



Probability and Random Process

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Functions of Two Random Variables



Two Functions of Two RVs

Suppose X and Y are jointly distributed RVs with joint PDF $f_{XY}(x, y)$ and

$$Z = G(X, Y)$$

$$W = H(X, Y)$$

Examples:

Rectangular-to-Polar conversion

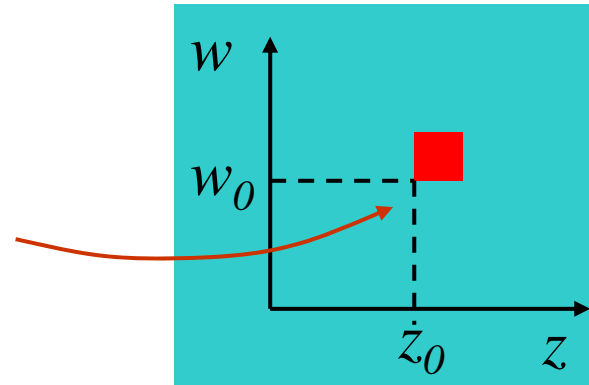
Rotation of Coordinates

Then one might wish to find $f_{ZW}(z, w)$

Consider the approximation:

$$P(\{z_0 < Z < z_0 + dz\} \cap \{w_0 < W < w_0 + dw\}) \\ \approx f_{zw}(z_0, w_0) dz dw$$

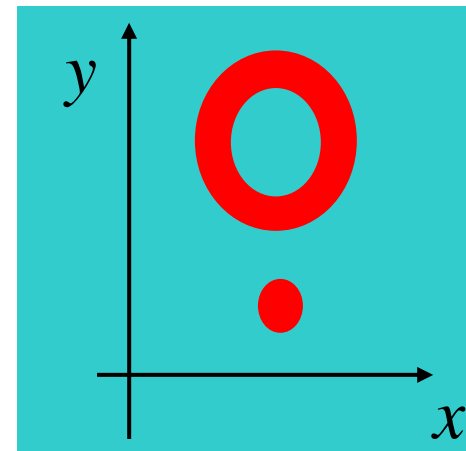
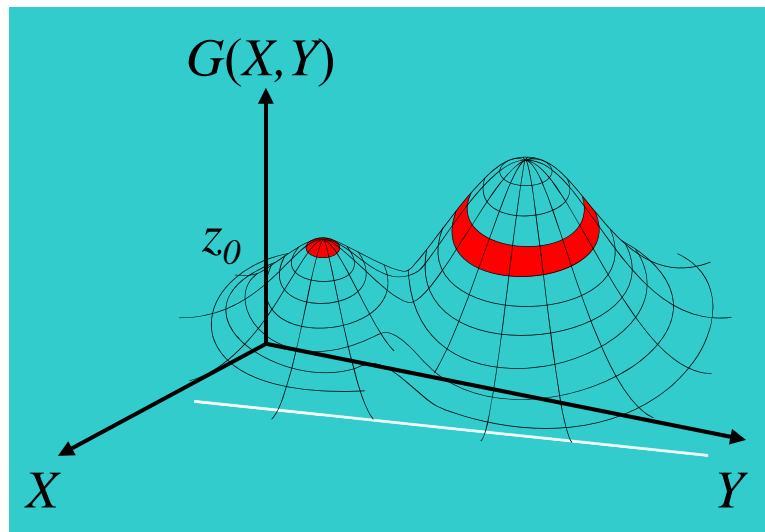
= The probability that (Z, W)
is in this small square



Graphical Example, Cont'd

First consider the event $\{z_0 < G(X, Y) < z_0 + dz\}$

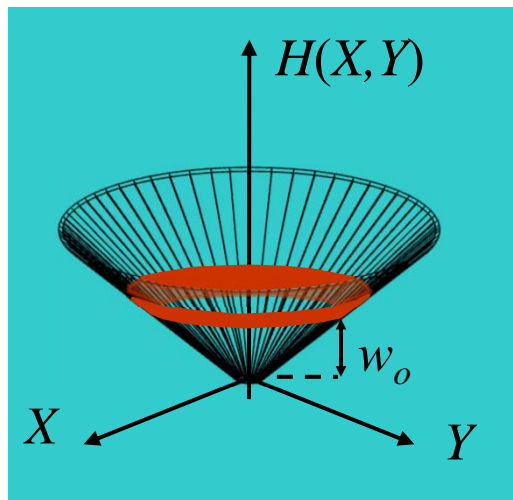
Example:



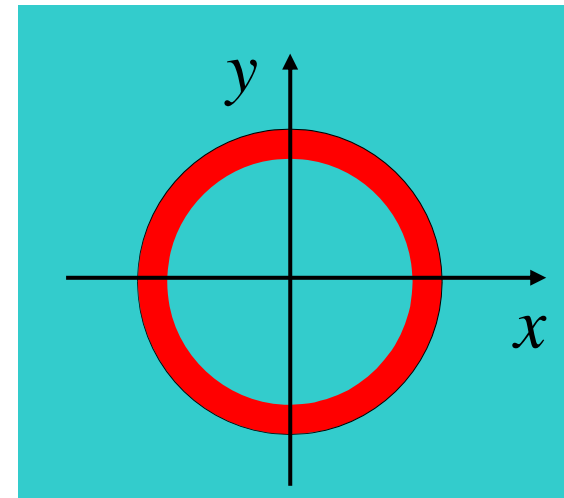
Event
Region

Graphical Example, Cont'd

Similarly, $\{w_0 < H(X, Y) < w_0 + dw\}$ also has an event region.



Event Region

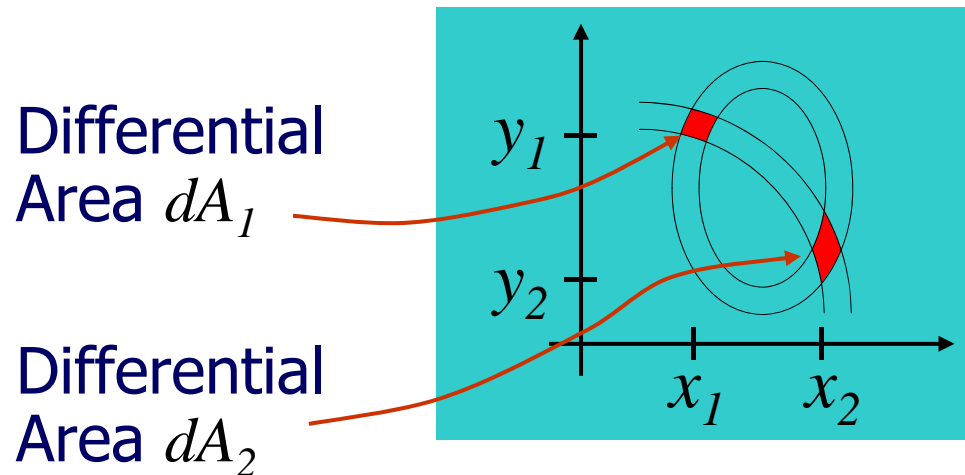


Graphical Example, Concluded

In this example,

$$P(\{z_0 < Z < z_0 + dz\} \cap \{w_0 < W < w_0 + dw\})$$

corresponds to the **intersection** of these two events regions.



(X_1, Y_1) and (X_2, Y_2)
are two solutions to
the equations:

$$z_0 = G(X, Y)$$

$$w_0 = H(X, Y)$$

Then we have

$$f_{zw}(z_0, w_0) dz dw \approx f_{XY}(x_1, y_1) dA_1 + f_{XY}(x_2, y_2) dA_2$$

It happens that

$$\frac{\partial A_j}{\partial z \partial w} \approx \frac{1}{|J(x_i, y_i)|},$$

where

$$J(x, y) = \det \begin{bmatrix} \frac{\partial z}{\partial x} & \frac{\partial z}{\partial y} \\ \frac{\partial w}{\partial x} & \frac{\partial w}{\partial y} \end{bmatrix} = \text{Jacobian of } G \text{ and } H$$

Recall the derivative in one RV case



General Formula

- In general, if there are n solutions to $z = G(X, Y)$ and $w = H(X, Y)$, (x_i, y_i) , $i = 1, 2, \dots, n$

- Then

$$f_{ZW}(z, w) = \sum_{i=1}^n \frac{f_{XY}(x_i, y_i)}{|J(x_i, y_i)|}$$

- This is similar to the formula for the function of a RV



Linear Example

Invertible Linear Transformation

$$\begin{bmatrix} Z \\ W \end{bmatrix} = \begin{bmatrix} a & b \\ c & d \end{bmatrix} \begin{bmatrix} X \\ Y \end{bmatrix}$$

Only one solution $A^{-1} = \frac{1}{|A|} A^*$

$$\begin{bmatrix} X \\ Y \end{bmatrix} = \begin{bmatrix} a & b \\ c & d \end{bmatrix}^{-1} \begin{bmatrix} Z \\ W \end{bmatrix} = \frac{1}{ad - cb} \begin{bmatrix} d & -b \\ -c & a \end{bmatrix} \begin{bmatrix} Z \\ W \end{bmatrix}$$

$$\text{So, } X = \frac{dZ - bW}{ad - cb} \quad Y = \frac{-cZ + aW}{ad - cb}$$

Jacobian

$$\det \begin{bmatrix} \frac{\partial z}{\partial x} & \frac{\partial z}{\partial y} \\ \frac{\partial w}{\partial x} & \frac{\partial w}{\partial y} \end{bmatrix} = \det \begin{bmatrix} a & b \\ c & d \end{bmatrix} = ad - bc$$

$$f_{zw}(z, w) = \frac{f_{xy} \left(\frac{dz - bw}{ad - cb}, \frac{-cz + aw}{ad - cb} \right)}{|ad - cb|}$$

S-plane to Z-plane mapping:

$$S = A + jB$$

$$Z = C + jD = e^{ST} = e^{(A + jB)T}$$

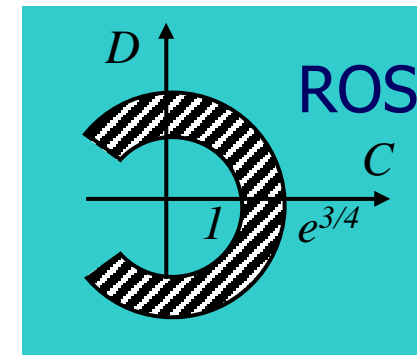
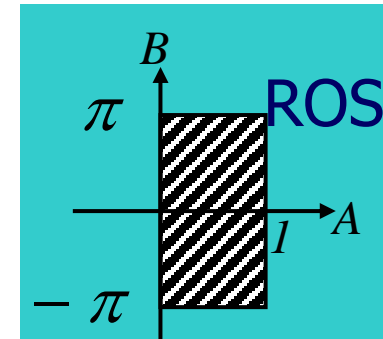
$$= e^{AT} e^{jBT} = e^{AT} \cos BT + je^{AT} \sin BT$$

$$\therefore C = e^{AT} \cos BT = G(A, B)$$

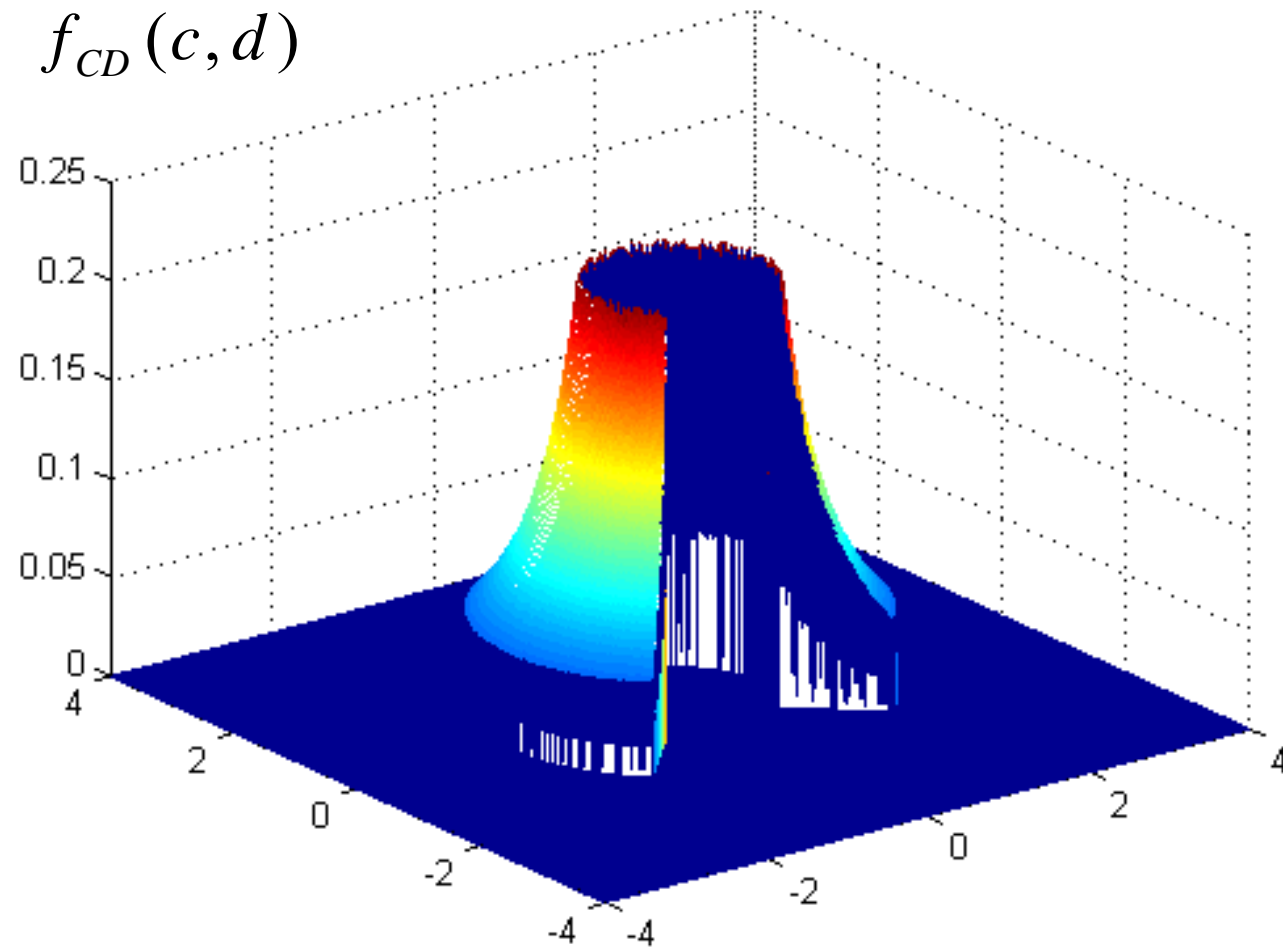
$$D = e^{AT} \sin BT = H(A, B)$$

$$\text{Let } f_{AB}(a, b) = \begin{cases} 1 & 0 < a < 1 \\ 2\pi & -\pi < b < \pi \end{cases}$$

and $T=3/4$



Nonlinear Example, Concluded





One Function of Two RVs

Two equally valuable approaches:

CDF Approach

Auxiliary RV approach

CDF approach: Given $f_{XY}(x, y)$ and $Z = G(X, Y)$.

Find $F_Z(z)$, then differentiate to get $f_Z(z)$.

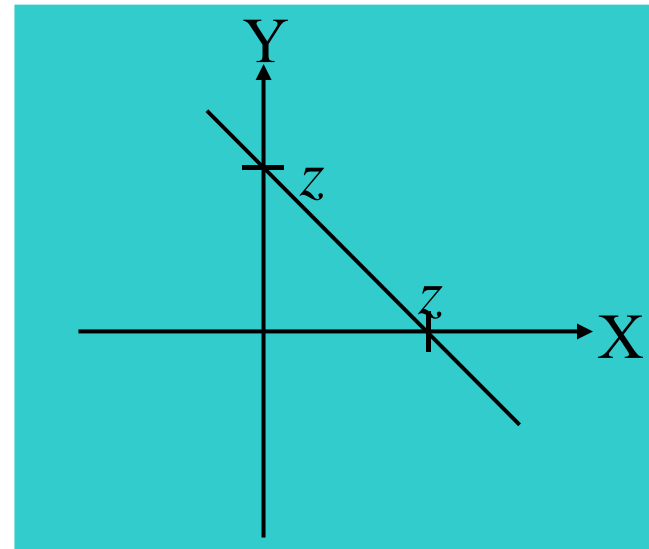
Auxiliary RV approach: Define an auxiliary or “dummy” RV W as either $W = X$ or $W = Y$. Use “two-functions-of-two-RVs” approach to get $f_{ZW}(z, w)$, then get marginal $f_Z(z)$.

Example: Adding RVs

$$Z = X + Y$$

$$\begin{aligned} F_z(z) &= P(Z \leq z) \\ &= P(X + Y \leq z) \end{aligned}$$

$$F_z(z) = \int_{-\infty}^{+\infty} \int_{-\infty}^{z-y} f_{XY}(x, y) dx dy$$



Recall Leibniz's Rule

If $\Phi(t) = \int_{a(t)}^{b(t)} f(x) dx,$

then $\Phi'(t) = f(b(t))b'(t) - f(a(t))a'(t)$

so $\frac{d}{dz} F_Z(z) = \int_{-\infty}^{+\infty} \left(\frac{d}{dz} \int_{-\infty}^{z-y} f_{XY}(x, y) dx \right) dy$

$\rightarrow f(z-y, y) - f(\infty, y) = 0$

$$f_Z(z) = \int_{-\infty}^{+\infty} f_{XY}(z-y, y) dy$$

Adding Independent RVs

Suppose, $Z=X+Y$ and X and Y are independent. Then

$$\begin{aligned} f_Z(z) &= \int_{-\infty}^{+\infty} f_{XY}(z-y, y) dy \\ &= \int_{-\infty}^{+\infty} f_X(z-y) f_Y(y) dy \\ &= f_X(z) * f_Y(z) \end{aligned}$$

↑
Convolution

★ REMEMBER

To **add** independent RVs, **convolve** their PDFs.



Example for discrete case

- Let X, Y be **non-negative integer valued** (discrete) r.v.'s that are **independent** and have pmfs $p_X(x)$ and $p_Y(y)$ respectively. Let $Z = X + Y$. Determine the pmf $p_Z(z)$.



Solution

$$p_Z(z) = P_r(Z = z) = P_r(X + Y = z)$$

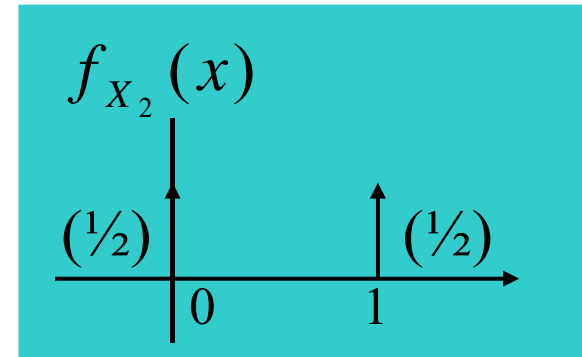
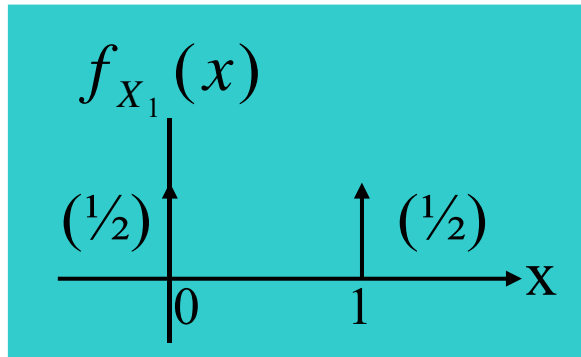
$$= \sum_{k=0}^z P_r(X = k, Y = z - k), \quad (\text{non-negative integer valued r.v.'s})$$

$$= \sum_{k=0}^z P_r(X = k) P_r(Y = z - k) \quad (\text{independence})$$

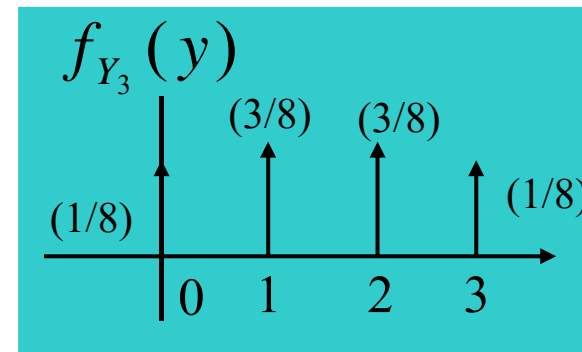
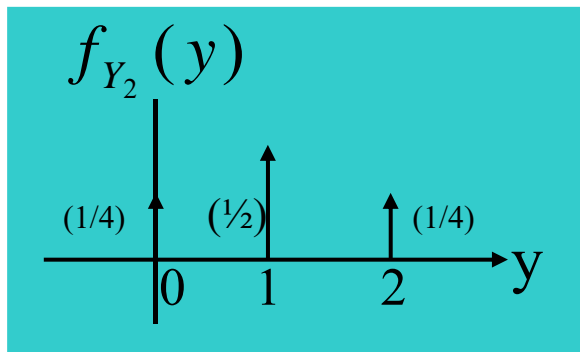
$$= \sum_{k=0}^z p_X(k) p_Y(z - k)$$

Adding Independent Bernoulli RVs

Let X_1 , X_2 and X_3 be iid Bernoulli RVs with $p=1/2$

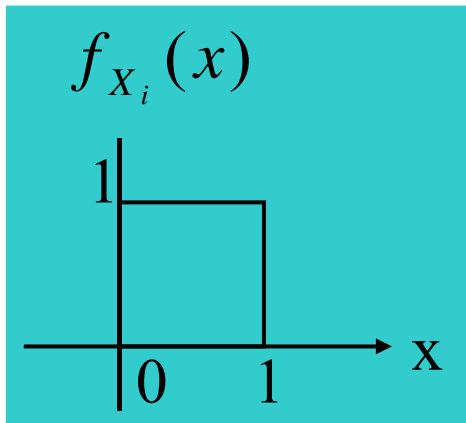


Let $Y_2 = X_1 + X_2$ and $Y_3 = Y_2 + X_3$

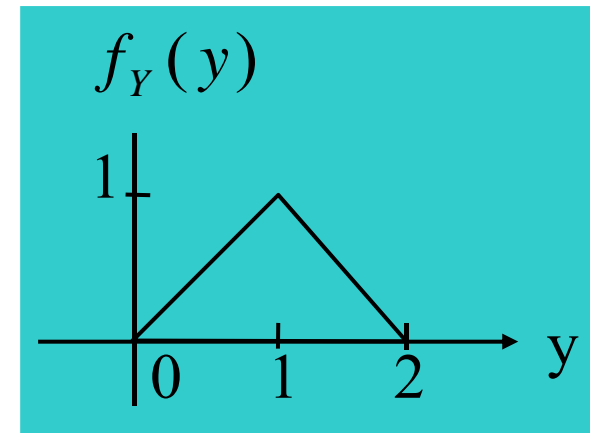


Adding Uniform RVs

Let X_1 and X_2 be i.i.d. with



Then $Y=X_1+X_2$ has the PDF ?





Example

- A random, continuous-valued signal X is transmitted over a channel subject to multiplicative, continuous-valued noise Y . The received signal is $Z = XY$. Find the cdf and density of Z if X and Y has a joint density $f_{XY}(x, y)$.

$$F_Z(z) = P(Z \leq z) = P(XY \leq z) = P((X, Y) \in A_z),$$

where $A_z := \{(x, y) : xy \leq z\}$ is partitioned into two disjoint regions, $A_z = A_z^+ \cup A_z^-$, as sketched in Figure 7.2. Next, since

$$F_Z(z) = P((X, Y) \in A_z^-) + P((X, Y) \in A_z^+),$$

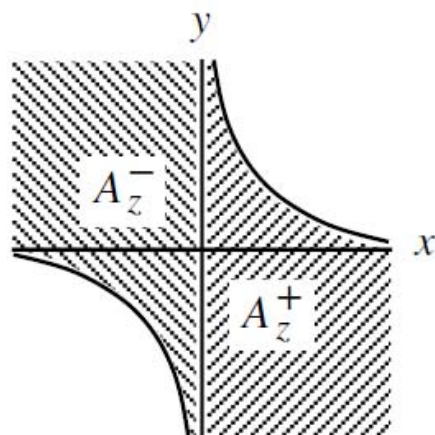


Figure 7.2. The curve is $y = z/x$. The shaded region to the left of the vertical axis is $A_z^- = \{(x, y) : y \geq z/x, x < 0\}$, and the shaded region to the right of the vertical axis is $A_z^+ = \{(x, y) : y \leq z/x, x > 0\}$. The sketch is for the case $z > 0$. How would the sketch need to change if $z = 0$ or if $z < 0$?



Solution

$$P((X,Y) \in A_z^+) = \int_0^\infty \left[\int_{-\infty}^{z/x} f_{XY}(x,y) dy \right] dx$$

and

$$P((X,Y) \in A_z^-) = \int_{-\infty}^0 \left[\int_{z/x}^\infty f_{XY}(x,y) dy \right] dx.$$

It follows that^b

$$f_Z(z) = \int_0^\infty f_{XY}\left(x, \frac{z}{x}\right) \frac{1}{x} dx - \int_{-\infty}^0 f_{XY}\left(x, \frac{z}{x}\right) \frac{1}{x} dx.$$

In the first integral on the right, the range of integration implies x is positive, and so we can replace $1/x$ with $1/|x|$. In the second integral on the right, the range of integration implies x is negative, and so we can replace $1/(-x)$ with $1/|x|$. Hence,

$$f_Z(z) = \int_0^\infty f_{XY}\left(x, \frac{z}{x}\right) \frac{1}{|x|} dx + \int_{-\infty}^0 f_{XY}\left(x, \frac{z}{x}\right) \frac{1}{|x|} dx.$$

Now that the integrands are the same, the two integrals can be combined to get

$$f_Z(z) = \int_{-\infty}^\infty f_{XY}\left(x, \frac{z}{x}\right) \frac{1}{|x|} dx.$$

Example

- Let $Y = \max(X_1, X_2)$, where X_1 , and X_2 are independent discrete r.v.'s with the given joint pmf

$$p_{X_1 X_2}(x_1, x_2) = p_{X_1}(x_1)p_{X_2}(x_2)$$

Let D_Y be the range space of Y .

$$D_Y = \{y_1, y_2, y_3, \dots\}, y_1 \leq y_2 \leq \dots$$

Compute the pmf of Y , i.e., $p_Y(y_i)$.

- $p_Y(y_i) = \Pr(Y = y_i) = \Pr(Y \leq y_i) - \Pr(Y \leq y_{i-1})$
- $\Pr(Y \leq y_i)$

$$\begin{aligned} &= \Pr(\max(X_1, X_2) \leq y_i) \\ &= \Pr(X_1 \leq y_i, X_2 \leq y_i) \\ &= P(\{s: X_1(s) \leq y_i, X_2(s) \leq y_i\}) \\ &= P(\{s: X_1(s) \leq y_i\} \cap \{s: X_2(s) \leq y_i\}) \\ &= P(\{s: X_1(s) \leq y_i\})P(\{s: X_2(s) \leq y_i\}) \text{ ---independent} \\ &= \Pr(X_1 \leq y_i) \Pr(X_2 \leq y_i) \\ &= [\sum_{x_i \leq y_i} p_{X_1}(x_i)][\sum_{x_i \leq y_i} p_{X_2}(x_i)] \end{aligned}$$



Auxiliary RV Example

Let $U = +\sqrt{XY}$, where X and Y are iid

$$f_X(x) = \begin{cases} \frac{1}{x^2} & x \geq 1 \\ 0 & \text{else} \end{cases} \quad f_Y(y) = \begin{cases} \frac{1}{y^2} & y \geq 1 \\ 0 & \text{else} \end{cases}$$

Let $V=X$ be the auxiliary RV.

1. The solution is:

$$X = V$$

$$Y = \frac{U^2}{V}$$

Auxiliary RV Example, Cont'd

2. Find Jacobian

$$J(x, y) = \det \begin{bmatrix} \frac{\partial u}{\partial x} & \frac{\partial u}{\partial y} \\ \frac{\partial v}{\partial x} & \frac{\partial v}{\partial y} \end{bmatrix}$$

$$\begin{aligned} U &= +\sqrt{XY} \\ V &= X \end{aligned}$$

$$= \det \begin{bmatrix} \text{don' t care} & \frac{1}{2} \sqrt{\frac{x}{y}} \\ 1 & 0 \end{bmatrix} = -\frac{1}{2} \sqrt{\frac{x}{y}}$$

3. Plug solution into Jacobian

$$X = V, \quad Y = \frac{U^2}{V}$$

$$-\frac{1}{2} \sqrt{\frac{x}{y}} = -\frac{1}{2} \sqrt{\frac{v}{\frac{u^2}{v}}} = -\frac{1}{2} \frac{v}{u}$$

4. PDF formula

$$f_{UV}(u, v) = \frac{f_{XY}\left(v, \frac{u^2}{v}\right)}{\left| -\frac{1}{2} \cdot \frac{v}{u} \right|} = \begin{cases} \frac{\frac{1}{v^2} \cdot \frac{v^2}{u^4}}{\left| -\frac{1}{2} \frac{v}{u} \right|} & v \geq 1, \quad \frac{u^2}{v} \geq 1 \\ 0 & o.w. \end{cases}$$

$$= \begin{cases} \frac{2}{u^3 v} & v \geq 1, \quad u^2 \geq v \\ 0 & o.w. \end{cases}$$

Plug arguments into
ROS for $f_{XY}(x, y)$

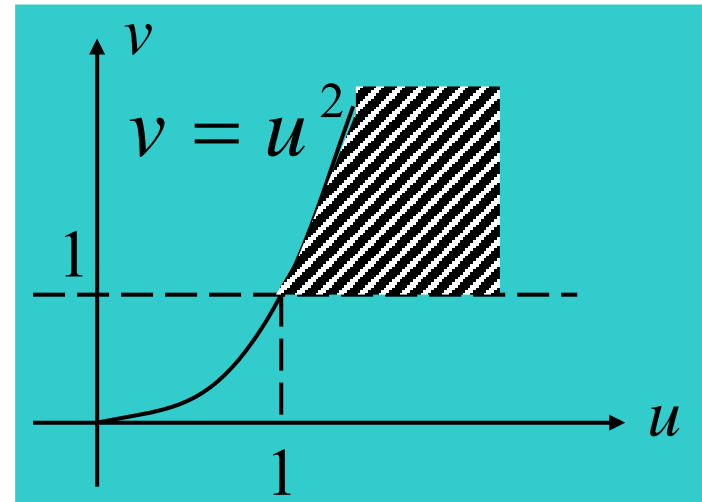
*$u > 0$ is understood
from the initial
definition.*

5. Find marginal $f_U(u)$.

Consider the ROS of $f_{UV}(u, v)$

$$f_U(u) = \int_{-\infty}^{+\infty} f_{UV}(u, v) dv = \int_1^{u^2} \frac{2}{u^3 v} dv$$

$$= \begin{cases} \frac{2 \ln(u^2)}{u^3} & u \geq 1 \\ 0 & o.w. \end{cases}$$





Expectation

- Given the joint pdf $f_{XY}(x, y)$, the **law of the unconscious statistician (LOTUS)** can easily be used to show that

$$E(g(X, Y)) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{XY}(x, y) dx dy$$

- Discrete case:

$$E(g(X, Y)) = \sum_{x, y} g(x, y) p_{XY}(x, y)$$



Example

X_1 and X_2 are discrete random variables. $Y = X_1 + X_2$.
 $E[Y] = ?$

$$\begin{aligned} E[Y] &= E[X_1 + X_2] = \sum_{x_1, x_2} (x_1 + x_2) p_{X_1 X_2}(x_1, x_2) \\ &= \sum_{x_1} \sum_{x_2} x_1 p_{X_1 X_2}(x_1, x_2) + \sum_{x_1} \sum_{x_2} x_2 p_{X_1 X_2}(x_1, x_2) \\ &= \sum_{x_1} x_1 \sum_{x_2} p_{X_1 X_2}(x_1, x_2) + \sum_{x_2} x_2 \sum_{x_1} p_{X_1 X_2}(x_1, x_2) \\ &= \sum_{x_1} x_1 p_{X_1}(x_1) + \sum_{x_2} x_2 p_{X_2}(x_2) \\ &= E[X_1] + E[X_2] \end{aligned}$$



Joint characteristic function

- For arbitrary random variables X and Y , their **joint characteristic function** is defined by

$$\varphi_{XY}(v_1, v_2) = E[e^{j(v_1X + v_2Y)}]$$

Joint characteristic function

If X and Y have joint pdf $f_{XY}(x, y)$, then

$$\varphi_{XY}(v_1, v_2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x, y) e^{j(v_1 x + v_2 y)} dx dy$$

which is just the **2D Fourier transform** of $f_{XY}(x, y)$ evaluated at $(-v_1, -v_2)$.

Using the **inverse Fourier transform**,

$$f_{XY}(x, y) = \frac{1}{(2\pi)^2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \varphi_{XY}(v_1, v_2) e^{-j(v_1 x + v_2 y)} dv_1 dv_2$$



Independence

- X and Y are **independent** if and only if their joint characteristic function factors into the product of the marginal characteristic functions

$$\varphi_{XY}(v_1, v_2) = \varphi_X(v_1) \varphi_Y(v_2)$$



Proof

If X and Y are independent

$$\begin{aligned}\varphi_{XY}(v_1, v_2) &= \mathbb{E} \left[e^{j(v_1 X + v_2 Y)} \right] \\ &= \mathbb{E} \left[e^{jv_1 X} \right] \mathbb{E} \left[e^{jv_2 Y} \right] \quad (\text{independence}) \\ &= \varphi_X(v_1) \varphi_Y(v_2)\end{aligned}$$

$$\begin{aligned} f_{XY}(x, y) &= \frac{1}{(2\pi)^2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \varphi_{XY}(v_1, v_2) e^{-j(v_1 x + v_2 y)} dv_1 dv_2 \\ &= \frac{1}{(2\pi)^2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \varphi_X(v_1) \varphi_Y(v_2) e^{-j(v_1 x + v_2 y)} dv_1 dv_2 \\ &= \left[\frac{1}{2\pi} \int_{-\infty}^{\infty} \varphi_X(v_1) e^{-jv_1 x} dv_1 \right] \left[\frac{1}{2\pi} \int_{-\infty}^{\infty} \varphi_Y(v_2) e^{-jv_2 y} dv_2 \right] \\ &= f_X(x) f_Y(y) \end{aligned}$$



Short Summary

Two functions of two RVs

Using Jacobian

$$f_{ZW}(z, w) = \sum_{i=1}^n \frac{f_{XY}(x_i, y_i)}{|J(x_i, y_i)|}$$

One function of two RVs

To add independent RVs, convolve their PDFs

CDF Approach

Auxiliary RV approach

Expectation



Thank You!