



Probability and Random Process

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- 2. Random Variables
 - Introduction to Random Variables
 - PMF and Discrete Random Variables
 - PDF and Continuous Random Variables
 - Gaussian CDF
 - Conditional Probability
 - Function of a RV
 - Expectation of a RV
 - Transform Methods and Probability Generating Function



PDF and Continuous Random Variables

The Probability Density Function

- The probability density function (PDF) is the derivative of the CDF, and denoted:

$$f_X(x) = \frac{d}{dx} F_X(x)$$

$$F_X(x) = \int_{-\infty}^x f_X(y) dy$$

- Some properties of $f_X(x)$ inherited from $F_X(x)$:

$$1. \lim_{x \rightarrow +\infty} F_X(x) = 1 \Rightarrow \int_{-\infty}^{+\infty} f_X(y) dy = 1$$

★REMEMBER PDFs have unit area.

$$2. F_X(x) \text{ non-decreasing} \Rightarrow f_X(x) \geq 0$$

$$3. P(a < X \leq b) = F_X(b) - F_X(a) = \int_a^b f_X(x) dx$$

PDFs must be **integrated** to get probability, hence the name “density”.

- Is $P(a \leq X \leq b) = F_X(b) - F_X(a)$?
 - If X is continuous, it is true, because $\Pr(X = a) = 0$.
 - If X is discrete, it may not be true.

- Does pdf $f_X(x)$ of a continuous r.v. need to be a continuous function?
 - No. e.g.. Exponential density is not continuous.(see later)

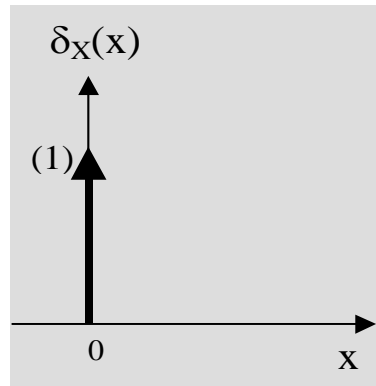
4. If $F_X(x)$ has a discontinuity at x_0 , then $f_X(x)$ has a delta function at x_0 with area equal to the height of the jump.

Dirac Delta Function Definition

- The function, denoted $\delta(x)$, has undefined (infinite) height and unit area. It is defined implicitly by the following equation:

$$U(x) = \int_{-\infty}^x \delta(y) dy,$$

where $U(x)$ is the unit step function.



It is OK to say:

$$\delta(x) = \frac{d}{dx} U(x)$$

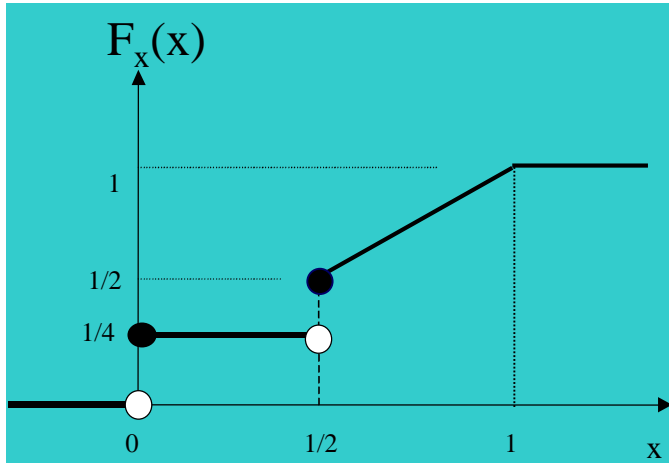
Dirac Delta Function Properties

- $\delta(x)$ is also called the “shifting function” because it “shifts out” one value from another function that multiplies it:

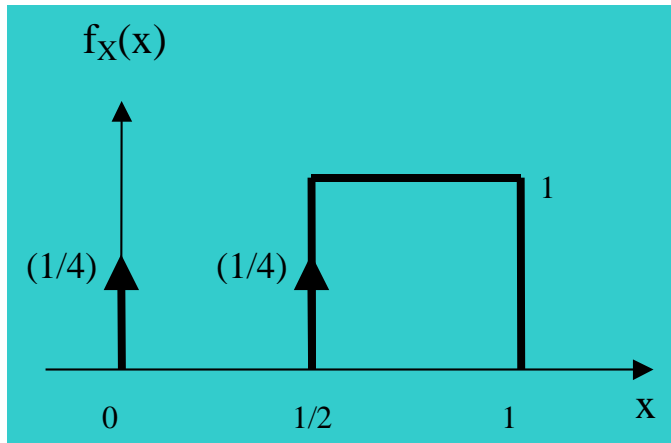
$$\int_{-\infty}^{+\infty} g(x)\delta(x-a)dx = g(a)$$

★ REMEMBER

A Dirac delta function only has meaning under an integral sign



$P(0 < X < 1/2) = \text{area under } f_X(x)$
 between 0 and $1/2$, but not
 including end points, so deltas
 are not included
 $= 0$



$$P(1/2 \leq X \leq 3/4) = \frac{1}{4} + \frac{1}{4} = \frac{1}{2}$$

*Area of impulse
 at $x=1/2$*

*Half the area
 of the
 rectangle*

- PDFs can be used for discrete RV's - these PDFs comprise only impulses.

Binomial PMF: $p_X(x) = \binom{n}{x} p^x q^{n-x} \quad x = 0, 1, \dots, n$

Binomial PDF: $f_X(x) = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k} \delta(x - k)$

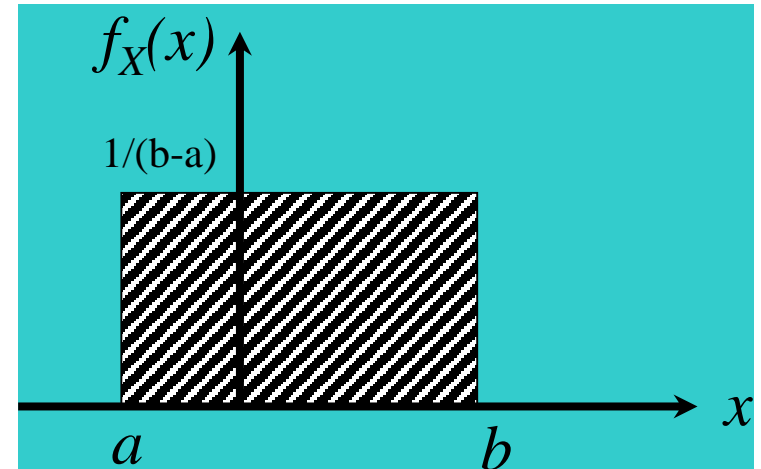


Some Special Continuous Random Variables

- Uniform
- Exponential
- Gaussian
- Rician
- Rayleigh
- Cauchy

$X \sim \text{uniform}[a, b]$ if $a < b$ and

$$f_X(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{o.w.} \end{cases}$$



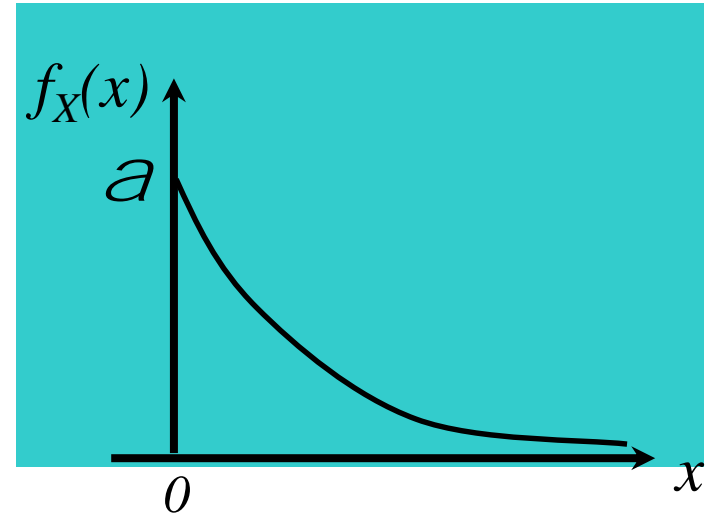
Shorthand: $X \sim U[a, b]$

Used to describe random quantities that we can bound, but otherwise know nothing about.

Ex: Phase of a radio frequency (RF) carrier at a receiver:
 $U[-\pi, \pi]$

$$X \sim \exp(\alpha)$$

$$f_X(x) = \begin{cases} \alpha e^{-\alpha x} & x \geq 0, \alpha > 0 \\ 0 & o.w. \end{cases}$$



- Used to describe times between randomly occurring events.
 - Time-to-failure
 - Inter-arrival times
- α is also called “rate parameter”
 - Relationship with Poisson r.v ?

CDF of Exponential R.V.

- The CDF of exponential R.V. is given by

$$F_X(x) = \begin{cases} 1 - e^{-\alpha x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

Gaussian (Normal)

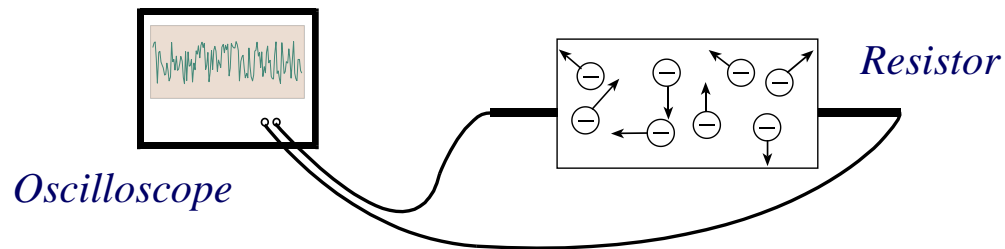
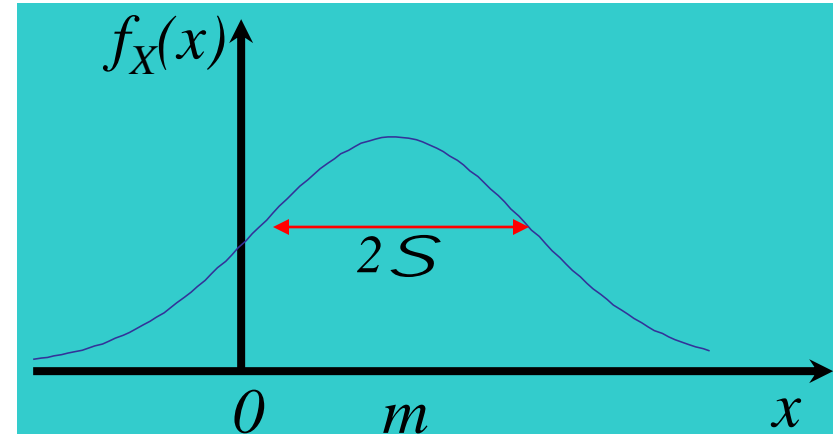
$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-m)^2}{2\sigma^2}}$$

Shorthand: $X \sim N(m, \sigma^2)$

$X \sim N(0,1)$ called standard normal.

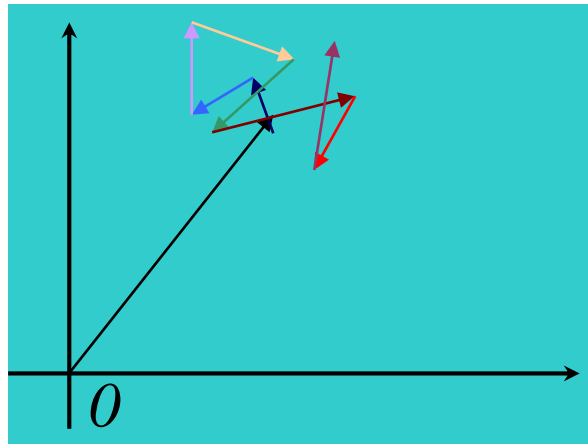
Used to model effects that are accumulations of large numbers of independent effects.

Ex: Thermal Noise is the voltage caused by the independent motions of thermally excited electrons in a resistance.



$$f_X(x) = \begin{cases} \frac{x}{\sigma^2} \exp\left\{-\frac{x^2 + A^2}{2\sigma^2}\right\} \cdot I_0\left(\frac{Ax}{\sigma^2}\right) & A \geq 0, x \geq 0 \\ 0 & x < 0 \end{cases}$$

X models the magnitude of the sum of one large known vector and lots of small random vectors.

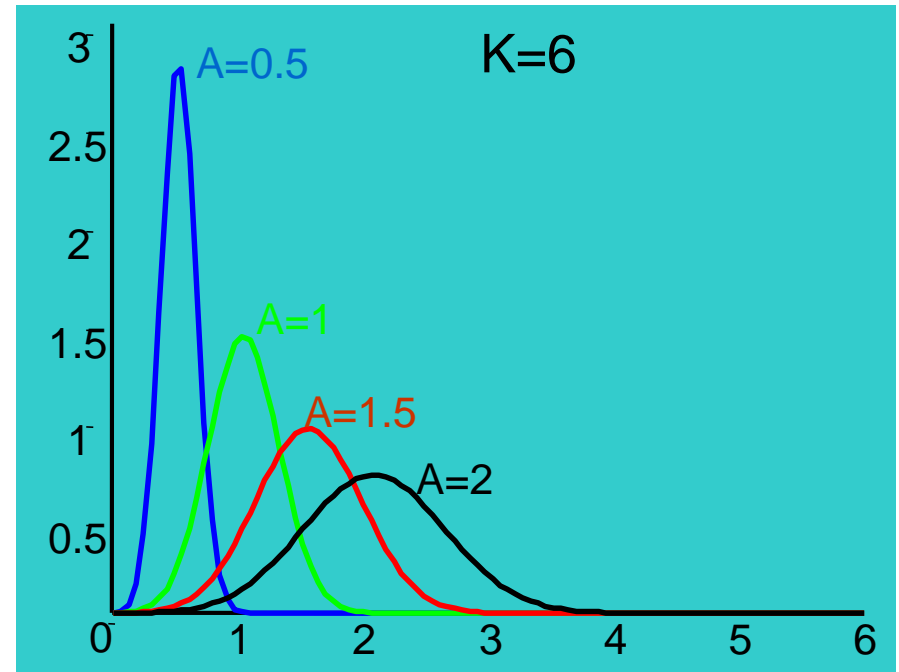
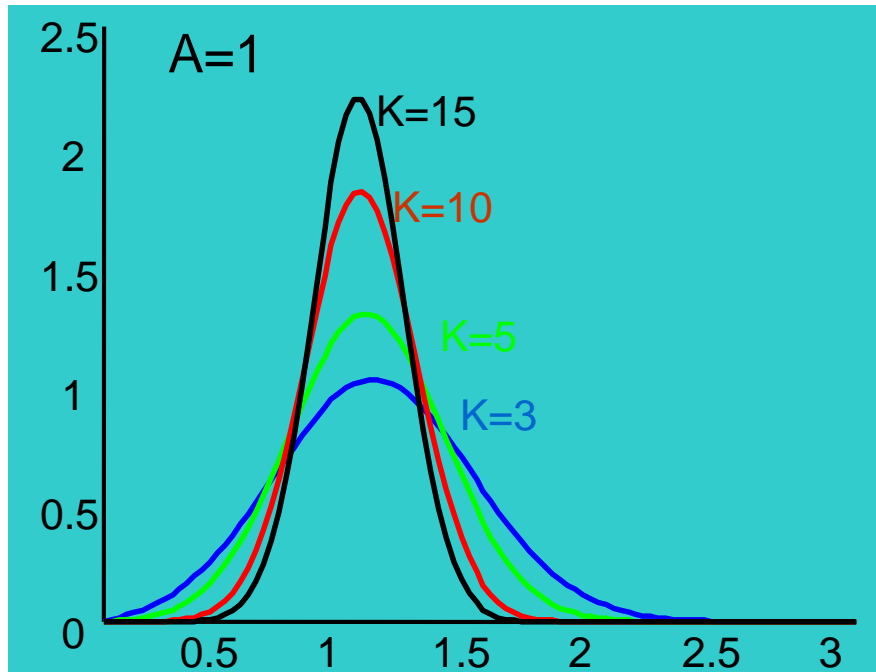


Bessel function of the first kind, order 0.

A is the length of the large vector.

In radio links with **line-of-sight (LOS)** and scattered paths, the signal envelope is Rician-distributed.

$K = A^2/2\sigma^2$ is called the “**K-factor**” or “Rician factor,” and represents the dominant component power over the average power of the sum of the other components.

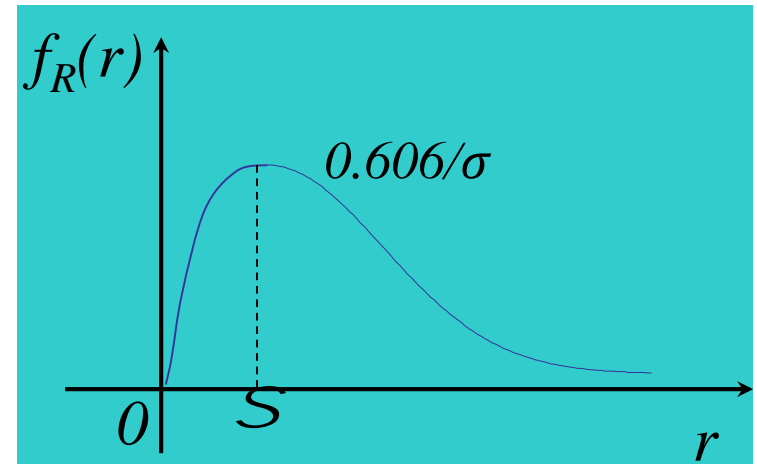


$$f_X(x) = \begin{cases} \frac{x}{\sigma^2} \exp\left\{-\frac{x^2}{2\sigma^2}\right\} & x \geq 0 \\ 0 & x < 0 \end{cases}$$

Special case of Rician when $A=0$.

Models envelopes in radio signals with no LOS path

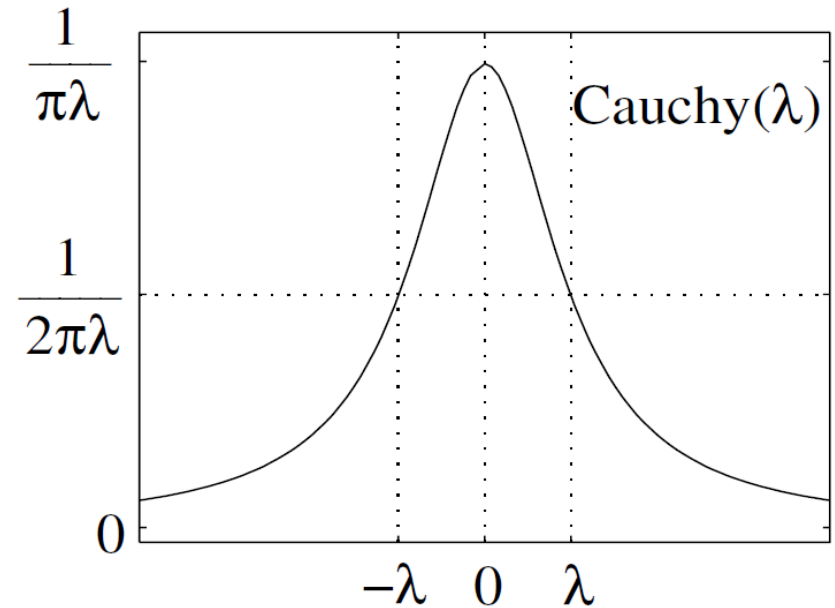
If $Z=X^2$, Z is an exponential RV.



- The Cauchy random variable with parameter $\lambda > 0$

$$f_X(x) = \frac{\lambda / \pi}{\lambda^2 + x^2}$$

- $X \sim \text{Cauchy}(\lambda)$



- The Cauchy random variable arises as the **tangent of a uniform random variable** and also as the **quotient of independent Gaussian random variables**.

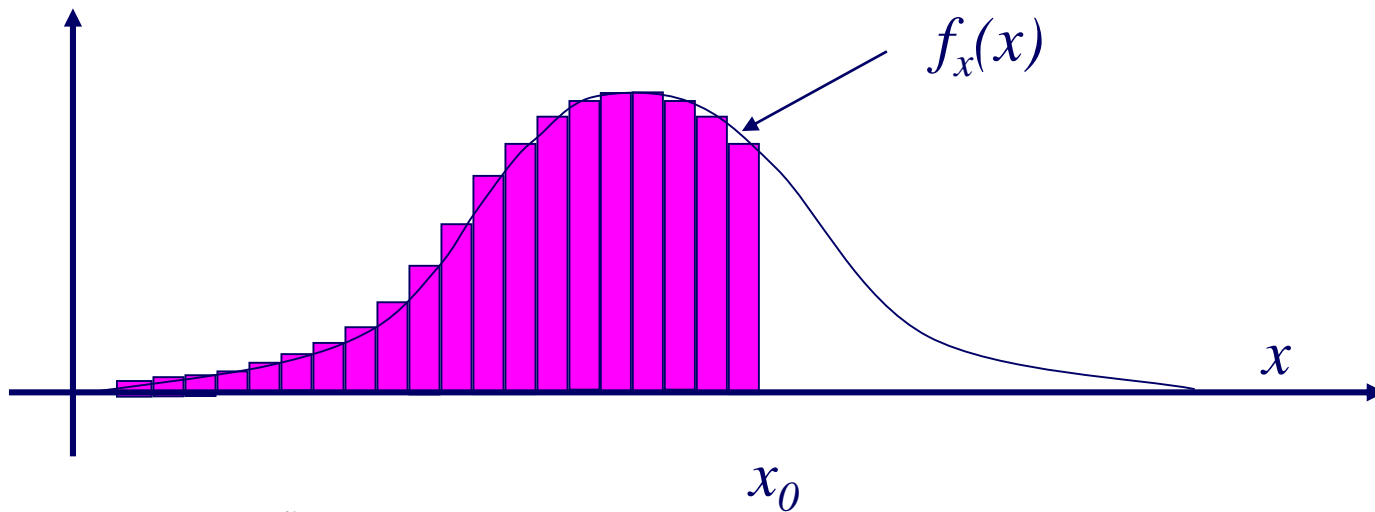
- PDFs are derivatives of CDFs
- Some continuous random variables
 - Uniform
 - Exponential
 - Gaussian
 - Rician
 - Rayleigh
 - Cauchy



Gaussian CDF

A Problem with the Gaussian PDF

★ **REMEMBER** The Gaussian PDF has no anti-derivative. In practice, it must be integrated numerically.



$$F_x(x_0) = \int_{-\infty}^{x_0} f_x(x) dx \approx \text{sum of rectangular areas}$$

Some Common Functions

- Several functions are commonly available in software packages, on calculators, and tabulated in books:

$\Phi(x)$ = CDF of the “standard normal” RV. The standard normal has zero mean and unit variance.

$$\Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy$$

$$Q(x) = 1 - \Phi(x) = \int_x^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy = \text{complement of } \Phi(x)$$

$\text{erf}(x)$ = error function

$\text{erfc}(x)$ = complementary error function

- Use change of variables in a CDF integral to make the integrand look like a standard normal pdf.

Ex : Let $X \sim N(5,9)$. Compute $P(X < 6.05)$.

Ex : Let $X \sim N(5,9)$. Compute $P(X < 6.05)$.

$$P(X < 6.05) = \int_{-\infty}^{6.05} \frac{1}{\sqrt{2\pi 9}} e^{-\frac{(x-5)^2}{2 \cdot 9}} dx$$

Let $y = \frac{x-5}{3}$. Then, $P(X < 6.05) =$

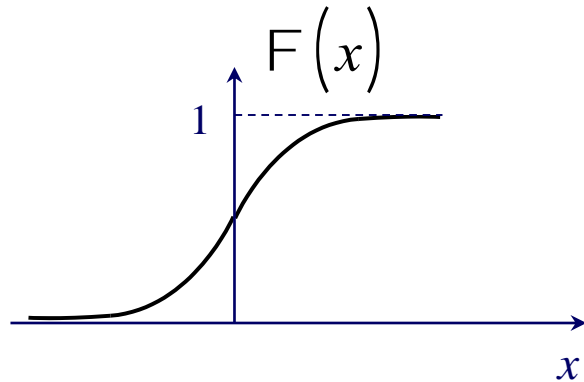
$$\int_{-\infty}^{\frac{6.05-5}{3}} \frac{1}{\sqrt{2\pi 9}} e^{-\frac{y^2}{2}} (3dy) = \int_{-\infty}^{0.35} \underbrace{\frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}}_{\text{std normal pdf}} dy = \Phi(0.35)$$

Use of $\Phi(x)$ Table

Part of numeric table for the Normal Distribution Function:

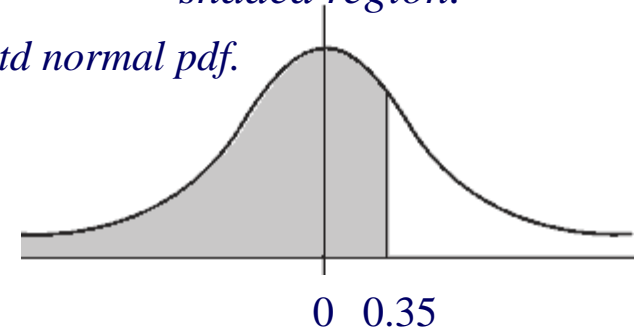
X	.00	0.01	0.02	.03	0.04	0.05
0.0	.5000	.5040	.5080	.5120	.5160	.5199
0.1	.5398	.5438	.5478	.5517	.5557	.5596
0.2	.5793	.5832	.5871	.5910	.5948	.5987
0.3	.6179	.6217	.6255	.6293	.6331	.6368
0.4	.6554	.6591	.6628	.6664	.6700	.6736
0.5	.6915	.6950	.6985	.7019	.7054	.7088
0.6	.7257	.7291	.7324	.7357	.7389	.7422
0.7	.7580	.7611	.7642	.7673	.7704	.7734
0.8	.7881	.7910	.7939	.7957	.7995	.8023
0.9	.8159	.8186	.8212	.8238	.8264	.8289

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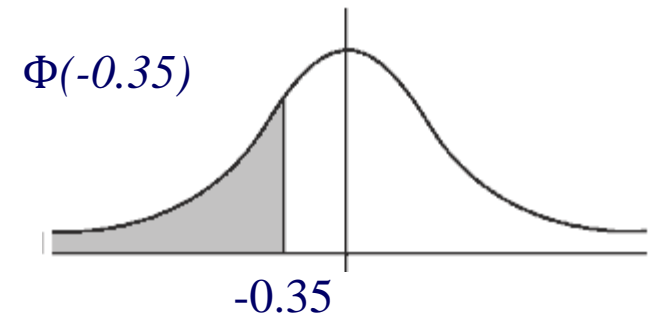


$\Phi(0.35)$ is the area of the shaded region.

Std normal pdf.



$$F(0.35) = 1 - F(-0.35) = 0.6368$$



- The CDF of a Gaussian random variable is given by

$$\begin{aligned} F(x) &= \int_{-\infty}^x \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(t-m)^2}{2\sigma^2}} dt \\ &= 1 - \int_x^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(t-m)^2}{2\sigma^2}} dt \\ &= 1 - \int_{\frac{x-m}{\sigma}}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}} du \\ &= 1 - Q\left(\frac{x-m}{\sigma}\right) \end{aligned}$$

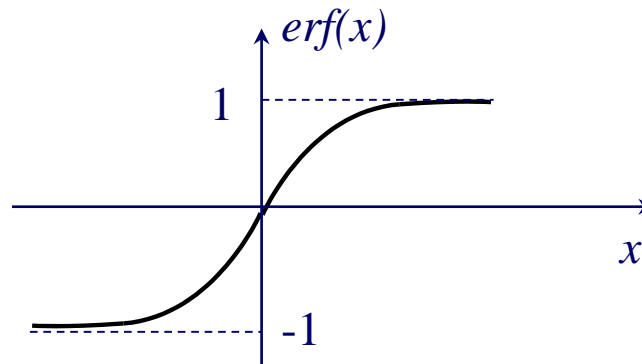
Using $Q(x)$ Table

Table of Q Function Values

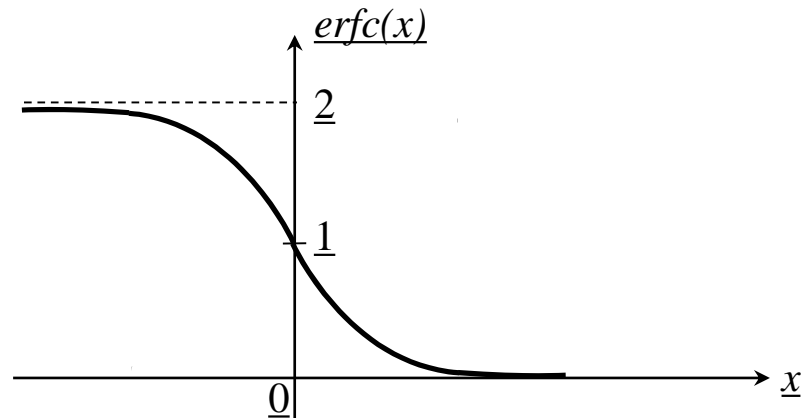
x	$Q(x)$	x	$Q(x)$	x	$Q(x)$	x	$Q(x)$
0	0.500000	1.8	0.035930	3.6	0.000159	5.4	3.3320×10^{-8}
0.1	0.460170	1.9	0.028717	3.7	0.000108	5.5	1.8990×10^{-8}
0.2	0.420740	2	0.022750	3.8	7.2348×10^{-5}	5.6	1.0718×10^{-8}
0.3	0.382090	2.1	0.017864	3.9	4.8096×10^{-5}	5.7	5.9904×10^{-9}
0.4	0.344580	2.2	0.013903	4	3.1671×10^{-5}	5.8	3.3157×10^{-9}
0.5	0.308540	2.3	0.010724	4.1	2.0658×10^{-5}	5.9	1.8175×10^{-9}
0.6	0.274250	2.4	0.008198	4.2	1.3346×10^{-5}	6	9.8659×10^{-10}
0.7	0.241960	2.5	0.006210	4.3	8.5399×10^{-6}	6.1	5.3034×10^{-10}
0.8	0.211860	2.6	0.004661	4.4	5.4125×10^{-6}	6.2	2.8232×10^{-10}
0.9	0.184060	2.7	0.003467	4.5	3.3977×10^{-6}	6.3	1.4882×10^{-10}
1	0.158660	2.8	0.002555	4.6	2.1125×10^{-6}	6.4	7.7689×10^{-11}
1.1	0.135670	2.9	0.001866	4.7	1.3008×10^{-6}	6.5	4.0160×10^{-11}
1.2	0.115070	3	0.001350	4.8	7.9333×10^{-7}	6.6	2.0558×10^{-11}
1.3	0.096800	3.1	0.000968	4.9	4.7918×10^{-7}	6.7	1.0421×10^{-11}
1.4	0.080757	3.2	0.000687	5	2.8665×10^{-7}	6.8	5.2309×10^{-12}
1.5	0.066807	3.3	0.000483	5.1	1.6983×10^{-7}	6.9	2.6001×10^{-12}
1.6	0.054799	3.4	0.000337	5.2	9.9644×10^{-8}	7	1.2799×10^{-12}
1.7	0.044565	3.5	0.000233	5.3	5.7901×10^{-8}	7.1	6.2378×10^{-13}

- erf(x) is a version of $\Phi(x)$ with odd symmetry, extreme values +1 and -1.

$$\begin{aligned}\operatorname{erf}(x) &= 2\Phi(\sqrt{2}x) - 1 \\ &= 2 \int_{-\infty}^x \frac{1}{\sqrt{\pi}} e^{-z^2} dz - 1 \\ &= \frac{2}{\sqrt{\pi}} \int_0^x e^{-z^2} dz\end{aligned}$$



$$\operatorname{erfc}(x) = 1 - \operatorname{erf}(x) = \frac{2}{\sqrt{\pi}} \int_x^{\infty} e^{-z^2} dz$$



Gaussian Integration

- Integrands that are scaled Gaussian pdfs must be numerically integrated. Learn to recognize the form:

$$Ke^{-ax^2+bx+c}$$

- K , b and c are any real numbers. “ a ” must be a positive real number.

Gaussian Integration Example

- Typical problem: Let a RV X have the following PDF:

$$f_X(x) = \begin{cases} Ke^{-4x^2+3x} & x > 2 \\ 0 & \text{ow} \end{cases}$$

What is K ? Use the fact that $\int_{-\infty}^{\infty} f_X(x)dx = 1$

$$K = \frac{1}{\int_2^{\infty} e^{-4x^2+3x} dx} = \frac{1}{e^{\frac{9}{16}} \sqrt{\frac{\pi}{4}} \cdot Q\left(\frac{13}{\sqrt{8}}\right)}$$

Can you derive it?

Gaussian Integration Example - I

To find K , use the property of PDFs: $\int_{-\infty}^{\infty} f_X(x) dx = 1$

$$\int_2^{\infty} K e^{-4x^2 + 3x} dx = 1$$

Complete the square in the exponent:

$$-4x^2 + 3x = -4\left(x^2 - \frac{3}{4}x\right) = -4\left[x^2 - \frac{3}{4}x + \left(\frac{3}{8}\right)^2 - \left(\frac{3}{8}\right)^2\right] = -4\left(x - \frac{3}{8}\right)^2 + 4\left(\frac{3}{8}\right)^2$$

Substitute into integral and use $e^{a+b} = e^a e^b$:

$$\int_2^{\infty} K e^{-4\left(x - \frac{3}{8}\right)^2 + 4\left(\frac{3}{8}\right)^2} dx = K e^{4\left(\frac{3}{8}\right)^2} \int_2^{\infty} e^{-4\left(x - \frac{3}{8}\right)^2} dx$$

Gaussian Integration Example - II

Now look at the exponent and match items with $-\frac{(x-m)^2}{2\sigma^2}$

$$-4\left(x - \frac{3}{8}\right)^2 = -\frac{(x-m)^2}{2\sigma^2}$$

$$\Rightarrow m = \frac{3}{8}, \quad \sigma^2 = \frac{1}{8}$$

Gaussian Integration Example - III

Construct proper coefficient to make integrand a Gaussian PDF (i.e. the $\frac{1}{\sqrt{2\pi\sigma}}$ factor)

$$\int_2^{\infty} K e^{-4x^2+3x} dx$$

$$= K e^{4\left(\frac{3}{8}\right)^2} \sqrt{2\pi \frac{1}{8}} \cdot \int_2^{\infty} \underbrace{\frac{1}{\sqrt{2\pi \frac{1}{8}}} e^{\frac{-4\left(x-\frac{3}{8}\right)^2}{2 \cdot \frac{1}{8}}}}_{\text{A Gaussian pdf for } X \sim N\left(\frac{3}{8}, \frac{1}{8}\right)} dx$$

Gaussian Integration Example - IV

The change of variables $y = \frac{x-m}{\sigma} = \frac{x - \frac{3}{8}}{\sqrt{\frac{1}{8}}}$

gives the $Q(x)$ form:

$$\int_2^{\infty} K e^{-4x^2+3x} dx = K e^{4\left(\frac{3}{8}\right)^2} \sqrt{2\pi \frac{1}{8}} \cdot Q\left(\frac{2-3/8}{\sqrt{1/8}}\right) = 1$$

$$\therefore K = \frac{1}{e^{\frac{9}{16}} \sqrt{\frac{\pi}{4}} \cdot Q\left(\frac{13}{\sqrt{8}}\right)}$$

- Gaussian CDF
- Some common functions
 - $\Phi(x)$ CDF of the standard Gaussian RV
 - Q function
 - Error function
 - Complementary error function
- Gaussian integration



Conditional Probability

Conditional Probability

- Let B be an event with $P(B) > 0$. For any event A , we define the conditional probability of A given B as:

$$P(A | B) = \frac{P(A \cap B)}{P(B)} = \frac{P(AB)}{P(B)}$$

$P(AB)$ is shorthand for $P(A \cap B)$

Conditioning a RV on an Event Involving Only that RV

- Here, we consider events of non-zero probability.
- Define the conditional CDF of X given the event B , assuming $P(B) > 0$, as:

$$F_{X|B}(x|B) = P(X \leq x | B) = \frac{P(\{X \leq x\} \cap B)}{P(B)}$$

- $F_{X|B}(x|B)$ is a valid CDF:
 1. $0 \leq F_{X|B}(x|B) \leq 1$
 2. $\lim_{x \rightarrow +\infty} F_{X|B}(x|B) = 1$
 3. $F_{X|B}(x|B)$ is non-decreasing

Proof of Property 2

$\{X \rightarrow \infty\}$ is the certain event Ω

$$\therefore \{X \leq +\infty\} \cap B = B$$

$$F_{X|B}(+\infty | B) = \frac{P(B)}{P(B)} = 1$$

Proof of Property 3

Let $x_1 \leq x_2$. Then,

$$\{X \leq x_2\} = \{X \leq x_1\} \cup \{x_1 < X \leq x_2\}$$

disjoint

Next, intersect both sides with the event B .

$$\{X \leq x_2\} \cap B = (\{X \leq x_1\} \cap B) \cup (\{x_1 < X \leq x_2\} \cap B)$$

still disjoint

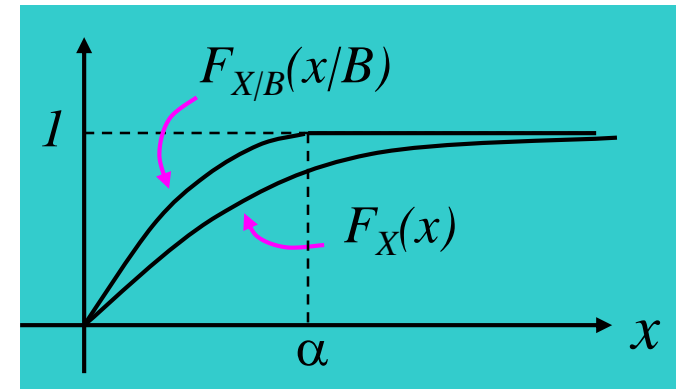
Take probability of both sides,

$$\begin{aligned} P(\{X \leq x_2\} \cap B) &= P(\{X \leq x_1\} \cap B) + \underbrace{P(\{x_1 < X \leq x_2\} \cap B)}_{\text{This term is } \geq 0} \\ \Rightarrow P(\{X \leq x_1\} \cap B) &\leq P(\{X \leq x_2\} \cap B) \\ \Rightarrow F_{X|B}(x_1) &\leq F_{X|B}(x_2) \end{aligned}$$

Example

Suppose X is an exponential RV
and let $B = \{X \leq \alpha\}$

$$\{X \leq x\} \cap B = \begin{cases} X \leq x & x < \alpha \\ X \leq \alpha & x \geq \alpha \end{cases}$$



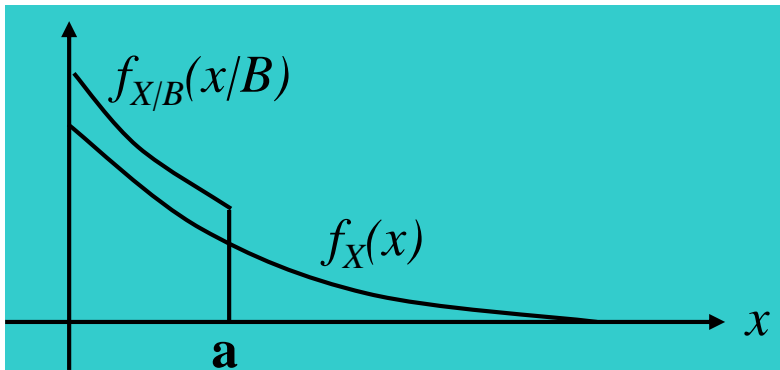
$$F_{X|B}(x | B) = \frac{P(\{X \leq x\} \cap B)}{P(B)} = \begin{cases} \frac{F_X(x)}{F_X(\alpha)} & x < \alpha \\ 1 & x \geq \alpha \end{cases}$$

Continuing
the example,

$$f_{X|B}(x | B) = \frac{d}{dx} F_{X|B}(x | B)$$

$$f_{X|B}(x | B) = \begin{cases} \frac{d}{dx} \frac{F_X(x)}{F_X(\alpha)} & x < \alpha \\ 0 & x \geq \alpha \end{cases}$$

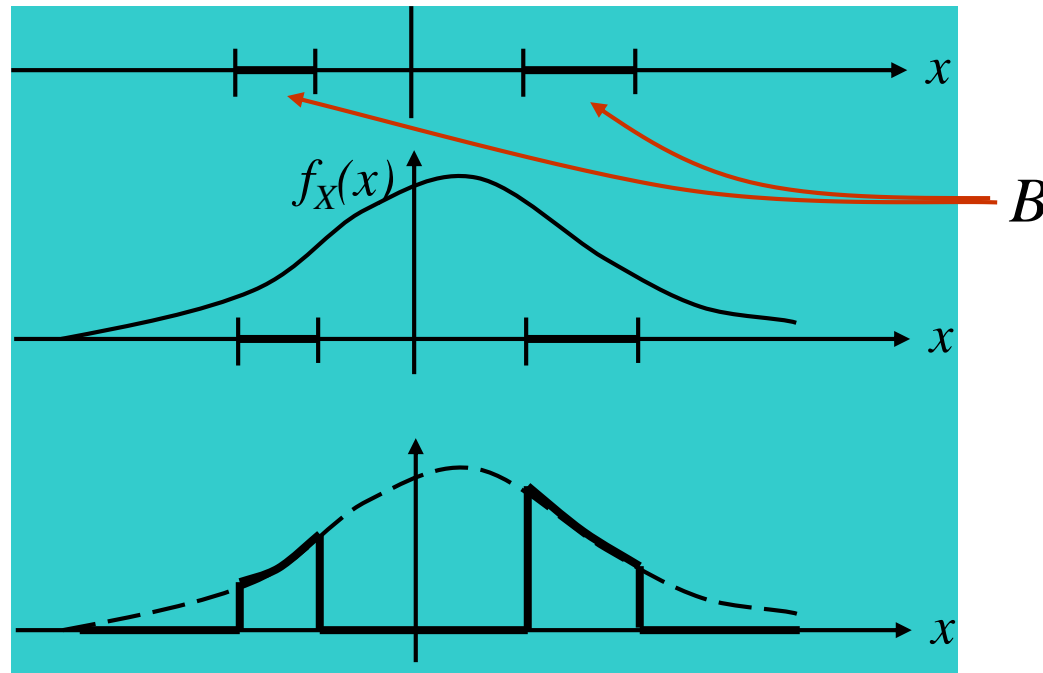
$$= \begin{cases} \frac{f_X(x)}{F_X(\alpha)} & x < \alpha \\ 0 & x \geq \alpha \end{cases}$$



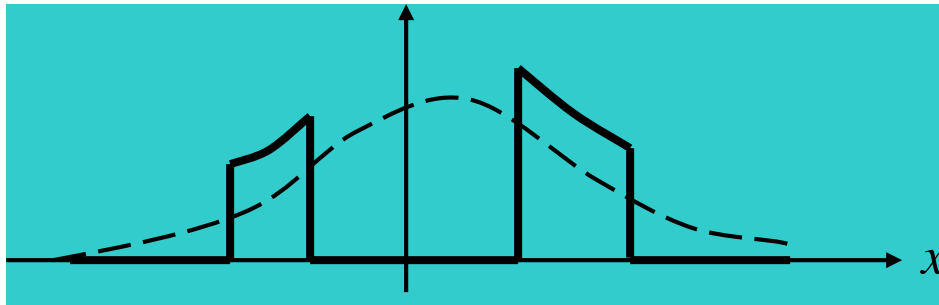
Graphical Interpretation

When B is a union of intervals on X , two steps are taken:

1. “zero-out” $f_X(x)$ everywhere but on these intervals.



2. Normalize (i.e. scale) zeroed-out version by its own area to get a unit-area function.



- Conditional probability
- Conditioning a RV on an event involving only that RV

$$F_{X|B}(x | B) = P(X \leq x | B) = \frac{P(\{X \leq x\} \cap B)}{P(B)}$$

$$f_{X|B}(x | B) = \frac{d}{dx} F_{X|B}(x | B)$$

- Graphical interpretation
 - Zero-out
 - Normalization



Thank You!