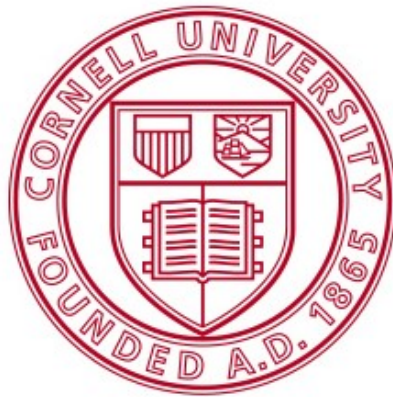


ECE 4110/ ECE 5110

Random Signals in Communication and Signal Processing



Homework 6

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P1. Solution:

$$(i) E\{(X_{n+1} - \alpha X_n)^2\} = \text{var}(X_{n+1} - \alpha X_n)$$

$$= \text{var}(X_{n+1}) - 2\alpha \text{var}(X_{n+1}, X_n) + \alpha^2 \text{var}(X_n)$$

$$= \alpha_0 - 2\alpha \alpha_1 + \alpha^2 \alpha_0 = (1 + \alpha^2) \alpha_0 - 2\alpha \alpha_1$$

$$\alpha_0 = \frac{1}{1-\alpha^2} \quad \alpha_1 = \frac{\alpha}{1-\alpha^2} \Rightarrow \alpha = \frac{\alpha_1}{\alpha_0}$$

$$\text{Thus, } \hat{X}_{n+1} = \frac{\alpha_1}{\alpha_0} X_n$$

$$(ii) E\{(X_{n+1} - \beta X_n - \gamma X_{n-1})^2\} = \text{var}(X_{n+1} - \beta X_n - \gamma X_{n-1}) \quad (\beta = \alpha_0, \gamma = \alpha_1)$$

$$= (1 + \beta^2 + \gamma^2) \alpha_0 - 2\beta \alpha_1 - 2\gamma \alpha_2 + 2\beta \gamma \alpha_1$$

$$= (1 + \beta^2 + \gamma^2) \alpha_0 + 2\beta(\gamma - 1) \alpha_1 - 2\gamma \alpha_2$$

$$\hat{X}_{n+1} = \sum_{i=0}^{\infty} a_i X_{n-i} \Rightarrow \begin{cases} \beta \alpha_0 + \gamma \alpha_1 = \alpha_1 \\ \beta \alpha_1 + \gamma \alpha_0 = \alpha_2 \end{cases}$$

$$\Rightarrow \begin{cases} \beta = \frac{\alpha_1(\alpha_0 - \alpha_2)}{\alpha_0^2 - \alpha_1^2} \\ \gamma = \frac{\alpha_0 \alpha_2 - \alpha_1^2}{\alpha_0^2 - \alpha_1^2} \end{cases}$$

$$\hat{X}_{n+1} = \frac{\alpha_1(\alpha_0 - \alpha_2)}{\alpha_0^2 - \alpha_1^2} X_n + \frac{\alpha_0 \alpha_2 - \alpha_1^2}{\alpha_0^2 - \alpha_1^2} X_{n-1}$$

(iii) From (i) and (ii) we have

$$E(X_{n+1} - \hat{X}_{n+1})^2 = (1 + \frac{\alpha_1^2}{\alpha_0^2}) \alpha_0 - 2 \frac{\alpha_1}{\alpha_0} \alpha_1 = \frac{\alpha_0^2 - \alpha_1^2}{\alpha_0}$$

$$E(X_{n+1} - \tilde{X}_{n+1})^2 = \frac{\alpha_0 \alpha_1^2 [\alpha_0 - \alpha_2]^2 + \alpha_0^3 [\alpha_2 - \alpha_1]^2 + [\alpha_0 - \alpha_1]^2 [\alpha_0 + \alpha_1]^2 + 2\alpha_1^2 (\alpha_0 - \alpha_1) [\alpha_0 \alpha_2 - \alpha_1^2]}{[\alpha_0^2 - \alpha_1^2]^2}$$

$$\Rightarrow D = \frac{[\alpha_1^2 - \alpha_0 \alpha_2]^2}{\alpha_0 [\alpha_0^2 - \alpha_1^2]}$$

(a) In this case, $\alpha_0 = \frac{1}{2}$, $\alpha_1 = 0$, $\alpha_2 = 0 \Rightarrow \hat{X}_{n+1} = \tilde{X}_{n+1} = 0$, $D = 0$

(b) In this case, $\alpha_0 = 1$, $\alpha_1 = \alpha$, $\alpha_2 = \alpha^2 \Rightarrow D = 0$

P2. Solution:

Let $\{Z_n: n = \dots, -1, 0, 1, \dots\}$ be independent R.V.s with zero mean and unit variances, define $X_n = \frac{Z_n + aZ_{n-1}}{\sqrt{1+a^2}}$ satisfies the autocovariance.

By the projection theorem, $X_n - \hat{X}_n$ is orthogonal to $\{X_{n-r}: r \geq 1\}$

Hence $E\{(X_n - \hat{X}_n)X_{n-r}\} = 0, r \geq 1$ and set $\hat{X}_n = \sum_{s=1}^{\infty} b_s X_{n-s}$ to obtain that

$$\begin{cases} \alpha = b_1 + b_2 \alpha \\ 0 = b_{s-1} \alpha + b_s + b_{s+1} \alpha, \text{ (for } s \geq 2) \end{cases} \text{ where } \alpha = \frac{a}{1+a^2}$$

$$\Rightarrow b_s = (-1)^{s+1} a^s$$

Hence $\hat{X}_n = \sum_{s=1}^{\infty} (-1)^{s+1} a^s X_{n-s}$, then the mean square error of prediction is

$$E\{(X_n - \hat{X}_n)^2\} = E\left\{\left(\sum_{s=0}^{\infty} (-a)^s X_{n-s}\right)^2\right\} = \frac{1}{1+a^2} E(Z_n^2) = \frac{1}{1+a^2}$$

$$E(\hat{X}_n) = 0$$

$$\text{cov}(\hat{X}_n, \hat{X}_{n-m}) = \sum_{r,s=1}^{\infty} b_r b_s \text{cov}(X_{n-r}, X_{n-m-s}), m \geq 0 \text{ depends on } m \text{ alone.}$$

Thus \hat{X} is weakly stationary.

~~P3~~

~~$$(a) p(t) = e^{-\frac{1}{2}t^2}$$~~

~~$$(b) p(t) = \frac{1}{2} \left(\frac{1}{1-it} + \frac{1}{1+it} \right) = \frac{1}{1+t^2}$$~~

P3. Solution:

$$R_X(\tau) = \int_{-\infty}^{\infty} S_X(f) e^{2j\pi f\tau} df$$

(a) For $S_X(f) = N(0,1) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{1}{2}(f)^2}$

Then $R_X(\tau) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}f^2} \cdot e^{2j\pi f\tau} df = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}f^2 + 2j\pi f\tau} df$
 $= \frac{1}{\sqrt{2\pi}} e^{-2\pi^2\tau^2} e^{-\pi^2\tau^2} \quad (\text{from table})$

(b) From table we know $\int_{-\infty}^{\infty} e^{-2\pi k|x|} e^{-2j\pi f\tau} dx = \frac{1}{\pi} \cdot \frac{k}{f^2 + k^2}$

Let $k = \frac{1}{2\pi}$, then we have $\int_{-\infty}^{\infty} e^{-|x|} e^{2j\pi f\tau} dx = \frac{4\pi}{1 + 4\pi^2 f^2}$

To calculate $R_X(\tau) = \int_{-\infty}^{\infty} S_X(f) e^{2j\pi f\tau} df$
 $= \int_{-\infty}^{\infty} e^{-|f|} e^{2j\pi f\tau} df \stackrel{f=-f}{=} \frac{4\pi}{1 + 4\pi^2 f^2}$

(b) For $S_X(f) = e^{-|f|} = \frac{4\pi}{1 + 4\pi^2 f^2}$

Then $R_X(\tau) = \int_{-\infty}^{\infty} e^{-|f|} e^{2j\pi f\tau} df = \int_{-\infty}^0 e^{(2j\pi\tau+1)f} df + \int_0^{\infty} e^{(2j\pi\tau-1)f} df$
 $= \frac{1}{2j\pi\tau+1} e^{(2j\pi\tau+1)f} \Big|_{-\infty}^0 + \frac{1}{2j\pi\tau-1} e^{(2j\pi\tau-1)f} \Big|_0^{\infty}$
 $= \frac{1}{2j\pi\tau+1} - \frac{1}{2j\pi\tau-1} = -\frac{4j\pi\tau}{4\pi^2\tau^2+1}$

P4. Solution:

$$(a) \hat{X}_t = aX_{t-1} + bX_{t-2} \quad Y_n = X_n + 2 \sum_{i=1}^{\infty} \beta^{i-1} X_{n-i} = X_n + 2X_{n-1} + 2\beta X_{n-2} + 2\beta^2 X_{n-3} \dots$$

$$Y_{n-1} = X_{n-1} + 2X_{n-2} + 2\beta X_{n-3} + 2\beta^2 X_{n-4} \dots$$

$$\text{Hence, } Y_n = X_n + \beta Y_{n-1} + (2-\beta)X_{n-1} = X_n + (2-\beta)X_{n-1} + \beta Y_{n-1}$$

$$\begin{aligned} E(Y_n Y_n) &= E[(X_n + (2-\beta)X_{n-1} + \beta Y_{n-1})(X_n + (2-\beta)X_{n-1} + \beta Y_{n-1})] \\ &= E(X_n^2) + (2-\beta)^2 E(X_{n-1}^2) + \beta^2 E(Y_{n-1}^2) + 2(2-\beta)E(X_n X_{n-1}) \\ &\quad + 2(2-\beta)\beta E(X_{n-1} Y_{n-1}) + 2\beta E(X_n Y_{n-1}) \end{aligned}$$

$$= 1 + (2-\beta)^2 + \beta^2 R_Y(0) + 2\beta(2-\beta)$$

$$\text{Then we have } (1-\beta^2) R_Y(0) = 1 + (2-\beta)^2 + 2\beta(2-\beta) = 1 + 2^2 - \beta^2$$

$$\Rightarrow C_Y(0) = R_Y(0) = \frac{1+2^2-\beta^2}{1-\beta^2}$$

$$\begin{aligned} \text{For } |k| \geq 1, \quad E(Y_{n+k} Y_n) &= E[(X_{n+k} + (2-\beta)X_{n+k-1} + \beta Y_{n+k-1}) Y_n] \\ &= E(X_{n+k} Y_n) + (2-\beta)E(X_{n+k-1} Y_n) + \beta E(Y_{n+k-1} Y_n) \end{aligned}$$

$$\text{For } |k| \geq 1, \text{ we have } C_Y(k) = \beta C_Y(k-1)$$

$$\text{For } k=1, \text{ we have } C_Y(1) = (2-\beta) + \beta C_Y(0)$$

$$\text{Hence, } C_Y(k) = \begin{cases} \frac{1+2^2-\beta^2}{1-\beta^2}, & \text{for } k=0 \\ \beta^{|k|-1} \left\{ \frac{2(1+2\beta-\beta^2)}{1-\beta^2} \right\}, & \text{for } k \neq 0 \end{cases}$$

$$\text{Set } \hat{Y}_{n+1} = \sum_{i=0}^{\infty} a_i Y_{n-i}, \text{ then we have } c_Y(k+1) = \sum_{i=0}^{\infty} a_i c_Y(k-i), \quad k \geq 0$$

$$\Rightarrow a_i = 2(\beta-2)^i, \text{ for } i \geq 0.$$

PS. Solution:

$$(a) \hat{X}_t = aX_{t_1} + bX_{t_2}$$

$$\begin{cases} E[(X_t - aX_{t_1} - bX_{t_2})X_{t_1}] = 0 \\ E[(X_t - aX_{t_1} - bX_{t_2})X_{t_2}] = 0 \end{cases} \Rightarrow \begin{cases} c(t-t_1) - a(c_0) - b(c(t_2-t_1)) = 0 \\ c(t-t_2) - a(c(t_1-t_2)) - b(c_0) = 0 \end{cases}$$

$$\Rightarrow \begin{cases} a = \frac{c_0(c(t-t_1) - c(t-t_2)(t_2-t_1))}{c_0^2 - c(t_2-t_1)^2} \\ b = \frac{c_0(c(t-t_2) - c(t-t_1)(t_2-t_1))}{c_0^2 - c(t_2-t_1)^2} \end{cases}$$

$$(b) \text{MSE} = E[(X_t - \hat{X}_t)^2] = E[(X_t - aX_{t_1} - bX_{t_2})^2]$$

$$= E(X_t^2) + a^2 E(X_{t_1}^2) + b^2 E(X_{t_2}^2) - 2bE(X_t X_{t_2}) - 2aE(X_t X_{t_1}) + 2abE(X_{t_1} X_{t_2})$$

$$= (1+a^2+b^2)c_0 - 2a(c(t-t_1)) - 2b(c(t-t_2)) + 2ab(c(t_1-t_2))$$

Then we can just substitute \hat{a} and \hat{b} from part (a) to get the result of MSE

$$(c) \text{ when } t=t_1, \text{ then } \text{MSE} = (1+a^2+b^2)c_0 - 2ac_0 + (2ab-2b)c(t_1-t_2)$$

$$\begin{cases} a = \frac{c_0^2 - c(t_1-t_2)^2}{c_0^2 - c(t_2-t_1)^2} = 1 \\ b = \frac{c_0(c(t_1-t_2) - c_0(c(t_1-t_2)))}{c_0^2 - c(t_2-t_1)^2} = 0 \end{cases} \Rightarrow \text{MSE} = (1+a^2+b^2-2a)c_0 + 2b(a-1)c(t_1-t_2) = 0$$

$$\text{when } t=t_2, \begin{cases} a = \frac{c_0(c(t_2-t_1) - c_0(c(t_2-t_1)))}{c_0^2 - c(t_2-t_1)^2} = 0 \\ b = \frac{c_0(c_0) - c(t_2-t_1)c(t_2-t_2)}{c_0^2 - c(t_2-t_1)^2} = 1 \end{cases}$$

$$\text{then } \text{MSE} = (1+a^2+b^2-2b)c_0 + 2a(b-1)c(t_1-t_2) = 0$$

The Answer is what we expect.

P6. Solution:

a) The power density function $S_x(f) = \mathcal{F}\{R_x(k)\} = \sum_{k=-\infty}^{\infty} R_x(k) e^{-j2\pi f k}$
 $\Rightarrow S_x(f) = \sum_{k=-\infty}^{\infty} 4\alpha^{|k|} e^{-j2\pi f k}$

For $|\alpha| < 1$, the sum involves an infinite geometric series and the series converges, then we have

~~$$S_x(f) = \frac{4(1-\alpha^2)e^{j2\pi f}}{2e^{j2\pi f} - 2e^{-j2\pi f} - 2 + e^{-j2\pi f}}$$~~

$$\begin{aligned} S_x(f) &= 4 \left(1 + \sum_{k=1}^{\infty} \alpha^k e^{-j2\pi f k} + \sum_{k=1}^{\infty} \alpha^k e^{j2\pi f k} \right) \\ &= 4 \left[1 + \sum_{k=1}^{\infty} (\alpha e^{-j2\pi f})^k + \sum_{k=1}^{\infty} (\alpha e^{j2\pi f})^k \right] \\ &= 4 \left(1 + \frac{\alpha e^{-j2\pi f}}{1 - \alpha e^{-j2\pi f}} + \frac{\alpha e^{j2\pi f}}{1 - \alpha e^{j2\pi f}} \right) \end{aligned}$$

~~$$(b) \text{ if } \alpha = 0.25, S_x(f) = \frac{4(1-\alpha^2)}{1-2\alpha\cos 2\pi f + \alpha^2}$$~~

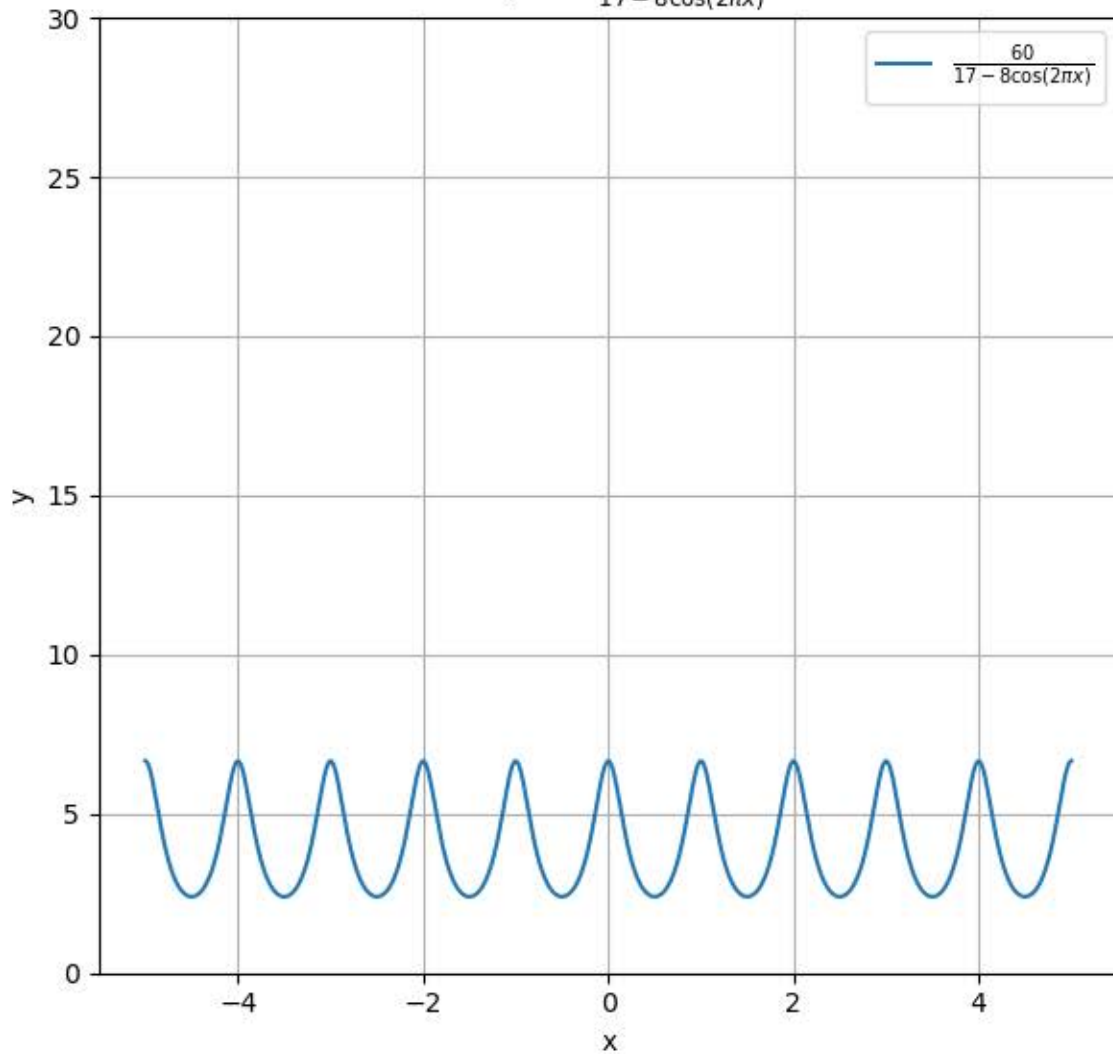
$$(b) \text{ if } \alpha = 0.25, S_x(f) = \frac{60}{17 - 8\cos 2\pi f}$$

$$\text{if } \alpha = 0.75, S_x(f) = \frac{28}{25 - 24\cos 2\pi f}$$

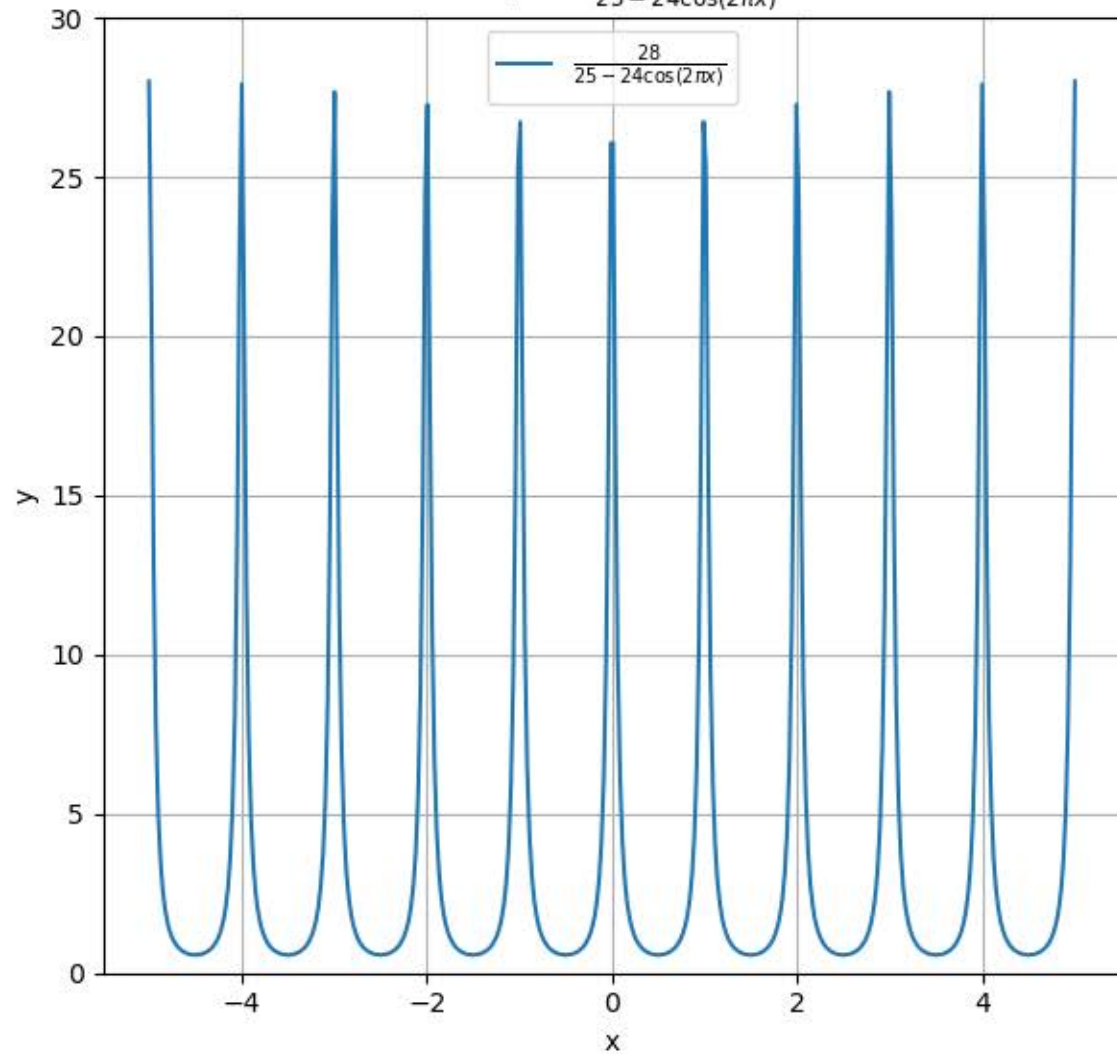
The figure is shown below.

for $|\alpha| < 1$, as α increase, ~~it effects~~ the strength of the signal increase. The figure becomes more "steep".

Graph of $\frac{60}{17 - 8\cos(2\pi x)}$



Graph of $\frac{28}{25 - 24\cos(2\pi x)}$



P7. Solution:

- (a) We can know from the problem that $R_X(k) = E(X_n X_{n-k}) = \begin{cases} 0, & \text{for } k \neq 0 \\ 1, & \text{for } k=0 \end{cases}$
Then we can get the spectral density function

$$S_X(f) = \int_{-\infty}^{\infty} R_X(k) e^{-j2\pi f k} dk = \int_{-\infty}^{\infty} 0 \cdot e^{-j2\pi f k} dk + 1 \cdot e^{-j2\pi f \cdot 0} d\omega = 1$$

- (b) Given that $Y_n = X_n + aX_{n-1}$, we can get the autocorrelation function

$$\begin{aligned} R_Y(k) &= E(Y_n Y_{n-k}) = E[(X_n + aX_{n-1})(X_{n-k} + aX_{n-k-1})] \\ &= E(X_n X_{n-k} + aX_n X_{n-k-1} + aX_{n-1} X_{n-k} + a^2 X_{n-1} X_{n-k-1}) \end{aligned}$$

Since X_n are uncorrelated, $E(X_n X_{n-a}) = 0$ for $a \neq 0$ and $E(X_n^2) = 1$

$$\text{Hence } R_Y(k) = \begin{cases} 1+a^2, & k=0 \\ a, & k=\pm 1 \\ 0, & \text{ow} \end{cases}$$

Then the spectral density function $S_Y(f)$ is the Fourier Transform of $R_Y(k)$

$$\begin{aligned} S_Y(f) &= \int_{-\infty}^{\infty} R_Y(k) e^{-j2\pi f k} dk \\ &= \int_{-\infty}^{\infty} 0 \cdot e^{-j2\pi f k} dk + (1+a^2) \cdot e^{-j2\pi f \cdot 0} d\omega + a \cdot e^{-j2\pi f \cdot 1} d\omega + a \cdot e^{j2\pi f \cdot 1} d\omega \\ &= 1+a^2 + a e^{-j2\pi f} + a e^{j2\pi f} \end{aligned}$$

(Sorry for write "2" in "a") .

P8. Solution:

$$Y_n = \alpha Y_{n-1} + X_n$$

$$\text{for } k=0, R_Y(k) = E[Y_n Y_n] = E[(\alpha Y_{n-1} + X_n)(\alpha Y_{n-1} + X_n)] \\ = E[\alpha^2 Y_{n-1}^2 + 2\alpha Y_{n-1} X_n + X_n^2] \quad (\sigma_X^2 = 1)$$

$$= \alpha^2 R_Y(0) + 1$$

$$\Rightarrow R_Y(0) = \alpha^2 R_Y(0) + 1 \Rightarrow R_Y(0) = \frac{1}{1-\alpha^2} = \frac{4}{3}$$

$$(\alpha = \frac{1}{2})$$

$$\text{For } k > 0, R_Y(k) = E[Y_{n+k} Y_n] = E[(\alpha Y_{n+k-1} + X_{n+k}) Y_n]$$

$$= \alpha E[Y_{n+k-1} Y_n] + E[X_{n+k} Y_n]$$

$$= \alpha R_Y(k-1)$$

$$\Rightarrow R_Y(k) = \alpha R_Y(k-1) \Rightarrow R_Y(k) = \frac{\alpha^{|k|}}{1-\alpha^2} = \frac{4}{3} \cdot \left(\frac{1}{2}\right)^{|k|}$$

$$S_Y(f) = \sum_{k=-\infty}^{\infty} R_Y(k) e^{-j2\pi f k} = \sum_{k=-\infty}^{\infty} \frac{4}{3} \cdot \frac{1}{2^{|k|}} \cdot e^{-j2\pi f k} = \frac{4}{3} \sum_{k=-\infty}^{\infty} \frac{1}{2^{|k|}} e^{-j2\pi f k}$$

$$= \frac{4}{3} \left(1 + \sum_{k=1}^{\infty} \frac{1}{2^k} e^{-j2\pi f k} + \sum_{k=-1}^{-\infty} \frac{1}{2^k} e^{-j2\pi f k} \right)$$

$$= \frac{4}{3} \left(1 + \sum_{k=1}^{\infty} \left(\frac{e^{-j2\pi f}}{2}\right)^k + \sum_{k=1}^{\infty} \left(\frac{e^{j2\pi f}}{2}\right)^k \right)$$

$$= \frac{4}{5 - 4\cos 2\pi f}$$

$$S_X(f) = \sigma_X^2 = 1$$

$$\text{Hence, } H(f) = \frac{S_Y(f)}{S_Y(f) + S_X(f)} = \frac{\frac{4}{5 - 4\cos 2\pi f}}{\frac{4}{5 - 4\cos 2\pi f} + 1} = \frac{4}{9 - 4\cos 2\pi f}$$