

16/02/2024 - Demostaciones

- $\text{Var}\{X\} = E\{(X - \mu_x)^2\} = E\{X^2\} - E^2\{X\}$  $= E\{X^2 - X\mu_x - X\mu_x + \mu_x^2\}$  $= E\{X^2\} - E\{2X\mu_x\} + E\{\mu_x^2\}$  $= E\{X^2\} - 2\mu_x E\{X\} + \mu_x^2$  $= E\{X^2\} - 2\mu_x^2 + \mu_x^2$  $= E\{X^2\} - \mu_x^2$  $= E\{X^2\} - (E\{X\})^2$
- $\text{Cov}\{X, Y\} = E_{x,y}\{(X - \mu_x)(Y - \mu_y)\} = E_{x,y}\{XY\} - E\{X\} E\{Y\}$  $= E_{x,y}\{(XY - X\mu_y - Y\mu_x + \mu_x\mu_y)\}$  $= E_{x,y}\{XY\} - \mu_y E\{X\} - \mu_x E\{Y\} + \mu_x\mu_y$  $= E_{x,y}\{XY\} - \mu_y\mu_x - \mu_x\mu_y + \mu_x\mu_y$  $= E_{x,y}\{XY\} - \mu_x\mu_y$  $= E_{x,y}\{XY\} - E\{X\} E\{Y\}$
- $\text{Cov}(X, Y) = E_{x,y}\{XY^T\} - E\{X\} E\{Y\}$ ,  $X, Y \in \mathbb{R}^D$   
\*  $\text{Cov}(X, X) = (X - E\{X\})(X - E\{X\})^T$ , \*  $E\{E\{X\}\} = E\{X\}$  $\text{Cov}(X, Y) = E\{(X - E\{X\})(Y^T - E\{Y\}^T)\}$  $= E\{XY^T - X E\{Y\}^T - Y^T E\{X\} + E\{X\} E\{Y\}^T\}$  $= E\{XY^T\} - E\{X\} E\{Y^T\} - E\{Y^T\} E\{X\} + E\{X\} E\{Y^T\}$  $= E\{XY^T\} - E\{X\} E\{Y^T\} - E\{Y^T\} E\{X\} + E\{X\} E\{Y^T\}$  $= E\{XY^T\} - E\{Y^T\} E\{X\}$

19/02/2024

## - Demostraciones

$$\bullet E\{X\} = \int_{-\infty}^{\infty} N(X|\mu, \sigma^2) X dx = \mu$$

$$= \int_{-\infty}^{\infty} X \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{|x-\mu|^2}{2\sigma^2}\right) dx$$

\* Quitar traslaciones + Sustitución  $Z = X - \mu$ ,  $x = z + \mu$   
 $dz = dx$

$$= \int_{-\infty}^{\infty} z + \mu \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{|z|^2}{2\sigma^2}\right) dz$$

$$= \int_{-\infty}^{\infty} z \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{|z|^2}{2\sigma^2}\right) dz + \int_{-\infty}^{\infty} \mu \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{|z|^2}{2\sigma^2}\right) dz$$

O Por simetría

$$= \mu \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{|z|^2}{2\sigma^2}\right) dz = \mu$$

Función de densidad

$$\bullet E\{X^2\} = \int_{-\infty}^{\infty} N(X|\mu, \sigma^2) X^2 dx = \mu^2 + \sigma^2$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{|x-\mu|^2}{2\sigma^2}\right) X^2 dx$$

\* Cambio de variable  $Z = \frac{x-\mu}{\sigma}$ ;  $dx = \sqrt{2\pi\sigma^2} dz$ ,  $x = \mu + \sqrt{2\sigma^2} z$

$$dx = \sqrt{2\sigma^2} dz; \quad x^2 = \mu^2 + 2\sigma^2 z^2 + 2\mu\sqrt{2\sigma^2} z$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi}} e^{-z^2} (\mu^2 + 2\sigma^2 z^2 + 2\mu\sqrt{2\sigma^2} z) \sqrt{2\sigma^2} dz$$

\* Se expande la integral en 3 integrales

$$= \mu^2 \int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi}} e^{-z^2} dz + 2\sigma^2 \int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi}} e^{-z^2} z^2 dz$$

Probabilidad

$$+ 2\mu\sqrt{2\sigma^2} \int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi}} e^{-z^2} z dz$$

$\Rightarrow 0$  ∵ Simetría

$$= \mu^2 + 2\sigma^2 \int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi}} e^{-z^2} z^2 dz$$

$$= \mu^2 + 2\sigma^2 \frac{1}{\sqrt{\pi}} \left( \frac{\sqrt{\pi}}{8} \right) = \mu^2 + \sigma^2 //$$

•  $\text{Var}[X] = \sigma^2$

$$= E[X^2] - E^2[X]; E[X^2] = \mu^2 + \sigma^2; E[X] = \mu$$

$$= \mu^2 + \sigma^2 - \mu^2$$

$$= \sigma^2 //$$

21/02/24 - Demostaciones

•  $\sigma_{ML}^2 = \frac{1}{N} \sum_{n=1}^N (X_n - \bar{X}_{ML})^2$

\* Se calcula el logaritmo de la función para simplificar cálculos.

$$\begin{aligned} \log(\sigma_{ML}^2) &= \log \left( \prod_{n=1}^N \frac{1}{\sqrt{2\pi\sigma_{ML}^2}} \exp \left( -\frac{(X_n - \bar{X}_{ML})^2}{2\sigma^2} \right) \right) \\ &= \log \left( \prod_{n=1}^N \frac{1}{\sqrt{2\pi\sigma_{ML}^2}} \right) + \log \left( \prod_{n=1}^N \exp \left( -\frac{(X_n - \bar{X}_{ML})^2}{2\sigma^2} \right) \right) \\ &= \log \left( \prod_{n=1}^N \frac{1}{(2\pi\sigma_{ML}^2)^{\frac{1}{2}}} \right) + \log \left( \exp \left( -\sum_{n=1}^N \frac{(X_n - \bar{X}_{ML})^2}{2\sigma^2} \right) \right) \\ &= -\frac{N}{2} \log(2\pi\sigma_{ML}^2) - \sum_{n=1}^N \frac{(X_n - \bar{X}_{ML})^2}{2\sigma^2} \end{aligned}$$

$$\log(\sigma_{ML}^2) = -\frac{N}{2} \log(2\pi) - \frac{N}{2} \log(\sigma_{ML}^2) - \frac{1}{2\sigma^2} \sum_{n=1}^N |X_n - \bar{X}_{ML}|^2$$

\* Se deriva respecto a  $\sigma^2$  y se iguala a cero

$$\begin{aligned} \frac{d}{d(\sigma^2)} (\log(\sigma_{ML}^2)) &= \frac{d}{d(\sigma^2)} \left( -\frac{N}{2} \log(2\pi) \right) \frac{d}{d(\sigma^2)} \left( -\frac{N}{2} \log(\sigma_{ML}^2) \right) \frac{d}{d(\sigma^2)} \left( -\frac{1}{2\sigma^2} \sum_{n=1}^N |X_n - \bar{X}_{ML}|^2 \right) \\ &= -\frac{N}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{n=1}^N (X_n - \bar{X}_{ML})^2 = 0 \end{aligned}$$

\* Se multiplica a ambos lados por  $(\sigma^2)^2$

$$= -\frac{N}{2\sigma^2} (\bar{x}^2) + \frac{1}{2(\sigma^2)^2} \sum_{n=1}^N (x_n - \mu_{ML})^2 = o(\sigma^2)^2$$

$$= -\frac{N}{2} \bar{x}_{ML}^2 + \frac{1}{2} \sum_{n=1}^N (x_n - \mu_{ML})^2 = 0$$

\* Se despeja  $\sigma_{ML}^2$

$$\sigma_{ML}^2 = \left(\frac{2}{N}\right) \frac{1}{2} \sum_{n=1}^N (x_n - \mu_{ML})^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \mu_{ML})^2 //$$

• Ejercicio: determinar el estimador con más información

$$x[n] = A + w_n; A \in \mathbb{R}^+ \text{ y } w_n \sim N(w_n | \mu, \sigma^2)$$

Sea:  $\hat{A}_1 = \frac{1}{N} \sum_{n=1}^N x_n$  Estimador 1;  $\hat{A}_2 = x_1$  Estimador 2

\* Se evalua el Bias

$$\text{Para } \hat{A}_1, b(\hat{A}_1) = E\{\hat{A}_1\} - A$$

$$= E\left\{ \frac{1}{N} \sum_{n=1}^N x[n] \right\} - A$$

$$= \frac{1}{N} \sum_{n=1}^N E\{A + w[n]\}$$

$$= E\{A\} + E\{w[n]\}$$

$$= \frac{1}{N} \sum_{n=1}^N A - A$$

$$= \frac{N}{N} A - A$$

$$\text{Para } \hat{A}_2$$

$$b(\hat{A}_2) = E\{\hat{A}_2\} - A$$

$$= E\{x[1]\} - A$$

$$= E\{A + w[1]\} - A$$

$$= A + E\{w[1]\} - A$$

$$= A - A$$

$$b(\hat{A}_2) = 0.$$

$$b(\hat{A}_1) = 0$$

\* No es posible concluir a partir del sesgo del estimador, por lo tanto se calcula la varianza

$$\text{Para } \hat{A}_1, \text{Var}\{\hat{A}_1\} = \text{Var}\left\{ \frac{1}{N} \sum_{n=1}^N x[n] \right\}$$

$$= \text{Var}\left\{ \frac{1}{N} \sum_{n=1}^N A + w[n] \right\}$$

$$= \text{Var}\left\{ \frac{1}{N} \sum_{n=1}^N A + \frac{1}{N} \sum_{n=1}^N w[n] \right\}$$

$$\begin{aligned}
 \text{Var}\left\{\frac{1}{N} \sum_{n=1}^N w[n]\right\} &= \text{Var}\left\{\frac{1}{N} w[1] + \dots + \frac{1}{N} w[N]\right\} \\
 &= \sum_{n=1}^N \text{Var}\left\{\frac{1}{N} w[n]\right\} \\
 &= \sum_{n=1}^N \frac{1}{N^2} \sigma^2 = \frac{N}{N^2} \sigma^2 = \frac{1}{N} \sigma^2
 \end{aligned}$$

$$\text{Var}\{A\} = \frac{\sigma^2}{N}$$

Para  $A_2$

$$\begin{aligned}
 \text{Var}\{A_2\} &= \text{Var}\{x[i]\} \\
 &= \text{Var}\{A + w[i]\} \\
 &= \text{Var}\{w[i]\} \\
 \text{Var}\{A_2\} &= \sigma^2
 \end{aligned}$$

El estimador  $A_1$  contiene mas información //

Ejercicio: Demuestre que:

$$E\{\mu_{mc}\} = \mu \quad \text{Con } X_n \sim N(X_n | \mu, \sigma^2)$$

$$E\{\sigma_{mc}^2\} = \left(\frac{N-1}{N}\right) \sigma^2$$

$$\begin{aligned}
 E\left\{\frac{1}{N} \sum_n (X_n - \mu_{mc})^2\right\} &= E\left\{\frac{1}{N} \sum_n (X_n^2 - 2X_n \mu_{mc} + \mu_{mc}^2)\right\} \\
 &= E\left\{\frac{1}{N} \sum_n X_n^2 - 2\mu_{mc} \frac{1}{N} \sum_n X_n + \frac{1}{N} \sum_n \mu_{mc}^2\right\} \\
 &= E\left\{\frac{1}{N} \sum_n X_n^2\right\} - 2 E\{\mu_{mc} \mu_{mc}\} + E\left\{\frac{1}{N} \sum_n \mu_{mc}^2\right\} \\
 &= \frac{1}{N} \sum_n E\{X_n^2\} - 2 E\{\mu_{mc}^2\} + E\{\mu_{mc}^2\}
 \end{aligned}$$

$$E\left\{\frac{1}{N} \sum_n (X_n - \mu)^2\right\} = \frac{1}{N} N E\{X_n^2\} - E\{\mu_{mc}^2\} = E\{X_n^2\} - E\{\mu_{mc}^2\}$$

$$\bullet \text{Var}\{X\} = E\{X^2\} - E\{X\}^2 \Rightarrow E\{X_n^2\} = \text{Var}\{X_n\} + E\{\mu_{m_n}^2\}$$

$$\bullet X_n \sim N(X_n | \mu, \sigma^2); E\{X_n^2\} = \sigma^2 + \mu^2$$

$$E\{X_{m_n}^2\} = \text{Var}\{M_{m_n}\} + E\{\mu_{m_n}\}^2$$

$$E\{\mu_{m_n}^2\} = \frac{\sigma^2}{N} + \mu^2$$

$$\bullet E\{\mu_{m_n}\} = E\left\{\frac{1}{N} \sum_n X_n\right\}$$

$$= \frac{1}{N} \sum_n E\{X_n\} = \frac{1}{N} \sum_n \mu = \frac{1}{N} NM$$

$$E\{\mu_{m_n}\} = \mu //$$

$$\bullet \text{Var}\{\mu_{m_n}\} = \text{Var}\left\{\frac{1}{N} \sum_n X_n\right\} = \frac{1}{N^2} \sum_n \text{Var}\{X_n\}$$

$$= \frac{1}{N^2} N \sigma^2 = \frac{\sigma^2}{N}$$

$$\bullet E\left\{\frac{1}{N} \sum_n (X_n - \mu_{m_n})^2\right\} = E\{X_n^2\} - E\{\mu_{m_n}^2\}$$

$$= \sigma^2 + \mu^2 - \left(\frac{\sigma^2}{N} + \mu^2\right)$$

$$= \sigma^2 - \frac{\sigma^2}{N} + \mu^2 - \mu^2$$

$$= \sigma^2 \left(1 - \frac{1}{N}\right)$$

$$E\left\{\frac{1}{N} \sum_n (X_n - \mu_{m_n})^2\right\} = \sigma^2 \frac{(N-1)}{N}$$

$$\sigma_{m_n}^2 = \sigma^2 \frac{(N-1)}{N} //$$

¿Cuál debiera ser la corrección sobre  $\sigma_{m_n}^2$  para evitar el Bias?

$$\frac{N}{N-1} \sigma_{m_n}^2 = \sigma^2$$

$$\frac{N}{N-1} \frac{1}{N} \sum_n (X_n - \mu_{m_n})^2 = \sigma^2$$

$$\frac{1}{N-1} \sum_n (X_n - \mu_{m_n})^2 = \sigma^2 //$$

26/02/24 -

Demostación

$$S^2 = I$$

\* Eigenvalores:  $X^T X = \Delta = V \Delta V$

\* Descomposición por valor Singular:  $X = U S V^*$

$$(U S V^*)^T (U S V^*)^T = V \Delta V^*$$

$$(V^*) (V S)^T U S V^* =$$

$$V S^* U^* U S V^* = U^* U = I$$

$$V S^* S V^* = V \Delta V^*$$

$$V S^2 V^* = V \Delta V^*$$

28/02/24 - Ejercicio

$$\text{ecmr}(Y, \phi w) = E\{||Y - \phi w||_2^2\} + \lambda ||w||_2^2$$

$$||Y - \phi w||_2^2 = [Y^T Y - 2Y^T \phi w + w^T \phi^T \phi w] \frac{1}{N}$$

$$* \langle Y - \phi w, Y - \phi w \rangle; \lambda ||w||_2^2 = \lambda \langle w, w \rangle = \lambda = w^T w$$

$$\frac{\partial}{\partial w} ||Y - \phi w||_2^2 = [-2\phi^T Y + 2\phi^T \phi w] \frac{1}{N} \in \mathbb{R}^Q$$

$$* \sum_{\partial w} ||w||_2^2 = 2\lambda w$$

$$[-\frac{2}{N} \phi^T Y + \frac{2}{N} \phi^T \phi w + 2\lambda w] \frac{N}{2} = 0 \left[ \frac{N}{2} \right]$$

$$\phi^T Y + \phi^T \phi w + 2\lambda w = 0$$

$$(\phi^T \phi + 2\lambda I)^{-1} (\phi^T \phi + 2\lambda I) w = (\phi^T \phi + 2\lambda I)^{-1} \phi^T Y$$

$$\hat{w} = (\phi^T \phi + 2\lambda I)^{-1} \phi^T Y$$

01/03/24 - Ejercicio

Asumiendo conjuntos datos i.i.d

$D = \{X_n \in \mathbb{R}^p, Y_n \in \mathbb{R}\}_{n=1}^N$  Encontrar  $\mathbf{W}$  y  $\sigma^2$  que

Maximizar  $J_a = \log$  verosimilitud  $* F(\mathbf{x} | \mathbf{w}) = \phi(\mathbf{w})$

$$\frac{d}{d\mathbf{w}} \log (P(Y) f(\mathbf{x} | \mathbf{w}, \sigma^2)) \in \mathbb{R}^q$$

$$= \log \left( \prod_{n=1}^N \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left( -\frac{\|Y_n - f(\mathbf{x}_n | \mathbf{w})\|^2}{2\sigma^2} \right) \right)$$

$$= \log \left( \prod_{n=1}^N \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left( -\frac{\|Y_n - \phi(\mathbf{w})\|^2}{2\sigma^2} \right) \right)$$

$$= \log \left( \prod_{n=1}^N \frac{1}{\sqrt{2\pi\sigma^2}} \right) + \log \left( \prod_{n=1}^N \exp \left( -\frac{\|\mathbf{x}_n - \phi(\mathbf{w})\|^2}{2\sigma^2} \right) \right)$$

$$= \log \left( \frac{1}{(2\pi\sigma^2)^{N/2}} \right) + \log \left( \exp \left( -\sum_{n=1}^N \frac{\|\mathbf{x}_n - \phi(\mathbf{w})\|^2}{2\sigma^2} \right) \right)$$

$$= \log \left( (2\pi\sigma^2)^{-\frac{N}{2}} \right) - \sum_{n=1}^N \frac{\|\mathbf{x}_n - \phi(\mathbf{w})\|^2}{2\sigma^2}$$

$$\stackrel{!}{=} -\frac{N}{2} \log (2\pi\sigma^2) - \frac{1}{2\sigma^2} \|\mathbf{y} - \phi(\mathbf{w})\|_2^2$$

$$= -\frac{N}{2} \{ \log (2\pi) + \log (\sigma^2) \} - \frac{1}{2\sigma^2} \|\mathbf{y} - \phi(\mathbf{w})\|_2^2$$

$$= \frac{d}{d\mathbf{w}} \left[ -\frac{N}{2} (\log (2\pi) + \log (\sigma^2)) \right] - \frac{d}{d\mathbf{w}} \left[ \frac{1}{2\sigma^2} \|\mathbf{y} - \phi(\mathbf{w})\|_2^2 \right]$$

$$= 0 - \frac{1}{2\sigma^2} [-2\phi^T \mathbf{y} + 2\phi^T \phi(\mathbf{w})]$$

$$= -\frac{1}{2\sigma^2} [-2\phi^T \mathbf{y} + 2\phi^T \phi(\mathbf{w})]$$

$$\hat{\mathbf{w}}_{ML} (= (\phi^T \phi)^{-1} \phi^T \mathbf{y})$$

\* Para  $\sigma^2$  se deriva respecto a  $\sigma^2$

$$\begin{aligned} & \frac{d}{d\sigma_n^2} \left[ -\frac{N}{2} (\log(2\pi) + \log(\sigma_n^2)) \right] - \frac{d}{d\sigma_n^2} \left[ \frac{1}{2\sigma_n^2} \|Y - \Phi w\|_2^2 \right] \\ &= \cancel{\frac{d}{d\sigma_n^2} \left[ -\frac{N}{2} \log(2\pi) \right]} + \frac{d}{d\sigma_n^2} \left[ \frac{N}{2} \log(\sigma_n^2) \right] - \frac{d}{d\sigma_n^2} \left[ \frac{1}{2\sigma_n^2} \|Y - \Phi w\|_2^2 \right] \\ &= -\frac{N}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \|Y - \Phi w\|_2^2 = 0 \end{aligned}$$

\* Se multiplica por  $\sigma^2$  y se despeja

$$-\frac{N}{2\sigma^2} (\sigma^2)^2 + \frac{1}{2(\sigma^2)^2} \|Y - \Phi w\|_2^2 = 0 \quad (\sigma^2)^2$$

$$-\frac{N}{2} \sigma_n^2 + \frac{1}{2} [-2\Phi^T Y + 2\Phi^T \Phi w] = 0$$

$$\sigma_n^2 = \frac{1}{N} [-2\Phi^T Y + 2\Phi^T \Phi w]$$

### Ejercicios

Encontrar los  $w$  que maximizan el log-MAP

$$\log(P(w | Y, \Phi, \sigma_n^2)) = \log \left( \prod_{n=1}^N (Y_n | \Phi(X_n)w, \sigma_n^2) \right) \prod_{q=1}^Q N(w_q | 0, \sigma_w^2)$$

$$= \left[ -\frac{N}{2} [\log(2\pi) + \log(\sigma_n^2)] - \frac{1}{2\sigma_n^2} \|Y - \Phi w\|_2^2 \right] A$$

$$+ \log \left( \prod_{q=1}^Q \left( \frac{1}{\sqrt{2\pi\sigma_w^2}} \exp \left( -\frac{\|w_q - 0\|_2^2}{2\sigma_w^2} \right) \right) \right)$$

$$= A + \log \left( \prod_{q=1}^Q \frac{1}{(2\pi\sigma_w^2)^{1/2}} \right) + \log \left( \prod_{q=1}^Q \exp \left( -\frac{\|w_q\|_2^2}{2\sigma_w^2} \right) \right)$$

$$= A + \log \left( \frac{1}{(2\pi\sigma_w^2)^{Q/2}} \right) + \log \left( \exp \left( -\sum_{q=1}^Q \frac{\|w_q\|_2^2}{2\sigma_w^2} \right) \right)$$

$$= A - \frac{Q}{2} [\log(2\pi) + \log(\sigma_w^2)] - \sum_{q=1}^Q \|w_q\|_2^2$$

$$* \|w\|_2^2 = \left( \sqrt{\sum_{q=1}^Q \|w_q\|_2^2} \right)^2 = \sum_{q=1}^Q \|w_q\|_2^2$$

$$\begin{aligned}
&= -\frac{N}{2} \left[ \log(2\pi) + \log(\sigma_p^2) \right] - \frac{1}{2\sigma_n^2} \| \mathbf{y} - \Phi \mathbf{w} \|_2^2 - \frac{\alpha}{2} \left[ \log(2\pi) + \log(\sigma_w^2) \right] \\
&\quad - \frac{1}{2\sigma_w^2} \| \mathbf{w} \|_2^2 \\
&= \text{Max} - \left[ \frac{2\sigma_n^2}{2\sigma_n^2} \| \mathbf{y} - \Phi \mathbf{w} \|_2^2 + \frac{2\sigma_w^2}{2\sigma_w^2} \| \mathbf{w} \|_2^2 \right] + c + e \\
&= \text{Min} \left[ \| \mathbf{y} - \Phi \mathbf{w} \|_2^2 + \frac{\sigma_n^2}{\sigma_w^2} \| \mathbf{w} \|_2^2 + c + e \right] \\
&= \hat{\mathbf{w}}_{MAP} = (\Phi^\top \Phi + \frac{\sigma_n^2}{\sigma_w^2} \mathbb{I})^{-1} \Phi^\top \mathbf{y} \rightarrow \hat{\mathbf{w}}_{MCRU} // 
\end{aligned}$$

\* MAP tiene la misma estructura que MCRU, En donde

$$\lambda = \frac{\sigma_n^2}{\sigma_w^2} //$$

Expresa la Solución de MAP mediante descomposición respectiva

$$\bullet \hat{\mathbf{w}}_{MAP} = (\Phi^\top \Phi + \frac{\sigma_n^2}{\sigma_w^2} \mathbb{I})^{-1} \Phi^\top \mathbf{y}$$

$$* \mathbf{X} = \mathbf{U} \mathbf{S} \mathbf{V}^* \quad \mathbf{X} = \Phi$$

$$((\mathbf{S} \mathbf{V}^*)^\top \mathbf{U}^\top) (\mathbf{U} \mathbf{S} \mathbf{V}^\top + \frac{\sigma_n^2}{\sigma_w^2} \mathbb{I})^{-1} (\mathbf{S} \mathbf{V}^*)^\top \mathbf{U}^\top \mathbf{y}$$

$$((\mathbf{V} \mathbf{S}^2 \mathbf{U}^\top) (\mathbf{U} \mathbf{S} \mathbf{V}^\top) + \frac{\sigma_n^2}{\sigma_w^2} \mathbb{I})^{-1} ((\mathbf{V} \mathbf{S}^2) \mathbf{U}^\top \mathbf{y}$$

$$(\mathbf{V} \mathbf{S}^2 \mathbf{V}^\top) + \frac{\sigma_n^2}{\sigma_w^2} \mathbb{I}^{-1} (\mathbf{V} \mathbf{S}^2 \mathbf{U}^\top \mathbf{y})$$