

2번문제

1. likelihood  $x_i | \theta \sim B(n, \theta)$

$$\begin{aligned}
 P(x_i | \theta) &= \prod_{i=1}^n n C_{x_i} \cdot (\theta)^{x_i} (1-\theta)^{n-x_i} \\
 &= \prod_{i=1}^n n C_{x_i} \times \theta^{\sum x_i} \cdot (1-\theta)^{n^2 - \sum x_i} \\
 &= \prod_{i=1}^n n C_{x_i} \times \exp\left(\sum x_i \cdot \log \theta + (n^2 - \sum x_i) \cdot \log(1-\theta)\right) \\
 &= \underbrace{\prod_{i=1}^n n C_{x_i}}_{h(x)} \times \underbrace{\exp\left(n^2 \log(1-\theta)\right)}_{g(\theta)} \times \underbrace{\exp\left(\sum x_i \cdot \log\left(\frac{\theta}{1-\theta}\right)\right)}_{\eta(x) \kappa(\theta)} \\
 &\rightarrow \text{지속족의 형태 만족.}
 \end{aligned}$$

$$P(x_i | \theta) \propto \theta^{\sum x_i} \cdot (1-\theta)^{n^2 - \sum x_i}$$

2. prior  $\theta \sim \text{Beta}(a, b)$

$$P(\theta) \propto \theta^{a-1} (1-\theta)^{b-1}$$

3. posterior

$$P(\theta | x) \propto \theta^{\sum x_i + a - 1} \cdot (1-\theta)^{n^2 - \sum x_i + b - 1}$$

$$\therefore \theta | x \sim \text{Beta}(\sum x_i + a, n^2 - \sum x_i + b)$$

3번문제

$$X \sim \text{NB}(r, p)$$

$$r \rightarrow \infty, p \rightarrow 1, r(1-p) \rightarrow \lambda \quad (\text{condition})$$

$$h(x) = \binom{r-1+x}{r-1} (1-p)^r p^x$$

$$\begin{aligned}
P(x) &= \binom{r+x-1}{x} (1-p)^r p^x \\
&= \frac{(r-1+x)!}{x! (r-1)!} \times \left( \frac{r}{r+x} \right)^r \times \left( \frac{x}{r+x} \right)^x \\
&= \frac{(r-1+x)!}{x! (r-1)!} \times e^{-\lambda} \times \frac{\lambda^x}{(r+x)^x} \\
&= e^{-\lambda} \times \frac{\lambda^x}{x!}
\end{aligned}$$


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variance of NB(r, p)

$$= \frac{rp}{(1-p)^2}$$

variance of pois( $\lambda$ )

$$= \lambda = \frac{rp}{1-p}$$

→ 음이항 분포의 분포가 더 큰 것을 볼 수 있다.

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$Y \sim B(n, p)$

$$\begin{aligned}
P(y) &= \binom{n}{y} \cdot p^y \cdot (1-p)^{n-y} \\
&= \frac{n!}{y! (n-y)!} \cdot p^y \cdot (1-p)^{n-y} \\
&= \frac{n!}{y! (n-y)!} \cdot \left( \frac{\lambda}{n} \right)^y \cdot \left( \frac{n-\lambda}{n} \right)^{n-y} \\
&= \frac{\lambda^y}{y!} \left( 1 - \frac{y}{n} \right)^{n-y} \\
&= \frac{\lambda^y}{y!} \cdot e^{-n \cdot \frac{y}{n}} = \frac{\lambda^y}{y!} \cdot e^{-\lambda}
\end{aligned}$$



