

### **Probability and Random Process**

#### Aimin Tang

The University of Michigan- Shanghai Jiao Tong University Joint Institute Shanghai Jiao Tong University

Dec. 3 2020



#### • 4. Random Process-II

- Introduction to Markov Processes
- Classifications of States and MCs
- Computing State Probabilities
- Continuous-time MC
- Ergodicity Theorems
- Series Expansions



## **Ergodicity Theorems**

- A process is ergodic in some statistic if that statistic can be estimated using a time average of an observation (i.e. using one outcome) of the process.
- If we weight the outcomes of a random process with their probabilities, we have the true mean of the RP

$$m_X(t) = \sum_i X(t, s_i) P(s_i)$$

- The mean is also called the ensemble average
  - The ensemble is another word for the collection of all outcomes



Next suppose we use only one outcome. Then we might compute its time average:

$$\langle X(t)\rangle_T = \frac{1}{2T} \int_{-T}^T X(t, s_i) dt$$

When does a limit of  $\langle X(t) \rangle_T$  equal  $m_X(t)$ ?



Let 
$$X(t) = A + \sin(\omega t + \theta)$$
,  $\theta \sim U[0,2\pi]$   $A \sim N(3,\sigma^2)$ 

A &  $\theta$  independent

$$E[X(t)] = 3$$

But

$$\lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} X(t, s_i) dt = A(s_i) \neq 3$$

With probability 1



We can derive the condition for

$$\lim_{T\to\infty} \langle X(t) \rangle_T = m_X$$

where l.i.m. is the limit in the mean-square convergence

$$\lim_{T \to \infty} E\left\{ \left( \left\langle X(t) \right\rangle_T - m_X \right)^2 \right\} = 0$$

• The RP X(t) satisfying this is "mean ergodic"

 By expanding the expectation on the previous slide, a process X(t) is mean-ergodic if and only if its autocovariance C<sub>X</sub>(t<sub>1</sub>, t<sub>2</sub>) satisfies:

$$\lim_{T\to\infty} \frac{1}{4T^2} \int_{-T}^{T} \int_{-T}^{T} C_X(t_1, t_2) dt_1 dt_2 \xrightarrow{T\to\infty} 0$$

If X(t) is WSS, this condition becomes

$$\frac{1}{2T} \int_{-2T}^{2T} C_X(\tau) \left( 1 - \frac{|\tau|}{2T} \right) d\tau \xrightarrow{T \to \infty} 0$$

Integrate over the time difference



Let  $X(t) = \eta + v(t)$  where  $C_v(\tau) = q\delta(\tau)$  and  $\eta$  is non-random.

$$C_{X}(\tau) = C_{v}(\tau)$$

$$\frac{1}{2T} \int_{-2T}^{2T} q \, \delta(\tau) \left(1 - \frac{|\tau|}{2T}\right) d\tau = \frac{q}{2T} \xrightarrow{T \to \infty} 0$$

Observations: Note that  $1 - \frac{|\tau|}{2T} \le 1$  for  $|\tau| \le 2T$ ,

a sufficient condition for mean-ergodicity is

$$\int_{-\infty}^{\infty} |C_X(\tau)| \, d\tau < \infty$$



We want to estimate  $R_X(\tau) = E\{X(t+\tau)X(t)\}$  by averaging  $Z(t) = X(t+\tau)X(t)$  over a long period of time

Apply previous result to Z(t):

$$\lim_{T\to\infty} \frac{1}{2T} \int_{-T}^{T} X(t+\tau)X(t)dt = R_X(\tau)$$

iff 
$$\frac{1}{4T^2} \int_{-T-T}^{T} C_Z(t_1, t_2) dt_1 dt_2 \xrightarrow{T \to \infty} 0$$

$$C_{Z}\left(t_{1},t_{2}\right)=E\left[X\left(t_{1}+\tau\right)X\left(t_{1}\right)X\left(t_{2}+\tau\right)X\left(t_{2}\right)\right]-R_{X}^{2}\left(\tau\right)$$



#### Suppose X(t) is Gaussian, zero-mean with

$$R_{X}(\tau) = e^{-\alpha|\tau|} \quad \alpha > 0$$

We observe that 
$$Z(t)$$
 is WSS. Let  $\gamma = t_1 - t_2$ , then  $C_Z(\gamma) = R_X^2(\gamma) + R_X(\gamma + \tau)R_X(\gamma - \tau)$ 
$$= e^{-\alpha 2|\gamma|} + e^{-\alpha [|\gamma + \tau| + |\gamma - \tau|]}$$

Check textbook: Papoulis "Probability, Random Variables and Stochastic Processes", Fourth edition, Page. 532



Use 
$$|\gamma| - |\tau| \le |\gamma + \tau|$$
  
and  $|\gamma| - |\tau| \le |\gamma - \tau|$   $\Rightarrow 2|\gamma| - 2|\tau| \le |\gamma + \tau| + |\gamma - \tau|$ 

$$C_{Z}(\gamma) \le e^{-\alpha 2|\gamma|} + e^{-\alpha \left[2|\gamma| - 2|\tau|\right]}$$
$$= e^{-\alpha 2|\gamma|} \left(1 + e^{\alpha 2|\tau|}\right)$$

#### Use the sufficient condition

$$\begin{split} &\int\limits_{-\infty}^{\infty} |C_{Z}(\gamma)| \, d\gamma \leq \left(1 + e^{\alpha 2|\tau|}\right) \int\limits_{0}^{\infty} 2e^{-\alpha 2|\gamma|} \, d\gamma \\ &= \left(1 + e^{\alpha 2|\tau|}\right) \frac{e^{-\alpha 2\gamma}}{-\alpha} \bigg|_{0}^{\infty} = \left(1 + e^{\alpha 2|\tau|}\right) \frac{1}{\alpha} < \infty \end{split}$$



- Determined condition for estimating an ensemble average using a time average
- Determined condition for estimating autocorrelation by applying the formula for the mean but replacing the autocovariance



## **Series Expansions**

In general, a continuous-time RP is an uncountable number of RVs

If the RP is limited in either time or frequency, or is mean-square periodic, then the RP can be specified by a countable number of RVs:

$$X(t) = \sum_{n=-\infty}^{\infty} a_n \phi_n(t)$$
 expansions

where the basis functions  $\phi_n(t)$  are not random and the coefficients  $a_n$ 's are a random sequence Often, the coefficients have properties that aid analysis

Ve501 2020-2021 Fall



- Sampling Theorem for band-limited RPs
- Fourier Series for mean square periodic RPs
- KL (Karhunen-Loéve) expansion for time-limited RPs

Ve501 2020-2021 Fall



 X(t) is bandlimited (BL) if it has finite power and if

$$S_X(\omega) = 0$$
 for  $|\omega| > B$ 

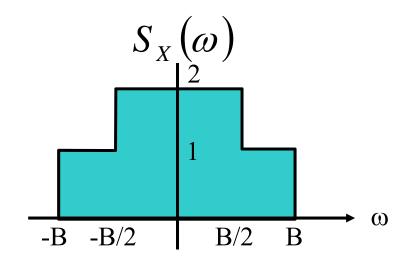
• Given this, the following expansion holds:

$$X(t) = \sum_{n=-\infty}^{\infty} X(nT) \frac{\sin[B(t-nT)]}{B(t-nT)} \quad \text{(m.s.)} \quad T = \frac{\pi}{B}$$

The samples are correlated in general



#### • Let X(t) have the PSD shown:



$$R_X(\tau) = \frac{B}{\pi} \operatorname{sinc}\left(\frac{B}{\pi}\tau\right) + \frac{B}{2\pi} \operatorname{sinc}\left(\frac{B}{2\pi}\tau\right)$$



The correlation between samples is

$$E\{X(nT)X(mT)\} = R_X([n-m]T)$$

$$= \begin{cases} \frac{3B}{2\pi} & n=m\\ \frac{B}{2\pi}\operatorname{sinc}([n-m]\frac{1}{2}) & n \neq m \end{cases}$$

Some pairs of samples are not orthogonal



• If X=Y (m.s.), then

$$E[(X-Y)^2]=0$$



 $\bullet$  X(t) is m.s. periodic with period T if

$$E\left[X(t+mT)-X(t)\right]^{2}=0$$

for all t and for all integers m

ullet Equivalently, X(t) is m.s. periodic iff

$$R_X(t_1 + mT, t_2 + nT) = R_X(t_1, t_2)$$

for every integer m and n



• Given that X(t) is m.s. periodic with period T, then the following m.s. expansion holds

$$X(t) = \sum_{n=-\infty}^{\infty} c_n e^{jn\omega_o t} \qquad c_n = \frac{1}{T} \int_0^T X(t) e^{-jn\omega_o t} dt$$
 for 
$$\omega_o = \frac{2\pi}{T}$$



## Correlation of Fourier Series Coefficients

• The correlation of the  $c_n$  is

$$E(c_n c_m^*) = \gamma_n \delta(\mathbf{n} - \mathbf{m})$$

Orthogonal coefficients!

• What is  $\gamma_n$ ?



#### **Mean Square of Coefficient**

$$E\{|c_{n}|^{2}\} = E\{\frac{1}{T^{2}}\int_{0}^{T}\int_{0}^{T}X(t)X(s)e^{-jn\omega_{0}t}e^{jn\omega_{0}s}dtds\}$$
$$= \frac{1}{T^{2}}\int_{0}^{T}\int_{0}^{T}R_{X}(t-s)e^{-jn\omega_{0}(t-s)}dtds$$

Both  $R_X(t-s)$  and  $e^{-jn\omega_0(t-s)}$  are periodic in t with period T, so the inner integral is invariant to s and equals  $T\gamma_n$ , where

$$\gamma_n = \frac{1}{T} \int_0^T R(\tau) e^{-jn\omega_0 \tau} d\tau$$

$$\therefore E[|c_n|^2] = \gamma_n$$



# Fourier Series Expansion of the Autocorrelation Function

 The previous slide implies that, for WSS Mean Square Periodic RPs,

$$R_X(\tau) = \sum_{n=-\infty}^{\infty} \gamma_n e^{jn\omega_o \tau}$$

Ve501 2020-2021 Fall



- Applies to time-limited RPs with finite autocorrelation
- Definition

$$X(t) = \sum_{n=-\infty}^{+\infty} c_n \phi_n(t) \qquad 0 < t < T$$

where  $\{\phi_1(t), \phi_2(t), ...\}$  is a complete ortho-normal (CON) set of deterministic functions

$$c_n = \int_0^T X(t)\phi_n^*(t)dt$$
, and  $E[c_n c_m^*] = \lambda_n \delta[n-m]$ 

Projection on the basis function



• "Complete" means that there are enough  $\phi(t)$  functions such that if g(t) is any deterministic, finite energy function over [0,T], then there exists a set of deterministic coefficients  $\alpha_n$  such that

$$\int_{0}^{T} \left| g(t) - \sum_{n=1}^{\infty} \alpha_{n} \phi_{n}(t) \right|^{2} dt = 0$$

"Orthonormal" means

$$\int_{0}^{T} \phi_{n}(t) \phi_{m}^{*}(t) dt = \delta[n-m]$$



• While there are many CON sets in general, for the KL expansion, the  $\phi_n(t)$ 's must satisfy

$$\int_{0}^{T} R_{X}(t_{1}, t_{2}) \phi(t_{2}) dt_{2} = \lambda \phi(t_{1})$$

- These  $\phi_n(t)$ 's are eigenfunctions of  $R_X(t_1, t_2)$  with corresponding eigenvalues  $\lambda_n$
- The  $\phi_n(t)$ 's are called basis functions



 There are several different kinds of series expansions that are possible, depending on how the RP is constrained (frequency, periodic, or time)

The desirable feature is for the coefficients of the expansion to be statistically orthogonal

29



## **Thank You!**

Ve501 2020-2021 Fall