9. Generative Classifiers STA3142 Statistical Machine Learning

Kibok Lee

Assistant Professor of Applied Statistics / Statistics and Data Science Mar {26, 28}, 2024

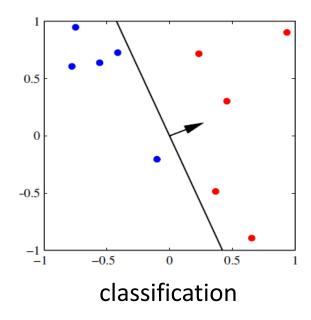


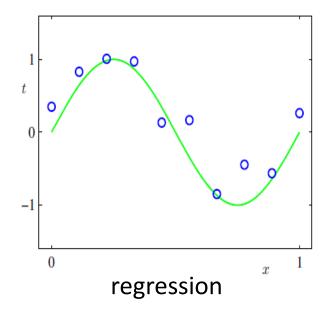
Assignment 1

- Due Friday 3/29, 11:59pm
- Topics
 - (Programming) NumPy basics
 - (Programming) Linear regression on a polynomial
 - (Math) Derivation and proof for linear regression
- Please read the instruction carefully!
 - Submit one <u>pdf</u> and one <u>zip</u> file separately
 - Write your code only in the designated spaces
 - Do not import additional libraries
 - ...
- If you feel difficult, consider to take option 2.

Recap: Supervised Learning

- Learning a function $h: \mathcal{X} \to \mathcal{Y}$
- Labels could be discrete or continuous
 - Discrete labels: classification
 - Continuous labels: regression





Classification Strategies

- Learning the distributions $p(C_k|x)$
 - Discriminative models: Directly model $p(C_k|x)$ and learn parameters from the training set.
 - Generative models: Learn class densities $p(x|C_k)$ and priors $p(C_k)$ to obtain $p(x,C_k)=p(x|C_k)p(C_k)$
- Nearest neighbor classification
 - Given query data x, find the closest training points and do majority vote.
- Discriminant functions
 - Learn a function h(x) that maps x onto some C_k .

Outline

- Generative models: Learn class densities $p(x|C_k)$ and priors $p(C_k)$ s.t. $p(x,C_k)=p(x|C_k)p(C_k)$
 - Gaussian Discriminant Analysis
 - Naïve Bayes Classifier

Probabilistic Generative Models

- Bayes' theorem reduces the classification problem $p(C_k|x)$ to estimating the distribution of the data.
- Density estimation problems are easy to learn from labeled training data.
 - Priors: $p(C_k)$
 - Class densities: $p(x|C_k)$
- Learning: Maximum likelihood estimation (MLE)
- Classification: Maximum a posteriori (MAP) estimation

$$\operatorname{argmax}_{k} p(C_{k}|x) = \operatorname{argmax}_{k} p(C_{k}, x)$$

Probabilistic Generative Models

• For two-class classification, Bayes' theorem says:

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)}$$

 The posterior is then expressed as the sigmoid of log odds:

$$p(C_1|\mathbf{x}) = \frac{1}{1 + \exp(-a)} = \sigma(a)$$

where
$$a = \ln \frac{p(C_1|\mathbf{x})}{p(C_2|\mathbf{x})} = \ln \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_2)p(C_2)}$$

Generative vs. Discriminative

- The **generative** approach is typically model-based and makes it possible to generate synthetic data from $p(x|C_k)$.
 - By comparing the synthetic data and real data, we get a sense of how good the generative model is.
- The **discriminative** approach typically has fewer parameters to estimate and have less assumptions about the data distribution (i.e., no p(x, ...)).
 - Linear (e.g., logistic regression) or quadratic (e.g., Gaussian discriminant analysis) in the input.
 - Less generative assumptions about the data (i.e., constructing the features may need prior knowledge)

Gaussian Discriminant Analysis

Gaussian Discriminant Analysis

- Prior distribution $p(C_k)$: Constant (e.g., Bernoulli)
- Likelihood $p(\mathbf{x}|C_k)$: Gaussian distribution

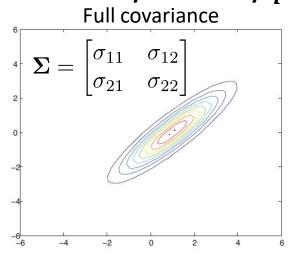
$$p(\mathbf{x}|C_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \mu_k)^T \mathbf{\Sigma}^{-1} (\mathbf{x} - \mu_k)\right\}$$

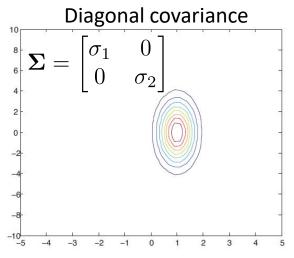
- Classification using Bayes' rule: $p(C_k|\mathbf{x}) = p(\mathbf{x}|C_k)p(C_k)/p(\mathbf{x})$
 - For two-class classification, $p(C_1|\mathbf{x}) = \sigma(a)$

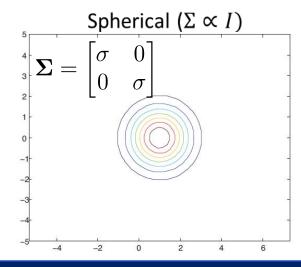
where
$$a = \ln \frac{p(C_1|\mathbf{x})}{p(C_2|\mathbf{x})} = \ln \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_2)p(C_2)}$$

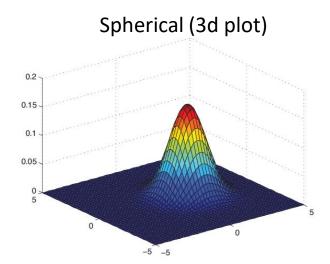
Examples of Gaussian Distributions

• Probability density $p(\mathbf{x})$ for 2-dim case



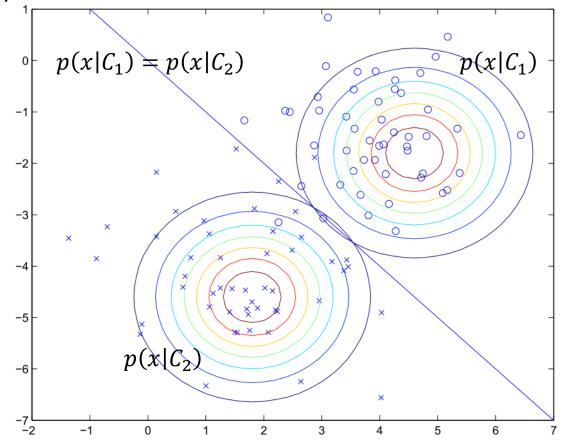






Example: Two-Class GDA

- GDA assumes the same covariance for all classes.
 - The decision boundary is linear.
 - e.g., two-class classification



Two-Class GDA Formulation

• We model $p(\mathbf{x}|C_k)$ as Gaussian distributions with the same covariance matrix.

$$p(\mathbf{x}|C_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \mu_k)^T \mathbf{\Sigma}^{-1} (\mathbf{x} - \mu_k)\right\}$$

Then, the posterior is derived as

$$p(C_1|\mathbf{x}) = \sigma(\mathbf{w}^T\mathbf{x} + w_0)$$

where
$$\mathbf{w} = \mathbf{\Sigma}^{-1}(\mu_1 - \mu_2)$$

$$w_0 = -\frac{1}{2}\mu_1^T \mathbf{\Sigma}^{-1} \mu_1 + \frac{1}{2}\mu_2^T \mathbf{\Sigma}^{-1} \mu_2 + \ln \frac{p(C_1)}{p(C_2)}$$

Two-Class GDA Derivation

$$\begin{split} P(x,C_1) &= P(x|C_1)P(C_1) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right\} P(C_1) \\ P(x,C_2) &= P(x|C_2)P(C_2) \\ &= \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(x-\mu_2)^T \Sigma^{-1}(x-\mu_2)\right\} P(C_2) \\ \log \frac{P(C_1|x)}{P(C_2|x)} &= \log \frac{P(C_1|x)}{1-P(C_1|x)} \qquad \text{``Log-odds''} \\ &= \log \frac{\exp\left\{-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right\}}{\exp\left\{-\frac{1}{2}(x-\mu_2)^T \Sigma^{-1}(x-\mu_2)\right\}} + \log \frac{P(C_1)}{P(C_2)} \\ &= \left\{-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right\} - \left\{-\frac{1}{2}(x-\mu_2)^T \Sigma^{-1}(x-\mu_2)\right\} + \log \frac{P(C_1)}{P(C_2)} \\ &= (\mu_1-\mu_2)^T \Sigma^{-1}x - \frac{1}{2}\mu_1^T \Sigma^{-1}\mu_1 + \frac{1}{2}\mu_2^T \Sigma^{-1}\mu_2 + \log \frac{P(C_1)}{P(C_2)} \end{split} \quad \text{Quadratic term canceled out because of the shared covariance} \end{split}$$

where
$$w_0 = -\frac{1}{2}\mu_1^T \Sigma^{-1} \mu_1 + \frac{1}{2}\mu_2^T \Sigma^{-1} \mu_2 + \log \frac{P(C_1)}{P(C_2)}$$

Extension to Multi-Class GDA

• For two-class classification, the posterior $p(C_k|\mathbf{x})$ is the sigmoid of the log odds

$$\sigma(a) = \frac{1}{1 + \exp(-a)}$$

$$a = \log\left(\frac{\sigma}{1-\sigma}\right) = \left(\Sigma^{-1}(\mu_1 - \mu_2)\right)^T x + w_0^{\frac{1}{\sigma}}$$

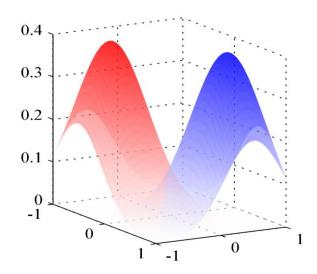
where
$$w_0 = -\frac{1}{2}\mu_1 \Sigma^{-1} \mu_1 + \frac{1}{2}\mu_2 \Sigma^{-1} \mu_2 + \log \frac{P(C_1)}{P(C_2)}$$

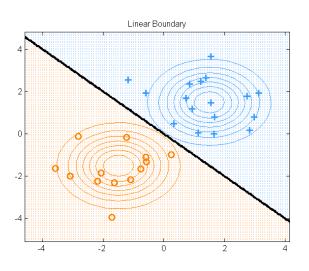
• For multi-class classification, the posterior $p(C_k|\mathbf{x})$ is softmax. $\exp(a_k)$

$$p_i = \frac{\exp(q_i)}{\sum_j \exp(q_j)}$$

GDA Decision Boundaries

- At decision boundary, we have $p(C_1|\mathbf{x}) = p(C_2|\mathbf{x})$
- Under the same covariance assumption, the boundary is linear.
 - Different priors $p(C_1)$, $p(C_2)$ does not change the linearity but shift it around.





Learning GDA via MLE

• Given training data $\{(\mathbf{x}^{(1)}, y^{(1)}), ..., (\mathbf{x}^{(N)}, y^{(N)})\}$ and a generative model with the shared covariance

• Priors:
$$p(y) = \phi^y (1 - \phi)^{1 - y}$$
 where $\phi = p(y = 1)$

Class densities:

$$p(\mathbf{x}|y=0) = \frac{1}{\sqrt{2\pi} |\Sigma|^{\frac{1}{2}}} \exp(-\frac{1}{2} (\mathbf{x} - \mu_0)^T \Sigma^{-1} (\mathbf{x} - \mu_0))$$

$$p(\mathbf{x}|y=1) = \frac{1}{\sqrt{2\pi} |\Sigma|^{\frac{1}{2}}} \exp(-\frac{1}{2} (\mathbf{x} - \mu_1)^T \Sigma^{-1} (\mathbf{x} - \mu_1))$$

Learning GDA via MLE

Maximum likelihood estimation (MLE):

$$\phi = \frac{1}{N} \sum_{i=1}^{N} \mathbf{1} \{ y^{(i)} = 1 \}$$

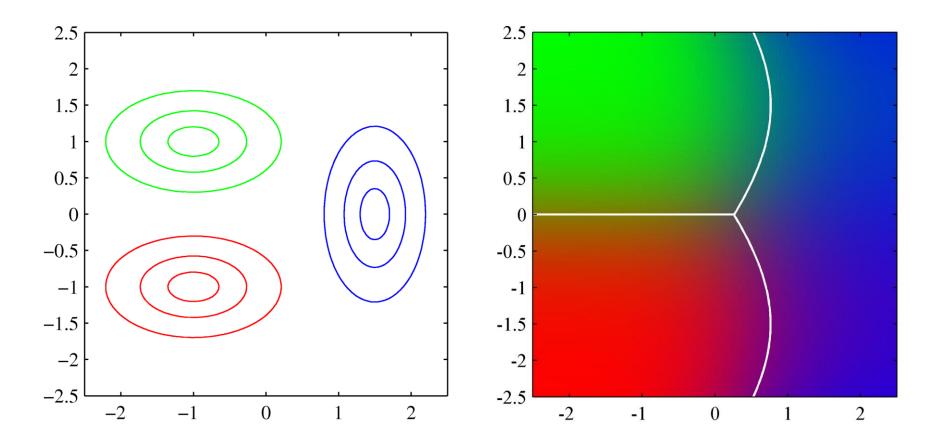
$$\mu_0 = \frac{\sum_{i=1}^{N} \mathbf{1} \{ y^{(i)} = 0 \} \mathbf{x}^{(i)}}{\sum_{i=1}^{N} \mathbf{1} \{ y^{(i)} = 0 \}}$$

$$\mu_1 = \frac{\sum_{i=1}^{N} \mathbf{1} \{ y^{(i)} = 1 \} \mathbf{x}^{(i)}}{\sum_{i=1}^{N} \mathbf{1} \{ y^{(i)} = 1 \}}$$

$$\sum_{i=1}^{N} \sum_{i=1}^{N} (\mathbf{x}^{(i)} - \mu_{y^{(i)}}) (\mathbf{x}^{(i)} - \mu_{y^{(i)}})^T$$

GDA with Different Covariance

 Decision boundaries are quadratic when each class has different covariance.



GDA vs. Logistic Regression

- GDA requires $O(M^2)$ learnable parameters
 - 2M parameters for the means of $p(\mathbf{x}|C_1)$ and $p(\mathbf{x}|C_2)$
 - M(M+1)/2 parameters for shared covariance matrix
 - Cf. Logistic regression requires only *M* parameters

- GDA has a strong modeling assumption and works well when the distribution follows the assumption.
 - Cf. Logistic regression has less parameters and is more flexible about data distribution.

Naïve Bayes Classifier

Naïve Bayes Classifier

- Prior distribution $p(C_k)$: Constant (e.g., Bernoulli)
- Likelihood $p(\mathbf{x}|C_k)$:

$$P(x_1, ..., x_M | C_k) = P(x_1 | C_k) \cdots P(x_M | C_k) = \prod_{j=1}^M P(x_j | C_k)$$

- Naïve Bayes assumption: Each coordinate of x is conditionally independent of other coordinates given the class label.
- Classification using Bayes' rule:

$$p(C_k|\mathbf{x}) = p(\mathbf{x}|C_k)p(C_k)/p(\mathbf{x})$$

For two-class classification,

$$P(C_1|\mathbf{x}) = \frac{P(C_1,\mathbf{x})}{P(\mathbf{x})} = \frac{P(C_1,\mathbf{x})}{P(C_1,\mathbf{x}) + P(C_2,\mathbf{x})}$$

Naïve Bayes Classifier

• Classification using Bayes' rule: $p(C_k|\mathbf{x}) = p(\mathbf{x}|C_k)p(C_k)/p(\mathbf{x})$

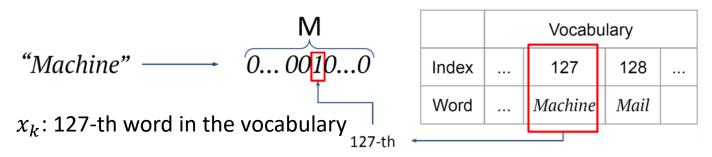
Classification is done by the MAP estimation:

$$\arg\max_k P(C_k|\mathbf{x}) = \arg\max_k P(C_k,\mathbf{x})$$

$$= \arg\max_k P(C_k)P(\mathbf{x}|C_k)$$
 Naïve Bayes Assumption
$$= \arg\max_k P(C_k)\prod_{i=1}^M P(x_j|C_k)$$

Example: Spam Mail Classification

- Label: y = 1 (spam), y = 0 (ham or non-spam)
- Features $\mathbf{x} = [x_1, x_2, ...]$
 - x_k : k-th word in a mail, where M is the vocabulary size
 - Each word is represented as one-hot encoding.



 Naïve Bayes assumption: Given a class label y, each word in a mail is an independent multinomial variable.

Naïve Bayes Classifier Formulation

- Prior: $P(\text{spam}) = Bernoulli(\phi)$
- Likelihood: $P(\text{word}|\text{spam}) = Multinomial(\mu_1^s, \dots, \mu_M^s)$ $P(\text{word}|\text{nonspam}) = Multinomial(\mu_1^{ns}, \dots, \mu_M^{ns})$
- Learning to find ϕ , μ^s , μ^{ns} that best fits the training data $\{(\mathbf{x}^{(1)}, y^{(1)}), ..., (\mathbf{x}^{(N)}, y^{(N)})\}$

$$\prod_{i=1}^{N} P(\mathbf{x}^{(i)}, y^{(i)}) \\
= \prod_{i=1}^{N} P(\mathbf{x}^{(i)} | y^{(i)}) P(y^{(i)}) \\
= \left(\prod_{i:y^{(i)}=1} P(\mathbf{x}^{(i)} | y^{(i)}) P(y^{(i)})\right) \left(\prod_{i:y^{(i)}=0} P(\mathbf{x}^{(i)} | y^{(i)}) P(y^{(i)})\right)$$

Spam

Ham (Non-spam)

$$\left(\prod_{i:y^{(i)}=1}P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)})
ight)$$
 $x_k^{(i)}$ k -th word

- Naïve Bayes assumption:
 - Prior: $P(\text{spam}) = Bernoulli(\phi)$ $P(y^{(i)} = 1) = \phi$

• Likelihood:
$$P(\operatorname{word}|\operatorname{spam}) = \operatorname{Multinomial}(\mu_1^s, \dots, \mu_M^s)$$

$$P(x^{(i)}|y^{(i)} = 1) = \prod_{k=1}^{\operatorname{len}(x^{(i)})} P(x_k^{(i)}|y^{(i)} = 1)$$

$$= \prod_{k=1}^{\operatorname{len}(x^{(i)})} \prod_{j=1}^{M} (\mu_j^s)^{I\left(x_k^{(i)} = j\text{-th word}\right)}$$

$$\left(\prod_{i:y^{(i)}=1} P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)})\right)$$

$$= \left(\prod_{i:y^{(i)}=1}^{N} \left(\prod_{k=1}^{len(x^{(i)})} \prod_{j=1}^{M} \left(\mu_j^s\right)^{I\left(x_k^{(i)}=j\text{-th word}\right)}\right) \phi\right)$$

$$\left(\prod_{i:y^{(i)}=1} P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)})\right) \\
= \left(\prod_{i:y^{(i)}=1}^{N} \left(\prod_{k=1}^{len(x^{(i)})} \prod_{j=1}^{M} (\mu_j^s)^{I(x_k^{(i)}=j-\text{th word})}\right) \phi\right) \\
= \left(\prod_{i:y^{(i)}=1}^{N} \prod_{k=1}^{len(x^{(i)})} \prod_{j=1}^{M} (\mu_j^s)^{I(x_k^{(i)}=j-\text{th word})}\right) \left(\prod_{i:y^{(i)}=1}^{N} \phi\right)$$

$$= \left(\prod_{i:y^{(i)}=1}^{N} \prod_{k=1}^{len(x^{(i)})} \prod_{j=1}^{M} (\mu_j^s)^{I(x_k^{(i)}=j-\text{th word})}\right) \left(\prod_{i:y^{(i)}=1}^{N} \phi\right)$$

$$\begin{pmatrix} \prod_{i:y^{(i)}=1} P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)}) \end{pmatrix}$$

$$= \begin{pmatrix} \prod_{i:y^{(i)}=1}^{N} \left(\prod_{k=1}^{len(x^{(i)})} \prod_{j=1}^{M} \left(\mu_{j}^{s}\right)^{I\left(x_{k}^{(i)}=j\text{-th word}\right)}\right) \phi \end{pmatrix}$$

$$= \begin{pmatrix} \prod_{i:y^{(i)}=1}^{N} \prod_{k=1}^{len(x^{(i)})} \prod_{j=1}^{M} \left(\mu_{j}^{s}\right)^{I\left(x_{k}^{(i)}=j\text{-th word}\right)} \end{pmatrix} \begin{pmatrix} \prod_{i:y^{(i)}=1}^{N} \phi \\ \prod_{i:y^{(i)}=1}^{N} \left(\prod_{j=1}^{N} \mu_{j}^{s}\right)^{\sum_{i:y^{(i)}=1}^{N} \sum_{k=1}^{len(x^{(i)})} I\left(x_{k}^{(i)}=j\text{-th word}\right)} \end{pmatrix} \begin{pmatrix} \prod_{i:y^{(i)}=1}^{N} \phi \end{pmatrix}$$

$$= \begin{pmatrix} \prod_{j=1}^{M} \left(\mu_{j}^{s}\right)^{\sum_{i:y^{(i)}=1}^{N} \sum_{k=1}^{len(x^{(i)})} I\left(x_{k}^{(i)}=j\text{-th word}\right)} \end{pmatrix} \begin{pmatrix} \prod_{i:y^{(i)}=1}^{N} \phi \\ \prod_{i:y^{(i)}=1}^{N} \phi \end{pmatrix}$$

$$\begin{pmatrix} \prod_{i:y^{(i)}=1} P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)}) \end{pmatrix}$$

$$= \begin{pmatrix} \prod_{i:y^{(i)}=1} \left(\prod_{k=1}^{N} \prod_{j=1}^{M} \left(\mu_{j}^{s}\right)^{I\left(x_{k}^{(i)}=j\text{-th word}\right)}\right) \phi \end{pmatrix}$$

$$= \begin{pmatrix} \prod_{i:y^{(i)}=1} \prod_{k=1}^{len(x^{(i)})} \prod_{j=1}^{M} \left(\mu_{j}^{s}\right)^{I\left(x_{k}^{(i)}=j\text{-th word}\right)} \end{pmatrix} \begin{pmatrix} \prod_{i:y^{(i)}=1} \phi \end{pmatrix}$$

$$= \begin{pmatrix} \prod_{j=1}^{M} \left(\mu_{j}^{s}\right)^{\sum_{i:y^{(i)}=1}^{N} \sum_{k=1}^{len(x^{(i)})} I\left(x_{k}^{(i)}=j\text{-th word}\right)} \end{pmatrix} \begin{pmatrix} \prod_{i:y^{(i)}=1} \phi \end{pmatrix}$$

$$= \begin{pmatrix} \prod_{j=1}^{M} \left(\mu_{j}^{s}\right)^{N_{j}^{spam}} \end{pmatrix} \text{count \# of j-th word in the vocabulary}$$

$$= \begin{pmatrix} \prod_{j=1}^{M} \left(\mu_{j}^{s}\right)^{N_{j}^{spam}} \end{pmatrix} \phi^{N^{spam}} \text{: Total \# of j-th word in spam emails}$$

$$N^{spam} \text{: Total \# of spam emails}$$

$$\left(\prod_{i:y^{(i)}=1}P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)})
ight)$$
 $x_k^{(i)}$ i -th mail $x_k^{(i)}$

- Naïve Bayes assumption:
 - Prior: $P(\text{spam}) = Bernoulli(\phi)$ $P(y^{(i)} = 1) = \phi$

• Likelihood:
$$P(\operatorname{word}|\operatorname{spam}) = \operatorname{Multinomial}(\mu_1^s, \dots, \mu_M^s)$$

$$P(x^{(i)}|y^{(i)} = 1) = \prod_{k=1}^{\operatorname{len}(x^{(i)})} P(x_k^{(i)}|y^{(i)} = 1)$$

$$= \prod_{k=1}^{\operatorname{len}(x^{(i)})} \prod_{j=1}^{M} (\mu_j^s)^{I\left(x_k^{(i)} = j\text{-th word}\right)}$$

Likelihood for ham (non-spam):

$$\left(\prod_{i:y^{(i)}=1} P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)})\right) \left(\prod_{i:y^{(i)}=0} P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)})\right)$$

- Naïve Bayes assumption:
 - Prior: $P(\text{spam}) = Bernoulli(\phi)$ $P(y^{(i)} = 0) = 1 - \phi$

• Likelihood:
$$P(\operatorname{word}|\operatorname{nonspam}) = \operatorname{Multinomial}(\mu_1^{ns}, \dots, \mu_M^{ns})$$

$$P(x^{(i)}|y^{(i)} = 0) = \prod_{k=1}^{len(x^{(i)})} P(x_k^{(i)}|y^{(i)} = 0)$$

$$= \prod_{k=1}^{len(x^{(i)})} \prod_{j=1}^{M} (\mu_j^{ns})^{I\left(x_k^{(i)} = j\text{-th word}\right)}$$

Putting together:

$$\prod_{i=1}^{N} P(\mathbf{x}^{(i)}, y^{(i)})$$

$$= \left(\prod_{i:y^{(i)}=1} P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)})\right) \left(\prod_{i:y^{(i)}=0} P(\mathbf{x}^{(i)}|y^{(i)})P(y^{(i)})\right)$$

$$= \left(\phi^{N^{spam}} \prod_{word j} (\mu_j^s)^{N_j^{spam}}\right) \left((1-\phi)^{N^{nonspam}} \prod_{word j} (\mu_j^{ns})^{N_j^{nonspam}}\right)$$

Log-likelihood

$$\begin{split} & \log P(\mathcal{D}) \\ &= & \log \prod_{i=1}^{N} P(x^{(i)}, y^{(i)}) \\ &= & N^{spam} \log \phi + \sum_{word j} N_{j}^{spam} \log \mu_{j}^{s} + N^{nonspam} \log (1 - \phi) + \sum_{word j} N_{j}^{nonspam} \log \mu_{j}^{ns} \end{split}$$

Log-likelihood

$$\begin{split} & \log P(\mathcal{D}) \\ &= & \log \prod_{i=1}^{N} P(x^{(i)}, y^{(i)}) \\ &= & N^{spam} \log \phi + \sum_{word \ j} N_{j}^{spam} \log \mu_{j}^{s} + N^{nonspam} \log (1 - \phi) + \sum_{word \ j} N_{j}^{nonspam} \log \mu_{j}^{ns} \end{split}$$

- Maximum (log-)likelihood estimation
 - Take the derivative of log-likelihood with respect to the parameters $\{\phi, \mu^s, \mu^{ns}\}$, and set it to zero.

• Find ϕ

$$\log P(\mathcal{D})$$

$$= N^{spam} \log \phi + \sum_{word j} N^{spam}_j \log \mu^s_j + N^{nonspam} \log (1 - \phi) + \sum_{word j} N^{nonspam}_j \log \mu^{ns}_j$$

$$\phi = \frac{N^{spam}}{N^{spam} + N^{nonspam}}$$

- Find μ^s (or similarly, μ^{ns})
 - $\{\mu^S\}$'s are NOT independent to each other; $\sum_{j=1}^M \mu_j^S = 1$
 - To only deal with variables independent to each other, we can express μ_M^s as $1-\sum_{i=1}^{M-1}\mu_i^s$



$$\sum_{word \ j=1}^{M} N_{j}^{spam} \log \mu_{j}^{s} = \sum_{word \ j=1}^{M-1} N_{j}^{spam} \log \mu_{j}^{s} + N_{M}^{spam} \log (1 - \sum_{j=1}^{M-1} \mu_{j}^{s})$$

- Find μ^s (or similarly, μ^{ns})
 - $\{\mu^S\}$'s are NOT independent to each other; $\sum_{j=1}^M \mu_j^S = 1$
 - To only deal with variables independent to each other, we can express μ_M^s as $1-\sum_{j=1}^{M-1}\mu_j^s$

$$\sum_{word \, j=1}^{M} N_j^{spam} \log \mu_j^s = \sum_{word \, j=1}^{M-1} N_j^{spam} \log \mu_j^s + N_M^{spam} \log (1 - \sum_{j=1}^{M-1} \mu_j^s)$$
 Take derivative
$$\frac{\partial}{\partial \mu_j^s} \left(\sum_{word \, j=1}^{M} N_j^{spam} \log \mu_j^s \right) = \frac{N_j^{spam}}{\mu_j^s} - \frac{N_M^{spam}}{1 - \sum_{j=1}^{M-1} \mu_j^s} = 0$$
 Constant w.r.t. j

- Find μ^s (or similarly, μ^{ns})
 - $\{\mu^S\}$'s are NOT independent to each other; $\sum_{j=1}^M \mu_j^S = 1$
 - To only deal with variables independent to each other, we can express μ_M^s as $1-\sum_{j=1}^{M-1}\mu_j^s$

$$\sum_{word \, j=1}^{M} N_j^{spam} \log \mu_j^s = \sum_{word \, j=1}^{M-1} N_j^{spam} \log \mu_j^s + N_M^{spam} \log (1 - \sum_{j=1}^{M-1} \mu_j^s)$$

$$\frac{\partial}{\partial \mu_j^s} \left(\sum_{word \, j=1}^{M} N_j^{spam} \log \mu_j^s \right) = \frac{N_j^{spam}}{\mu_j^s} - \frac{N_M^{spam}}{1 - \sum_{j=1}^{M-1} \mu_j^s} = 0$$

$$\frac{a}{b} = \frac{c}{d} = \frac{a+c}{b+d}$$

$$\frac{N_j^{spam}}{\mu_j^s} = \text{constant} = \frac{\sum_{j=1}^{M} N_j^{spam}}{\sum_{j=1}^{M} \mu_j^s} = \sum_{j=1}^{M} N_j^{spam} = N^{spam}$$

- Find μ^s (or similarly, μ^{ns})
 - $\{\mu^S\}$'s are NOT independent to each other; $\sum_{j=1}^M \mu_j^S = 1$
 - To only deal with variables independent to each other, we can express μ_M^S as $1-\sum_{i=1}^{M-1}\mu_i^S$

$$\sum_{word j=1}^{M} N_{j}^{spam} \log \mu_{j}^{s} = \sum_{word j=1}^{M-1} N_{j}^{spam} \log \mu_{j}^{s} + N_{M}^{spam} \log (1 - \sum_{j=1}^{M-1} \mu_{j}^{s})$$

$$\frac{\partial}{\partial \mu_{j}^{s}} \left(\sum_{word j=1}^{M} N_{j}^{spam} \log \mu_{j}^{s} \right) = \frac{N_{j}^{spam}}{\mu_{j}^{s}} - \frac{N_{M}^{spam}}{1 - \sum_{j=1}^{M-1} \mu_{j}^{s}} = 0$$

$$\therefore \quad \mu_{j}^{s} = \frac{N_{j}^{spam}}{\sum_{i} N_{j}^{spam}}$$

Summary

$$\begin{split} P(spam) &= \phi = \frac{N^{spam}}{N^{spam} + N^{nonspam}} \\ P(word = j | spam) &= \mu_j^s = \frac{N_j^{spam}}{\sum_j N_j^{spam}} \\ P(word = j | non - spam) &= \mu_j^{ns} = \frac{N_j^{nonspam}}{\sum_j N_j^{nonspam}} \end{split}$$

- N^{spam}: Total # of spam emails
- N^{nonspam}: Total # of ham (non-spam) emails
- N_i^{spam} : Total # of j-th word in spam emails
- $N_j^{nonspam}$: Total # j-th word in ham (non-spam) emails

Laplace Smoothing

- Maximum likelihood is problematic when a specific word count is 0.
 - Leads to probability of the specific word 0
- Solution: Put imaginary counts for each word
 - Prevent zero probability estimates (overfitting)
 - Add "1" as imaginary count for each word

$$P(spam) = \phi = \frac{N^{spam}}{N^{spam} + N^{nonspam}}$$

$$P(word = j | spam) = \mu_j^s = \frac{N_j^{spam} + 1}{\sum_j N_j^{spam} + M}$$

$$P(word = j|non - spam) = \mu_j^{ns} = \frac{N_j^{nonspam} + 1}{\sum_j N_j^{nonspam} + M}$$

Laplace Smoothing

- Maximum likelihood is problematic when a specific word count is 0.
 - Leads to probability of the specific word 0
- Solution: Put imaginary counts for each word
 - Prevent zero probability estimates (overfitting)
 - Add "1" as imaginary count for each word

$$P(spam) = \phi = \frac{N^{spam}}{N^{spam} + N^{nonspam}}$$

If we smooth prior as well:

$$P(spam) = \phi = \frac{N^{spam} + 1}{N^{spam} + N^{nonspam} + 2}$$

(Don't have to do this)

Next: Other Classifiers