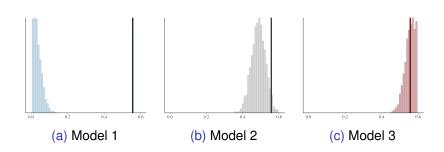
- Direct measurements of PM 2.5 from ground monitors at 2980 locations
- · High-resolution satellite data of aerosol optical depth



Posterior predictive checking – marginal predictive distributions



Posterior predictive checking – test statistic (skewness)



Posterior predictive checking – test statistic (median for groups)

