Week 3 HW HyeondoOh

Week 3 HW States

Date

1 BDA Ch2 ex8

Normal distribution with unknown mean: a random sample of n students is drawn from a large population, and their weights are measured. The average weight of the n sampled students is $\overline{y} = 150$ pounds. Assume the weights in the population are normally distributed with unknown mean θ and known standard deviation 20 pounds. Suppose your prior distribution for θ is normal with mean 180 and standard deviation 40.

- (a) Give your posterior distribution for θ. (Your answer will be a function of n.)
- (b) A new student is sampled at random from the same population and has a weight \(\tilde{y}\) pounds. Give a posterior predictive distribution for \(\tilde{y}\). (Your answer will still be function of \(n.\))
- (c) For n=10, give a 95% posterior interval for θ and a 95% posterior predictive
- (d) Do the same for n = 100

11 Sample -> 9= 150

110 ~ N(0,20) pri 0~N(180 40)

$$\mathcal{M} : \frac{1}{2^{n}} + \frac{1}{2^{n}} \mathcal{M}_{0} + \frac{1}{2^{n}} + \frac{1}{2^{n}} \mathcal{J}$$

$$Z_n^{\lambda} = \frac{1}{Z_0^{\lambda} + \frac{n}{G^{\lambda}}} = \frac{1}{\frac{1}{1600} + \frac{n}{400}} = \frac{1600}{400}$$

b) posterior prodictive dice for g

E(y | y) = Mr = 6000 + 180

posteria interval for 0, 9 c) n=(0

9) N=100

) Two-par-Normal data with consulte prior dist ONAN of marginal posterior dist" 95

$$\frac{\text{Hike lihrod}}{\text{p(y| M. o')}} = \frac{\pi}{14} \left(\frac{1}{2\pi o'} e^{-\frac{(y_1 - M)^2}{2 o'}} \right)$$

$$\propto \frac{1}{\sqrt{1}} \exp \left\{ -\frac{\Xi(y_1 - M)^2}{2 o'} \right\}$$

P(M. 5) = p(M(5) P(0)

* P(o=14) marginal post dist

Ploy = P(0) P(40) = P(0) \ P(4100) P(M10) dM

$$\frac{\left(\frac{1}{\sigma^{2}}\right)^{\frac{1}{2}}}{\left(\frac{1}{\sigma^{2}}\right)^{\frac{1}{2}}} \exp\left\{-\frac{1}{2\sigma^{2}}\left(\frac{1}{\sigma^{2}}\right)^{\frac{1}{2}}\right\} \exp\left\{-\frac{1}{2\sigma^{2}}\right\} \exp\left\{-\frac{1}{2\sigma^{2}}\right\} \exp\left\{-\frac{1}{2\sigma^{2}}\right\} \exp\left\{-\frac{1}{2\sigma^{2}}\right\} \exp\left\{-\frac{1}{2\sigma^{2}}\right\} \exp\left\{-\frac{1}{2\sigma^{2}}\right\} \exp\left\{-\frac{1}{2\sigma^{2}}\right\} \exp\left\{-\frac{$$

 $\alpha\left(\frac{1}{L}\right)^{\frac{1}{2}+1} \cdot \exp\left[-\frac{\gamma \alpha_{s}}{2\alpha_{s}}\right] \cdot \left(\frac{\alpha_{s}}{2}\right)^{\frac{1}{2}} \cdot \exp\left(-\frac{\gamma \alpha_{s}}{2\alpha_{s}}(1-1)\zeta_{s}\right) \cdot \left(\frac{\alpha_{s}}{2}\right)^{\frac{1}{2}}$

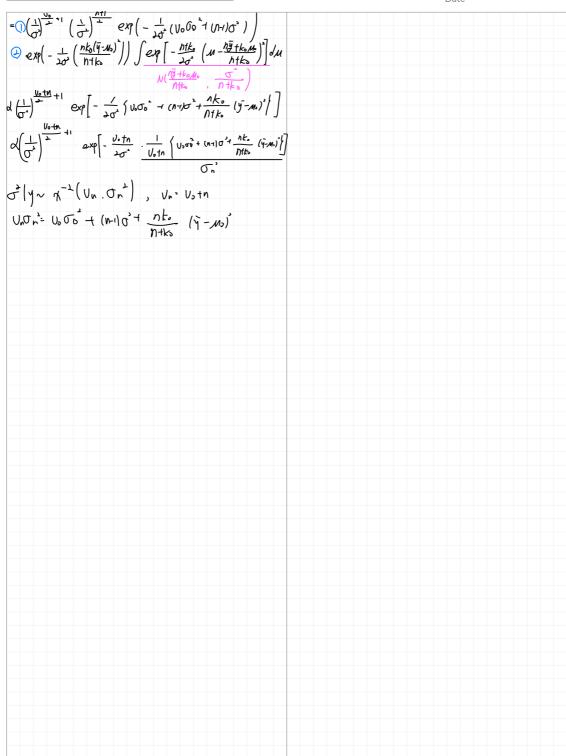
= [ext -] (ntko)M - 2(ny+ kom) M + ko M + ny /] dm

$$= \int \exp \left[-\frac{n + k_0}{2\sigma^2} \left[M^2 - 2 \frac{n \overline{y} + k_0 M_0}{n + k_0} M + \left(\frac{n \overline{y} + k_0 M_0}{n + k_0} \right)^2 \right] dM$$

$$= \frac{1}{2\sigma^2} \left[k_0 M_0^2 + n \overline{y}^2 - \frac{(n \overline{y} + k_0 M_0)^2}{n + k_0} \right] dM$$

= exp = 10 (rkd y-m)) PHR = ntto (n-ny+tom)] dm

Date

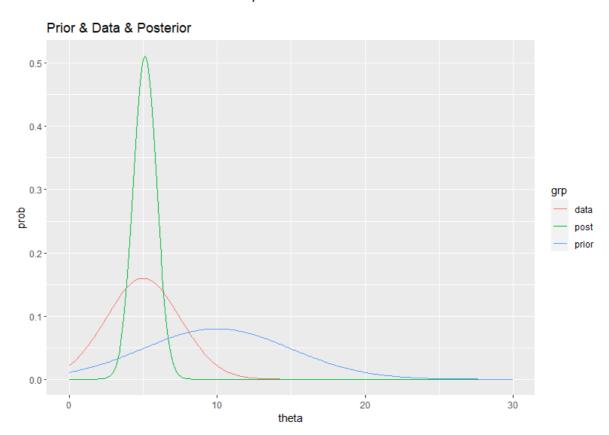


3. R코드

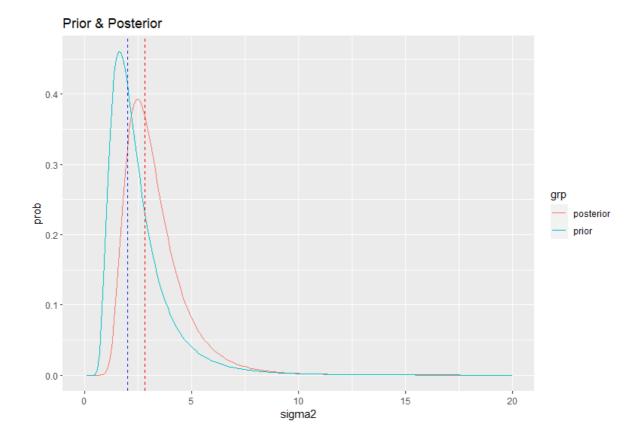
One par model

```
install.packages('ggplot2')
install.packages('tidyr')
install.packages("ggpubr")
library(ggplot2)
library(ggpubr)
library(tidyr)
## One parameter model
## Normal model with unknown mu
## prior
mu_0 = 10
tau_0 = 5
## data
mu = 5
sd = 2.5
n = 10
## posterior(parameter update)
mu_n = ((1/tau_0^2)/(1/tau_0^2+n/sd^2))*mu_0+
  (n/sd^2/(1/tau_0^2+n/sd^2))*mu
tau_n = sqrt(1/(1/tau_0^2+n/sd^2))
title = "Prior & Data & Posterior"
theta = seq(0,30,0.1)
p = data.frame(theta = theta,
               prior = dnorm(theta, mu_0, tau_0),
               post = dnorm(theta, mu_n, tau_n),
               data = dnorm(theta, mu, sd)
)%>% gather(grp, prob, -theta) %>%
  ggplot(aes(x=theta, y=prob, color=grp))+geom_line()+labs(title=title)
ggarrange(p)
## Normal model with unknown sigma
# prior
sigma_0 = 2
nu_0 = 9
# data1
data = rnorm(5, 7, 3)
mu = mean(data)
sigma = var(data)
n = length(data)
# posterior
nu_n = nu_0 + n
sigma_n = (nu_0*sigma_0^2+sum((data-mu)^2))/nu_n
```

Normal model with unknown μ



Normal model with unknown σ^2



Two par model

```
## Two parameter model
# data
D = c(1.64, 1.70, 1.72, 1.74, 1.82, 1.82, 1.82, 1.90, 2.08)
n = length(D); xbar = mean(D); s2 = var(D)
# prior
mu0 = 1.9; kappa0 = 1; s20 = 0.01; nu0 = 1
# posterior
kappa1 = kappa0 + n
nu1 = nu0 + n
mu1 = (kappa0 * mu0 + n * xbar) / kappa1
s21 = (1/ nu1) * (nu0*s20 + (n-1)*s2 + (kappa0*n/kappa1)*(xbar-mu0)^2)
# visualize
prior = function(theta, sigma2)
  dnorm(theta, mu0, sqrt(sigma2/kappa0)) * dsinvchisq(sigma2, nu0, s20)
posterior = function(theta, sigma2)
  dnorm(theta, mu1, sqrt(sigma2/kappa1)) * dsinvchisq(sigma2, nu1, s21)
dsinvchisq = function(theta,v,tau2)
  ((v*tau2)^{(v/2)})/gamma(v/2)*(1/theta)^{(v/2+1)*exp(-v*tau2/(2*theta))}
mu = seq(1.6, 2, length.out = 100)
sigma2 = seq(0.001, 0.04, length.out = 100)
cmu = rep(mu, each = length(sigma2))
csigma2 = rep(sigma2, length(mu))
prr_dens = mapply(prior, cmu, csigma2)
post_dens = mapply(posterior, cmu, csigma2)
title1 = "Joint prior"
```

```
p1 = data.frame(mu = cmu, sigma2 = csigma2, dens = prr_dens) %>%
  ggplot(aes(x=cmu, y=sigma2))+
  geom_raster(aes(fill = dens, alpha= dens), interpolate= T)+
  geom_contour(aes(z= dens), color = 'black', size= 0.2)+
  scale_fill_gradient(low= "cornflowerblue", high= "cornflowerblue", guide= F)+
  scale_alpha(range=c(0,1), guide=F)+
  labs(title=title1)
title2 = "Joint posterior"
p2 = data.frame(mu = cmu, sigma2 = csigma2, dens = post_dens) %>%
  ggplot(aes(x=cmu, y=sigma2))+
  geom_raster(aes(fill = dens, alpha= dens), interpolate= T)+
  geom_contour(aes(z= dens), color = 'black', size= 0.2)+
  scale_fill_gradient(low= "cornflowerblue", high= "cornflowerblue", guide= F)+
  scale_alpha(range=c(0,1), guide=F)+
  labs(title=title2)
ggarrange(p1, p2)
```

