



Probability and Random Process

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- 4. Random Process-II
 - Introduction to Markov Processes
 - Classifications of States and MCs
 - Computing State Probabilities
 - Continuous-time MC
 - Ergodicity Theorems
 - Series Expansions



Computing State Probabilities

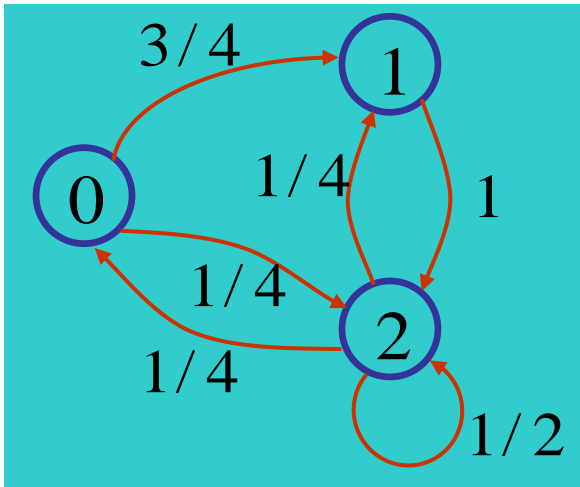
Computing State Probabilities

- When it exists, the steady state distribution (i.e. probability vector) can be found:
 1. Repeated propagation of the state vector using a computer program until it is sufficiently converged
 2. As the eigenvector of \mathbf{P}^T corresponding to the unity eigenvalue, normalized by the sum of its elements
 3. Direct solving of the equations:

$$\begin{bmatrix} \pi_0 & \pi_2 & \cdots & \pi_{N-1} \end{bmatrix} = \begin{bmatrix} \pi_0 & \pi_2 & \cdots & \pi_{N-1} \end{bmatrix} \mathbf{\Pi}$$

$$\text{and} \quad \sum_{i=0}^{N-1} \pi_i = \mathbf{1}$$

Example



$$\mathbf{\Pi} = \begin{bmatrix} P_{00} & P_{01} & P_{02} \\ P_{10} & P_{11} & P_{12} \\ P_{20} & P_{21} & P_{22} \end{bmatrix} = \begin{bmatrix} 0 & 3/4 & 1/4 \\ 0 & 0 & 1 \\ 1/4 & 1/4 & 1/2 \end{bmatrix}$$

- All pairs of states communicate, therefore this MC is **irreducible**.
- It is **aperiodic** (because $p_{22} > 0$).
- Because the state space is finite, the states must be **recurrent non-null**.

⇒ Therefore, the steady state probability vector exists.

Example, Cont'd

$$\begin{bmatrix} \pi_0 & \pi_1 & \pi_2 \end{bmatrix} = \begin{bmatrix} \pi_0 & \pi_1 & \pi_2 \end{bmatrix} \begin{bmatrix} 0 & 3/4 & 1/4 \\ 0 & 0 & 1 \\ 1/4 & 1/4 & 1/2 \end{bmatrix}$$

$$\pi_0 = \frac{\pi_2}{4} \quad (1)$$

$$\pi_1 = \frac{3}{4} \pi_0 + \frac{\pi_2}{4} \quad (2)$$

$$\pi_2 = \frac{\pi_0}{4} + \pi_1 + \frac{\pi_2}{2} \quad (3)$$

$$\pi_0 + \pi_1 + \pi_2 = 1 \quad (4)$$

Substitute (1) into (2) and (4):

$$\pi_1 = \frac{3}{4} \pi_0 + \pi_0 \quad (5)$$

$$\pi_0 + \pi_1 + 4\pi_0 = 1 \quad (6)$$

Substitute (5) into (6):

$$\frac{7}{4} \pi_0 + 5\pi_0 = 1 \Rightarrow \pi_0 = \frac{4}{27}$$

Example, Cont'd

Now, substitute back into (5): $\pi_1 = \frac{7}{27}$

Now use (4) to get: $\pi_2 = 1 - \frac{(4+7)}{27} = \frac{16}{27}$

$$\therefore \boldsymbol{\pi}^{(\infty)} = \left[\frac{4}{27}, \frac{7}{27}, \frac{16}{27} \right]$$

- Suppose we want $\rho^{(n)}$ for some very large, but finite, n ?
- We can calculate $\rho^{(n)}$ directly in closed form using Z-transforms.
- Take Z-transform of both sides of $\boldsymbol{\pi}^{(n+1)} = \boldsymbol{\pi}^{(n)}\mathbf{\Pi}$

$$Z\{\boldsymbol{\pi}^{(n+1)}\} = Z\{\boldsymbol{\pi}^{(n)}\mathbf{\Pi}\}$$

Let $G(z) = Z\{\boldsymbol{\pi}^{(n)}\}$

What is $Z\{\rho^{(n+1)}\}$?

Z-Transform Details

$$Z[\pi^{(n+1)}] = \sum_{n=0}^{\infty} \pi^{(n+1)} z^{-n} = z \sum_{n=0}^{\infty} \pi^{(n+1)} z^{-(n+1)}$$

Let $k = n + 1$,
$$= z \sum_{k=1}^{\infty} \pi^{(k)} z^{-(k)} = z[G(z) - \pi^{(0)}]$$

So
$$Z[\pi^{(n+1)}] = Z[\pi^{(n)} \Pi]$$

$$z[G(z) - \pi^{(0)}] = G(z) \Pi$$

$$G(z)[z\mathbf{I} - \Pi] = z\pi^{(0)}$$

$$G(z) = z\pi^{(0)}[z\mathbf{I} - \Pi]^{-1}$$

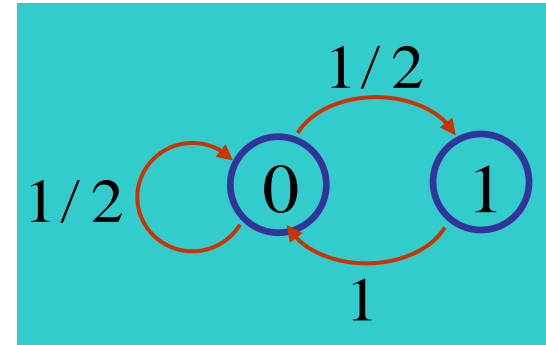
Then **inverse Z-transform** is performed on $G(z)$ for $\pi^{(n)}$

Example

$$\mathbf{\Pi} = \begin{bmatrix} 1/2 & 1/2 \\ 1 & 0 \end{bmatrix}$$

Suppose $\boldsymbol{\pi}^{(0)} = [1/4 \quad 3/4]$

We want an expression for $\rho^{(n)}$



$$G(z) = z \begin{bmatrix} 1/4 & 3/4 \end{bmatrix} [z\mathbf{I} - \mathbf{\Pi}]^{-1}$$

$$[z\mathbf{I} - \mathbf{\Pi}]^{-1} = \left[\begin{bmatrix} z & 0 \\ 0 & z \end{bmatrix} - \begin{bmatrix} 1/2 & 1/2 \\ 1 & 0 \end{bmatrix} \right]^{-1} = \begin{bmatrix} z - 1/2 & -1/2 \\ -1 & z \end{bmatrix}^{-1}$$

$$= \frac{\begin{bmatrix} z & 1/2 \\ 1 & z - 1/2 \end{bmatrix}}{z(z - 1/2) - 1/2} = \frac{\begin{bmatrix} z & 1/2 \\ 1 & z - 1/2 \end{bmatrix}}{(z - 1)(z + 1/2)}$$

Example, Cont'd

Partial Fraction Expansion on each matrix term yields:

$$[z\mathbf{I} - \mathbf{\Pi}]^{-1} = \begin{bmatrix} \left(\frac{2/3}{z-1} + \frac{1/3}{z+1/2} \right) & \left(\frac{1/3}{z-1} - \frac{1/3}{z+1/2} \right) \\ \left(\frac{2/3}{z-1} - \frac{2/3}{z+1/2} \right) & \left(\frac{1/3}{z-1} + \frac{2/3}{z+1/2} \right) \end{bmatrix}$$

$$\begin{aligned} G(z) &= z \begin{bmatrix} 1/4 & 3/4 \end{bmatrix} [z\mathbf{I} - \mathbf{\Pi}]^{-1} \\ &= \begin{bmatrix} \left(\frac{2z/3}{z-1} - \frac{5z/12}{z+1/2} \right) & \left(\frac{z/3}{z-1} + \frac{5z/2}{z+1/2} \right) \end{bmatrix} \\ \boldsymbol{\pi}^{(n)} &= \begin{bmatrix} \frac{2}{3} - \left(-\frac{1}{2} \right)^n \frac{5}{12}, & \frac{1}{3} + \left(-\frac{1}{2} \right)^n \frac{5}{12} \end{bmatrix} \end{aligned}$$

Partial Fraction Details

$$\frac{H(z)}{(z-1)(z+1/2)} = \frac{A}{z-1} - \frac{B}{z+1/2} = \frac{zA + A/2 + Bz - B}{(z-1)(z+1/2)}$$

$$\text{For } H(z) = z \Rightarrow \left. \begin{array}{l} A + B = 1 \\ A/2 - B = 0 \end{array} \right\} \quad A = \frac{2}{3}, \quad B = \frac{1}{3}$$

$$\text{For } H(z) = 1/2 \Rightarrow \left. \begin{array}{l} A + B = 0 \\ A/2 - B = 1/2 \end{array} \right\} \quad A = \frac{1}{3}, \quad B = -\frac{1}{3}$$

$$\text{For } H(z) = 1 \Rightarrow \left. \begin{array}{l} A + B = 0 \\ A/2 - B = 1 \end{array} \right\} \quad A = \frac{2}{3}, \quad B = -\frac{2}{3}$$

$$\text{For } H(z) = z - 1/2 \Rightarrow \left. \begin{array}{l} A + B = 1 \\ A/2 - B = -1/2 \end{array} \right\} \quad A = \frac{1}{3}, \quad B = \frac{2}{3}$$

Example, Cont'd

$$\boldsymbol{\pi}^{(n)} = \left[\frac{2}{3} - \left(-\frac{1}{2}\right)^n \frac{5}{12}, \frac{1}{3} + \left(-\frac{1}{2}\right)^n \frac{5}{12} \right]$$

Check: $\boldsymbol{\pi}^{(0)} = \left[\frac{2}{3} - \frac{5}{12}, \frac{1}{3} + \frac{5}{12} \right] = \left[\frac{3}{12}, \frac{9}{12} \right]$ ✓

$$\boldsymbol{\pi}^{(1)} = \left[\frac{2}{3} + \frac{5}{24}, \frac{1}{3} - \frac{5}{24} \right] = \left[\frac{7}{8}, \frac{1}{8} \right]$$

$$= \boldsymbol{\pi}^{(0)} \mathbf{\Pi} = \left[\frac{1}{4}, \frac{3}{4} \right] \begin{bmatrix} 1/2 & 1/2 \\ 1 & 0 \end{bmatrix} = \left[\frac{7}{8}, \frac{1}{8} \right]$$
 ✓

- Compute steady state distribution
 - Check if steady state probability vector exists
- Calculate non-steady state distribution
 - By using the Z-transform

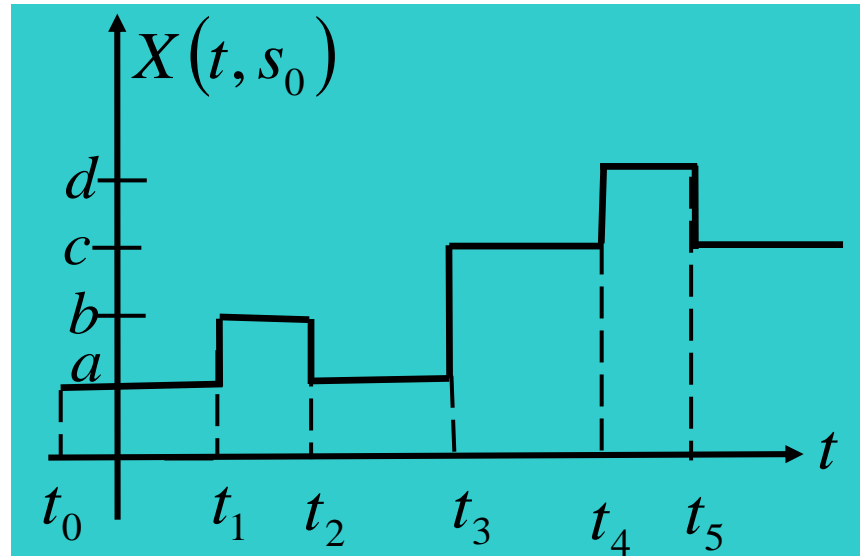


Continuous-time MC

Continuous-Time Discrete-State

- In a **continuous-time discrete-state** MC, or CTMC, a transition can occur *at any instant of time*.
- Instead of a probability transition matrix, as in discrete-time MCs, there is a **transition rate matrix**.
- A CTMC can be specified by its underlying **point process** (the list of transition times) and the **embedded MC** (the list of values corresponding to the transitions).

Example



For this realization, the point process is:

$$\{\dots, t_0, t_1, t_2, t_3, t_4, \dots\}$$

And the embedded MC is:

$$\{\dots, a, b, a, c, d, \dots\}$$

A Stepping Stone to Rates: Transition Probabilities

- A CTMC, $X(t)$, may also be described by its initial state probabilities

$$p_i(t_0) = P(X(t_0) = a_i)$$

Where a_i is the value associated with state i , and its **transition probabilities**

$$\pi_{ij}(t_1, t_2) = P(X(t_2) = a_j \mid X(t_1) = a_i)$$

- As in the DTMC case, these quantities satisfy:

$$\sum_j \pi_{ij}(t_1, t_2) = 1, \quad \sum_i p_i(t_1) \pi_{ij}(t_1, t_2) = p_j(t_2)$$

and the Chapman-Kolmogorov equations

$$\pi_{ij}(t_1, t_3) = \sum_r \pi_{ir}(t_1, t_2) \pi_{rj}(t_2, t_3), \quad t_1 < t_2 < t_3$$

Homogenous CTMCs

The CTMC is homogenous if its transition probabilities depend only on the difference $t = t_2 - t_1$:

$$\pi_{ij}(\tau) = P(X(t + \tau) = a_j \mid X(t) = a_i)$$

The Chapman-Kolmogorov equations become

$$\pi_{ij}(\tau + \alpha) = \sum_r \pi_{ir}(\tau) \pi_{rj}(\alpha)$$

or in matrix form:

$$\mathbf{\Pi}(\tau + \alpha) = \mathbf{\Pi}(\tau) \mathbf{\Pi}(\alpha)$$

And the state probability vector, $\mathbf{p}(t)$, satisfies:

$$\mathbf{p}(\tau + t) = \mathbf{p}(t) \mathbf{\Pi}(\tau)$$

- The derivative-from-the-right of the transition probability matrix

$$\mathbf{\Lambda} = \left. \frac{d}{d\tau} \mathbf{\Pi}(\tau) \right|_{\tau=0^+} = \begin{bmatrix} \lambda_{11} & \lambda_{12} & \cdots & \lambda_{N1} \\ \lambda_{21} & \ddots & & \vdots \\ \vdots & & & \\ \lambda_{N1} & \cdots & & \lambda_{NN} \end{bmatrix}$$

is the **transition rate matrix**.

- λ_{ij} indicates the rate, in terms of probability mass per unit time, that probability mass moves out of state i and into state j (λ_{ij} can be < 0).

Conservation of Mass

- If probability mass is not flowing out of a state, then it must be staying in
- Recall $\sum_j \pi_{ij}(\tau) = 1$
- Differentiation of this yields $\sum_j \lambda_{ij} = 0$
- $-\lambda_{ii}$ is the rate that mass moves out of state i . It does not include mass coming in from other states.

$$-\lambda_{ii} = \sum_{j \neq i} \lambda_{ij}$$

- All Markov processes share the interesting property that the time it takes for a change of state is an **exponentially distributed random variable**
- All CTMCs stay in given state for an exponentially distributed period of time, with parameter $-\lambda_{ii}$

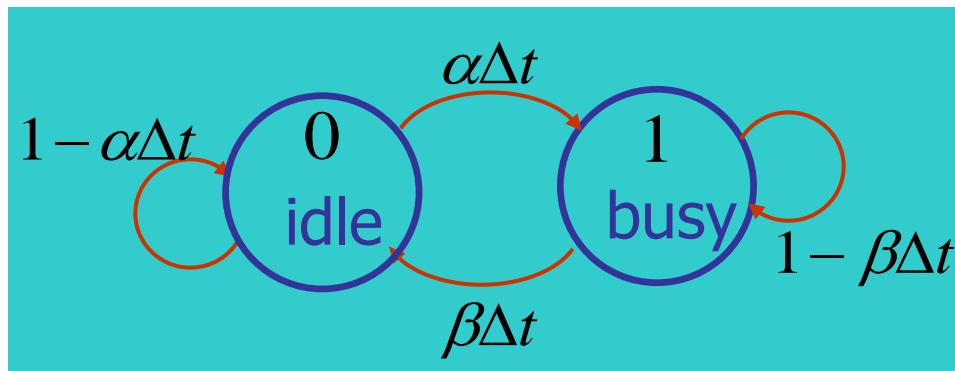
Global Balance Equations

- In steady state, rate of flow into a state must equal the rate of flow out of a state.
- Consider state j over a very small interval Δt :

Into j :

$$\left. \begin{array}{l} \text{Into } j: \sum_{\substack{i \\ j \neq i}} \Delta t \lambda_{ij} \rho_i \\ \text{Out of } j: \sum_{\substack{i \\ j \neq i}} \Delta t \lambda_{ji} \rho_j \end{array} \right\} \sum_{\substack{i \\ j \neq i}} \Delta t \lambda_{ij} \rho_i = \rho_j \sum_{\substack{i \\ j \neq i}} \Delta t \lambda_{ji}$$

Simple Queuing System [Ex. 8.15 from Kleinrock]



Global Balance Equations

GBE, repeated:

$$\begin{bmatrix} \lambda_{00} & \lambda_{01} \\ \lambda_{10} & \lambda_{11} \end{bmatrix} = \begin{bmatrix} -\alpha & \alpha \\ \beta & -\beta \end{bmatrix}$$

$$\sum_{\substack{i \\ j \neq i}} \Delta t \lambda_{ij} \rho_i = \rho_j \sum_{\substack{i \\ j \neq i}} \Delta t \lambda_{ji}$$

$$j=0: \quad \lambda_{10} \rho_1 = \rho_0 \lambda_{01} \quad \Rightarrow \quad \beta \rho_1 = \rho_0 \alpha \quad (1)$$

$$j=1: \quad \lambda_{01} \rho_0 = \rho_1 \lambda_{10} \quad \Rightarrow \quad \text{Same (no help)}$$

$$\text{Can use:} \quad \rho_0 + \rho_1 = 1 \quad (2)$$

Eliminate r_1 by substituting $r_1 = 1 - r_0$

$$\beta(1 - \rho_0) = \rho_0 \alpha \quad \Rightarrow \quad \beta = (\alpha + \beta) \rho_0 \quad \Rightarrow \quad \begin{cases} \rho_0 = \frac{\beta}{\alpha + \beta} \\ \rho_1 = \frac{\alpha}{\alpha + \beta} \end{cases}$$

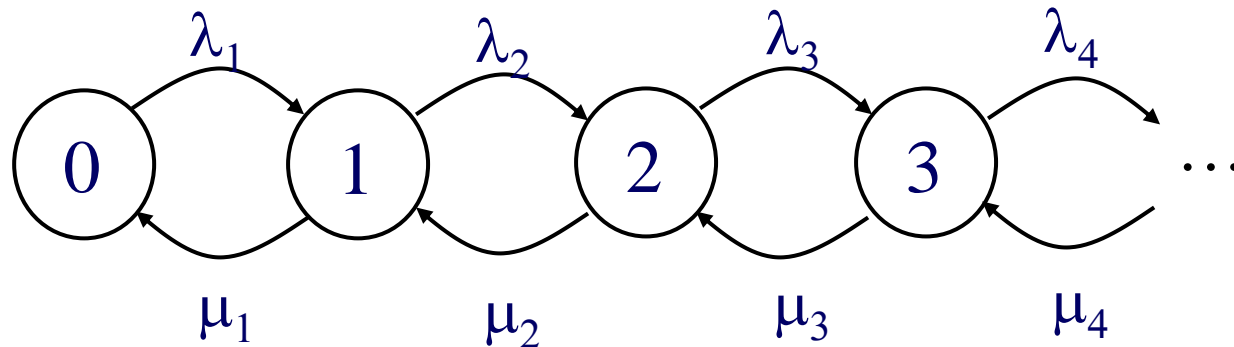


Continuous-Time Discrete-State Examples

- Birth and Death Process
- M/M/1 Single-Server Queue
- M/M/c Multi-Server Queue

Birth and Death Process

- The birth–death process can go only from state n to state $n-1$ or $n+1$ in one transition
 - The state transitions are of only two types: "birth", which increases the state variable by one and "death", which decreases the state by one
- Graphical interpretation: any process with the following state diagram



GBE for Birth-Death

$$p_j = r_j p_{j-1} \quad \text{and} \quad p_j = \prod_{k=1}^j r_k p_0, \quad \text{with} \quad r_k = \frac{\lambda_k}{\mu_k}$$

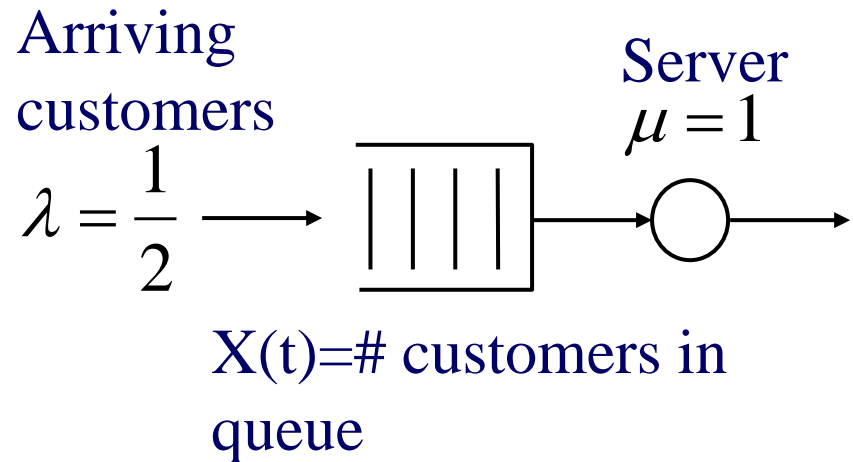
Let $R_j = \prod_{k=1}^j r_k$ then the condition $\sum_{j=0}^{\infty} p_j = 1$ yields

$$p_0 = \frac{1}{\sum_{j=0}^{\infty} R_j} \quad \text{and} \quad p_j = \frac{R_j}{\sum_{j=0}^{\infty} R_j}$$

Queue Notation

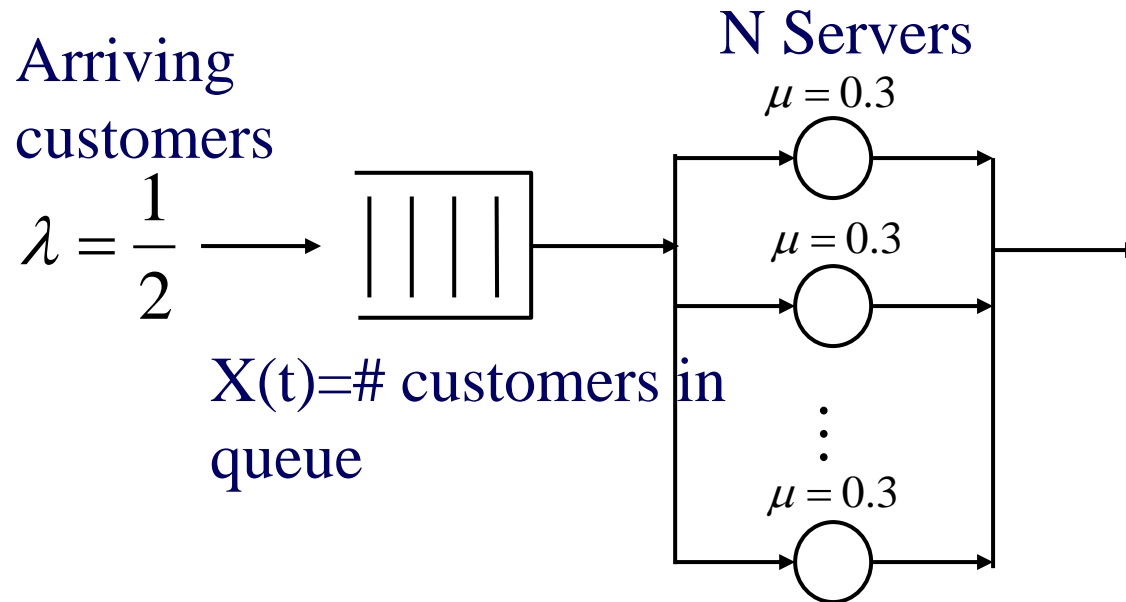
- a/b/c
 - a indicates customer arrival distribution.
M=Poisson
 - b indicates distribution of service times.
M=Exponential
 - c=Number of servers

M/M/1



Multiple Servers

M/M/N

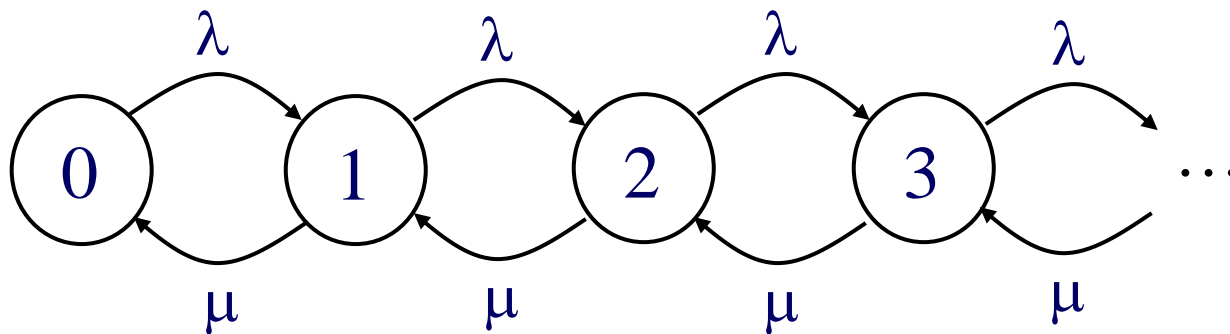


- Arrivals occur according to a **Poisson** process with average rate λ .
- Inter-arrivals are iid **exponential** with mean $1/\lambda$.
- Customers are served, one at a time, with iid **exponentially** distributed service times with parameter μ

$$f_{\tau_s}(t) = \begin{cases} \mu e^{-\mu t} & t \geq 0 \\ 0 & t < 0 \end{cases} \quad E\{\tau_s\} = \frac{1}{\mu}$$

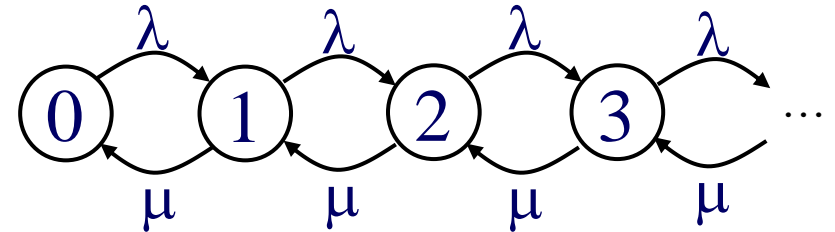
M/M/1 State Diagram

- State space: the number of customers in the system
 - Including any currently in service



M/M/1 is a special case of birth-death process, with the identical arriving and serving rates

Global Balance Equations for M/M/1



State 0: $p_0\lambda = p_1\mu$

Can rewrite as: $p_0\lambda - p_1\mu = 0$

State 1: $p_1\lambda + p_1\mu = p_0\lambda + p_2\mu$

Can rewrite as: $p_1\lambda - p_2\mu = p_0\lambda - p_1\mu = 0$

State j: $p_j\lambda + p_j\mu = p_{j-1}\lambda + p_{j+1}\mu$

Can rewrite as: $p_j\lambda - p_{j+1}\mu = p_{j-1}\lambda - p_j\mu = 0$

GBE Concluded

$$p_{j-1}\lambda - p_j\mu = 0 \quad \text{implies} \quad p_j = \frac{\lambda}{\mu} p_{j-1}$$

$$\text{or} \quad p_j = \left(\frac{\lambda}{\mu}\right)^j p_0$$

$$\text{Thus, requiring} \quad \sum_{j=0}^{\infty} p_j = \sum_{j=0}^{\infty} \left(\frac{\lambda}{\mu}\right)^j p_0 = 1$$

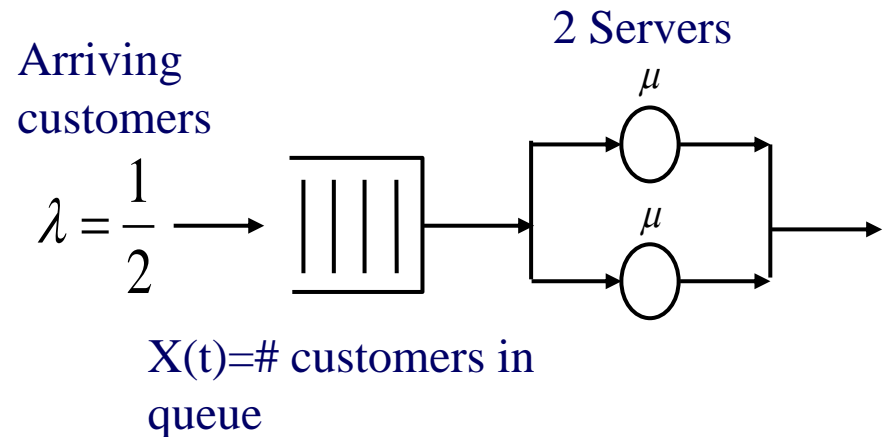
$$\text{yields} \quad p_0 = 1 - \frac{\lambda}{\mu} \quad \text{only if} \quad \frac{\lambda}{\mu} < 1$$

Must be able to
serve them
at a higher rate
than they arrive

M/M/c: Multi-server Queue

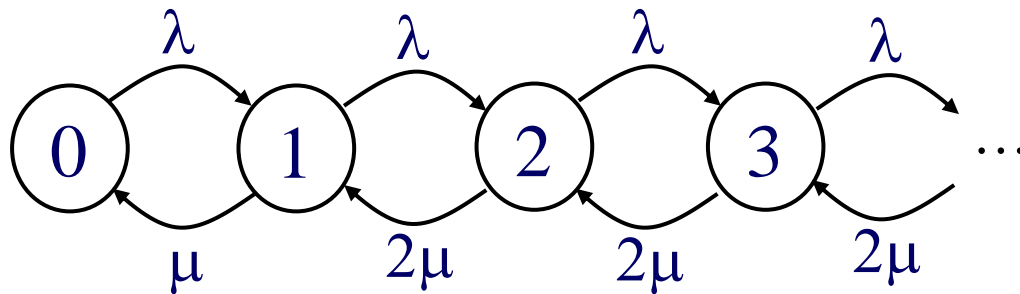
- Customer is served by first available server
- Let τ_i = time until server i is available
- Then time until next server is available is $X = \min(\tau_1, \tau_2, \dots, \tau_n)$

M/M/2



M/M/c State Diagram

- Since each of the servers works at rate μ , the total departure rate will be 2μ
- The state diagram becomes



- Continuous-time Markov Chains (CTMCs) can change state at any time.
 - The time between entering state i and leaving it is exponentially distributed with parameter $- \lambda_{ii}$
- Global Balance Equations equalize mass flow into and out of a state, assuming steady state conditions.
- Birth-death processes
 - M/M/c queues



Thank You!