

## Exercise 8.15 and 8.20

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### 0.1 8.15

a. Show that the reliability  $\varnothing$  of Figure 8.2 can be expressed as:

$$\varnothing(X_1, \dots, X_n) = X_1 X_5 \max(X_2, X_3 X_4)$$

$X_i$  : The indicator that component i is working, then  $X_i$  is either the value

$$X_i = \begin{cases} 0 & \text{not working} \\ 1 & \text{working} \end{cases}$$

As shown in the Figure 8.2, in order to make the system to work, component 1 and component 5 must work as well; Between component 1 and 5, the system depends on components 2,3 and 4, here there are three paths:

$$\begin{cases} 2 \\ 2 \text{ and } 4 \\ 3 \text{ and } 4 \end{cases}$$

Because  $\max(X_2, X_3 X_4)$  is same with  $X_2$   
then

$$\varnothing(X_1, \dots, X_5) = X_1 \max(X_2, X_3 X_4) X_5 = X_1 X_5 \max(X_2, X_3 X_4)$$

b. Consider the representation  $X_3 = 1_{\{U \leq \theta\}}$  and show that IPA is biased for  $\varnothing$  (8.28)

Show IPA is biased:

$$\begin{aligned} E[X_3(\theta)] &= P(U \leq \theta) = \theta \\ \Rightarrow \frac{\partial}{\partial \theta} [E(X_3(\theta))] &= 1 \end{aligned}$$

However  $X_3(\theta, v)$  is piecewise constant function, jump at  $v = \theta$

So for  $\forall w$ , such that  $U(w) \neq \theta$

we have  $\frac{\partial}{\partial \theta} X_3(\theta, v) = 0$

As  $P(V(w) = \theta) = 0$  (when  $U(w) \neq \theta$ )

then  $E[X_3'(\theta, v)] = 0$ ,

then  $1 = \frac{\partial}{\partial \theta} [E(X_3(\theta))] \neq E[\frac{\partial}{\partial \theta} X_3(\theta, v)] = 0$

then IPA is biased function  $X_3$ , and consequently is biased as well for  $\varnothing$  (8.28)

**c. Calculate the SF estimator for (8.28) and show that it is unbiased for (8.28)**

$X_i$  is Bernoulli ( $p_i$ ), eg:

$$\begin{cases} P(X_i = 1) = p_i \\ P(X_i = 0) = 1 - p_i \end{cases}$$

the pdf of  $X_i$  is  $f(X_i, p_i) = p_i^{x_i}(1 - p_i)^{1-x_i}$ ,  $X_i \in 0, 1$

The density function is given by  $f_n \theta = P_3$

$$L(\theta|x) = f(x, \theta) = \theta^{x_3}(1 - \theta)^{1-x_3} \pi_{i \neq 3} p_i^{x_i}(1 - p_i)^{1-x_i}, x_i \in \{0, 1\}$$

$$\mathcal{O}(X_1, \dots, X_5) = X_1 X_5 \max(X_2, X_3 X_4)$$

then  $\mathcal{O}(X_1, \dots, X_5) = 1$  only if one of the following

$$\begin{cases} X = (1, 1, 1, 1, 1) \\ X = (1, 1, 0, 1, 1) \\ X = (1, 1, 1, 0, 1) \\ X = (1, 1, 0, 0, 1) \\ X = (1, 0, 1, 1, 1) \end{cases}$$

$$L(\theta|X) = \theta[p_1 p_2 p_4 p_5 + p_1 p_2 p_5 + p_1 p_2 p_4 p_5] + (1 - \theta)[p_1 p_2 p_4 p_5 + p_1 p_2 p_5]$$

$$\log L(\theta|X) = \log[\theta[p_1 p_2 p_4 p_5 + p_1 p_2 p_5 + p_1 p_2 p_4 p_5] + (1 - \theta)[p_1 p_2 p_4 p_5 + p_1 p_2 p_5]]$$

$$= \log \theta + \log(1 - \theta) + \log[p_1 p_2 p_4 p_5 + p_1 p_2 p_5 + p_1 p_2 p_4 p_5] + \log[p_1 p_2 p_4 p_5 + p_1 p_2 p_5]$$

The score function will be

$$S(\theta|X) = \frac{\partial \log(L(\theta|X))}{\partial \theta} = \frac{1}{\theta} - \frac{1}{1 - \theta}$$

unbiased  $L(\theta|X) = f(x, \theta) = \theta C_1 + (1 - \theta) C_2$ , where

$$C_1 = p_1 p_2 p_4 p_5 + p_1 p_2 p_5 + p_1 p_2 p_4 p_5$$

$$C_2 = p_1 p_2 p_4 p_5 + p_1 p_2 p_5$$

(i) Easy to see that  $f(X, \theta)$  is differentiable in  $\theta \in (0, 1)$ , using the bounding condition theorem

$$\frac{\partial}{\partial \theta} f(X, \theta) = C_1 - C_2$$

$$\text{let } k(n) = \begin{cases} C_1 - C_2 & n = 0 \\ 0 & \text{else} \end{cases} \quad n = 1$$

$$\rightarrow \sum_{n \in v} V(n) K(n) = (V(0) + V(1))(C_1 - C_2) < \infty$$

then (ii) is satisfied

Using theorem 8.3 then SF and MVD estimate function 8.28 are unbiased.

**d. Calculate the MVD estimator for (8.28) and show that it is unbiased for (8.28)**

$$\frac{d}{d\theta} |_{\theta=\theta_0} \int h(x) f_{\theta}(x) dx = \int h(x) S(\theta_0, x) f_{\theta_0}(x) dx = C_{\theta_0} (\int h(x) f_{\theta_0}^+(x) dx - \int h(x) f_{\theta_0}^-(x) dx)$$

In particular when

$$S(\theta_0, x) f_{\theta_0}(x) = C_{\theta_0} (f_{\theta_0}^+(x) - f_{\theta_0}^-(x))$$

From tabel 7.1  $C_{\theta_0} = 1$

$$f_{\theta_0}^+(x) = \text{Dirac}(0) = S_0$$

$$f_{\theta_0}^-(x) = \text{Dirac}(1) = S_1$$

then

$$\frac{\partial}{\partial \theta} \int h(x) f_{\theta}(x) dx = \frac{\partial}{\partial \theta} [\theta S(0) + (1 - \theta) S(1)] = h(0) - h(1)$$

while yields the unbiased MVD estimator

$$D^{MVD}(\theta) = h(0) - h(1)$$

Same proof in c, as the MVD was derived from SF

## 0.2 8.20

**(a) Calculate the Score Function for this distribution.**

Because  $f_{\theta}(x) = (x^2/2\theta^3)e^{-x/\theta}$

And  $S(\theta, x) = \frac{\partial}{\partial \theta} \log(f_{\theta}(x))$

So:

$$\begin{aligned} S(\theta, x) &= \frac{\partial}{\partial \theta} \log((x^2/2\theta^3)e^{-x/\theta}) = \frac{\frac{\partial}{\partial \theta} f_{\theta}(x)}{f_{\theta}(x)} \\ &= \frac{x^2}{2\theta^3} e^{-x/\theta} \frac{x}{\theta^2} - e^{-x/\theta} \frac{3x^2}{2\theta^4} = e^{-x/\theta} \frac{x^2}{2\theta^4} \left( \frac{x}{\theta} - 3 \right) \end{aligned}$$

**(b) Verify the conditions of Theorem 8.3. What family of functions  $v(\cdot)$  can be used? Is  $h$  in the family?**

Given  $h(x) = \max(0, x - s)$  and  $f_{\theta}(x) = (x^2/2\theta^3)e^{-x/\theta}$

(i) It is obvious that  $f_{\theta}(x)$  is differentiable with respect to  $\theta$  for all  $x \in S$

(ii)