

The Role of Oil Prices in Russia's Economy

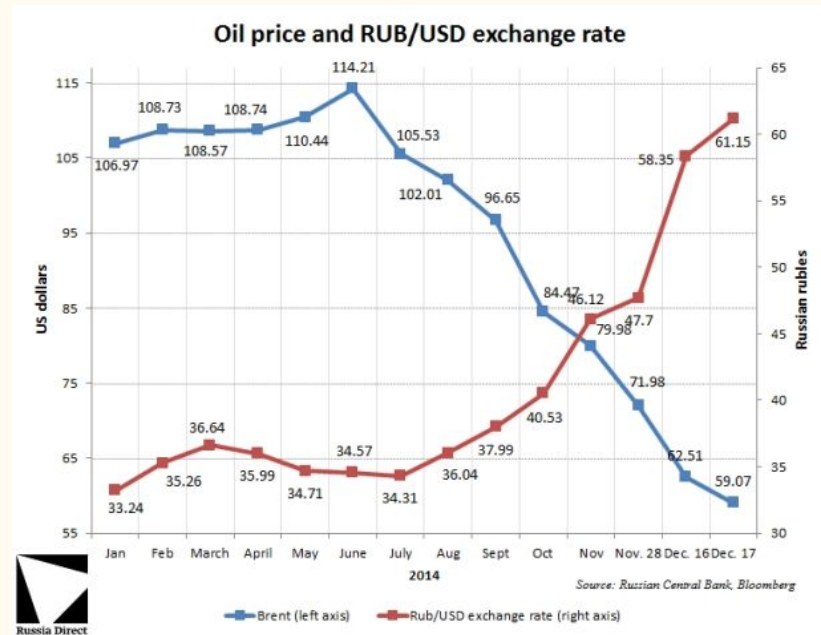
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Introduction

- Economy is highly dependent
 - In 2013, oil and gas was 68% of Russian exports
- Oil and gas is nearly half of the Russian government's budget
- Oil price directly impacts the Ruble price:
 - Cheap oil causes depreciation
 - Correlation strengthened as oil production increased
- Crucial factor in geopolitics
 - May impact military spending



Economic Impact of the Oil Price



- As cheap oil weakens the Russian economy, this causes the ruble to depreciate, leading to inflation.
- Causes high volatility in foreign exchange markets.
- Currency weakness accelerated inflation to a 13-year high of 16.9 percent in March 2015.
- Annual consumer-price growth eased to 12.9 percent in December 2015.

Data set

2003 Q1 - 2016 Q3

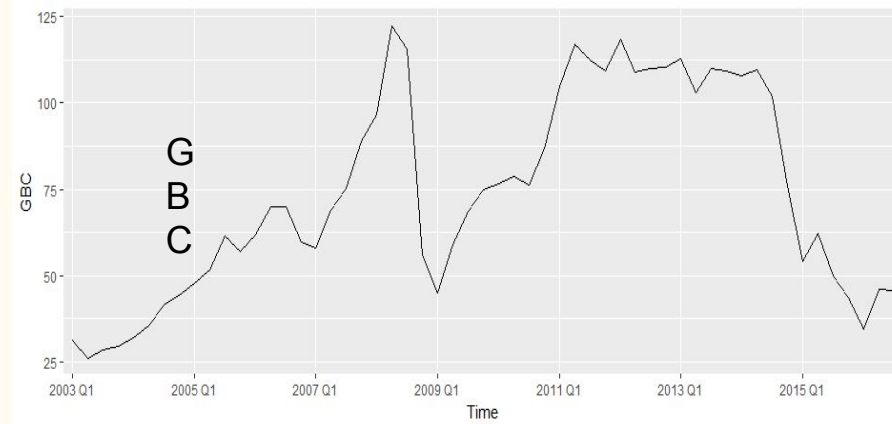
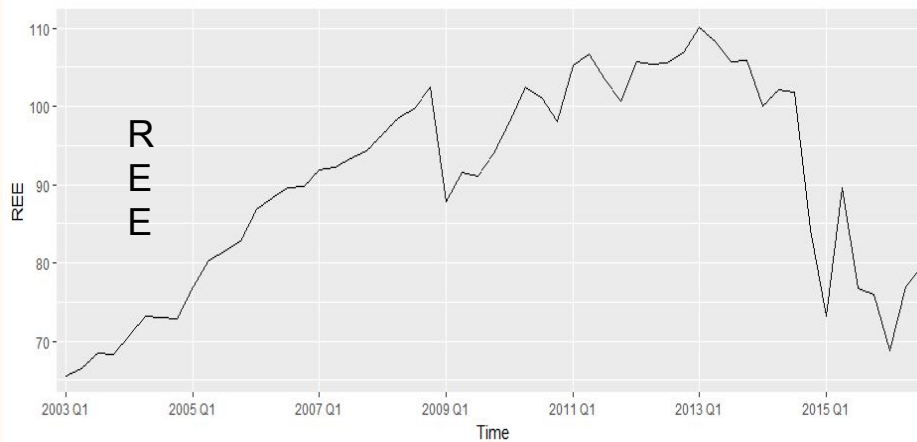
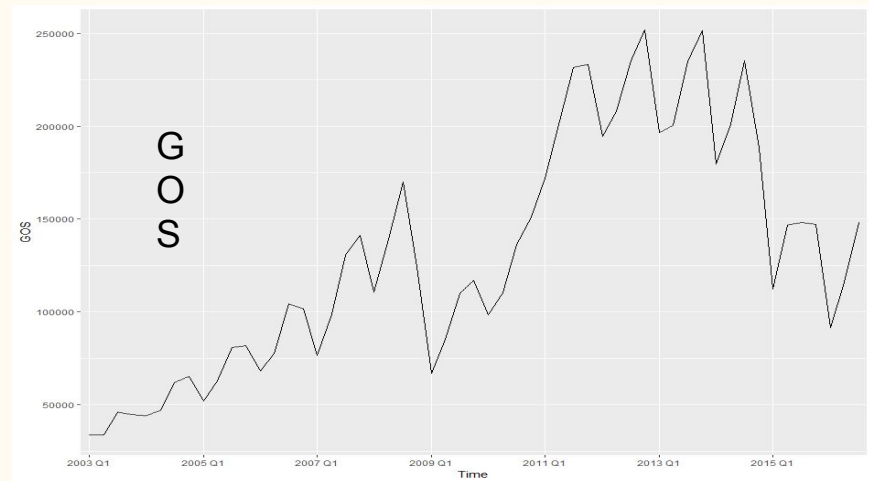
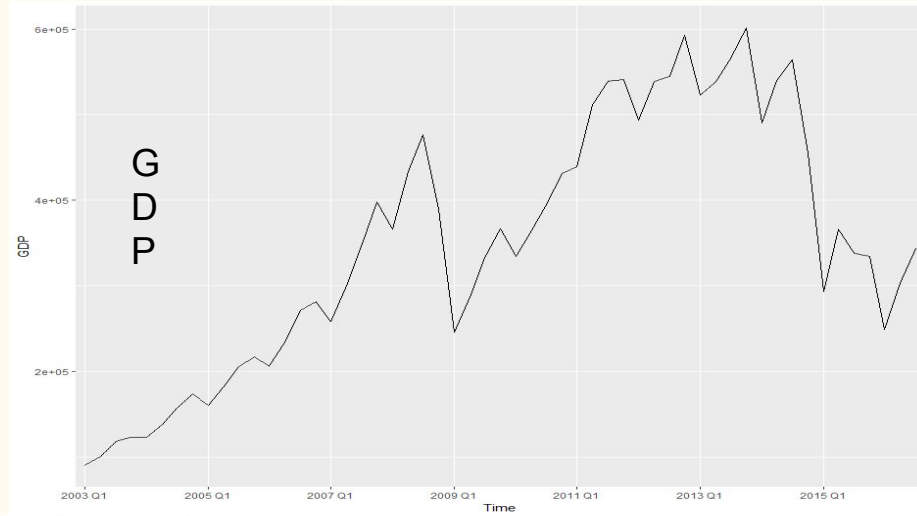
Gross Domestic Product (GDP)

Gross Operating Surplus (GOS)

Real Effective Exchange Rates (REE)

Global Price of Brent Crude (GBC)

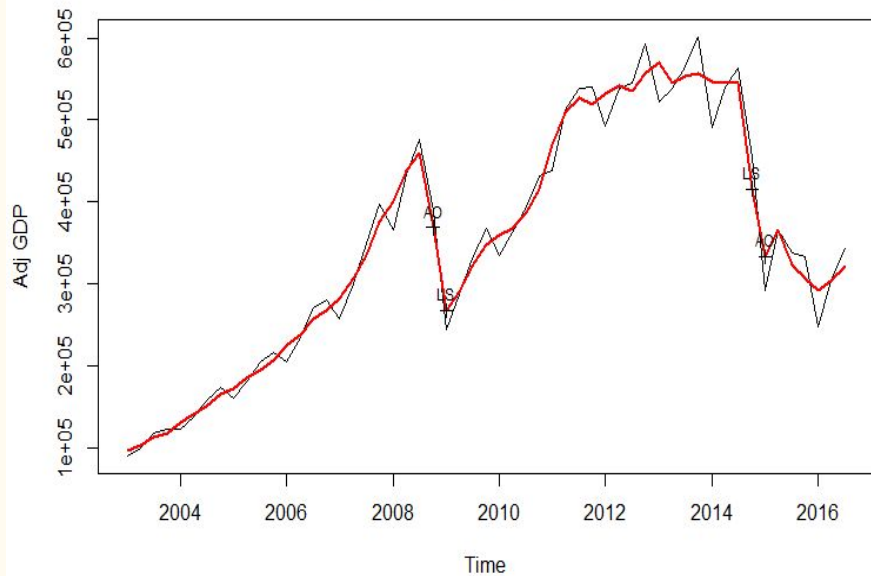




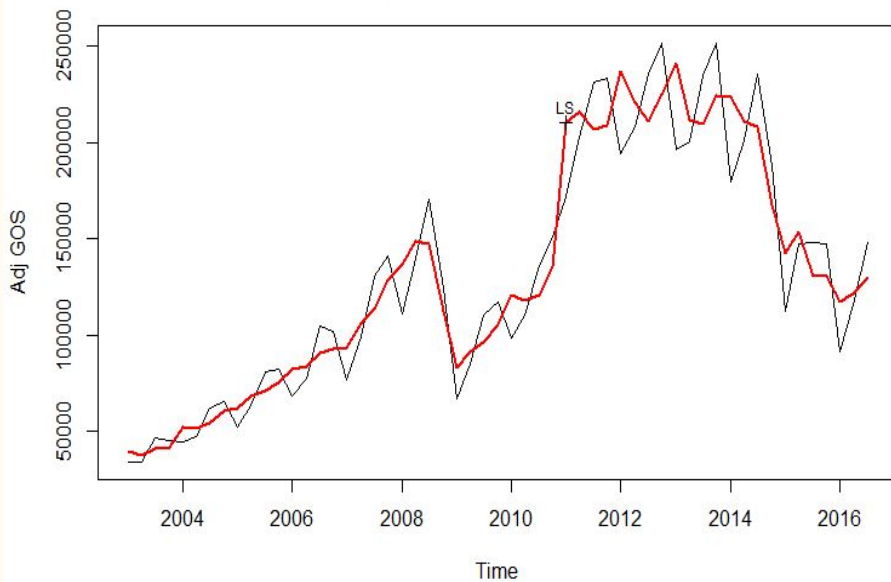
Seasonal adjustment

- by using X-13ARIMA-Seats

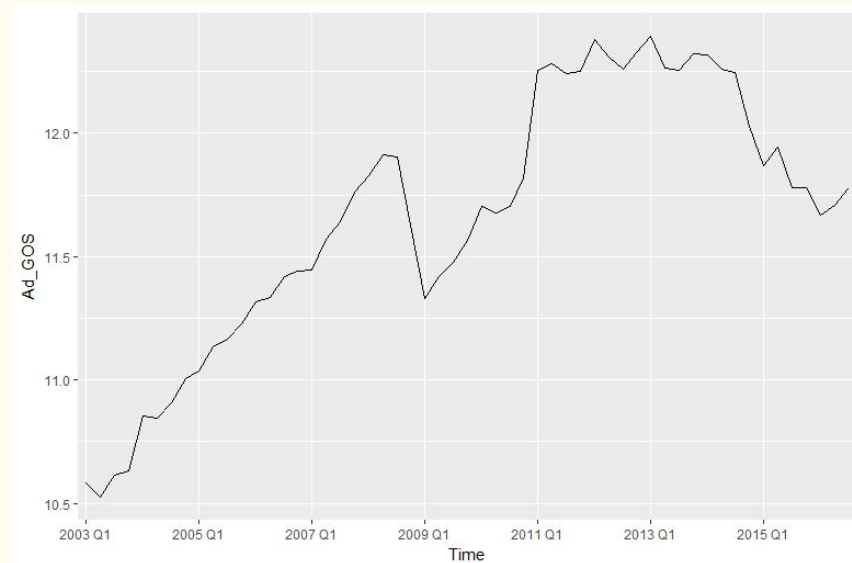
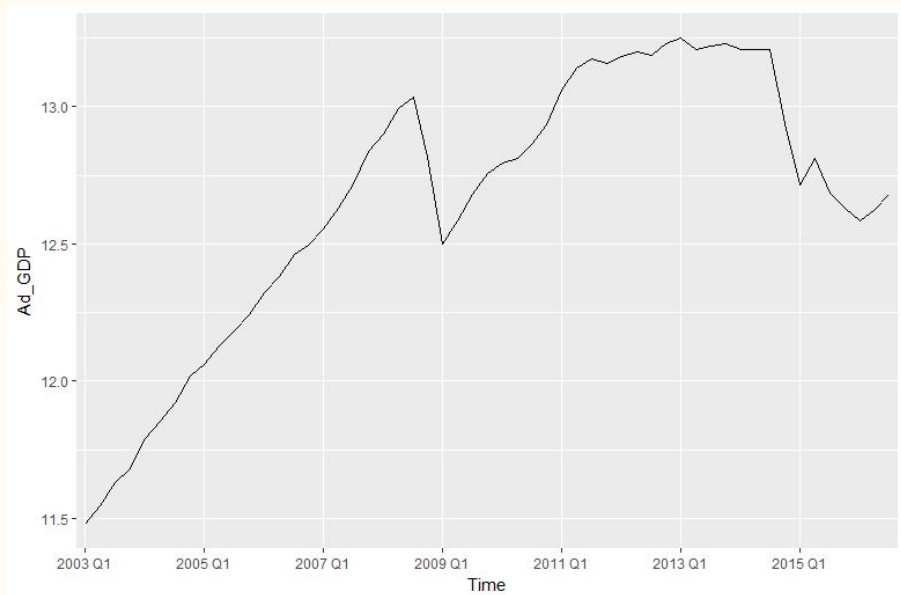
Original and Adjusted Series



Original and Adjusted Series



Log Transformation on GDP and GOS



Unit root tests

Augmented Dickey-Fuller test: All series are I(1)

	GDP	Diff_GDP	GOS	Diff_GOS	REE	Diff_REE	GBC	Diff_GBC
None	-1.832	-4.726***	-1.618	-4.445***	-1.565	-5.959***	-2.116	-5.490***
Constant	1.886	7.494***	1.445	6.627***	1.680	11.990***	1.889	10.07***
Con. + Trend	2.729	11.240***	2.038	9.940***	2.512	17.980***	2.825	15.098***

Both X_{it} & Y_t are $I(1)$



$$Y_t - \beta_i X_{it} = \varepsilon_t \sim I(0)$$

Cointegration-Johansen test

Lag Selection: Varselect()

AIC(n)	HQ(n)	SC(n)	FPE(n)
2	2	1	2

Johansen Test: find one cointegration relationship

	test	10pct	5pct	1pct
r <= 3	3.61	7.52	9.24	12.97
r <= 2	10.57	17.85	19.96	24.60
r <= 1	21.17	32.00	34.91	41.07
r = 0	58.16	49.65	53.12	60.16

Vector Error Correction Model

Vector Error Correction Model

$$\Delta y_t = \Pi y_{t-1} + \Gamma_1 \Delta y_{t-1} + \dots + \Delta y_{t-p+1} + \varepsilon_t$$

Only have one cointegration relationship, then,

$$\Delta y_t = \alpha \beta^T y_{t-1} + \Gamma_1 \Delta y_{t-1} + \dots + \Delta y_{t-p+1} + \varepsilon_t$$

where α is 4×1 and β^T is 1×4

Vector Error Correction Model

Then

$$\Delta Y_t = \begin{bmatrix} -0.260 \\ -0.259 \\ 0.020 \\ -0.572 \end{bmatrix} [1.000 \quad -0.757 \quad -2.983 \quad 0.892] Y_{t-1} + \begin{bmatrix} 0.402 & -0.019 & -1.085 & 0.383 \\ 0.607 & -0.132 & -1.488 & 0.406 \\ 0.191 & 0.082 & -0.552 & 0.047 \\ 0.261 & -0.004 & -1.696 & 0.519 \end{bmatrix} \Delta Y_{t-1}$$

Test residuals,

	GDP	GOS	REE	GBC
ADF Test	-4.7003***	-4.2066***	-4.1282***	-5.2647***
Normality Test	0.1631***	0.6431***	0.0041	0.0002
Ljung-Box Test	0.8729	0.9267	0.5577	0.971

Cointegration- Engle and Granger test

1. Test if there is Unit root in each component series individually by using ADF.
 2. Test cointegration among the components, i.e., to test whether αY_t is $I(0)$.
- If $Y_t = (Y_{1t}, Y_{2t}, \dots, Y_{mt})$ is cointegrated, αY_t is $I(0)$ where $\alpha = (\alpha_1, \alpha_2, \dots, \alpha_m)$. Then, $(1/\alpha_1)\alpha$ is also a cointegrated vector where $\alpha_1 \neq 0$

Cointegration- Engle and Granger test

- EG consider the regression model for Y_{1t}

$$Y_{1t} = \delta D_t + \phi_1 Y_{2t} + \dots + \phi_{m-1} Y_{mt} + \varepsilon_t$$

where D_t : deterministic terms.

- Check whether residual term is $I(1)$ or $I(0)$
 - If $I(1)$, then Y_t is not cointegrated.
 - If $I(0)$, then Y_t is cointegrated with a normalized cointegrating vector
- Steps:
 - Run OLS. Get estimate $\hat{\alpha} = (1, \hat{\phi}_1, \dots, \hat{\phi}_{m-1})$
 - Use residual for unit root testing.

$$\alpha' = (1, \phi_1, \dots, \phi_{m-1})$$

Cointegration- Engle and Granger test

Is cointegrated?

	GDP	GOS	REE	GBC
GDP		No	No	Yes
GOS	No		No	No
REE	No	No		Yes
GBC	Yes	No	Yes	

Causality test of GDP and GBC

Granger Causality Null Hypothesis	P-value
GDP do not Granger-cause GBC	0.7794***
GBC do not Granger cause GDP	0.0051

- GBC granger causes Russia GDP.
- At the same time, Russia GDP does not granger cause GBC

Error Correction Model of GDP and GBC

Select Lags

AIC(n)	HQ(n)	SC(n)	FPE(n)
2	2	2	2

Then,

$$\Delta \text{GDP}_t = 1.3997 - 0.1592(\text{GDP}_{t-1} - 1.058511\text{GBC}_{t-1}) + 0.4184\Delta \text{GBC}_{t-1}$$

The residual of ECM

Test	Value
ADF	-5.0995***
Normality Test	0.0001
Ljung-Box Test	0.9994***

By using the ADF test, we can find out that the residual of ECM follow the $I(0)$.

Thus, the long run are stationary.

Comparison of Engle-Granger Test and Johansen Test

First of all, the two methods are essentially different, and may disagree on inferences from the same data. The Engle-Granger two-step method for estimating the VEC model, first estimating the cointegrating relation and then estimating the remaining model coefficients, differs from Johansen's approach.

Secondly, the cointegrating relations estimated by the Engle-Granger approach may not correspond to the cointegrating relations estimated by the Johansen approach, especially in the presence of multiple cointegrating relations. It is important, in this context, to remember that cointegrating relations are not uniquely defined, but depend on the decomposition $C = AB'$ of the impact matrix.

Vector Autoregressive Model

- Reduced form

$$\begin{bmatrix} X_{1t} \\ X_{2t} \\ \vdots \\ X_{Kt} \end{bmatrix} = \begin{bmatrix} \phi_{11} & \cdots & \phi_{1K} \\ \vdots & \ddots & \vdots \\ \phi_{K1} & \cdots & \phi_{KK} \end{bmatrix} \begin{bmatrix} X_{1,t-1} \\ X_{2,t-1} \\ \vdots \\ X_{K,t-1} \end{bmatrix} + \begin{bmatrix} \varepsilon_{1,t} \\ \varepsilon_{2,t} \\ \vdots \\ \varepsilon_{K,t} \end{bmatrix}, \begin{bmatrix} \varepsilon_{1,t} \\ \varepsilon_{2,t} \\ \vdots \\ \varepsilon_{K,t} \end{bmatrix} \sim N(0, \Sigma).$$

- Structural form

$$\begin{bmatrix} 1 & 0 & \cdots & 0 \\ l_{21}^* & 1 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ l_{K1}^* & l_{K2}^* & \cdots & 1 \end{bmatrix} \begin{bmatrix} X_{1t} \\ X_{2t} \\ \vdots \\ X_{Kt} \end{bmatrix} = \begin{bmatrix} \phi_{11}^* & \cdots & \phi_{1K}^* \\ \vdots & \ddots & \vdots \\ \phi_{K1}^* & \cdots & \phi_{KK}^* \end{bmatrix} \begin{bmatrix} X_{1,t-1} \\ X_{2,t-1} \\ \vdots \\ X_{K,t-1} \end{bmatrix} + \begin{bmatrix} \varepsilon_{1,t}^* \\ \varepsilon_{2,t}^* \\ \vdots \\ \varepsilon_{K,t}^* \end{bmatrix}, \text{cov}(\varepsilon_{i,t}^*, \varepsilon_{j,t}^*) = 0$$

- ARDL

$$\begin{cases} X_{1t} = \phi_{11}^* X_{1,t-1} + \cdots + \phi_{1K}^* X_{K,t-1} + \varepsilon_{1,t}^* \\ X_{2t} = \phi_{21}^* X_{1,t-1} + \cdots + \phi_{2K}^* X_{K,t-1} - l_{21}^* X_{1t} + \varepsilon_{2,t}^* \\ X_{3t} = \phi_{31}^* X_{1,t-1} + \cdots + \phi_{3K}^* X_{K,t-1} - l_{31}^* X_{1t} - l_{32}^* X_{2t} + \varepsilon_{3,t}^* \\ \vdots \\ X_{Kt} = \phi_{K1}^* X_{1,t-1} + \cdots + \phi_{KK}^* X_{K,t-1} - \sum_{i=1}^{K-1} l_{Ki}^* X_{it} + \varepsilon_{K,t}^* \end{cases}$$

Vector Autoregressive After Difference

Select lag

AIC(n)	HQ(n)	SC(n)	FPE(n)
4	1	1	4

VAR with lag 1

$$\begin{aligned}GDP_t &= 0.0137 + 0.2608GDP_{t-1} - 0.0751GOS_{t-1} - 0.2160REE_{t-1} + 0.3493GBC_{t-1} \\GOS_t &= 0.0133 + 0.6007GDP_{t-1} - 0.1079GOS_{t-1} - 1.0184REE_{t-1} + 0.2790GBC_{t-1} \\REE_t &= -0.0047 + 0.1723GDP_{t-1} + 0.00951GOS_{t-1} - 0.1514REE_{t-1} + 0.1003GBC_{t-1} \\GBC_t &= 0.0119 - 0.1878GDP_{t-1} + 0.06151GOS_{t-1} + 0.0187REE_{t-1} + 0.3949GBC_{t-1}\end{aligned}$$

Breusch-Godfrey test

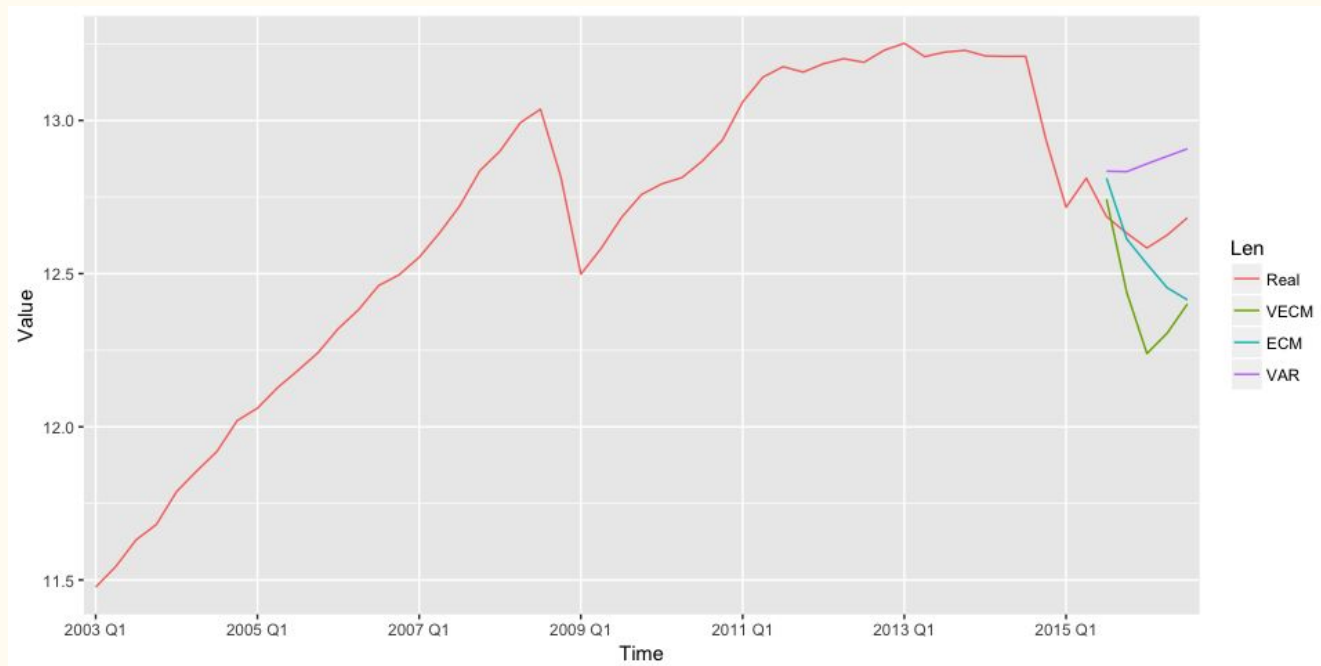
H0: There is no ACF of any order up to 10

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Breusch-Godfrey LM test
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data:  Residuals of VAR object Var_Mdl
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Chi-squared = 94.911, df = 80, p-value = 0.1221
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Predict GDP



Conclusion

- We examined how sensitive Russia's output and fiscal policy are to changes in international oil prices and the real exchange rate of the rouble. The findings support the prevailing common view that both of these factors play a major role in the Russian economy.
- More interesting, we found Russia's GDP and international oil prices has a long-run stationary relationship.